The different hypoglycemic effects between East Asian and non-Asian type 2 diabetes patients when treated with SGLT-2 inhibitors as an add-on treatment for metformin: a systematic review and metaanalysis of randomized controlled trials

Xianzhi Li^{1,2,*}, Qianping Zhang^{3,*}, Xiaojun Zhou^{1,2}, Siyi Guo⁴, Shan Jiang⁴, Yuhan Zhang², Ruzhen Zhang², Jianjun Dong⁴, Lin Liao^{1,2}

¹Department of Endocrinology and Metabology, The First Affiliated Hospital of Shandong First Medical University & Shandong Provincial Qianfoshan Hospital, Shandong Key Laboratory of Rheumatic Disease and Translational Medicine, Shandong Institute of Nephrology, Ji-nan, Shandong 250014, China

²Department of Endocrinology and Metabology, Shandong Provincial Qianfoshan Hospital, Shandong University, Shandong Key Laboratory of Rheumatic Disease and Translational Medicine, Shandong Institute of Nephrology, Ji-nan, Shandong 250014, China

³Division of Endocrinology, Department of Internal Medicine, Dezhou Municipal Hospital, Dezhou, Shandong 253000, China

⁴Department of Endocrinology, Qilu Hospital, Cheeloo College of Medicine, Shandong University, Ji-nan, Shandong 250012, China

*Equal contribution

Correspondence to: Lin Liao, Jianjun Dong; email: liaolin@sdu.edu.cn, cwc ll@sdu.edu.cnKeywords: SGLT-2 inhibitors, metformin, Asian, type 2 diabetes, randomized controlled trial, meta-analysisReceived: January 26, 2021Accepted: March 27, 2021Published: May 11, 2021

Copyright: © 2021 Li et al. This is an open access article distributed under the terms of the <u>Creative Commons Attribution</u> <u>License</u> (CC BY 3.0), which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited.

ABSTRACT

Aims: To investigate the efficacy and safety of SGLT-2 inhibitors as an add-on treatment for metformin between Asian and non-Asian T2DM.

Methods: A systematic literature search of PubMed, EMBASE, and the Cochrane Library was performed through August 2020 with the following keywords: Sodium-Glucose Transporter 2 Inhibitors, Sodium Glucose Transporter 2 Inhibitors, SGLT2 inhibitor, SGLT-2 inhibitors, type 2 diabetes, and randomized controlled trials. Double-blinded RCTs comparing SGLT-2 inhibitors as an add-on treatment for metformin and metformin monotherapy in adults with type 2 diabetes were included. A random effects model was used to calculate overall effect sizes.

Results: 5 RCTs with 1193 Asian patients and 7 RCTs with 2098 non-Asian patients were investigated. The improvement in HbA1c and fasting blood glucose in the Asian patients (WMD, -0.73%; 95% Cl, -1.01% to -0.46%, p < 0.01; WMD, -1.51; 95% Cl, -1.81 to -1.21, p < 0.01, respectively) were both significantly better than in the non-Asians (WMD, -0.45%; 95% Cl, -0.62% to -0.29%, p < 0.01; WMD, -1.03; 95% Cl, -1.27 to -0.78, p < 0.01, respectively). The effect of weight loss was similar in the non-Asian patients and Asian patients. There was little difference in the improvement of systolic blood pressure between them. The risk of serious adverse events was not significantly increased between the Asian and non-Asian patients.

Conclusion: SGLT-2 inhibitors as an add-on treatment for metformin are more efficacious in East Asian T2DM patients than in non-Asian T2DM patients without an additional risk of severe adverse events.

INTRODUCTION

T2DM is one of the most prevalent diseases in the world and more than half of the patients are in Asia [1, 2]. So, it is important for a hypoglycemic agent to have a good efficacy and safety in Asian T2DM. According to Chinese Diabetes Society guidelines [3], if lifestyle intervention cannot reach the target blood glucose, metformin is often recommended as a first-line pharmacological treatment for patients with T2DM [4, 5]. However, as type 2 diabetes progresses, metformin monotherapy often fails to provide sustained and beneficial glycemic control [6].

The guidelines of both European Association for the Study of Diabetes (EASD) and the American Diabetes Association (ADA) recommend SGLT-2 inhibitors as a combination therapy with other antihyperglycemic drugs at any stage of T2DM and when metformin is not tolerated or ineffective, SGLT-2 inhibitors was recommend as an acceptable alternative to metformin [7–8]. There were similar recommendations in clinical practice guidelines for T2DM in Asian countries [9–11].

SGLT-2 inhibitors are a new type of oral antidiabetic drug that can selectively inhibit renal glucose reabsorption, thereby increasing urinary glucose excretion and reducing hyperglycemia. Furthermore, SGLT-2 inhibitors do not increase the incidence of hypoglycemia because their hypoglycemic mechanism is independent of insulin [12].

At present, as the prevalence of T2DM substantially increases in the Asian population [11], the efficacy and safety of this novel therapeutic regimen for Asian T2DM patients have also attracted increasing attention. There were a series of studies that evaluated the efficacy and safety of the SGLT-2 inhibitors as an add-on treatment for metformin in Asian patients with T2DM [13–17]. However, no studies have been conducted to compare the differences in efficacy and safety of SGLT-2 inhibitors as an add-on treatment for metformin between Asian and non-Asian patients up to now. Due to differences in dietary patterns or habits, body mass index, and genetic and racial backgrounds between Asian and non-Asian population [18], the effect on glycemic control and body weight control as well as the adverse effects of SGLT-2 inhibitors as an add-on treatment for metformin may be different between patients with T2DM in these two populations. Thus, we performed this systematic review and metaanalysis to evaluate the differences in efficacy and safety of this therapeutic regimen between Asian and non-Asian patients with T2DM. Furthermore, we discussed the potential causes of these conflicting results.

MATERIALS AND METHODS

This systematic review and meta-analysis was conducted and reported according to the Meta-Analysis of Observational Studies in Epidemiology (MOOSE) guidelines [19]. We searched the included studies on the basis of the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement [20].

Databases and search strategy

A systematic literature search was carried out using three online databases, EMBASE, PubMed, and the Cochrane Library, to look for randomized clinical trials directly comparing SGLT-2 inhibitors as an add-on treatment for metformin with metformin monotherapy in patients with T2DM up to August 2020, with no restrictions on the year of publication. The following Medical Subject Headings (MeSH) terms and text words were used in different search combinations: Sodium Glucose Transporter 2 Inhibitors, Sodium-Glucose Transporter 2 Inhibitors, SGLT-2 Inhibitors, SGLT2 inhibitor, type 2 diabetes, and randomized controlled trials.

These articles were qualified by two authors independently reviewing and cross-checking. A manual search was also performed for potentially relevant eligible studies from the reference lists of the included papers. Reviewers resolved their disagreements by consensus.

Study selection and criteria

Inclusion criteria were as follows: (i) studies on patients with type 2 diabetes; (ii) trials examining SGLT-2 inhibitors as an add-on treatment for metformin by using a metformin monotherapy control group; (iii) trials with a duration greater than 12 weeks; (iv) trials with at least one baseline and posttreatment efficacy or safety outcomes of interest; (v) randomized control trials; and (vi) trials that reported at least one dispersion measure [confidence interval (CI), standard deviation (s.d), or standard error of mean (s.e.m)] for treatment groups of interest.

Exclusion criteria were as follows: (i) studies not published in English; (ii) animal studies; (iii) studies investigating non-SGLT-2 inhibitors as an add-on treatment for metformin interventions; (iv) letters and comments, reviews, and meta-analyses; and (v) studies lacking necessary data for efficacy and safety analyses.

Data extraction and quality assessment

Two reviewers scanned the studies which qualified for this meta-analysis independently and extracted the corresponding data to a pre-designed form. The reviewers evaluated the study together to reach an agreement, if there was a discrepancy.

The following data were extracted: research characteristics (first author, publication year, trial duration, proportion of Asian subjects, number of study subjects); subjects' baseline characteristics (average age, sex ratio, diabetes duration, drug dosage, hemoglobin A1c (HbA1c,%), fasting plasma glucose (FPG, mmol/L), body mass index (BMI), systolic blood pressure (SBP, mmHg); outcomes (change in HbA1c, change in FPG, change in body weight (BW, kg), change in SBP, serious adverse events (SAEs, SAEs were any untoward medical events that resulted in death, were life-threatening, resulted in persistent or significant disability/incapacity, required hospitalization or prolonged hospitalization and other medically important events [16]). The same dose groups were selected to achieve maximum equivalence for research that used multiple doses of an SGLT-2 inhibitor [21].

We used the Cochrane Collaboration's risk of bias tool with Review Manager version 5.3 to evaluate the quality of the included RCTs [22]. The quality evaluation criteria included (i) random sequence generation; (ii) allocation concealment; (iii) blinding of subjects and personnel; (iv) blinding of outcome assessment; (v) incomplete outcome data; (vi) selective reporting; and (vii) other bias. Three levels were used to evaluate the risk bias of each study: "high risk", "unclear risk", and "low risk". The results of the evaluations were placed in a risk of bias summary. Details are shown in Figure 1.

Data synthesis and analysis

The primary outcome was the difference of the two therapeutic regimens in the HbA1c change from baseline between Asian and non-Asian T2DM patients. Other outcomes included (i) change from baseline in FPG; (ii) change from baseline in BW; (iii) change from baseline in SBP; and (iv) incidence of SAEs between Asian and non-Asian T2DM patients.

