# Fruit and vegetable intake and prostate cancer risk in the European Prospective Investigation into Cancer and Nutrition (EPIC) 

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Abbreviations: BMI: body mass index; CIs: confidence intervals; EPIC: European Prospective Investigation into Cancer and Nutrition; FFQ: food-frequency questionnaire; ICD: International Statistical Classification of Diseases, Injuries and Causes of Death; HRs: hazard ratios; ORs: odds ratios; PCa: prostate cancer; PSA: prostate-specific antigen; TNM: tumor-node-metastasis; UK: United Kingdom; WCRF/AICR: World Cancer Research Fund/American Institute for Cancer Research
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#### Abstract

Several dietary factors have been studied in relation to prostate cancer; however, most studies have not reported on subtypes of fruit and vegetables or tumor characteristics, and results obtained so far are inconclusive. This study aimed to examine the prospective association of total and subtypes of fruit and vegetable intake with the incidence of prostate cancer overall, by grade and stage of disease, and prostate cancer death. Lifestyle information for 142,239 men participating in the European Prospective Investigation into Cancer and Nutrition from 8 European countries was collected at baseline. Multivariable Cox regression models were used to estimate hazard ratios (HRs) and $95 \%$ confidence intervals (Cls). After an average follow-up time of 13.9 years, 7,036 prostate cancer cases were identified. Compared with the lowest fifth, those in the highest fifth of total fruit intake had a significantly reduced prostate cancer risk ( $\mathrm{HR}=0.91 ; 95 \% \mathrm{CI}=0.83-0.99$; $p$-trend $=0.01$ ). No associations between fruit subtypes and prostate cancer risk were observed, except for citrus fruits, where a significant trend was found ( $\mathrm{HR}=0.94 ; 95 \% \mathrm{Cl}=0.86-1.02 ; p$-trend $=0.01$ ). No associations between total and subtypes of vegetables and prostate cancer risk were observed. We found no evidence of heterogeneity in these associations by tumor grade and stage, with the exception of significant heterogeneity by tumor grade ( $p_{\text {heterogeneity }}<0.001$ ) for leafy vegetables. No significant associations with prostate cancer death were observed. The main finding of this prospective study was that a higher fruit intake was associated with a small reduction in prostate cancer risk. Whether this association is causal remains unclear.


## What's new?

The role of diet in prostate-cancer etiology is uncertain, and associations may vary by tumor characteristics. In this prospective, longitudinal study, the authors examined the association of total and subtypes of fruit and vegetable intake with the overall incidence of prostate cancer. They then analyzed incidence by grade, stage of disease, and prostate-cancer death. They found that higher fruit intake was associated with a small reduction in prostate cancer risk, and that this association did not differ by tumor characteristics.

Prostate cancer is the second most frequently diagnosed cancer in men worldwide. ${ }^{1}$ Due to the wide international variation in prostate cancer incidence, environmental and lifestyle factors, such as dietary factors, have been proposed as possible risk factors for the disease. ${ }^{2}$ However, the role of diet in the development of prostate cancer is uncertain, and associations may vary depending on the grade or stage of the disease. ${ }^{3}$

Several dietary factors have been studied in relation to prostate cancer, among them fruit and vegetable intake, ${ }^{4}$ but studies vary in design, most of them do not report results subdivided by grade and stage of prostate cancer or fruit and
vegetable subtypes, and the results obtained so far are inconclusive. ${ }^{4-9}$ High fruit and vegetable intake may be associated with reduced deoxyribonucleic acid (DNA) oxidation, cell damage, and low-grade inflammation, as well as increasing the activity of detoxification enzymes. ${ }^{10}$ Furthermore, different types of fruit and vegetables consumed may provide different bioactive compounds with different actions on prostate cancer. ${ }^{4,11}$ In recent years, special attention has been paid to the possible protective effect of the consumption of cruciferous vegetables and tomatoes on prostate cancer, although to date results are not conclusive. ${ }^{12,13}$ Regarding fruit consumption, most prospective studies have found no association with
prostate cancer development. ${ }^{4-6,9}$ The latest meta-analysis from the World Cancer Research Fund/American Institute for Cancer Research (WCRF/AICR) stated that no conclusion could be reached on whether fruit and vegetable consumption is associated with prostate cancer risk ${ }^{3}$; this metaanalysis did not differentiate between grade and stage of the disease (stage and grade are grouped together as advanced/ high grade or non-advanced/low grade), and only included a limited number of fruit and vegetable subtypes (namely cruciferous vegetables and tomatoes); therefore, more studies are needed.

The aim of this study was to examine the association of fruit and vegetable consumption with prostate cancer risk in the European Prospective Investigation into Cancer and Nutrition (EPIC). In this cohort, based on analyses of the first 1,104 cases, it was previously reported that total consumption of fruits and vegetables was not related to the development of total prostate cancer. ${ }^{14}$ However, there were insufficient data to investigate whether the association varied by tumor characteristics, whether there was an association with prostate cancer death, or to investigate the associations with different fruit and vegetable subtypes. Therefore, our aim is to report results from an extension of this earlier work, including now 7,036 prostate cancer cases, as well as to analyze the main fruit and vegetable subtypes, and to describe whether any associations differ by tumor grade or stage, and prostate cancer death.

## Material and Methods

## Subjects and study design

EPIC is an ongoing multicenter, prospective cohort study investigating the relationships of dietary and lifestyle factors with cancer and other chronic diseases carried out in 23 centers in 10 European countries: Denmark, France, Germany, Greece, Italy, The Netherlands, Norway, Spain, Sweden and the United Kingdom (UK). Participants were mostly recruited from the general population. The full cohort comprises 519,978 participants including 153,457 men, most aged $35-70$ years, recruited between 1992 and 2000. The details of the study design used in the EPIC study have been described elsewhere. ${ }^{15}$

A study flowchart is shown in Supporting Information Figure 1. Since only women were recruited in France, Norway, Naples (Italy) and Utrecht (The Netherlands), these centers where not included in the current study, and data from 19 centers in 8 countries were included in this analysis. A total of 367,898 women were excluded. We also excluded men who were diagnosed with cancer (except non-melanoma skin cancer) before recruitment ( $n=3,972$ ), those with missing dates of prostate cancer diagnosis $(n=14)$ or follow-up ( $n=1,433$ ), those aged $<20$ years at recruitment ( $n=2$ ), as well as those who had no non-dietary or dietary data, or men with an extreme energy intake in relation to estimated requirement $(n=5,766) .{ }^{16}$ Finally, a total of 142,239 men were available for analysis.

