BMJ Open

Determinants of immunisation coverage of children aged 12–59 months in Indonesia: a cross-sectional study

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To cite: Herliana P, Douiri A. Determinants of immunisation coverage of children aged 12–59 months in Indonesia: a cross-sectional study. *BMJ Open* 2017;**7**:e015790. doi:10.1136/bmjopen-2016-015790

► Prepublication history for this paper is available online. To view these files, please visit the journal online (http://dx.doi.org/10.1136/bmjopen-2016-015790).

Received 30 December 2016
Revised 23 August 2017
Accepted 8 September 2017

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ABSTRACT

Objectives Despite the adoption of WHO's Expanded Programme on Immunisation in Indonesia since 1977, a large proportion of children are still completely unimmunised or only partly immunised. This study aimed to assess factors associated with low immunisation coverage of children in Indonesia.

Setting Children aged 12–59 months in Indonesia.

Participant The socioeconomic characteristics and immunisation status of the children were obtained from the most recent Demographic and Health Survey, the 2012 Indonesia Demographic and Health Survey. Participants were randomly selected through a two-stage stratified sampling design. Data from 14 401 children aged 12–59 months nested within 1832 census blocks were included in the analysis. Multilevel logistic regression models were constructed to account for hierarchical structure of the data.

Results The mean age of the children was 30 months and they were equally divided by sex. According to the analysis, 32% of the children were fully immunised in 2012. Coverage was significantly lower among children who lived in Maluku and Papua region (adjusted OR: 1.94; 95% CI 1.42 to 2.64), were 36–47 months old (1.39; 1.20 to 1.60), had higher birth order (1.68; 1.28 to 2.19), had greater family size (1.47; 1.11 to 1.93), whose mother had no education (2.13; 1.22 to 3.72) and from the poorest households (1.58; 1.26 to 1.99). The likelihood of being unimmunised was also higher among children without health insurance (1.16; 1.04 to 1.30) and those who received no antenatal (3.28; 2.09 to 5.15) and postnatal care (1.50; 1.34 to 1.69).

Conclusions Socioeconomic factors were strongly associated with the likelihood of being unimmunised in Indonesia. Unimmunised children were geographically clustered and lived among the most deprived population. To achieve WHO target of protective coverage, public health interventions must be designed to meet the needs of these high-risk groups.

Strengths and limitations of this study

- Our study investigated, for the first time, the factors associated with routine immunisation coverage of children in Indonesia using data from the most recent demographic and health survey.
- The large sample size allowed us to analyse many potential predictors simultaneously and produce reliable estimates.
- We used multilevel modelling to account for the hierarchical structure of the data.
- However, we could only build a two-level model (ie, children nested within census blocks) instead of the ideal three-level model (ie, children within households nested within census blocks) because there was no household identifier in the dataset, as it may compromise the participants' anonymity.
- The selection of variables included in this study also relied on the information available from the dataset.

BACKGROUND

In 1974, the WHO initiated the Expanded Programme on Immunisation (EPI) with the goal of providing universal immunisation for all children.[1] The first diseases targeted were diphtheria, tetanus, pertussis, polio, measles and tuberculosis.[1] New and increasingly sophisticated vaccines have become available, and more children than ever before are being vaccinated today.[2 3] Global coverage increased from 74% in 2000 to 86% in 2014.[4] As a result, the annual number of child deaths fell from 9.6 million in 2000 to 5.9 million in 2015.[1 4] Immunisation drives this reduction in child mortality and the collective recognition has led to the development of the Global Vaccine Action Plan (GVAP), a framework to help countries achieve universal child immunisation by 2020.[3] The target, as stated in the United Nations Sustainable Development Goals, is to end preventable child deaths by 2030.[5]

Despite this progress, vaccine-preventable diseases are still responsible for 1.5 million child deaths each year.[6] Almost 18.7 million children were not given routine immunisation in 2014, and 75% of them live in only 10 countries in Africa and Asia.[4] Although some regions have successfully maintained a high level of immunisation coverage, there are pockets of unimmunised children which induce the continuous spread of diseases and outbreaks.[2] This highlights the fact that global coverage may hide variability between

countries. It also suggests that the achievements are still fragile. Should this trend continue, the goals of providing universal immunisation for all children by 2020 and ending vaccine-preventable deaths by 2030 could not be achieved, and the cost of such failure would be close to 26 million deaths.[3]

One of the 10 countries that are home to the highest number of unimmunised children is Indonesia.[4] Indonesia is a lower middle-income country located in Southeast Asia.[7] It has an estimated population of over 255 million in 2015, 10% of whom are children under the age of five.[8] Child mortality rate in Indonesia currently stands at 27 deaths per 1000 births and ranks 101st out of 175 countries.[9] Approximately 36% of child deaths were caused by infectious diseases.[10] For most of these diseases, vaccines are available to prevent child deaths.

The Indonesian Ministry of Health (MOH), which organises public health matters within the Indonesian government, has adopted and implemented the EPI guidelines since 1977 through a routine immunisation programme that is compulsory for all children.[11] Even so, a large number of young children in Indonesia are still either completely unimmunised or only partly immunised. In 2013, the MOH has reported that only 59.2% of children were fully immunised.[11] There were also striking gaps within the country as coverage was as low as 29.2% at a certain area in Indonesia.[11] These figures were well below the 90% advised threshold that is required to maintain herd immunity and prevent the spread of diseases.[3] As the fourth most populous country in the world with a great proportion of young children, the risk of large and uncontrollable outbreaks in Indonesia is more likely than ever.

To significantly increase coverage in Indonesia, a strategy proposed by GVAP is to identify and engage the unimmunised children.[3] These children are often the ones carrying a heavier burden of diseases.[3] There is particular concern that diseases may thrive when unimmunised children are residentially segregated from immunised children.[2] It is therefore critical to know who they are, where they live and what factors might have contributed to their unimmunised status to ascertain where greater efforts are needed.

While administrative and geographic barriers may contribute to low coverage in a country with such a large population,[12] GVAP explicitly highlights the importance of socioeconomic factors in determining coverage.[3] Theory suggests that factors such as income level, employment status and education are major determinants of healthcare use,[13] and a growing body of empirical evidence advances such association. The socioeconomic characteristics attached to routine immunisation coverage, and the extent these factors may play a role, vary by country.[12 14–24] However, no such research has been done in Indonesia.

