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Surgical Decision Making in Brain Hemorrhage

New Analysis of the STICH, STICH II, and STITCH(Trauma) Randomized Trials

Barbara A. Gregson, PhD; Patrick Mitchell, FRCS; A. David Mendelow, FRCS(SN)

Background and Purpose—The STICH (Surgical Trial in Lobar Intracerebral Haemorrhage) I and II trials randomized patients with spontaneous intracerebral hemorrhage (ICH) to early surgery or initial conservative treatment. Both were nonsignificant; possibly because surgery has minimal effect on recovery, or because surgery benefits some and harms others. We introduce a new nonparametric method of analysis. The method is then applied to data from a third trial, STITCH(Trauma) (Surgical Trial in Traumatic Intracerebral Haemorrhage), which addressed a similar surgical question in head-injured patients.

Methods—Data from 1541 patients from the STICH trials were analyzed using (1) standard meta-analysis of prognosis-based dichotomized outcome and prespecified standard subgroups of Glasgow Coma Scale (GCS): 3–8, 9–12, and 13–15; (2) new nonparametric regression of ranked Extended Glasgow Outcome Scale against ranked GCS and ranked volume; and (3) analysis (1) repeated using categories identified by analysis (2).

Results—Standard meta-analysis showed more favorable outcomes, although nonsignificant, with surgery if presenting GCS was 9–12 (spontaneous ICH odds ratio, 0.70 [95% CI, 0.48–1.03; $P=0.07$]; traumatic odds ratio, 0.48 [95% CI, 0.18–1.26; $P=0.14$]). Ranked analysis showed a similar pattern of results for both spontaneous and traumatic ICH. Surgery was harmful for small lesions with increasing benefit for larger volumes. With GCS, surgery had little effect at either ends of the spectrum but suggested a beneficial effect in the range 10 to 13 (identified graphically). Repeating the meta-analysis with this categorization showed significant benefit for surgery (spontaneous odds ratio, 0.71 [95% CI, 0.51–1.00; $P=0.05$]; traumatic odds ratio, 0.16 [95% CI, 0.05–0.51; $P=0.002$]).

Conclusions—The nonsignificant results observed in the STICH trials are because of mixing patients who benefit from surgery with those who are harmed. Patients with a GCS 10–13 or a large ICH are likely to benefit from surgery. Our analysis showed a similar effect on traumatic ICH/contusion data and promises to be a valuable tool.

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Key Words: cerebral hemorrhage ■ conservative treatment ■ decision making ■ Glasgow Outcome Scale ■ surgery

Randomized controlled trials are the prevailing paradigms of clinical research, and surgical evacuation of intracerebral hematomas (ICHs) has been the subject of 16.^{1–16} The more powerful trials and meta-analyses^{9,11,17,18} have repeatedly failed to displace the null hypothesis though have often shown nonsignificant benefits. That notwithstanding, there is an ongoing interest in the question with further trials ongoing or planned (MISTIE III [Minimally Invasive Surgery and rtPA for Intracerebral Hemorrhage Evacuation], CLEAR IV [Clot Lysis: Evaluating Accelerated Resolution of Intraventricular Hemorrhage], and STITCH III [International Surgical Trial in IntraCerebral Haemorrhage]).

The null results of trials and meta-analyses led to the hypothesis that the effect of surgery, if any, is too small to detect

with current trial data. The amount of randomized data now in hand has limited the size of such an overall benefit to the extent that, were it there, it would be of doubtful clinical significance. These analyses used an assumption common to much clinical research: there is a uniform response to a treatment for the population under study. This assumption is particularly weak with surgical treatments because in addition to the patient-specific variables that affect medical treatments, there is also variation between surgeons. Mechanistic lines of argument have led to a different theory, one commonly held by surgeons and one that predates all trial data. This theory is that response to surgery is not uniform from one case of ICH to another but varies with clot volume and clinical condition.

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Correspondence to Barbara A. Gregson, PhD, Neurosurgical Trials Group, Institute of Neuroscience, Newcastle University, Wolfson Research Centre, Campus for Ageing and Vitality, Westgate Rd, Newcastle upon Tyne, NE4 5PL. Email barbara.gregson@ncl.ac.uk

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In this article, we present a combined analysis of trial data specifically aimed at testing the hypothesis: Is the neutrality of trial data a consequence of little effect or a consequence of mixing cases who benefit from surgery with those that are harmed?

Two trials used similar methodology to investigate whether patients had a better outcome with a policy of early surgery to evacuate spontaneous ICH compared with initial conservative treatment. The STICH (Surgical Trial in Lobar Intracerebral Haemorrhage) included patients with supratentorial ICH of at least 2 cm in diameter and with a Glasgow Coma Scale (GCS) of at least 5.⁹ In STICH II (Surgical Trial in Lobar Intracerebral Haemorrhage), these criteria were refined to include only patients with a lobar hemorrhage within 1 cm of the cortex surface of between 10 and 100 mL and where the Motor GCS was 5 or 6 and the Eye score of the GCS 2 or more.¹¹

Patients were excluded from both if there was an icthorrhagic lesion (aneurysm or arteriovenous malformation) underlying the hemorrhage, or if they had severe preexisting comorbidity that would mean that they could not achieve a Good Recovery on the Glasgow Outcome Scale (GOS) at 6 months even if they made a full recovery from the hemorrhage. Neither trial displaced the null hypothesis but both showed a nonsignificant benefit with surgical evacuation.

