



Review article

Assessing the efficacy of Alpha₁-Proteinase inhibitor (human) augmentation therapy for Alpha₁-Antitrypsin deficiency – Related emphysema: Challenges and opportunities

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ARTICLE INFO

Keywords:

Alpha₁-antitrypsin
 Alpha₁-antitrypsin deficiency
 Alpha₁-proteinase inhibitor
 Emphysema
 Systematic review
 Lung density
 Spirometry
 Exacerbations

ABSTRACT

Clinical benefit of Alpha₁-Proteinase Inhibitor (Human) (A₁-PI) products for Alpha₁-antitrypsin deficiency (AATD) is uncertain, based on a systematic review of observational studies and randomized controlled trials (RCTs) in AATD of Alpha₁-Proteinase Inhibitor (Human) (A₁-PI) products. At the recommended dose, A₁-PI products raise its serum concentration but do not normalize levels. Observational studies suggest A₁-PI might modestly slow progression of airflow limitation in patients with intermediate airflow obstruction, a finding not confirmed by three placebo-controlled RCTs of limited power, which showed non-significant rates of forced expiratory volume in 1 s (FEV₁) change favoring placebo. These RCTs found trends favoring A₁-PI in loss of high-resolution computerized tomographic (HRCT) lung density. While two meta-analyses of HRCT lung density change in RCTs achieved significance favoring A₁-PI arms, clinical benefit remains uncertain. HRCT lung density measurements don't distinguish changes in measured density due to fluid shifts into and out of the lungs and changes in lung inflammation from those due to progressive loss of alveolar mass. A meta-analysis of RCTs found exacerbations significantly increased in A₁-PI groups compared to placebo. No RCTs have shown favorable effects of A₁-PI on mortality, FEV₁, 6-min walking distance, quality of life, change in diffusion capacity of carbon monoxide (DL_{CO}), or exacerbation frequency. A fourth RCT comparing two dose regimens of A₁-PI is underway. RCTs have not provided evidence of clinical benefit in terms of how patients feel, function, or survive. Results have implications for the design of future clinical trials of A₁-PI and potentially other products targeting AATD-associated emphysema.

1. Introduction

The objective of this article is to provide a systematic review of observational studies and randomized controlled trials that have attempted to evaluate the efficacy of Alpha₁-Proteinase Inhibitor (Human) augmentation Therapy for Alpha₁-Antitrypsin Deficiency – Related Emphysema (AATD-RE), and to discuss their implications for the design of future studies. Alpha₁-antitrypsin deficiency (AATD) is an autosomal codominant genetic disorder caused by mutations in the gene coding for alpha₁-proteinase inhibitor (A₁-PI)/alpha₁-antitrypsin (AAT). A₁-PI, among other activities, inhibits lung neutrophil elastase (NE), and helps protect lung parenchyma from proteolytic destruction [1–6]. The most common manifestations of severe AATD are emphysema and liver disease [1,7]. AATD phenotypes, corresponding serum levels and comparative risk of emphysema are listed in Table 1. Normal (Pi*MM) serum AAT ranges

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<https://doi.org/10.1016/j.heliyon.2024.e31183>

Received 2 October 2023; Received in revised form 4 May 2024; Accepted 12 May 2024

Available online 23 May 2024

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from ~20 to 53 μM [8]. Severe AATD is defined by having a serum antigenic A₁-PI level <11 μM . Approximately two-thirds of never-smokers and ~90 % of ever-smokers with severe AATD eventually develop emphysema [7,9].

AATD-RE is characterized by dyspnea, reduced exercise tolerance/capacity, airflow limitation, and acute exacerbations of COPD. In patients with AATD-RE, progressive loss of alveoli can lead to respiratory failure and death or the need for lung transplantation to sustain life.

1.1. Augmentation therapy for AATD-related lung disease

The only therapeutic agents approved for AATD-RE are Alpha₁-Proteinase Inhibitor (A₁-PI) (Human) products, manufactured from pools of human plasma [17–20]. Therapy with A₁-PI is referred to as “augmentation therapy,” because it augments the blood and lung levels of A₁-PI [2,21]. At the currently recommended weekly dose, the steady-state trough level does not achieve normal levels [22].

The US Food and Drug Administration (FDA) licensed the first A₁-PI product in 1987 based on a demonstration of A₁-PI-induced changes in biochemical biomarker endpoints, prior to the formal introduction in 1992 of an accelerated approval pathway based on a surrogate endpoint. Manufacturers of US-licensed A₁-PI augmentation therapy products have entered into postmarketing commitments (PMCs) [23] to conduct randomized, controlled clinical trials (RCTs) to evaluate potential long-term benefit using clinically-meaningful endpoints [22,24,25]. Data suggest that the current A₁-PI dose regimen may be suboptimal [26–31]. Manufacturers have agreed to study a higher dose to ascertain if it is associated with [improved] efficacy and safety. No PMCs of A₁-PI products have been completely fulfilled, and only one is currently in progress. PMC trials may be informative regarding the relationships between A₁-PI serum levels and emphysema progression, plus exacerbation risk in patients receiving A₁-PI products at currently recommended and higher doses.

The first A₁-PI approval was based on demonstrating (a) a rise from pre-therapy baseline and maintenance of serum A₁-PI trough level to > 11 mM and (b) a rise from baseline in lung epithelial lining fluid (ELF) A₁-PI concentration calculated from broncho-alveolar lavage (BAL) fluid sampling. The history of the hypothetical therapeutic target serum level of 11 mM A₁-PI is discussed in online Supplement A to this article.

Clinical benefit has yet to be demonstrated for A₁-PI products. Their effects, if any, on symptoms, function/exercise tolerance, the incidence, severity, and/or duration of exacerbations, and/or the progression of emphysema or airways obstruction have not been demonstrated in adequate and well-controlled clinical trial(s) [17,32–35]. The current FDA-approved indication for A₁-PI products is listed in Table S1.

1.2. Pharmacokinetics of A₁-PI

Population pharmacokinetic parameters for antigenic serum A₁-PI during chronic dosing with Zemaira brand A₁-PI 60 mg/kg IV weekly, as reported by Tortorici et al., included clearance of 45.2 mL/day, volume of distribution of 10.0 L, half-life of 6.8 days, C_{avg} of 20.9 μM (90 % CI 14.8–26.7), C_{max} of 28.5 μM (90 % CI 13.3–44.7) and C_{trough} of 16.2 μM (90 % CI 11.1–22.6) [36].

There is substantial inter-subject variability in A₁-PI trough levels in persons receiving augmentation therapy at the recommended weekly dose of 60 mg/kg [22]. Published literature and data contained in FDA-approved package inserts of A₁-PI products indicate that most patients receiving regular weekly 60 mg/kg infusions of licensed A₁-PI products do not achieve normal serum levels of A₁-PI and may continue to experience progression of COPD and exacerbations [17,32–34,18,37–39]. Additional information on the pharmacokinetics of A₁-PI is provided in online supplement B.

1.3. Natural history studies of AATD

Several authors have examined the relationship between baseline FEV₁ and subsequent rate of change of FEV₁ in patients with AATD not receiving A₁-PI [40–43]. The fastest deterioration of airflow obstruction tends to occur in patients with intermediate degrees of airflow limitation at baseline, such as those with baseline FEV₁ approximately 30–65 % of predicted [40–43]. Among usual, AAT-replete COPD patients, those with very severe airflow limitation at baseline as a group have slower subsequent decline in FEV₁

Table 1
Normal and Selected AATD Phenotypes and their Serum A₁-PI Levels [1,8,10–13].

Phenotype	Risk of Emphysema	Serum Antigenic A ₁ -PI Levels (μM)
PI*ZZ	Very High	2–7
PI*SZ	Intermediate ^a	9–23
PI*MZ	Low but greater than normal ^b	17–33
PI*MM (Normal)	Normal (Low)	20–53

The nomenclature for AATD phenotypes relates to the results of isoelectric focusing of patients’ serum. The Z, S, and M allelic variants each have a different isoelectric point, giving rise to bands of different mobility. “PI” stands for proteinase inhibitor.

^a The odds ratio for risk of emphysema for PI*SZ has been estimated in a meta-analysis to be 3.26-fold compared to that of normal PI*MM individuals (95 % CI 1.24–8.57) [14].

