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Etiologic effects and optimal intakes of foods and nutrients for risk of cardiovascular diseases and diabetes: Systematic reviews and meta-analyses from the Nutrition and Chronic Diseases Expert Group (NutriCoDE)

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Abstract

Background

Dietary habits are major contributors to coronary heart disease, stroke, and diabetes. However, comprehensive evaluation of etiologic effects of dietary factors on cardiometabolic outcomes, their quantitative effects, and corresponding optimal intakes are not wellestablished.

Objective

To systematically review the evidence for effects of dietary factors on cardiometabolic diseases, including comprehensively assess evidence for causality; estimate magnitudes of etiologic effects; evaluate heterogeneity and potential for bias in these etiologic effects; and determine optimal population intake levels.

Methods

We utilized Bradford-Hill criteria to assess probable or convincing evidence for causal effects of multiple diet-cardiometabolic disease relationships. Etiologic effects were quantified from published or *de novo* meta-analyses of prospective studies or randomized clinical trials, incorporating standardized units, dose-response estimates, and heterogeneity by age and other characteristics. Potential for bias was assessed in validity analyses. Optimal intakes were determined by levels associated with lowest disease risk.



all data in the study and had final responsibility for the decision to submit for publication.

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Results

We identified 10 foods and 7 nutrients with evidence for causal cardiometabolic effects, including protective effects of fruits, vegetables, beans/legumes, nuts/seeds, whole grains, fish, yogurt, fiber, seafood omega-3s, polyunsaturated fats, and potassium; and harms of unprocessed red meats, processed meats, sugar-sweetened beverages, glycemic load, *trans*-fats, and sodium. Proportional etiologic effects declined with age, but did not generally vary by sex. Established optimal population intakes were generally consistent with observed national intakes and major dietary guidelines. In validity analyses, the identified effects of individual dietary components were similar to quantified effects of dietary patterns on cardiovascular risk factors and hard endpoints.

Conclusions

These novel findings provide a comprehensive summary of causal evidence, quantitative etiologic effects, heterogeneity, and optimal intakes of major dietary factors for cardiometabolic diseases, informing disease impact estimation and policy planning and priorities.

Introduction

Cardiometabolic diseases including coronary heart disease (CHD), stroke, and type 2 diabetes are leading causes of morbidity and mortality globally [1]. In 2011, the United Nations highlighted suboptimal diet as one of the principal drivers of these diseases [2]. Our collaborative work in the Global Burden of Diseases (GBD) Study demonstrated that 8 of the top 20 risk factors for lost disability-adjusted life-years globally were dietary factors; and that several more of the top 20 were strongly diet-related, including high blood pressure, body mass index (BMI), fasting plasma glucose, and total cholesterol [3]. In sum, suboptimal diet is one of the leading preventable causes of death and disability in the US and globally [3–6].

To determine the impact of specific dietary factors on cardiometabolic diseases and inform priorities for intervention and prevention, it is crucial to understand the strength of evidence on causality, the magnitudes of disease-specific etiologic effects (e.g., relative risks [RRs]), the heterogeneity in these effects by underlying individual characteristics such as age or sex, and the optimal levels of consumption for reducing risk. Yet, these key questions have not previously been systematically assessed nor comparably reviewed for CHD, stroke, and diabetes. Though some evidence on diet and cardiometabolic diseases has been previously assessed, no contemporary investigation comprehensively evaluated multiple dietary factors while also including qualitative assessment of evidence for causality [7], quantitative assessments of etiologic dose-responses [7] and optimal consumption levels [7–9].

To address these gaps in knowledge, we systematically reviewed the evidence for effects of dietary factors on cardiometabolic diseases, including comprehensively assess evidence for causality; estimate magnitudes of etiologic effects focusing on dose-responses rather than simple categorical comparisons; evaluate heterogeneity and potential for bias in these etiologic effects; and determine optimal population intake levels. We hypothesized that certain individual dietary components would have probable or convincing evidence for causal effects on cardiometabolic diseases; and that magnitudes of estimates would be reasonably unbiased based on validity analyses.

