

ORIGINAL RESEARCH

Influence of the large-Z effect during contact between butterfly sister species

Erik D. Nelson  | Qian Cong | Nick V. Grishin

Department of Biophysics, University of Texas Southwestern Medical Center, Dallas, TX, USA

Correspondence

Erik D. Nelson, Department of Biophysics, University of Texas Southwestern Medical Center, 6001 Forest Park Rd., Room ND10.124, Dallas, TX 75235–9050, USA. Email: nelsonerikd@gmail.com

Funding information

National Institutes of Health, Grant/Award Number: GM127390

Abstract

Recently diverged butterfly populations in North America have been found to exhibit high levels of divergence on the Z chromosome relative to autosomes, as measured by fixation index, F_{st} . The pattern of divergence appears to result from accumulation of incompatible alleles, obstructing introgression on the Z chromosome in hybrids (i.e., the large-Z effect); however, it is unknown whether this mechanism is sufficient to explain the data. Here, we simulate the effects of hybrid incompatibility on interbreeding butterfly populations using a model in which populations accumulate cross-incompatible alleles in allopatry prior to contact. We compute statistics for introgression and population divergence during contact between model populations and compare our results to those for 15 pairs of butterfly species interbreeding along a suture zone in central Texas. Time scales for allopatry and contact in the model are scaled to glacial and interglacial periods during which real populations evolved in isolation and contact. We find that the data for butterflies are explained well by an otherwise neutral model under slow fusion conditions. In particular, levels of divergence on the Z chromosome increase when interacting clusters of genes are closely linked, consistent with clusters of functionally related genes in butterfly genomes.

KEYWORDS

gene cluster, hybrid incompatibility, introgression, sex chromosome, speciation

1 | INTRODUCTION

Recent studies comparing divergent sister populations of butterflies have revealed elevated levels of divergence on the Z chromosome relative to autosomes (Cong et al., 2019; Kronforst et al., 2013). To explain this result, it has been suggested that the observed patterns of divergence are caused by the accumulation of postzygotic incompatibilities, obstructing introgression on the Z chromosome in hybrids (see, e.g., Figure 5 of Cong et al., 2019). However, a number of other factors can contribute to this effect, including higher rates of adaptation on the Z chromosome, changing population sizes, and

differing rates of reproductive success for male and female butterflies (Van Belleghem et al., 2017). As a result, it is of interest to know the "bare" contribution of hybrid incompatibilities to the extent of divergence in autosomes and Z chromosomes—for example, in an otherwise "neutral" model where these factors are absent. In this work, we develop such a model and compare our results to those obtained by Cong et al. (Cong et al., 2019) for 15 closely related species of butterfly interbreeding along a suture zone in central Texas.

The Texas suture zone is formed by emigration of butterfly species from glacial refugia along coastal and inland regions of Mexico and the southern United States, extending from the Yucatan

This is an open access article under the terms of the Creative Commons Attribution License, which permits use, distribution and reproduction in any medium, provided the original work is properly cited.

© 2021 The Authors. *Ecology and Evolution* published by John Wiley & Sons Ltd.

peninsula to the tip of Florida (see Figures S1 and S2). The species sampled by Cong et al. diverged on the order of 1 million years ago (Zhang et al., 2019) and have, as a result, experienced multiple periods of glacial cooling and interglacial warming. During glacial periods, central Texas was subjected to severe decreases in temperature (Annan & Hargreaves, 2013), which would have caused drastic, if not total isolation of sister species in southeastern and southwestern refugia; during the most recent warming period, sister species migrated into Texas, while major portions of their populations remained in refugial regions, isolated from the suture zone by large distances. To determine the influence of hybrid incompatibility with the Z chromosome during contact, we will at first neglect the effects of isolation by distance and consider a generic model of secondary contact (Geneva et al., 2015; Harris & Nielsen, 2016) in which a population divides, and the resulting sister populations evolve for a period in allopatry while accumulating hybrid incompatibilities and later begin to interbreed. We then compare statistics for introgression and population divergence for gene sequences in our model to those obtained for real populations by Cong et al. (2019).

To represent the state space for pairs of sister populations, Cong et al. employed two basic statistics: (a) the index of gene flow, I_{gf} , an extension of the indicator function G_{min} developed by Geneva et al. (Geneva et al., 2015), defined as the fraction of independent sequence windows along a genome with $G_{min} \leq G_0$, where G_0 is a threshold for introgression, and (b) the fixation index, or relative divergence function, F_{st} (Bhatia et al., 2013); I_{gf} measures the fraction of sequence windows where introgression has occurred, while F_{st} measures the degree of genetic difference between populations (details are provided in the Methods section). Multiple genomic samples were collected from each sister population, and separate index values were computed for autosomes and Z chromosomes. The results are shown in Figure 1; data points in this figure describe index values for sister organisms that have been classified as different species in the literature (green), more closely related organisms for which classification is uncertain (yellow), and samples of the same species (red). When populations are compared through their autosomes (Figure 1a), the data exhibit a continuous pattern across the entire range of index values; however, for the Z chromosome (Figure 1b), the data obtained from samples of the same species (red) are separated from those of closely related species by a gap of "missing" values, suggestive of a sudden transition (Kronforst et al., 2013; Nosil et al., 2017). For different species (green and yellow data points), relative divergence (F_{st}) values for the Z chromosome are always larger than those for autosomes (Figure 2). However, the fraction of divergent nucleotide positions in the Z chromosomes of sister species is similar to those for autosomes (see Figure 5a of reference Cong et al., 2019), indicating similar absolute rates of divergence, inconsistent with a "faster-Z" effect (Avila et al., 2014; Meisel & Connollon, 2013). In accord with these results, Cong et al. have argued that the pattern of data in Figure 1 reflects the influence of negative fitness interactions between autosomes and Z chromosomes in hybrids during periods of interbreeding—in other words, a "large-Z" effect (Muirhead & Presgraves, 2016; Van Belleghem et al., 2017).

Our goal in this work is to determine whether this mechanism can explain the data for butterflies—specifically, the gap in Figure 1b, and the large differences, $\Delta F = F_Z - F_A$, between F_{st} values for autosomes and Z chromosomes shown in Figure 2. To accomplish this, we simulate populations of model butterflies under conditions that scale to those experienced by real populations during glacial isolation and contact. Chromosomes in our model consist of adjacent gene segments of identical length (see Appendix A, Figure A1). Mutation rates and rates of crossing over within gene segments are scaled to estimates for *Drosophila*, and *Heliconius* butterflies, and rates of crossing over between gene segments are varied to reflect the typical separation, or degree of linkage between genes on butterfly chromosomes. Mutations are individually neutral. Hybrid incompatibility in the model occurs as a result of negative fitness interactions between mutant alleles that rise to fixation in different populations during allopatry, similar to the model described by Orr (Orr, 1995). Interactions are pairwise and connect loci in autosomes to loci in the Z chromosome(s). The fitness cost for a pair of interacting loci resembles the "pathway" model described by Lindke and Buerkle (Lindtke & Buerkle, 2015). Depending on the strength of the interactions and the migration rate, the model leads either to fusion or continued divergence during contact. Here, we focus primarily on fusion conditions.