We used the combined STICH and STICH II data sets to investigate the relationship between outcome and the effect modification of variables that may influence the response to surgery following spontaneous ICH: neurological condition quantified by GCS and clot volume. We undertook standard meta-analysis with standard prespecified data grouping and supplemented this with a new nonparametric ranked analysis that displays the whole data set and avoids the problems of comparing groups with arbitrarily selective boundary criteria.

In view of the novelty of our analysis method, we also applied it to another data set to fully test its utility. The STITCH(Trauma) trial is the first and only trial on surgical evacuation of traumatic brain contusions (traumatic ICH) of 10 mL or more.¹⁰ It used a similar methodology to the STICH trials with no restriction in GCS. It was curtailed early because of research funding restrictions. One hundred seventy out of a planned 840 cases were recruited. It also had a nonsignificant improvement in primary outcome with surgery but it also showed a large, statistically significant reduction in death rate in the surgical group.

Methods

The data used in this analysis will be submitted to the VISTA database in the future but is available from the authors ([online-only Data Supplement](#)).

The first STICH trial randomized 1033 patients; STICH II randomized 601; outcome was not available for 69 patients from STICH, 18 from STICH II. A further 6 patients are lost to analyses involving the extended GOS (GOSE) where there was insufficient information to code whether the outcome was upper or lower severe disability (1 case), upper or lower moderate disability (4 cases), or upper or lower good recovery (1 case). This analysis, therefore, relies on 1541 patients who were randomized to early surgery or initial conservative treatment.

STITCH(Trauma) randomized 170 patients and outcome was available for 167.

Meta-Analysis

Meta-analysis was undertaken using Review Manager (5.3). A dichotomized prognosis-based outcome was used to improve the

sensitivity of the trial.¹⁹ The prognostic variables were age, GCS at randomization and volume of hematoma. For patients with a good prognosis, which was defined as a value of 27.692 or greater on the prognostic algorithm ($10 \times \text{GCS} - \text{age} - [0.64 \times \text{volume}]$), a favorable outcome was defined as good recovery or moderate disability on the GOSE and unfavorable outcome as severe disability, vegetative, or dead on the GOSE. For patients with a poor prognosis, a value below 27.692 on the algorithm, favorable outcome was defined as good recovery, moderate disability, or upper severe disability and unfavorable outcome as lower severe disability, vegetative, or death. Thus, patients with a poor prognosis who became housebound but could be left alone for at least 8 hours were regarded as having a favorable outcome. Meta-analyses are reported for all spontaneous and traumatic ICH patients separately and a subgroup analysis of GCS using the standard prespecified bandings: GCS 3–8, 9–12, and 13–15.

Ranked Analysis

Each individual was allocated ranks on the GOSE and GCS scales and on lesion volume so that for each scale their rank was the number of subjects with equal or lower values to themselves, including themselves. Then nonparametric regression of ranked GOSE versus another ranked variable was used. Lowess cubic kernel curves were fitted and span selected by inspection using Matlab R2015a. The simplest nonparametric regression method is the moving average. This involves deciding on a proportion of the data set to use, say for example 20%, and then starting at the left-hand side of the x axis and plotting the average of the left most 20% of data points. Then moving one data point right, averaging this 20% and so on across to the right hand most 20%. The proportion of data points used, 20% in this case, is the span. We used local linear regression with the Lowess function but the same principle of span applies. Short spans give good independence between neighboring regions of the data set but a lot of random wiggling of the regression line. Long spans smooth out these wiggles but have the disadvantage that one region of the curve can significantly affect neighboring regions and this can obscure fine details in the underlying relationship. Spans are selected by inspection to be the shortest that give reasonable suppression of random wiggles. For the present study spans were varied between 0% and 100% and the largest sensitivity of P values found. We plotted this relationship separately for the initial conservative and early surgery groups on the same graph. Bootstrapping was used to test the significance of the area between the lines.

The nonparametric regression lines give a visual estimate of the likely difference between outcomes for the treatments under consideration. They can be viewed as the number of points on the outcomes scale that can be expected for a particular patient's characteristics when plotted on the independent variable (x) scale. The limitation of nonparametric statistics with regard to CIs led us to analyze these differences using a new statistic and that is the area between the curves. This is a direct measure of the difference between the positions of the 2 curves summed over the entire range of the independent variable being considered. How clinically relevant this statistic is has yet to be determined as it has not previously been used but it is readily related to the outcome scales.

Bootstrapping is a numerical method that makes no parametric assumptions. It is based on deriving a distribution for the statistic under study by doing the above-described analysis repeatedly on the data, but with the distinction that instead of dividing the patients into 2 groups according to how they were allocated in the trials, the patients are divided at random. Then after a million iterations enough area measurements have been generated to create a frequency distribution of areas which is centered on zero and is approximately normal. This distribution can then be used to compare with the area found on the initial data analysis and so determine what proportion of randomly generated areas are less than or more than the area we have found on analysis. This test is inherently 2-tailed because the area is not dependent on which of the 2 regression lines is the higher at any particular point.

We performed a sensitivity analysis to determine how sensitive the P values were to the robustness feature in the MATLAB Lowess function and to the selected span. In general, with short spans there is more random noise or wiggling in the regression curves with correspondingly larger areas between the curves and any systematic trend will be more

easily masked by these larger areas so shorter spans will lead to lower *P* values. At the other end of the spectrum, 100% span will give *P* values identical to a Wilcoxon rank-sum test unless the regressions lines cross. The main area of interest was in spans between 35% and 65%.

