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# The association between immunerelated adverse events and the prognosis of solid cancer patients treated with immunotherapy: a systematic review and meta-analysis

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# Abstract

**Background:** Immune-related adverse events (irAEs) are common during immune checkpoint inhibitor (ICI) treatment and reported to be associated with good survival. This study evaluated the association between onset timing of irAEs and survival of cancer patients treated with ICIs. **Methods:** Databases including PubMed, Embase, and the Cochrane library were systematically searched to retrieve clinical studies assessing the relationship between irAEs and survival in cancer patients with ICIs. The overall response rate for treatment response and hazard ratio (HR) for overall survival (OS) and progression-free survival (PFS) were calculated using RevMan 5.3. Subgroup analysis in terms of cancer type, ICIs type, region, specific irAEs, accordingly.

**Results:** A total of 34 studies were included. The HRs for OS and PFS in cancer patients with *versus* without irAEs were 0.57 [95% confidence interval (CI): 0.44, 0.74; p < 0.0001], and 0.50 (95% CI: 0.37, 0.67; p < 0.00001), respectively. The odds ratio for overall response in cancer patients with irAEs was 4.72 (95% CI: 3.48, 6.40; p < 0.00001) compared with those without irAEs. Subgroup analyses suggested that the prognostic role of irAEs was associated with cancer types and region, but not irAEs types. The landmark analysis of OS revealed that there is a non-proportional (early) effect of irAEs on OS in ICI-treated cancer patients (landmark >12 weeks, HR<sub>0S</sub> = 1.08; 95% CI: 0.89, 1.30; p = 0.46).

**Conclusion:** Our findings suggest that the occurrence of irAEs could be a prognostic factor for cancer patients who were treated with ICIs.

*Keywords:* biomarker, immune-related adverse events, immunotherapy, meta-analysis, survival

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#### Introduction

Immunotherapy is one of the most promising treatment strategies against various tumors.<sup>1,2</sup> Immune checkpoint inhibitors (ICIs) have been widely used in clinical practice,<sup>2</sup> with improved response rates and prolonged survival of cancer patients.<sup>2,3</sup> However, not all the patients gain benefits due to numerous difficulties, such as tumor heterogeneity and host immunity status.<sup>3,4</sup>

Therefore, it is essential to explore potential prognostic factors to predict who could have better outcome from immunotherapy.

As a result of enhanced or improved host immunity by ICIs, immune-related adverse events (irAEs) are often reported in clinical trials.<sup>5–11</sup> Results of several meta-analyses suggest that the commonly affected organs are skin, endocrine,

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gastrointestinal tract and liver.<sup>12–14</sup> Recent research attention has focused on the relation between irAEs and treatment outcomes.<sup>10,11,15</sup> Results from these trials show that cancer patients with irAEs have better efficacy than those without irAEs.<sup>10,16</sup> However, there are also other views against a positive relation between irAEs and overall survival (OS) of ICI-treated cancer patients.<sup>17–19</sup> In addition, the onset time of irAEs varied between individual patients and ICI agents, making it difficult to examine the actual prognostic role of irAEs in survival of cancer patients. Therefore, whether irAEs can serve as a biomarker for immunotherapy is still in debate.

In this study, we systematically searched databases to identify clinical studies assessing the effects of irAEs on treatment outcomes and survival of cancer patients treated with ICIs, and aimed to evaluate the relation between irAEs and efficacy and survival in cancer patients receiving ICIs.

# **Material and methods**

This study was performed according to the PRISMA and the Cochrane handbook guide-lines. This study was not registered.

# Search strategy

Electronic databases including Embase, PubMed, and the Cochrane library were searched until February 2019. Search terms were "Immunerelated adverse events or irAEs or irAE or treatment related adverse events," "cancer or tumor or neoplasm," "immune checkpoint inhibitors or immune checkpoint blockades or PD-1 inhibitors or PD-L1 inhibitors or CTLA-4 inhibitors." These terms were used in different combinations. There were no language restrictions during the search.

# Inclusion and exclusion criteria

Clinical studies including randomized controlled trials, and retrospective studies were included. Inclusion criteria: (1) Cancers were diagnosed by sufficient clinical evidence, such as pathology or cytology; (2) patients were treated with ICIs including PD-1, PD-L1, and/or CTLA-4 inhibitors alone or in combination; (3) survival data [OS, progression-free survival (PFS), and/or time to treatment failure (TTF)] in cancer patients with *versus* without irAEs were reported; treatment

response measures were reported in the included studies; (4) if results from the same patient sources were published by different journals, the study with the most complete or up-to-date data was included; (5) eligible studies were not only fulltext, but also abstract, conference meeting presentation, and unpublished literature.

Exclusion criteria: Studies were excluded if (1) insufficient data on baseline information, efficacy or survival; (2) reviews, animal studies, comments, survey, and guidelines.

### Study selection and data extraction

The screen for eligible studies was conducted by two researchers (HX and DC), independently. Any inconsistency was solved through discussion. The following information was extracted: first name of the first author, publication time, region, number of patients, sex (male), age, cancer type, immunotherapy agent, reported specific irAEs, objective response rate (ORR), hazard ratios (HRs) of irAEs *versus* no irAEs for OS, PFS, and TTF based on landmark analysis or not.

# Quality assessment

To evaluate the quality of the retrospective studies, the Newcastle–Ottawa scale (NOS) method was introduced.<sup>20</sup> According to the protocol of the NOS, three major aspects are focused on during evaluation: selection, comparability of the cohort, and evaluation of the results. According to the instruction of the NOS, a maximum of four stars, two stars, and three stars can be given to the selection, the comparability, and the results assessment, respectively. A good quality study was defined as having six or more stars.