## Assessment of dietary intake and other predictor variables

At baseline, information was collected on lifestyle, health status, socio-demographic characteristics, anthropometry and medical history. ${ }^{15}$ Information on dietary intake during the year before recruitment was collected by country- or centerspecific validated dietary questionnaires. Most centers used self-administered dietary questionnaires, however, in Greece, Ragusa (Italy), and Spain participants were interviewed by trained staff members. In Malmö (Sweden) dietary intake was assessed using a modified diet history method which combined information from a self-administered diet questionnaire, a 7 -day food registration and a 1-hr interview. With the purpose of improving the comparability of dietary data across the participating centers, dietary intakes from the questionnaires were calibrated using a standardized, computer-based, $24-\mathrm{hr}$ dietary recall method in an $8 \%$ random sample of the whole EPIC cohort. Information on validation of the dietary questionnaires has been published previously. ${ }^{17}$

On the basis of this information, daily fruit (fresh fruit only) and vegetable (potatoes and dried beans not included) consumption was based on the food group classification previously used for the EPIC cohort. ${ }^{18}$ In this study, fruit and vegetables were in turn classified into sub-groups. The subtypes of fruit considered were citrus fruits (e.g., oranges, lemons), apples and pears, and bananas. For vegetables, the following sub-groups were used: cruciferous vegetables (e.g., broccoli, cabbage), leafy vegetables (e.g., spinach, lettuce), fruiting vegetables (e.g., tomato, sweet pepper, eggplant), tomatoes (raw, cooked, sauce) and root vegetables (e.g., carrot, beetroot).

## Endpoints

The main source of information on prostate cancer cases was population-based cancer registries. In Germany and Greece follow-up was based on a combination of methods, including health insurance records, cancer and pathology registries, as well as active follow-up through participants or relatives; selfreported incident cancers were verified through medical records. Follow-up began at the date of recruitment and was censored at the date of last known contact, or at the date of diagnosis of cancer, death, emigration or the end of the follow-up period, whichever came first. A total of 7,036 men developed prostate cancer (code: C61) according to the 10th Revision of the International Statistical Classification of Diseases, Injuries and Causes of Death (ICD). ${ }^{19}$

Information on stage [tumor-node-metastasis (TNM) staging code] and grade of prostate cancer (based on Gleason sum) was collected from each center, where possible. Grade was stratified as low-intermediate (Gleason score of $<8$, or grade coded as well, moderately, or poorly differentiated; $n=3,757$ ) or high (Gleason score of $\geq 8$, or grade coded as undifferentiated; $n=726$ ) grade. Localized stage included those confined within the prostate and with no metastases at

Table 1. Baseline characteristics of 142,239 men in EPIC (1992-2013) according to observed fruit and vegetable intake


Percentages do not sum to $100 \%$ due to missing data.
${ }^{1}$ Values are means (SD).
diagnosis (TNM staging score of $\leq T_{2}$ and $N_{0} / N_{x}$ and $M_{0}$, or stage coded in the recruitment center as localized; $n=2,641$ ). Advanced cases included tumors that had spread beyond the prostate at diagnosis ( $\mathrm{T}_{3}-\mathrm{T}_{4}$ and/or $\mathrm{N}_{1}-\mathrm{N}_{3}$ and/or $\mathrm{M}_{1}$, and/or stage coded in the recruitment center as metastatic; $n=1,389)$. Fatal cases were those who died of prostate cancer $(n=936)$.

## Statistical analysis

Hazard ratios (HRs) and 95\% confidence intervals (CIs) were estimated using Cox proportional hazards models using age as the underlying time variable. The date of last follow-up ranged from January 2011 in Germany to October 2013 in Spain. All analyses were stratified by center and age $(<50$, $50-54.9,55-59.9,60-64.9,65-69.9$ and $\geq 70$ years) at recruitment. To check for violation of the proportional hazards assumption we used time-varying covariates and Schoenfeld residuals, which indicated no evidence of deviation from the proportional hazards assumption. Fruit and vegetable intakes (g/day) were divided into fifths based on the distribution in
the EPIC cohort and also modeled as continuous variables in increments of $100 \mathrm{~g} /$ day (approximately equivalent to one portion). Tests for linear trend were performed using a pseudo-continuous variable equal to the median value in each fifth of intake divided by 100 . All models were adjusted for educational level (no degree or equivalent, degree or equivalent, unknown), smoking status (never, former, current, unknown), marital status (married or cohabiting, not married or cohabiting, unknown), diabetes (no, yes, unknown), physical activity (inactive, moderately inactive, moderately active, active, unknown ${ }^{20}$ ), height $(<170,170-174,175-$ $179, \geq 180 \mathrm{~cm}$, unknown), BMI ( $<22.5,22.5-24.9,25-$ $29.9, \geq 30 \mathrm{~kg} / \mathrm{m}^{2}$, unknown), and total energy intake (fifths). Participants with missing values were assigned an "unknown" category; $<3 \%$ of values were missing for each covariate, with the exception of marital status, for which $30 \%$ of values were missing.