For a given set of conditions (i.e., interaction strength, migration rate, time in allopatry, etc.), we conduct multiple simulations in parallel to generate statistical profiles for index values and other quantities of interest. We first show that a purely neutral model, in which hybrid interactions are turned off during contact, is unlikely to explain the data for ΔF (we refer to this situation as the "null model" below). During allopatry, F_Z is slightly larger than F_A , as expected, due the smaller population size, and hence higher substitution rate for Z chromosomes (Van Belleghem et al., 2017) (populations maintain an approximately 1:1 sex ratio in the model); results for the null model during contact are similar to those for populations evolving in allopatry. However, as hybrid interactions are increased, profiles for ΔF begin to resemble the pattern of data in Figure 2, particularly when the rate of crossing over between genes is small, reflecting closely linked genes on butterfly chromosomes. In this case, which would perhaps correspond to interacting clusters of functionally related genes (Cong et al., 2016; McDonald & Rosbash, 2001), we find that the model can explain the large values of ΔF shown in Figure 2 under realistic conditions. The index of gene flow, I_A , agrees (on average) with the data in Figure 1a during both allopatry and contact, and the model also leads to a "statistical" gap in I_Z analogous to that Figure 1b during contact. However, data for I_Z typically exceed the values in Figure 1b for intermediate values of F_Z , suggesting a missing feature in the model. We return to this point in the Discussion section later below.

2 | METHODS

We simulate model populations in three phases: (a) equilibration of an initial, ancestral population, (b) division of the population and

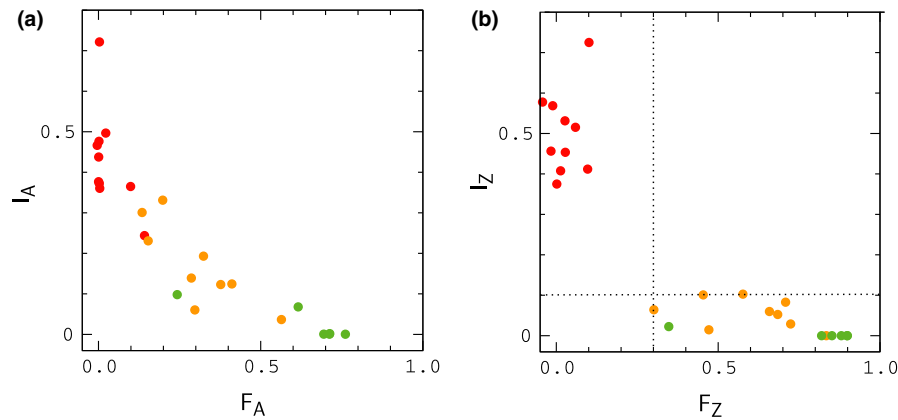


FIGURE 1 Index of gene flow versus index of fixation for autosomes (a) and Z chromosomes (b) of sister species sampled by Cong et al. Data for I_A and I_Z are multiplied by a factor of 4 to remove a scaling factor used in their work. Data points describe pairs of organisms that have been classified as different species in the literature (green), closely related organisms for which classification is uncertain (yellow), and organisms of the same species (red). The dotted lines in panel (B) are included simply to guide the eye. Index values were computed from sequence windows of about 1 kb in length

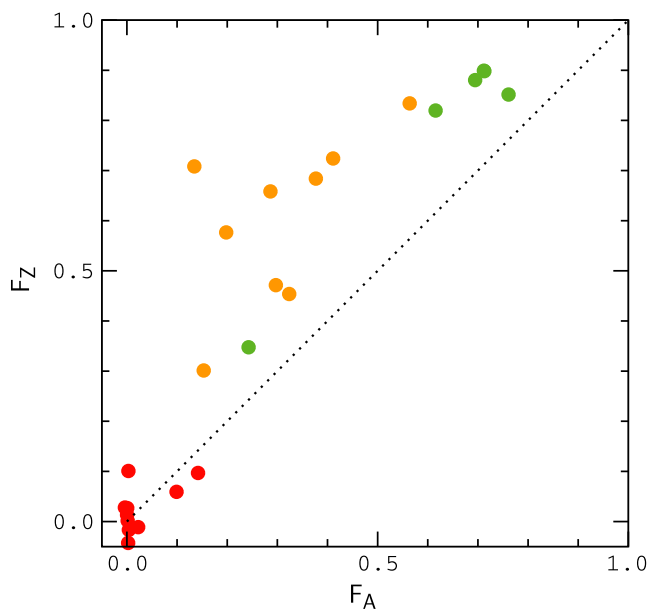


FIGURE 2 Correlation between F_Z and F_A values from Figure 1. The dotted line $F_Z = F_A$ is included to guide the eye

evolution of sister populations in allopatry, and (c) contact between sister populations, subject to hybrid interactions between mutant alleles acquired in allopatry. In each phase, populations evolve by plain Wright-Fisher dynamics with random mating between male and female individuals (Gillespie, 2004). In each generation, mutations occur within genes at a rate μ per gene per generation. Pairs of male and female individuals are then selected at random according to fitness for mating. Male genomes undergo explicit meiosis, in which chromosomes are duplicated, and the resultant chromatids undergo random crossing over (Veller et al., 2018) with separate rates, r and r' , for crossing over within and between gene segments (meiosis is achiasmatic in model females, consistent with butterfly reproduction; Edelman et al., 2019). A single offspring is generated

from each mating event by random union of male and female gametes, and the procedure is repeated until the original population is replenished. Accordingly, populations maintain a roughly 1:1 ratio of male and female individuals. During contact, an equal number of offspring (with mean $N\epsilon$, where N is the size of each population and ϵ is the migration rate) are selected at random from each population to undergo migration, and the selected offspring are exchanged between populations.

We consider two possible evolutionary scenarios; the scenarios are the same, except for the initial population size and the transition into allopatry: In scenario (i), the initial population is cloned, while in scenario (ii), the initial population divides into equal parts. We first describe scenario (i) and then describe the differences between scenarios (i) and (ii).

In scenario (i), an initial population of size N is equilibrated for Δt_E generations. Let $g_i = (g_i, g'_i)$ denote the allelic state of a diploid locus l in a genome g . All genomes in the initial population begin with $g_i = 0$ uniformly. Mutation events during a simulation act to assign the maternal (g_i) or paternal (g'_i) state of a locus to 1. All mutations are individually neutral. After equilibration, loci l that have fixed in the population for the mutant allele type are returned to their initial states, $g_i = 0$. The population is then duplicated, and the resultant sister populations (each of size N) evolve in allopatry for a period Δt_A . At the end of this period, loci that have fixed for the mutant allele across both populations are returned to their initial states. Several pairs of loci are then selected to participate in hybrid fitness interactions (see below), and the two populations evolve in contact for a period Δt_C subject to fitness costs incurred due to interactions formed by various allele combinations at the selected loci.

In scenario (ii), the entire procedure is the same, except that the initial population has size $2N$ before dividing into sister populations of size N (in this case, the equilibration period is twice as long). In both scenarios, the W chromosome acts only to determine the sex of an individual.

During allopatry, mutant alleles are lost, or rise toward fixation in each population via genetic drift. Loci that are nearly fixed for the mutant allele type in one population are usually far from fixation in the other. Let $p_{i,\gamma}$ denote the frequency of the mutant allele type at locus i in population γ , with $\gamma = 1, 2$. To describe the cost of hybridization, we select a small number of loci in autosomes for which $p_{i,1} \sim 1$ and $p_{i,2} \sim 0$ to interact negatively with loci in the Z chromosome(s) for which $p_{j,1} \sim 0$ and $p_{j,2} \sim 1$ (see Appendix A). We then repeat this process with the population subscripts interchanged, selecting an equal number of loci in autosomes with $p_{i,2} \sim 1$ and $p_{i,1} \sim 0$ to interact negatively with loci in the Z chromosome(s) for which $p_{j,2} \sim 0$ and $p_{j,1} \sim 1$ (typically, $p_{i,\gamma} = 0, 1$ for selected loci in our simulations, but in general, this will depend on the mutation rate, the amount of time spent in allopatry, and the number of selected loci).