# Statistical analysis

The RevMan 5.3 software was used to combine the individual HR and its related 95% confidence interval (CI). GraphPad Prism 6 was used to draw plots. Engauge software was used to extract survival rate at various time-points from survival plots. Q test and  $I^2$  statistic were introduced to calculate the heterogeneity among the included studies. A significant heterogeneity was considered if p < 0.1 or  $I^2 > 50\%$ , and the random-effects model was used. If p > 0.1 or  $I^2 < 50\%$ , the fixedeffect model was used. For the pooled estimate, it was considered statistically significant if p < 0.05.



**Figure 1.** Flow chart of identifying eligible studies and characteristics of irAEs. (A) Flow chart of identifying eligible studies; (B) individual incidence of irAEs among included studies; and (C) onset time (median days) of irAEs in individual studies.

To detect the impact of time on the prognostic role of irAEs, we used landmark data from individual studies to perform the meta-analysis. We also calculated the odds ratio (OR) of survival rates at 2, 4, 6, 8, 10, 12, 18, 24, and 30 months using the individual data (number of death, number at risk) of the included studies. The OR was calculated as following: (death events/number at risk)<sub>non-irAE</sub>/ (death events/number at risk)<sub>irAE</sub>.

The individual HRs of irAEs *versus* no irAEs were extracted. For studies that did not present HRs directly, reported methods<sup>21,22</sup> were used to calculate the HR. The overall HR<1 indicated that appearance of irAEs was associated with better outcomes for cancer patients treated with ICIs. If HR>1, it indicated patients with irAEs had poor outcomes. Treatment response rates after ICIs treatment were also extracted to determine the influence of irAEs on treatment efficacy of ICIs. Subgroup analyses of survival were performed with regard to cancer type, ICIs type, region, specific irAEs, and number of irAEs. Funnel plot was used

to detect the publication bias, and a p < 0.05 suggested that there was a significant publication bias.

### Results

### Search results

A total of 760 relevant articles were retrieved after the preliminary search. After removing 118 duplications, the title and abstract of the remaining 642 studies were screened; 601 of them were discarded as they were animal studies, comments, reviews, and brief reports. After reading the full text, a further seven articles were excluded because of insufficient data, and the remaining 34 studies with 5840 patients were considered as eligible studies.<sup>10,11,15–19,23–49</sup> Figure 1A shows the details of the literature screen and selection.

### Baseline characteristics of included studies

Table 1 presents the baseline characteristics of the included studies. Of these studies 94% were

Table 1. Baseline ch	naracteristics of ir	ncluded stu	dies.						
Author name	Design	irAEs+	irAEs-	Region	Age	Sex, male	Patient type	Treatment	Outcomes
Ali et al. <sup>25</sup>	retrospective	7	33	Europe	66 [46–88]	22	advanced NSCLC	Nivolumab	efficacy, irAE
Arbour <i>et al.</i> <sup>29</sup>	retrospective		640	North America	29-93	232	advanced NSCLC	Single-agent PD- (L)1 inhibitor	efficacy, OS, PFS, irAE
Dumenil <i>et al.</i> <sup>43</sup>	retrospective	47	20	Europe	69	46	advanced NSCLC	Nivolumab	efficacy, OS, PFS, irAE
Faje <i>et al</i> . <sup>26</sup>	retrospective	64	216	North America	63	67	Melanoma	Ipilimumab	OS, TTF, irAE
Freeman and Weber <sup>38</sup>	retrospective	87	56	North America	٨A	NA	advanced melanoma	Nivolumab	0S, irAE
Freeman-Keller et al. <sup>19</sup>	retrospective	101	47	North America	17-90	87	Melanoma	Nivolumab	efficacy, OS, PFS, irAE
Fucà <i>et al.</i> <sup>44</sup>	retrospective	151		Europe	65	89	advanced NSCLC	PD-1 or PD-L1 inhibitors	efficacy, OS, PFS, irAE
Fujii et al. <sup>31</sup>	retrospective	98	192	North America	59 [19–86]	136	advanced cancer	Immunotherapy drug	efficacy, OS, PFS, irAE
Nakamura <i>et al.</i> <sup>35</sup>	retrospective	6	26	Asia	40-85	18	Melanoma	Nivolumab	efficacy, OS, PFS, irAE
Haratani <i>et al.</i> <sup>11</sup>	retrospective	69	65	Asia	68 [33–85]	06	advanced NSCLC	Nivolumab	efficacy, OS, PFS, irAE
Horvat <i>et al.</i> <sup>18</sup>	retrospective	254	77	North America	65 (21–93)	182	Melanoma	Ipilimumab	0S, TTF, irAE
Hua <i>et al.</i> <sup>49</sup>	prospective	17	50	Europe	54 (20–74)	NA	advanced melanoma	Pembrolizumab	efficacy, OS, irAE
Indini et al. <sup>47</sup>	retrospective	102	71	Europe	62 [18–85]	107	advanced melanoma	Nivolumab or Pembrolizumab	efficacy, OS, PFS, irAE
Judd <i>et al.</i> <sup>17</sup>	retrospective	64	160	North America	65	101	advanced cancer	Nivolumab or Pembrolizumab	efficacy, OS, PFS
Kim et al. <sup>33</sup>	retrospective	19	39	Asia	63 [49–68]	43	advanced NSCLC	Nivolumab or Pembrolizumab	efficacy, OS, PFS, irAE
Kothari <i>et al.</i> <sup>28</sup>	retrospective	28	147	North America	68 [33–88]	NA	advanced NSCLC	Nivolumab	efficacy, OS, PFS, irAE
Lisberg <i>et al.</i> <sup>34</sup>	retrospective	39	58	North America	32-83	50	NSCLC	Pembrolizumab	efficacy, OS, PFS, irAE
									(Continued)