Methods

Evidence for causality

Our methods for evaluating strength of evidence for causal diet-chronic disease relationships were reported [10]. We searched for dietary factors with evidence for causal effects on total cardiovascular disease (CVD), CHD, stroke, or diabetes. Given paucity of evidence from randomized controlled trials, our primary determinations were based on Bradford-Hill criteria [11] graded independently and in duplicate (RM, DM), including evidence on strength/consistency, temporality, coherence, specificity, analogy, plausibility, biological gradient, and supportive experimental data (Text A in S1 File). In our final analysis, we conservatively only included factors which were determined to have probable or convincing evidence for causal effects. Based on our and other recent reviews [12], many dietary factors were evaluated and determined not to achieve these criteria for causality; e.g., a leading candidate not achieving sufficient evidence was coffee, and others were extra-virgin olive oil, monounsaturated fat, cocoa, and tea (Text B in S1 File). We also qualitatively considered concordance of our conclusions with other criteria for causality of diet-chronic disease relationships as probable or convincing, such as from the World Health Organization (WHO) and WCRF/AICR [13-15]. Overall, we elected to be conservative in our approach, excluding rather than including dietary factors with borderline judgments on at least probable causal evidence. As evidence continues to accrue, we hope to update this investigation in future years using similar standardized methods. For the present work focused on diet, we did not evaluate alcohol which is often considered separately as a potentially addictive substance, is implicated in accidental deaths, and for which health effects have been evaluated [16].

Literature searches for etiologic effects

For each identified diet-disease relationship, we performed multiple systematic searches of PubMed through 1/May/2015 to identify meta-analyses of randomized controlled trials (RCTs) or prospective cohort studies evaluating these specific dietary factors and total CVD, CHD including subtypes (fatal, nonfatal), stroke including subtypes (ischemic, hemorrhagic), or diabetes. For sodium and sugar-sweetened beverages (SSBs), we also reviewed effects on blood pressure (BP) and obesity, respectively, based on RCTs demonstrating primary effects on these risk pathways. Our detailed protocol for identifying studies on etiologic effects of dietary habits on chronic diseases has been reported [10]. Search terms and results are provided in S1 File (Text B, Table A, Figure A). For each search, one investigator screened all electronically identified titles and abstracts and, for all articles selected for full-text review, further hand-searched the citation lists and also the first 20 "related articles" in PubMed. These searches were supplemented with additional expert contacts to identify all potentially relevant articles.

For a few dietary factors for which evidence for causal effects on specific cardiometabolic outcomes was identified, recently published meta-analyses were unavailable. For these diet-disease relationships, we performed *de novo* meta-analyses according to PRISMA guidelines (S2 File) [17]. These included systematic searches of online databases and hand-searching of reference lists and related citations. For each meta-analysis, titles and abstracts of identified studies were screened by one investigator, and relevant full-texts were reviewed independently and in duplicate by two investigators. Protocols for these meta-analyses are provided in S1 File (Text C-D).

Study inclusion

Published meta-analyses were eligible if including RCTs or prospective cohorts of the identified diet-disease relationship of interest. Whenever possible, we prioritized meta-analyses that characterized dose-responses using all available data (as opposed to comparisons of extreme categories, e.g., high vs. low). Meta-analyses including only retrospective case-control studies were excluded due to greater potential for selection bias, recall bias, and reverse causation. When more than one meta-analysis was identified for any diet-disease relationship, we included the dose-response analysis with the greatest number of studies and clinical events. When recent meta-analyses were identified, they were not updated. We only included published, peer-reviewed meta-analyses; or performed *de novo* meta-analyses with all methods presented. For new meta-analyses, we included all RCTs and prospective cohorts that assessed the diet-disease relationship of the interest (Text C in S1 File). Studies were excluded if they only reported crude estimates, lasted less than 3 months, or focused on special populations (e.g., comparisons of vegetarians vs. non-vegetarians).

Data extraction

For each published meta-analysis, we extracted data independently and in duplicate using a standardized electronic spreadsheet on definitions of dietary factors and outcomes, numbers of studies included, pooled risk estimates and corresponding uncertainty, study designs, sample sizes, numbers of events, mean ages of participants at baseline, lengths of follow-up, ranges of intake, statistical methods, evidence for bias, and control for confounders in individual studies. In most cases, all required data were not reported in the original meta-analyses and were therefore extracted from the original individual studies cited in the meta-analysis.

For new meta-analyses, data were extracted independently and in duplicate from each identified individual study using a standardized electronic spreadsheet. Data were extracted on author name, contact information, publication year, study name, location, design, population (age, sex, race, sample size), follow-up duration, exposure/intervention (definition, assessment, categories, dose in each category), outcome (definition, ascertainment), analysis method, covariates, number of events, and the risk estimate and its corresponding uncertainty in each exposure/intervention category. For each meta-analysis, we standardized the risk estimates to a common standardized serving size to enable comparability across studies.