The meta-analyses with fixed effects were performed by computing odds ratios (ORs) and 95% CIs for the outcomes of dichotomous variables (incidence of SAEs), and weighted mean differences (WMDs) and 95% CIs were used for continuous variables (change from baseline in HbA1c, FPG, BW, and SBP) [23]. Cochran's Q (chi-squared) test was used to quantify the heterogeneity, and the I² statistic was used to evaluate the extent of inconsistency [24]: I² > 50% was considered substantial heterogeneity [25]. Funnel plots was used to assess publication bias (Supplementary Figures 6, 7). The overall quality of evidence for each outcome was assessed by the Grading of Recommendations Assessment, Development and Evaluation (GRADE) approach (Supplementary Figures 1–5). Analysis of all data was by Review Manager (RevMan 5.3) statistical software and GRADEprofiler software (version 2015, Evidence Prime Inc, McMaster University, Canada).

RESULTS

Literature search

A total of 879 potential articles were identified from the three databases. After removing duplicates, 628 articles remained. Of these, 611records were excluded based on titles and abstracts, resulting in the selection of 17 articles and then a more detailed assessment of their qualifications. Five papers were further discarded for the following reasons: (i) The duration of three studies [26-28] (78-112 weeks) were much longer than the others (12-26 weeks); (ii) The proportion of Asian patients in the paper presented by Merker et al. was about 50%, and it could not be divided into any subgroup [29]; (iii) In the paper presented by Amin et al., the proportion of males in the experimental group was significantly higher than that in the control group, which might affect experimental result [30]. Finally, 12 randomized controlled trials were included. No additional eligible studies were found by searching clinical trial registration website. (the search process is summarized in Figure 2).

Study characteristics

These 12 selected studies were published from 2010 to 2018 and included a total of 3291 patients with T2DM. The characteristics of the studies are shown in Supplementary Tables 1 and 2. In addition, the Asian subjects in this study mainly refer to patients in East Asia, such as China and Japan.

For Asian subjects, 5 trials (1193 patients) were included [13–17]. Among these, 1 trial involved treatment with canagliflozin, 1 with dapagliflozin, and 3 with ipragliflozin as an add-on treatment for metformin (Supplementary Table 1). All control groups were treated with metformin monotherapy.

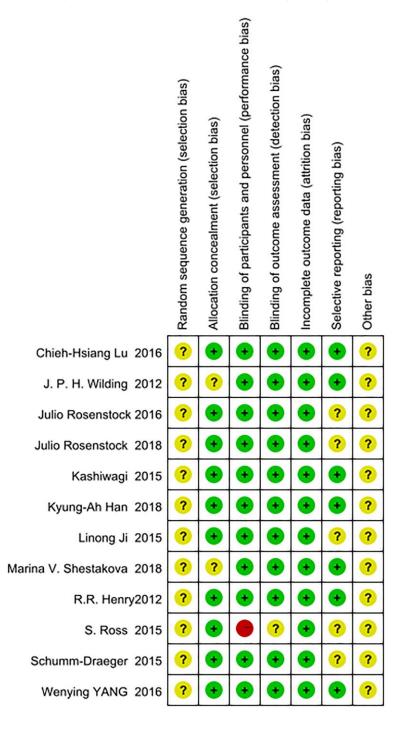
For non-Asian subjects, 7 trials (2098 patients) were included [31–37]. Among these, 1 trial used canagliflozin, 2 trials used dapagliflozin, 1 trial used empagliflozin, 2 trials used ertugliflozin and 2 trials used ipragliflozin as an add-on treatment for metformin (Supplementary Table 2). All control groups were treated with metformin monotherapy.

In these trials, the SGLT-2 inhibitors were used at the recommended dose, and metformin was used at no less than 1500 mg daily. The duration of SGLT-2 inhibitors as an add-on treatment for metformin was 12 to 26 weeks.

Efficacy

SGLT-2 inhibitors significantly improved HbA1c (%) for both Asian and non-Asian subjects as additional

treatments for metformin compared with metformin monotherapy. Moreover, the decrease in HbA1c for the Asian patients (WMD, -0.73%; 95% CI, -1.01% to -0.46%, p < 0.01) was significantly greater than that for the non-Asian patients (WMD, -0.45%; 95% CI, -0.62% to -0.29%, p < 0.01), although with high heterogeneities of I² = 87% and 78%, respectively (Figure 3A). Therefore, a sensitivity analysis by random effect model was conducted to exclude the data leading to significant heterogeneity of the Asian



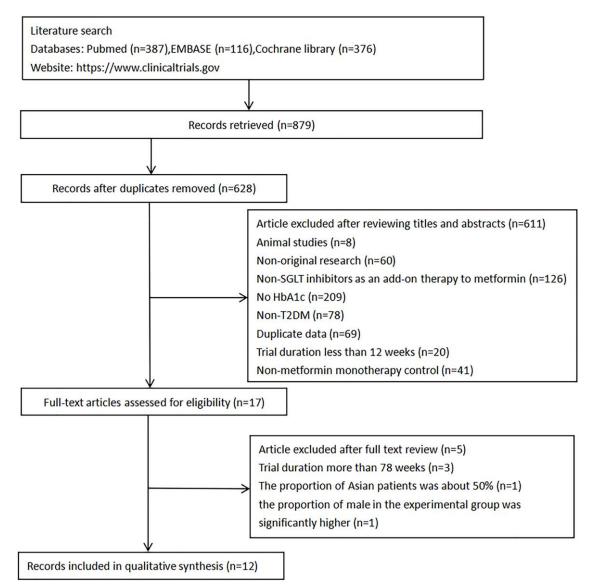


subgroup (Ipragliflozin in combination with metformin for the treatment of Japanese patients with type 2 diabetes [15]) and non-Asian subgroup (Effect of ertugliflozin on glucose control, body weight, blood pressure and bone density in type 2 diabetes mellitus inadequately controlled on metformin monotherapy [35]), and still found that HbA1c was significantly reduced more in Asian patients than in non-Asian patients: (WMD, -0.60; 95% CI, -0.75 to -0.46, p <0.01) vs (WMD, -0.39; 95% CI, -0.51 to -0.27, p <0.01) without a significant heterogeneity (I² = 39% and 49%, respectively). (Figure 3B). Test for subgroup differences: I² = 79.6% (p < 0.05). This indicated that the heterogeneity did not affect the results.

SGLT-2 inhibitors significantly decreased FPG, BW, and SBP when used as additional treatments to metformin compared with metformin monotherapy.

Similarly to the HbA1c results, SGLT-2 inhibitors reduced FPG (mmol/L) significantly more in Asian patients than in non-Asian patients as an add-on treatment for metformin: (WMD, -1.51; 95% CI, -1.81 to -1.21, p < 0.01) and (WMD, -1.03; 95% CI, -1.27 to -0.78, p < 0.01), respectively (Figure 4). The heterogeneity among all trials in the two subgroups was not significant (I² = 36% and 45%, respectively). Test for subgroup differences: I² = 83% (p < 0.01).

Comparison of body weight changes between Asian and non-Asian T2DM patients corrected by metformin monotherapy indicated no difference between the two groups: (WMD, -1.69; 95% CI, -1.98 to -1.41, p < 0.01) vs (WMD, -1.68; 95% CI, -1.97 to -1.39, p < 0.01). Details are shown in Figure 5. Test for subgroup differences: $I^2 = 0$ (p > 0.05).





There was no significant difference in the reduction of systolic blood pressure (mmHg) between the Asian and non-Asian patients: (WMD, -3.50; 95% CI, -5.30 to -1.70, p < 0.01) vs (WMD, -3.39; 95% CI, -5.73 to -1.06, p < 0.01), respectively. The total heterogeneity among all trials was not significant (I² = 28%). Test for subgroup differences: I² = 0 (p > 0.05). (Figure 6).

Safety

SGLT-2 inhibitors as an add-on treatment for metformin did not increase the risk of serious adverse events in Asians (OR = 0.90, 95% CI, 0.43 to 1.88, p = 0.78) or in non-Asians (OR = 0.93, 95% CI, 0.50 to

1.72, p = 0.81) compared with metformin monotherapy (Figure 7).

DISCUSSION

Our systematic review evaluated the clinical efficacy and safety of SGLT-2 inhibitors as an add-on treatment for metformin in comparison with metformin monotherapy in Asian and non-Asian T2DM patients. Our results indicated that the improvement in HbA1c and fasting blood glucose in Asian patients were superior to that of non-Asian patients and confirmed the hypothesis that efficacy on glucose control differs between Asian and non-Asian T2DM patients.

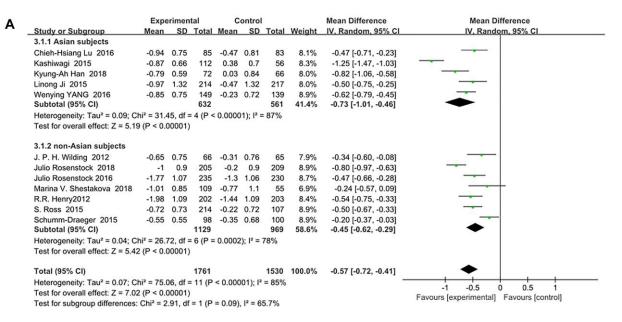


Figure 3A. Forest plot of the weighted mean difference in the change of HbA1c from baseline.