Tests for heterogeneity of trends for case-defined characteristics [histologic grade (low-intermediate or high), tumor stage (localized or advanced), age at diagnosis ( $<65$ years,

Table 2. Multivariable-adjusted hazard ratios $(95 \% \mathrm{CI})$ for total prostate cancer by fifths of observed fruit intake in 142,239 men in EPIC (1992-2013)

|  | Fifths of observed fruit intake |  |  |  |  | $p$-trend ${ }^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 |  |
| Total fruit, g/day | $\leq 66.7$ | $>66.7$ to $\leq 123.6$ | $>123.6$ to $\leq 197.3$ | $>197.3$ to $\leq 320.0$ | >320.0 |  |
| Cases, $n$ | 1,420 | 1,540 | 1,556 | 1,419 | 1,101 |  |
| HR (95\% CI) | 1 ref | 1.04 (0.97-1.12) | 1.02 (0.95-1.10) | 1.01 (0.94-1.09) | 0.93 (0.86-1.02) | 0.04 |
| Adjusted HR (95\% CI) ${ }^{1}$ | 1 ref | 1.03 (0.95-1.10) | 1.00 (0.93-1.08) | 0.99 (0.91-1.07) | 0.91 (0.83-0.99) | 0.01 |
| Citrus fruit, g/day | $\leq 5.8$ | $>5.8$ to $\leq 14.3$ | $>14.3$ to $\leq 36.8$ | $>36.8$ to $\leq 78.4$ | $>78.4$ |  |
| Cases, $n$ | 1,551 | 1,494 | 1,528 | 1,341 | 1,122 |  |
| HR (95\% CI) | 1 ref | 1.08 (1.00-1.16) | 1.04 (0.97-1.12) | 0.96 (0.90-1.04) | 0.96 (0.88-1.04) | 0.03 |
| Adjusted HR (95\% CI) ${ }^{1}$ | 1 ref | 1.06 (0.99-1.14) | 1.02 (0.95-1.10) | 0.95 (0.88-1.02) | 0.94 (0.86-1.02) | 0.01 |
| Apple/pear, g/day | $\leq 9.0$ | $>9.0$ to $\leq 28.0$ | $>28.0$ to $\leq 62.0$ | $>62.0$ to $\leq 116.4$ | >116.4 |  |
| Cases, $n$ | 1,238 | 1,465 | 1,383 | 1,436 | 1,514 |  |
| HR (95\% CI) | 1 ref | 1.09 (1.01-1.18) | 1.04 (0.96-1.13) | 1.07 (0.99-1.17) | 1.04 (0.96-1.13) | 0.9 |
| Adjusted HR (95\% CI) ${ }^{1}$ | 1 ref | 1.08 (1.00-1.17) | 1.03 (0.95-1.12) | 1.06 (0.97-1.15) | 1.02 (0.94-1.11) | 0.7 |
| Banana, g/day | $\leq 0.6$ | $>0.6$ to $\leq 7.1$ | $>7.1$ to $\leq 15.0$ | $>15.0$ to $\leq 43.0$ | >43.0 |  |
| Cases, $n$ | 1,270 | 1,357 | 1,430 | 1,496 | 1,483 |  |
| HR (95\% CI) | 1 ref | 1.10 (1.01-1.20) | 1.09 (1.01-1.19) | 1.10 (1.01-1.19) | 1.07 (0.99-1.17) | 0.8 |
| Adjusted HR (95\% CI) ${ }^{1}$ | 1 ref | 1.09 (1.00-1.19) | 1.08 (0.99-1.17) | 1.08 (0.99-1.17) | 1.05 (0.96-1.14) | 0.8 |

Cox regression analysis. All models are adjusted for age (underlying time variable) and stratified by recruitment center and age at recruitment.
${ }^{1}$ Additionally adjusted for educational level (no degree, degree, unknown), smoking status (never, former, current, unknown), marital status (married, not married, unknown), diabetes (yes, no, unknown), physical activity (inactive, moderately inactive, moderately active, active, unknown), height ( $<170,170-174,175-179, \geq 180 \mathrm{~cm}$, unknown), body mass index ( $<22.5,22.5-24.9,25-29.9, \geq 30 \mathrm{~kg} / \mathrm{m}^{2}$, unknown), and total energy intake (fifths).
${ }^{2} p$-values for trend were obtained using a pseudo-continuous variable equal to the median value in each fifth of intake.
$\geq 65$ years), and time between blood collection and diagnosis ( $<5$ years, $\geq 5$ years)] were performed. For this, we fitted stratified Cox models based on competing risks and compared the risk coefficients and standard errors in the subgroups of interest after excluding cases of unknown grade or stage. ${ }^{21}$ For the non-case-defined factors [age at recruitment ( $<65$ years, $\geq 65$ years), BMI ( $<25 \mathrm{~kg} / \mathrm{m}^{2}, \geq 25 \mathrm{~kg} / \mathrm{m}^{2}$ ), and country (Denmark, Germany, Greece, Italy, The Netherlands, Spain, Sweden, UK)], the test for heterogeneity was assessed by using likelihood ratio tests to compare the Cox models with and without interaction terms for the dietary variable and the relevant factor. All analyses were performed using Stata version 14.1 (Stata Corporation, College Station, TX), all tests of significance were two-sided, and a $p$ values $<0.05$ was considered statistically significant.

## Results

After an average of 13.9 years of follow-up, 7,036 men were diagnosed with prostate cancer among the 142,239 participants included in this study. The median age at prostate cancer diagnosis was 68 years (range, 41-95 years). Baseline details of participants are shown in Table 1. Some characteristics varied by fruit and vegetable consumption. For example, men in the highest fifths of fruit and vegetable intake were older at recruitment and at diagnosis, less likely to
smoke, more likely be diabetic, were shorter and with a higher BMI and energy intake.

The estimated HRs for total prostate cancer risk across fifths of fruit intake, overall and for specific types of fruit are shown in Table 2. The only statistically significant association was observed for total fruit intake: compared with men in the lowest fifth of total fruit, the men in the highest fifth of total fruit intake had a HR 0.91 ( $95 \%$ CI: 0.83-0.99; ptrend $=0.01$ ). There was also some evidence of a significant trend for citrus fruits (HR in highest fifth $=0.94 ; 0.86-1.02$; $p$-trend $=0.01$ ) although the risk in the highest fifth of intake was not statistically significant. No associations with total prostate cancer risk were observed for other subtypes of fruits (apples and pears, and bananas).

