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REGULAR RESEARCH ARTICLE

Antidepressive Effect of Antipsychotics in the Treatment of Schizophrenia: Meta-Regression Analysis of Randomized Placebo-Controlled Trials

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Abstract

Background: Antipsychotics improve the positive symptoms of schizophrenia. However, little is known about the extent of antidepressive effects of antipsychotics and their correlation with effects on other symptom domains in schizophrenia. The aim was to investigate whether antidepressive effects of antipsychotics have a significant correlation with the effects on specific symptom domains of schizophrenia.

Methods: Electronic databases were searched to identify eligible studies that reported antidepressive effects of antipsychotics for the treatment of adult patients with schizophrenia in double-blind, randomized placebo-controlled trials (RCTs). Mean change from baseline in depressive symptoms was meta-analyzed, and the correlation with the effects on other symptom domains was examined through meta-regression analysis.

Results: Thirty-five RCTs (13890 patients) were included in this meta-analysis. Overall, antipsychotics showed greater efficacy than placebo in reducing depressive symptoms, with small to medium effect sizes (standardized mean difference = -0.27, 95% confidence interval -0.32 to -0.22, P < .001). All the antipsychotics, except for chlorpromazine, haloperidol, and ziprasidone, were associated with significantly greater decreases in depressive symptoms compared with placebo (standardized mean difference = -0.19 to -0.40). A higher antidepressive effect was significantly correlated with a higher improvement in Positive and Negative Syndrome Scale/Brief Psychiatric Rating Scale total, positive, and negative, and Positive and Negative Syndrome Scale-general psychopathology symptoms (β =.618, P < .001; β =.476, P < .001; β =.689, P < .001; β =.603, P < .001, respectively). **Conclusions:** Second-generation antipsychotics (except for ziprasidone) were associated with small to medium effects sizes on improvement in depressive symptoms among adult patients with schizophrenia. The antidepressive effect of antipsychotics was significantly correlated with improvement in other symptom domains, with the highest correlation observed for improvement in negative symptoms.

PROSPERO registration number: CRD42019133015

Key Words: Antipsychotics, antidepressants, schizophrenia, psychopathology, meta-regression

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Significance Statement

To our knowledge, this study is the first comprehensive meta-analysis that focused on the antidepressive effects of antipsychotics in adult patients with acute schizophrenia. In the present systematic review and meta-analysis of 35 double-blind, placebo-controlled RCTs with a total of 13890 patients, we found that treatment with second-generation antipsychotics except for ziprasidone was associated with significant improvement of depressive symptoms with small to medium effect size. Furthermore, there was a significant correlation between improvement in depressive symptoms and in psychopathological domain, indicating that some of the reduction in depressive symptoms may be related to the improvement in other symptoms of schizophrenia, in particular negative symptoms.

Despite the extensive use of second-generation antipsychotics in the treatment of schizophrenia, depressive symptoms are still frequently observed and remain a therapeutic challenge in clinical practice. Our findings could renew interest in clinical research on the true antidepressive effect of antipsychotics.

Introduction

A significant proportion of patients with schizophrenia exhibit depressive symptoms during their disease course (Siris, 2000; Siris and Bench, 2003; Conley et al., 2007; Buckley et al., 2009; Majadas et al., 2012). One possible explanation of such depressive symptoms was the early concept of "akinetic depression" that attributed such depression to the extrapyramidal symptoms induced by first-generation antipsychotic (FGA) medications (Van Putten and May, 1978). Another explanation was the notion of "revealed depression" that hypothesized that depression is an inherent part of the illness with the depressive symptoms hidden from clinicians when psychotic symptoms are pronounced, only to be revealed when such psychotic symptoms remit (Hirsch, 1982). A further idea was that of "post-psychotic" depression that hypothesized that reactive depression ensures when individuals with schizophrenia recover and gain insight into their illness and life situation (McGlashan and Carpenter, 1976). Depressive symptoms in schizophrenia have also been thought to be peripheral symptoms caused by comorbid mental disorders (Siris, 2000; Siris and Bench, 2003).

The presence of depressive symptoms in patients with schizophrenia is characterized by an impairment in occupational functioning (Sands and Harrow, 1999), a decreased quality of life (Reine et al., 2003; Sim et al., 2004), and an increased risk of relapse or hospitalization (Mandel et al., 1982). More importantly, depressive symptoms in schizophrenia are closely related to a higher risk of suicide (Bottlender et al., 2000). Thus, depressive symptoms are a major cause of health, social, and financial burden of schizophrenia (Knapp et al., 2004). The frequency of clinically relevant depressive symptoms is greater in the acute phase than in maintenance phase of schizophrenia (Mulholland and Cooper, 2000). Patients with schizophrenia in a stable phase frequently present depressive symptoms despite not being diagnosed with depression (Majadas et al., 2012). Such depressive symptoms in patients with schizophrenia are therefore likely underrecognized and undertreated. Currently available antipsychotics improve the positive symptoms of schizophrenia (Leucht et al., 2017). However, little is known about the extent of antidepressive effect of antipsychotics and about their relationship to the effect on other symptom domains in those with this disorder. Multiple antipsychotics are now approved for use in multiple countries for treatment of treatment-resistant major depressive disorder, and the use of such drugs for treating depression has increased substantially (Gerhard et al., 2014). This further raises the question of how effective these agents are in the treatment of depression in the context of schizophrenia.

We hypothesized that antipsychotics are associated with significant improvement of depressive symptoms in patients with schizophrenia. We aimed to conduct a systematic review and meta-analysis of acute-phase, randomized placebo-controlled trials (RCTs) reporting the efficacy of antipsychotics on depressive symptoms in schizophrenia and to investigate the relationship between antidepressive effect of antipsychotics and effects on the other symptom domains of schizophrenia.

Methods

The meta-analysis followed PRISMA guidelines for reporting meta-analyses of RCTs (Moher et al., 2009). The review protocol was registered at the international prospective register of systematic reviews registration number CRD42019133015.

Study Selection

We selected published and unpublished double-blind RCTs of antipsychotics for adult patients with acute schizophrenia or related disorders; all were placebo controlled. We included all antipsychotics licensed in at least 1 country, except clozapine, as it may be a more efficacious drug (Huhn et al., 2019), and so pooling it with the other compounds would not have been appropriate (only 1 clozapine arm with 16 patients from 1 study [Honigfeld, 1984] had to be excluded on this basis making the impact of this decision negligible). We excluded intramuscular formulations because these are used primarily either for emergency use (short-acting intramuscular drugs) or for relapse prevention (long-acting depot drugs).

Data Sources

We conducted a literature search without language restrictions—using MEDLINE/PubMed, Cochrane library, Scopus, and Embase from database inception (last search: May 18, 2019)—for RCTs of patients with acute schizophrenia. We also searched for unpublished studies, such as conference proceedings and clinical trial registries (http://clinicaltrials.gov/). Search terms included synonyms of (1) schizophrenia AND (2) controlled AND (3) randomized AND (4) clinical trial AND (5) antipsychotic drug. Hand searches of reference lists of relevant publications were also conducted. Because the current investigation was a systematic review and meta-analysis utilizing only secondary sources, informed consent and approval by an ethics board were not required.