Mean Difference Mean Difference Experimental Control В Study or Subgroup Mean SD Total Mean SD Total Weight IV, Random, 95% C IV, Random, 95% CI 3.6.1 Asian subjects Chieh-Hsiang Lu 2016 -0.94 0.75 -0.47 0.81 -0.47 [-0.71, -0.23] 85 83 9.4% Kyung-Ah Han 2018 -0.79 0.59 72 0.03 0.84 66 9.2% -0.82 [-1.06, -0.58] Linong Ji 2015 -0.97 1.32 214 -0.47 1.32 217 9.0% -0.50 [-0.75, -0.25] -0.62 [-0.79, -0.45] Wenying YANG 2016 -0.85 0.75 149 -0.23 0.72 139 11.9% Subtotal (95% CI) 520 505 39.5% -0.60 [-0.75, -0.46] Heterogeneity: Tau² = 0.01; Chi² = 4.94, df = 3 (P = 0.18); l² = 39% Test for overall effect: Z = 8.33 (P < 0.00001) 3.6.2 non-Asian subjects J. P. H. Wilding 2012 -0.65 0.75 66 -0.31 0.76 65 8.7% -0.34 [-0.60, -0.08] 1.07 Julio Rosenstock 2016 -1 77 235 -1.3 1.06 230 11.0% -0.47 [-0.66, -0.28] Marina V. Shestakova 2018 -1.01 0.85 109 -0.77 1.1 55 6.6% -0.24 [-0.57, 0.09] R.R. Henry2012 -1.98 1.09 202 -1.44 1.09 203 10.3% -0.54 [-0.75, -0.33] S. Ross 2015 -0.72 0.73 214 -0.22 0.72 107 12.0% -0.50 [-0.67, -0.33] Schumm-Draeger 2015 -0.55 0.55 98 -0.35 0.68 100 11.8% -0.20 [-0.37, -0.03] 924 Subtotal (95% CI) 760 -0.39 [-0.51, -0.27] 60.5% Heterogeneity: Tau² = 0.01; Chi² = 9.82, df = 5 (P = 0.08); l² = 49% Test for overall effect: Z = 6.35 (P < 0.00001) Total (95% CI) 1444 1265 100.0% -0.47 [-0.59, -0.36] Heterogeneity: Tau² = 0.02; Chi² = 23.74, df = 9 (P = 0.005); l² = 62% -1 -0.5 ò 0.5 Test for overall effect: Z = 8.44 (P < 0.00001) Favours [experimental] Favours [control] Test for subgroup differences: Chi² = 4.90, df = 1 (P = 0.03), l² = 79.6%

Figure 3B. Forest plot of the weighted mean difference in the change of HbA1c from baseline. (sensitivity analysis).

The reasons for this discrepancy may lie in the following aspects. First, the body weight of Asians is smaller than non-Asians [38]. So, the drug doses (amount of drug/kg) of SGLT-2 inhibitors are higher in patients of the Asian group. Within a certain range, the hypoglycemic effects of SGLT-2 inhibitors are positively correlated with the dose [39, 40]. The second reason might be the different genetic and ethnic backgrounds of Asians and non-Asians, Asians have higher insulin resistance than non-Asians [41, 42] and SGLT2 inhibitors could attenuates inflammation and insulin resistance [43, 44], and could improve insulin

resistance more in Asians than non-Asians. Besides, several biological differences have been identified between patients of Asian and non-Asian patients, which might cause the difference in pharmacokinetics and pharmacodynamics of SGLT-2 inhibitor between the two groups [45].

The weight decrease corrected by metformin monotherapy for the non-Asian patients was similar to that for the Asian patients. The mechanism of weight loss caused by SGLT-2 inhibitors may be due to glycosuria, which in turn leads to energy as well as

| | Expe | rimen | tal | с | ontrol | | | Mean Difference | Mean Difference |
|---|------------|-----------|----------|-----------------------|----------------------|-------|--------|----------------------|--|
| Study or Subgroup | Mean | SD | Total | Mean | SD | Total | Weight | IV, Random, 95% CI | IV, Random, 95% CI |
| 5.1.1 Asian subjects | | | | | | | | | 53 |
| Chieh-Hsiang Lu 2016 | -1.34 | 1.98 | 87 | -0.32 | 1.72 | 83 | 8.7% | -1.02 [-1.58, -0.46] | |
| Kashiwagi 2015 | -1.23 | 1.48 | 112 | 0.59 | 1.53 | 56 | 9.9% | -1.82 [-2.31, -1.33] | |
| Kyung-Ah Han 2018 | -1.2 | 1.76 | 72 | 0.42 | 1.78 | 66 | 8.2% | -1.62 [-2.21, -1.03] | |
| Wenying YANG 2016 | -1.48 | 1.54 | 151 | 0.03 | 1.53 | 140 | 12.3% | -1.51 [-1.86, -1.16] | |
| Subtotal (95% CI) | | | 422 | | | 345 | 39.1% | -1.51 [-1.81, -1.21] | • |
| Heterogeneity: Tau ² = 0.03; C | hi² = 4.67 | 7, df = : | 3 (P = (| 0.20); l ² | = 36% | | | | |
| Test for overall effect: Z = 9.80 | 0 (P < 0.0 | 00001) | | | | | | | |
| 5.1.2 non-Asian subjects | | | | | | | | | |
| J. P. H. Wilding 2012 | -0.79 | 1.76 | 66 | -0.06 | 1.77 | 65 | 8.0% | -0.73 [-1.33, -0.13] | |
| Julio Rosenstock 2016 | -2.9 | 1.5 | 235 | -1.9 | 1.5 | 230 | 13.9% | -1.00 [-1.27, -0.73] | |
| Marina V. Shestakova 2018 | -1 | 1.79 | 109 | -0.44 | 2.76 | 55 | 5.7% | -0.56 [-1.36, 0.24] | |
| R.R. Henry2012 | -3.35 | 2.03 | 209 | -1.93 | 2.05 | 207 | 11.6% | -1.42 [-1.81, -1.03] | |
| S. Ross 2015 | -1.3 | 1.46 | 214 | 0 | 2.07 | 107 | 10.7% | -1.30 [-1.74, -0.86] | |
| Schumm-Draeger 2015 | -1.33 | 1.46 | 99 | -0.56 | 1.54 | 101 | 11.1% | -0.77 [-1.19, -0.35] | |
| Subtotal (95% CI) | | | 932 | | | 765 | 60.9% | -1.03 [-1.27, -0.78] | • |
| Heterogeneity: Tau ² = 0.04; C | hi² = 9.04 | 1, df = 1 | 5 (P = (|).11); l ² | = 45% | 5 | | | |
| Test for overall effect: Z = 8.29 | 9 (P < 0.0 | 00001) | | | | | | | |
| Total (95% CI) | | | 1354 | | | 1110 | 100.0% | -1.20 [-1.44, -0.97] | • |
| Heterogeneity: Tau ² = 0.08; C | hi² = 23.9 | 91, df = | 9 (P = | 0.004) | ; l ² = 6 | 2% | | | |
| Test for overall effect: Z = 10.0 | | - | | , | | | | | -2 -1 0 1 2 |
| Test for subgroup differences: | | | · | = 0.02) | , ² = 8 | 3.0% | | | Favours [experimental] Favours [control] |

Figure 4. Forest plot of the weighted mean difference in the change of FPG from baseline.

| | Eve | erimen | tal | c | ontrol | | | Mean Difference | Mean Difference |
|--|-----------------------|---------|-----------|-----------|----------------------|--------|--------|----------------------|--|
| Chudu an Cubanaun | | | | | | | Mainha | | |
| Study or Subgroup | Mean | 50 | Total | Mean | 50 | Total | weight | IV, Random, 95% CI | IV, Random, 95% CI |
| 7.1.1 Asian subjects | | | | | | | | | |
| Chieh-Hsiang Lu 2016 | -2.93 | | 87 | | 2.43 | 83 | 8.6% | -1.23 [-1.93, -0.53] | |
| Kashiwagi 2015 | -2.33 | 1.8 | 112 | -0.63 | 1.68 | 56 | 13.8% | -1.70 [-2.25, -1.15] | |
| Kyung-Ah Han 2018 | -1.96 | 1.94 | 72 | -0.23 | 1.69 | 66 | 11.5% | -1.73 [-2.34, -1.12] | |
| Wenying YANG 2016 | -2.6 | 2.19 | 140 | -0.7 | 2.12 | 151 | 17.1% | -1.90 [-2.40, -1.40] | |
| Subtotal (95% CI) | | | 411 | | | 356 | 51.1% | -1.69 [-1.98, -1.41] | • |
| Heterogeneity: Tau ² = 0.00 | 0; Chi ² = | 2.38, 0 | df = 3 (| P = 0.50 |); ² = | 0% | | | |
| Test for overall effect: Z = | | | | | | | | | |
| 7.1.2 non-Asian subjects | | | | | | | | | |
| J. P. H. Wilding 2012 | -2.1 | 2.65 | 66 | -0.48 | 2.65 | 65 | 5.1% | -1.62 [-2.53, -0.71] | |
| Julio Rosenstock 2016 | -3.2 | 3.08 | 237 | -1.9 | 3.08 | 237 | 13.7% | -1.30 [-1.85, -0.75] | |
| R.R. Henry2012 | -3.33 | 3.47 | 209 | -1.36 | 3.46 | 208 | 9.5% | -1.97 [-2.64, -1.30] | |
| S. Ross 2015 | -2.89 | 2.63 | 214 | -0.97 | 2.58 | 107 | 11.6% | -1.92 [-2.52, -1.32] | |
| Schumm-Draeger 2015 | -2.79 | 2.41 | 99 | -1.12 | 2.52 | 101 | 9.0% | -1.67 [-2.35, -0.99] | |
| Subtotal (95% CI) | | | 825 | | | 718 | 48.9% | -1.68 [-1.97, -1.39] | • |
| Heterogeneity: Tau ² = 0.00 | 0: Chi ² = | 3.16. | df = 4(1) | P = 0.53 | 3); ² = | 0% | | | |
| Test for overall effect: Z = | | | • | | | | | | |
| Total (95% CI) | | | 1236 | | | 1074 | 100.0% | -1.69 [-1.89, -1.48] | ◆ |
| Heterogeneity: Tau ² = 0.00 | 0: Chi ² = | 5.54. | df = 8 () | P = 0.70 | 0); ² = | 0% | | | + + + + |
| Test for overall effect: Z = | | | | | | | | | -4 -2 0 2 4 |
| Test for subgroup differen | • | | | 1 (P = 0) | 94) 12 | = 0% | | | Favours [experimental] Favours [control] |
| reaction adogroup different | 003. 011 | - 0.0 | | - (0 | | - 0 /0 | | | |

Figure 5. Forest plot of the weighted mean difference in the change of body weight from baseline.

water loss through osmotic diuresis [46, 47]. Another possible cause is the reduction in total body adipose tissue volume, which similarly involves subcutaneous and visceral deposits [48].