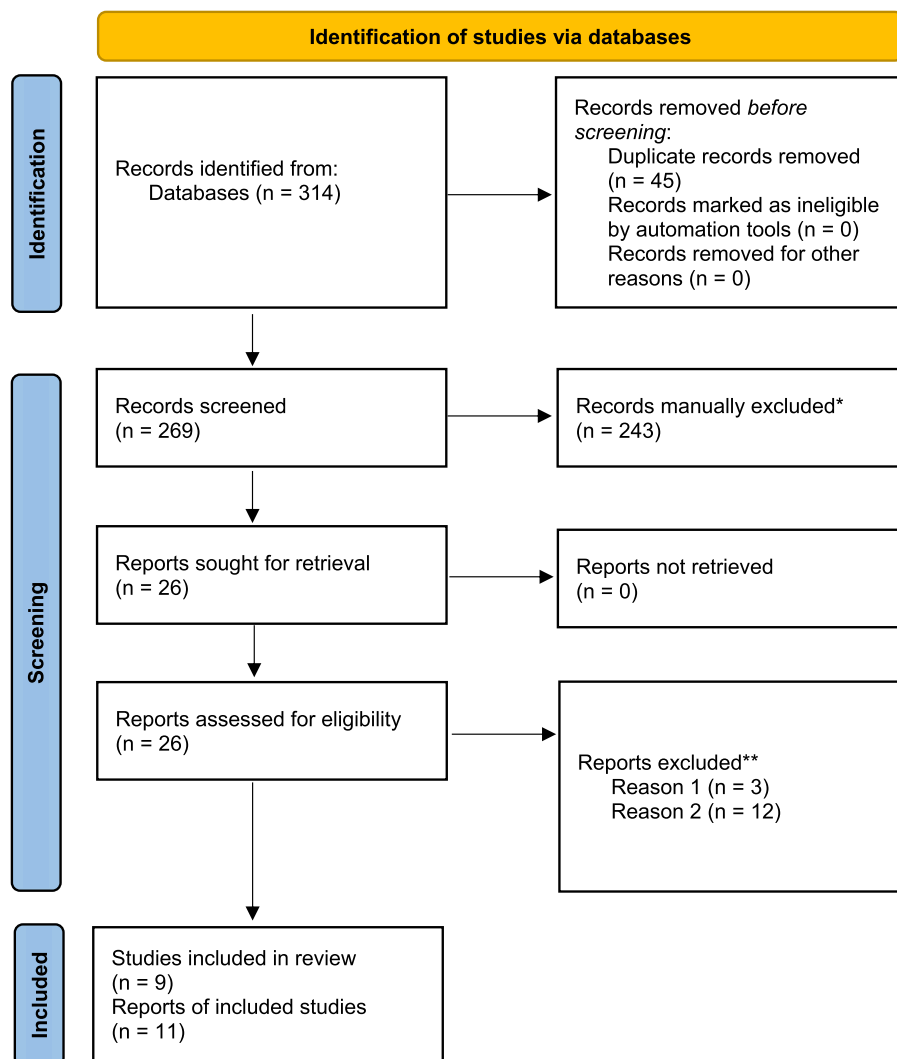
^b The risk of hospitalization for COPD in Denmark for PI*MZ individuals was found to be 2.2-fold greater than that of PI*MM individuals [8,15]. The rate of decline in DL_{CO} among PI*MZ individuals also has been found to be more rapid than for PI*MM individuals [16].

than patients with less severe FEV₁ impairment [44,45]. These data may be relevant when considering possible enrichment strategies when crafting subject inclusion criteria in trials using FEV₁ as a primary or co-primary endpoint.

2. Methods

Drs. L. Ross Pierce and Gavin Imperato, with the assistance of Ms. Gwendolun Halford, conducted literature searches in 2022 comprising each of the bullets below using PubMed, Embase, Cochrane Central, and [ClinicalTrials.gov](https://www.clinicaltrials.gov) and performed a systematic review using the PRISMA 2020 Checklist [46] as follows.

A. Prisma Flow Diagram for Search 1 (Observational Studies)



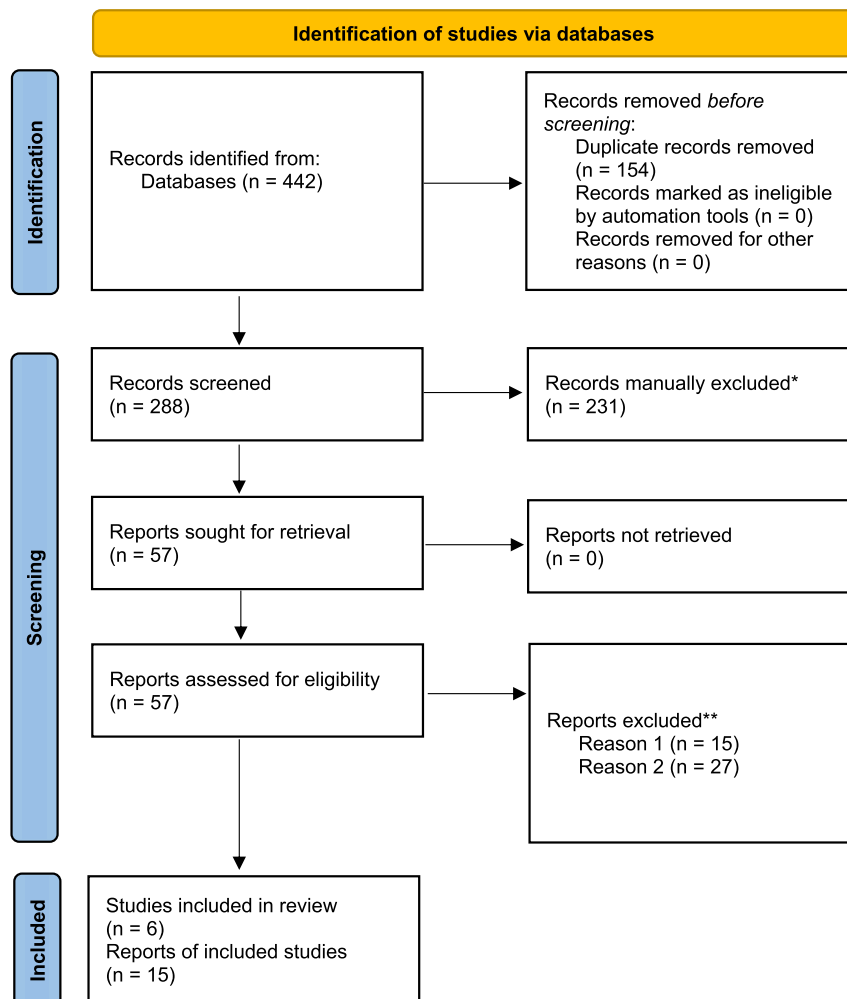
*No automation tools were used for excluding records deemed irrelevant to the screening topic.

**Reason 1 for excluding reports: Study randomized, not observational.

Reason 2 for excluding reports: Otherwise not meeting criteria.

Fig. 1. Prisma Flow Diagram Describing Process for Selection of Studies Included in this Review A. Prisma Flow Diagram for Search 1 (Observational Studies).

B. Prisma Flow Diagram for Search 2 (Randomized Controlled Trials)



*No automation tools were used for excluding records deemed irrelevant to the screening topic.

**Reason 1 for excluding reports: Study randomized, but not placebo – or dose – controlled.

Reason 2 for excluding reports: Study not randomized.

Fig. 1. (continued).

- Observational Studies comparing AATD patients who received vs. those who did not receive Alpha₁-Proteinase Inhibitor augmentation therapy from the year 1987 to present, English language only. Studies comprising fewer than 50 subjects or over an observation period of <1 year were excluded.
- Randomized controlled trials involving patients with AATD and lung disease/COPD/emphysema and Alpha₁-Proteinase Inhibitor augmentation therapy from 1987 to present, all languages. Studies comparing different manufacturers' A₁-PI products and not including a placebo group or dose comparison were excluded.

Papers from the searches were screened and relevant articles reviewed independently by Drs. Pierce and Imperato and studies that met the above criteria were included in the review/discussion. All results compatible with each output domain were sought. Relevant data from each included report was directly transcribed by Dr. Pierce into this review. Included studies were assessed for sources of bias, including baseline imbalances in covariates that could affect outcome measures, the potential for regression to the mean, potential differences in standard of care between comparison groups, and other potential confounding factors. The range of the magnitudes of potential bias among the studies comprising this review is large. The review was not registered.

High Level Specific literature search Strategy Terms and Boolean operators were as follows:

Observational Studies Search:

[efficacy OR effect OR effectiveness OR lung function decline OR lung density decline OR diffusion capacity of carbon monoxide decline OR exercise capacity decline OR FEV₁ decline OR mortality OR survival OR [longitudinal] follow-up] AND [Alpha1-Proteinase Inhibitor OR Alpha1-Antitrypsin OR augmentation therapy] AND alpha1-antitrypsin deficiency AND [COPD OR emphysema or lung disease] NOT cross-sectional, NOT animal.

Randomized Controlled Trials Search:

Randomized Controlled Trial AND [Alpha1-Proteinase Inhibitor (Human) OR Alpha1-Antitrypsin] AND Alpha1-Antitrypsin Deficiency.

The complete list of search terms and Boolean Operators for the above searches is given in [Supplement E](#).

3. Results

We identified nine observational studies meeting protocol/search criteria comparing AATD patients who received vs. those who did not receive Alpha₁-Proteinase Inhibitor augmentation therapy.

We identified six randomized placebo – and/or dose – comparison controlled trials involving patients with AATD and lung disease/ COPD/emphysema and Alpha₁-Proteinase Inhibitor augmentation therapy. Three randomized placebo – controlled trials were completed and one is ongoing. Two trials used a randomized crossover design to compare two dosage regimens.

Prisma flow diagrams describing the study selection process for searches 1 (observational studies) and 2 (RCTs) are presented in [Fig. 1](#).