Evidence synthesis

Data synthesis utilized published results when dose-response meta-analyses were performed in the published report, or categorical comparisons when such findings were unavailable; and otherwise, when possible, data from the individual original articles in each meta-analysis to perform new dose-response meta-analyses. For *de novo* meta-analyses, we extracted data on each individual study as described above and performed random-effects two-step generalized least squares for trend estimation (GLST command in STATA) [18, 19]. This method utilizes all available data to compute study-specific dose-response estimates based on the natural log relative risk (RR) in each category of intake and pools these to estimate an overall RR for a standardized serving and frequency of intake. We assessed between-study heterogeneity using Cochran's Q and I² statistic. I² values of 25–50%, 50–75%, and >75% were considered to represent low, moderate, and high heterogeneity, respectively [20]. Potential for publication bias was explored statistically using Begg's test [21] and by visual inspection of funnel plots. All analyses were conducted with STATA 14.0 software (StataCorp).

Heterogeneity in etiologic effects

Proportional effects (RRs) of major cardiometabolic risk factors have been shown to decline with age [3, 22]. To quantify and incorporate this effect modification by age, we evaluated the proportional differences in RRs for major diet-related cardiometabolic risk factors, including

systolic blood pressure (SBP), body mass index (BMI), fasting plasma glucose and total cholesterol, across 6 age groups from 25–34 to 75+ years (Text F & Figure B in <u>S1 File</u>) [22–25]. Because proportional differences between adjacent age groups were quite similar across these four risk factors, we applied the mean proportional differences in RR by age across all risk factors to the dietary RRs, anchored at the mean age at event of each diet-disease pair. In applying these to diet, we used Monte Carlo simulations to estimate the uncertainty in the age-distributed log RRs, sampling from the distribution of log RRs at the age at event. Based on 1,000 simulations, we utilized the 2.5th and 97.5th percentiles to derive the 95% uncertainty interval, hereafter described as the 95% confidence interval (CI). We also reviewed the findings of meta-analyses to consider potential effect modification by sex and, where relevant, other factors such as race, hypertensive status, and BMI.

Optimal intakes

To permit comparable quantitative assessment of impacts on disease, we characterized the optimal population consumption levels of each dietary factor for risk of cardiometabolic diseases [10, 26]. Optimal levels were selected primarily based on risk (observed consumption levels associated with lowest disease risk in meta-analyses) with further considerations of feasibility (observed national mean consumption levels in nationally representative surveys worldwide) [27–32] and consistency with other assessments (existing major dietary guideline reports) [33–35]. Because populations inevitably have a range of consumption levels, we utilized a normal distribution around each optimal intake level with standard deviation (SD) equaling 10% of the mean, consistent with optimal distribution ranges of metabolic risk factors [3, 36–39].

Assessment of validity and bias

Estimated etiologic effects could be limited by confounding (typically causing overestimation of effects) and measurement error (typically causing underestimation of effects). Measurement error was generally not addressed in most studies, although some utilized serial measures of diet. To reduce bias from confounding nearly all identified observational studies evaluating etiologic effects utilized multivariable adjustment for major demographic factors and, in many cases, other dietary factors. Yet, we recognized that clustering of dietary patterns could still cause unmeasured confounding, e.g., from clustering of healthful factors such as fruits, vegetables, and whole grains and inverse correlations of these with harmful factors such as SSBs or processed meats. Thus, even with multi-variable adjustment, our final calculated etiologic effects from studies of an individual dietary component might overestimate its effects, as compared with the true effect when the dietary component is consumed as part of an overall diet pattern.

To assess potential bias from dietary pattern effects, we performed 3 validity analyses (Tables S4-S6 in S1 File), based on: (a) prospective long-term observational studies evaluating overall dietary patterns and clinical cardiovascular events; (b) randomized controlled feeding trials evaluating overall dietary patterns and cardiovascular risk factors (LDL-cholesterol, SBP); and (c) a large RCT evaluating overall dietary patterns and clinical cardiovascular events. For each, we compared the observed effect from the dietary pattern study to the estimated RR calculated by jointly considering the individual etiologic effects (RRs) for each dietary component in that pattern.