No association was found between total prostate cancer risk and vegetable intake, overall ( $\mathrm{HR}=1.02,0.93-1.12$ ) or by subtypes of vegetables (cruciferous vegetables, leafy vegetables, fruiting vegetables, tomatoes, root vegetables; Table 3).

The associations of fruit and vegetable intake with risk for total prostate cancer, prostate cancer subdivided by grade and stage of disease, and for prostate cancer death, using both the observed and calibrated intakes, are shown in Tables 4 and 5. Results for observed and calibrated intake were similar in direction. There was a weak significant association between observed and calibrated total fruit intake and risk of total prostate cancer; an increase of $100 \mathrm{~g} /$ day was

Table 3. Multivariable-adjusted hazard ratios ( $95 \% \mathrm{Cl}$ ) for prostate cancer by fifth of observed vegetable intake in 142,239 men in EPIC (1992-2013)

|  | Fifths of observed vegetable intake |  |  |  |  | $p$-trend ${ }^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 |  |
| Total vegetables, g/day | $\leq 82.8$ | $>82.8$ to $\leq 126.4$ | $>126.4$ to $\leq 182.1$ | $>182.1$ to $\leq 281.6$ | >281.7 |  |
| Cases, $n$ | 1,526 | 1,451 | 1,457 | 1,549 | 1,053 |  |
| HR (95\% CI) | 1 ref | 1.04 (0.96-1.12) | 1.01 (0.93-1.09) | 1.07 (0.99-1.16) | 1.04 (0.95-1.14) | 0.4 |
| Adjusted HR (95\% CI) ${ }^{1}$ | 1 ref | 1.03 (0.96-1.11) | 0.99 (0.92-1.07) | 1.05 (0.97-1.14) | 1.02 (0.93-1.12) | 0.6 |
| Cruciferous vegetables, g/day | $\leq 3.0$ | $>3.0$ to $\leq 9.2$ | $>9.2$ to $\leq 18.7$ | $>18.7$ to $\leq 36.6$ | >36.6 |  |
| Cases, $n$ | 1,457 | 1,287 | 1,253 | 1,190 | 1,207 |  |
| HR (95\% CI) | 1 ref | 1.06 (0.97-1.15) | 1.05 (0.96-1.15) | 1.07 (0.98-1.16) | 1.06 (0.96-1.18) | 0.4 |
| Adjusted HR (95\% CI) ${ }^{1}$ | 1 ref | 1.06 (0.97-1.15) | 1.04 (0.95-1.13) | 1.05 (0.96-1.15) | 1.06 (0.96-1.17) | 0.5 |
| Leafy vegetables, g/day | $\leq 1.6$ | $>1.6$ to $\leq 6.0$ | $>6.0$ to $\leq 15.1$ | $>15.1$ to $\leq 36.9$ | >36.9 |  |
| Cases, $n$ | 1,773 | 1,453 | 1,218 | 1,034 | 916 |  |
| HR (95\% CI) | 1 ref | 1.03 (0.96-1.11) | 1.09 (1.00-1.18) | 1.04 (0.95-1.14) | 1.09 (0.97-1.22) | 0.3 |
| Adjusted HR (95\% CI) ${ }^{1}$ | 1 ref | 1.01 (0.94-1.09) | 1.06 (0.97-1.16) | 1.01 (0.92-1.11) | 1.06 (0.95-1.19) | 0.4 |
| Fruiting vegetables, g/day | $\leq 21.6$ | $>21.6$ to $\leq 36.5$ | $>36.5$ to $\leq 56.3$ | $>56.3$ to $\leq 97.8$ | >97.8 |  |
| Cases, $n$ | 1,642 | 1,590 | 1,465 | 1,365 | 974 |  |
| HR (95\% CI) | 1 ref | 1.07 (1.00-1.15) | 1.05 (0.97-1.13) | 1.03 (0.95-1.11) | 1.04 (0.95-1.14) | 0.8 |
| Adjusted HR (95\% CI) ${ }^{1}$ | 1 ref | 1.06 (0.99-1.14) | 1.04 (0.96-1.12) | 1.01 (0.94-1.10) | 1.03 (0.94-1.13) | 0.9 |
| Tomatoes, g/day | $\leq 9.0$ | $>9.0$ to $\leq 18.9$ | $>18.9$ to $\leq 30.8$ | $>30.8$ to $\leq 67.3$ | $>67.3$ |  |
| Cases, $n$ | 1,612 | 1,540 | 1,443 | 1,480 | 961 |  |
| HR (95\% CI) | 1 ref | 1.05 (0.97-1.13) | 1.04 (0.96-1.12) | 1.01 (0.93-1.09) | 1.09 (0.98-1.21) | 0.2 |
| Adjusted HR (95\% CI) ${ }^{1}$ | 1 ref | 1.04 (0.97-1.12) | 1.03 (0.95-1.11) | 1.00 (0.92-1.08) | 1.08 (0.97-1.20) | 0.2 |
| Root vegetables, g/day | $\leq 3.8$ | $>3.8$ to $\leq 8.9$ | $>8.9$ to $\leq 16.3$ | $>16.3$ to $\leq 35.5$ | >35.5 |  |
| Cases, $n$ | 1,416 | 1,371 | 1,353 | 1,334 | 1,562 |  |
| HR (95\% CI) | 1 ref | 1.01 (0.93-1.09) | 1.02 (0.94-1.10) | 1.01 (0.93-1.09) | 1.04 (0.96-1.13) | 0.3 |
| Adjusted HR (95\% CI) ${ }^{1}$ | 1 ref | 1.00 (0.92-1.08) | 1.00 (0.92-1.08) | 0.99 (0.92-1.07) | 1.02 (0.94-1.11) | 0.5 |