Let $f(\mathbf{g}_i, \mathbf{g}_j)$ denote the log fitness cost for a pair of selected loci, (i, j) . We define f as follows: If both loci are homozygous for the mutant allele, $f = 4s$; if one locus is homozygous for the mutant allele and one locus is heterozygous, $f = 2s$; and if both loci are heterozygous, $f = s/4$. For all other combinations, $f = 0$. The fitness of a genome is then defined as

$$w = \exp - \sum_{(i,j)} f(\mathbf{g}_i, \mathbf{g}_j), \quad (1)$$

where the sum extends over pairs of selected loci. Here, we assume that mutant loci on the single Z chromosome of a female genome are dominant and act as homozygous loci on the two Z chromosomes of a male genome. This condition, and the fact that f is smaller than $2s$ when both loci are heterozygous, ensures that hybrid females are typically less fit than hybrid males, consistent with Haldane's rule and the analysis of Cong et al. (2019). The model for f is the same as the "pathway" model used by Lindtke and Buerkle (2015) to describe hybrid interactions among autosomal loci, except for the factor $f = s/4$ when both loci are heterozygous (in this case, $f = 0$ in the pathway model).

Genomes in our simulations contain three sets of chromosomes; autosomes carry three genes, and Z chromosomes carry six genes, each of length L loci, as shown in Appendix A, Figure A1. Since interacting loci are required to have $p_{i,\gamma} \sim 1$ in one of the diverging populations, we are somewhat limited in regard to the number of interacting loci we can select to define w in a given simulation. Here, we select six pairs of loci in each simulation for which the differences between $p_{i,1}$ and $p_{i,2}$ above are largest. Unless otherwise noted, we select loci that connect the first pair of autosomes to the Z chromosome(s).

The dynamical parameters of the simulations are chosen so that their scaled values ($N\mu$, Nr , and Nr') agree in order of magnitude with values obtained for *Heliconius* and *Drosophila*, an organism used to infer the biochemical functions of butterfly genes (Cong et al., 2019). Estimates for the point mutation rate in *Drosophila* are in the range of about $10^{-10} - 10^{-9}$ per generation (Halligan & Keightley, 2009; Keightley & Eyre-Walker, 2000; Keightley et al., 2014); here, we have assumed a genome size for *Drosophila* of 180 Mbp to compute point

mutation rates from ref. Halligan and Keightley (2009). Assuming an effective population size for *Drosophila* of 10^6 individuals (Keightley, Ness, et al., 2014; Li et al., 1999), and a typical gene length of about 1770 bp (Keightley & Eyre-Walker, 2000), we obtain a scaled mutation rate of $N\mu \approx 0.18 - 1.8$ per gene per generation. Here, we adopt a value $2N\mu = 1$ in the lower range of these estimates (during publication, we became aware of a significantly larger estimate for *Heliconius* (Keightley, Pinharanda, et al., 2014), and we discuss the possible effect of this alternative later below). The typical length of a chromosome in *Heliconius* is about 20 Mbp, and the crossover rate per chromosome is about $r \sim 1$ per generation (Edelman et al., 2019); if we assume that the typical length of a gene in *Heliconius* is the same as in *Drosophila*, we obtain a crossover rate per gene of about $r \sim 10^{-4}$ per generation and a scaled rate of about $Nr \sim 100$ (Van Belleghem et al., 2017).

Chromosomes are described as strings of characters in our C++ code. During reproduction, chromosome strings are copied and recombined many thousands of times, making it costly to simulate butterfly genes explicitly. The fraction of mutant alleles participating in model chromosomes is typically on the order of a several percent for the time scales considered here (see ref. Cong et al., 2019, for comparison). Consequently, the probability that a mutation attempt is repeated at the same locus during allopatry or contact is small. Thus, in order to reduce the computational cost of our simulations, we use "compressed" genes of length $L = 100$ loci to represent genes on butterfly chromosomes. We simulate populations of $N = 10^4$ individuals for various values of the parameters s , ϵ , and r' . The morphologies (sex organs, wing color patterns, etc.) of butterfly sister specimens are similar, suggesting that prezygotic barriers to introgression may be small. Accordingly, we explore a broad range of migration rates, $0.1 \leq N\epsilon \leq 10$. Interaction strengths are varied in the range, $0.01 \leq s \leq 0.1$, including the null model, $s = 0$. Our main findings are summarized below. Links to the data and C++ code used to conduct the simulations are provided in the Data Accessibility section.

3 | RESULTS

To begin our investigation, we explore the time dependence of the statistics G_{\min} and F_{st} for populations evolving in allopatry for comparison with the results of Geneva et al. (Geneva et al., 2015).

To define these objects, let $d_i^{\mu\nu} = |g_i^\mu - g_i^\nu|$ denote the difference (Hamming distance) between genomes g^μ and g^ν at (haploid) locus i , and let

$$d_{\lambda, \lambda + \Delta}^{\mu\nu} = \sum_{l=\lambda}^{\lambda + \Delta} d_l^{\mu\nu} \quad (2)$$

denote the distance between g^μ and g^ν for a window of loci, $[\lambda, \lambda + \Delta]$. Assume that we have sampled a small number of genomes

FIGURE 3 Mean values of G_{min} and F_{st} for autosomal genes as a function time since diverging in allopatry under scenarios (i) (green) and (ii) (maroon). Averages are computed from 128 replicate simulations with $N = 10^4$, $2\mu = 10^{-4}$, and $r, r' = 10^{-2}$. The plots are precise polynomial fits to the averages. Plots of F_{st} computed from four and ten samples per population are essentially identical

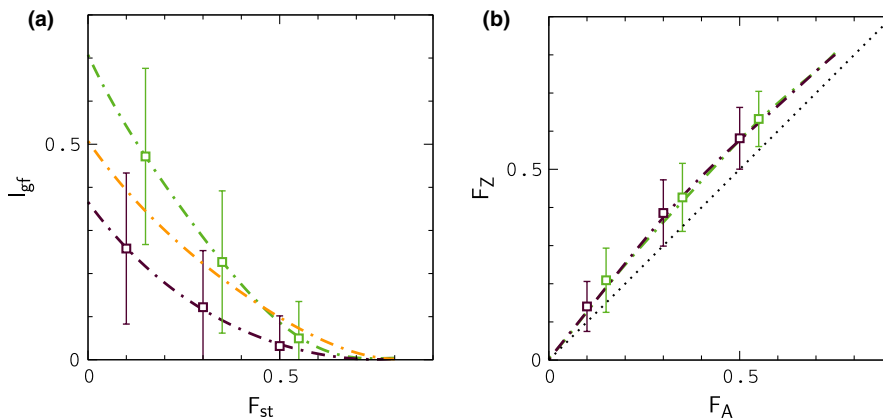
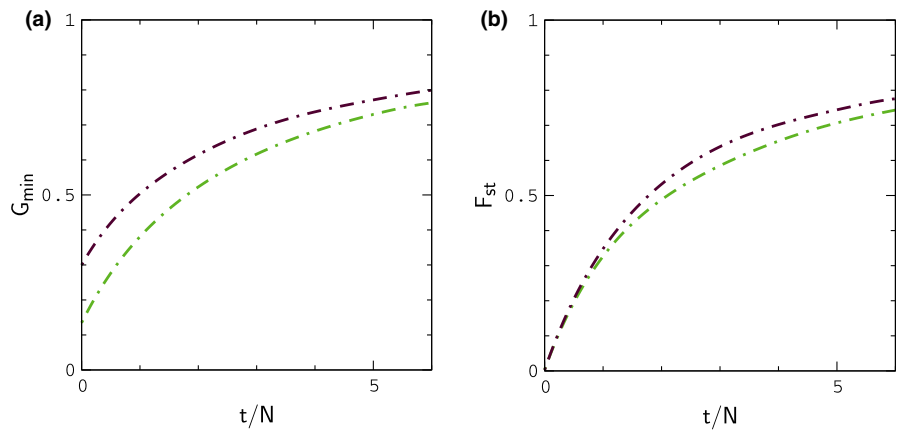