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Table 1. (Continued									
Author name	Design	irAEs+	irAEs-	Region	Age	Sex, male	Patient type	Treatment	Outcomes
Margiotta <i>et al.</i> <sup>46</sup>	retrospective	163		North America	NA	NA	advanced cancer	PD-1 or PD-L1 inhibitors	efficacy
Mian <i>et al.<sup>27</sup></i>	retrospective	510	348	North America	69	NA	Melanoma	Ipilimumab	OS, irAE
Fujisawa <i>et al.</i> 37	retrospective	47	13	Asia	63 (31–85)	30	advanced melanoma	Nivolumab, followed with Ipilimumab	efficacy, OS, irAE
Owen <i>et al</i> . <sup>30</sup>	retrospective	24	66	North America	67	NA	NSCLC	Nivolumab	0S, irAE
Pawel <i>et al.</i> <sup>40</sup>	retrospective	132	293	North America	NA	NA	advanced NSCLC	Atezolizumab	OS
Ricciuti <i>et al.</i> <sup>48</sup>	retrospective	85	110	Europe	63 [30–84]	128	advanced NSCLC	Nivolumab	efficacy, OS, PFS, irAE
Rogado <i>et al.</i> <sup>16</sup>	retrospective	10	30	Europe	NA	AN	advanced NSCLC	Nivolumab	efficacy, OS, PFS, irAE
Santini <i>et al.</i> <sup>24</sup>	retrospective	20	18	North America	42-84	20	advanced NSCLC	PD-1 or PD-L1 inhibitors	efficacy, OS, PFS, irAE
Sato <i>et al.</i> <sup>15</sup>	retrospective	11	27	Asia	69 [49–86]	28	advanced NSCLC	Nivolumab	efficacy, OS, PFS, irAE
Scott and Pennell <sup>41</sup>	retrospective	210		North America	68 [60-74]	113	advanced NSCLC	Nivolumab	0S, irAE
Shah <i>et al.</i> <sup>45</sup>	retrospective	141		North America	NA	AN	advanced NSCLC	PD-1 or PD-L1 inhibitors	efficacy
Taniguchi <i>et al.</i> <sup>42</sup>	retrospective	201		Asia	68 [27–87]	135	NSCLC	Nivolumab	efficacy, OS, PFS, irAE
Teraoka <i>et al.</i> <sup>23</sup>	prospective	19	24	Asia	70 (50–82)	27	advanced NSCLC	Nivolumab	efficacy, OS, PFS, irAE
Toi <i>et al.</i> <sup>36</sup>	retrospective	29	41	Asia	68	61	advanced NSCLC	Nivolumab	efficacy, OS, PFS, irAE
Toi <i>et al.</i> <sup>10</sup>	retrospective	28	42	Asia	68 [36–88]	61	advanced NSCLC	Nivolumab	efficacy, OS, PFS, irAE
Wen <i>et al.</i> <sup>39</sup>	retrospective	35	17	Asia	53 (20–78)	31	advanced melanoma	Nivolumab, Ipilimumab	efficacy, OS, PFS, irAE
Zimmerman et al. <sup>32</sup>	retrospective	39	87	North America	58	AN	advanced melanoma	Ipilimumab	efficacy, OS, PFS, irAE
irAE, immune-relate ligand 1; PFS, progre	d adverse events; N ssion-free survival.	IA, not availat	ble; NSCLC, 1	non-small-cell lung	cancer; 0S, d	overall sur	vival; PD-1, programmed c	ell death protein-1; PD-	-L1, Programmed death-

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retrospective and only two were prospective studies. The publication years ranged from 2013 to 2019. Most of the cases were diagnosed with nonsmall-cell lung carcinoma (NSCLC) or melanoma. The incidences of irAEs among included studies ranged from 16% to 85% (Figure 1B), with an overall incidence of 44.89% (1994/4442). The median days of irAEs onset ranged from 29 to 126 (Figure 1C). The reporting regions included Asia (n=10), North America (n=17), and Europe trials<sup>10,15–17,19,23,25,26,32–36,48,49</sup> Fourteen (n=7).reported outcomes of tumor response, 20 studies<sup>11,16-19,23,26-28,31,33-40</sup> showed OS data, and 15 studies exhibited PFS data.<sup>10,11,15-18,23,28,31,33-36,47,48</sup>

### NOS quality assessment result

As most of the included trials were retrospective studies, the NOS method was applied to assess the overall quality of these studies. As shown in Table 2, five studies<sup>27,30,32,40,46</sup> had five stars, showing high risk of bias and were considered as low to moderate quality. The main reasons lowering the overall quality were selection and outcome bias. The data of incidence of irAEs from these studies were used for overall calculation of irAEs occurrence, but excluded from meta-analysis.

# Results of meta-analysis

OS. The impact of irAEs on OSin cancer patients treated with ICIs was assessed in 20 studies.<sup>11,16–19,23,26–28,31,33–40</sup> Among them, 1731 cancer patients presented at least one of the reported irAEs, and 1991 cases were absent from irAEs. The random-effect model was applied (p < 0.01,  $I^2 = 90\%$ ). The combined result (Figure 2A) showed that patients with irAEs had a significantly reduced risk of mortality compared with no irAEs group (HR=0.57; 95% CI: 0.44, 0.74; p < 0.0001).

The prognostic role of irAEs is dependent on onset time of irAEs. Next, we examined the impact of onset time of irAEs on the prognostic role of irAEs. The timings of landmark analyses in the included studies ranged from 6 weeks to 20 weeks, and we classified the studies into two groups ( $\leq 12$  and >12 weeks). As shown in Figure 2B, the patients with any irAEs still had a better OS (HR=0.60; 95% CI: 0.41, 0.89; p=0.01) than those without irAEs in the  $\leq 12$  subgroup analysis. However, when the landmark timing extended to >12 weeks, there was no significant difference between OS in patients with any irAEs versus no irAEs (HR=1.00; 95% CI: 0.84, 1.19; p=0.98).