4. Non-interventional epidemiology studies of A₁-PI augmentation therapy

4.1. NHLBI registry study

The NHLBI epidemiologic registry study explored the possible effects of augmentation therapy with ProLasta brand A₁-PI over a period of 3.5–7 years on mortality and FEV₁ rate of change in severe AATD [39,47]. The pre-specified primary endpoints of the NHLBI Registry study were comparisons between participants who received augmentation therapy with A₁-PI versus those who did not receive A₁-PI of (1) mortality among registry participants with baseline FEV₁ below 30 %, and (2) FEV₁ rate of change among registry participants whose baseline FEV₁ was greater than 30 % [47]. *The publication reporting the results of the NHLBI registry does not provide the results of either the mortality or the FEV₁ primary endpoints for the prespecified primary analysis subgroups of participants with baseline FEV₁ below or above 30 % of predicted, respectively.* Rather, the results of *post-hoc* analyses using different FEV₁ strata are reported. Inspection of the results for the reported FEV₁ strata suggest that both prespecified primary endpoints likely failed [39].

As shown in [Table 2](#), multivariate survival analysis of NHLBI Registry participants who received versus those who did not receive A₁-PI therapy did not demonstrate a significant survival advantage with A₁PI among patients whose FEV₁ was below 35 % of predicted. However, multivariate survival analysis of the entire registry cohort did suggest a survival advantage with augmentation therapy (RR 0.64, 95 % CI 0.43–0.94, $p = 0.02$ (unadjusted for multiple *post-hoc* comparisons)). This result was driven by the participants with an intermediate degree of airway obstruction (FEV₁ 35 %–49 %) [39].

As shown in [Table 3](#), multivariate analysis of FEV₁ rate of change in the entire NHLBI Registry cohort of participants who received versus those who did not receive A₁-PI therapy did not find a statistically significant advantage with augmentation therapy (difference in rates of change in FEV₁ = 4.2 mL/year, 95 % CI -5.7 – 14.2, $p = 0.40$, unadjusted). Nor did multivariate analysis of FEV₁ rate of change demonstrate a statistically significant difference with A₁-PI among patients whose FEV₁ was 35 %–79 % of predicted (difference = 13.6 mL/year, 95 % CI -4.1 – 31.1, $p = 0.13$, unadjusted). However, multivariate *post-hoc* analysis of the difference in rates of change in FEV₁ among patients whose FEV₁ was 35 %–49 % in participants who had received versus those who had not received A₁-PI therapy suggested the possibility of benefit with augmentation therapy (difference = 26.8 mL/year, 95 % CI 2.8–50.9, $p = 0.03$,

Table 2

Multivariate survival analysis^a showing relative risk of death by mean FEV₁ strata in NHLBI registry prospective epidemiologic study [39].

Mean FEV ₁ % Predicted ^b	Number of Subjects	Relative Risk of Death ^c (Therapy vs. Control)	RR 95 % CI ^c	P Value ^d
<35 %	482	0.83	0.52, 1.33	0.44
35–49 %	217	0.21	0.09, 0.50	<0.001
>50 %	349	0.75	0.22, 2.56	0.64
Overall	1048	0.64	0.43, 0.94	0.02

^a Multivariate Survival Analysis included receipt of a lung transplant treated as a time-varying covariate. The baseline or first available post-bronchodilator measurement of FEV₁% predicted was used as a covariate in the survival models, using the following staging strata: FEV₁% predicted <35 % [Stage III], 35–49 % [Stage II], 50–79 % [Stage I], and ≥ 80 % [Normal]). Only subjects who were successfully contacted 6 or more months following enrollment were included in the survival analysis.

^b Mean FEV₁ predicted over the duration of each subjects' trial participation. Analyses using baseline FEV₁ to define strata yielded similar results, according to the publication.

^c It is possible that these point estimates for relative risk of death and corresponding confidence intervals overstate, or understate, the true mortality effect (if any) of augmentation therapy at the currently recommended dose because of differences in baseline factors (unaccounted for) between augmented and non-augmented patients other than their augmentation therapy status.

^d Reported p values are not adjusted for multiple comparisons, so they may be considered to lack validity.

Table 3

Multivariate analysis^a of FEV₁ decline: Mean FEV₁ decline (mL/yr) by FEV₁% predicted and augmentation therapy status in NHLBI registry prospective epidemiologic study [39].

Mean FEV ₁ % Predicted ^b	Difference ^c in FEV ₁ Slopes: Receiving vs. Not Receiving A ₁ -PI ^c	95 % CI ^d	P Value ^e
< 35 %	2.6	−111.3, 16.5	0.71
35–49 %	26.8	2.8, 50.9	0.03
30–64 %	18	2, 34	0.03
50–79 %	7.5	−14.7, 29.6	0.50
≥80 %	−23.8	−50.9, 3.3	0.09
35–79 %	13.6	−4.1, 31.1	0.13
Overall	4.2	−5.7, 14.2	0.40

^a Covariates in the multivariate analysis linear mixed effects model included mean FEV₁% predicted, calculated from all available visits, bronchodilator responsiveness, coded as whether the subject ever versus never had a bronchodilator response (defined as postbronchodilator increase in FEV₁ of at least 200 mL and 12 % over the prebronchodilator value at any visit, and the cumulative time (since enrollment) for which each subject had received augmentation therapy at each follow-up visit (as a time-dependent covariate).

^b Mean FEV₁ predicted over the duration of each subjects' trial participation. Analyses using baseline FEV₁ to define strata yielded similar results, according to the publication. For example, the difference in mean FEV₁ rates of decline with and without augmentation therapy for the subgroup with baseline FEV₁ 35–49 % of predicted was 22 mL/yr, $p = 0.04$.

^c A positive difference in slopes implies a slower rate of decline for subjects receiving A₁-PI augmentation therapy compared with those not receiving augmentation therapy.

^d It is possible that these point estimates and corresponding confidence intervals overstate, or even understate, the true effect (if any) of augmentation therapy at the currently recommended dose on FEV₁ rate of change because of differences in baseline factors (unaccounted for) between augmented and non-augmented patients other than their augmentation therapy status.

^e Reported p values are not adjusted for multiple comparisons, so they may be considered to lack validity.

unadjusted for multiple *post-hoc* comparisons) [39]. A similar finding was obtained in an analysis of registry patients whose average FEV₁ was 30–64 % of predicted (difference in rates of change in FEV₁ = 18 mL/year, 95 % CI 2–34, $p = 0.03$, unadjusted) [39]. As *post-hoc analyses, these results in subgroups with intermediate degrees of airflow limitation should be considered for hypothesis generation only*;

Table 4

Differences in FEV₁ slopes between Patients Receiving and Not Receiving A₁-PI in Non-Randomized Observational Studies.

Study, Year of Publication/ Reference No./Design	Overall n (A ₁ -PI)/n (no A ₁ - PI)/Duration (A ₁ -PI)/ Duration (no A ₁ -PI) ^a (years)	Overall Study Δ (FEV ₁) mL/year \pm SD or % predicted	Overall 95 % CI (mL/ year) or p Value ^b for FEV ₁ Difference	Subgroup FEV ₁ 30–65 % Δ (FEV ₁) mL/ year \pm SD	Subgroup FEV ₁ 30–65 % 95 % CI (mL/year) or p Value ^b
NHLBI Registry ^c 1998/Ref 28/ Prospective	581/317/3.5–7 yrs	4.2 mL/year	−4.9–13.3	13.6	−3.3–30.5
Seersholm 1995/Ref 31/ Retrospective	198/97/3.2 years/5.8 years (means)	22 mL/year	$p = 0.02$	21	$p = 0.04$
Wencker 2001 /Ref 38/Retrospective	96/4.0 years/4.2 years (means)	14.6 \pm 61.4 mL/year	$p = 0.019$	11.6 \pm 48.8	$p = 0.066$
Tonelli 2009/Ref 29 Unstated (Presumably Retrospective)	124/40 + 26 ^d /3.5 years (mean, combined)	43.9 mL/year	$p = 0.046$	54.0	$p = 0.07$
Schouten 2021/39/ Retrospective	128/246 /8.6/8.6 years (mean)	−0.085 % predicted	−1.144 to 0.717 % predicted $p = 0.71$	NR ^e	NR
National German Registry 2017/40 Unstated (Presumably Retrospective)	85/15 /4.9 years (mean, combined)	33.8 mL/year	NS ^f	NR	NR
Spanish National Database 2018/41/Retrospective	77/45 /8 years (mean, combined)	NR	$p = 0.675$	NR	NR

^a Durations in years are provided for studies in which different cohorts of patients receiving and not receiving A₁-PI augmentation therapy are compared or the same cohort is compared during periods in which the patients were not receiving, then receiving augmentation therapy.

^b 95 % confidence intervals (CI) for the difference in FEV₁ rate of change between patients receiving and not receiving augmentation therapy (slope of those receiving minus slope of those not receiving A₁-PI) are provided where reported; otherwise, nominal p values as reported, which have not been corrected for multiple comparisons are given. The latter are not considered valid, due to lack of correction for multiple comparisons.

^c Values for the NHLBI Registry Study are taken from Ref 26 (Chapman meta-analysis) which reports the difference in FEV₁ slopes for the stratum of patients with baseline FEV₁ from 30–65 % of predicted.

^d Results are reported for the combined original cohort of 124 patients on A₁-PI and 40 patients not on A₁-PI, plus an additional 26 patients who were not included in the original analysis because their smoking status was initially missing (but later located). The breakdown of these additional 26 patients by treatment status was not provided in the publication.

^e NR = not reported.

^f NS = not statistically significant ($p \geq 0.05$).

the pre-specified co-primary efficacy endpoint for FEV₁ was unlikely to have been met among participants whose baseline FEV₁ was >30 %, and augmentation therapy did not appear to have any effect in the overall study population.