Meta-Analysis Using New Cut Points

This ranked regression system has the advantage that it can show when the lines cross each other, that is, when one treatment is superior over one range of a variable and the other is superior over another range. So, for example, we found surgery to be superior for large clot volumes and conservative to be superior for small. The crossing point of the regression lines gives the cut point between these ranges.

These cut points could be used to identify the appropriate regions to use for meta-analysis bands. A further post hoc meta-analysis was undertaken for GCS using the new cut points identified.

Each study obtained ethical approval before randomizing any patients. The trial protocols were published.²⁰⁻²²

Results

Meta-Analysis

Figure 1A through 1D shows the results of the meta-analyses. All the studies individually show a nonsignificant trend

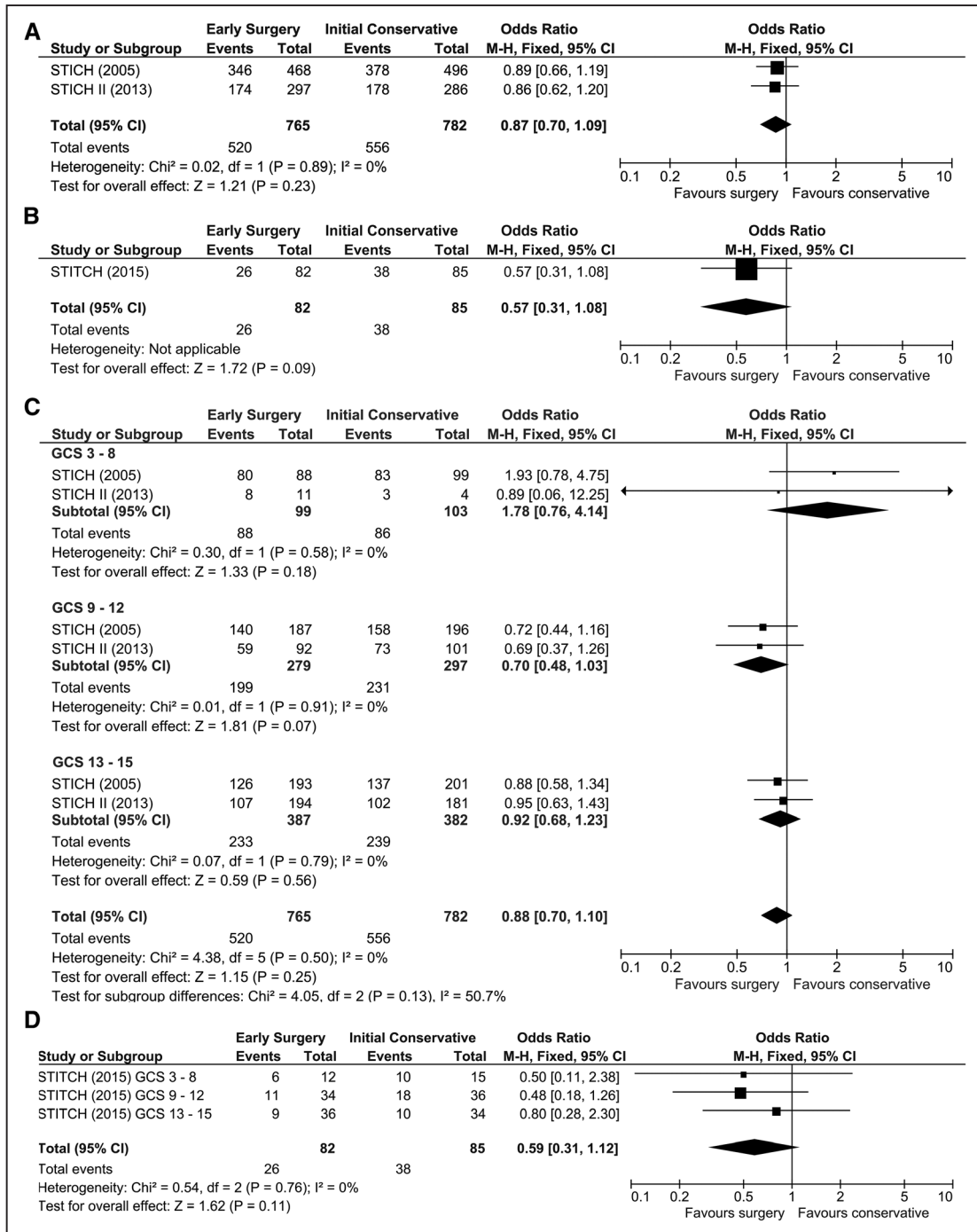


Figure 1. Forest plots. **A**, All spontaneous intracerebral hemorrhage (ICH) patients. **B**, All traumatic ICH patients. **C**, Spontaneous ICH by Scale (GCS) with standard grouping (<9, 9–12, and 13–15). **D**, Traumatic ICH by GCS with standard grouping (<9, 9–12, and 13–15).