FIGURE 4 Index distributions for the simulations in Figure 3. Data points denote binned averages and error bars indicate the widths of the distributions of values for selected bins. Broken lines are precise polynomial fits to $\langle I_A \rangle (t)$ versus $\langle F_A \rangle (t)$, and $\langle F_Z \rangle (t)$ versus $\langle F_A \rangle (t)$ as a function of time, where braces denote averaging over simulations (a plot of $\langle I_Z \rangle (t)$ versus $\langle F_Z \rangle (t)$ for scenario (ii) (yellow) is included for comparison). The large widths for I_A reflect the small number of gene sequences considered in the model

from each population. For a given window of loci, G_{min} is then defined as the ratio (Geneva et al., 2015),

$$G_{min} = \min d_{\lambda, \lambda + \Delta}^{\mu\nu} / \overline{d_{\lambda, \lambda + \Delta}^{\mu\nu}}^{1.2} \tag{3}$$

where $\min d_{\lambda, \lambda + \Delta}^{\mu\nu}$ and $\overline{d_{\lambda, \lambda + \Delta}^{\mu\nu}}$ are the minimum and average distances between sequences sampled from different populations; the fixation index, or relative divergence is defined as (Geneva et al., 2015),

$$F_{st} = 1 - \frac{\overline{d_{\lambda, \lambda + \Delta}^{\mu\nu}}^{-1} + \overline{d_{\lambda, \lambda + \Delta}^{\mu\nu}}^{-2}}{2 \overline{d_{\lambda, \lambda + \Delta}^{\mu\nu}}^{1.2}} \tag{4}$$

where, for example, $\overline{d_{\lambda, \lambda + \Delta}^{\mu\nu}}^{-1}$ is the average distance between sequences sampled from population 1 with $\mu \neq \nu$. Below, we compute G_{min} for individual gene sequences (see Appendix A), and we compute F_{st} by averaging the numerator and denominator of the fraction in Equation (3) over gene sequences, as recommended by Bhatia et al. (2013). For a given window, the probability of obtaining a value of G_{min} that is less than a given value of G_0 increases with the number of samples used to compute G_{min} . For this reason, we compute G_{min} by sampling four genomes from each population, as in the method used for butterfly

genomes, and we compute F_{st} by sampling ten genomes from each population (this last step is intended as a means to reduce noise in plots like Figure 1). The introgression measure, I_{gf} , is defined as the fraction of gene segments with $G_{min} \leq 0.25$ [see Figure 4a of Geneva et al. (2015)]. Except for the number of samples used to compute F_{st} , our approach is the same as that used by Cong et al. (2019) (to be more precise, when larger numbers of specimens were available for a pair of species, Cong et al. determined the index values by repeatedly sampling 4 pairs specimens from each species at random, and averaging the results).

In Figure 3, we plot F_{st} and G_{min} for autosomal genes as a function of time diverging in allopatry under scenarios (i) and (ii). The results can be compared with Figure 2 of Geneva et al. (2015). Although a direct comparison is not possible (Geneva et al. average simulations of a single sequence window over a range of μ and r values), our results behave as expected for the lower values of μ and r used in our simulations (the transition to allopatry in Geneva et al. is analogous to scenario (i)). Interestingly, there is a noticeable difference in the plots of G_{min} for duplication and division of populations in allopatry, and scenario (ii) leads to closer agreement with butterfly data for I_{gf} . Results for I_{gf} and F_{st} corresponding to the simulations in Figure 3 are shown in Figure 4; to compare our results to the butterfly data, we

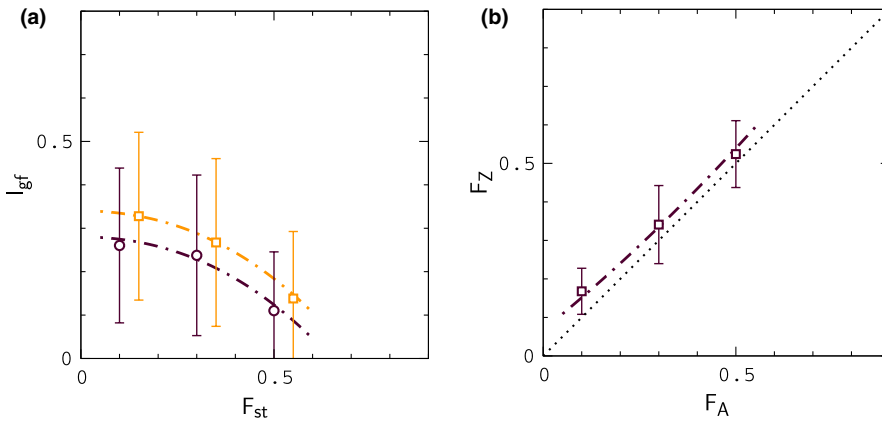


FIGURE 5 Index distributions for a purely neutral (null) model during secondary contact. Averages for l_Z versus F_A (l_Z versus F_Z) in panel (a) are indicated by circles (squares). In both panels, error bars indicate the widths of the distributions for selected bins. The simulation parameters are the same as in Figures 3 and 4, with $N_e = 1.5$, $\Delta t_A = 2N$, and $\Delta t_C = N$. Dashed lines in the figure are cubic fits to the averages and are simply intended to guide the eye

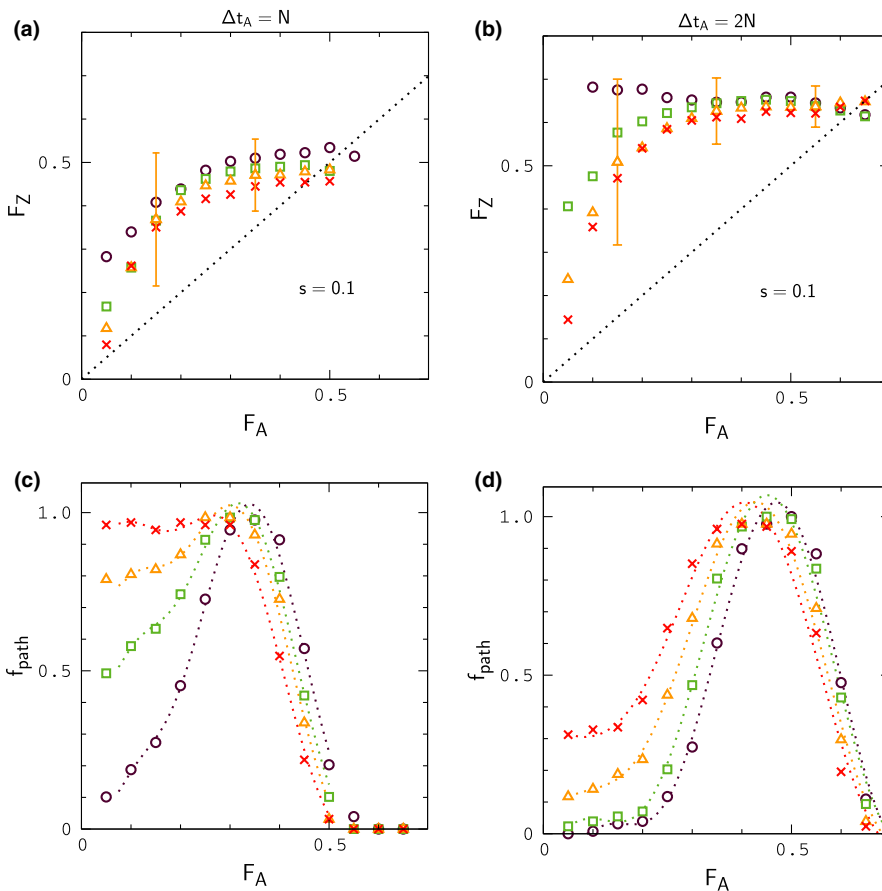


FIGURE 6 Study of F_Z versus F_A during contact for different periods in allopatry. Data points in panels (a) and (b) denote binned averages of F_Z for migration rates $N_e = 1.5$ (circles), 2.5 (squares), 4 (triangles), and 6 (crosses). Each set of points is the result of 128 replicate simulations with $N = 10^4$, $2\mu = 10^{-4}$, $r, r' = 10^{-2}$, sampled for $\Delta t_C = N$ generations. For clarity, the distribution widths for F_Z are indicated only for $N_e = 4$. Data points in panels (c) and (d) describe the fraction of simulation paths that have reached a given bin for F_A at least once during contact

plot binned averages of the index values sampled at regular points during the simulations in Figure 3. Note that the results for F_{st} are unlikely to explain the large values of ΔF in Figure 2 under either scenario.