In addition, both in Asian and non-Asian T2DM patients, SBP decreased significantly following treatment with SGLT-2 inhibitors as an add-on treatment for metformin, and there was no significant difference between the two groups. According to current data, the decrease in blood pressure may be due

to the combination of diuresis, nephron remodeling, and weight loss [47, 49].

Regarding safety, the incidence of serious adverse events from SGLT-2 inhibitors as an add-on treatment for metformin, both in Asian and non-Asian T2DM patients, indicated no significant change compared with metformin monotherapy. There was also no significant difference in the incidence of severe adverse events between the Asian and non-Asian T2DM patients with SGLT-2 inhibitors as an add-on treatment for metformin [50, 51].

| Exp | eriment | al | c | ontrol | | | Mean Difference | Mean Difference |
|------------------------|--|---|---|--|--|---|---|---|
| Mean | SD | Total | Mean | SD | Total | Weight | IV, Random, 95% CI | IV, Random, 95% CI |
| | | | | | | | | |
| -6.8 | 14.1 | 87 | -1.3 | 12.9 | 83 | 8.9% | -5.50 [-9.56, -1.44] | |
| -1.2 | 11.7 | 112 | 2.4 | 13.7 | 56 | 8.5% | -3.60 [-7.79, 0.59] | |
| -2.35 | 10.47 | 74 | -1.14 | 9.76 | 68 | 12.3% | -1.21 [-4.54, 2.12] | |
| -2.5 | 14.38 | 141 | 1.8 | 12.13 | 128 | 13.2% | -4.30 [-7.47, -1.13] | |
| | | 414 | | | 335 | 42.8% | -3.50 [-5.30, -1.70] | ◆ |
| 0; Chi ² : | = 3.00, (| df = 3 (| P = 0.39 | 9); l² = 0 | % | | | |
| : 3.80 (P | = 0.00 | 01) | | | | | | |
| s | | | | | | | | |
| -3.8 | 10.8 | 66 | -0.5 | 10.9 | 65 | 10.3% | -3.30 [-7.02, 0.42] | |
| -2.2 | 9.2 | 237 | -0.3 | 9.2 | 237 | 29.0% | -1.90 [-3.56, -0.24] | |
| -3.8 | 10.24 | 214 | 1.6 | 11.38 | 107 | 17.9% | -5.40 [-7.96, -2.84] | |
| | | 517 | | | 409 | 57.2% | -3.39 [-5.73, -1.06] | ◆ |
| 6; Chi ² : | = 5.11, 0 | df = 2 (| P = 0.08 | B); I ² = 6 | 1% | | | |
| = 2.85 (P | = 0.00 | 4) | | | | | | |
| | | 931 | | | 744 | 100.0% | -3.37 [-4.70, -2.04] | ◆ |
| 37: Chi ² : | = 8.31. | df = 6 (| P = 0.22 | 2): $ ^2 = 2$ | 8% | | | |
| | | • | | ,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,, | | | | -10 -5 0 5 10 |
| | | | | | 0.01 | | | Favours [experimental] Favours [control] |
| | Mean -6.8 -1.2 -2.35 -2.5 -3.80 (P s -3.8 -2.2 -3.8 -2.2 -3.8 -2.2 -3.8 -2.2 -3.8 -2.2 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -2.5 -3.8 -3.8 -2.5 -3.8 -2.5 -3.8 -3.8 -2.5 -3.8 -3.8 -2.5 -3.8 -3.8 -2.5 -3.8 -3.8 -3.8 -3.8 -3.8 -3.8 -3.8 -3.8 | Mean SD -6.8 14.1 -1.2 11.7 -2.35 10.47 -2.5 14.38 00; Chi ² = 3.00, 0 3.80 (P = 0.000 s -3.8 10.8 -2.2 9.2 -3.8 10.24 46; Chi ² = 5.11, 0 2.85 (P = 0.000 47; Chi ² = 8.31, 0 4.96 (P < 0.000 | -6.8 14.1 87 -1.2 11.7 112 -2.35 10.47 74 -2.5 14.38 141 414 10; Chi ² = 3.00, df = 3 (: 3.80 (P = 0.0001) s -3.8 10.8 66 -2.2 9.2 237 -3.8 10.24 214 517 :6; Chi ² = 5.11, df = 2 (: 2.85 (P = 0.004) 931 :7; Chi ² = 8.31, df = 6 (: 4.96 (P < 0.00001) | Mean SD Total Mean -6.8 14.1 87 -1.3 -1.2 11.7 112 2.4 -2.35 10.47 74 -1.14 -2.5 14.38 141 1.8 414 -2.5 14.38 141 1.8 90; Chi ² = 3.00, df = 3 (P = 0.39 3.80 (P = 0.0001) 35 -3.8 10.8 66 -0.5 -2.2 9.2 237 -0.3 -3.8 10.24 214 1.6 517 517 517 516; Chi ² = 5.11, df = 2 (P = 0.004) 931 52.85 (P = 0.004) 931 57; Chi ² = 8.31, df = 6 (P = 0.22) 4.96 (P < 0.00001) | Mean SD Total Mean SD -6.8 14.1 87 -1.3 12.9 -1.2 11.7 112 2.4 13.7 -2.35 10.47 74 -1.14 9.76 -2.5 14.38 141 1.8 12.13 414 .00 Chi² = 3.00, df = 3 (P = 0.39); l² = 0 3.80 (P = 0.0001) S -3.8 10.8 66 -0.5 10.9 -2.2 9.2 237 -0.3 9.2 -3.8 10.24 214 1.6 11.38 517 .6; Chi² = 5.11, df = 2 (P = 0.08); l² = 6 .2.85 (P = 0.004) 931 6; Chi² = 8.31, df = 6 (P = 0.22); l² = 2 .4.96 (P < 0.00001) | Mean SD Total Mean SD Total -6.8 14.1 87 -1.3 12.9 83 -1.2 11.7 112 2.4 13.7 56 -2.35 10.47 74 -1.14 9.76 68 -2.5 14.38 141 1.8 12.13 128 414 335 0; Chi² = 3.00, df = 3 (P = 0.39); l² = 0% :3.80 (P = 0.0001) s -3.8 10.8 66 -0.5 10.9 65 -2.2 9.2 237 -0.3 9.2 237 -3.8 10.24 214 1.6 11.38 107 517 409 517 409 517 409 i6; Chi² = 5.11, df = 2 (P = 0.08); l² = 61% :2.85 (P = 0.004) 931 744 i7; Chi² = 8.31, df = 6 (P = 0.22); l² = 28% 528 538 538 538 | Mean SD Total Mean SD Total Weight -6.8 14.1 87 -1.3 12.9 83 8.9% -1.2 11.7 112 2.4 13.7 56 8.5% -2.35 10.47 74 -1.14 9.76 68 12.3% -2.5 14.38 141 1.8 12.13 128 13.2% 414 335 42.8% 414 335 42.8% 00; Chi ² = 3.00, df = 3 (P = 0.39); l ² = 0% 3.80 (P = 0.0001) 5 -3.8 10.8 66 -0.5 10.9 65 10.3% -2.2 9.2 237 -0.3 9.2 237 29.0% -3.8 10.24 214 1.6 11.38 107 17.9% 517 409 57.2% 517 409 57.2% 66; Chi ² = 5.11, df = 2 (P = 0.08); l ² = 61% 2.85 (P = 0.004) 931 744 100.0% 67; Chi ² = 8.31, df = 6 (P = 0.22); l ² = 28% </td <td>Mean SD Total Mean SD Total Weight IV, Random, 95% Cl -6.8 14.1 87 -1.3 12.9 83 8.9% -5.50 [-9.56, -1.44] -1.2 11.7 112 2.4 13.7 56 8.5% -3.60 [-7.79, 0.59] -2.35 10.47 74 -1.14 9.76 68 12.3% -1.21 [4.54, 2.12] -2.5 14.38 141 1.8 12.13 128 13.2% -4.30 [-7.47, -1.13] 414 335 42.8% -3.50 [-5.30, -1.70] 00 00; Chi² = 3.00, df = 3 (P = 0.39); l² = 0% -3.80 (P = 0.0001) -3.80 (P = 0.0001) s -3.8 10.8 66 -0.5 10.9 65 10.3% -3.30 [-7.02, 0.42] -2.2 9.2 237 -0.3 9.2 237 29.0% -1.90 [-3.56, -0.24] -3.8 10.24 214 1.6 11.38 107 17.9% -5.40 [-7.96, -2.84] 517 409</td> | Mean SD Total Mean SD Total Weight IV, Random, 95% Cl -6.8 14.1 87 -1.3 12.9 83 8.9% -5.50 [-9.56, -1.44] -1.2 11.7 112 2.4 13.7 56 8.5% -3.60 [-7.79, 0.59] -2.35 10.47 74 -1.14 9.76 68 12.3% -1.21 [4.54, 2.12] -2.5 14.38 141 1.8 12.13 128 13.2% -4.30 [-7.47, -1.13] 414 335 42.8% -3.50 [-5.30, -1.70] 00 00; Chi ² = 3.00, df = 3 (P = 0.39); l ² = 0% -3.80 (P = 0.0001) -3.80 (P = 0.0001) s -3.8 10.8 66 -0.5 10.9 65 10.3% -3.30 [-7.02, 0.42] -2.2 9.2 237 -0.3 9.2 237 29.0% -1.90 [-3.56, -0.24] -3.8 10.24 214 1.6 11.38 107 17.9% -5.40 [-7.96, -2.84] 517 409 |