towards favorable outcome with surgery. Figure 1A shows that the combined result from meta-analysis of the spontaneous ICH trials gives an overall odds ratio (OR) of 0.87 (95% CI, 0.70–1.09; $P=0.23$). Figure 1B shows that for traumatic ICH the OR is 0.57 (95% CI, 0.31–1.08; $P=0.09$). Figure 1C and 1D is the subgroup analyses looking at the 3 bands of the GCS. There is a wide variation in the ORs observed for the GCS band 3 to 8 with the results dominated by STICH showing no treatment effect. For GCS band 13 to 15 (the most conscious patients), the results are similar between the 2 studies giving a combined OR of 0.92 (95% CI, 0.68–1.23; $P=0.56$). For the intermediate GCS band 9

to 12, both studies give similar results with the combined OR for the spontaneous studies is 0.70 (95% CI, 0.48–1.03; $P=0.07$). In traumatic ICH, there is a tendency for improved outcome with surgery in both GCS band 3 to 8 (OR, 0.50 [95% CI, 0.11–2.38; $P=0.38$]) and GCS 9 to 12 band (OR, 0.48 [95% CI, 0.18–1.26; $P=0.14$]) but again no treatment effect in GCS band 13 to 15.

Ranked Analysis

Figure 2A through 2D shows the results of the ranked analyses. In these graphs, all patients are located on the y axis according to their GOSE ranking (where GOSE varies

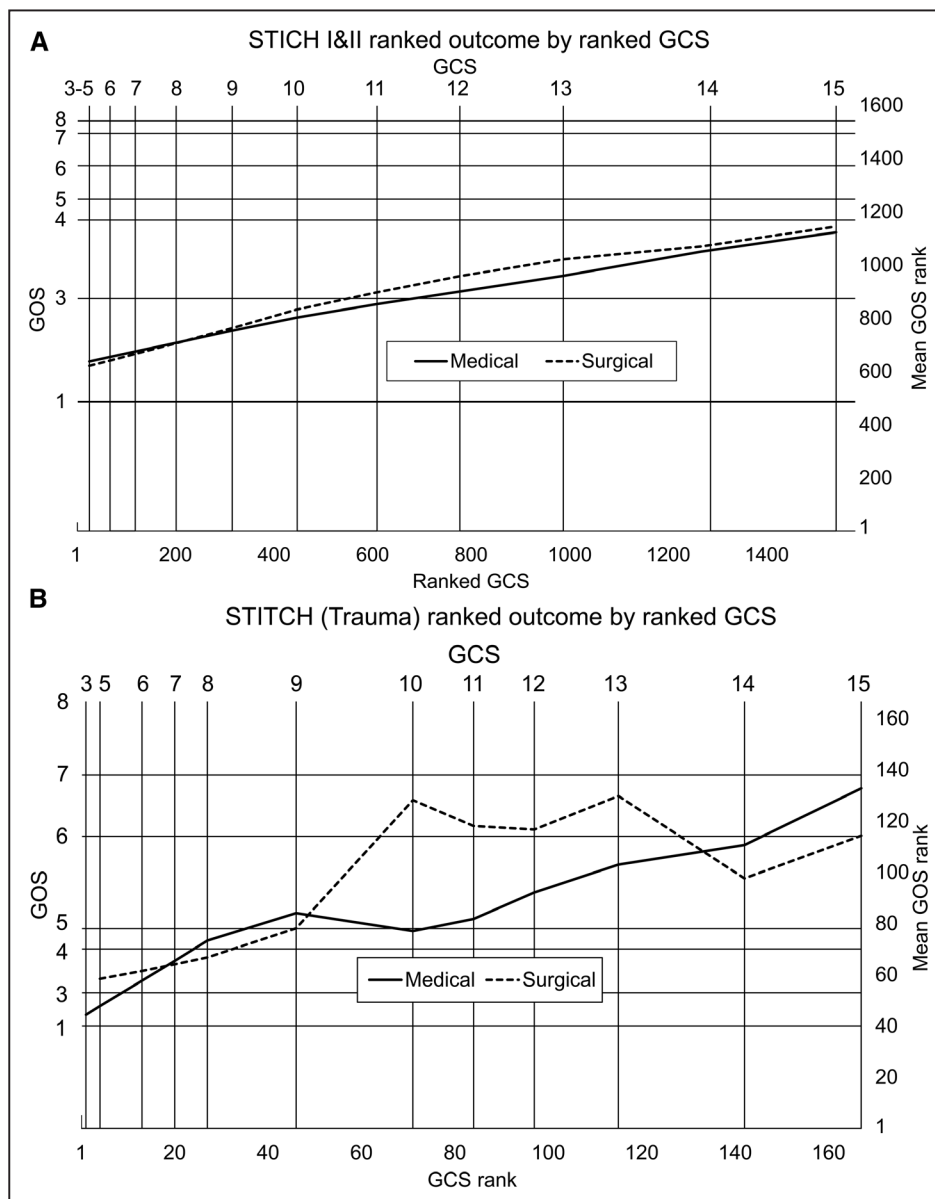


Figure 2. In these figures in all cases the total population, including both early surgery and initial conservative patients are ranked according to the parameters of Glasgow Coma Scale (GCS) at presentation and clot volume for the x axis and Extended Glasgow Outcome Scale (GOSE) for the y axis. A case's ranking is equal to the number of cases below or equal to that case in the parameter involved so that, for example, in the combined STICH I and II trials (Surgical Trial in Lobar Intracerebral Haemorrhage), there were 485 patients with a GOSE of 1 (died) and they were all given a ranking of 485 shown by the horizontal line. Only 2 patients had a GOSE of 2 (vegetative) so the grid line for this outcome cannot be resolved from that for GOSE 1. The x axis is simply all patients ranked. Equal spacing on the axes equals equal numbers of patients. On the y axis nonparametric regression values are plotted using Matlab 8.5.0.197613 (R2015a) malowess with robustness and spans given below. Spans were determined by inspection. The area between the lines was used as the statistic for P value calculations. **A**, Data combined from the STICH and STICH II trials plotting ranked presenting GCS against a nonparametric regression of ranked GOSE. Span was 0.4, $P=0.4$. **B**, Data from the STICH(Trauma) trial (Surgical Trial in Traumatic Intracerebral Haemorrhage) plotting ranked presenting GCS against nonparametric regression of ranked extended GOS. Span is 0.4, $P=0.012$. (Continued)