To further evaluate the impact of time on prognostic role of irAEs, we used individual HRs from included studies with landmark analysis to draw a scatter plot (Figure 2C), and found that the HR of irAE versus no irAE was increasing over time (linear regression,  $R^2 > 0.4$ ). Next, the number of death and number at risk at various time-points extracted from the survival curve of the included studies were used to calculate a series of OR for OS and PFS, aiming to assess the influence of time on the prognostic effect of irAEs. As shown in Figure 2D, the prognostic effects of irAEs with regard to OS and PFS were decreased over time. At 2 months, the ORs of irAEs versus no irAEs for OS and PFS were 6.27 and 4.15, whereas they were 1.91 and 2.95 at 12 months, respectively. Together, these results showed that the association between irAEs and survival in cancer patients treated with ICIs was changing over time.

Specific irAEs on OS. Next, we assessed the prognostic effect of specific irAEs on OS in cancer patients treated with ICIs. There were five studies11,26,33,35,48 that reported the impact of endocrine adverse events on survival, and seven studies11,19,23,35,39,48,49 showed survival data of patients suffering skin and vitiligo events. As shown in Supplemental Figure 1A, the randomeffect model was used (p < 0.1). The pooled results showed that patients with endocrine adverse events had a 61% reduction in risk of death (HR = 0.39; 95% CI: 0.27, 0.56; p < 0.0001) compared with patients without these events. Patients that presented with skin rash or vitiligo also had a significantly lower risk of mortality (HR=0.48; 95% CI: 0.28, 0.84; p=0.009) compared with no irAEs group. When combining the data of these two groups, the overall HR was 0.43 (p=0.0003) with a low risk of heterogeneity  $(p=0.50, I^2=0\%).$ 

Number of *irAEs* on OS. Two studies<sup>19,34</sup> evaluated the impact of number of *irAEs* on OS. The results from these studies suggested that increased number of *irAEs* may be associated with better survival when comparing with those without *irAEs* or lower number of *irAEs*. As there were differences in the statistical methods, it was not appropriate to perform meta-analysis. By performing the Cox proportional hazards regression models, the study by Lisberg *et al.* found that increasing 
 Table 2. Quality assessment of included studies.

Author	Select	ion			Comparability	Outco	me		Score
	A	В	С	D	E	F	G	н	_
Ali et al. <sup>25</sup>	$\stackrel{\wedge}{\simeq}$	$\stackrel{\wedge}{\bowtie}$	$\stackrel{\wedge}{\simeq}$	$\overleftrightarrow$	**	☆	$\overleftrightarrow$	$\stackrel{\wedge}{\bowtie}$	9
Arbour <i>et al.</i> <sup>29</sup>	*	$\stackrel{\wedge}{\asymp}$	$\overset{\wedge}{\Join}$	$\overleftrightarrow$	**	☆	${\sim}$		9
Dumenil <i>et al</i> . <sup>43</sup>	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\asymp}$	$\stackrel{\wedge}{\asymp}$		**	☆	${\sim}$		7
Faje <i>et al.</i> <sup>26</sup>		$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\asymp}$		**	☆	${\sim}$		8
Freeman and Weber <sup>38</sup>	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\bowtie}$	$\stackrel{\wedge}{\asymp}$		**	☆	${\sim}$		8
Freeman-Keller <i>et al</i> . <sup>19</sup>		$\stackrel{\wedge}{\bowtie}$	$\stackrel{\wedge}{\bowtie}$		**	☆	${\leftrightarrow}$		7
Fucà <i>et al</i> . <sup>44</sup>	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\asymp}$	$\stackrel{\wedge}{\asymp}$		**	☆	${\sim}$		7
Fujii <i>et al</i> . <sup>31</sup>		$\stackrel{\wedge}{\asymp}$	$\stackrel{\wedge}{\asymp}$	$\overleftrightarrow$	**	☆	${\sim}$		8
Fujisawa <i>et al.</i> 37		$\stackrel{\wedge}{\bowtie}$	$\stackrel{\wedge}{\bowtie}$	$\overleftrightarrow$	**	☆	${\leftrightarrow}$	$\stackrel{\wedge}{\bowtie}$	9
Haratani <i>et al</i> . <sup>11</sup>		$\stackrel{\wedge}{\bowtie}$	$\stackrel{\wedge}{\bowtie}$	$\overleftrightarrow$	**	☆	${\leftrightarrow}$		8
Horvat <i>et al</i> . <sup>18</sup>	$\stackrel{\wedge}{\simeq}$	$\stackrel{\wedge}{\bowtie}$	$\stackrel{\wedge}{\bowtie}$	$\overleftrightarrow$	**	☆	${\leftrightarrow}$	$\stackrel{\wedge}{\bowtie}$	9
Hua et al.49		$\stackrel{\wedge}{\sim}$	$\overset{\wedge}{\Join}$		**	☆	${\sim}$		7
Indini <i>et al.</i> 47	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\asymp}$	$\stackrel{\wedge}{\asymp}$		**	☆	${\sim}$		8
Judd et al. <sup>17</sup>		$\stackrel{\wedge}{\asymp}$	$\stackrel{\wedge}{\asymp}$	$\overleftrightarrow$	**	☆	${\sim}$		9
Kim et al. <sup>33</sup>	$\stackrel{\wedge}{\simeq}$	$\stackrel{\wedge}{\bowtie}$	$\stackrel{\wedge}{\bowtie}$	$\overleftrightarrow$	**	☆	${\leftrightarrow}$	$\stackrel{\wedge}{\bowtie}$	9
Kothari <i>et al.</i> <sup>28</sup>		$\stackrel{\wedge}{\bowtie}$	$\stackrel{\wedge}{\bowtie}$		**	☆	${\leftrightarrow}$		7
Lisberg <i>et al</i> . <sup>34</sup>	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\asymp}$	$\stackrel{\wedge}{\bowtie}$		**	☆	$\stackrel{\sim}{\sim}$		7
Margiotta <i>et al.</i> <sup>46</sup>		$\stackrel{\wedge}{\asymp}$	$\stackrel{\wedge}{\asymp}$		**	☆	${\sim}$		7
Mian <i>et al</i> . <sup>27</sup>	$\stackrel{\wedge}{\simeq}$	$\stackrel{\wedge}{\bowtie}$			**	☆	${\leftrightarrow}$		6
Nakamura <i>et al.</i> <sup>35</sup>		$\stackrel{\wedge}{\bowtie}$	$\stackrel{\wedge}{\bowtie}$	$\overleftrightarrow$	**	☆	${\leftrightarrow}$		9
Owen <i>et al</i> . <sup>30</sup>	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\asymp}$	$\stackrel{\wedge}{\asymp}$		**	☆	${\sim}$		7
Pawel <i>et al</i> . <sup>40</sup>		$\stackrel{\wedge}{\asymp}$	$\overset{\wedge}{\Join}$		**	☆	${\sim}$		7
Ricciuti <i>et al</i> .48	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\overset{\wedge}{\Join}$	$\stackrel{\wedge}{\sim}$	**	☆	$\stackrel{\sim}{\sim}$	$\stackrel{\wedge}{\bowtie}$	9
Rogado <i>et al</i> . <sup>16</sup>		$\stackrel{\wedge}{\asymp}$	$\stackrel{\wedge}{\asymp}$		**	☆	${\sim}$		7
Santini <i>et al.</i> <sup>24</sup>	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\asymp}$	$\stackrel{\wedge}{\asymp}$		**	☆	${\leftrightarrow}$		7
Sato <i>et al.</i> <sup>15</sup>	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\overleftrightarrow$	**	☆	$\overleftrightarrow$		8
Scott and Pennell <sup>41</sup>	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	**	☆	$\stackrel{\wedge}{\sim}$		9
Shah et al.45	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$			☆	$\overleftrightarrow$		6
Taniguchi <i>et al.</i> <sup>42</sup>	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	**	☆	$\stackrel{\wedge}{\sim}$		9
Teraoka <i>et al.</i> <sup>23</sup>	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\overleftrightarrow$	**	☆	$\overleftrightarrow$		9
Toi <i>et al</i> . <sup>10</sup>	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$		☆	$\stackrel{\wedge}{\sim}$		8
Toi et al. <sup>36</sup>	$\overleftrightarrow$	$\stackrel{\wedge}{\sim}$	$\stackrel{\wedge}{\sim}$		**	☆	$\overleftrightarrow$		7
Wen <i>et al.</i> <sup>39</sup>	$\overleftrightarrow$	$\overset{\sim}{\sim}$	$\overset{\sim}{\sim}$		**	☆	$\overleftrightarrow$	☆	8
Zimmerman et al. <sup>32</sup>	${\sim}$	$\stackrel{\sim}{\sim}$	$\stackrel{\sim}{\sim}$		**	☆	$\overleftrightarrow$		7