Figure 6. Forest plot of the weighted mean difference in the change of systolic blood pressure from baseline.

| | Experim | | Contr | | | Odds Ratio | Odds Ratio |
|---|---------------------------|-----------|------------|-------------------|--------|---------------------|--|
| Study or Subgroup | Events | Total | Events | Total | Weight | M-H, Random, 95% C | M-H, Random, 95% Cl |
| 6.1.1 Asian subjects | | | | | | | |
| Kyung-Ah Han 2018 | 7 | 74 | 6 | 68 | 17.1% | 1.08 [0.34, 3.39] | _ |
| Linong Ji 2015 | 5 | 223 | 4 | 226 | 12.7% | 1.27 [0.34, 4.80] | |
| Wenying YANG 2016 | 3 | 152 | 6 | 145 | 11.4% | 0.47 [0.11, 1.90] | |
| Subtotal (95% CI) | | 449 | | 439 | 41.2% | 0.90 [0.43, 1.88] | + |
| Total events | 15 | | 16 | | | | |
| Heterogeneity: Tau ² = 0.00; C | hi ² = 1.20, d | if = 2 (P | = 0.55); | ² = 0% | | | |
| Test for overall effect: Z = 0.28 | 8 (P = 0.78) | | | | | | |
| 6.1.2 non-Asian subjects | | | | | | | |
| J. P. H. Wilding 2012 | 1 | 68 | 1 | 66 | 2.9% | 0.97 [0.06, 15.84] | |
| Julio Rosenstock 2018 | 7 | 205 | 8 | 209 | 21.0% | 0.89 [0.32, 2.50] | |
| Julio Rosenstock 2016 | 7 | 237 | 7 | 237 | 19.8% | 1.00 [0.35, 2.90] | _ _ |
| Marina V. Shestakova 2018 | 1 | 38 | 1 | 14 | 2.8% | 0.35 [0.02, 6.03] | · · · · · · · · · · · · · · · · · · · |
| R.R. Henry2012 | 3 | 211 | 4 | 208 | 9.8% | 0.74 [0.16, 3.33] | |
| Schumm-Draeger 2015 | 2 | 99 | 0 | 101 | 2.4% | 5.21 [0.25, 109.80] | |
| Subtotal (95% CI) | | 858 | | 835 | 58.8% | 0.93 [0.50, 1.72] | • |
| Total events | 21 | | 21 | | | | |
| Heterogeneity: Tau ² = 0.00; C | hi ² = 1.81, c | if = 5 (P | = 0.87); | 2 = 0% | | | |
| Test for overall effect: Z = 0.24 | 4 (P = 0.81) | | 8 | | | | |
| Total (95% CI) | | 1307 | | 1274 | 100.0% | 0.92 [0.57, 1.47] | • |
| Total events | 36 | | 37 | | | | |
| Heterogeneity: Tau ² = 0.00; C | hi ² = 3.00, d | if = 8 (P | = 0.93); | ² = 0% | | | |
| Test for overall effect: Z = 0.36 | | | | | | | 0.001 0.1 1 10 100 |
| Test for subgroup differences: | | | (P = 0.96) |), $ ^2 = 0$ | 0% | | Favours [experimental] Favours [control] |

Figure 7. Forest plot of the odds ratio in the incidence of serious adverse events.

There are limitations in this systematic review and meta-analysis. First, the trials investigating non-Asian individuals recruited Asians may cause bias in results, and we discarded trials whose proportion of Asian patients more than 20% (eg., Merker et al. was about 50% [29]) from non-Asian group to minimize this bias. Second, only Chinese and Japanese subjects (5 studies with 1193 subjects) were included since trial from other Asian countries was not found from the databases, which may be diverse from China and Japan population and may influence the outcome of the study, we therefore labeled "Asian patients" as "East Asian patients" in the conclusion to minimize the possible bias. Third, the duration of SGLT-2 inhibitors as an add-on treatment for metformin in this meta-analysis was 12 to 26 weeks, which was short-term and the longterm effect needs further study. Fourth, the statistical heterogeneity (I²) was relatively high for the comparison of the change in HbA1c ($I^2 = 87\%$ and 78%, respectively). The different types of SGLT-2 inhibitors used and the different treatment times may have contributed to the high level of heterogeneity. Because of the high heterogeneity, we used random effect model to conduct a sensitivity analysis excluding the data leading to significant heterogeneity of each group. Moreover, the findings may be restricted to some brands of SGLT-2 inhibitors, because SGLT-2 inhibitors used in East Asian patients were canagliflozin, dapagliflozin and ipragliflozin, while in non-Asian patients were canagliflozin, dapagliflozin, ipragliflozin, empagliflozin and ertugliflozin. Therefore, given the above limitations, our results should be interpreted with caution.

CONCLUSIONS

According to this systematic review and metaanalysis, when SGLT-2 inhibitors are used as an addon treatment for metformin, the decrease in HbA1c and fasting blood glucose in East Asian T2DM patients was significantly greater than in non-Asian patients compared with metformin monotherapy. There was no significant difference in the change in systolic blood pressure and body weight between the two groups.

Moreover, there was no significant difference in severe adverse events between the East Asian and non-Asian T2DM patients when treated with SGLT-2 inhibitors as an add-on treatment for metformin. Therefore, this treatment regimen could be more efficacious in East Asian T2DM patients than in non-Asian T2DM patients in short term without an additional risk of severe adverse events. Additional clinical trials with larger sample sizes and more available data with different races may better clarify these issues.

AUTHOR CONTRIBUTIONS

LL and XZL conceived and designed the systematic review protocol; SSG, XJZ, SJ performed the literature search and study selection; QPZ, YHZ, RZZ performed the data extraction and quality assessment; XZL, JJD performed the statistical analyses; XZL, LL and JJD drafted the manuscript and revised it critically.

All authors gave final approval for submitted manuscript content and agreed to be accountable for all aspects of the work in ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

CONFLICTS OF INTEREST

All authors declare no conflicts of interest.

FUNDING

This work was supported by the National Natural Science Foundation of China, No.81770822.

REFERENCES

- 1. IDF. IDF DIABETES ATLAS 2019. https://www.diabetesatlas.org/en/
- Ogurtsova K, da Rocha Fernandes JD, Huang Y, Linnenkamp U, Guariguata L, Cho NH, Cavan D, Shaw JE, Makaroff LE. IDF Diabetes Atlas: Global estimates for the prevalence of diabetes for 2015 and 2040. Diabetes Res Clin Pract. 2017; 128:40–50. <u>https://doi.org/10.1016/j.diabres.2017.03.024</u> PMID:<u>28437734</u>
- Jia W, Weng J, Zhu D, Ji L, Lu J, Zhou Z, Zou D, Guo L, Ji Q, Chen L, Chen L, Dou J, Guo X, et al, and Chinese Diabetes Society. Standards of medical care for type 2 diabetes in China 2019. Diabetes Metab Res Rev. 2019; 35:e3158. <u>https://doi.org/10.1002/dmrr.3158</u> PMID:30908791
- Buse JB, Wexler DJ, Tsapas A, Rossing P, Mingrone G, Mathieu C, D'Alessio DA, Davies MJ. 2019 Update to: Management of Hyperglycemia in Type 2 Diabetes, 2018. A Consensus Report by the American Diabetes Association (ADA) and the European Association for the Study of Diabetes (EASD). Diabetes Care. 2020; 43:487–93.

https://doi.org/10.2337/dci19-0066 PMID:<u>31857443</u>

 Holman RR, Paul SK, Bethel MA, Matthews DR, Neil HA. 10-year follow-up of intensive glucose control in type 2 diabetes. N Engl J Med. 2008; 359:1577–89. https://doi.org/10.1056/NEJMoa0806470 PMID:<u>18784090</u>

- 6. Zhong X, Lai D, Ye Y, Yang X, Yu B, Huang Y. Efficacy and safety of empagliflozin as add-on to metformin for type 2 diabetes: a systematic review and metaanalysis. Eur J Clin Pharmacol. 2016; 72:655–63. <u>https://doi.org/10.1007/s00228-016-2010-8</u> PMID:<u>26832915</u>
- Inzucchi SE, Bergenstal RM, Buse JB, Diamant M, Ferrannini E, Nauck M, Peters AL, Tsapas A, Wender R, Matthews DR. Management of hyperglycaemia in type 2 diabetes, 2015: a patient-centred approach. Update to a position statement of the American Diabetes Association and the European Association for the Study of Diabetes. Diabetologia. 2015; 58:429–42.