In this study, we used data from the 2012 Indonesia Demographic and Health Survey (IDHS) which collected information on both the immunisation status and the socioeconomic characteristics of Indonesian children under 5 years of age. Our aim was to identify the socioeconomic factors associated with routine immunisation coverage of children in Indonesia. The results should help in identifying susceptible subgroups of the population that require additional resources and focused attention.

METHODS

Data source

This study is a secondary data analysis of the most recent DHS in Indonesia. The IDHS is conducted routinely by the national statistics authority Statistics Indonesia, in collaboration with the National Population and Family Planning Board, the Indonesian MOH and ICF International.[25] Studies on its quality suggest that DHS is nationally representative, with little evidence of systematic bias.[26]

Data were collected from 7 May to 31 July 2012. Participants were selected through a two-stage stratified sampling design. The primary sampling unit was the census block (CB) and the complete list of households in each CB became the basis for second-stage sampling. A total of 46 024 households were chosen as the sample. From 44 302 occupied households, 45 607 women aged 15–49 were successfully interviewed, yielding a response rate of 96%.

The Women's Questionnaire included questions about the woman's background characteristics and her children aged under 5, for whom immunisation and health data were collected. The dataset had one record for every child of each interviewed woman, born in the 5 years preceding the survey. Data were obtained for 18 021 children.

Outcome variable

The outcome variable in the analysis was the child's immunisation status. Information on immunisation status was collected from two sources, the health card or health book shown to the interviewer or if unavailable from the mother's report. The health card or health book was available 85.77% of the time.

The outcome variable was categorised as 'fully immunised' if they had received the full schedule of routine immunisation and otherwise 'unimmunised', regardless of the source of the information. Routine immunisation referred to three doses of DTP vaccines, four doses of polio vaccine, one dose of measles vaccine, one dose of Bacillus Calmette–Guérin (BCG) vaccine and four doses of hepatitis B vaccine, scheduled to be received by the age of 12 months.[11] The proportion of children who had been fully immunised defined immunisation coverage.[27]

In a small number of cases, where health cards were unavailable and mothers indicated that they did not know about the immunisation status (1.51%), the child was considered as not fully immunised. The fact that mothers responded 'don't know' is likely to reflect that the child was not fully immunised[12 28] and fits better in the 'unimmunised' category.

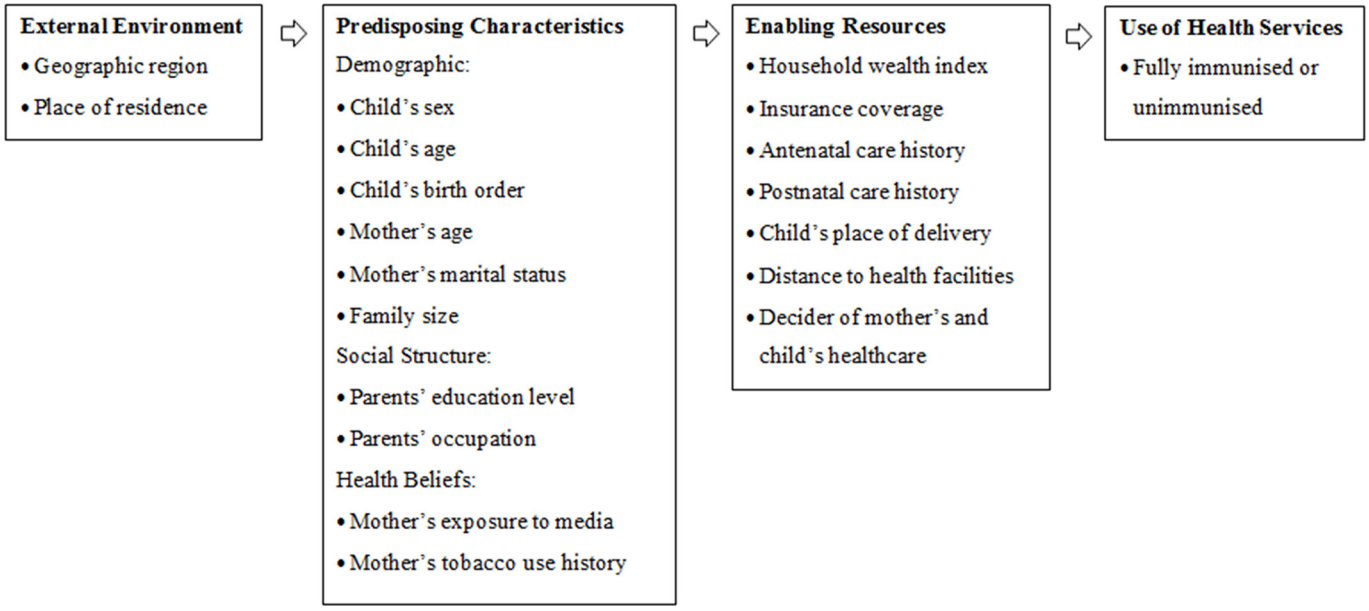

Figure 1 Theoretical framework of factors potentially associated with immunisation coverage of children in Indonesia, informed by Andersen's Behavioural Health Model.

Independent variables

Selection of independent variables was based on the literature review and variables available in the dataset. Twenty-two independent variables were identified as potential factors and Andersen's Behavioural Health Model[13] was used as a framework to group the factors into three main groups: external environment, predisposing and enabling factors (figure 1). The model has been commonly used to examine factors associated with health service use, including immunisation uptake.[21 29]

Predisposing characteristics consist of demographic factors, social structure such as educational attainment and occupation and health beliefs which involve health-related knowledge and behaviours.[13] Enabling resources are related to individuals' personal and community support which enable them to use health services, reflected by income level, insurance coverage and other factors that could affect one's access to health services.[13] Lastly, external environment incorporates wider social and environmental determinants of health.[13]

Categorisation of continuous variables and description of categorical variables were undertaken according to the literature. The child's age (12–59 months) was categorised into groups at 1-year intervals. Similarly, the mother's age (15–49 years) was categorised into groups at 5-year intervals. The child's birth order and family size were also categorised into groups based on previously published literatures.