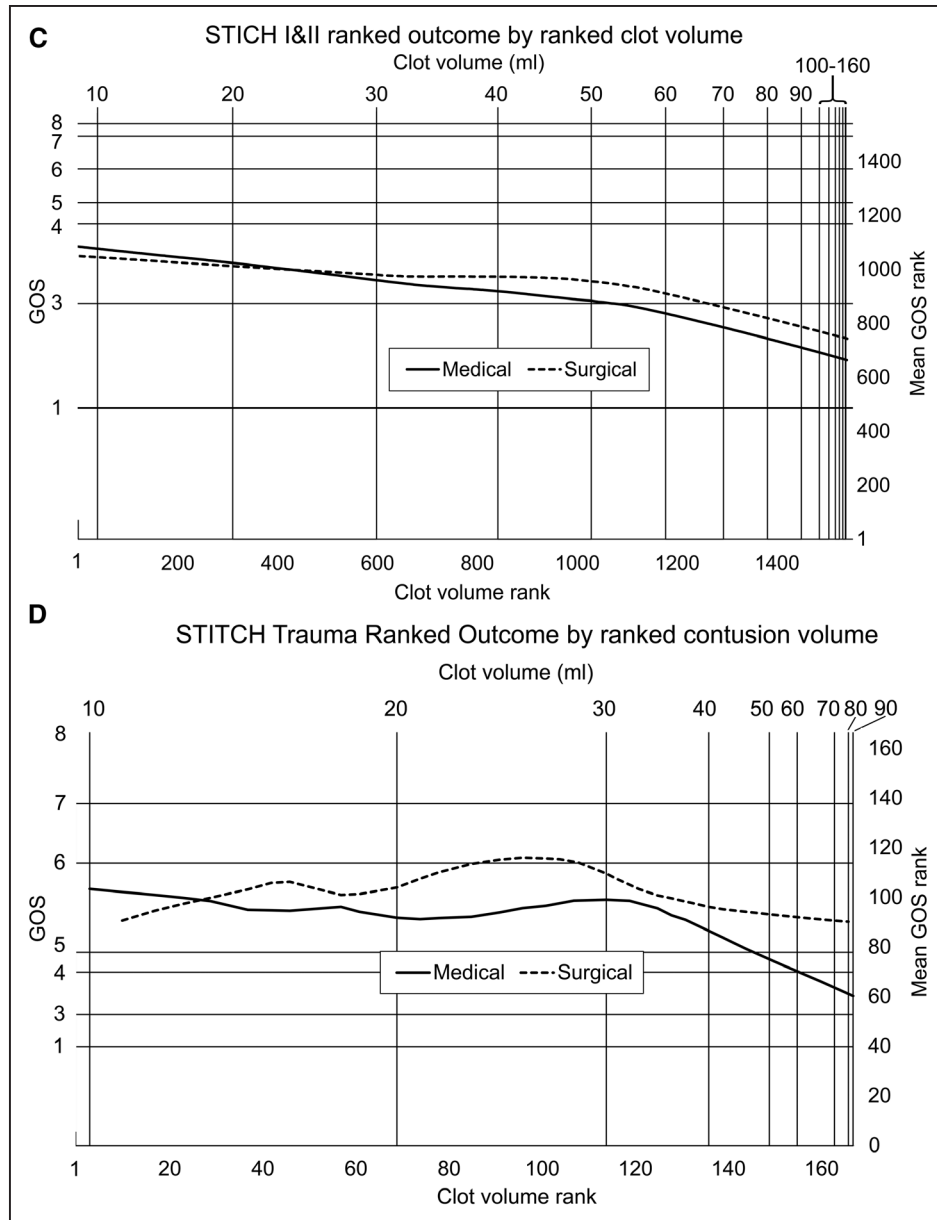


Figure 2 Continued. **C**, Data combined from the STICH and STICH II trials (Surgical Trial in Lobar Intracerebral Haemorrhage), plotting ranked presenting clot volume against nonparametric regression of ranked GOSE. Span was 0.55, $P=0.042$. **D**, Data from the STITCH(Trauma) trial (Surgical Trial in Traumatic Intracerebral Haemorrhage) plotting ranked presenting contusion volume against nonparametric regression of ranked GOS. Span is 0.4, $P=0.013$.

between 1 for dead and 8 for upper good recovery) and on the x axis according to their parameter ranking (GCS or volume). The plots are nonparametric regression of patients in the 2 groups: surgical and conservative. It is of interest that despite the combined STICH data set being much larger and of a different pathology to the STITCH(Trauma) data set, both show the same basic relationships. For presenting GCS, surgery may be more beneficial in an intermediate range and not at the extremes. Though this pattern was common to both data sets it was not significant for spontaneous ICH (spontaneous $P=0.4$, traumatic $P=0.012$). The maximum benefit from surgery seems to be for patients with a GCS of 10 to 13. For hematoma volume, surgery is detrimental for small volumes and beneficial for large volumes with benefit increasing the larger the volume.