https://doi.org/10.1007/s00125-014-3460-0 PMID:<u>25583541</u>

- American Diabetes Association. 9. Pharmacologic Approaches to Glycemic Treatment: *Standards of Medical Care in Diabetes-2019*. Diabetes Care. 2019 (Suppl 1); 42:S90–102. <u>https://doi.org/10.2337/dc19-S009</u> PMID:30559235
- 9. Khoo CM, Deerochanawong C, Chan SP, Matawaran B, Sheu WH, Chan J, Mithal A, Luk A, Suastika K, Yoon KH, Ji L, Man NH, Pollock C. Use of sodium-glucose co-transporter-2 inhibitors in Asian patients with type 2 diabetes and kidney disease: An Asian perspective and expert recommendations. Diabetes Obes Metab. 2021; 23:299–317. https://doi.org/10.1111/dom.14251 PMID:33155749
- Malaysian Clinical Practice Guidelines. Management of type 2 diabetes mellitus. <u>http://www.acadmed.</u> <u>org.my/view file.cfm?fileid=763</u>. Updated 2015. Accessed April 2, 2019.
- Deerochanawong C, Chan SP, Matawaran BJ, Sheu WH, Chan J, Man NH, Suastika K, Khoo CM, Yoon KH, Luk A, Mithal A, Linong J. Use of sodium-glucose cotransporter-2 inhibitors in patients with type 2 diabetes mellitus and multiple cardiovascular risk factors: An Asian perspective and expert recommendations. Diabetes Obes Metab. 2019; 21:2354–67. <u>https://doi.org/10.1111/dom.13819</u> PMID:<u>31264765</u>
- Gallo LA, Wright EM, Vallon V. Probing SGLT2 as a therapeutic target for diabetes: basic physiology and consequences. Diab Vasc Dis Res. 2015; 12:78–89. <u>https://doi.org/10.1177/1479164114561992</u> PMID:<u>25616707</u>
- 13. Ji L, Han P, Liu Y, Yang G, Dieu Van NK, Vijapurkar U, Qiu R, Meininger G. Canagliflozin in Asian patients

with type 2 diabetes on metformin alone or metformin in combination with sulphonylurea. Diabetes Obes Metab. 2015; 17:23–31. <u>https://doi.org/10.1111/dom.12385</u> PMID:25175734

- 14. Han KA, Chon S, Chung CH, Lim S, Lee KW, Baik S, Jung CH, Kim DS, Park KS, Yoon KH, Lee IK, Cha BS, Sakatani T, et al. Efficacy and safety of ipragliflozin as an add-on therapy to sitagliptin and metformin in Korean patients with inadequately controlled type 2 diabetes mellitus: A randomized controlled trial. Diabetes Obes Metab. 2018; 20:2408–15. <u>https://doi.org/10.1111/dom.13394</u> PMID:29862619
- Kashiwagi A, Kazuta K, Goto K, Yoshida S, Ueyama E, Utsuno A. Ipragliflozin in combination with metformin for the treatment of Japanese patients with type 2 diabetes: ILLUMINATE, a randomized, double-blind, placebo-controlled study. Diabetes Obes Metab. 2015; 17:304–08. <u>https://doi.org/10.1111/dom.12331</u> PMID:24919820
- Lu CH, Min KW, Chuang LM, Kokubo S, Yoshida S, Cha BS. Efficacy, safety, and tolerability of ipragliflozin in Asian patients with type 2 diabetes mellitus and inadequate glycemic control with metformin: Results of a phase 3 randomized, placebo-controlled, doubleblind, multicenter trial. J Diabetes Investig. 2016; 7:366–73. <u>https://doi.org/10.1111/jdi.12422</u> PMID:27330723
- Yang W, Han P, Min KW, Wang B, Mansfield T, T'Joen C, Iqbal N, Johnsson E, Ptaszynska A. Efficacy and safety of dapagliflozin in Asian patients with type 2 diabetes after metformin failure: A randomized controlled trial. J Diabetes. 2016; 8:796–808.

https://doi.org/10.1111/1753-0407.12357 PMID:<u>26589253</u>

- Ma RC, Chan JC. Type 2 diabetes in East Asians: similarities and differences with populations in Europe and the United States. Ann N Y Acad Sci. 2013; 1281:64–91. <u>https://doi.org/10.1111/nyas.12098</u> PMID:<u>23551121</u>
- Stroup DF, Berlin JA, Morton SC, Olkin I, Williamson GD, Rennie D, Moher D, Becker BJ, Sipe TA, Thacker SB. Meta-analysis of observational studies in epidemiology: a proposal for reporting. Meta-analysis Of Observational Studies in Epidemiology (MOOSE) group. JAMA. 2000; 283:2008–12. https://doi.org/10.1001/jama.283.15.2008

PMID:10789670

- 20. Moher D, Shamseer L, Clarke M, Ghersi D, Liberati A, Petticrew M, Shekelle P, Stewart LA, and PRISMA-P Group. Preferred reporting items for systematic review and meta-analysis protocols (PRISMA-P) 2015 statement. Syst Rev. 2015; 4:1. <u>https://doi.org/10.1186/2046-4053-4-1</u> PMID:<u>25554246</u>
- 21. Li J, Gong Y, Li C, Lu Y, Liu Y, Shao Y. Long-term efficacy and safety of sodium-glucose cotransporter-2 inhibitors as add-on to metformin treatment in the management of type 2 diabetes mellitus: A metaanalysis. Medicine (Baltimore). 2017; 96:e7201. https://doi.org/10.1097/MD.000000000007201 PMID:<u>28682870</u>
- 22. Cumpston M, Li T, Page MJ, Chandler J, Welch VA, Higgins JP, Thomas J. Updated guidance for trusted systematic reviews: a new edition of the Cochrane Handbook for Systematic Reviews of Interventions. Cochrane Database Syst Rev. 2019; 10:ED000142. <u>https://doi.org/10.1002/14651858.ED000142</u> PMID:<u>31643080</u>
- 23. DerSimonian R, Laird N. Meta-analysis in clinical trials revisited. Contemp Clin Trials. 2015; 45:139–45.
 <u>https://doi.org/10.1016/j.cct.2015.09.002</u>
 PMID:<u>26343745</u>
- 24. Xu HC, Pang YC, Chen JW, Cao JY, Sheng Z, Yuan JH, Wang R, Zhang CS, Wang LX, Dong J. Systematic Review and Meta-analysis of the Change in Ghrelin Levels After Roux-en-Y Gastric Bypass. Obes Surg. 2019; 29:1343–51. <u>https://doi.org/10.1007/s11695-018-03686-3</u> PMID:30684171
- 25. Higgins JP, Thompson SG, Deeks JJ, Altman DG. Measuring inconsistency in meta-analyses. BMJ. 2003; 327:557–60. <u>https://doi.org/10.1136/bmj.327.7414.557</u> PMID:<u>12958120</u>
- 26. Bailey CJ, Gross JL, Hennicken D, Iqbal N, Mansfield TA, List JF. Dapagliflozin add-on to metformin in type 2 diabetes inadequately controlled with metformin: a randomized, double-blind, placebo-controlled 102week trial. BMC Med. 2013; 11:43. <u>https://doi.org/10.1186/1741-7015-11-43</u> PMID:<u>23425012</u>
- Bolinder J, Ljunggren Ö, Johansson L, Wilding J, Langkilde AM, Sjöström CD, Sugg J, Parikh S. Dapagliflozin maintains glycaemic control while reducing weight and body fat mass over 2 years in patients with type 2 diabetes mellitus inadequately controlled on metformin. Diabetes Obes Metab. 2014; 16:159–69.

https://doi.org/10.1111/dom.12189 PMID:23906445

 Ferrannini E, Berk A, Hantel S, Pinnetti S, Hach T, Woerle HJ, Broedl UC. Long-term safety and efficacy of empagliflozin, sitagliptin, and metformin: an active-controlled, parallel-group, randomized, 78-week open-label extension study in patients with type 2 diabetes. Diabetes Care. 2013; 36:4015–21.

https://doi.org/10.2337/dc13-0663 PMID:24186878

- 29. Merker L, Häring HU, Christiansen AV, Roux F, Salsali A, Kim G, Meinicke T, Woerle HJ, Broedl UC, and EMPA-REG EXTEND MET investigators. Empagliflozin as add-on to metformin in people with Type 2 diabetes. Diabet Med. 2015; 32:1555–67. <u>https://doi.org/10.1111/dme.12814</u> PMID:26031566
- Amin NB, Wang X, Jain SM, Lee DS, Nucci G, Rusnak JM. Dose-ranging efficacy and safety study of ertugliflozin, a sodium-glucose co-transporter 2 inhibitor, in patients with type 2 diabetes on a background of metformin. Diabetes Obes Metab. 2015; 17:591–98. <u>https://doi.org/10.1111/dom.12460</u> PMID:25754396
- Schumm-Draeger PM, Burgess L, Korányi L, Hruba V, Hamer-Maansson JE, de Bruin TW. Twice-daily dapagliflozin co-administered with metformin in type 2 diabetes: a 16-week randomized, placebocontrolled clinical trial. Diabetes Obes Metab. 2015; 17:42–51.

https://doi.org/10.1111/dom.12387 PMID:25200570

PMID:22413962

- 32. Hadjadj S, Rosenstock J, Meinicke T, Woerle HJ, Broedl UC. Initial Combination of Empagliflozin and Metformin in Patients With Type 2 Diabetes. Diabetes Care. 2016; 39:1718–28. <u>https://doi.org/10.2337/dc16-0522</u> PMID:27493136
- Henry RR, Murray AV, Marmolejo MH, Hennicken D, Ptaszynska A, List JF. Dapagliflozin, metformin XR, or both: initial pharmacotherapy for type 2 diabetes, a randomised controlled trial. Int J Clin Pract. 2012; 66:446–56. <u>https://doi.org/10.1111/j.1742-1241.2012.02911.x</u>
- 34. Ross S, Thamer C, Cescutti J, Meinicke T, Woerle HJ, Broedl UC. Efficacy and safety of empagliflozin twice daily versus once daily in patients with type 2 diabetes inadequately controlled on metformin: a 16-week, randomized, placebo-controlled trial. Diabetes Obes Metab. 2015; 17:699–702.