In the remaining figures, we describe results for contact between populations under scenario (ii). To schedule the simulations, we assume that periods of contact are comparable to those of real populations during interglacial warming periods. The time scale for glacial or interglacial periods in North America over the last million years is roughly between 10^4 and 10^5 years. To calibrate the model to real-time scales, we assume, consistent with our choice of parameters, that N generations in the model correspond to N_e generations for

butterfly populations, where N_e is the effective population size for butterflies. Then, solving for α in the expression $\alpha N_e \tau = \Delta \tau$, where τ is the generation time and $\Delta \tau$ is the length of an glacial period, the corresponding period of contact in the model is αN . Although data for N_e is unavailable for the species in Figure 1, we can obtain a rough idea of how N_e varies over time and among species from the study of *Heliconius* populations by Van Belleghem et al. (2017) (for *Drosophila*, see Sprengelmeyer et al., 2019). Below, we focus our attention on values of $N_e \tau$ on the order of 10^5 years, consistent with the lower range of N_e values in Van Belleghem et al. (2017), in which case, N generations in the model corresponds in order of magnitude to the length of a glacial or interglacial period for butterflies.

FIGURE 7 Study of F_Z versus F_A during contact for decreasing interaction strengths. The simulation parameters are the same as those listed in Figure 6 except where indicated in panels (a) and (b)

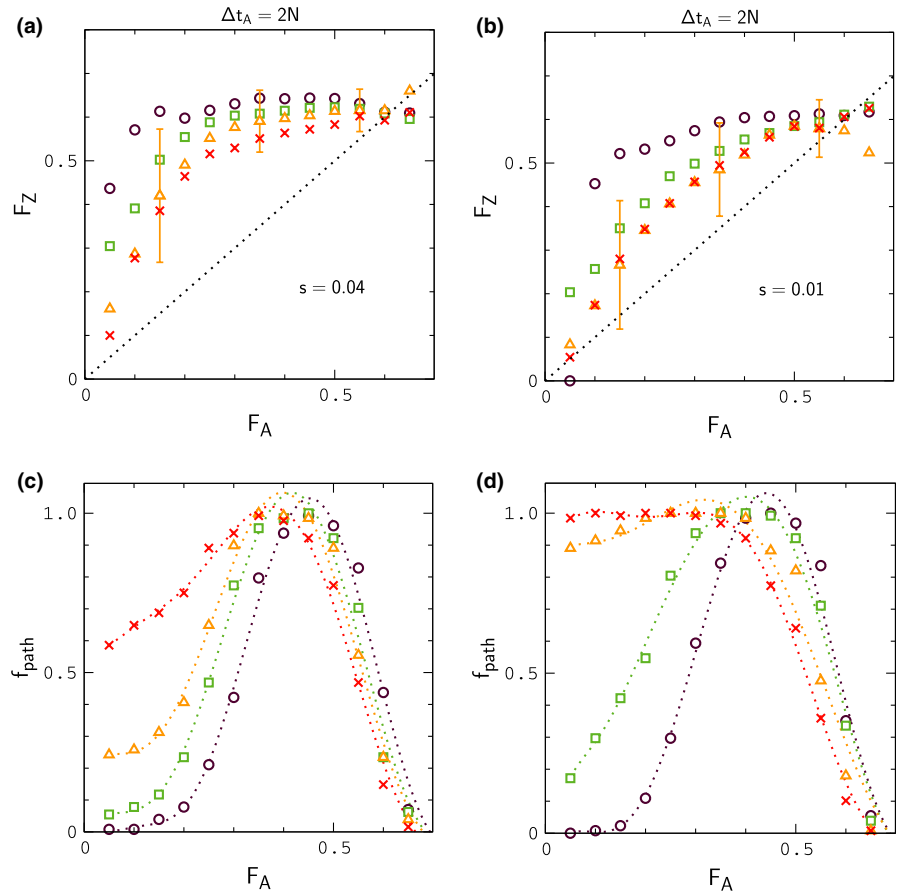
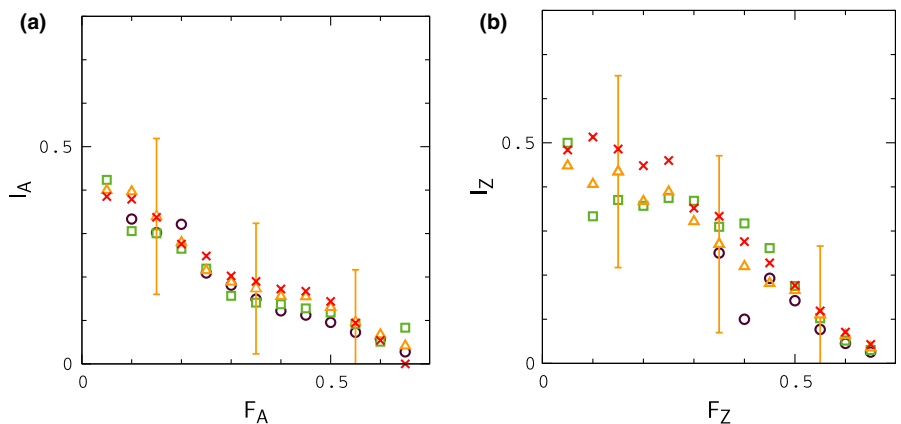


FIGURE 8 Averages for I_A versus F_A and I_Z versus F_Z obtained from the simulations in Figure 6b. Error bars indicate the widths of distributions for $N_E = 4$. The widths reflect the relatively small numbers of gene sequences used to determine I_A and I_Z in the model, as noted above



The results are summarized in Figures 5–7. In these simulations, we focus on the case of closely linked genes, $r' = r$, as in Figure 4; higher rates of crossing over between genes ($r' > r$) are explored later below. Data sets (circles, squares, etc.) in each figure are obtained from 128 replicate simulations sampled every 100 generations. As in Figure 4, we compute binned averages of the index values sampled during the period of interest in order to compare our results to Figures 1 and 2. In this case, each data point represents an average over snapshots of the populations as they evolve during contact, so that the values collected in each bin are sampled at different times during the simulations. An alternative might be to plot averages for the simulations at specific points in time, as we did, for

example, with $\langle F_Z \rangle (t)$ versus $\langle F_A \rangle (t)$ in Figure 4. However, during contact, it is often the case that some simulations proceed steadily toward fusion while others initially diverge. As a result, averages such as $\langle F_A \rangle (t)$ are very noisy, and plots of $\langle F_Z \rangle (t)$ versus $\langle F_A \rangle (t)$ do not accurately reflect the result of a random sample of populations during contact. Thus, short of constructing movies of the index distributions for each set of conditions, the present approach seems sufficient to express the results.