Cox regression analysis. All models are adjusted for age (underlying time variable) and stratified by recruitment center and age at recruitment.
${ }^{1}$ Additionally adjusted for educational level (no degree, degree, unknown), smoking status (never, former, current, unknown), marital status (married, not married, unknown), diabetes (yes, no, unknown), physical activity (inactive, moderately inactive, moderately active, active, unknown), height ( $<170,170-174,175-179, \geq 180 \mathrm{~cm}$, unknown), body mass index ( $<22.5,22.5-24.9,25-29.9, \geq 30 \mathrm{~kg} / \mathrm{m}^{2}$, unknown), and total energy intake (fifths).
${ }^{2} p$-values for trend were obtained using a pseudo-continuous variable equal to the median value in each fifth of intake.
associated with lower risk of total prostate cancer [3\% lower risk ( $95 \% \mathrm{CI}=0.95-0.99$ ) for observed and $4 \% ~(95 \% \mathrm{CI}=0.94-$ 0.99 ) lower risk for calibrated intake]. An increase of $100 \mathrm{~g} /$ day in citrus fruit intake was related to lower total prostate cancer risk ( $\mathrm{HR}=0.92,0.86-0.98$ for observed and $\mathrm{HR}=0.88,0.80-$ 0.97 for calibrated intake; $p$-trend $=0.01$ and 0.009 , respectively). There was no evidence of heterogeneity in separate analyses by grade and stage, with the exception of significant heterogeneity in the association by tumor grade for leafy vegetables for both observed intake ( $p_{\text {heterogeneity }}<0.001$; $\mathrm{HR}=0.94,0.76-1.17$ for low-intermediate grade and $\mathrm{HR}=2.66$, 1.54-4.58 for high grade cancer) and calibrated intake ( $p_{\text {heterogeneity }}<0.001 ; \mathrm{HR}=0.97,0.61-1.55$ for low-intermediate grade and $\mathrm{HR}=8.98,2.79-28.97$ for high grade cancer). We observed no associations between fruit or vegetable intake (both overall and for specific subtypes) and prostate cancer death.

There was no significant heterogeneity for the association between fruit or vegetable intake and total prostate cancer risk when subdivided by age at recruitment ( $<65$ years, $\geq 65$ years), age at diagnosis ( $<65$ years, $\geq 65$ years), by time between recruitment and diagnosis ( $<5$ years, $\geq 5$ years), BMI ( $<25 \mathrm{~kg} / \mathrm{m}^{2}, \geq 25 \mathrm{~kg} / \mathrm{m}^{2}$ ), and country (Denmark, Germany, Greece, Italy, The Netherlands, Spain, Sweden, UK) (Supporting Information Tables 1 and 2).

## Discussion

In this large European prospective study, a higher intake of total fruit was associated with a small reduction in prostate cancer risk, while vegetable consumption was not related to prostate cancer risk. When consumption was analyzed according to fruit and vegetable subtypes, we found that citrus fruits were weakly associated with a reduced risk of

Table 4. Multivariable-adjusted hazard ratios ( $95 \% \mathrm{Cl}$ ) for prostate cancer per unit increase (per $100 \mathrm{~g} /$ day) of fruit intake in 142,239 men in EPIC (1992-2013)

|  | No. of cases | Observed intake |  |  | Calibrated intake |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | HR (95\% CI) ${ }^{1}$ | $P$-trend ${ }^{2}$ | $P$ for het. ${ }^{3}$ | HR (95\% CI) ${ }^{1}$ | $P$-trend ${ }^{2}$ | $P$ for het. ${ }^{3}$ |
| Total fruit |  |  |  |  |  |  |  |
| Total PCa | 7,036 | 0.97 (0.95-0.99) | 0.01 |  | 0.96 (0.94-0.99) | 0.006 |  |
| Grade |  |  |  |  |  |  |  |
| Low-intermediate | 3,757 | 0.98 (0.96-1.01) | 0.3 |  | 0.98 (0.94-1.01) | 0.2 |  |
| High | 726 | 0.98 (0.92-1.04) | 0.4 | 0.8 | 0.97 (0.90-1.05) | 0.4 | 0.8 |
| Stage |  |  |  |  |  |  |  |
| Localized | 2,641 | 0.96 (0.93-0.99) | 0.02 |  | 0.95 (0.91-0.99) | 0.02 |  |
| Advanced | 1,389 | 0.99 (0.95-1.04) | 0.7 | 0.3 | 0.99 (0.93-1.04) | 0.6 | 0.3 |
| PCa death | 936 | 0.98 (0.92-1.03) | 0.4 |  | 0.97 (0.91-1.04) | 0.5 |  |
| Citrus fruit |  |  |  |  |  |  |  |
| Total PCa | 7,036 | 0.92 (0.86-0.98) | 0.01 |  | 0.88 (0.80-0.97) | 0.009 |  |
| Grade |  |  |  |  |  |  |  |
| Low-intermediate | 3,757 | 0.90 (0.82-0.99) | 0.03 |  | 0.85 (0.75-0.97) | 0.02 |  |
| High | 726 | 0.89 (0.77-1.15) | 0.6 | 0.7 | 0.93 (0.70-1.25) | 0.6 | 0.6 |
| Stage |  |  |  |  |  |  |  |
| Localized | 2,641 | 0.90 (0.81-1.00) | 0.05 |  | 0.84 (0.72-0.98) | 0.03 |  |
| Advanced | 1,389 | 0.90 (0.77-1.05) | 0.2 | 0.9 | 0.86 (0.69-1.08) | 0.2 | 0.9 |
| PCa death | 936 | 0.92 (0.77-1.11) | 0.4 |  | 0.89 (0.69-1.16) | 0.4 |  |
| Apple/pear |  |  |  |  |  |  |  |
| Total PCa | 7,036 | 0.99 (0.95-1.03) | 0.7 |  | 0.99 (0.93-1.04) | 0.6 |  |
| Grade |  |  |  |  |  |  |  |
| Low-intermediate | 3,757 | 1.01 (0.95-1.07) | 0.8 |  | 1.01 (0.93-1.08) | 0.9 |  |
| High | 726 | 1.00 (0.88-1.14) | 0.9 | 0.9 | 1.00 (0.85-1.17) | 0.9 | 0.9 |
| Stage |  |  |  |  |  |  |  |
| Localized | 2,641 | 0.98 (0.90-1.05) | 0.5 |  | 0.98 (0.89-1.07) | 0.6 |  |
| Advanced | 1,389 | 1.04 (0.95-1.15) | 0.4 | 0.3 | 1.04 (0.92-1.17) | 0.5 | 0.4 |
| PCa death | 936 | 0.94 (0.83-1.05) | 0.3 |  | 0.93 (0.81-1.07) | 0.3 |  |
| Banana |  |  |  |  |  |  |  |
| Total PCa | 7,036 | 0.99 (0.91-1.08) | 0.8 |  | 1.01 (0.85-1.19) | 0.9 |  |
| Grade |  |  |  |  |  |  |  |
| Low-intermediate | 3,757 | 0.97 (0.86-1.10) | 0.6 |  | 0.99 (0.78-1.26) | 0.9 |  |
| High | 726 | 1.06 (0.81-1.38) | 0.7 | 0.6 | 1.13 (0.67-1.90) | 0.6 | 0.6 |
| Stage |  |  |  |  |  |  |  |
| Localized | 2,641 | 0.95 (0.82-1.10) | 0.5 |  | 0.94 (0.71-1.25) | 0.7 |  |
| Advanced | 1,389 | 0.92 (0.76-1.12) | 0.4 | 0.8 | 0.90 (0.62-1.32) | 0.6 | 0.8 |
| PCa death | 936 | 1.07 (0.86-1.34) | 0.5 |  | 1.16 (0.75-1.80) | 0.5 |  |