Note: "Selection" part includes A: representativeness of cases, B: selection of controls, C: exposure ascertainment, and D: no death when investigation begin. "Comparability" part includes E: comparable on confounders. "Outcome" part includes F: outcome assessment, G: adequate follow-up, and H: loss to follow-up rate. The total score is equal to the total number of stars.



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numbers of irAEs was associated with a trend toward improved OS (unadjusted HR=0.77; p=0.079 and adjusted HR=0.72; p=0.088).<sup>34</sup> The study of Freeman-Keller *et al.* showed that compared with patients who reported two or fewer irAEs, OS benefit was observed in those with three or more irAEs (HR=0.53; p < 0.001).<sup>19</sup>

High-grade irAEs and OS. Only two studies<sup>17,31</sup> reported survival data of patients with high-grade irAEs (grade 3 or higher) *versus* no irAEs. These patients were diagnosed with various types of cancer. However, the heterogeneity between these two studies was low, as indicated by the test (p=0.92,  $I^2=0\%$ ). The pooled result failed to determine a significant role for high-grade irAEs in predicting the survival of cancer patients treated with ICIs, though there was a trend (HR=0.61; 95% CI: 0.38, 1.00; p=0.05).

# PFS

Fifteen studies<sup>10,11,15-18,23,28,31,33-36,47,48</sup> reported the PFS of cancer patients with or without irAEs (Figure 3A). A significant heterogeneity between included studies was found (p < 0.01,  $I^2 = 84\%$ ), and the random-effect model was introduced to minimize the impact of differences. The overall result showed that patients with irAEs had a lower risk of disease progression (HR=0.50; 95% CI: 0.37, 0.67; p < 0.00001) when compared with those without irAEs. We also performed a metaanalysis based on the data from landmark analysis (Figure 3B). The pooled result showed that occurrence of irAEs was associated with better PFS (HR=0.68; 95% CI: 0.55, 0.85; p=0.0006) in patients receiving ICIs, though some of the included studies showed negative conclusions.

### ORR

Fourteen studies<sup>10,15–17,19,23,25,26,32–36,48,49</sup> reported ORRs in ICIs-treated cancer patients presenting irAEs *versus* no irAEs. The random-effect model was used as the subgroup  $I^2$  was 59.3%. As presented in Figure 3C, appearance of irAEs in cancer patients was associated with an improved ORR when compared with those without irAEs (OR=4.72; 95% CI: 3.48, 6.40; p < 0.00001), indicating the ICIs efficacy was almost five times better in patients with irAEs.