In Figure 5, we plot I_{gf} and F_{st} for a purely neutral, or null model in which interactions are turned off ($s = 0$) during contact. In this example, populations evolve in allopatry for $\Delta t_A = 2N$ generations and remain in contact for $\Delta t_C = N$ generations. The mean value of F_A just

prior to contact is $F_A \approx 0.5$. The results are similar to those obtained during allopatry in Figure 4, and, accordingly, the null model is unlikely to explain the data in Figure 2; note that I_Z is shifted upward from I_A by an amount similar to that in Figure 4a.

In Figures 6 and 7, we explore the effect of hybrid interactions on ΔF . For these simulations, populations evolve in allopatry for a period of either $\Delta t_A = N$ or $\Delta t_A = 2N$ generations and remain in contact for $\Delta t_C = N$ generations; the mean values of F_A just prior to contact are $F_A \approx 0.3$ and $F_A \approx 0.5$, respectively. Data points in upper panels of the figures denote binned averages for a particular choice of s and ϵ ; lower panels describe the fraction of simulations contributing to each data point—or, more precisely, the fraction of simulations for a given s and ϵ that have visited a given bin for F_A at least once. The numbers of samples contributing to each data point are shown in Figures S3 and S4.

As is evident by inspection of Figures 6b and 7a, large values of ΔF , consistent with the largest values in Figure 2, can occur at low to moderate frequency when hybrid interactions are sufficiently strong. For weaker interactions (Figure 7b), when small values of F_A are more frequent, data for F_Z usually remain above the point $F_Z \approx 0.3$, the smallest value of F_Z for different species in Figure 1, when $F_A \approx 0.15$, the smallest value of F_A for different species. In this case, samples drawn at random from simulations with $F_A \gtrsim 0.15$ are unlikely to occur in the gap region of missing F_Z values in Figure 1. For all of the conditions considered in Figures 6 and 7, the results for I_{gf} are similar to those obtained for the null model in Figure 5; results for I_{gf} corresponding to the simulations in Figure 6b are shown in Figure 8; in reading this figure, note that if $F_A \gtrsim 0.15$, then typically $F_Z > 0.3$ according to the results in Figures 6 and 7, in which case I_Z significantly smaller than its "same species" value, analogous to Figure 1b. In addition, for many of the simulation sets in Figures 6 and 7, samples of F_A less than 0.15 are infrequent. In these situations, limited random samples of the simulations are likely to result in a pattern of data for I_A and I_Z resembling the patterns for different species in Figure 1. For lower rates of migration, $N\epsilon \lesssim 1$, leading to slow fusion or continued divergence during contact, I_A and I_Z begin to resemble Figure 1 explicitly, concurrent with large values of ΔF , as shown in Appendix B. Finally, in Appendix C, we show that larger rates of crossing over between genes, $r' > r$ lead to smaller values of ΔF under fusion conditions.

4 | DISCUSSION

The model seems to capture the data in Figures 1 and 2 rather well, with the exception of results for I_Z , which are larger than those in Figure 1b for intermediate values of F_Z . Much of this discrepancy appears to result from the smaller population size for Z chromosomes relative to autosomes in the model. To see this, note that plots of I_Z and I_A for the null model (Figure 5) are similar to those when hybrid interactions are included (Figure 8). In allopatry, and in the null model, I_Z is shifted upward from I_A by a similar amount. However, in these situations, the only distinguishing factor between

the dynamics of genes on autosomes and Z chromosomes is population size. Thus, for example, we would expect I_Z to approach I_A (i.e., which is similar to I_Z in Figure 1) if males were more abundant than females in the model.

Another issue is the estimate used for the mutation rate, $N\mu$. The recent estimate for *Heliconius* noted above (Keightley, Pinharanda, et al., 2014) is several times larger than the estimate used in our simulations. Larger mutation rates would lead to more rapid divergence of populations in allopatry and different conditions for fusion and continued divergence following contact. However, it is worth recalling that the value of $N_e\tau$ used to estimate the length of glacial periods for the model is in the lower range of values for *Heliconius*. (Van Belleghem et al., 2017). Larger and perhaps more realistic values of N_e would lead to shorter glacial periods (i.e., smaller values of α in the relation $\alpha N_e\tau = \Delta\tau$ above), which would act to compensate for an increased mutation rate in allopatry. In addition, populations would have less time to interbreed during contact, leading to plots that more closely resemble those for slow fusion in Appendix B.

Finally, it is important to remark that index values computed for a pair of species will depend on where the specimens are collected. The locations of specimens studied in this work often extend over thousands of kilometers on either side of the suture zone (see, e.g., Figure S1). In these distant regions of the landscape, sister populations evolve in greater isolation and, hence, diverge at a higher rate. As a result, the index values obtained by Cong et al. reflect an average over individuals diverging at different rates. In addition, large regions of the landscape on either side of the suture zone are fragmented, consisting of loosely distributed patches of resources on which butterfly numbers can vary dramatically (see, e.g., McIntire et al., 2013; O'Hara, 2005; Schultz & Crone, 2001). The environments on either side of the suture zone are different, which has probably led to some level of divergent adaptation (e.g., discordant mating cycles (Cong et al., 2016), mate preferences (Kronforst et al., 2013), and environment preferences), limiting the rates of interbreeding and gene flow in some complex way (Edelaar et al., 2008; Flaxman et al., 2014; M'Gonigle et al., 2012). Given the present scale of computing power, and the increasing ease of obtaining genetic information, it would be worthwhile to develop software capable of modeling the properties above for realistic population sizes and genome structures (Haller & Messer, 2019). The present work is the first step toward this goal.

ACKNOWLEDGMENTS

It is a pleasure to thank Jing Zhang and two anonymous reviewers for helpful comments during the completion of this work. This study is supported in part by a grant (to NVG) from the National Institutes of Health (GM127390).

CONFLICT OF INTEREST

None declared.

AUTHOR CONTRIBUTION

Erik D. Nelson: Conceptualization (lead); Data curation (lead); Formal analysis (lead); Investigation (lead); Methodology (lead); Software

(lead); Writing-original draft (lead); Writing-review & editing (lead). **Qian Cong:** Conceptualization (supporting); Data curation (supporting); Investigation (supporting); Visualization (supporting); Writing-original draft (supporting). **Nick V. Grishin:** Funding acquisition (lead); Resources (lead); Supervision (supporting); Writing-original draft (supporting).

DATA AVAILABILITY STATEMENT

C++ code used to generate the data: <https://cloud.biohpc.swmed.edu/index.php/s/kSQPenOQPDXTx7Q>