Meta-Analysis Using New Cut Points

The meta-analysis of the spontaneous and traumatic ICH data sets undertaken using the new banding for GCS identified by the rank analysis is shown in Figure 3A and 3B. The OR for GCS band 10 to 13 in the spontaneous data set is 0.71 (95% CI, 0.51–1.00; $P=0.05$) and in the traumatic data set is 0.16 (95% CI, 0.05–0.51; $P=0.002$).

Discussion

Analysis of the data from 1541 randomized and monitored spontaneous ICH patients suggests that patients with an intermediate presenting GCS in the range of 9 to 12 may benefit from surgery whereas those with higher or lower GCS may not. This finding is consistent with surgical opinion which generally holds that in patients with a high GCS, the risks

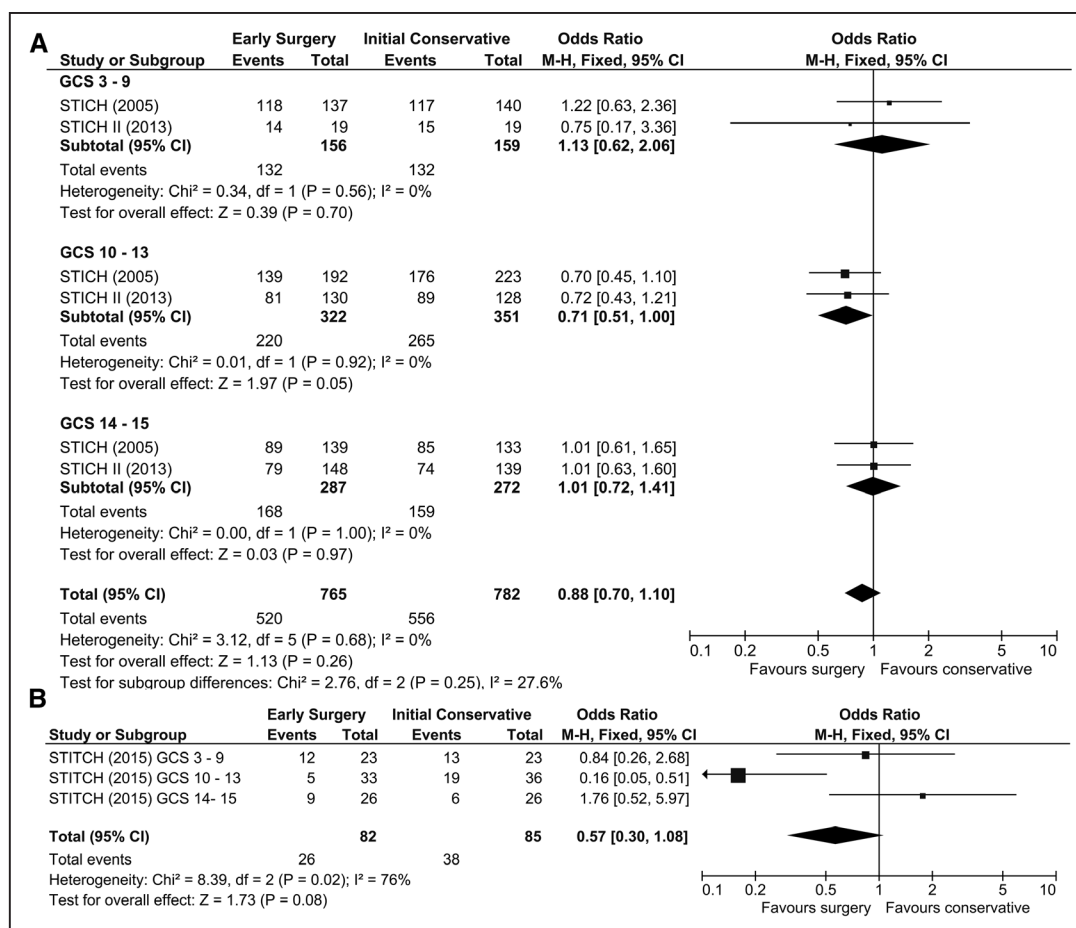


Figure 3. Forest plots using new categorization of Glasgow Coma Scale (GCS; ≤9, 10–13, and 14–15). **A**, Spontaneous intracerebral hemorrhage (ICH) studies. **B**, Traumatic ICH study.

of surgery are poorly justified by small potential benefit; and patients with a low GCS have a poor prognosis with or without surgery which consequently has little to offer.

The dichotomized method used for the primary analysis of the trials limits our ability to study how particularly GCS and volume interact with surgery. At the extremes of both GCS and volume, any effect of surgery would be expected to lie within the extreme ranges of the GOSE (death, upper good recovery) rather than between them and dichotomized outcomes that split the scale in the middle will not pick these up. To do so, we need to use the full GOSE rather than dichotomize it but this leads to a problem with the common method of relating variables—regression. Regression is a parametric method that relies on the assumption that variables are ordinal (each point on the scale is strictly less than the one below it and greater than the one above it) and uniform (and by the same amount). The ordinal assumption can be made for GOSE (because for each point on the scale, a patient at that point can do everything at the point below them plus something else) but not the uniform assumption.