Among NHLBI Registry patients who received augmentation therapy, only 51 % initially received the currently recommended 60 mg/kg weekly dosage regimen. This dropped to 33 % at last report. Many subjects received a double dose every two weeks or a quadruple dose every four weeks, resulting in higher peak and lower trough A₁-PI levels [39].

As a non-randomized trial, the investigators could not exclude the possibility that any observed differences may have been due to factors other than augmentation therapy. They noted that “a definitive conclusion will require a randomized controlled trial.” [39].

Re-analysis of the mortality data from the NHLBI Registry study, published by Rahaghi et al. [48], corrected database errors and included additional mortality status ascertainment. The original analysis plan described in the 1998 NHLBI Registry publication was altered in the Rahaghi analysis, in that patients who did not have contact with the study six months or more following enrollment were included in the Rahaghi analysis, unlike in the 1998 publication. The Rahaghi re-analysis states that the original statistical plan to exclude these participants “will always stand as the primary analysis of mortality since it was *pre-hoc*.” The re-analysis found significantly lower mortality among subjects ever on augmentation therapy compared to those never on A₁-PI therapy within each decile of baseline FEV₁% predicted between 10 % and 60 % [48].

FEV₁ results of the NHLBI Registry Study are compared with those of smaller observational studies comparing patients receiving vs. not receiving A₁-PI in Table 4.

4.2. Seersholm 1995 study

A retrospective epidemiological study compared the rates of decline in FEV₁ between 198 ex-smoker PI*ZZ patients in Germany who received Prolastin brand A₁-PI over a mean of 3.2 years and 97 ex-smoker Danish PI*ZZ patients who did not receive A₁-PI and were followed over a mean of 5.8 years [42]. Patients who had at least two FEV₁ measurements obtained at least one year apart were included in the analysis. Attempts were made to adjust for differences between cohorts in the proportions of subjects of each sex, follow-up time, and initial FEV₁. Overall, patients in Germany who received Prolastin had a significantly slower rate of decline in FEV₁ as compared to patients in Denmark who did not receive A₁-PI (means of 53 ± 38 vs. 75 ± 60 mL/year, respectively, difference 22 mL/year, p = 0.02). This difference was largely accounted for by patients whose baseline FEV₁ was between 30 and 65 % of predicted, in whom the mean rates of decline in FEV₁ were 62 ± 25 vs. 83 ± 49 mL/year (difference 21 mL/year, p = 0.04, unadjusted) for the augmented versus unaugmented patients, respectively. Differences in rates of FEV₁ decline between patients in Germany and Denmark whose FEV₁ baseline values were <30 or >65 % of predicted were not significant (p = 0.6 and 0.7, respectively). German patients whose baseline FEV₁ was >65 % of predicted were required to demonstrate a rate of decline in FEV₁ of at least 120 mL/year, but this requirement was not listed for the Danish cohort. Thus, these data are subject to regression to the mean for the German cohort, which could have biased the results. Differences in standard-of-care between the two cohorts in different countries may have also confounded the results.

4.3. Wencker 2001 study

Wencker et al. published a retrospective observational study comparing the rates of change of FEV₁ before and after weekly administration of Prolastin brand A₁-PI at 60 mg/kg in 96 severe AATD subjects in Germany (62 males and 34 females) of mean age 44.3 years and baseline FEV₁ < 65 % or, if baseline FEV₁ > 65 %, who had a documented rate of decline in FEV₁ of >120 mL/yr [49]. The study included overlapping data of German subjects who received A₁-PI in the Seersholm study [42] and was limited by its retrospective nature, lack of randomization, lack of a concurrent control group, lack of parallel design, lack of blinding, and expected regression to the mean in the subgroup of patients with baseline FEV₁ > 65 % of predicted who were selected for analysis based on having a pre-augmentation therapy observed rate change in FEV₁ of >120 mL/yr. Subjects were followed for a mean of 50.7 (SD 34.3) months prior to augmentation therapy and for 35.7 (SD 23.3) months during augmentation therapy.

Table 5

FEV₁ slopes (mL/year) among patients prior to and during A₁-PI augmentation therapy according to baseline FEV₁ strata in the Wencker 2001 study [49].

Initial FEV ₁ % Predicted ^a	n	Mean Slope FEV ₁ (mL/year) Prior to A ₁ -PI Augmentation	Mean Slope FEV ₁ (mL/year) During A ₁ -PI Augmentation	Difference Between Augmented and Non-Augmented Periods (SD)	P Value ^c
<30 %	25	-15.3	-19.0	-3.7 (48.6)	NS
30-65 %	60	-49.3	-37.8	11.6 (48.8)	0.066
Overall^b	96^b	-49.2^b	-34.3^b	14.9^b (61.4)	0.019^b

^a Initial FEV₁% Predicted refers to the FEV₁% predicted at the start of the period of observation without A₁-PI augmentation therapy. Results for the baseline FEV₁ stratum FEV₁% predicted >65 % are not shown because they are subject to regression to the mean due to the additional requirement for this stratum only that patients be fast decliners (rate of FEV₁ decline of >120 mL/yr).

^b FDA considers the overall results unreliable due to inclusion of data for 11 subjects with baseline FEV₁% predicted who were also required to be fast decliners (rate of FEV₁ decline of >120 mL/yr) and whose results were therefore subject to regression to the mean.

^c Reported p values are not adjusted for multiple comparisons, so they may be considered to lack validity.

Wencker et al. observed a mean slowing of 14.9 (SD 61.4) mL/yr in FEV₁ rate of change after initiation of augmentation therapy in the overall cohort ($p = 0.019$, unadjusted for multiplicity, Table 5). For subjects ($n = 60$) with an initial FEV₁ 30–65 % of predicted at the start of unaugmented observation, the declining slope of FEV₁ was reduced (less negative) by a mean of 11.6 (SD 48.8) mL/yr ($p = 0.066$, unadjusted) during A₁-PI administration [49].

The above results use the initial FEV₁ to stratify subjects; different results were obtained when FEV₁ at the time of initiation of augmentation therapy was used for stratification [49]. Because of limitations in study design, inconsistent results within FEV₁ strata, depending on whether the initial or final FEV₁ during the unaugmented observation period was used for defining subgroups, and potential regression to the mean to have biased results, this study should be interpreted with great caution. Nevertheless, this study is included in a published meta-analysis of the effects of augmentation therapy in AATD-related COPD [41].

4.4. Tonelli 2009 study

Tonelli et al. published an observational study involving 164 Pi*ZZ AATD patients enrolled in the Alpha-1 Foundation DNA and Tissue Bank study [40]. Change in FEV₁ (Δ FEV₁) was defined as the initial FEV₁ minus the FEV₁ obtained by random effects model. The change in FEV₁ was compared between patients who were receiving A₁-PI augmentation therapy at the time of enrollment ($n = 124$) and those who were not ($n = 40$) overall and for subgroups based on initial FEV₁ % of predicted. An analysis of mortality at 5 years used logistic regression with age, gender, baseline FEV₁, presence of COPD and smoking status as covariates.

The mean follow-up time was 41.7 (SEM \pm 2.6; range: 6 to 268) months. The augmented patients were older on average, more commonly had COPD with lower baseline FEV₁ (43 % (SEM \pm 2 %) versus 77 % (SEM \pm 5 %)), and had a higher number of individuals who required oxygen therapy. The study was underpowered to detect any difference in survival between the groups [40].

As shown in Table 6, when adjusted by age at baseline, sex, smoking status and baseline FEV₁ % of predicted, the difference in mean overall change in FEV₁ between these groups was 47.6 mL/year ($p = 0.05$) favoring augmentation therapy. Similar results were obtained when 26 patients in whom the smoking history was initially missing but later located were added to the analysis [40].

Patients with FEV₁ >65 % had higher rates of FEV₁ decline if they received augmentation therapy (Δ FEV₁ augmented: -108.7 ± 17.3 mL/year versus Δ FEV₁ non-augmented: -29.2 ± 15.3 mL/year; for an overall difference of -79.5 mL/year, $P < 0.001$). A trend towards a beneficial effect of A₁-PI was observed in the group with an initial FEV₁ 30 %–65 % (difference in Δ FEV₁: 54 mL/year, $P = 0.07$) [40].

4.5. Additional studies of FEV₁ rate of change among patients receiving vs. not receiving A₁-PI

The Shouten 2021 retrospective study [49], the National German Registry study [50], and the Spanish National Database study [51] involved 128/246, 85/15, and 77/45 Pi*ZZ patients who received/did not receive A₁-PI, respectively. All three studies failed to demonstrate a significant difference in rate of change of FEV₁ between patients who received and did not receive A₁-PI. Further details of these studies are provided in online Supplement C.