Following IDHS protocol,[25] household wealth index was constructed based on household amenities and assets (radio, television, refrigerator, bicycle, motorcycle or car) and dwelling characteristics (electricity, flooring, roofing, water source, toilet facilities and sleeping arrangements). It was categorised into quintiles from poorest to richest. In the absence of direct information on household income or expenditures, wealth index is considered a robust measure of household income level.[30] Insurance coverage represented any health insurance provided through social security or local government, by employer, privately purchased or other insurance. Antenatal care represented any pregnancy-related care provided by skilled health personnel or traditional birth attendants during the pregnancy, irrespective of the type of provider and the number of visits. Similarly, postnatal care represented any examination by skilled health personnel or traditional birth attendants within 2 months of the child's birth, irrespective of the type of provider and the number of visits.

The 33 provinces in Indonesia were categorised into six island-based regions.[25] The child's place of delivery was classified into three categories: home, public health institution and private health institution. Public health institution included public hospitals, public clinics, health centres, village health posts and delivery posts. Private health institution included private hospitals, private clinics, maternity hospitals, maternity home and also private practices of obstetrician, general practitioner, nurse, midwife and village midwife.

Statistical analysis

The original dataset comprised of 18 021 children aged 0–59 months distributed among 1840 CBs. For the purpose of the analysis, we excluded 3620 children who were under 1 year old because they were not old enough to have received the full schedule of routine immunisation in Indonesia. The final sample, therefore, contained 14 401 children from 1832 CBs. From this, we had 656 children (4.6%) with missing immunisation status because they were no longer alive at the time of the survey, leaving complete observations of 13 745 children

(95.4%). Given the small number of missing values, we used complete case analysis and no sensitivity analysis was required.

Data analysis was conducted using STATA V.14 software. Frequency and percentage were used to report baseline characteristics of the children. Cross-tabulation was undertaken to demonstrate the proportion of different categories with respect to immunisation status. The immunisation status as outcome variable was coded into 0 for 'fully immunised' and 1 for otherwise 'unimmunised'.

Univariate analysis was used to separately evaluate the effect of each independent variable on the outcome variable. Test of trends across ordered groups were evaluated. Variables with a univariate p Value of less than 0.2 were then selected as candidates for the multivariate analysis.

Multilevel logistic regression was used to estimate immunisation status in multivariate context while accounting for clustering. Model fitting using residuals were checked. A two-level model was used for the multivariate analysis (ie, children nested within CBs). This was run using the meqrlogit command in STATA V.14, a method based on maximum likelihood and robust to missing values. Associations between independent variables and the likelihood of children being unimmunised were assessed simultaneously. The results were expressed as adjusted OR with 95% CI.

RESULTS

Descriptive statistics

A total of 14 401 children from 1832 CBs were included in the analysis. Our result showed that only 31.5% (95% CI 30.7% to 32.3%) of the children aged 12–59 months had been fully immunised at the time of the survey. The baseline characteristics of sample were presented in table 1.

The mean age of the children was 30 months and they were equally divided by sex. More than half of them were second to fourth born. The mothers were 25 to 29 years old on average and almost all were married at the time of the survey. Most of the families had five to nine household members.

Majority of the mothers were secondary school graduates. Although educational attainment was approximately equal for both parents, nearly half of the mothers did not work. A large proportion of the mothers were exposed to media at least once a week and almost all reported that they did not smoke around the time of the survey.

In terms of enabling resources, half of the children lived in the poorer and poorest households. Additionally, almost two-thirds of the children were not covered by health insurance. While only a small proportion were born without antenatal care, much more children were born without postnatal care. Nearly half of the children were delivered at home although most mothers reported that distance to health facilities were not a big problem. Lastly, the majority of mothers reported that they were involved in the decision-making process of their own healthcare as well as their children's.

Table 1 Baseline characteristics of sample (n=14 401)

Characteristics	Frequency*	Percentage (%)
Immunisation status		
Fully immunised	4331	31.5
Unimmunised	9414	68.5
External environment		
Geographic region		
Sumatera	4061	29.5
Java	3079	22.4
Bali and Nusa Tenggara	1220	9.0
Kalimantan	1447	10.5
Sulawesi	2381	17.3
Maluku and Papua	1557	11.3
Place of residence		
Urban	6307	45.9
Rural	7438	54.1
Predisposing characteristics		
Child's sex		
Male	7092	51.6
Female	6653	48.4
Child's age (months)		
12–23	3501	25.5
24–35	3413	24.8
36–47	3378	24.6
48–59	3453	25.1
Child's birth order		
1st	5929	35.9
2nd–4th	7533	54.8
≥5th	1283	9.3
Mother's age (years)		
15–19	262	1.9
20–24	2381	17.3
25–29	3928	28.6
30–34	3454	25.2
35–39	2410	17.5
40–44	1104	8.0
45–49	206	1.5
Mother's marital status		
Married	13 168	95.8
Living with partner	176	1.3
Widowed	118	0.8
Divorced	231	1.7
No longer living together	43	0.3
Never in union	9	0.1
Family size (number of household members)		
≤4	5314	38.6
5–9	7637	55.6

Continued

Table 1 Continued

Characteristics	Frequency*	Percentage (%)
≥10	794	5.8
Mother's educational level		
Higher	1819	13.2
Secondary	7221	52.6
Primary	4291	31.2
No education	414	3.0
Father's educational level		
Higher	1740	12.7
Secondary	7438	54.2
Primary	4204	30.6
No education	311	2.3
Don't know	24	0.2
Mother's occupation		
Professional	1018	7.4
Agricultural	1855	13.5
Industrial	1571	11.4
Clerical, services and sales	3236	23.6
Did not work	6052	44.1
Do not know	2	0.0
Father's occupation		
Professional	1336	9.8
Agricultural	3550	25.9
Industrial	4884	35.6
Clerical, services and sales	3709	27.0
Did not work	225	1.6
Do not know	12	0.1
Mother's exposure to media (newspaper, magazine, radio or television)		
At least once a week	11528	83.9
Less than once a week	1527	11.1
Not at all	686	5.0
Mother's tobacco use history		
Smokes nothing	13317	96.9
Uses tobacco	424	3.1
Enabling resources		
Household wealth index		
Richest	2108	15.3
Richer	2276	16.6
Middle	2504	18.2
Poorer	2722	19.8
Poorest	4135	30.1
Covered by health insurance		
Yes	5580	40.6
No	8156	59.4
Antenatal care		
Received some care	10861	96.2
Received no care	640	3.8
Postnatal care		
Received some care	7395	65.7
Received no care	3813	33.8
Don't know	53	0.5
Child's place of delivery		
Home	6325	46.2
Public health institution	2527	18.4
Private health institution	4823	35.2
Other	28	0.2
Distance to health facilities		
Not a big problem	11915	86.9
Big problem	1792	13.1
Maternal healthcare decision-making		
By herself	4758	35.7
Jointly with husband	6567	49.3
Husband alone	1972	14.7
By others	34	0.3
Child healthcare decision-making		
By herself	4497	36.3
Jointly with husband	1407	50.5
Husband alone	6255	11.4
By others	225	1.8

*Total number varies between categories because of missing values.