PCa: prostate cancer.
Cox regression analysis. All models are stratified by center and age at recruitment and adjusted for age (underlying time variable), educational level (no degree, degree, unknown), smoking status (never, former, current, unknown), marital status (married, not married, unknown), diabetes (yes, no, unknown), physical activity (inactive, moderately inactive, moderately active, active, unknown), height ( $<170,170-174,175-179, \geq 180 \mathrm{~cm}$, unknown), body mass index ( $<22.5,22.5-24.9,25-29.9, \geq 30 \mathrm{~kg} / \mathrm{m}^{2}$, unknown), and total energy intake (fifths).
${ }^{1} \mathrm{HR}(95 \% \mathrm{Cl})$ estimated per $100 \mathrm{~g} /$ day unit increase in fruit intake.
${ }^{2} p$-values for trend were obtained using a pseudo-continuous variable equal to the median value in each fifth of intake.
${ }^{3} p$ values from test for heterogeneity for the associations of fruit intake with risk of prostate cancer categorized according to prostate tumor grade (low-intermediate or high) and stage (localized or advanced).
Low-intermediate grade (Gleason score of $<8$, or grade coded as well, moderately, or poorly differentiated). High grade (Gleason score of $\geq 8$, or grade coded as undifferentiated). Localized stage (TNM staging score of TO-T2 and NO/Nx and MO, or stage coded in the recruitment center as localized). Advanced stage (T3-T4 and/or N1-N3 and/or M1, and/or stage coded in the recruitment center as metastatic).

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Table 5. Multivariable-adjusted hazard ratios ( $95 \% \mathrm{Cl}$ ) for prostate cancer per unit increase (per $100 \mathrm{~g} /$ day) of vegetable intake in 142,239 men in EPIC (1992-2013)