4.6. Sandhaus 2020 study

A retrospective analysis of two prospectively followed Pi*ZZ (“or worse” genotype) AATD patient cohorts with emphysema in two countries differing in access to A₁-PI augmentation therapy was undertaken [52]. The U.S. cohort comprised patients receiving A₁-PI and followed in AlphaNet’s Disease Management and Prevention program with monthly telephone interviews. Mortality data was derived from the US National Death Index and direct notification by family. The U.K. control group comprised A₁-PI naïve patients from the AATD U.K. Registry who were prospectively followed with annual medical review and physiological measurements. Mortality data for this cohort were derived from a central NHS database, general practitioners or by direct notification by family. Quality of Life was assessed in both groups annually with the St. George’s Respiratory Questionnaire (SGRQ). Subjects ($n = 655$) were matched based on age, sex, baseline year, and smoking history. Overall 10-year survival rate was longer in the A₁-PI (U.S.) group compared to matched controls from the U.K., with rates of 80 % (95 % CI 76.7 to 83.4) and 68.5 % (95 % CI 64.2 to 73.1), respectively ($p < 0.001$). Lung transplantation within 5 years of baseline assessment occurred in 13.3 % (95 % CI 7.54 to 18.7) of the U.S. treatment group and 58.5 % (95 % CI 40.3 to 71.2) of the U.K. control group. Annual worsening of SGRQ total was on average 1.3 points worse/year in control group patients compared to those receiving A₁-PI (95 % CI 0.41 to 2.19, $p = 0.004$). Features of the AlphaNet Program other than A₁-PI augmentation therapy and other differences in care between the U.S. and U.K. may have affected the reported outcomes of this

Table 6

Δ FEV₁ (mL/year) among patients receiving and not receiving A₁-PI augmentation therapy according to baseline FEV₁ strata in the Tonelli study [40].

Initial FEV ₁ % Predicted	Augmented		Non-Augmented		p Value ^a
	n	Δ FEV ₁ (mL/year)	n	Δ FEV ₁ (mL/year)	
<30 %	30	0.9	3	20.1	0.59
30–65 %	79	2.08	10	−51.9	0.07
>65 %	15	−108.7	27	−29.2	0.0006
Overall	124	10.6	40	−37.0	0.05

^a Reported nominal p values are unadjusted for multiple comparisons, so they may be considered to lack validity.

retrospective cross-country comparison [52].

4.7. Chapman meta-analysis

Chapman et al. published a meta-analysis of four observational studies and one randomized, placebo-controlled trial of A₁-PI augmentation therapy [41]. This meta-analysis of 1509 patients should be interpreted with caution due to the predominant inclusion of observational studies subject to bias from unaccounted covariates. This meta-analysis includes the Wencker 2001 retrospective within-patient comparison [49] and the Seersholm study [42] which compared FEV₁ rates of change between patients in Germany receiving augmentation and patients in Denmark not receiving augmentation therapy. Both studies were subject to regression to the mean among German patients with baseline FEV₁ > 65 % at baseline, who were required to be fast decliners. As shown in Table 7 below, among all patients included in the meta-analysis, the mean decline in FEV₁ was slower by 13.4 mL/year (95 % CI 1.5–25.3 mL/year) among patients receiving A₁-PI compared to those who were not. This overall effect was said to reflect predominantly the results of patients with baseline FEV₁ of 35–65 % of predicted, whose absolute difference in average rates of FEV₁ decline was 17.9 mL/yr (95 % CI 9.6–26.1 mL/yr). Differences with and without augmentation therapy among patients whose baseline FEV₁ were <30 % or >65 % were not significant. The statistical significance in the overall pooled cohort is lost after removing either the Wencker or Seersholm observational studies, but the statistical significance is retained for the baseline FEV₁ 30–65 % of predicted subgroup after removal of the Wencker study. The point estimate for the single randomized, placebo-controlled trial included in the meta-analysis shows a small, non-significant trend in rate of change of FEV₁ favoring the placebo group [32].

4.8. Barros-Tizon 2012 exacerbation study

Barros-Tizon et al. conducted a retrospective, multicenter, observational study of AATD patients in Spain whose records were available for 18 months prior to and 18 months while on A₁-PI augmentation therapy with either Trypsone (Grifols) or Prolastin (Talecris) [53]. Of 127 patients included in the study, only 6 % received A₁-PI weekly, with 17 % receiving A₁-PI biweekly and 76 % receiving A₁-PI every 3 weeks. (Every- 2- and 3- week regimens are associated with higher peak serum levels of A₁-PI). The study's primary objective was to examine whether A₁-PI augmentation therapy was associated with a reduction of both the number of exacerbations and the percentage of patients experiencing exacerbations, defined using GOLD criteria [54] and requiring a change in medical regimen.

Overall, mean exacerbations per patient were reduced from 1.2 ± 1.6 prior to augmentation therapy to 1.0 ± 2.2 ($p < 0.01$) during the 18 months following initiation of A₁-PI treatment. Among patients who had at least one exacerbation during the 18 months prior to initiation of augmentation therapy, mean exacerbations per patient were reduced from 2.0 ± 1.6 to 1.4 ± 2.7 ($p < 0.01$) during augmentation therapy [53].

Augmentation therapy was associated with a reduction in the percentage of patients who experienced one or more exacerbations from 59.1 % to 44.1 % in the overall population ($p < 0.005$). The difference between the two periods in the percentage of patients who experienced one or more severe exacerbations was not statistically significant (46.7 % vs. 55.4 %, respectively). Multivariate analysis with imputation of missing values for lung function and laboratory variables suggested that unaugmented patients would be expected to be at 1.4- to 4.2-fold more risk of exacerbation than patients receiving A₁-PI [53]. Courses of systemic antibiotics were significantly reduced during the period of augmentation therapy compared with the prior period (means: 13 vs. 28, respectively, $p < 0.05$) [53].

Limitations of the study include its retrospective nature, lack of randomization and lack of concurrent control, heterogeneity of A₁-PI dosing regimen and agent, and lack of masking which may have contributed to an expectation of benefit and influenced subjects' reporting of exacerbations.

5. Randomized controlled trials of augmentation therapy

The results of three published double-blind, placebo-controlled RCTs of augmentation therapy, each testing a different A₁-PI product, are shown in Table 8.

Table 7

FEV₁ slope (mL/year) among patients receiving and not receiving A₁-PI augmentation therapy according to baseline FEV₁ strata (Chapman meta-analysis) [41].

Initial FEV ₁ % Predicted	Augmented		Non-Augmented		Slope Difference* (95 % CI)
	n	Mean Slope FEV ₁ (mL/year)	n	Mean Slope FEV ₁ (mL/year)	
<30 %	454	−30.6	180	−30.9	1.8 (−7, 10.5)
30–65 %	398	−50.8	263	−67.9	17.9 (9.6, 26.1)
>65 %	43	−92.1	175	−97.2	3.5 (−49, 55.9)
Overall	924	−48.0	681	−59.4	13.4 (1.5, 25.3)

*Slope differences are as reported in the publication and are noted to not exactly correspond to the differences in mean slopes between augmented and non-augmented subjects as reported in this table and the publication.

Table 8
Randomized, double-blind, placebo-controlled trials of A₁-PI in severe AATD-Associated emphysema.^a

Reference	A ₁ PI Product Manufacturer	Number of Subjects Randomized	Adjusted HRCT Lung Density Outcome	FEV ₁ Slope Outcome	Exacer-bations Outcome
20 Dirksen A et al., 1999	Laboratoire Biologique du Fractionnement (LBF)	56	PD15 ^b p = 0.07	p = 0.20 ^b favored placebo	NM ^d
21 Dirksen A et al., 2009 [EXACTLE]	Grifols	77	PD15 ^b p = 0.068	NS ^c favored placebo	NS ^c favored placebo
22,23,24 Chapman KR et al., 2015 [RAPID]	CSL Behring	180	PD15 ^b p = 0.06	NS ^c favored placebo	NS ^c favored placebo

^a Lung-Volume-Adjusted High Resolution Computerized Tomographic (HRCT) Lung Density, FEV₁ Slope, and total Exacerbation Incidence Outcomes are given as unadjusted 2-sided p values for the difference between A₁-PI and placebo arms.

^b Primary Efficacy Endpoint, 2-sided p values.

^c Secondary Efficacy Endpoint, 2-sided p value.

^d NM = Not Measured.

5.1. Dirksen 1999 RCT

This RCT randomized 56 Danish and Dutch ex-smokers with PI*ZZ AATD and FEV₁ 30%–80% of predicted to LBF (French) A₁-PI 250 mg/kg or albumin 625 mg/kg IV every four weeks for at least 3 years [32]. The primary endpoint, FEV₁ rate of decline, was non-significantly slower in the placebo group compared to the A₁-PI group (p = 0.20). The rate of decline in lung density by lung-volume-adjusted CT (assessed by the PD15, the 15th percentile of lung density voxels), a secondary endpoint, was numerically slower in the A₁-PI group (mean 1.5 ± 0.41 g/L/year) compared with the placebo group (mean 2.6 ± 0.41 g/L/year), with a difference of 1.07 g/L/year between treatment groups (p = 0.07, unadjusted). Exacerbation rates and mortality were not reported [32].

5.2. Dirksen 2009 [EXACTLE] RCT

This exploratory RCT, sponsored by Prolastin's manufacturer, randomized 77 PI*ZZ AATD Danish, Swedish, and U.K. subjects with emphysema 1:1 to Prolastin brand A₁-PI or placebo at the weekly IV dose of 60 mg/kg for a treatment period of 2–2.5 years [33]. The primary analysis of the primary endpoint, lung-volume-adjusted rate of change in lung density by CT, was numerically slower in the A₁-PI group (mean −1.384 g/L/year) compared with the placebo group (mean −2.241 g/L/year), with an estimated treatment difference of 0.857 g/L/year (p = 0.068). As in the 1999 Dirksen study, the mean rate of decline in FEV₁ was numerically faster in the A₁-PI group than in the placebo group, but this difference was not statistically significant. The annual incidence of exacerbations trended higher in the A₁-PI group (2.55 ± 2.14) than in the placebo group (2.19 ± 1.33), (difference NS), but, in a *post-hoc* analysis the percentage of severe exacerbations among those with known severity was lower in the A₁-PI group (13/194, 6.7%) compared to that in the placebo group (21/155, 13.5%, p value for difference 0.013). A statistically significant weak correlation was observed between changes in lung-volume-adjusted lung density and FEV₁ (Correlation coefficient 0.216, p = 0.007) [33].