Data for the paper: <https://cloud.biohpc.swmed.edu/index.php/s/WpLWiKo9rzTF88X>

ORCID

Erik D. Nelson  <https://orcid.org/0000-0001-8345-227X>

REFERENCES

- Annan, J. D., & Hargreaves, J. C. (2013). A new global reconstruction of temperature changes at the last glacial maximum. *Climate of the Past*, 9, 367–376. <https://doi.org/10.5194/cp-9-367-2013>
- Avila, V., de Procé, S. M., Campos, J. L., Borthwick, H., Charlesworth, B., & Betancourt, A. J. (2014). Faster-x effects in two *Drosophila* lineages. *Genome Biology and Evolution*, 6, 2968–2982. <https://doi.org/10.1093/gbe/evu229>
- Bhatia, G., Patterson, N., Sankararaman, S., & Price, A. L. (2013). Estimating and interpreting FST: The impact of rare variants. *Genome Research*, 23, 1514–1521. <https://doi.org/10.1101/gr.154831.113>
- Cong, Q., Shen, J., Borek, D., Robbins, R. K., Otwinowski, Z., & Grishin, N. V. (2016). Complete genomes of hairstreak butterflies, their speciation, and nucleo-mitochondrial incongruence. *Scientific Reports*, 6, 24863. <https://doi.org/10.1038/srep24863>
- Cong, Q., Zhang, J., & Grishin, N. (2019). Genomic determinants of speciation in butterflies. <https://www.biorxiv.org/content/10.1101/837666v1>
- Edelaar, P., Siepielski, A. M., & Clobert, J. (2008). Matching habitat choice causes directed gene flow: A neglected dimension in evolution and ecology. *Evolution*, 62, 2462–2472. <https://doi.org/10.1111/j.1558-5646.2008.00459.x>
- Edelman, N. B., Frandsen, P. B., Miyagi, M., Clavijo, B., Davey, J., Dikow, R. B., García-Accinelli, G., Van Belleghem, S. M., Patterson, N., Neafsey, D. E., Challis, R., Kumar, S., Moreira, G. R. P., Salazar, C., Chouteau, M., Counterman, B. A., Papa, R., Blaxter, M., Reed, R. D., ... Mallet, J. (2019). Genomic architecture and introgression shape a butterfly radiation. *Science*, 366, 594–599. <https://doi.org/10.1126/science.aaw2090>
- Flaxman, S. M., Wacholder, A., Feder, J. L., & Nosil, P. (2014). Theoretical models of the influence of genomic architecture on the dynamics of speciation. *Molecular Ecology*, 23, 4074–4088. <https://doi.org/10.1111/mec.12750>
- Geneva, A. J., Muirhead, C. A., Kingan, S. B., & Garrigan, D. (2015). A new method to scan genomes for introgression in a secondary contact model. *PLoS One*, 10, e0118621. <https://doi.org/10.1371/journal.pone.0118621>
- Gillespie, J. H. (2004). *Population genetics, a concise guide*. Johns Hopkins University Press.
- Haller, B. C., & Messer, P. W. (2019). Slim 3: Forward genetic simulations beyond the wright-fisher model. *Molecular Biology and Evolution*, 36, 632–637. <https://doi.org/10.1093/molbev/msy228>
- Halligan, D. L., & Keightley, P. D. (2009). Spontaneous mutation accumulation studies in evolutionary genetics. *Annual Review of Ecology and Systematics*, 40, 151–172. <https://doi.org/10.1146/annurev.ecolsys.39.110707.173437>
- Harris, K., & Nielsen, R. (2016). The genetic cost of neanderthal introgression. *Genetics*, 203, 881–891. <https://doi.org/10.1534/genetics.116.186890>
- Keightley, P. D., & Eyre-Walker, A. (2000). Deleterious mutations and the evolution of sex. *Science*, 290, 331–333. <https://doi.org/10.1126/science.290.5490.331>
- Keightley, P. D., Ness, R. W., Halligan, D. L., & Hadrill, P. R. (2014). Estimation of the spontaneous mutation rate per nucleotide site in a *Drosophila melanogaster* full-sib family. *Genetics*, 196, 313–320. <https://doi.org/10.1534/genetics.113.158758>
- Keightley, P. D., Pinharanda, A., Ness, R. W., Simpson, F., Dasmahapatra, K. K., Mallet, J., Davey, J. W., & Jiggins, C. D. (2014). Estimation of the spontaneous mutation rate in *Heliconius melpomene*. *Molecular Biology and Evolution*, 32, 239–243. <https://doi.org/10.1093/molbev/msu302>
- Kronforst, M. R., Hansen, M. E. B., Crawford, N. G., Gallant, J. R., Zhang, W., Kulathinal, R. J., Kapan, D. D., & Mullen, S. P. (2013). Hybridization reveals the evolving genomic architecture of speciation. *Cell Reports*, 5, 666–677. <https://doi.org/10.1016/j.celrep.2013.09.042>
- Li, Y., Satta, Y., & Takahata, N. (1999). Paleo-demography of the *Drosophila melanogaster* subgroup: Application of the maximum likelihood method. *Genes & Genetic Systems*, 74, 117–127. <https://doi.org/10.1266/ggs.74.117>
- Lindtke, D., & Buerkle, C. A. (2015). The genetic architecture of hybrid incompatibilities and their effect on barriers to introgression in secondary contact. *Evolution*, 69, 1987–2004. <https://doi.org/10.1111/evo.12725>
- M'Gonigle, L. K., Mazucco, R., Otto, S. P., & Dieckmann, U. (2012). Sexual selection enables long-term coexistence despite ecological equivalence. *Nature*, 484, 506–509. <https://doi.org/10.1038/nature10971>
- McDonald, M. J., & Rosbash, M. (2001). Microarray analysis and organization of circadian gene expression in *Drosophila*. *Cell*, 107, 567–578. [https://doi.org/10.1016/S0092-8674\(01\)00545-1](https://doi.org/10.1016/S0092-8674(01)00545-1)
- McIntire, E. J., Rompre, G., & Severns, P. M. (2013). Biased correlated random walk and foray loop: Which movement hypothesis drives a butterfly metapopulation? *Oecologia*, 172, 293–305. <https://doi.org/10.1007/s00442-012-2475-9>
- Meisel, R. P., & Connollon, T. (2013). The faster-x effect: Integrating theory and data. *Trend in Genetics*, 29, 537–544. <https://doi.org/10.1016/j.tig.2013.05.009>
- Muirhead, C. A., & Presgraves, D. C. (2016). Hybrid incompatibilities, local adaptation, and the genomic distribution of natural introgression between species. *American Naturalist*, 187, 249–261. <https://doi.org/10.1086/684583>
- Nosil, P., Feder, J. L., Flaxman, S. M., & Gompert, Z. (2017). Tipping points in the dynamics of speciation. *Nature Ecology and Evolution*, 1, 1. <https://doi.org/10.1038/s41559-016-0001>
- O'Hara, R. B. (2005). Comparing the effects of genetic drift and fluctuating selection on genotype frequency changes in the scarlet tiger moth. *Proceedings of the Royal Society B*, 272, 211–217. <https://doi.org/10.1098/rspb.2004.2929>
- Orr, H. A. (1995). The population genetics of speciation: The evolution of hybrid incompatibilities. *Genetics*, 139, 1805–1813. <https://doi.org/10.1093/genetics/139.4.1805>
- Schultz, C. B., & Crone, E. E. (2001). Edge mediated dispersal behavior in a prairie butterfly. *Ecology*, 82, 1879–1892. [https://doi.org/10.1890/0012-9658\(2001\)082\[1879:EMDBIA\]2.0.CO;2](https://doi.org/10.1890/0012-9658(2001)082[1879:EMDBIA]2.0.CO;2)
- Sprengelmeyer, Q. D., Mansourian, S., Lange, J. D., Matute, D. R., Cooper, B. S., Jirle, E. V., Stensmyr, M. C., & Pool, J. E. (2019). Recurrent collection of *Drosophila melanogaster* from wild African environments and genomic insights into species history. *Molecular Biology and Evolution*, 37, 627–638. <https://doi.org/10.1093/molbev/msz271>

- Van Belleghem, S. M., Baquero, M., Papa, R., Salazar, C., McMillan, W. O., Counterman, B. A., Jiggins, C. D., & Martin, S. H. (2017). Patterns of z chromosome divergence among *Heliconius* species: the importance of historical demography. *Molecular Ecology*, 27, 3852–3872. <https://doi.org/10.1111/mec.14560>
- Veller, C., Kleckner, N., & Nowak, M. A. (2018). A rigorous measure of genome-wide genetic shuffling that takes into account crossover positions and Mendel's second law. *Proceedings of the National Academy of Sciences of the United States of America*, 116, 1659–1668. <https://doi.org/10.1073/pnas.1817482116>
- Zhang, J., Cong, Q., Shen, J., Opler, P. A., & Grishin, N. (2019). Genomics of a complete butterfly continent. <https://www.biorxiv.org/content/10.1101/829887v1>

SUPPORTING INFORMATION

Additional supporting information may be found online in the Supporting Information section.