We introduced subgroup analysis to evaluate whether the prognostic role of irAEs related to treatment response was independent of cancer types or not. The results showed that the relation of irAEs and efficacy was independent of cancer types. The HRs for NSCLC, melanoma and other cancers were 6.53 (95% CI: 4.22, 10.10; p < 0.0001), 3.86 (95% CI: 2.26, 6.58; p < 0.00001), and 2.70 (95% CI: 1.30, 5.61; p = 0.008), respectively.

### Subgroup analysis of OS and PFS

*Cancer type.* The prognostic role of irAEs on OS (HR=0.58; 95% CI: 0.39, 0.87; p=0.008) and PFS (HR=0.45; 95% CI: 0.33, 0.61; p < 0.00001) were significant in NSCLC patients treated with ICIs (Supplemental Figure 2A and 2B). Similar results were observed for OS (HR=0.68; 95% CI: 0.53, 0.87; p=0.002), but not for PFS (HR=0.51; 95% CI: 0.23, 1.15; p=0.10) in melanoma patients.

Drug type. The common agents were Nivolumab and Ipilimumab. The subgroup analysis showed that irAEs was a significant predictor for OS (HR=0.58 for Nivolumab; HR=0.64 for Ipilimumab; HR=0.48 for other, p < 0.05 for all) and PFS (HR=0.45 for Nivolumab; HR=0.57 for other, p < 0.01 for all), suggesting the prognostic role of irAEs was not dependent on the types of the immune checkpoint blockades (Supplemental Figure 3A and 3B).

Region. The studies were reported from Asia, North America and Europe. The subgroup analysis showed that OS of cancer patients from Asia (HR=0.51; 95% CI: 0.31, 0.83; p=0.006) and North America (HR=0.57; 95% CI: 0.41, 0.78; p = 0.0004) was better if they had irAEs, but not Europe (HR = 0.70; 95% CI: 0.32, 1.57; p = 0.39), indicating that the predictive role of irAEs on OS was dependent on region (Supplemental Figure 4A). As shown in Supplemental Figure 4B, the prognostic role of irAEs still worked on PFS in patients in Asia (HR=0.35; 95% CI: 0.27, 0.46; p < 0.00001), but the prognostic role of irAEs on PFS was not significant in cancer patients from North America (HR=0.87; 95% CI: 0.65, 1.16; *p*=0.33) and Europe (HR=0.62; 95% CI: 0.31, 1.23; *p*=0.17).

Assessment of publication bias. The funnel analysis of the included studies was conducted using OS, PFS and ORR data. The symmetry of the funnel graph is good (Supplemental Figure 5), suggesting that the results are less likely to be affected by publication bias.