5.3. Chapman 2015 (RAPID) RCT

This RCT, sponsored by Zemaira's manufacturer, randomized 180 U.S. and EU non-smoker and ex-smoker subjects with AATD and FEV₁ values 35–70% of predicted 1:1 to Zemaira brand A₁-PI or placebo at 60 mg/kg IV weekly for 2 years [34]. The RAPID trial had a substantial imbalance in baseline lung density that favored the A₁-PI treatment arm. Baseline lung density was 3.2 g/L greater in the placebo group than in the A₁-PI group. Data from the placebo subjects in the trial (and to a lesser extent in the A₁-PI arm subjects) have shown a faster rate of decline of adjusted lung density among subjects with higher baseline lung density values [55]. Eighteen subjects in the placebo group and 9 in the A₁-PI group discontinued the study prematurely; discontinuations due to an adverse event or death numbered 7 versus 2, respectively, in placebo and A₁-PI arms). The primary analysis of the primary endpoint, lung-volume-adjusted rate of change in lung density by CT using scan data obtained both at the two inspiratory states, total lung capacity (TLC) and functional residual capacity (FRC), was numerically slower in the A₁-PI group (mean −1.45 g/L/yr) compared with the placebo group (mean −2.19 g/L/yr), with a treatment difference of 0.62 g/L/yr (2-sided p = 0.06) [34]. Statistically combining CT scan data from scans taken at both TLC and FRC had been proposed by the manufacturer and accepted by FDA. Only one of multiple prespecified supplementary analyses of adjusted lung density rate of change (using only CT data obtained at TLC) achieved statistical significance (without adjustment for multiple analyses). Some have suggested that this particular supplementary analysis should be relied upon instead of the prespecified primary analysis of the primary efficacy endpoint for concluding the product was effective, because the primary analysis of the primary endpoint combined CT data taken at both FRC and TLC and because *unadjusted* lung density measurements made at FRC are more variable than those made at TLC. Such a *post-hoc* approach for drawing inferences regarding efficacy is not statistically justified. Adjusting lung density by lung volume substantially reduces variability [56–59], and the combined FRC/TLC adjusted lung density measurements for the prespecified primary endpoint were slightly *less variable* than those for the supplementary analysis using only adjusted lung density measurements taken at TLC. While changes in air trapping can theoretically

affect CT lung density measurements taken at FRC more than TLC, correction of lung density values for lung volume accounts for changes in lung volume due to air trapping. The RAPID trial's pre-specified primary endpoint combining adjusted TLC and FRC lung density data was an appropriate and acceptable primary analysis approach (which, nevertheless, did not achieve statistical significance).

Subgroup analyses of the between-treatment-group mean annual rates of change in adjusted lung density at TLC from the intent-to-treat population of the RAPID trial indicate a several-fold difference between men and women [60]. While the overall difference between Zemaira and placebo arms in rates of decline in adjusted lung density was 0.74 g/L per year, the difference among females ($n = 79$) was 1.45 g/L per year. Higher between-treatment-group differences were also seen for subjects with higher baseline A₁-PI serum levels and subjects with BMI >30 kg/m² ($n = 21$, difference in P15 rate of change = 2.21 g/L per year). While these subgroup analyses are exploratory, they merit further study. The observation of a greater between-treatment group difference among obese subjects may relate to expected higher blood A₁-PI levels in subjects dosed on a mg/kg basis who have a lower percent lean body mass.

5.4. RAPID trial extension

In the RAPID trial extension phase non-U.S. subjects were offered open-label Zemaira brand A₁-PI for two years [35]. One hundred forty subjects participated: seventy-six who had been randomized to Zemaira (early-start group) and 64 who had been randomized to placebo (late-start group) during the prior two-year RAPID trial. One hundred thirty-one subjects completed the RAPID trial extension with 121 having complete lung density by HRCT data. In a *post-hoc* subgroup analysis of late-start group subjects who had HRCT lung density data, these subjects experienced a slowing in their mean rate of change of lung volume-adjusted lung density by HRCT at TLC from -2.26 g/L per year during their two years of placebo treatment in the RAPID trial to -1.26 g/L per year during their two years of open-label Zemaira treatment in the RAPID extension [35].

Using a single mixed-effects regression model with combined data from the RAPID trial and its open-label extension, an inflection point in the mean rates of lung density decline was seen when the late-start group began treatment with open-label Zemaira, with a reduction in the mean rate of lung density by HRCT loss during the trial extension of 0.52 g/L per year (95 % CI 0.22–0.83 g/L per year). A slight acceleration in the mean rate of decline in adjusted lung density by HRCT at TLC in early start subjects after entering the RAPID trial extension was not significant (change from -1.51 g/L per year during the RAPID trial to -1.63 g/L per year during the RAPID trial extension) [35]. The less negative slope in the mean rate of lung density decline in the late-start group subjects after starting open-label Zemaira is suggestive of a treatment effect of Zemaira; however, the results are tempered by the *post-hoc* nature of the analysis, the lack of a concurrent control group continuing placebo treatment during the RAPID trial extension, lack of blinding during the extension, and the modest absolute difference in the delayed start subgroup in the mean rates of lung density decline between the RAPID trial and its extension, which is of uncertain clinical significance.

5.5. Non-HRCT outcomes in the rapid trial

As in the two Dirksen studies, the mean rate of decline in FEV₁ in the RAPID trial was numerically slightly faster in the A₁-PI group compared to the placebo group, but this difference was not statistically significant (NS) [34]. The annual incidence of serious exacerbations was also numerically greater in the A₁-PI group than in the placebo group in RAPID, (differences NS). Non-statistically significant trends in changes in exercise capacity by shuttle walk test, incidence of subjects with dyspnea reported as an adverse event, hospitalizations for exacerbations, duration of hospitalizations, and change in diffusion capacity of carbon monoxide (DL_{CO}) all favored the placebo group [34]. Combining data from the double-blind RAPID trial and its open label extension, a statistically significant weak correlation was observed between rate of change in adjusted lung density by CT and change in FEV₁ ($r = 0.286$) and change in FEV₁ % predicted ($r = 0.0338$) [35]. This finding is difficult to interpret, given that in the double-blind RAPID trial the non-statistically significant trend in change in FEV₁ numerically favored the placebo group [34].

6. Exacerbation outcomes in randomized, placebo-controlled trials of A₁-PI augmentation therapy

A meta-analysis of two completed randomized, double-blind, placebo-controlled trials [33,34] of IV A₁-PI products found a statistically significant increase in exacerbation frequency in the A₁-PI arms compared to placebo (difference: 0.29 exacerbations per subject-year, 95 % CI 0.04–0.54, $P = 0.02$) [61]. As noted, in a *post-hoc* analysis performed in one of these two trials reporting exacerbation severity data, the frequency of severe exacerbations in the subset of A₁-PI arm exacerbation episodes for which severity data were available was lower compared to the corresponding subgroup of the placebo arm [33]. In the other trial in which serious exacerbation data were analyzed (RAPID), a non-statistically significant numerical increase in exposure-adjusted serious exacerbation rate was observed in the A₁-PI arm (0.146 serious exacerbations per subject-yr) compared to that in the placebo arm (0.115 serious exacerbations per subject year, ratio A₁-PI: placebo 1.256, 95 % CI 0.46–3.45) [17]. In another *post-hoc* analysis of a subset of subjects completing a 2-year open-label extension phase of the RAPID trial, subjects in both randomization arms experienced further increases (60 % and ~100 % in late and early start groups, respectively) in exposure-adjusted serious exacerbation rates after switching from placebo or double-blind A₁-PI to open-label A₁-PI [17]. Only the increase in serious exacerbation rate in the early start group originally randomized to double-blind A₁-PI and then switched to open-label A₁-PI during the extension phase achieved nominal statistical significance (95 % CI for the ratio of rates during the two periods: 1.21–3.67) [17]. The 95 % CI for the ratio of serious exacerbation rates observed in the placebo arm of the RAPID trial after vs. before crossover to open-label Zemaira treatment during the extension study was 0.80–3.03) [17].

7. Meta-analyses of randomized, placebo-controlled trials of A₁-PI augmentation therapy

Gøtzsche and Johansen of the Cochrane Group published a meta-analysis of the three randomized, double-blind, placebo-controlled trials of A₁-PI augmentation therapy described above [62]. The planned primary efficacy endpoint of the meta-analysis, mortality, wasn't possible, as mortality was only reported for the RAPID trial. Combining data from all three trials, the secondary outcome variable in the meta-analysis, lung density by HRCT, deteriorated significantly less in the augmentation therapy group than in the placebo group, (mean difference 0.86 g/L per year, 95 % CI 0.31 to 1.42; $p = 0.002$). Similar results were obtained by Edgar et al. in a separate meta-analysis of the same three trials (mean difference 0.79 g/L per year, 95 % CI 0.29–1.29; $p = 0.002$) [61]. Across the three trials, FEV₁ deteriorated more in the treatment group than in the placebo group, but there was no significant between-group difference (SMD -0.19, 95 % CI -0.42 to 0.05; $p = 0.12$ in Gøtzsche and Johansen [62] difference in mean FEV₁ % predicted -0.56, 95 % CI -1.41 – 0.09, $p = 0.20$ in Edgar et al. [61]). On average, carbon monoxide diffusion capacity deteriorated more in the treatment group than in the placebo group, but there was no significant between-group difference (SMD -0.11, 95 % CI -0.35 to 0.12; $p = 0.34$ in Gøtzsche and Johansen [62]; SMD -0.11, 95 % CI -0.33 to 0.11; $p = 0.34$ in Edgar et al. [61]).