Data Extraction

Data were extracted independently by ≥2 reviewers (I.M., T.N., K.H.) experienced in conducting literature searches and data extraction. Disagreements were resolved by consensus. Foreign language papers were translated by bilingual speakers, and data extraction was double checked by 2 investigators (T.N., K.H.) using Google Translate (http://translate.google.com/).

The primary efficacy outcome was set as mean change from baseline to endpoint in depressive symptom scale scores, as represented by the change in the Montgomery Åsberg Depression Rating Scale (MADRS) (Montgomery and Åsberg, 1979), Hamilton Rating Scale for Depression (HAM-D) (Hamilton, 1960), or Calgary Depression Scale for Schizophrenia (CDSS) (Addington et al., 1993). If changes in those scale scores were not available, the change in the score on the Positive and Negative Syndrome Scale (PANSS)-Anxiety/Depression factor (Kay et al., 1987) or on the Brief Psychiatric Rating Scale (BPRS)-Depression cluster (Overall and Gorham, 1962) was used for the calculation of the effect size. Secondary outcomes included symptom severity (positive, negative, general, and total symptoms), Clinical Global Impressions-Severity Illness scale (CGI-S) (Guy, 1976), global functioning, quality of life, all-cause discontinuation at study endpoint, inefficacy-related discontinuation, and intolerabilityrelated discontinuation.

All eligible trials were assessed for methodological quality using the Cochrane Collaboration's tool for assessing risk of bias (Higgins et al., 2011). We extracted data on study design and patient, illness, and treatment components.

Data Synthesis

All data were double-entered into and meta-analyzed with Comprehensive Meta-Analysis Version 3 (BioStat; Englewood, New Jersey) using a random effects model, as heterogeneity among studies was expected (DerSimonian and Laird, 1986). Continuous outcomes were expressed as the standardized mean difference (SMD) using the inverse variance method, each with their 95% confidence intervals (CIs). For simplicity, we adjusted effect sizes so that SMDs < 0 indicated superiority of antipsychotic drug, independent of whether a lower value (e.g., depressive symptom) or higher value (e.g., functioning, quality of life) was a positive outcome. Dichotomous outcomes were expressed as relative risk (RR) each with their 95% CIs. RR values <1 indicated superiority of antipsychotic treatment for negative outcomes such as discontinuation due to any cause. Numberneeded-to-harm (NNH) was calculated when risk differences were significant.

Several articles reported trials with 2 or more experimental groups and only 1 placebo group. To avoid counting the placebo patients twice, we followed the recommendation of the Cochrane Collaboration and divided the placebo group equally into 2 (or more) groups with smaller sample sizes so that the total numbers of participants added up to the original size of the group. We thereby avoided an inflation of sample size, which would lead to an increase of type I errors and thus overoptimistically small SDs. We explored study heterogeneity using the χ^2 test of homogeneity and I² statistics, with P<.05 and I²>50%, respectively, indicating significant heterogeneity. All analyses were 2-tailed with an α of .05. In the primary analyses, antipsychotics and placebo were compared at study end point.

We also conducted subgroup and exploratory maximum likelihood random-effects meta-regression analyses of the primary outcome to identify potential moderators or mediators. Subgroup analyses were based on (1) antipsychotic drug type (FGA vs second-generation antipsychotics [SGA]), (2) neuroscience-based nomenclature (NbN) (Zohar et al., 2015), (3) publication year (published 1999 or before, from 2000 to 2009, 2010 or later), (4) country, (5) study sample size ($n = \le 400$ vs >400), (6) mean age (coded in these ranges: 30-35, 35-40, 40-45, 45-50 years), (7) included patients (schizophrenia, schizophrenia and schizoaffective disorder or schizoaffective disorder), and (8) rating scale (MADRS, HAM-D, or CDSS vs the PANSS anxiety/depression scale or BPRS depression cluster). The NbN categories were defined as follows: M1: receptor antagonists (D2) clopenthixol, fluphenazine, haloperidol, perphenazine, pimozide, pipotiazine, sulpiride, and trifluoperazine; M2: receptor antagonists (D2, 5-HT2) chlorpromazine, iloperidone, loxapine, lurasidone, olanzapine, sertindole, thioridazine, ziprasidone, and zotepine; M3: receptor partial agonists (D2, 5-HT1A) aripiprazole, brexpiprazole, and cariprazine; M4: receptor antagonists (D2, 5-HT2, NE, α 2) as enapine, paliperidone, and risperidone; and M5: receptor antagonist (D2, 5-HT2) and reuptake inhibitor (NET) quetiapine.

The secondary outcome measures were regressed on the primary outcome measures to better understand the clinical impact of the primary outcome measures. Finally, we inspected funnel plots. The regression test by (Egger et al., 1997) and the trim-and-fill method by Duval and Tweedie, (2000) were used to examine the presence of publication bias. In this large exploratory set of analyses, no adjustments were made to the P values for the multiple comparisons; therefore, these P values should be interpreted with caution.

Results

Results of the Search

We identified 9572 records and included 35 RCTs with a total of 13890 participants. The PRISMA flow diagram of the search and screening process is presented in supplementary Figure 1; detailed characteristics of individual included studies are summarized in Table 1.

Description of Included Studies

Studies were published between 1984 and 2019. Twenty-three studies were conducted as multinational clinical trials. Eight studies were conducted in the United States, 2 were in Japan, 1 was in Germany, and 1 was not clear about the study location. Thirty-three studies were sponsored by industry, and sponsors for the remaining 2 studies were not clear or not mentioned. The median number of participants per study was 444 (range: 39–674), and the mean duration was 5.7 weeks (range: 3–7). Thirty-four studies (97.1%) used different versions of DSM and 1 used Research Diagnostic Criteria as the basis for diagnosis. The mean age of participants was 39.2 years (SD=2.9), 65.5% of participants were men, and 46.3% of patients were White.

The number of studies with each individual antipsychotic were lurasidone (7), olanzapine (7), paliperidone (6), quetiapine (5), brexpiprazole (4), haloperidol (4), risperidone (4), asenapine (3), aripiprazole (2), ziprasidone (2), cariprazine (1), and chlorpromazine (1).