(A)			An	y irAEs	No irAEs		Hazard I	Ratio	Hazard Ratio
Study or Subgroup	log[Hazard Ra	tio]	SE	Total	Tota	l Weight	IV, Rando	<u>m, 95% Cl</u>	IV, Random, 95% Cl
Fujii2018	-0.91	163 0	.3537	98	192	2 6.7%	0.40 [0	0.20, 0.80]	
Haratani2018	-0.64	144 0	.3081	69	65	5 7.3%	0.52 [0	0.29, 0.96]	_ <b></b> _
Horvat2015	-0.02	108 0. 755 0	3021	254	44	+ 9.8% I 7.4%	0.96 [0	0.78, 1.18]	]
Judd2017	-0.21	107 0	.1282	64	160	9.6%	0.81 [0	0.63, 1.04]	
Kim2017	-0.96	676 0	.4104	19	39	5.9%	0.38 [0	0.17, 0.85]	
Kothari2017	0.26	624 0	.6014	28	147	4.0%	1.30 [0	0.40, 4.23]	
Lisberg2018	-0.95	589 0	.2267	39	58	8 8.4%	0.38 [0	0.25, 0.60]	
Nakamura2016	-1.42	271	0.398	9	26	6.1%	0.24 [0	0.11, 0.52]	
Ricciuti2019	-0.85	310 0	.1594	85	110	) 9.3% ) 5.9%	1.50.00	0.30, 0.56]	
Sato2018	-2.30	026 0	.6675	10	27	7 3.5%	0.10 [0	0.03. 0.371	
Teraoka2017	-1.04	198 0	.1139	19	24	9.7%	0.35 [0	.28, 0.44]	-
Toi2018	-0.8	344 0	.3657	28	42	6.5%	0.43 [0	0.21, 0.88]	<b>-</b> _
T-4-1 (05% OI)				005	4005	400.00/	0 50 50	07 0 071	
Hotorogonoity: Tau <sup>2</sup> = 0	22: Chi2 - 70 8	e df -	13 (P < 0	835	1035	100.0%	0.50 [0	.37, 0.67]	
Test for overall effect: Z	r = 4.51 (P < 0.0)	0001	13 (F < 0	.00001),	1 - 04 /0				0.05 0.2 1 5 20
		,							Favours Any irAEs Favours No irAEs
( <b>B</b> )			۵	nv ir∆F	s No ir AF	e	Hazard	Ratio	Hazard Ratio
Study or Subaroup	log[Hazard R	atiol	SE	Tot	al To	tal Weight	IV. Fixe	d. 95% CI	IV. Fixed, 95% CI
Haratani2018	-0.6	6125	0.2975	e	59	65 14.0%	0.54 [0.	.30. 0.971	
Kothari2017	0.2	2624	0.5473	2	.8 1	47 4.1%	1.30 [0.	.44, 3.80]	
Lisberg2018	-0.2	2877	0.1417	3	9	58 61.8%	0.75 [0.	.57, 0.99]	
Ricciuti2019	-0.5	5978	0.2606	8	85 1	10 18.3%	0.55 [0.	.33, 0.92]	
Sato2018	-1	.273	0.8426	1	1	27 1.7%	0.28 [0.	.05, 1.46] -	•
Total (95% CI)				23	2 4	07 100 0%	0 88 0	55 0 851	•
Heterogeneity: $Chi^2 = i$	4 23. df = 4 (P =	0.38)	$1^2 = 6\%$	20		07 100.07	0.00 [0.		
Test for overall effect:	Z = 3.45 (P = 0.	0006)	,. 070						0.1 0.2 0.5 1 2 5 10
		,							Favours Any IrAEs Favours No IrAEs
$(\mathbf{C})$	Any irAE	s	No irA	Es		Odds F	Ratio		Odds Ratio
Study or Subgroup	D Events 1	Total	Events	Total	Weight	M-H, Fixe	ed. 95% C	I	M-H, Fixed, 95% Cl
1.1.1 NSCLC									
Ali2016	5	7	7	33	1.8%	9.29 [1.4	47, 58.47]		· · · · · · · · · · · · · · · · · · ·
Kim2017	6	19	4	39	4.7%	4.04 [0.9	98, 16.65]		
Lisberg2018	10	26	10	71	8.7%	3.81 [1.3	35, 10.73]		
Ricciuti2019	37	85	11	110	14.2%	6.94 [3.2	26, 14.78]		
Rogado2018	6	10	8	30	4.2%	4.13 [0.9	92, 18.52]		
Sato2018	7	11	2	27	1.1%	21.88 [3.29	9, 145.24]		
Teraoka2017	9	27	2	16	4.4%	3.50 [0.6	65, 18.85]		
Toi2017	17	29	3	41	2.7%	17.94 [4.4	48, 71.93]		
Subtotal (95% CI)		214		367	41.8%	6.53 [4.2	2, 10.10]		
I otal events	97	(n – o	47	00/					
Heterogeneity: Chi-	= 6.14,  at = 7	(P = 0	0.52); I* =	0%					
rest for overall elle	Cl. Z = 0.42 (P	< 0.00	5001)						
1.1.2 Melanoma									
Freeman-Keller2016	<u> </u>	41	14	71	16.4%	2 61 [1	11 6 141		<b>_</b> _
Fuiisawa2018	16	28	5	42	4.5%	9 87 12 9	38 32 651		
Hua2016	12	17	14	50	5.5%	6.17 [1.8	34, 20,741		
Zimmerman2013	11	58	4	38	10.3%	1.99 [0	.58. 6.781		
Subtotal (95% CI)		144		201	36.6%	3.86 [2	26, 6.58]		•
Total events	55		37						
Heterogeneity: Chi <sup>2</sup>	= 4.87, df = 3	(P = 0	.18); I <sup>2</sup> =	38%					
Test for overall effect	ct: Z = 4.95 (P	< 0.00	0001)						
1.1.3 Multiple canc	ers								
Fujii2018	3	12	15	252	2.7%	5.27 [1.2	29, 21.51]		
Judd2017	16	64	12	96	18.9%	2.33 [1	.02, 5.34]		
Subtotal (95% CI)		76		348	21.6%	2.70 [1	.30, 5.61]		
Total events	19	·	27	00/					
Heterogeneity: Chi <sup>2</sup>	= 0.99, df = 1	(P = 0	0.32); l <sup>2</sup> =	0%					
lest for overall effe	ct: Z = 2.66 (P	= 0.00	JQ)						
Total (95% CI)		434		916	100.0%	4,72 [3	48, 6 401		◆
Total evente	171	104	111	010	//	4.12 [0			· ·
Heterogeneity: Chi <sup>2</sup>	= 16.15 df = 1	3 (P =	= 0.24)· I	= 20%					
Test for overall effe	ct: Z = 10.00 (F	P < 0.0	00001)	-070				0.01	0.1 1 10 100
Test for subaroup d	ifferences: Chi	<sup>2</sup> = 4.9	91. df = 2	(P = 0.0	09). I <sup>2</sup> = 59	9.3%		F	avours No IrAEs Favours Any irAEs

**Figure 3.** Combined analysis of prognostic effect of any immune-related adverse events (irAEs) *versus* no irAEs on progression-free survival (PFS) and efficacy in cancer patients treated with immune checkpoint inhibitors. A, The association between irAEs and PFS; B, Assessing the association between irAEs and PFS based on landmark analysis results; C, Objective response rates in cancer patients with or without irAEs when treated with immune checkpoint inhibitors. Subgroup analysis was performed with regard to cancer types (non-small-cell lung carcinoma, melanoma, and other cancers).

# Discussion

This meta-analysis confirmed that patients with irAEs had better survival and treatment response from immune checkpoint blockades when compared with patients without irAEs. The results also indicated that the predictive value of irAEs may be dependent on onset timing of irAEs, cancer type, region, but independent of ICI type. With regard to specific irAEs, we found skin reaction and endocrine adverse events were associated with better OS than those without these events. For number and grade of irAEs, it was suggested that patients with more irAEs and higher grade irAEs may have better OS.

In recent years, several biomarker candidates have emerged and some of them are promising, such as PD-1,<sup>1,2</sup> tumor mutation burden,<sup>50</sup> sex,<sup>51</sup> and baseline neutrophil to lymphocyte ratio (NLR).<sup>20</sup> Biomarkers such as PD-1 and tumor mutation burden are accurate and reliable, as supported by high-quality clinical studies.<sup>1,2,52</sup> But these biomarkers are usually expensive and complex. Sex and NLR are recently suggested to be prognostic markers of immunotherapy in cancer patients. A meta-analysis by Conforti et al. found that the difference in efficacy between men and women treated with ICIs was significant (p = 0.0019), and they concluded that the magnitude of immunotherapy benefit was sexdependent.<sup>51</sup> Our previous study suggested that baseline NLR was also a reliable and feasible biomarker in advanced NSCLC patients treated with Nivolumab, though this finding was limited to NSCLC.<sup>20</sup> In this study, we found that irAEs was associated with better outcomes of ICIs, and this relation was dependent on time, cancer types and region. The degree and number of irAEs also impacted the OS of patients treated with ICIs.