For prospective cohorts evaluating overall diet patterns and CVD events [40–44], the observed multivariable-adjusted RR in each category (e.g., quintile) of the dietary pattern was compared to the estimated effect calculated by combining the reported differences in each

individual dietary component (e.g., fruits, nuts) across each category of the diet pattern with our estimated individual etiologic effect (RR) for that dietary component, assuming a multiplicative relation between RRs for individual components. We focused on foods and excluded overlapping components (e.g., we included whole grains, fruits, and vegetables; and excluded dietary fiber); we also assumed no benefits from differences in other dietary factors (e.g., coffee) in the pattern for which we had not determined a causal etiologic effect.

For randomized controlled feeding trials of dietary patterns and CVD risk factors, we performed inverse-variance-weighted meta-regression across all of the treatment arms of three large, well-established dietary pattern trials [45-47] to estimate the independent effects of five individual dietary components, when consumed as part of an overall dietary pattern, on SBP and LDL-cholesterol. We evaluated achieved changes in fruits, vegetables, nuts, whole grains, and fish simultaneously as independent variables, with changes in SBP or LDL-C as the dependent variable. For each dietary component, we then calculated how the identified change in SBP and LDL-C from the meta-regression would alter cardiovascular risk, based on the established relationship between SBP and LDL-C and clinical events [48-52] assuming independent, multiplicative effects of SBP and LDL-C. These effects, calculated based only on how each dietary component altered SBP and LDL-C in randomized controlled feeding trials of diet patterns, were then compared to our estimated etiologic effect on cardiovascular events for that dietary component. We recognized that the calculated effects based on the feeding trial results might be conservative, as they presume that the summed CVD benefits of these dietary factors are attributable only to effects on SBP and LDL-C, when in reality other pathways of benefit likely exist.

Lastly, we compared the observed vs. estimated risk using findings from the PREDIMED trial, a large RCT evaluating the effects of two overall dietary patterns on CVD incidence [53]. The estimated risk reductions were calculated by combining the observed differences in individual dietary components achieved in the trial with our estimated quantitative effects for each dietary component, assuming multiplicative effects of each individual component.

Results

Dietary factors with evidence for causality

We identified 10 foods and 7 nutrients with probable or convincing evidence for causal effects on specific cardiometabolic outcomes (Table 1). Among different criteria, the strength of association was most variable, and coherence, temporality, and biologic gradient were least variable (Table 2).

Our systematic searches to evaluate etiologic effects for these 17 foods and nutrients identified 896 potentially relevant meta-analyses or reviews articles, of which 23 were finally included in our estimates (Table A & Figure A in <u>S1 File</u>), including 1 de novo meta-analysis for 4 diet-disease relationships (Text E & Tables B-C in <u>S1 File</u>). We did not find sufficient probable or convincing evidence of causal etiologic effects on cardiometabolic diseases of many other dietary factors of interest, for example dietary cholesterol, plant omega-3 fats, monounsaturated fats, eggs, poultry, tea, coffee, or cocoa.

Etiologic effects on cardiovascular diseases

Sixteen of the identified dietary factors had evidence for causal effects on CVD (Table 1). Among different clinical events, fruits, fish, and fiber were most frequently studied in relation to CHD (16 cohorts each) (Table 3). The total numbers of people in each meta-analysis of clinical events ranged from about 140,000 for *trans*-fats and CHD to about 820,000 for fruits and CHD. The largest number of total events was for processed meats and CHD (21,308 events);



Table 1. Dietary factors and cardiometabolic outcomes with probable or convincing evidence for an etiologic relationship¹.

| Dietary Risk Factor | Cardiovascular Outcomes | Metabolic Outcomes |
|--|--|----------------------------|
| Foods | | |
| Low fruits ² | CHD, ischemic stroke, hemorrhagic stroke | |
| Low vegetables ³ | CHD, ischemic stroke, hemorrhagic stroke | |
| Low beans/legumes ³ | CHD | |
| Low nuts/seeds | CHD | Diabetes |
| Low whole grains | CVD, CHD | Diabetes |
| High red meats, unprocessed ⁴ | | Diabetes |
| High processed meats ⁵ | CHD | Diabetes |
| Low fish/seafood | CHD (fatal) | |
| Low yogurt | | Diabetes |
| High sugar-sweetened beverages | CHD | Diabetes, BMI ⁶ |
| Nutrients | | |
| Low polyunsaturated fats (replacing either carbohydrates or saturated fats) ⁷ | CHD | |
| Low seafood omega-3 fats ⁸ | CHD (fatal) | |
| High <i>trans</i> -fats | CHD | |
| Low dietary fiber | CVD, CHD, stroke | Diabetes |
| High glycemic load | CHD, stroke | Diabetes |
| High dietary sodium | CVD (fatal), systolic BP ⁹ | |
| Low dietary potassium | Stroke | |

¹ See <u>Table 2</u> for details on assessments of causality of each relationship.