Efficacy of Antipsychotics for Depressive Symptoms

Across 46 comparisons based on 35 RCTs, antipsychotics were more efficacious than placebo for depressive symptoms (46

	Design								
	blinding		Patients. No.						
	country		(randomized) treatment						Denression rating
Study	sponsor	Duration	arms (n)	Population	Mean age (y)	% Male	% White	Mean dose (mg/d) or range	scale
Borison, 1996	DB NR	Q	109 QUE (54) PBO (55)	Schizophrenia (DSM-III-R) 18–60	36.5	89.9	60.6	307	BPRS depression cluster score
Cantillon, 2017	Fharma DB International Pharma	7	234 234 ARI 15 (20) [RP5063 15 (58)] [RP5063 30 (59)] [RP5063 50 (58)]	y of age Schizoaffective schizoaffective disorder (DSM-IV- TR) 18–65 y of age	35.9	76.4	5.15	ARI 15 fixed dose	CDSS
Canuso, 2010a	DB International Pharma	Q	PBO (39) 316 PAL Lower-dose (109) PAL Higher-dose (100)	Schizoaffective disorder (DSM-IV) 18–65 y of age	37.2	64.8	46.5	Lower dose 5.7 Higher dose 11.6	HAM-D 21ª
Canuso, 2010b	DB International	9	FBO (107) 311 PAL (216) PPO (95)	Schizoaffective disorder (DSM-IV)	37.6	55.9	51.3	8.6	HAM-D 21ª
Correll, 2015	Friatina DB International Pharma	Q	FBC (32) 636 [BRE 0.25 (90)] BRE 2 (182) BRE 4 (180)	16–05 y of age Schizophrenia (DSM-IV-TR) 18–65 y of age	40.1	63.1	66.5	BRE 0.25, 2, and 4 fixed dose	PANSS- Anxiety/ Depression
Corrigan, 2004	DB International Pharma	Q	467 467 OLA 15 (93) [SON 10 (99)] [SON 60 (91)]	Schizophrenia (DSM-IV) 18–65 y of age	37.6	65.7	31.7	OLA 15 fixed dose	CDSS
Cutler, 2010	DB US Pharma	Q	FBO (87) 565 QUE XR 400 (114) QUE XR 800 (105) QUE XR 800 (113) QUE IR 800 (116) BDO (117)	Schizophrenia (DSM-IV) 18–65 y of age	41.4	71.5	32.5	QUE 400, 600, and 800 fixed dose	PANSS- Anxiety/ Depression
Daniel, 1999	DB International Pharma	Q	202 (117) 302 ZIP 80 (106) ZIP 160 (104) PBO (92)	Schizophrenia or schizoaffective disorder (DSM- III-R) 18 y or older	36.6	71.2	68.2	ZIP 80 and 160 fixed dose	MADRS

Table 1. Characteristics of Included Studies

	Design								
	blinding		Patients. No.						
	country		(randomized)						Danxassion rating
Study	sponsor	Duration	arms (n)	Population	Mean age (y)	% Male	% White	Mean dose (mg/d) or range	scale
Davidson, 2007	DB International Pharma	ω	618 OLA 10 (128) PAL 3 (127) PAL 9 (125) [PAL 15 (115)] PBO (173)	Schizophrenia (DSM-IV) 18 y or older	36.8	89	67	OLA 10 PAL 3, 9, and 15 fixed dose	PANSS- Anxiety/ Depression
Durgam, 2015	DB International Pharma	Q	617 ARI 10 (152) ARI 10 (152) CAR 6 (157) PRO (153)	Schizophrenia (DSM-IV-TR) 18–60 y of age	38.5	63.2	64.0	ARI 10 CAR 3 and 6 fixed dose	PANSS-Anxiety/ Depression
Egan, 2013	DB International Pharma	4	216 216 OLA 15 (47) [MK-8998 (86)] PBO (83)	Schizophrenia (DSM-IV-TR) 18–55 y of age	36.8	58.3	99.5	OLA 15 fixed dose	MADRS
Hirayasu, 2010	DB Japan Pharma	Q	220 (22) 321 OLA 10 (47) PAL 6 (136) PBO (138)	Schizophrenia (DSM-IV) 20 y or older	45.3	50.6	0	OLA 10 PAL 6 fixed dose	PANSS-Anxiety/ Depression
Higuchi, 2019	DB International Pharma	Q	460 LUR 40 (131) LUR 80 (131) RIS 4 (65) PRO (133)	Schizophrenia (DSM-IV) 18–75 y of age	45.6	59.1	0	LUR 40 and 80 RIS 4 fixed dose	PANSS-Anxiety/ Depression
Honigfeld, 1984	DB US NR	4	39 39 [CLO (16)] CP (15) PRO (8)	Schizophrenia (clinical diagnosis)	NR	NR	NR	[CLO 800] CP 1333	BPRS depression cluster score
Ishigooka, 2018	DB Japan Pharma	Q	459 459 BRE 1 (115) BRE 2 (115) BRE 2 (113) PRO (116)	Schizophrenia (DSM-IV-TR) 18–65 y of age	44.3	47.5	0	BRE 1, 2, and 4 fixed dose	PANSS-Anxiety/ Depression
Kahn, 2007	DB International Pharma	٥	588 QUE IR 400 (123) QUE XR 400 (113) QUE XR 600 (113) QUE XR 800 (121) PBO (118)	Schizophrenia (DSM-IV) 18–65 y of age	34.2	60.2	59.2	QUE 400, 600, and 800 fixed dose	PANSS-Anxiety/ Depression

Table 1. Continued

Continued
Table 1.

	Design								
	blinding		Datiants No						
	country		(randomized) treatment						Depression rating
Study	sponsor	Duration	arms (n)	Population	Mean age (y)	% Male	% White	Mean dose (mg/d) or range	scale
Kane, 2010	DB International Pharma	ى	458 ASE 10 (114) ASE 20 (106) HAL 8 (115) PRO (123)	Schizophrenia (DSM-IV-TR) 18 y or older	37 to 40 years across the 4 treatment groups	52% to 68% across the 4 treatment groups	59% to 64% across the 4 treatment groups	ASE 10 and 20 HAL 8 fixed dose	CDSS
Kane, 2015	DB International Pharma	Q	674 BRE 1 (120) BRE 2 (186) BRE4 (184) PBO (184)	Schizophrenia (DSM-IV-TR) 18–65 y of age	38.4	62.8	60.4	BRE 1, 2, and 4 fixed dose	PANSS-Anxiety/ Depression
Keck, 1998	DB US Pharma	4	139 ZIP 40 (44) ZIP 120 (47) PBO (48)	Schizophrenia or schizoaffective disorder (DSM- III-R) 18–64 y of age	39.4	79.1	71.9	ZIP 40 and 120 fixed dose	BPRS depression cluster score ^b
Kinoshita, 2016	DB International Pharma	Q	532 ASE 10 (176) ASE 20 (182) PBO (174)	Schizophrenia (DSM-IV-TR) 20–64 y of age	41.4	48.1	0	ASE 10 and 20 fixed dose	PANSS-Anxiety/ Depression
Klieser and Lehmann, 1989	DB Germany NR	ε	75 HAL 20 (22) [AMIT 150 (20)] [TRA 400 (17)] PBO (16)	Schizophrenia (DSM-III)	42.6°	40.8°	NR	HAL 20 fixed dose	HAM-D
Landbloom, 2017	DB Intemational Pharma	Q	360 [ASE 5 (98)] ASE 10 (113) OLA 15 (46) PBO (103)	Schizophrenia (DSM-IV-TR) 18 y or older	40.6	58.5	71.4	ASE 5 and 10 OLA 15 fixed dose	PANSS-Anxiety/ Depression
Lieberman, 2016	DB US Pharma	4	335 RIS 4 (82) [ITT 007 60 (84)] [ITT 007 120 (84)] PBO (85)	Schizophrenia (DSM-IV-TR- Clinical Trials Version) 18–55 y	40.1	82.6	18.6	RIS 4 fixed dose	CDSS ^d
Loebel, 2013	DB Intemational Pharma	٥	488 LUR 80 (125) LUR 160 (121) QUE 600 (120) PBO (122)	Schreia (DSM-IV-TR) 18–75 y of age	37.2	68.3	56.6	LUR 80 and 160 QUE 600 fixed dose	MADRS