To our knowledge, this is the first landmarkbased meta-analysis evaluating the relation between any irAEs and efficacy of ICIs and survival in various cancers with positive findings. Previously, there was a meta-analysis<sup>53</sup> evaluating the prognostic role of vitiligo development on survival of patients with stage III–IV melanoma receiving immunotherapy. They found that vitiligo development was significantly related with both better OS (HR=0.25, p=0.003) and PFS (HR=0.51, p=0.005), when compared with those without vitiligo development.<sup>53</sup> In 2016, an abstract by Prince *et al.*<sup>54</sup> suggested that AEs with checkpoint inhibitors did not predict for improved OS. It seems that there are still disagreements even after pooling individual data, and the connection between irAEs and survival fails to reach an agreement. The former study only focused on the specific vitiligo event in melanoma patients, while the latter abstract did not mention the survival data. In this study, we included eligible studies as much as possible, without limitations to cancer type, ICIs type, and region. Instead of focusing on vitiligo, we checked the differences between OS of patients with any irAEs versus no irAEs, and found the occurrence of any irAEs was a beneficial indicator for ICIs treatment in terms of OS, PFS and ORR. These findings demonstrate that irAEs are associated with better efficacy of immunotherapy, indicating irAEs could serve as an indicator of immunotherapy efficacy. Interestingly, we found that the prognostic role of irAEs may be dependent on region. The better survival benefit for Asian patients may be explained by the incidence of irAEs in this population. The study by Yang et al. suggested that the incidence of irAEs in Asian patients could be as high as 90%, possibly related to T-cell aggregation.55 Of note, we also aimed to address the association between specific irAEs and prognosis of patients receiving ICIs treatments. However, we focused only on skin and endocrine toxicities in particular due to lack of sufficient data on other irAEs.

Though a positive link is found between irAEs and efficacy and survival of cancer patients treated with ICIs, it is still not convincing in determining irAEs as a prognostic factor for immunotherapy. The timing of occurrence of irAEs varied between individuals. Patients may already exhibit favorable benefits from ICIs before the appearance of irAEs, or experience irAEs after several cycles of treatments. A treatment landmark-based study may reduce the influence of the above factor. Indeed, after performing meta-analyses based on data from landmark analysis, we find the prognostic role of irAEs is not significant when a longer timing (>12 weeks) is applied, suggesting irAEs may be a time-dependent prognostic factor.

However, a better way to assess the effect of onset time of irAEs on survival outcomes is accessing and evaluating individual data. It is suggested that irAEs are associated with antibody production and memory immune responses. Both early and late onset of irAEs should be associated with better survival upon immunotherapy. In contrast, our results found that irAEs that occurred after 12 weeks were not significantly associated with improved outcomes. We suggest that there is a non-proportional (early) effect of irAEs on PFS and OS, and this might partially explain the lack of effect of irAEs on OS when the landmark time is >12 weeks. Another reason may be the limited number of studies reporting >12 weeks landmark analysis. With regard to time-point, there were a few time-points used for landmark analysis in the included studies, such as 6 weeks (n=2), 9 weeks (n=1), 12 weeks (n=4), 14 weeks (n=1), 16 weeks (n=2), and 20 weeks (n=1). The time-point of 12 weeks was used as a cut-off as it is located in the middle of multiple cut-off values and may represent the actual impact of irAEs on survival (with relevantly sufficient studies for analysis). In fact, there is currently a lack of best cut-off time-point for landmark analysis in these patients. The impact of discontinuous ICIs on survival is not well known, though there is evidence supporting that re-challenge with ICIs is still an effective way to control malignant diseases.<sup>24,56</sup> For these retreated patients, it is also not known whether irAEs are still associated with improved outcomes or not.

There are several limitations within our study. First, although we included 34 studies, they were almost all retrospective trials with small numbers of participants, making inevitable baseline differences. Indeed, the baseline characteristics of these eligible studies differed from each other, such as number, age, sex of participants, disease type and stage, treatments, irAE definition, and outcome measurement. To minimize the impacts of these factors on survival, we used subgroup analyses in terms of cancer types, ICIs types, and region. Second, there may be publication bias. This may be related to the published literature that was mostly positive results. The ones that the analysis failed to include could be gray literature, such as unpublished literature, unpublished results due to negative results, special reports, etc. Third, risks of selection, reporting, and outcome bias existed within the included studies. Patients with treatment response experienced more cycles of ICIs, which may result in increased risk of irAEs. In addition, patients with rapid progression and irAEs after short-term ICIs treatment may be not included in the original study. Of note, there are differences between anti-PD-1 and anti-CTLA-4 in terms of the safety profile, outcomes and mechanisms. In general, anti-CTLA-4 therapy alone yields higher toxicity compared with anti-PD-1/ L1 therapy alone. When the two combined together, there generally has been an even higher rate of irAEs as well as higher response. These factors were bound to affect the results of our study. In view of the above defects and problems, it is suggested that the results of this study should be applied with caution.

### Conclusion

Our findings suggest that irAEs is a time-dependent prognostic factor for cancer patients treated with ICIs. Further research and clinical trials are needed to verify our findings.

### Authors' contributions

XX performed the study selection, data extraction and analysis, and manuscript writing.

HX performed the study selection, evaluated the quality, data extraction and analysis.

WG conducted the data extraction and quality assessment.

JL conducted the data extraction and quality assessment.

DC designed, examined the data and monitored and supervised the process of this study.

### **Conflict of interest statement**

The authors declare that there is no conflict of interest.

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# Availability of data and material

All data generated or analyzed during this study are included in this published article and its supplementary information files.

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# Supplemental material

Supplemental material for this article is available online.

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