8. Randomized dose comparison studies of A₁-PI augmentation therapy

8.1. Campos 2013 [SPARK] study

The 2013 SPARK study conducted by Campos et al. [63], compared the safety and pharmacokinetics of Prolastin-C brand A₁-PI at weekly 120 mg/kg and 60 mg/kg intravenous doses in 30 subjects with severe AATD who were randomized to receive one of the two weekly doses for 8 weeks before being crossed over to the alternate dose for an additional 8 weeks. Mean trough antigenic A₁-PI levels were 27.7 μM following 120 mg/kg weekly Prolastin-C and 17.3 μM following the 60 mg/kg weekly dose. Twenty-three subjects on the 60 mg/kg dose reported 69 treatment-emergent adverse events (TEAEs) and 18 subjects on the 120 mg/kg dose reported 43 TEAEs. No subjects reported serious AEs and no subject withdrew due to an AE. Seven subjects (23 %) experienced a total of 9 exacerbations while on the 60 mg/kg dose and 5 subjects reported a total of 6 exacerbations while on the 120 mg dose [63].

8.2. Campos 2019 study

The study published in 2019 by Campos et al. [31] was an open-label, 3-period crossover study of ten subjects (nine ZZ and one SZ genotype) with severe AATD-related emphysema who received four weeks of standard 60 mg/kg weekly A₁-PI (Zemaira) therapy, followed by 4 weeks of 120 mg/kg weekly A₁-PI therapy, followed by a return to 4 weeks of 60 mg/kg weekly A₁-PI administration. Bronchoalveolar lavage (BAL) and plasma samples were obtained for A₁-PI levels and other biomarkers at baseline and at the end of each 4-week study period while clinically stable. The study enrolled subjects who had been receiving A₁-PI at the standard weekly intravenous dose of 60 mg/kg, as well as at least one long-acting bronchodilator and an inhaled corticosteroid [31].

The primary outcome was change in pulmonary inflammatory markers measured in BAL fluid. Secondary endpoints were improvement in serum/plasma inflammatory markers and elastin degradation markers in serum and BAL fluid [31].

Trough antigenic AAT levels increased from $16.7 \pm 2.3 \mu\text{M}$ following standard A₁-PI dosing at the end of the first period to $27.2 \pm 5.0 \mu\text{M}$ by the end of the 2nd (120 mg/kg weekly “double dose”) period, and returned to $16.0 \pm 2.6 \mu\text{M}$ at the end of the 3rd period after resumption of the standard dose. Trough AAT levels were within the normal range on the 120 mg/kg weekly regimen. BAL fluid neutrophil elastase decreased from $3.32 \pm 0.86 \text{ nM}$ on standard dose A₁-PI to $1.61 \pm 0.29 \text{ nM}$ ($p = 0.008$) on the double dose regimen, while BAL fluid cathepsin G, another lung protease, fell from $247.0 \pm 60.4 \text{ nM}$ to $56 \pm 15.3 \text{ nM}$ ($p = 0.005$). Mean trough neutrophil elastase and cathepsin G levels were slightly but not statistically significantly higher at the end of period 3 resumption of the standard dose A₁-PI regimen, suggesting a carry-over effect from the prior period of double dose administration. Plasma levels of Aa-Val³⁶⁰, a biomarker of neutrophil elastase activity, fell significantly from $10.15 \pm 1.19 \text{ nM}$ after SD therapy to $7.89 \pm 0.57 \text{ nM}$ on double dose therapy ($p = 0.016$), with a subsequent rise a month after single dose therapy was resumed. The markers of elastin degradation, trough BAL fluid desmosine and isodesmosine, fell from $3.74 \pm 3.7 \text{ pg/mg protein}$ on SD therapy to $0.65 \pm 0.34 \text{ pg/mg protein}$ after double dose A₁-PI administration ($p = 0.050$). The lower BAL fluid levels were maintained after resumption of the standard dose, again suggesting a carry-over effect. Plasma desmosine/isodesmosine levels were unchanged on double dose A₁-PI therapy [31].

Levels of several, but not all markers of airway inflammation (cytokines, chemokines, and growth factors) also fell by the end of period 2 double dosing, including tumor necrosis factor α (TNFα), interleukin-17 (IL-17), granulocyte-macrophage colony-stimulating factor (GM-CSF), macrophage migration inhibitor factor (MIF), macrophage inflammatory protein 1a chemokine (C-C motif) ligand 3 (MIP1aCCL3), macrophage colony-stimulating factor (M-CSF), interferon γ (IFNγ), IL-2, IL-12p40, IL-10, IL-3, Eotaxin/CCL11, IL-9, IL-4, and basic fibroblast growth factor (bFGF). At the end of period 3, some of these markers returned to levels observed at the end of the initial standard dose period 1, but others remained depressed, suggesting a possible carry-over effect. One month of double-dose therapy appeared to downregulate cytokines that affect the Janus kinase signal transducer and activator of transcription proteins (JAK-STAT), T cell receptor signaling pathways, and those that affect cytokines in macrophage migration, eosinophil recruitment, humoral and adaptive immunity, neutrophil activation, and cachexia [31].

These results suggest that patients receiving standard dose A₁-PI augmentation therapy can have ongoing protease activity, elastin degradation, and lung inflammation that is ameliorated after one month of “double dosing” with 120 mg/kg weekly intravenous dosing. The observation of a possible carry-over effect for several inflammatory biomarkers and protease activities suggests there could be a role for studying alternating periods of higher and standard doses in patients with severe AATD – related lung disease [31].

The authors concluded that “increasing AAT levels into the normal range may provide additional clinical benefits and have a more robust impact on clinical outcomes in subjects with AAT deficiency requiring augmentation therapy.” [31].

8.3. SPARTA trial

The Study of ProlAstin-C Randomized Therapy with Alpha-1 augmentation (SPARTA, NCT01983241) [25], is a randomized, double-masked, parallel, placebo-controlled, dose comparison phase 3 trial in which 339 subjects with severe AATD-RE are being randomized 1:1:1 to receive placebo, 60 mg/kg weekly Prolastin-C brand A₁-PI, or 120 mg/kg weekly Prolastin C for three years. The primary endpoint is the rate of change of lung volume-adjusted whole lung HRCT lung density (PD15) at TLC. Secondary endpoints include severe exacerbations of COPD, rate of change of PD15 of the basal lung region, changes from baseline in FEV₁, change from baseline in Saint George’s Respiratory Questionnaire, and change from baseline in EuroQol (Quality of Life)-5 Demension-5 Level. The study began November 2013 and was projected to be completed June 2025 (total duration 12 years). The trial has enrolled subjects with severe AATD having allelic combinations of ZZ, SZ, Z(null), (null)(null), S(null), or “at-risk” alleles with screening visit FEV₁ > 30 % and <80 % of predicted and FEV₁/FVC, <70 %, DL_{CO} < 60 % predicted OR evidence of pulmonary emphysema on CT scan. Exclusion criteria include receipt of more than one month of A₁-PI augmentation therapy within the past six months, history of an exacerbation within five weeks prior to screening, and history of lung or liver transplant [25].

9. Discussion

Findings from epidemiologic studies in AATD-RE suggest that the rate of decline in FEV₁ may be greatest among unaugmented patients with intermediate degrees of airflow limitation and that the possible benefit of A₁-PI augmentation therapy, if any, might be concentrated in this subgroup [39–43,64]. A post-hoc analysis of the largest prospective observational study of A₁-PI use in AATD [39] provides weak evidence that it may slow the rate of decline in FEV₁ in the subgroup of subjects with FEV₁ 30–65 % of predicted. While the study authors reported statistical significance for FEV₁ change in favor of augmentation therapy among subjects with FEV₁ 35–49 % and 30–65 %, these results lack statistical significance after adjustment for multiple comparisons and should be considered for hypothesis generation only. The slowing in mean decline in FEV₁ in this subgroup receiving A₁-PI was less than half that required to slow the decline to that attributable only to aging in AAT-replete individuals. There is no evidence that the study met its prespecified primary endpoints for either FEV₁ change or mortality, although an updated *post-hoc* analysis by different authors [48] suggested that mortality might be improved among patients with baseline FEV₁% predicted between 10 % and 60 %, with the largest mortality reduction in patients whose FEV₁ was 35–49 %. The authors cautioned that this mortality result may have been confounded by differences in important baseline variables that were not captured/accounted for in the analyses.

The data from several of the smaller observational studies included in this review were taken from national registries in AATD whose main objectives were not the evaluation of the efficacy of A₁-PI augmentation therapy per se. For this reason, in addition to bias resulting from physician and patient decisions regarding whether to undergo augmentation therapy and possible imbalances in measured and unmeasured covariates, the results of these observational studies results must be interpreted with caution.