	blinding		Patients, No.						
	country		(randomized) treatment						Depression rating
Study	sponsor	Duration	arms (n)	Population	Mean age (y)	% Male	% White	Mean dose (mg/d) or range	scale
Loebel, 2016	DB International Pharma	۵	412 LUR 80 or 160 (199) [LUR 20 (101)] PBO (112)	Schizophrenia (DSM-IV-TR) 18–75 y of age	40.8	63.7	72.7	LUR 80 or 160 At first 2 wk, ER continued to take LUR 80 mg/d, while ENR were re- randomized in 1:1 ratio either to continue 80 mg/d or receive 160 mg/d for	MADRS
Marder, 2007	DB US Pharma	و	444 OLA 10 (110) PAL 6 (112) PAL 12 (112) PAO 14 10	Schizophrenia (DSM-IV) 18 y or older	41.6	74	43	remaining 4 wk OLA 10 PAL 6 and PAL 12 fixed dose	PANSS- Anxiety/ Depression
Marder, 2016	DB International Pharma	Q	468 468 BRE (151) QUE (154) PRO (163)	Schizophrenia (DSM-IV-TR) 18–65 y of age	40.6	56.9	75.2	BRE 2-4 QUE 400-800	PANSS- Anxiety/ Depression
Meltzer, 2004	DB US Pharma	Ś	$\begin{array}{c} \text{HAL 10 (98)} \\ \text{HAL 10 (98)} \\ \text{[5-HT]}_{2a/2c} \text{-Ant.} \\ \text{[74]} \\ \text{[NK_3-Ant. (70)]} \\ \text{[NK_3-Ant. (70)]} \\ \text{[CB_1-Ant. (72)]} \\ \text{[CB_1-Ant. (69)]} \\ \text{[NTS_1-Ant. (69)]} \end{array}$	Schizophrenia or schizoaffective disorder (DSM-IV) 18–64 y of age	36.6	73.8	48.9	HAL 10 fixed dose	CDSS
Meltzer, 2011	DB Intemational Pharma	Q	478 478 LUR 40 (120) LUR 120 (119) OLA 15 (123) DED (116)	Schizophrenia (DSM-IV-TR) 18–75 y of age	37.7	78.0	35.7	LUR 40 and 120 OLA 15 fixed dose	MADRS
Nakamura, 2009	DB US Pharma	9	180 180 LUR 80 (90) PRO (90)	Schizophrenia (DSM-IV) 18–64 y of age	40.8	76.7	33.9	LUR 80 fixed dose	MADRS
Nasrallah, 2013	DB International Pharma	Q	500 500 LUR 40 (125) LUR 80 (123) LUR 120 (124) PBO (128)	Schizophrenia (DSM-IV) 18–75 y of age	38 8	69.5	49.5	LUR 40, 80, and 120 fixed dose	MADRS

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Table 1. Continued

Design

Continued	
Table 1.	

	Design								
	blinding		Patients. No.						
	country		(randomized) treatment						Denression rating
Study	sponsor	Duration	arms (n)	Population	Mean age (y)	% Male	% White	Mean dose (mg/d) or range	scale
Potkin, 2015	DB US Pharma	۵	356 LUR 40 (69) LUR 20 (71) [LUR 20 (71)] HAL 10 (73) PBO (72)	Schizophrenia (DSM-IV) 18–64 y of age	41.2	74.2	4 <u>9</u> .9	LUR 20, 40, and 80 HAL 10 fixed dose	MADRS
	1								
NCT01175135	DB International Pharma	4	259 RIS 6 (37) [PF-02545920 10 (74)] [PF-02545920 30 (74)] PBO (74)	Schizophrenia (DSM-IV) 18–65 y of age	41.9	69.4	NR	RIS 6 fixed dose	PANSS-Anxiety/ Depression
ncluded studies	= 35. n = 13 890.		~		Mean age: 39.2 +	2.9 V			
Country: interna (studies=2, n=	tional (studies = . 780), Germany (:	23, n=10 387) studies=1, n=	, US (studies =8, n = 25: =75), NR (studies = 1, n.	39), Japan =109).	% Male: 65.5 ± 10				
Sponsorship: ph	armaceuticals (s	tudies=33, n=	=13 776), NR (studies=	: 2, n=114).	% White: 46.3 ± 2	5.2			
study duration: .	mean=5.7 ± 0.9 1	weeks (range	3-7)		The number of p: Antipsychotics: a cariprazine (stu (studies=7), ols (studies=4), zir	atients per study ripiprazole (stud idies = 1), chlorpi inzapine (studie rasidone (studie	: median = 444 (rar ies= 2), asenapine omazine (study= : s= 7), paliperidone s= 2)	nge 39–674) (studies=3), brexpiprazole (studi 1), haloperidol (studies=4), lurasi (studies=6), quetiapine (studies	es=4), done =5), risperidone
Abbreviations: AMI	T, amitriptyline; AR	'I, aripiprazole; ⁄	ASE, asenapine; BPRS, brie	f psychiatry rating scale; B	RE, brexpiprazole; CA	R, cariprazine; CDS!	5, Calgary Depression	Scale for Schizophrenia; CLO, clozapir	ie; CP, chlorpromazine;

DB, double blind; DSM-IV-TR, Diagnostic and Statistical Manual of Mental Disorders, Fourth Edition, Taxt Revision; HAL, haloperidol; HAM-D, Hamilton Rating Scale for Depression; LUR, lurasidone; MADRS, Montgomery Åsberg Depression Rating Scale; NR, not reported; OLA, olanzapine; PANSS; Positive and Negative Syndrome Scale; PBO, placebo; QUE, quetiapine; RIS, risperidone; SON, sonepiprazole; TRA, trazodone; ZIP, ziprasidone. PAL, ER, DSM-III-R ? aAnti-depressive effect was assessed in patients with baseline HAM-D21 ≥ 16. bAnti-depressive effect was assessed in patients with baseline BPRS depression cluster score ≥18.

cIncluding 45 patients with major depressive disorder.

dAnti-depressive effect was assessed in patients with baseline CDSS > 6. Drug groups or numbers in squared brackets were not used in any analysis, neither the primary one nor in a sensitivity analysis.

comparisons, 11197 participants, SMD: -0.27, 95% CI: -0.32 to -0.22, P<.001; heterogeneity, τ^2 =.01, I^2 =31.3%, Q=65.5, df=45, P=.025) (Figure 1). All the antipsychotics, except for chlorpromazine, haloperidol, and ziprasidone, were associated with significantly greater decreases in depressive symptoms compared with placebo (SMD=-0.19 to -0.40) (Figure 1; supplementary Table 1).