|  | No. of cases | Observed intake |  |  | Calibrated intake |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | HR (95\% CI) ${ }^{1}$ | $p$-trend ${ }^{2}$ | $p$ for het. ${ }^{3}$ | HR (95\% CI) ${ }^{1}$ | $P$-trend ${ }^{2}$ | $P$ for het. ${ }^{3}$ |
| Total vegetables |  |  |  |  |  |  |  |
| Total PCa | 7,036 | 1.01 (0.98-1.03) | 0.6 |  | 1.01 (0.95-1.07) | 0.7 |  |
| Grade |  |  |  |  |  |  |  |
| Low-intermediate | 3,757 | 0.99 (0.95-1.03) | 0.5 |  | 0.97 (0.89-1.05) | 0.5 |  |
| High | 726 | 1.08 (1.00-1.17) | 0.06 | 0.05 | 1.19 (1.00-1.41) | 0.06 | 0.04 |
| Stage |  |  |  |  |  |  |  |
| Localized | 2,641 | 1.02 (0.98-1.07) | 0.3 |  | 1.05 (0.95-1.16) | 0.4 |  |
| Advanced | 1,389 | 1.02 (0.96-1.08) | 0.6 | 0.9 | 1.03 (0.90-1.18) | 0.7 | 0.8 |
| PCa death | 936 | 1.05 (0.98-1.13) | 0.2 |  | 1.11 (0.95-1.30) | 0.2 |  |
| Cruciferous vegetables |  |  |  |  |  |  |  |
| Total PCa | 6,394 | 1.06 (0.90-1.23) | 0.50 |  | 1.08 (0.76-1.54) | 0.7 |  |
| Grade |  |  |  |  |  |  |  |
| Low-intermediate | 3,281 | 0.97 (0.77-1.21) | 0.8 |  | 0.78 (0.46-1.34) | 0.4 |  |
| High | 647 | 1.26 (0.75-2.09) | 0.4 | 0.3 | 2.10 (0.58-7.68) | 0.3 | 0.2 |
| Stage |  |  |  |  |  |  |  |
| Localized | 2,085 | 1.03 (0.78-1.36) | 0.8 |  | 0.90 (0.47-1.70) | 0.7 |  |
| Advanced | 1,303 | 1.04 (0.72-1.50) | 0.8 | 0.9 | 1.11 (0.43-2.89) | 0.8 | 0.7 |
| PCa death | 882 | 1.30 (0.86-1.96) | 0.2 |  | 1.72 (0.65-4.53) | 0.3 |  |
| Leafy vegetables |  |  |  |  |  |  |  |
| Total PCa | 6,394 | 1.07 (0.91-1.25) | 0.4 |  | 1.19 (0.86-1.66) | 0.3 |  |
| Grade |  |  |  |  |  |  |  |
| Low-intermediate | 3,281 | 0.94 (0.76-1.17) | 0.6 |  | 0.97 (0.61-1.55) | 0.9 |  |
| High | 647 | 2.66 (1.54-4.58) | <0.001 | $<0.001$ | 8.98 (2.79-28.97) | $<0.001$ | $<0.001$ |
| Stage |  |  |  |  |  |  |  |
| Localized | 2,085 | 1.10 (0.82-1.48) | 0.5 |  | 1.37 (0.74-2.54) | 0.3 |  |
| Advanced | 1,303 | 1.19 (0.78-1.79) | 0.4 | 0.8 | 1.53 (0.66-3.54) | 0.3 | 0.8 |
| PCa death | 882 | 1.26 (0.79-2.00) | 0.3 |  | 1.95 (0.79-4.83) | 0.1 |  |
| Fruiting vegetables |  |  |  |  |  |  |  |
| Total PCa | 7,036 | 1.00 (0.95-1.06) | 0.9 |  | 1.00 (0.92-1.10) | 0.9 |  |
| Grade |  |  |  |  |  |  |  |
| Low-intermediate | 3,757 | 1.01 (0.94-1.09) | 0.8 |  | 1.01 (0.90-1.15) | 0.8 |  |
| High | 726 | 1.03 (0.86-1.24) | 0.7 | 0.8 | 1.07 (0.80-1.44) | 0.6 | 0.7 |
| Stage |  |  |  |  |  |  |  |
| Localized | 2,641 | 1.03 (0.94-1.14) | 0.5 |  | 1.04 (0.89-1.21) | 0.6 |  |
| Advanced | 1,389 | 1.03 (0.89-1.19) | 0.7 | 0.9 | 1.06 (0.85-1.33) | 0.6 | 0.8 |
| PCa death | 936 | 1.07 (0.91-1.25) | 0.4 |  | 1.12 (0.88-1.43) | 0.3 |  |
| Tomatoes |  |  |  |  |  |  |  |
| Total PCa | 7,036 | 1.05 (0.97-1.15) | 0.2 |  | 1.09 (0.96-1.24) | 0.2 |  |
| Grade |  |  |  |  |  |  |  |
| Low-intermediate | 3,757 | 1.01 (0.90-1.14) | 0.8 |  | 1.03 (0.86-1.23) | 0.7 |  |
| High | 726 | 1.10 (0.82-1.48) | 0.5 | 0.6 | 1.30 (0.84-2.00) | 0.2 | 0.3 |
| Stage |  |  |  |  |  |  |  |
| Localized | 2,641 | 1.00 (0.86-1.16) | 0.9 |  | 1.03 (0.83-1.28) | 0.8 |  |
| Advanced | $1,389$ | 1.07 (0.86-1.33) | 0.6 | 0.6 | 1.18 (0.86-1.63) | 0.3 | 0.5 |

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Table 5. Multivariable-adjusted hazard ratios ( $95 \% \mathrm{CI}$ ) for prostate cancer per unit increase (per $100 \mathrm{~g} /$ day) of vegetable intake in 142,239 men in EPIC (1992-2013) (Continued)

|  | No. of cases | Observed intake |  |  | Calibrated intake |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | HR (95\% CI) ${ }^{1}$ | $p$-trend ${ }^{2}$ | $p$ for het. ${ }^{3}$ | HR (95\% CI) ${ }^{1}$ | $P$-trend ${ }^{2}$ | $P$ for het. ${ }^{3}$ |
| PCa death | 936 | 1.20 (0.93-1.54) | 0.2 |  | 1.38 (0.97-1.97) | 0.07 |  |
| Root vegetables |  |  |  |  |  |  |  |
| Total PCa | 7,036 | 1.05 (0.91-1.21) | 0.5 |  | 1.05 (0.74-1.49) | 0.8 |  |
| Grade |  |  |  |  |  |  |  |
| Low-intermediate | 3,757 | 0.95 (0.77-1.17) | 0.6 |  | 0.76 (0.45-1.27) | 0.3 |  |
| High | 726 | 1.43 (0.92-2.22) | 0.1 | 0.09 | 2.10 (0.67-6.54) | 0.2 | 0.1 |
| Stage |  |  |  |  |  |  |  |
| Localized | 2,641 | 1.07 (0.83-1.37) | 0.6 |  | 0.99 (0.54-1.82) | 0.9 |  |
| Advanced | 1,389 | 0.89 (0.64-1.23) | 0.5 | 0.4 | 0.67 (0.29-1.54) | 0.3 | 0.5 |
| PCa death | 936 | 1.07 (0.74-1.55) | 0.7 |  | 1.27 (0.50-3.22) | 0.6 |  |

PCa: prostate cancer.
Cox regression analysis. All models are stratified by center and age at recruitment and adjusted for age (underlying time variable), educational level (no degree, degree, unknown), smoking status (never, former, current, unknown), marital status (married, not married, unknown), diabetes (yes, no, unknown), physical activity (inactive, moderately inactive, moderately active, active, unknown), height ( $<170,170-174,175-179, \geq 180 \mathrm{~cm}$, unknown), body mass index ( $<22.5,22.5-24.9,25-29.9, \geq 30 \mathrm{~kg} / \mathrm{m}^{2}$, unknown), and total energy intake (fifths).
${ }^{1} \mathrm{HR}(95 \% \mathrm{Cl})$ estimated per $100 \mathrm{~g} /$ day unit increase in vegetable intake.
${ }^{2} p$-values for trend were obtained using a pseudo-continuous variable equal to the median value in each fifth of intake.
${ }^{3} p$ values from test for heterogeneity for the associations of vegetable intake with risk of prostate cancer categorized according to prostate tumor grade (low-intermediate or high) and stage (localized or advanced).
Low-intermediate grade (Gleason score of $<8$, or grade coded as well, moderately, or poorly differentiated). High grade (Gleason score of $\geq 8$, or grade coded as undifferentiated). Localized stage (TNM staging score of TO-T2 and NO/Nx and MO, or stage coded in the recruitment center as localized). Advanced stage (T3-T4 and/or N1-N3 and/or M1, and/or stage coded in the recruitment center as metastatic).
prostate cancer. These associations did not differ by tumor grade or stage, with the exception of leafy vegetables, for which a positive association with high grade prostate cancer was found.