Notwithstanding their limitations and variability of results and study power, taken as a whole, epidemiology studies of A₁-PI at the weekly IV dose of 60 mg/kg vs. untreated patients provide weak evidence suggesting potential benefit in slowing FEV₁ rate of decline in subjects with FEV₁ baseline values from ~30 % to ~65 %, with little apparent benefit in subjects with milder or more severe airflow obstruction at baseline [39–43,64]. However, this finding has not been confirmed by three underpowered, randomized, placebo-controlled trials [32–34].

An 11.5 % absolute reduction in mortality with A₁-PI was observed in a *retrospective* cross-country analysis by Sandhaus et al. [52], as well as a substantial reduction in lung transplantation and a slight reduction in the rate of deterioration in SGRQ. However, differences in care between the U.S. and U.K. may have affected the reported outcomes of this retrospective cross-country comparison.

The retrospective crossover study by Barros-Tizon et al. [53] found a 15 % absolute reduction during augmentation therapy in the proportion of patients who had exacerbations, but the A₁-PI regimen consisted of weekly dosing in only 6 % of patients and there was no concurrent control group or masking. The higher peak serum A₁-PI level associated with the higher doses given at less frequent than weekly intervals in that study might have yielded different results from what may occur with standard weekly dosing at 60 mg/kg.

While none of the three randomized, placebo-controlled trials of three different A₁-PI products achieved statistical significance for their prespecified primary efficacy endpoints (FEV₁ in the first trial and PD15 lung-volume-adjusted lung density by HRCT in the second and third trials), two published meta-analyses of the PD15 results of these studies achieved statistical significance favoring the A₁-PI groups [61,62]. Because the first RCT [32] used a markedly different dosage schedule from the other trials (monthly vs. weekly, resulting in a higher peak and lower trough serum A₁-PI concentration, I question the appropriateness of including the first study in these meta-analyses. The RAPID trial [34], had a substantial imbalance in baseline lung density that appears to have favored the A₁-PI treatment arm. Baseline lung density was not included in the model used for analysis of the primary endpoint in the RAPID trial [34], so the degree to which the reported p values would have changed by its inclusion in the model remains unknown. I encourage the manufacturer to make public the raw data from the trial, so that the effect of this baseline imbalance could be independently assessed.

Other potential confounders in the RCTs of A₁-PI products include differential dropout rates, potential differences over the course of the trials between randomization groups in subjects experiencing fluid shifts into and out of the lungs and/or changes in inflammatory cell burden as from exacerbations which would affect measured lung density independent of any changes in lung parenchymal mass, and differences in the proportion of men/women between treatment arms, if the underlying rate of lung density loss differs between sexes as is well-established in the case of FEV₁. An analysis of the effect of baseline demographics on HRCT outcomes in the

RAPID trial suggests a large sex difference in the between-treatment-group differences in rate of decline of lung volume-adjusted lung density (PD15) [60].

The between-group differences in mean rates of change in PD15 lung-volume adjusted lung density over the course of the RCTs are of uncertain clinical significance. The heterogeneity of published studies and limited longitudinal data have made it difficult to determine the magnitude of change in adjusted lung density by HRCT rate of change that would constitute the minimum clinically important difference (MCID) [65]. Nevertheless, a MCID for HRCT lung density of -2.89 g/L has been proposed by one group [66,67]. The mean differences observed between A₁-PI and placebo groups in adjusted lung density in RAPID were less than half of this estimated MCID as proposed by Crossley et al., and less than one-fifth of the within-subject standard-deviation of 3.42 g/L per yr. for repeated lung density measurements reported in the ECLIPSE study [45].

There are no data from *interventional* clinical trials to indicate that a difference in HRCT lung density is predictive of any clinical benefit to patients. To date, no RCT has shown both a statistically significant result for HRCT lung density changes and a statistically significant positive result for any other non-HRCT-related clinical endpoint [32–34]. In the largest RCT comparing A₁-PI with placebo (RAPID) [34], trends in non-HRCT lung density secondary clinical endpoints consistently favored placebo, whereas trends in HRCT lung density favored the augmentation therapy group. Although one randomized controlled trial [33] and one subgroup analysis of another randomized controlled trial [34] combined with its open-label extension study [35] have shown weak correlations between change in HRCT lung density and change in FEV₁, in both trials point estimates for change in FEV₁ numerically favored the placebo groups rather than the A₁-PI therapy groups.

For a biomarker such as HRCT lung density to be considered a validated surrogate endpoint suitable for use as a single primary endpoint in Phase 3 clinical trials intended to support marketing, the endpoint should be “supported by a clear mechanistic rationale and clinical data providing strong evidence that an effect on the surrogate endpoint predicts a specific clinical benefit [as measured by how patients feel, function or survive] ... Generally, required evidence includes a combination of a clear mechanistic rationale and in most cases, data from multiple randomized clinical trials showing that the effect on the surrogate endpoint predicts the effect on the clinical outcome of primary interest.” [68].

As for non-lung-density efficacy endpoints, none of the RCTs provided evidence that the A₁-PI augmentation therapy products at the doses administered produced significant favorable effects on mortality, exacerbation rate, FEV₁ rate of change, DL_{CO} rate of change, exercise capacity, or symptoms such as dyspnea. A meta-analysis of RCTs two and three [33,34] which measured exacerbation frequency found a significant excess of exacerbations in the A₁-PI arms compared to placebo [61]. A post-hoc analysis of the exposure-adjusted serious exacerbation rate found that the rate doubled during open-label A₁-PI administration years 3 and 4 of the RAPID trial extension compared to the rate in the active A₁-PI arm during double-blind years 1 and 2 of the RAPID trial (95 % CI for the ratio rates during the two periods: 1.21–3.67) [17]. The serious exacerbation rate observed in the placebo arm of the RAPID trial increased by 60 % after crossover to open-label Zemaira treatment during the extension study (95 % CI for the ratio of the rates: 0.80–3.03) [17].

After the licensure of the first A₁-PI product, FDA determined that there is no adequate scientific basis to conclude that the historic 11 μM target for serum A₁-PI is an appropriate therapeutic target for augmentation therapy [24]. For this reason, following a unanimous recommendation of the FDA Blood Products Advisor Committee (BPAC) in 2009 [24], FDA has requested manufacturers of new plasma-derived A₁-PI products to conduct pre-licensure a randomized, concurrently controlled, double-blind clinical trial using clinically-meaningful endpoint(s) that reflect how patients feel, function, or survive, or the forced expiratory volume in 1 s (FEV₁) rate of change as the primary efficacy endpoint to provide substantial evidence of effectiveness [28].

10. Conclusion/recommendations

Studies evaluating the clinical efficacy of A₁-PI augmentation therapy in emphysema associated with AATD, taken as a whole, are inconclusive. Further validation of lung volume-adjusted lung density by HRCT in interventional trials should be undertaken before this marker is used as a sole primary endpoint of future clinical trials. This would include demonstrating, in the context of one or more adequate and well-controlled randomized interventional trials, that differences in the rates of change in adjusted lung density predict concordant differences in changes in one or more endpoints directly reflective of how patients feel, function, or survive [68,69] and acceptance of a reasonably-established minimum clinically important difference. For future trials, the author encourages use of adaptive trial designs [70] and enrichment strategies [71,72,] to enroll subjects who are most likely to benefit from augmentation therapy. For example, limiting enrollment to fast decliners in FEV₁ % predicted would improve the likelihood that future studies may be able to determine whether higher doses of A₁-PI than currently recommended are clinically beneficial to patients [72,73].

Funding source

This work was supported by Midnight Sun Technologies, LLC, 1577 Spring Hill RD Ste 320, Vienna, VA 22182 USA. Midnight Sun Technologies, LLC played no role in the collection, analysis and interpretation of data, in the writing of the report, or in the decision to submit the article for publication.

Declarations

Review and/or approval by an ethics committee was not needed for this study because, as a systematic review, no data for this project were collected directly from human subjects.

Informed consent was not required for this study because no data for this project were collected directly from human subjects. All data was taken from published literature in which the respective authors had obtained informed consent.

Data availability

Underlying data for this systematic review may be available from the authors of the literature reports which formed the basis of this review and are not available from the corresponding author.

Disclaimer

This article reflects the views of the author and should not be construed to represent FDA's views or policies. The author's comments are an informal communication and represent his own best judgment. These comments do not bind or obligate FDA.

CRediT authorship contribution statement

L. Ross Pierce: Writing – review & editing, Writing – original draft, Visualization, Validation, Supervision, Resources, Project administration, Methodology, Funding acquisition, Formal analysis, Data curation, Conceptualization.

Declaration of competing interest

The authors declare the following financial interests/personal relationships which may be considered as potential competing interests: L. Ross Pierce, M.D. reports financial support was provided by Midnight Sun Technologies, LLC.

Acknowledgements

The author would like to thank Drs. Prateek Shukla and Million Tegenge for their reviews of the manuscript, and Drs. Gavin Imperato, Lei Xu, Tejashri Purohit-Sheth, and Wilson Bryan for encouragement and support. The author also thanks Dr Gavin Imperato for his assistance in screening literature search output for studies meeting selection criteria for inclusion in this review, as well as Ms. Gwendolyn Halford for assistance with the literature searches.

Appendix A. Supplementary data

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.heliyon.2024.e31183>.

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