Meta-Regression Analyses

Meta regression analyses showed that the improvement of PANSS/BPRS total score, positive symptom, negative symptom, PANSS-general psychopathology scale score, and CGI-S score significantly predicted the improvement of depressive symptoms with slightly higher correlation with negative symptoms (β =0.618, 95% CI = 0.433 to 0.802, P<.001; β =0.476, 95% CI = 0.294 to 0.657, P<.001; β =0.689, 95% CI = 0.475 to 0.903, P<.001; β =0.603, 95% CI = 0.302 to 0.904, P<.001; and β =0.585, 95% CI = 0.343 to 0.826, P<.001; respectively) (Figure 2a-c; supplemental Figure 2a-b).

The improvement of depressive symptoms measured with PANSS/BPRS-depression significantly correlated with the improvement of PANSS/BPRS total score, positive, and negative symptoms, and PANSS-general psychopathology scale score with similar coefficients as found with the primary outcome (β =0.610, 95% CI = 0.349 to 0.870, P<.001; β =0.441, 95% CI = 0.219 to 0.663, P<.001; β =0.688, 95% CI = 0.388 to 0.989, P<.001; and β =0.667, 95% CI = 0.350 to 0.983, P<.001, respectively) (supplementary Figure 2c–f).

The improvement of depressive symptoms measured with MADRS, HAM-D, or CDSS significantly correlated with the improvement of PANSS/BPRS total score, positive, and negative symptoms (β = 0.622, 95% CI = 0.360 to 0.884, P < .001; β = 0.548, 95% CI = 0.261 to 0.834, P<.001; β = 0.689, 95% CI = 0.380 to 0.998, P<.001, respectively) (supplementary Figure S3a-c); however, it was not significantly correlated with PANSS-general psychopathology scale score (supplementary Figure 3d). The effect size for the improvement of depressive symptoms also had a significant correlation with the improvement of global functioning $(\beta = 0.650, 95\% \text{ CI} = 0.020 \text{ to } 1.281, P = .043)$ (supplementary Figure 4). This effect was independent from mean age of participants, percentage of male or white, and publication year (supplementary Figures 5-8). The difference between antipsychotics and placebo decreased with higher baseline PANSS depression score, but baseline MADRS total score was not a statistically significant moderator (supplementary Figures 9 and 10).

Subgroup Analyses

The summary of subgroup analyses for depressive symptoms is presented in Figure 3. Significant superiority of antipsychotics over the placebo remained in SGA (SMD=-0.28, 95% CI = -0.33 to -0.23, P<.001), but not in FGA (SMD=-0.08, 95% CI = -0.24 to 0.07, P=.294).

Superiority of antipsychotics over the placebo remained in M2-M5 subgroups of NbN (M2, 17 comparisons, 4134 participants, SMD = -0.23, 95% CI = -0.32 to -0.13, P<.001; M3, 7 comparisons, 2752 participants, SMD = -0.25, 95% CI = -0.34 to -0.15, P<.001; M4, 12 comparisons, 3245 participants, SMD = -0.37, 95% CI = -0.44 to -0.29, P<.001, M5, 5 comparisons, 1768 participants, SMD = -0.32, 95% CI = -0.48 to -0.16, P<.001), but not in M1 subgroup of NbN (5 comparisons, 624 participants, SMD = 0.08, 95% CI = -0.24 to 0.07, P = .294). Results for each individual antipsychotic drug separately are appended online (see supplementary Table 1).

Superiority of antipsychotics was moderated by publication year. In RCTs published until 1999 (5 comparisons, 538 participants, SMD = -0.08, 95% CI = -0.27 to 0.12, P = .443) consisting of chlorpromazine (1 RCT), haloperidol (1 RCT), quetiapine (1 RCT), and ziprasidone (2 RCTs), antipsychotics were not significantly superior to the placebo. However, this was not the case in the newer RCTs published from 2000 to 2009 (9 comparisons, 2760 participants, SMD=-0.34, 95% CI = -0.43 to -0.26, P<.001) or after 2010 (32 comparisons, 9225 participants, SMD=-0.26, 95% CI = -0.32 to -0.20, P<.001), which included 3/5 M1 antipsychotics studies (Figure 3). Significant superiority of antipsychotics was also demonstrated in RCTs conducted in mixed countries, the United States, and Japan (31 comparisons, 9557 participants, SMD=-0.29, 95% CI = -0.35 to -0.22, P<.001; 10 comparisons, 1921 participants, SMD=-0.20, 95% CI = -0.30 to -0.10, P<.001; 3 comparisons, 903 participants, SMD = -0.32, 95% CI = -0.56 to -0.08, P = .010, respectively).

Regarding sample size, significant superiority of antipsychotics over the placebo remained in both larger (>400 participants) and smaller RCTs (<400 participants). Though not statistically significant, the effect sizes tended to be higher in younger patients compared with older patients with the target symptoms for which antipsychotic effects might be most expected (SMD = -0.37 [mean age 30–35 years] and -0.30 [mean age 35–40 years] compared with -0.22 [mean age 40–45 years] and -0.24 [mean age 45–50 years]). Significant superiority of antipsychotics was also confirmed in other examined subpopulations, with similar effect sizes (Figure 3). When analyses were restricted to standard depression scales (HAMD, MADRS, CDSS), effects were similar to that obtained with PANSS and BPRS depression scales (Figure 3, bottom row).

Secondary Outcomes

A summary of the pooled results is presented in supplementary Tables 1b-h and 2. Compared with placebo, antipsychotics were associated with significant improvement of overall symptoms (43 comparisons, 12048 participants, SMD = -0.41, 95% CI = -0.48to -0.35, P < .001), positive symptoms (40 comparisons, 12106 participants, SMD = -0.42, 95% CI = -0.49 to -0.35, P < .001), negative symptoms (41 comparisons, 12017 participants, SMD = -0.32, 95% CI = -0.37 to -0.26, P < .001), PANSS general psychopathology score (17 comparisons, 5160 participants, SMD = -0.39, 95% CI = -0.49 to -0.30, P < .001), CGI-S (37 comparisons, 10 482 participants, SMD = -0.33, 95% CI = -0.39 to -0.28, P < .001), global functioning (9 comparisons, 2725 participants, SMD = -0.39, 95% CI = -0.48 to -0.31, P < .001), and quality of life (3 comparisons, 1029 participants, SMD = -0.47, 95% CI = -0.60 to -0.34, P < .001) (supplementary Table 1b-h).

There were significant differences between antipsychotics and the placebo in all-cause discontinuation (45 comparisons, 12882 participants, RR=0.78, 95% CI = 0.73 to 0.84, P<.001, NNH=-11) and inefficacy-related discontinuation (43 comparisons, 12931 participants, RR=0.52, 95% CI = 0.47 to 0.57, P<.001, NNH=-10). However, no significant difference was observed between antipsychotics and the placebo in intolerability related discontinuation (supplementary Table 2).

Publication Bias and Risk of Bias Assessment

No publication bias was detected: the funnel plots were symmetrical (Egger's regression test for depressive symptoms: intercept = 0.85, 95% CI = -0.52 to 2.22, P = .218, and the trim and fill method yielded almost identical effect size (supplementary Figure 11).