Although in the current large prospective study we observed an inverse association between total fruit consumption and prostate cancer development, the WCRF/AICR meta-analysis, which included a total of 16 prospective studies with 26,671 cases of prostate cancer, showed a null association between fruit consumption and total prostate cancer risk [Relative Risk per $100 \mathrm{~g} /$ day intake 1.00 ( $95 \%$ CI $0.99-$ 1.01)] in 2014. ${ }^{3}$ Only one small prospective study ( 139 incident prostate cancer cases) on fruit consumption and prostate cancer risk has been published since then, which also showed no association. ${ }^{22}$ If there is an association between fruit consumption and prostate cancer risk it might be due to the high content of vitamins (such as vitamin C) and phytochemicals (such as phenolic compounds and carotenoids), which may have anti-carcinogenic properties. ${ }^{10,23}$ However, men who have high fruit intake differ in several respects from men with a low fruit intake and therefore residual confounding cannot be excluded. In this study, total vegetable consumption was not associated with overall prostate cancer risk, which is in agreement with WCRF/AICR meta-analysis. ${ }^{3}$

When fruit and vegetables were divided into subtypes, we found a weak inverse association between citrus fruit and incidence of prostate cancer, which was significant when citrus fruit was introduced continuously per $100 \mathrm{~g} /$ day
increments. Only one ${ }^{24}$ out of the five ${ }^{5,8,9,24,25}$ studies that have assessed the association between citrus fruit and prostate cancer risk has also found a significant inverse association, and no association between circulating concentrations of the citrus biomarker $\beta$-cryptoxanthin and prostate cancer risk was observed in a pooled analysis of 10 prospective studies. ${ }^{26}$ Our non-significant associations between several vegetable subtypes and risk of prostate cancer are in line with previous prospective studies which have also found no association between overall prostate cancer risk and cruciferous ${ }^{9,12,27}$ or leafy vegetable intake. ${ }^{8,9,28,29}$ Also, although early reports linked frequent consumption of tomatoes, tomato products or lycopene (a carotenoid from tomatoes) with lower risk of overall prostate cancer, ${ }^{13,25}$ our study and the latest meta-analysis from WCRF/AICR did not support this association, ${ }^{3}$ and nor did findings from a pooled analysis of blood lycopene concentrations and overall prostate cancer risk (although there was statistically significant heterogeneity by stage of disease, and the odds ratios (ORs) for aggressive disease for the highest compared with the lowest fifth of lycopene was $0.65(95 \%$ CI: $0.46,0.91$; p-trend $=0.032) .{ }^{26}$ However, it should be highlighted that not all studies have divided fruit and vegetable intake in the same subtypes.

As far as we are aware, no other large prospective study has examined the association of fruit and vegetable intake with prostate cancer risk separately by both grade and stage of the tumor, and only two studies have analyzed this association with prostate cancer death as the outcome. ${ }^{7,30}$ The
latest WCRF/AICR meta-analysis showed the associations of fruit and vegetable intake with aggressive prostate cancer, a categorization which combined high grade, advanced stage, and prostate cancer mortality. ${ }^{3}$ In the current study we found no association between fruit or vegetable intake and prostate cancer death, or any evidence that any association differed by tumor grade or stage, with the exception of the association of leafy vegetables by tumor grade, with a positive association being limited to high grade prostate cancer only. While leafy vegetables are good sources of folate and circulating folate has been related to high grade prostate cancer, ${ }^{31}$ this association may be a chance finding given the multiple tests. To our knowledge, one previous prospective study has found a positive association between leafy vegetables intake and nonlocalized or high grade prostate cancer risk, ${ }^{9}$ but this study did not evaluate associations separately for grade and stage of the disease.

Some strengths and limitations of the present study should be considered. The major strengths of this study include the prospective design, the large number of total prostate cancer cases and prostate cancer deaths and the large amount of grade and stage information, and the reliable identification of prostate cancer cases through cancer registries and/or verified medical records. The Gleason grade was based on data available from biopsies and surgical pathology and there may be some misclassification because of changes in grading over time. The dietary questionnaires in all EPIC centers were validated and dietary intakes were calibrated using measures from a standardized $24-\mathrm{hr}$ diet recall method, with the aim of correcting for over and under-estimation of dietary intake. ${ }^{17}$ We were able to look at eight subgroups of fruit and vegetables, although we were not able to look at further subtypes, such as grapes and berries, because median intakes were very low. A limitation of this study was that fruit and vegetable intake was estimated using dietary
assessment questionnaires only at baseline, and their consumption may have changed during follow-up and resulted in exposure misclassification. However, if this is the case, it would have introduced non-differential misclassification, which tends to bias associations towards the null association. Moreover, as with every observational study, we cannot exclude the possibility of residual confounding by other potential risk factors, such as prostate-specific antigen (PSA) testing which was not available in our cohort. A recent study in the UK has found that those who consume a higher amount of fruit are more likely to have a PSA test ${ }^{32}$; since those who have had a PSA test are more likely to be diagnosed with prostate cancer, it is unlikely that adjustment for this variable would have changed the associations found. Finally, the increment unit used in our study was $100 \mathrm{~g} /$ day in order to provide comparability with previous publications including the WCRF meta-analysis; however, it should be acknowledged that although this increment is appropriate for major food groups, it is relatively large for some specific food subgroups.

The findings from this large prospective study in European men suggest that higher total fruit consumption may be associated with a small reduction in prostate cancer risk. Some weak evidence of an inverse association between citrus fruit and overall prostate cancer risk, and a positive association between leafy vegetables with high grade prostate cancer was observed. More data are needed from large observational studies with long-term follow-up, fruit and vegetable subtypes, and prostate cancer risk by grade and stage of the tumor and prostate cancer mortality before conclusions on risk can be drawn.

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