Univariate analysis

The association between each independent variable and the likelihood of being unimmunised was investigated one by one. The result were shown in table 2.

Geographic region came out as a significant predictor of immunisation coverage in our univariate analysis. The majority, one-third, of children who were fully immunised lived in Java, while the lowest coverage was reported in Maluku and Papua. The odds of being unimmunised were almost threefold among children who lived in Maluku and Papua (OR 2.80; 95% CI 2.42 to 3.24). On the contrary, we found that children from Bali and Nusa Tenggara had the least likelihood of being unimmunised (OR 0.86; 95% CI 0.75 to 0.99). Our univariate analysis also showed that children from rural areas were significantly more likely to be unimmunised compared with their urban counterparts (OR 1.39; 95% CI 1.30 to 1.50).

Although coverage was approximately equal for both sexes, the child's age and birth order were significantly associated with coverage. Older children were more likely to be unimmunised compared with the youngest ones. The odds of being unimmunised among the older

Table 2 Univariate analysis results for factors associated with low immunisation coverage of children aged 12–59 months in Indonesia

Characteristics	Status (%)				Unadjusted OR (95% CI)	p Value
	Fully immunised		Unimmunised			
External environment						
Geographic region						
Sumatera	1135	(26.2)	2926	(31.8)	1.68 (1.52 to 1.86)	0.000
Java	1215	(28.1)	1864	(19.8)	1	
Bali and Nusa Tenggara	525	(12.1)	695	(7.4)	0.86 (0.75 to 0.99)	0.032
Kalimantan	490	(11.3)	957	(10.2)	1.27 (1.12 to 1.45)	0.000
Sulawesi	672	(15.5)	1709	(18.2)	1.66 (1.48 to 1.86)	0.000
Maluku and Papua	294	(6.8)	1263	(13.4)	2.80 (2.42 to 3.24)	0.000
Place of residence						
Urban	2232	(51.5)	4075	(43.3)	1	
Rural	2099	(48.5)	5339	(56.7)	1.39 (1.30 to 1.50)	0.000
Predisposing characteristics						
Child's sex						
Male	2255	(52.1)	4837	(51.4)	1	
Female	2076	(47.9)	4577	(48.6)	1.03 (0.96 to 1.10)	0.455
Child's age (months)						
12–23	1246	(28.8)	2255	(24.0)	1	
24–35	1066	(24.6)	2347	(24.9)	1.22 (1.10 to 1.34)	0.000
36–47	1011	(23.3)	2367	(25.1)	1.30 (1.17 to 1.43)	0.000
48–59	1008	(23.3)	2445	(26.0)	1.34 (1.21 to 1.48)	0.000
Child's birth order						
First	1675	(38.7)	3254	(34.6)	1	
Second–fourth	2413	(55.7)	5120	(54.4)	1.29 (1.21 to 1.37)	0.000
≥5th	243	(5.6)	1040	(11.0)	1.41 (1.27 to 1.57)	0.000
Mother's age (years)						
15–19	67	(1.5)	195	(2.1)	1	
20–24	704	(16.2)	1677	(17.8)	0.82 (0.61 to 1.10)	0.178
25–29	1219	(28.2)	2709	(28.8)	0.76 (0.57 to 1.02)	0.064
30–34	1166	(26.9)	2288	(24.3)	0.67 (0.51 to 0.90)	0.007
35–39	815	(18.8)	1595	(16.9)	0.67 (0.50 to 0.90)	0.007
40–44	301	(7.0)	803	(8.5)	0.92 (0.67 to 1.25)	0.579
45–49	59	(1.4)	147	(1.6)	0.86 (0.57 to 1.29)	0.458
Mother's marital status						
Married	4159	(96.0)	9009	(95.7)	1	
Living with partner	50	(1.2)	126	(1.3)	1.16 (0.84 to 1.62)	0.368
Widowed	37	(0.9)	81	(0.9)	1.01 (0.68 to 1.49)	0.958
Divorced	70	(1.6)	161	(1.7)	1.06 (0.80 to 1.41)	0.678
No longer living together	11	(0.3)	32	(0.3)	1.34 (0.68 to 2.67)	0.400
Never in union	4	(0.0)	5	(0.1)	0.58 (0.15 to 2.15)	0.413
Family size (number of household members)						
≤4	1746	(40.3)	3568	(37.9)	1	
5–9	2381	(55.0)	5256	(55.8)	1.08 (1.00 to 1.16)	0.044
≥10	204	(4.7)	590	(6.3)	1.42 (1.20 to 1.68)	0.000