² Excluding 100% juices.

³ Excluding vegetable juices, starchy vegetables such as potatoes or corn, and salted or pickled vegetables. Because certain beans/legumes (e.g., black beans, lentils) were commonly included as vegetables in many of the identified studies, the etiologic effects identified for vegetables should be considered as representing the effect of vegetables including beans/legumes. We also evaluated etiologic effects of beans/legumes separately.

⁴ Beef, lamb, or pork; excluding poultry, fish, eggs, and processed meat.

⁵ Any meat preserved by smoking, curing, or salting or addition of chemical preservatives, such as bacon, salami, sausages, hot dogs, or processed deli or luncheon meats, and excluding fish or eggs.

⁶ In addition to the effect of sugar-sweetened beverages (SSBs) on adiposity (body mass index, BMI), evidence from prospective studies suggested an additional, BMI-independent effect of SSBs on incidence of type 2 diabetes and coronary heart disease (CHD).

⁷ Reported effects are nearly identical for polyunsaturated fats replacing carbohydrates or saturated fats.

⁸ Eicosapentaenoic acid (EPA) + docosahexaenoic acid (DHA).

⁹ We identified concordant evidence for direct effects on fatal cardiovascular disease (CVD) and systolic blood pressure (BP).

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the fewest, for fruits or vegetables and hemorrhagic stroke (1,535 events). Across the metaanalyses for CVD, the median age at event was 61.1 years (range: 50 to 72.2 years). Relative risks were generally modest, ranging from 0.73–0.95 per daily serving of protective foods, and 1.12–1.37 for harmful foods. Some of the larger effects were for fruits and hemorrhagic stroke (RR:0.73 per daily serving), nuts/seeds and fatal CHD (0.76 per 4 servings/week), beans/ legumes and CHD (0.77 per daily serving), and processed meats and CHD (1.37 per daily serving). Dietary sodium increased BP with a monotonic dose-response, with identified heterogeneity in this effect by age, race, and hypertensive status; with consistent evidence for higher risk of fatal CVD comparing high vs. low intakes. Conversely, dietary potassium was linked to lower risk of stroke, with a RR of 0.87 per 1,000 mg/d.

Etiologic effects on diabetes

Only 8 identified dietary factors had probable or convincing evidence for causal effects on diabetes (Table 1), including protective effects of nuts/seeds, whole grains, yogurt, and dietary