		Total					95%	6 CI	- P.
Antipsychotics	Study	(n)				SMD	Lower limit	Upper limit	value
	Durgam 2015	299	_ - -			-0.41	-0.64	-0.18	<0.001
ARI	Cantillon 2017	50		_		-0.32	-0.94	0.30	0.317
	ARI total	349				-0.40	-0.61	-0.18	<0.001
	Kinoshita 2016	525	_ —			-0.48	-0.66	-0.30	<0.001
	Kane 2010	336				-0.25	-0.48	-0.03	0.026
ASE	Landbloom 2017	210				-0.11	-0.38	0.16	0.439
	ASE total	1,071				-0.30	-0.51	-0.08	0.007
	Kane 2015	657	_ _			-0.33	-0.50	-0.15	<0.001
	Correll 2015	536				-0.15	-0.33	0.03	0.108
BRE	Marder 2016	309				-0.12	-0.34	0.10	0.290
	Ishigooka 2018	447				-0.11	-0.32	0.11	0.320
	BRE total	1,949	-			-0.19	-0.29	-0.08	<0.001
64.0	Durgam 2015	454				-0.36	-0.56	-0.17	<0.001
CAR	CAR total	454				-0.36	-0.56	-0.17	<0.001
CP.	Honigfeld 1984	20			→ →	0.76	-0.16	1.69	0.105
CP	CP total	20				0.76	-0.16	1.69	0.105
	Meltzer 2004	191				-0.16	-0.45	0.12	0.262
	Potkin 2015	143		-		-0.10	-0.43	0.23	0.557
HAL	Kane 2010	234				-0.09	-0.35	0.17	0.488
	Klieser and Lehmann 1989	36			_	0.01	-0.65	0.66	0.986
	HAL total	604	-			-0.11	-0.27	0.05	0.179
	Loebel 2013	366				-0.58	-0.80	-0.36	<0.001
	Loebel 2016	286				-0.36	-0.61	-0.12	0.003
	Nakamura 2009	169				-0.36	-0.66	-0.05	0.021
	Meltzer 2011	350				-0.09	-0.31	0.13	0.432
LUR	Higuchi 2019b	439				-0.09	-0.29	0.11	0.397
	Higuchi 2019a	383				-0.07	-0.29	0.14	0.499
	Nasrallah 2013	486		-		-0.02	-0.22	0.19	0.876
	Potkin 2015	205		_		0.01	-0.28	0.30	0.954
	LUR total	2,684	-			-0.19	-0.34	-0.04	0.011
	Hirayasu 2010	184				-0.52	-0.86	-0.18	0.002
	Davidson 2007	246				-0.43	-0.69	-0.18	0.001
	Meltzer 2011	235				-0.37	-0.63	-0.11	0.005
014	Marder 2007	210				-0.33	-0.61	-0.06	0.017
054	Corrigan 2004	178				-0.18	-0.48	0.11	0.229
	Egan 2013	123		-		-0.14	-0.51	0.23	0.460
	Landbloom 2017	144		—		-0.03	-0.38	0.33	0.879
	OLA total	1,320	-			-0.31	-0.43	-0.19	<0.001
	Kane 2007	500				-0.55	-0.76	-0.35	<0.001
	Canuso 2010b	210	———			-0.47	-0.76	-0.17	0.002
	Hirayasu 2010	272				-0.39	-0.63	-0.15	0.001
PAL	Canuso 2010a	201				-0.35	-0.65	-0.05	0.022
	Davidson 2007	366	— .			-0.34	-0.56	-0.12	0.003
	Marder 2007	327				-0.21	-0.45	0.02	0.071
	PAL total	1,876	-			-0.39	-0.49	-0.29	<0.001
	Loebel 2013	236				-0.61	-0.87	-0.35	<0.001
	Kahn 2007	573	— •—			-0.37	-0.57	-0.16	<0.001
QUE	Marder 2016	309				-0.28	-0.51	-0.06	0.014
	Cutler 2010	544				-0.22	-0.43	-0.02	0.035
	Borison 1996	106		—		0.00	-0.38	0.38	1.000
	QUE total	1,768	-			-0.32	-0.48	-0.16	<0.001
	Lieberman 2016	25	• · · · · · · · · · · · · · · · · · · ·	-		-0.45	-1.25	0.35	0.271
	Geffen 2012	184				-0.39	-0.68	-0.10	0.009
RIS	NCT01175135	89				-0.28	-0.74	0.18	0.236
	Higuchi 2019a	193		_		-0.03	-0.33	0.27	0.846
	RIS total	491	-			-0.24	-0.43	-0.05	0.016
	Keck 1998	87				-0.28	-0.72	0.15	0.206
ZIP	Daniel 1999	289	-+-			-0.13	-0.38	0.12	0.325
	ZIP total	376				-0.16	-0.38	0.05	0.138
		44.407				-0.27	-0.32	-0.22	<0.001

Figure 1. Effect sizes of all studies included in the analysis, forest plot. AP, antipsychotic drug; ARI, aripiprazole; ASE, asenapine; BRE, brexpiprazole; CP, chlorpromazine; HAL, haloperidol; CI, confidence interval; LUR, lurasidone; OLA, olanzapine; PAL, paliperidone; PBO, placebo; QUE, quetiapine; RIS, risperidone; SMD, standardized mean difference; ZIP, ziprasidone. SMDs < 0 favor the antipsychotic treatment.



Figure 2. Meta-regression analysis of improvement in depressive symptoms on improvement in clinical symptoms. BPRS, Brief Psychiatric Rating Scale; CGI-S, Clinical Global Impressions–Severity illness scale; PANSS, Positive and Negative Syndrome Scale; SMD, standardized mean difference. (a) Correlation between improvement in depressive symptoms and improvement in PANSS/BPRS total score. (b) Correlation between improvement in depressive symptoms and improvement in positive symptoms. (c) Correlation between improvement in negative symptoms.