Continued

Table 2 Continued

Characteristics	Status (%) Fully immunised		Unimmunised		Unadjusted OR (95% CI)	p Value
Mother's educational level						
Higher	756	(17.5)	1063	(11.3)	1	
Secondary	2451	(56.6)	4770	(50.7)	1.38 (1.25 to 1.54)	0.000
Primary	1081	(25.0)	3210	(34.1)	2.11 (1.88 to 2.37)	0.000
No education	43	(0.9)	371	(3.9)	6.14 (4.41 to 8.53)	0.000
Father's educational level						
Higher	717	(16.6)	1023	(10.9)	1	
Secondary	2508	(58.0)	4930	(52.5)	1.38 (1.24 to 1.53)	0.000
Primary	1054	(24.4)	3150	(33.5)	2.09 (1.86 to 2.36)	0.000
No education	42	(1.0)	269	(2.9)	4.49 (3.20 to 6.30)	0.000
Don't know	3	(0.0)	21	(0.2)	4.91 (1.46 to 16.5)	0.010
Mother's occupation						
Professional	428	(9.9)	590	(6.3)	1	
Agricultural	405	(9.4)	1450	(15.4)	2.60 (2.20 to 3.07)	0.000
Industrial	480	(11.1)	1091	(11.6)	1.65 (1.40 to 1.94)	0.000
Clerical, services and sales	1069	(24.7)	2167	(23.0)	1.47 (1.27 to 1.70)	0.000
Did not work	1944	(44.9)	4108	(43.7)	1.53 (1.34 to 1.76)	0.000
Father's occupation						
Professional	520	(12.0)	816	(8.7)	1	
Agricultural	809	(18.7)	2741	(29.2)	2.16 (1.89 to 2.47)	0.000
Industrial	1584	(36.7)	3300	(35.1)	1.33 (1.17 to 1.50)	0.000
Clerical, services and sales	1350	(31.2)	2359	(25.1)	1.11 (0.98 to 1.27)	0.102
Did not work	58	(1.4)	167	(1.8)	1.83 (1.33 to 2.52)	0.000
Don't know	2	(0.0)	10	(0.1)	3.19 (0.70 to 14.6)	0.136
Mother's exposure to media (newspaper, magazine, radio or television)						
At least once a week	3814	(88.1)	7714	(82.0)	1	
Less than once a week	373	(8.6)	1154	(12.2)	1.53 (1.35 to 1.73)	0.000
Not at all	142	(3.3)	544	(5.8)	1.89 (1.57 to 2.29)	0.000
Mother's tobacco use history						
Smokes nothing	4246	(98.0)	9071	(96.4)	1	
Uses tobacco	85	(2.0)	339	(3.6)	1.87 (1.47 to 2.37)	0.000
Enabling resources						
Household wealth index						
Richest	914	(21.1)	1194	(12.7)	1	
Richer	834	(19.2)	1442	(15.3)	1.32 (1.17 to 1.49)	0.000
Middle	883	(20.4)	1621	(17.2)	1.41 (1.25 to 1.58)	0.000
Poorer	848	(19.6)	1874	(19.9)	1.69 (1.50 to 1.90)	0.000
Poorest	852	(19.7)	3283	(34.9)	2.95 (2.63 to 3.31)	0.000
Covered by health insurance						
Yes	1993	(46.0)	3587	(38.1)	1	
No	2336	(54.0)	5820	(61.9)	1.38 (1.29 to 1.49)	0.000
Antenatal care						
Received some care	3668	(99.0)	7193	(94.8)	1	
Received no care	38	(1.0)	394	(5.2)	5.29 (3.78 to 7.39)	0.000

Continued

Table 2 Continued

Characteristics	Status (%) Fully immunised		Unimmunised		Unadjusted OR (95% CI)	p Value
Postnatal care						
Received some care	2732	(73.8)	4663	(61.7)	1	
Received no care	958	(25.9)	2855	(37.8)	1.75 (1.60 to 1.90)	0.000
Don't know	14	(0.3)	39	(0.5)	1.63 (0.88 to 3.01)	0.117
Child's place of delivery						
Home	1376	(31.8)	4949	(52.8)	1	
Public health institution	1041	(24.1)	1486	(15.9)	0.40 (0.36 to 0.44)	0.000
Private health institution	1905	(44.0)	2918	(31.1)	0.43 (0.40 to 0.46)	0.000
Other	6	(0.1)	22	(0.2)	1.02 (0.41 to 2.52)	0.967
Distance to health facilities						
Not a big problem	3885	(89.9)	8030	(85.6)	1	
Big problem	438	(10.1)	1354	(14.4)	1.50 (1.33 to 1.68)	0.000
Maternal healthcare decision-making						
By mother herself	1461	(34.7)	3297	(36.1)	1	
Jointly with husband	2193	(52.1)	4374	(47.9)	0.88 (0.82 to 0.96)	0.003
Husband alone	543	(12.9)	1429	(15.7)	1.17 (1.04 to 1.31)	0.010
By others	10	(0.3)	24	(0.3)	1.06 (0.51 to 2.23)	0.870
Child healthcare decision-making						
By mother herself	1469	(37.0)	3028	(36.0)	1	
Jointly with husband	2015	(50.8)	4240	(50.4)	1.12 (0.99 to 1.28)	0.076
Husband alone	424	(10.7)	983	(11.7)	1.02 (0.94 to 1.11)	0.621
By others	59	(1.5)	166	(1.9)	1.36 (1.01 to 1.85)	0.045

children ranged from 1.22 to 1.34. Similarly, children who were not first born had significantly higher chance of being unimmunised. The odds of being unimmunised increased as the child's age and birth order increased (p<0.000).

We found that children whose mothers were 30–39 years old at the time of the survey were less likely to be unimmunised (OR 0.67; 95% CI 0.50 to 0.90). However, there was no clear trend across the age groups. We also found that children who came from bigger families were significantly more likely to be unimmunised. The likelihood increased by 8% up to 42%. As the number of household members increased, the likelihood of a child to be unimmunised increased (p<0.000).

Although their marital status was not a significant predictor of coverage, each parent educational attainment was significantly associated with coverage. As parents' educational attainment increased, the likelihood of being unimmunised decreased (p<0.000). Hence, children from uneducated parents had the highest odds of being unimmunised. Those whose mothers had no education were at least six times more likely to be unimmunised (OR 6.14; CI 95% 4.41 to 8.53). Likewise, children whose fathers were uneducated had greater than fourfold chance of being unimmunised (OR 4.49; 95% CI 3.20 to 6.30).

Additionally, parents' occupation, mother's exposure to media and mother's tobacco use history were significantly associated with coverage. Across the occupational groups, children whose parents worked in agriculture had the highest odds of being unimmunised. Children whose mothers worked in agriculture were 2.6 times more likely to be unimmunised (OR 2.60; 95% CI 2.20 to 3.07), while children whose fathers worked in agriculture were 2.16 times more likely to be unimmunised (OR 2.16; 95% CI 1.89 to 2.47). Regarding mother's exposure to media, the child's likelihood of being unimmunised increased as the frequency of media exposure decreased (p<0.000). Finally, children whose mothers smoked tobacco around the time of the survey had 87% higher chance of being unimmunised (OR 1.87; 95% CI 1.47 to 2.37).

We found that as the household wealth index increased, the likelihood of being unimmunised decreased (p<0.000). Hence, children from poorest households had the highest odds of being unimmunised (OR 2.95; 95% CI 2.63 to 3.31). We also found that children who had no health insurance were significantly more likely to be unimmunised compared with those who had insurance (OR 1.38; 95% CI 1.29 to 1.49).

Our univariate analysis indicated that antenatal and postnatal care visits were significant predictors of

coverage in Indonesia. Our results showed that children who were born without antenatal care were at least five times more likely to be unimmunised (OR 5.29; 95% CI 3.78 to 7.39). Likewise, those who were born without postnatal care were 75% more likely to be unimmunised (OR 1.75; 95% CI 1.60 to 1.90).