| Fields Fields Field < | Dietary Factor | Cardiometabolic Outcome | Strength | Consistency | Temporality | Coherence | Specificity | Analogy | Plausibility | Biological Gradient | Experiment |
|--|--|----------------------------|---|---|---|---|---|-----------------------|---|---|---|
| Fuls CHO +++ ++++ ++++ +++++ +++++ +++++ +++++ +++++ +++++ +++++ +++++ +++++ +++++ +++++ +++++ +++++< +++++< +++++< +++++< +++++ +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< +++++< </td <td>Foods</td> <th></th> <td></td> <td></td> <td></td> <td></td> <td></td> <td></td> <td></td> <td></td> <td></td> | Foods | | | | | | | | | | |
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| CHO i | | Hemorrhagic stroke | +++++++++++++++++++++++++++++++++++++++ | ++++ | ++++ | ++++++ | ‡ | +++5 | ++++ | +++++++++++++++++++++++++++++++++++++++ | ++ ²³ |
| International production in the product of | Vegetables | CHD | + | +++ | ++++ | ++++++ | ‡ | +++5 | ++++ | ŧ | ++ ²³ |
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| $ \begin{array}{ c c c c c c c c c c c c c c c c c c c$ | Nuts/seeds | CHD (fatal) | +++++++++++++++++++++++++++++++++++++++ | ++++ | ++++ | +++++++++++++++++++++++++++++++++++++++ | ‡ | +++2 | ++++ | ŧ | +++++ |
| Minde grains Diabates i++ i++ i++ i++ i+++ | | CHD (non-fatal) | ++++++ | ++++ | + | ++++ | + | +++5 | +++++++++++++++++++++++++++++++++++++++ | +++ | ++++ |
| | | Diabetes | ‡ | ++++ | ++++ | ++++++ | ‡ | +++5 | ++++ | ŧ | +++ |
| $ \begin{array}{ c c c c c c c c c c c c c c c c c c c$ | Whole grains | CVD | + | ++++ | ++++ | ++++ | ‡ | 9+++ | ++++ | ‡ | +++ |
| | | CHD | + | ++++ | ++++ | +++++++++++++++++++++++++++++++++++++++ | ‡ | 9+++ | ++++ | ŧ | +++ ²⁴ |
| | | Diabetes | ‡ | ++++ | +++++++++++++++++++++++++++++++++++++++ | ++++++ | ‡ | 9+++ | +++++++++++++++++++++++++++++++++++++++ | + | +++ ²⁴ |
| | Red meats, unprocessed | Diabetes | ‡ | +++++++++++++++++++++++++++++++++++++++ | +++++++++++++++++++++++++++++++++++++++ | +++++++++++++++++++++++++++++++++++++++ | +++++++++++++++++++++++++++++++++++++++ | ×+++ | ‡ | +++++ | ++++ |
| | Processed meats | CHD | +++++++++++++++++++++++++++++++++++++++ | ++++ | +++++++++++++++++++++++++++++++++++++++ | ++++++ | ‡ | 8+++ | +++++++++++++++++++++++++++++++++++++++ | + | +++++ |
| | | Diabetes | +++++ | ++++ | ++++ | ++++++ | ‡ | +++7 | + | +++++ | +++ |
| Sugar-sweetened Bolymass index $\mu + \mu$ | Yogurt | Diabetes | ‡ | +++ | ++++ | ++++++ | +++++ | 6++ | ++ | +++++++++++++++++++++++++++++++++++++++ | ++ ²⁵ |
| | Sugar-sweetened | Body mass index | +++ | ++++ | ++++ | ++++++ | +++++ | +++ ^{10, 11} | ++++ | ŧ | +++++ |
| $\begin{tabular}{ c c c c c c c c c c c c c c c c c c c$ | beverages | CHD | ‡ | ++++ | ‡ | ++++++ | ‡ | ++10 | ‡ | +++++++++++++++++++++++++++++++++++++++ | +++ |
| Intrinsical set in the set of | | Diabetes | +++++++++++++++++++++++++++++++++++++++ | ++++ | ++++ | ++++ | +++ | +++ | +++++ | + | +++++ |
| | Nutrients | | - | | | | | | | | |
| | PUFAs | CHD | + | +++ | ++++ | +++++ | ++++ | +12 | ++++ | # | +++++++++++++++++++++++++++++++++++++++ |
| | Seafood w-3s (fish/ seafood) ² | CHD (fatal) | ‡ | +++++++++++++++++++++++++++++++++++++++ | +++++++++++++++++++++++++++++++++++++++ | +++++++++++++++++++++++++++++++++++++++ | +++++++++++++++++++++++++++++++++++++++ | +++ | +++++++++++++++++++++++++++++++++++++++ | +++ ²² | +++++++++++++++++++++++++++++++++++++++ |
| $ \begin{array}{llllllllllllllllllllllllllllllllllll$ | Trans-fats | CHD | +++++++++++++++++++++++++++++++++++++++ | ++++ | ++++ | +++++++++++++++++++++++++++++++++++++++ | ++++ | ++ | ++++ | + | +++++++++++++++++++++++++++++++++++++++ |
| $ \begin{array}{llllllllllllllllllllllllllllllllllll$ | Dietary fiber | CVD | +++++++++++++++++++++++++++++++++++++++ | ++++ | ++++ | +++++++++++++++++++++++++++++++++++++++ | ‡ | +++ ¹⁵ | +++++ | ŧ | +++++ |
| $ \begin{array}{llllllllllllllllllllllllllllllllllll$ | | CHD | +++++++++++++++++++++++++++++++++++++++ | +++ | ++++ | ++++ | + | +++ ¹⁵ | +++++++++++++++++++++++++++++++++++++++ | +++ | +++ |
| $ \begin{array}{l lllllllllllllllllllllllllllllllllll$ | | Stroke | ‡ | ++++ | ++++ | +++++++++++++++++++++++++++++++++++++++ | ‡ | +++ ¹⁵ | ++++ | ŧ | +++ |
| $ \begin{array}{c ccccccccccccccccccccccccccccccccccc$ | | Diabetes | ‡ | ++++ | ++++ | +++++++++++++++++++++++++++++++++++++++ | ‡ | +++ | +++++ | ŧ | +++++ |
| $ \begin{array}{c ccccccccccccccccccccccccccccccccccc$ | Glycemic load | CHD | +++3 | +++3 | ++++ | ++++++ | ‡ | ++10 | ++ | ++ ²¹ | ++++ |
| Diabetes + +++< | | Stroke | ‡ | ++++ | ++++ | ++++++ | ‡ | +17 | ++ ²⁰ | ++ ²¹ | ++++ |
| Sodium Systolic BP ++4 +++ +++ +++ +19 +++ | | Diabetes | + | ++ | +++ | ++++ | + | +++1 | ++++ | ++ ²¹ | +++ |
| CVD (fatal) ++ ++ ++ +++ +++ +++ +++ +++ +++ +++ | Sodium | Systolic BP | ++ 4 | ++++ | +++ | +++++ | ++++ | +19 | ++++ | ++++ | +++++ |
| - | | CVD (fatal) | ‡ | +++ | ++++ | ++++++ | ‡ | +++ ¹⁸ | ++++ | + | ++++ |