									95%	6 CI	в	Hetero	geneity
Variables	Subgroup	N	n					SMD	Lower limit	Upper limit	value	P- value	 ²
	FGA	5	624		-•			-0.08	-0.24	0.07	0.294	0.462	0.0
AP type	SGA	41	11,899		•			-0.28	-0.33	-0.23	<0.001	0.002	43.1
	M1	5	624		-0	_		-0.08	-0.24	0.07	0.294	0.462	0.0
	M2	17	4,134		\bigcirc			-0.23	-0.32	-0.13	<0.001	0.008	51.4
NbN	M3	7	2,752		\bigcirc			-0.25	-0.34	-0.15	<0.001	0.227	26.4
	M4	12	3,245		\bigcirc			-0.37	-0.44	-0.29	<0.001	0.392	5.4
	M5	5	1,768		-0-			-0.32	-0.48	-0.16	<0.001	0.075	52.9
	1999 or before	5	538		-•	_		-0.08	-0.27	0.12	0.443	0.357	8.7
Publication	2000-2009	9	2,760					-0.34	-0.43	-0.26	<0.001	0.380	6.6
year	2010 or after	32	9,225					-0.26	-0.32	-0.20	<0.001	0.002	47.4
	Mixed countries	31	9,557		\bigcirc			-0.29	-0.35	-0.22	<0.001	0.001	49.8
	US	10	1,921		-0-			-0.20	-0.30	-0.10	<0.001	0.423	1.7
Country	Japan	3	903	-				-0.32	-0.56	-0.08	0.010	0.070	62.4
	Germany	1	36	-			_	0.01	-0.65	0.66	0.986	-	-
	NR	1	106					0.00	-0.38	0.38	1.000	-	-
Sample	>400	27	9,776					-0.28	-0.34	-0.22	<0.001	0.001	53.6
size	≤400	19	2,747		•			-0.23	-0.32	-0.14	<0.001	0.208	20.2
	30-35	2	757		-0-			-0.37	-0.54	-0.21	<0.001	0.915	0.0
	35-40	22	6,200		\bigcirc			-0.30	-0.38	-0.23	<0.001	0.005	49.4
Age	40-45	17	4,514		\bigcirc			-0.22	-0.29	-0.15	<0.001	0.276	15.1
	45-50	4	1,032		-0			-0.24	-0.47	-0.02	0.034	0.037	64.6
	NR	1	20		-		\longrightarrow	0.76	-0.16	1.69	0.105	-	-
Included	SCZ only	38	11,439					-0.27	-0.33	-0.21	<0.001	0.001	47.4
patient	SCZ+SzAD or SzAD	8	1,084					-0.23	-0.38	-0.08	0.003	0.252	22.3
Rating	MADRS, HAM-D, CDSS	21	4,106		0			-0.26	-0.34	-0.17	<0.001	0.014	44.8
scale	PANSS or BPRS depression	25	7,990		\bigcirc			-0.27	-0.34	-0.21	<0.001	0.010	44.2
			-1	l.00 -0	.50 0.	00 0.	50 1.0	0					

Favors AP < Favors PBO

Figure 3. Summary of subgroup analyses for depressive symptoms. AP, antipsychotic drug; BPRS, Brief Psychiatric Rating Scale; CDSS, Calgary Depression Scale for Schizophrenia; FGA, first-generation antipsychotic drug; HAM-D, Hamilton Rating Scale for Depression; MADRS, Montgomery Åsberg Depression Rating Scale; NbN, neuroscience-based nomenclature; PANSS, Positive and Negative Syndrome Scale; PBO, placebo; SCZ, schizophrenia; SGA, second-generation antipsychotic drug; SMD, standardized mean difference; SzAD, schizoaffective disorder. SMDs lower than 0 favor the antipsychotic treatment. Bold case indicates that results of the subgroup were statistically significant. P values <.05 indicate statistical significance. (a) M1–M5 are drug mechanisms of action according to the neuroscience-based nomenclature; (NbN). M1: receptor antagonists (D₂) clopenthixol, fluphenazine, haloperidol, perphenazine, pimozide, pipotiazine, sulpiride, and trifluoperazine. M2: receptor antagonists (D₂, 5-HT₂) chlorpromazine, iloperidone, loxapine, lurasidone, olanzapine, sertindole, thioridazine, ziprasidone, and zotepine. M5: receptor antagonists (D₂, 5-HT₂) and reuptake inhibitor (NET) quetiapine.

Detailed results of Cochrane risk of bias assessment are presented in supplementary Table 3. Although we included only double-blind RCTs, the reports often did not provide enough details on the sequence generation or allocation concealment. The explanation for the proper blinding methods and the functioning of the blinding were also insufficiently reported.

Discussion

This meta-analysis included 35 double-blind, placebocontrolled RCTs involving 13890 patients. To our knowledge, this is the first study that focused on the antidepressive effects of antipsychotics and investigated their various aspects in patients with acute schizophrenia.

Our findings suggest that treatment with antipsychotics is associated with significant improvement of depressive symptoms in patients with schizophrenia, but the effect size was small to medium. These results are consistent with a previous meta-analysis (Leucht et al., 2017), though the results are not applicable to all individual antipsychotic drugs. In the current meta-analysis, ziprasidone, haloperidol, and chlorpromazine were not significantly separate from placebo regarding efficacy on depressive symptoms. There was no significant difference between antipsychotics of the M1 group of NbN (i.e., FGA) and placebo in improvement of depressive symptoms. It should be noted, however, that the relatively smaller number of studies (n=5) of FGAs limits statistical power for detecting effects within that subgroup. Two RCTs of ziprasidone (Keck et al., 1998; Daniel et al., 1999) incorporated into the analysis showed significant improvement of psychiatric symptoms but not antidepressive effects. These results suggest that the mechanism of action on receptors other than dopamine D_o receptors, including serotonin receptors, is related to the antidepressive effects of certain antipsychotic drugs. In other subgroup analysis, there was no statistically significant antidepressive effect by antipsychotics in 5 studies published before 1999, in which ziprasidone, chlorpromazine, and haloperidol are included. Among the antipsychotics found here to show significantly greater improvement in depressive symptoms compared with placebo among patients with schizophrenia, the range in effect sizes across agents was similar to that found in a meta-analysis of the antidepressant effects of SGAs as augmentation agents for use with treatment-resistant depression (SMD range=-0.19 to -0.40 found here, range = -0.27 to -0.43 reported in Zhou et al., 2015).

Except for the difference related to individual antipsychotic drug, the overall favorable efficacy of antipsychotics on depressive symptoms was confirmed by additional sensitivity analyses indicating that effects were independent of influence from a number of moderators and potential confounders. Effects were similar when standard depression scales were separately examined. In the subgroup analysis based on the mean age of participants, the efficacy of antipsychotics was established in all age ranges with numerically higher antidepressive effect in studies of younger participants. Similarly, in the meta-regression analysis of antidepressive effect of antipsychotics and the mean age of participants, there was a tendency for antipsychotics to have greater efficacy in younger participants, although this correlation was not significant. These results may suggest the importance of pharmacological treatment for depressive symptoms beginning at an earlier stage of schizophrenia, though more research is needed to investigate age effects.

In the meta-regression analysis using PANSS/BPRS total, subscale scores, and CGI-S, statistically significant positive correlations with improvement of depressive symptoms were observed with slightly higher correlation for negative symptoms. These results suggest that some of the reduction in depressive symptoms by antipsychotics may be related to the improvement in other symptoms of schizophrenia, in particular negative symptoms. Thus, it is possible that the apparent antidepressive effects of antipsychotics on depression could be in fact be in part an effect on negative symptoms that is being picked up in depression scores. Conversely, it is also possible that negative symptoms may be a manifestation of depression. The overlap between some negative symptoms and depressive symptoms would be consistent with this explanation (Krynicki et al., 2018). To this point, a number of studies have found that negative symptoms cluster into 2 distinct domains: avolition-apathy and expressive deficit (Galderisi et al., 2018). It is possible that the correlation between change in depressive symptoms and change in negative symptoms is a function of 1 of these 2 domains rather than both. One study found that cognitive therapy, a treatment known to impact depressive symptoms, improved avolition–apathy but had no effect on measures of emotional expression in patients with schizophrenia (Grant et al., 2012). Future research is needed to sort out the effects of antipsychotics on the relation between change in depressive symptoms and different dimensions of negative symptoms.