In terms of access to health services, we found that children who were born in health institution were significantly less likely to be unimmunised compared with those who were born at home. Specifically, children who were born at public health institution had the least likelihood of being unimmunised (OR 0.40; 95% CI 0.36 to 0.44). In addition, children whose mothers think that distance to health facilities was a big problem had 50% higher chance of being unimmunised (OR 1.50; 95% CI 1.33 to 1.68).

Multivariate analysis

Out of the 22 independent variables, child's sex and mother's marital status were excluded. table 3 summarised the significant results of our multilevel logistic regression analysis between the remaining 20 independent variables and the likelihood of being unimmunised.

After accounting for the other remaining variables, geographic region and place of residence were significantly associated with coverage. The likelihood of being unimmunised was highest among children who lived in Maluku and Papua. Children who lived in this region were almost twice as likely to be unimmunised compared with those who lived in Java (AOR 1.94; 95% CI 1.42 to 2.64). Similarly, children who lived in Sumatera had considerably higher odds of being unimmunised (AOR 1.51; 95% CI 1.24 to 1.83). In contrast, children from Bali and Nusa Tenggara were less likely to be unimmunised (AOR 0.71; 95% CI 0.54 to 0.94). Those who lived in rural areas were also less likely to be unimmunised compared with their urban counterparts (AOR 0.82; 95% CI 0.69 to 0.96).

The likelihood of being unimmunised differed significantly across the age groups. Older children were more likely to be unimmunised compared with those in the youngest age group. The odds ranged from 1.24 (95% CI 1.08 to 1.42) to 1.39 (95% CI 1.20 to 1.60). Of all age groups, children aged 36–47 months had the highest odds of being unimmunised (AOR 1.39; 95% CI 1.20 to 1.60).

The child's birth order and family size were also significantly correlated with immunisation status. As a child's birth order or family size increased, the likelihood of being unimmunised also increased. A second child was 18% more likely to be unimmunised compared with a first child (AOR 1.18; 95% CI 1.03 to 1.35), while a fifth child had 68% higher chance of being unimmunised (AOR 1.68; 95% CI 1.28 to 2.19). Accordingly, children who came from bigger families had higher likelihood of being unimmunised. Those who lived in households with 10 or more family members were 47% more likely to be unimmunised (AOR 1.47; 95% CI 1.11 to 1.93).

Table 3 Multivariate analysis results for factors significantly associated with low immunisation coverage of children aged 12–59 months in Indonesia

Characteristics	Adjusted OR (95% CI)	p Value
External environment		
Geographic region		
Sumatera	1.51 (1.24 to 1.83)	0.000
Java	1	
Bali and Nusa Tenggara	0.71 (0.54 to 0.94)	0.016
Maluku and Papua	1.94 (1.42 to 2.64)	0.000
Place of residence		
Urban	1	
Rural	0.82 (0.69 to 0.96)	0.013
Predisposing characteristics		
Child's age (months)		
12–23	1	
24–35	1.24 (1.08 to 1.42)	0.002
36–47	1.39 (1.20 to 1.60)	0.000
48–59	1.36 (1.17 to 1.58)	0.000
Child's birth order		
1st	1	
2nd–4th	1.18 (1.03 to 1.35)	0.016
≥5th	1.68 (1.28 to 2.19)	0.000
Family size (number of household members)		
≤4	1	
≥10	1.47 (1.11 to 1.93)	0.006
Mother's educational level		
Higher	1	
No education	2.13 (1.22 to 3.72)	0.008
Father's occupation		
Professional	1	
Clerical, services and sales	0.82 (0.67 to 1.00)	0.047
Enabling resources		
Household wealth index		
Richest	1	
Poorer	1.30 (1.06 to 1.59)	0.011
Poorest	1.58 (1.26 to 1.99)	0.000
Covered by health insurance		
Yes	1	
No	1.16 (1.04 to 1.30)	0.010
Antenatal care		
Received some care	1	
Received no care	3.28 (2.09 to 5.15)	0.000
Postnatal care		
Received some care	1	
Received no care	1.50 (1.34 to 1.69)	0.000

Continued

Table 3 Continued

Characteristics	Adjusted OR (95% CI)	p Value
Child's place of delivery		
Home	1	
Public health institution	0.55 (0.47 to 0.64)	0.000
Private health institution	0.62 (0.54 to 0.72)	0.000
Maternal healthcare decision-making		
By herself	1	
Jointly with husband	0.86 (0.76 to 0.96)	0.010

Children whose mothers had no education were at least twice as likely to be unimmunised than those whose mothers were high school graduates or higher (AOR 2.13; 95% CI 1.22 to 3.72). Similarly, the odds of being unimmunised were significantly higher among the poorer (AOR 1.30; 95% CI 1.06 to 1.59) and the poorest (AOR 1.58; 95% CI 1.26 to 1.99). Also, those without health insurance were more likely to be unimmunised (AOR 1.16; 95% CI 1.04 to 1.30).

The odds of being unimmunised were strikingly higher among children without antenatal or postnatal care. Children who were born without antenatal care were more than three times as likely to be unimmunised (AOR 3.28; 95% CI 2.09 to 5.15). Likewise, those who had no postnatal care had a 50% higher chance of being unimmunised (AOR 1.50; 95% CI 1.34 to 1.69). Additionally, children who were born in health institution were less likely to be unimmunised compared with those who were born at home (AOR 0.55; 95% CI 0.47 to 0.64). Furthermore, children whose parents jointly decided on maternal healthcare and whose fathers worked in clerical, services and sales were significantly less likely to be unimmunised (AOR 0.86; 95% CI 0.76 to 0.96 and AOR 0.82; 95% CI 0.67 to 1.00, respectively).

DISCUSSION

Main findings

Our study investigated, for the first time, the factors associated with routine immunisation coverage of children aged 12–59 months in Indonesia, using data from 2012 IDHS. Our analysis revealed that only 31.5% of the children had been fully immunised. After accounting for all confounders, 13 factors were significantly associated with low coverage in Indonesia: geographic region, place of residence, child's age, child's birth order, family size, mother's education, father's occupation, household wealth index, insurance coverage, antenatal care, postnatal care, child's place of delivery and maternal healthcare decision-making.

There are discrepancies between the coverage level reported by the officials and the one discovered in this study. In 2012, the Indonesian MOH reported coverage level of 86.8%.[31] The coverage level determined through 2012 IDHS is therefore much lower than that contained in the official report.