The GCS is neither clearly ordinal nor uniform. It is made up of 3 subscales each of which can reasonably be regarded as ordinal but combining them means that it is not true that for every point on the GCS patients have worse coma than for the point above. However, in practice we consider the GCS to

be ordinal because its correlation with outcome in large data sets is ordinal.

We, therefore, use a regression method based on rank that only requires the ordinal and not the uniform assumption. Each subject is allocated a rank on the GOSE and GCS scales and on lesion volume. Then nonparametric regression of ranked GOSE versus another ranked variable is plotted.

Significance was tested with bootstrapping. Comparing areas under the 2 lines is related to the Wilcoxon ranked-sum test. This is fair in cases where the lines do not cross. Where they do cross it is not satisfactory because significance will be lost by combining a disadvantage in one region with a benefit in another irrespective of how large the effect is and whether or not it is real. Therefore, we use the statistic of area between the lines.

The standard meta-analysis found that patients with an intermediate presenting GCS in the range of 9 to 12 appeared to benefit from surgery whereas those with higher or lower GCS did not. The ranked analysis suggested that this benefit accrued for patients with a GCS in the range 10 to 13 and that this may be a better window to investigate in future trials as it was derived directly from the data rather than depending on an arbitrarily selected GCS range. The post hoc analysis confirmed that a GCS range of 10 to 13 identified those patients most likely to benefit from surgery.

For hematoma volume, the finding was of a negative effect of surgical evacuation at low volume passing through neutrality to a positive effect as volumes increased. Again this is consistent with surgical opinion which generally holds that the effect of surgery in this situation is mediated by decompression and that the greater the volume of clot the greater the potential for decompression such that at small volumes, the risk of surgery is not justified but this reverses with larger volumes.

The STICH and STITCH(Trauma) trials were similar in design but they addressed different pathologies, in view of which it is striking how similar are the patterns seen in both with the ranked analysis. We consider that it is possible that this pattern reflects general principles that apply across neurological conditions and neurosurgical operations. Those being the larger the surgical target the greater the potential benefit of surgery and with mild disease the risks of surgery outweigh the benefits. An idea for which there is also support from animal models.²³

Limitations

The nonparametric analysis is an exploratory post hoc analysis but has been undertaken on a large data set and offers better hypotheses for future trials in spontaneous and traumatic ICH.

The *P* values of the ranked analysis represent the probability that allocating cases from the whole population to groups at random would give an area between the lines of equal to or greater than that obtained by allocating cases according to the group they were randomized to in the trials. We consider all these *P* values to be marginal. A limitation of the method we used is that to some extent the *P* value is dependent on the selected span. At low span, more random wiggles in the lines increases the area between them. This reduces the proportion of that area that is because of systematic differences and so increases *P* values. At high span, key features of the line such as the boundary and crossing positions are more prone to bias. The conventional method of span selection, and the one we used, is by inspection. Regression plots are examined and the lowest span which leads to a reasonable suppression of random scatter is selected. It should be noted, however, that relatively small changes in span can affect the *P* value so that, for example, in the case of Figure 2B, changing the span from 0.55 to 0.4 changes the *P* value to 0.055. For this reason, we describe *P* values as marginal rather than showing definite significance or otherwise.

Conclusions

Meta-analysis favors the hypothesis that the neutral result of spontaneous ICH trials is because of the mixing of systematic positive and negative effects rather than there being a uniform small or absent effect. We identified that patients with an intermediate GCS (10–13) and a large spontaneous ICH may be more likely to benefit from surgery and that this might be a more appropriate categorization of the GCS for future studies of ICH. Our ranked analysis method shows similar effects on traumatic ICH/contusion data and may prove to be a valuable tool in assessing the effects of treatments.