https://doi.org/10.1111/dom.12469 PMID:25827441

 Rosenstock J, Frias J, Páll D, Charbonnel B, Pascu R, Saur D, Darekar A, Huyck S, Shi H, Lauring B, Terra SG. Effect of ertugliflozin on glucose control, body weight, blood pressure and bone density in type 2 diabetes mellitus inadequately controlled on metformin monotherapy (VERTIS MET). Diabetes Obes Metab. 2018; 20:520–29.

https://doi.org/10.1111/dom.13103 PMID:28857451

36. Shestakova MV, Wilding JPH, Wilpshaar W, Tretter R, Orlova VL, Verbovoy AF. A phase 3 randomized placebo-controlled trial to assess the efficacy and safety of ipragliflozin as an add-on therapy to metformin in Russian patients with inadequately controlled type 2 diabetes mellitus. Diabetes Res Clin Pract. 2018; 146:240–50.

https://doi.org/10.1016/j.diabres.2018.10.018 PMID:<u>30391333</u>

- 37. Wilding JP, Ferrannini E, Fonseca VA, Wilpshaar W, Dhanjal P, Houzer A. Efficacy and safety of ipragliflozin in patients with type 2 diabetes inadequately controlled on metformin: a dose-finding study. Diabetes Obes Metab. 2013; 15:403–09. <u>https://doi.org/10.1111/dom.12038</u> PMID:<u>23163880</u>
- Chan JC, Malik V, Jia W, Kadowaki T, Yajnik CS, Yoon KH, Hu FB. Diabetes in Asia: epidemiology, risk factors, and pathophysiology. JAMA. 2009; 301:2129–40. <u>https://doi.org/10.1001/jama.2009.726</u>

PMID:<u>19470990</u>

 Komoroski B, Vachharajani N, Boulton D, Kornhauser D, Geraldes M, Li L, Pfister M. Dapagliflozin, a novel SGLT2 inhibitor, induces dose-dependent glucosuria in healthy subjects. Clin Pharmacol Ther. 2009; 85:520–26.

https://doi.org/10.1038/clpt.2008.251 PMID:<u>19129748</u>

- 40. Yang L, Li H, Li H, Bui A, Chang M, Liu X, Kasichayanula S, Griffen SC, Lacreta FP, Boulton DW. Pharmacokinetic and pharmacodynamic properties of single- and multiple-dose of dapagliflozin, a selective inhibitor of SGLT2, in healthy Chinese subjects. Clin Ther. 2013; 35:1211–22.e2. https://doi.org/10.1016/j.clinthera.2013.06.017 PMID:23910664
- Elsamadony A, Yates KF, Sweat V, Yau PL, Mangone A, Joseph A, Fierman A, Convit A. Asian Adolescents with Excess Weight are at Higher Risk for Insulin Resistance than Non-Asian Peers. Obesity (Silver Spring). 2017; 25:1974–79.

https://doi.org/10.1002/oby.22003 PMID:28941205

- Knight TM, Smith Z, Whittles A, Sahota P, Lockton JA, Hogg G, Bedford A, Toop M, Kernohan EE, Baker MR. Insulin resistance, diabetes, and risk markers for ischaemic heart disease in Asian men and non-Asian in Bradford. Br Heart J. 1992; 67:343–50. <u>https://doi.org/10.1136/hrt.67.5.343</u> PMID:<u>1389712</u>
- Xu L, Nagata N, Nagashimada M, Zhuge F, Ni Y, Chen G, Mayoux E, Kaneko S, Ota T. SGLT2 Inhibition by Empagliflozin Promotes Fat Utilization and Browning and Attenuates Inflammation and Insulin Resistance by Polarizing M2 Macrophages in Diet-induced Obese Mice. EBioMedicine. 2017; 20:137–49. <u>https://doi.org/10.1016/j.ebiom.2017.05.028</u> PMID:<u>28579299</u>
- 44. Xu L, Nagata N, Chen G, Nagashimada M, Zhuge F, Ni Y, Sakai Y, Kaneko S, Ota T. Empagliflozin reverses obesity and insulin resistance through fat browning and alternative macrophage activation in mice fed a high-fat diet. BMJ Open Diabetes Res Care. 2019; 7:e000783.

https://doi.org/10.1136/bmjdrc-2019-000783 PMID:<u>31749970</u>

45. Lim LL, Tan AT, Moses K, Rajadhyaksha V, Chan SP. Place of sodium-glucose cotransporter-2 inhibitors in East Asian subjects with type 2 diabetes mellitus: Insights into the management of Asian phenotype. J Diabetes Complications. 2017; 31:494–503.

https://doi.org/10.1016/j.jdiacomp.2016.10.008 PMID:27866701

- 46. Lee PC, Ganguly S, Goh SY. Weight loss associated with sodium-glucose cotransporter-2 inhibition: a review of evidence and underlying mechanisms. Obes Rev. 2018; 19:1630–41. https://doi.org/10.1111/obr.12755 PMID:30253050
- 47. Scheen AJ. Pharmacodynamics, efficacy and safety of sodium-glucose co-transporter type 2 (SGLT2) inhibitors for the treatment of type 2 diabetes mellitus. Drugs. 2015; 75:33–59. https://doi.org/10.1007/s40265-014-0337-y
 PMID:25488697
- 48. Lundkvist P, Pereira MJ, Katsogiannos P, Sjöström CD, Johnsson E, Eriksson JW. Dapagliflozin once daily plus exenatide once weekly in obese adults without diabetes: Sustained reductions in body weight, glycaemia and blood pressure over 1 year. Diabetes Obes Metab. 2017; 19:1276–88. <u>https://doi.org/10.1111/dom.12954</u>

PMID:<u>28345814</u>

- 49. Maliha G, Townsend RR. SGLT2 inhibitors: their potential reduction in blood pressure. J Am Soc Hypertens. 2015; 9:48–53. https://doi.org/10.1016/j.jash.2014.11.001 PMID:25537461
- 50. Chen Z, Li G. Sodium-Glucose Co-Transporter 2 Inhibitors Compared with Sulfonylureas in Patients with Type 2 Diabetes Inadequately Controlled on Metformin: A Meta-Analysis of Randomized Controlled Trials. Clin Drug Investig. 2019; 39:521–31.

https://doi.org/10.1007/s40261-019-00781-w PMID:<u>31041606</u>

51. Mishriky BM, Tanenberg RJ, Sewell KA, Cummings DM. Comparing SGLT-2 inhibitors to DPP-4 inhibitors as an add-on therapy to metformin in patients with type 2 diabetes: A systematic review and metaanalysis. Diabetes Metab. 2018; 44:112–20. <u>https://doi.org/10.1016/j.diabet.2018.01.017</u> PMID:<u>29477373</u>

SUPPLEMENTARY MATERIALS

Supplementary Figures

| Patient or population: patients with T: Settings: Intervention: SGLT2-i plus MET Comparison: MET monotherapy | 2DM | | | | | |
|---|--|---|--------------------------------|------------------------------------|---------------------------------------|----------|
| Outcomes | Illustrative com Assumed risk MET monotherapy | Corresponding risk | Relative effect (95% CI) | No of Participants (studies) | Quality of the evidence (GRADE) | Comments |
| HbA1c (sensitivity analysis) HbA1c Follow-up: 12-26 weeks | | The mean hba1c (sensitivity analysis) in the intervention groups was 0.47 lower (0.54 to 0.41 lower) | | 2709 (10 studies) | ⊕⊕⊕⊝ moderate ¹ | |
| HbA1c (sensitivity analysis) - Asian subjects HbA1c Follow-up: 18-24 weeks | | The mean hba1c (sensitivity analysis) - asian subjects in the intervention groups was 0.61 lower (0.71 to 0.5 lower) | | 1025 (4 studies) | ⊕⊕⊕⊝ moderate ¹ | |
| HbA1c (sensitivity analysis) - non- Asian subjects HbA1c Follow-up: 12-26 weeks | | The mean hba1c (sensitivity analysis) - non-asian subjects in the intervention groups was 0.4 lower (0.48 to 0.31 lower) | | 1684 (6 studies) | ⊕⊕⊕⊝ moderate ¹ | |

*The basis for the assumed risk (e.g. the median control group risk across studies) is provided in footnotes. The corresponding risk (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% Cl).

CI: Confidence interval;

GRADE Working Group grades of evidence

High quality: Further research is very unlikely to change our confidence in the estimate of effect.

Moderate quality: Further research is likely to have an important impact on our confidence in the estimate of effect and may change the estimate.

Low quality: Further research is very likely to have an important impact on our confidence in the estimate of effect and is likely to change the estimate.

Very low quality: We are very uncertain about the estimate.