While our study analysed cross-sectional survey data, the official report used administrative data which are commonly employed to assess immunisation coverage in low-resource settings.[32] The estimate is obtained by dividing the number of doses administered at health services by the expected target population.[32 33] Although this is readily available, results can be unreliable, particularly when there are uncertainties surrounding the total number of age-eligible children.[32 34]

The discrepancy between estimates obtained from administrative and survey data have also been reported in the past.[34–37] Administrative estimates tend to be higher than those obtained from the survey,[33] which is observed in our finding as well. Comparisons of administrative and survey estimates are made more complicated by the fact that the number of age-eligible children included in each analysis differs.[33] The estimate from administrative data includes children aged 0–11 months, while the survey usually includes children aged up to 59 months.[33 34] The coverage from MOH report was of children aged 0–11 months, because they are the youngest group eligible to receive the full schedule of routine immunisation. Measles vaccine, for example, is the last one on the schedule and is given starting at the age of 9 months. However, it could be administered up to the age of 12 months.[38] There are also booster campaign and backlog fighting initiative for children up to 3 years of age, as well as other supplemental immunisation activities which targeted children aged 9–59 months. This is all part of routine immunisation programme in Indonesia.[38] Therefore, estimates from administrative data would not have covered the entire target population of routine immunisation coverage. This indicates a weakness in the surveillance system and highlights the need of quality assurance of immunisation data.

Factors associated with immunisation coverage

After accounting for all observed confounders, geographic region was significantly associated with coverage. The six geographic regions used in our analysis represented the six largest islands in Indonesia. Each has its own population density, religious affiliation and political situation, economic potential and level of development. Our analysis suggested that children from the Maluku and Papua region had the highest odds of being unimmunised. The Maluku and Papua region is located in the easternmost part of Indonesia and is economically deprived. It is the largest yet least developed region with ongoing conflicts. Eligible children most likely lived in remote areas without access to health services. It is therefore not surprising that we found these children to have the highest likelihood of being unimmunised. Our research confirms that geographical disparities may contribute to low coverage, particularly in developing countries with a large population.[12] Similar findings were reported from India[38] and Nigeria.[14]

Children from urban areas have been reported to have better immunisation status compared with their rural counterparts.[30] By contrast, our results revealed that children who lived in rural areas were less likely to be unimmunised. Although health services are better and more easily accessible in urban areas compared with rural areas,[28] this fact likely masks the extent of urban poverty.[30] Estimates suggest that one-third of urban populations in developing countries are actually living in slums.[39] With limited access to health services and poor quality of life, it is certainly likely that urban children had higher odds of being unimmunised. Unfortunately, we lacked information to distinguish between urban areas with higher socioeconomic status and the slums. Further research in this field could assist strategic planning and resource allocation.

Our analysis revealed that children of older age groups were significantly more likely to be unimmunised compared with those in the youngest group. In other words, later birth years were associated with better coverage. It may indicate a positive trend of the immunisation programme performance over the years.[40] In the 5 years preceding the survey, the Indonesian government showed strong commitment towards immunisation programme. In line with global and national commitment to reduce the number of preventable child deaths, there were sharp increase in central government's budget for immunisation programme. Between the year of 2007 and 2008 alone, it increased by 40%.[41] In 2010, immunisation programme became a national priority under Presidential Instructions No 1 and No 3.[41] Among the key performance indicators was acceleration of coverage, which gradually increased between the year 2007 and 2012.[11 41] Our finding suggested that immunisation policy development in Indonesia might have played a role in improving coverage.

As the birth order increases, the likelihood of a child being unimmunised increases. A possible explanation is that parents may have developed confidence in their child's healthcare as a result of years of experience from previous children and could dismiss the importance of immunisation.[42 43] On the contrary, it could be that the first-born experienced adverse reaction to immunisation, leading the parents to believe that immunisation was risky.[43]

Consistently, children who came from larger families were more likely to be unimmunised. The number of household members has been linked with health outcome in many developing countries. As the number of family members increases, the quality of care they receive decreases.[28 42] This is because limited family resources are spread more sparsely, reducing the level of health investment received by each household member.

Our data revealed that children whose mothers had no education were at least twice as likely to be unimmunised compared with those whose mothers were high school graduates. This indicates that maternal education is a major determinant of immunisation coverage in Indonesia. The obvious explanation is that literacy and educational attainment facilitate understanding of the recommended immunisation schedule.[40] This suggests that improving the programme to achieve the target of herd immunity might be helpful only in the short term. It highlights the need for a long-term investment in human capital, especially in Indonesian women.[28]

Children whose fathers work in clerical, services or sales were less likely to be unimmunised compared with children of professionals. This is unexpected, given that people who work in clerical, services or sales are usually of a lower socioeconomic status and may find it difficult to obtain permission for work leave to enable their children to be immunised.[14] Nonetheless, our result confirmed previous finding which reported similar association in Bangladesh.[16] Fathers who were professionals were significantly less likely to have their children fully immunised, as they tend to work long hours and are too preoccupied to be involved in their child's healthcare.

Wealth is a well-established indicator of access to health services in many countries regardless of income groups. Our analysis indicated that children from poorer and poorest households were more likely to be unimmunised. Given that immunisation services are available free of charge in Indonesia, the indirect cost of immunisation may be the relevant factor instead. Lost work days and transport costs could deter parents from enabling their child to be immunised.[44 45] The likelihood of being unimmunised was also higher among children without health insurance. This is reasonable because health insurance alleviate the burden of out-of-pocket spending, including indirect cost of immunisation. Most studies from developing countries have reported that health insurance has a positive impact on increasing healthcare use.[46]

The odds of being unimmunised were considerably higher among children without antenatal and postnatal care. Children who were born without antenatal care were at least three times more likely to be unimmunised. Likewise, children who did not receive postnatal care had a 50% greater chance of being unimmunised (AOR 1.50; 95% CI 1.34 to 1.69). This finding reflects the importance of information received by mothers during antenatal and postnatal care. Their visits might have equipped them with the necessary knowledge on child immunisation. In Indonesia, at least four antenatal visits are recommended during pregnancy. However, this service has been underused[29] and the negative implication of missed opportunities for immunisation coverage is almost certain.