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Disclosures

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References

- Auer LM, Deinsberger W, Niederkorn K, Gell G, Kleinert R, Schneider G, et al. Endoscopic surgery versus medical treatment for spontaneous intracerebral hematoma: a randomized study. *J Neurosurg.* 1989;70:530–535. doi: 10.3171/jns.1989.70.4.0530
- Batjer HH, Reisch JS, Allen BC, Plaizier LJ, Su CJ. Failure of surgery to improve outcome in hypertensive putaminal hemorrhage. A prospective randomized trial. *Arch Neurol.* 1990;47:1103–1106.
- Chen X, Wu J, Zhou X, YZhang Y, Wang Z, Qin Z, et al. The randomized multicentric prospective controlled trial in the standardized treatment of hypertensive intracerebral hematomas: the comparison of surgical therapeutic outcomes with conservative therapy. *Chinese J Clin Neurosci.* 2001;4:365–368.
- Chen X, Yang H, Cheng Z. A prospective randomised trial of surgical and conservative treatment for hypertensive intracerebral haemorrhage. *Acta Acad Med Shanghai.* 1992;19:237–240.
- Hattori N, Katayama Y, Maya Y, Gatherer A. Impact of stereotactic hematoma evacuation on activities of daily living during the chronic period following spontaneous putaminal hemorrhage: a randomized study. *J Neurosurg.* 2004;101:417–420. doi: 10.3171/jns.2004.101.3.0417
- Hosseini H, Leguerinel C, Hariz M, Melon E, Palfi S, Deck P, et al. Stereotactic aspiration of deep intracerebral haematomas under computed tomographic control, a multicentric prospective randomised trial. *Cerebrovasc Dis.* 2003;16S4:57.
- Juvela S, Heiskanen O, Poranen A, Valtonen S, Kuurne T, Kaste M, et al. The treatment of spontaneous intracerebral hemorrhage. A prospective randomized trial of surgical and conservative treatment. *J Neurosurg.* 1989;70:755–758. doi: 10.3171/jns.1989.70.5.0755
- McKissock W, Richardson A, Taylor J. Primary intracerebral haemorrhage: a controlled trial of surgical and conservative treatment in 180 unselected cases. *Lancet.* 1961;2:221–226.
- Mendelow AD, Gregson BA, Fernandes HM, Murray GD, Teasdale GM, Hope DT, et al; STICH Investigators. Early surgery versus initial conservative treatment in patients with spontaneous supratentorial intracerebral haematomas in the International Surgical Trial in Intracerebral Haemorrhage (STICH): a randomised trial. *Lancet.* 2005;365:387–397. doi: 10.1016/S0140-6736(05)17826-X
- Mendelow AD, Gregson BA, Rowan EN, Francis R, McColl E, McNamee P, et al; STITCH(Trauma) Investigators. Early Surgery versus Initial Conservative Treatment in Patients with Traumatic Intracerebral Hemorrhage (STITCH[Trauma]): the first randomized trial. *J Neurotrauma.* 2015;32:1312–1323. doi: 10.1089/neu.2014.3644
- Mendelow AD, Gregson BA, Rowan EN, Murray GD, Gholkar A, Mitchell PM; STICH II Investigators. Early surgery versus initial conservative treatment in patients with spontaneous supratentorial lobar intracerebral haematomas (STICH II): a randomised trial. *Lancet.* 2013;382:397–408. doi: 10.1016/S0140-6736(13)60986-1
- Morgenstern LB, Frankowski RF, Shedden P, Pasteur W, Grotta JC. Surgical treatment for intracerebral hemorrhage (STICH): a single-center, randomized clinical trial. *Neurology.* 1998;51:1359–1363.
- Pantazis G, Tsitsopoulos P, Mihas C, Katsiva V, Stavrianos V, Zymaris S. Early surgical treatment vs conservative management for spontaneous supratentorial intracerebral hematomas: a prospective randomized study. *Surg Neurol.* 2006;66:492–501; discussion 501. doi: 10.1016/j.surneu.2006.05.054

14. Teernstra OP, Evers SM, Lodder J, Leffers P, Franke CL, Blaauw G; Multicenter Randomized Controlled Trial (SICHPA). Stereotactic treatment of intracerebral hematoma by means of a plasminogen activator: a multicenter randomized controlled trial (SICHPA). *Stroke*. 2003;34:968–974. doi: 10.1161/01.STR.0000063367.52044.40
15. Wang WZ, Jiang B, Liu HM, Li D, Lu CZ, Zhao YD, et al. Minimally invasive craniopuncture therapy vs. conservative treatment for spontaneous intracerebral hemorrhage: results from a randomized clinical trial in China. *Int J Stroke*. 2009;4:11–16. doi: 10.1111/j.1747-4949.2009.00239.x
16. Zuccarello M, Brott T, Derex L, Kothari R, Sauerbeck L, Tew J, et al. Early surgical treatment for supratentorial intracerebral hemorrhage: a randomized feasibility study. *Stroke*. 1999;30:1833–1839.
17. Gregson BA, Broderick JP, Auer LM, Batjer H, Chen XC, Juvela S, et al. Individual patient data subgroup meta-analysis of surgery for spontaneous supratentorial intracerebral hemorrhage. *Stroke*. 2012;43:1496–1504. doi: 10.1161/STROKEAHA.111.640284
18. Prasad K, Mendelow AD, Gregson BA. Surgery for primary supratentorial intracerebral haemorrhage. *Cochrane Database of Syst Rev*. 2008;CD000200. doi: 10.1002/14651858.CD000200.
19. Murray GD, Barer D, Choi S, Fernandes H, Gregson B, Lees KR, et al. Design and analysis of phase III trials with ordered outcome scales: the concept of the sliding dichotomy. *J Neurotrauma*. 2005;22:511–517. doi: 10.1089/neu.2005.22.511
20. Protocol 99PRT/7:International STICH (Surgical Trial in IntraCerebral Haemorrhage). *Lancet* <https://www.thelancet.com/protocol-reviews/99PRT-7> July 14, 2003. Accessed March 5, 2019.
21. Gregson BA, Rowan EN, Mitchell PM, Unterberg A, McColl EM, Chambers IR, et al. Surgical trial in traumatic intracerebral hemorrhage (STITCH(Trauma)): study protocol for a randomized controlled trial. *Trials*. 2012;13:193. doi: 10.1186/1745-6215-13-193
22. Mendelow AD, Gregson BA, Mitchell PM, Murray GD, Rowan EN, Gholkar AR; STICH II Investigators. Surgical trial in lobar intracerebral haemorrhage (STICH II) protocol. *Trials*. 2011;12:124. doi: 10.1186/1745-6215-12-124
23. Mendelow AD. Mechanisms of ischemic brain damage with intracerebral hemorrhage. *Stroke*. 1993;24(suppl 12):I115–I117; discussion I118.