¹ Funnel diagrams of HbA1c

Supplementary Figure 1. GRADE approach to assess the overall confidence for HbA1c.

| SGLT2-i plus MET comp | ared to MET monothe | erapy for T2DM | | | |
|--|---|---|-----------------------------|---------------------------------|---|
| Patient or population: pat Settings: Intervention: SGLT2-i plus Comparison: MET monothe | MET | | | | |
| Outcomes | Illustrative compara Assumed risk MET monotherapy | tive risks* (95% Cl) Corresponding risk SGLT2-i plus MET | Relative effect (95% CI) | No of Participants (studies) | Quality of the evidence Comments (GRADE) |
| FPG FPG Follow-up: 12-26 weeks | | The mean fpg in the intervention groups was 1.2 lower (1.34 to 1.07 lower) | | 2464 (10 studies) | ⊕⊕⊕⊕ high |
| FPG - Asian subjects FPG Follow-up: 18-24 weeks | | The mean fpg - asian subjects in the intervention groups was 1.51 lower (1.75 to 1.28 lower) | | 767 (4 studies) | ⊕⊕⊕⊕ high |
| FPG - non-Asian subjects FPG Follow-up: 12-26 weeks | | The mean fpg - non-asian subjects in the intervention groups was 1.04 lower (1.21 to 0.88 lower) | | 1697 (6 studies) | ⊕⊕⊕ high |

risk in the comparison group and the relative effect of the intervention (and its 95% Cl).

CI: Confidence interval;

GRADE Working Group grades of evidence

High quality: Further research is very unlikely to change our confidence in the estimate of effect. Moderate quality: Further research is likely to have an important impact on our confidence in the estimate of effect and may change the estimate.

Low quality: Further research is very likely to have an important impact on our confidence in the estimate of effect and is likely to change the estimate.

Very low quality: We are very uncertain about the estimate.

Supplementary Figure 2. GRADE approach to assess the overall confidence for FPG.

SGLT2-i plus MET compared to MET monotherapy for T2DM

Patient or population: patients with T2DM

Settings: Intervention: SGLT2-i plus MET

Comparison: MET monotherapy

| Outcomes | Illustrative compara | ative risks* (95% CI) | | lo of Participants | The company of the second second | Comments |
|--|----------------------|---|------------|--------------------|----------------------------------|----------|
| | Assumed risk | Corresponding risk | (95% CI) (| (studies) | evidence (GRADE) | |
| | MET monotherapy | SGLT2-i plus MET | | | | |
| Body Weight Body Weight Follow-up: 12-26 weeks | | The mean body weight in the intervention groups was 1.69 lower (1.89 to 1.48 lower) | | 2310 9 studies) | ⊕⊕⊕⊕ high | |
| Body Weight - Asian subjects Body Weight Follow-up: 18-24 weeks | | The mean body weight - asian subjects in the intervention groups was 1.69 lower (1.98 to 1.41 lower) | | 767 (4 studies) | ⊕⊕⊕⊕ high | |
| Body Weight - non-Asian subjects Body Weight Follow-up: 12-26 weeks | | The mean body weight - non-asian subjects in the intervention groups was 1.68 lower (1.97 to 1.39 lower) | | 1543 5 studies) | ⊕⊕⊕⊕ high | |

The basis for the assumed risk (e.g. the median control group risk across studies) is provided in footnotes. The corresponding risk (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% Cl).

CI: Confidence interval;

GRADE Working Group grades of evidence

High quality: Further research is very unlikely to change our confidence in the estimate of effect.

Moderate quality: Further research is likely to have an important impact on our confidence in the estimate of effect and may change the estimate.

Low quality: Further research is very likely to have an important impact on our confidence in the estimate of effect and is likely to change the estimate.

Very low quality: We are very uncertain about the estimate.

Supplementary Figure 3. GRADE approach to assess the overall confidence for Body Weight.

| SGLT2-i plus MET com | pared to MET monotherapy for T2DM | 1 |
|----------------------|-----------------------------------|---|
|----------------------|-----------------------------------|---|

Patient or population: patients with T2DM

Settings: Intervention: SGLT2-i plus MET

Comparison: MET monotherapy

| Outcomes | Illustrative compara Assumed risk MET monotherapy | tive risks* (95% CI) Corresponding risk SGLT2-i plus MET | Relative effect (95% CI) | No of Participants (studies) | Quality of the evidence Comments (GRADE) |
|---|---|---|-----------------------------|---------------------------------|---|
| SBP SBP Follow-up: 12-26 weeks | | The mean sbp in the intervention groups was 3.16 lower (4.21 to 2.1 lower) | | 1675 (7 studies) | ⊕⊕⊕⊕ high |
| SBP - Asian subjects SBP Follow-up: 18-24 weeks | | The mean sbp - asian subjects in the intervention groups was 3.5 lower (5.3 to 1.7 lower) | | 749 (4 studies) | ⊕⊕⊕⊕ high |
| SBP - non-Asian subjects SBP Follow-up: 12-26 weeks | | The mean sbp - non-asian subjects in the intervention groups was 2.98 lower (4.28 to 1.68 lower) | | 926 (3 studies) | ⊕⊕⊕⊕ high |

The basis for the assumed risk (e.g. the median control group risk across studies) is provided in footnotes. The corresponding risk (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: Confidence interval;

GRADE Working Group grades of evidence

High quality: Further research is very unlikely to change our confidence in the estimate of effect.

Moderate quality: Further research is likely to have an important impact on our confidence in the estimate of effect and may change the estimate.

Low quality: Further research is very likely to have an important impact on our confidence in the estimate of effect and is likely to change the estimate. Very low quality: We are very uncertain about the estimate.

Supplementary Figure 4. GRADE approach to assess the overall confidence for SBP.

SGLT2-i plus MET compared to MET monotherapy for T2DM

Patient or population: patients with T2DM

Settings: Intervention: SGLT2-i plus MET

Comparison: MET monotherany

| Outcomes | Illustrative comparativ Assumed risk MET monotherapy | ve risks* (95% Cl) Corresponding risk SGLT2-i plus MET | Relative effect (95% CI) | No of Participants (studies) | Quality of the evidence (GRADE) | Comments |
|---|--|--|-----------------------------|---------------------------------|------------------------------------|----------|
| Serious Adverse events | Study population | | OR 0.93 | 2581 | ⊕⊕⊕⊕ | |
| SAEs Follow-up: 12-26 weeks | 29 per 1000 | 27 per 1000 (17 to 42) | (0.59 to 1.48) | (9 studies) | high | |
| | Moderate | | | | | |
| | 30 per 1000 | 28 per 1000 (18 to 44) | | | | |
| Serious Adverse events - Asian subjects | Study population | | OR 0.89 (0.43 to 1.84) | 888 | ⊕⊕⊕⊕ high | |
| SAEs Follow-up: 18-24 weeks | 36 per 1000 | 33 per 1000 (16 to 65) | | (3 studies) | | |
| | Moderate | | | | | |
| | 41 per 1000 | 37 per 1000 (18 to 73) | | | | |
| Serious Adverse events - non-Asian subjects | Study population | | OR 0.96 | 1693 | ⊕⊕⊕⊕ high | |
| SAEs Follow-up: 12-26 weeks | 25 per 1000 | 24 per 1000 (13 to 43) | (0.52 to 1.76) | (6 studies) | | |
| | Moderate | | | | | |
| | 24 per 1000 | 23 per 1000 (13 to 41) | | | | |

sis for the assumed risk (e.g. the median control group risk across s provided in footnotes. The corresponding risk (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% Cl).

CI: Confidence interval; OR: Odds ratio;

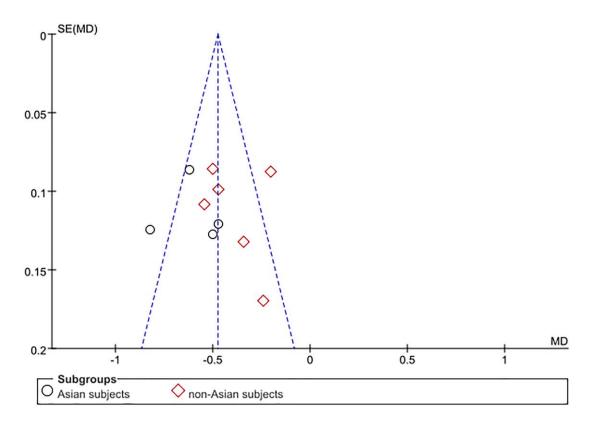
GRADE Working Group grades of evidence

High quality: Further research is very unlikely to change our confidence in the estimate of effect.

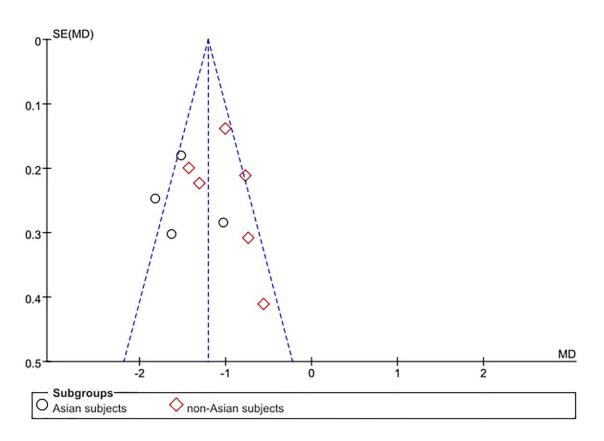
Moderate quality: Further research is likely to have an important impact on our confidence in the estimate of effect and may change the estimate.

Low quality: Further research is very likely to have an important impact on our confidence in the estimate of effect and is likely to change the estimate. Very low quality: We are very uncertain about the estimate.

Supplementary Figure 5. GRADE approach to assess the overall confidence for SAEs.



Supplementary Figure 6. Funnel diagrams of HbA1c.



Supplementary Figure 7. Funnel diagrams of FPG.

Supplementary Tables

Please browse Full Text version to see the data of Supplementary Tables 1 and 2.

Supplementary Table 1. Characteristics of the randomized controlled trials included in the Asian subgroup.

Supplementary Table 2. Characteristics of the randomized controlled trials included in the non-Asian subgroup.