There was a significant association between a child's place of delivery and immunisation coverage. Children who were born in public or private health institution were less likely to be unimmunised compared with those who were born at home. This is most likely because children who were born at health facilities were vaccinated, or were given recommendation to be vaccinated, immediately after birth. Furthermore, a study from Kenya has shown that women who deliver at home or unassisted may have

a distrust of modern medicine and a stronger preference for traditional remedies.[47] By extension, they could have a sceptical view about childhood immunisation.[48]

Our analysis also showed that children who were born in private health institution had greater odds of being unimmunised relative to those who were born in public health institution (AOR 0.62; 95% CI 0.54 to 0.72 and AOR 0.55; 95% CI 0.47 to 0.64, respectively). In Indonesia, private health institution do not benefit from government's healthcare funding, although they do operate under the ministerial decree to deliver routine immunisation. Consequently, there is no financial incentive for private health institution to ensure that children are fully immunised. Therefore, strengthening the implementation of the ministerial decree for private health institution may help in improving immunisation coverage.

Children whose parents jointly decide on maternal healthcare were less likely to be unimmunised. This emphasises the importance of family support in using health services, confirming what had been outlined by Andersen in his theoretical framework.[13] The combination of both mother's autonomy and father's involvement in the decision-making process seemed to be essential. This suggests that interventions which educate and involve fathers might have the potential to increase immunisation coverage.[49]

Although our findings were consistent with reports from other lower middle income countries, we found that several factors were not significant predictors of coverage in Indonesia. Despite reports from India, a child's sex did not affect coverage in Indonesia. This is consistent with studies from Nigeria undertaken by Antai[14] and Adebiyi.[50] It appears that gender could predict immunisation status only if the child is from a society where gender inequality is prevalent.[50] We also found no correlation between a mother's age and her child's immunisation status. Previous studies have reported that the odds of a child being unimmunised is greater for both younger and older mothers, suggesting a U-shaped association.[28] However, this association might be mitigated by patterns of other coexisting variables in our analysis, such as the child's birth order and the mother's level of education.

Strengths and limitations

To our knowledge, this study was the first to identify factors associated with routine immunisation coverage of children in Indonesia. We used the 2012 IDHS dataset, which was the most recent one. The large sample size allowed us to analyse many potential predictors simultaneously. It also increased the validity of our results. Furthermore, we used multilevel modelling to account for the hierarchical structure of the data. We have also adjusted our analysis to meet the local context and produce reliable estimates. However, our results should be considered in the light of potential limitations.

As with other secondary analysis of cross-sectional survey data, caution should be exercised in inferring causality between the socioeconomic factors and immunisation coverage. In addition, the nature of our data source and analysis potentially limit generalisability. There is a need to verify the validity of the observed associations using longitudinal data.

Information on a child's immunisation status was subject to bias, because we included mother's report as a source of information. As such, we relied on the mother's ability to recall her child's immunisation status accurately. Nonetheless, mother's report is considered a valid measure of coverage in the absence of a health card, especially in developing countries.[51] We therefore believe that our reliance on mother's report is reasonable and not likely to have introduced bias into our study.

The selection of variables included in this study relied on the information available from the dataset. Other potential predictors that were previously identified in lower middle-income setting, such as ethnicity and religion, could not be assessed in this study. Categorisation of original responses from the survey might have also influenced the results.

The 2012 IDHS selected participants through a two-stage stratified sampling design. The primary sampling unit was the CBs and the complete list of households in each CB became the basis for second-stage sampling. However, there was no household identifier in the dataset as it may compromise the participants' anonymity. Therefore, we could only build a two-level model (ie, children nested within CBs) instead of a three-level model (ie, children within households nested within CBs). We recognise that children living in the same household could have shared similar health characteristics, which reflects parent-specific knowledge or beliefs on immunisation.[12] However, our analysis of variables that served as a proxy of parent-specific knowledge or beliefs (ie, mother's exposure to media and mother's tobacco use history) emerged as being insignificant. Therefore, we have good reason to believe that this limitation is unlikely to have any impact on the validity of our analysis.

Finally, we classified immunisation status into 'fully immunised' and 'unimmunised' based on whether the child received full schedule of immunisation or otherwise. While other studies have used three distinct categories: fully immunised, partly immunised and completely unimmunised, we dichotomised our outcome variable and did not distinguish partly immunised from completely unimmunised. This is because our study focused on factors associated with the coverage of routine immunisation, which is the complete uptake of recommended vaccination represented by the fully immunised. Reasons for Indonesian children being partly immunised and completely unimmunised might differ, and future research can potentially address this question.

CONCLUSION

In this study, we examined variables that contribute to a child's immunisation status in Indonesia. Our results suggested that immunisation coverage is suboptimal

due to socioeconomic factors. Among the demographic groups, children who lived in Maluku and Papua region and children from the poorest households have the lowest coverage. We also identified maternal education and antenatal care visits as key factors that policy-makers can target to improve immunisation coverage in Indonesia.

Beyond mapping trend of coverage nationally, we recommend regular monitoring and evaluation of coverage at province and district levels. This is important to identify high-risk areas and implement targeted activities in the communities. Increasing awareness and financial support for deprived households with more than one child may help reduce the indirect cost and motivate parents to immunise their children. Promoting equal access to education, encouraging institutional deliveries and scaling up use of antenatal and postnatal care may significantly improve coverage in Indonesia.

Acknowledgements We are grateful to the ICF International for granting us access to the datasets and to the Indonesia Endowment Fund for Education (*Lembaga Pengelola Dana Pendidikan*) for funding PH a master scholarship at the Department of Primary Care and Public Health Sciences, King's College London. This analysis was part of PH dissertation.

Contributors PH and AD participated in the design of the study. PH performed the analysis and prepared the manuscript. AD provided data analysis advice and revision of the manuscript. All authors read and approved the final manuscript.

Funding This study was funded by Indonesia Endowment Fund for Education.

Competing interests PH had financial support from LPDP for the submitted work, no financial relationships with any organisations that might have an interest in the submitted work in the previous three years; no other relationships or activities that could appear to have influenced the submitted work.

Ethics approval This study did not require ethical approval as it used unidentifiable secondary data. Permission to use the dataset was obtained from ICF International, who obtained approval to conduct IDHS in 2012. No identifiable information was included in the dataset and no attempt was made to identify any individual interviewed in the survey.

Provenance and peer review Not commissioned; externally peer reviewed.

Data sharing statement The electronic datasets analysed in this study are available for legitimate research purposes from the Measure DHS website.

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