Several other factors may also be responsible for the significant correlations among positive symptoms, negative symptoms, and depressive symptoms. Patients with high levels of treatment adherence may have substantially greater changes on all of these symptom measures compared with patients with relatively lower treatment adherence. Similarly, differences among individuals in the metabolism of the drugs may produce differences in clinical improvement on all measures. In addition, it should be noted that rating scales can be prone to halo effects or correlational error whereby raters tend to let their general impression of a patient's severity influence their ratings on all items (i.e., a more severely psychotic patient is also more likely to be rated high on items unrelated to psychosis including depression). Such halo effects can increase the correlations among the scales.

Although our results suggest that lower baseline PANSSdepression scale predicted greater difference between antipsychotics and placebo in decrease of depressive symptom scale score, the coefficient was small and such a significant correlation was not observed between baseline MADRS total score and depressive symptom improvement. While previous studies (Conley et al., 2007; Perkins et al., 2008) have identified comorbid depression as a predictor of poor antipsychotic treatment outcome, a participant-level metaanalysis (Furukawa et al., 2015) reported that the difference in symptom reduction between antipsychotics and placebo increased as the baseline severity increased. The association between baseline severity and symptom improvement by antipsychotic treatment may be different between depression and other symptom domains, although the reasons for these controversial results remain unknown. Further studies are needed to investigate the association between baseline severity and depressive symptom improvement by antipsychotics in schizophrenia.

The relation of extrapyramidal symptoms and depression in schizophrenia also needs to be considered. Depression measures cannot always discriminate between depressive symptoms and extrapyramidal (or negative) symptoms. However, a systematic review of depression measures (Lako et al., 2012) found that the CDSS was the most accurate scale for differentiating depression from extrapyramidal and negative symptoms in patients with schizophrenia. Nonetheless, a study in Chinese patients with schizophrenia found that severe levels of akathisia and dyskinesia were significantly associated with greater severity of depressive symptoms (Li et al., 2018). It is possible that severe extrapyramidal symptoms may affect depression symptom severity.

Our results need to be discussed with reference to previous meta-analyses for treatments of depression and negative symptoms in schizophrenia. In 1 meta-analysis (Fusar-Poli et al., 2015), it was reported that treatment with antidepressants was associated with significant improvement of negative symptoms. In another meta-analysis (Helfer et al., 2016), a significant effect of antidepressant augmentation for both depressive and negative symptoms was found. Our results are in line with these previous meta-analyses in the relation of improvement in depressive symptoms and in negative symptoms. In the more recent meta-analysis (Galling et al., 2018), antidepressant augmentation was not superior to placebo for depressive symptom reduction although significantly superior to placebo for negative symptom reduction.

Considering that Galling et al. (Galling et al., 2018) focused on the effects of antidepressants, depressive symptoms in schizophrenia may be qualitatively different from major depressive disorder. As mentioned, there are some important commonalities between depression and negative symptoms in schizophrenia (Andreasen and Olsen, 1982; Carpenter et al., 1985; Siris et al., 1988). Decline in motivation, concentration difficulty, social withdrawal, and anhedonia are common features to both depression and negative symptoms. Other features differentiate these 2 domains of symptoms: feelings of guilt, suicidal ideation, and hopelessness are specifically recognized as depressive symptoms (Lindenmayer et al., 1991; Bermanzohn and Siris, 1992).

Although almost all the SGAs showed significant antidepressive effect in our meta-analysis and the improvement of depressive symptom significantly correlated with the improvement of clinicians' subjective perception of clinical improvement, as measured with the CGI-S, with similar coefficient as found with the PANSS scores, care should be taken as to whether these effects are clinically meaningful. Given that it is difficult to clinically distinguish depression and negative symptoms in schizophrenia, some of the observed antidepressive effects of SGAs may include the improvement of negative symptoms at least in part. On the other hand, a previous study suggested that psychosocial factors such as loss and social isolation, rather than biological symptoms, were core features for post-psychotic depression (Sandhu et al., 2013). To address these psychosocial factors, nonpharmacological intervention such as psychoeducation and social support may also be needed to treat depression in schizophrenia.

Several limitations should be noted. First, included studies are not primarily aimed to improve depressive symptoms of schizophrenia patients, and this might contribute to the heterogeneity of results. Second, the type of pretreatment and pretrial wash out of the medication might have an influence on the mood of the patient, but no studies reported information on the type and dose of pretrial medications. For this reason, we could not analyze pretreatment drug information as a covariate. Third, there are no clinical trials investigating antidepressive effect of antipsychotics in drug-naïve or in the first-episode schizophrenia patients. This was disappointing considering that depression is prevalent in first-episode schizophrenia (Bustamante et al., 1994; Addington et al., 1998; Jäger et al., 2007), and the effect of antipsychotics on first episode patients is greater than that on multiple episode patients (Jäger et al., 2007). Fourth, no studies reported whether psychotherapy or psychological treatment was applied. If these nonpharmacological treatments were present yet not balanced across treatment groups, these would be confounding factors. Fifth, the causality between improvement of depressive symptoms and that of other psychopathological domains is unknown. Improvement of depressive symptoms may be based on improvement of other psychopathological domains, especially negative symptoms. Alternatively, it may be that improvement of other psychopathological domains was observed due to improvement of depressive symptoms, or reciprocal causation may be occurring between these domains of symptoms. Finally, rating scales for depressive symptoms used in included trials were disparate. The most commonly used rating scale for depressive symptoms was the PANSS-anxiety/ depression, which is not designed to assess the depressive symptoms of schizophrenia patients and is not able to distinguish depressive from anxiety and negative symptoms. Though subgroup analyses constrained to studies using standard depression scales yielded similar results as found with the PANSSanxiety/depression scale, as mentioned, such rating scales can be prone to halo effects. Caution is needed to interpret the results of this meta-analysis because of these halo effects or correlational error.

In the present systematic review and meta-analysis, SGA therapy, except for ziprasidone, demonstrates small to medium treatment effects sizes on depressive symptoms in adult patients with schizophrenia. There was a significant correlation between improvement in depressive symptoms and in psychopathological domain, indicating that some of the reduction in depressive symptoms may be related to the improvement in other symptoms of schizophrenia, in particular negative symptoms. Depressive symptoms are still frequently observed in clinical practice despite the fact that the SGAs have been widely used in the treatment of schizophrenia many years. Taking this into consideration, the results may be overestimating the true antidepressive effect of antipsychotics. While current findings are encouraging in that some antipsychotics are effective for the treatment of depressive symptoms in schizophrenia, further sufficiently large, parallel-group RCTs with the treatment of depressive symptoms as the primary focus would be required to confirm the results of our meta-analysis.

Supplementary Materials

Supplementary materials accompanying this paper are available at International Journal of Neuropsychopharmacology online.

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Statement of Interest

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