How to cite this article: Nelson, E. D., Cong, Q., & Grishin, N. V. (2021). Influence of the large-Z effect during contact between butterfly sister species. *Ecology and Evolution*, 11, 11615–11626. <https://doi.org/10.1002/ece3.7785>

APPENDIX A

HYBRID INTERACTION MODEL

To model hybrid interactions, we select a small number of loci in autosomes for which $p_{i,1} \sim 1$ and $p_{i,2} \sim 0$ to interact negatively with loci in the Z chromosome(s) for which $p_{j,1} \sim 0$ and $p_{j,2} \sim 1$; we then repeat this process with the population subscripts interchanged, selecting an equal number of loci in autosomes with $p_{i,2} \sim 1$ and $p_{i,1} \sim 0$ to interact negatively with loci in the Z chromosome(s) for which $p_{j,2} \sim 0$ and $p_{j,1} \sim 1$. Figure A1 illustrates a pair of such interactions in a male F1 hybrid genome formed at the time of contact between diverged populations in our model. Chromosomes from populations 1 and 2 are colored red and green (respectively) with gene segments denoted by shaded blocks, and selected loci in genes denoted by vertical lines. Typically, $p_{i,j} = 1, 0$ for selected loci, however, this is not always true for the simulations in Figure 6a due to the shorter time spent in allopatry.

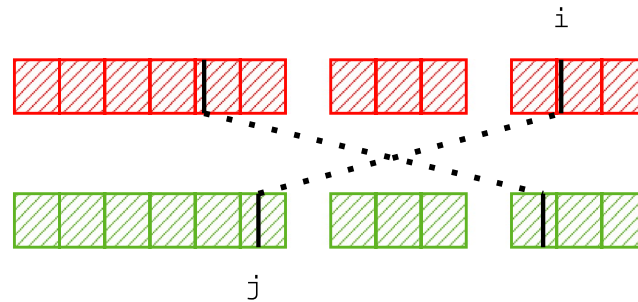


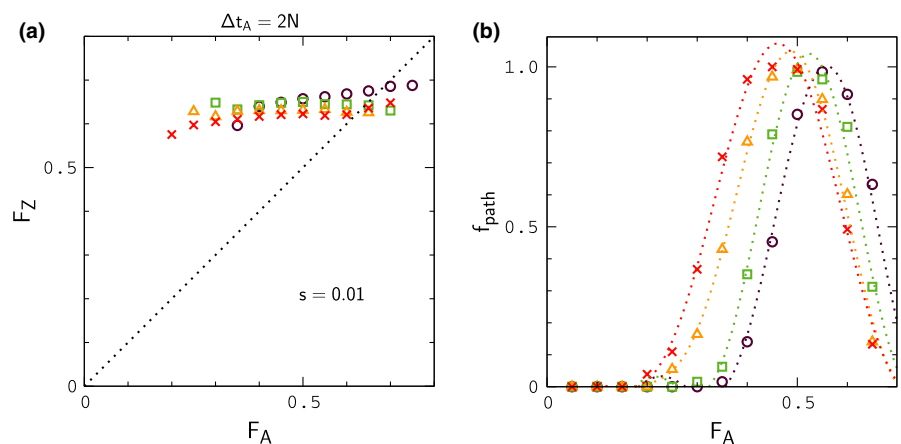
FIGURE A1 Hybrid genome formed at the point of contact between model populations. Chromosomes from populations 1 and 2 are colored red and green, with gene segments indicated by shaded blocks. The illustration shows a pair of interactions (dotted lines) connecting selected alleles (solid lines) in autosomes (right) to selected alleles in the Z chromosomes (left) in a male F1 hybrid

APPENDIX B

LOWER MIGRATION RATES

As the migration rate $N\epsilon$ is decreased, a point is reached where plots of I_A versus F_A and I_Z versus F_Z explicitly resemble those in Figure 1, concurrent with large values of ΔF (Figures B1 and B2).

FIGURE B1 Study of F_Z versus F_A for low migration rates. Data points in panel (a) denote averages of F_Z for migration rates $N\epsilon = 0.125$ (circles), 0.25 (squares), 0.5 (triangles), and 0.75 (crosses). Each set of averages is computed from 128 replicate simulations with $N = 10^4$, $2\mu = 10^{-4}$, $r, r' = 10^{-2}$ and $\Delta t_C = N$. Panel (b) describes the fraction of simulation paths that reach a given bin for F_Z



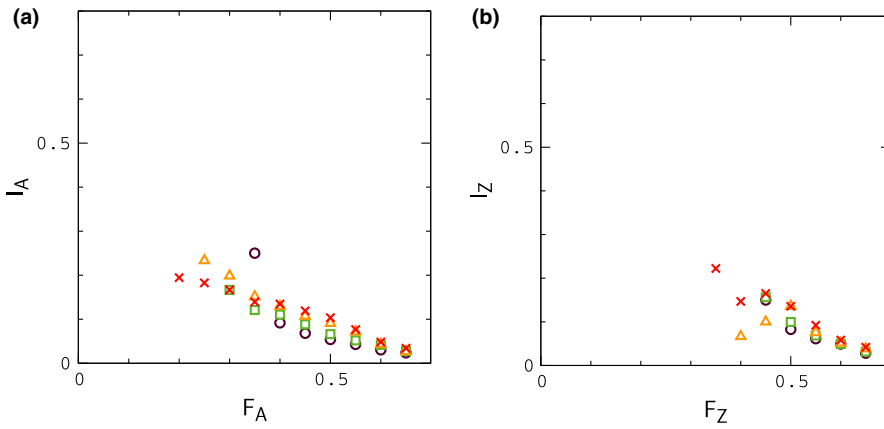


FIGURE B2 Averages for I_A versus F_A and I_Z versus F_Z obtained from the simulations in Figure B1

APPENDIX C

WEAK LINKAGE BETWEEN GENES

As the rate of crossing over between genes is increased from its value for closely linked genes, $r' = r$, differences between F_Z versus F_A decrease (Figure C1).

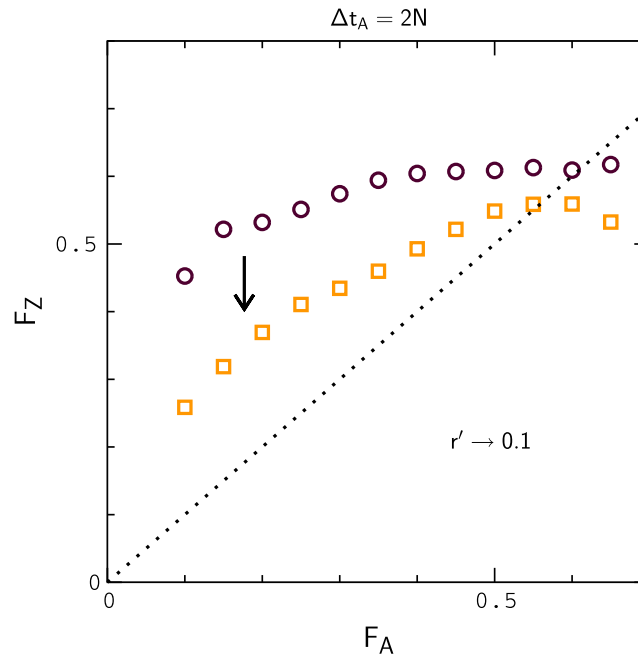


FIGURE C1 Study of F_Z versus F_A as the rate of crossing over between gene sequences is increased. Data points denote averages of F_Z versus F_A for crossover rates $r' = 0.01$ (circles), and $r' = 0.1$ (squares). Averages are computed from 128 replicate simulations with $s = 0.01$, $N = 10^4$, $2\mu = 10^{-4}$, $r = 10^{-2}$, and $\Delta t_C = N$