| lietary Factor | Cardiometabolic Outcome | Strength | Consistency | Temporality | Coherence | Specificity | Analogy | Plausibility | Biological Gradient | Experiment |
|----------------------------|-------------------------------------|-----------------|--------------------|-------------------------------------|-----------------|---------------------|-------------------|-------------------|------------------------|-----------------|
| Potassium | Stroke | + | +++ | +++ | +++ | ++++ | +++ ¹⁸ | +++ | +++ | +++++ |
| To score each Bradford-F | Hill criterion, the following ge | eneral princi | oles were utilize | d, focusing on | evidence fron | n meta-analys | ies of prosp | bective cohort | studies and/or rand | omized |
| mitations: + Emerging evi | dence from a few studies or | r conflicting r | esults from seve | relatively lew l eral studies;—(| criterion not m | et. Definition | s for each c | of the nine crite | ria and adaptations | to the |
| eneral scoring system we | re as follows: Strength: ma | agnitude of a | issociation, inclu | uding RRs for p | rotective fact | ors of >0.9 (+ | , 0.8–0.89 | (++), or <0.8 (| +++); and for harmfi | ul factors, of |
| 1.11 (+), 1.25 (++), and > | 1.25 (+++). Since magnitud | le is directly | dependent on bo | oth the selected | d serving size | and frequend | y of consul | mption, we util | ized serving sizes n | nost similar to |
| tandard dietary guideline | s and frequencies of consur | nption repre- | senting modest, | standardized (| differences in | intake (e.g., 1 | serving/d | of fruit) that ar | e easily communica | ted and could |
| e feasibly achieved by an | intervention. Consistency | : associatior | n is repeatedly o | bserved in diffe | erent populati | ons and circul | nstances, i | ncluding ≥80° | % of included study- | specific |
| stimates being in the exp | sected direction (+++); \geq 60 - | · <80% (++); | ≥40 - <60% (+) | ; and <40% (no | ot meeting crit | eria). Tempo | rality: expo | osure precede | s outcome. Becaus | e all evidence |
| as based on longitudinal | studies, this was a necessa | IN criterion (| +++); when rela | tively few over: | all studies we | re available (< | 5), we grad | ded this criteri | on conservatively as | ‡ |
| :oherence: interpretation | of association does not cor | oflict with kno | own natural histo | ory and biology | of the diseas | e, for exampl | e based on | pathways of c | lisease occurrence | and laboratory |
| ndings on the dietary fact | or. Specificity: exposure lir | nked to a spe | ecific outcome. I | Because many | nutritional fac | stors can plau | sibly have | diverse effects | and influence multi | ple |
| utcomes, scoring was ba: | sed on three principles: 1) d | lietary factor | influences a me | chanism/pathv | vays known t | o cause the or | utcome; 2) | dietary factor | not associated with | nultiple other, |
| nrelated non-communica | ole diseases (e.g., multiple | cancers, chr | onic obstructive | pulmonary dis | ease (COPD | i); 3) dietary a | ssociation | has additional | specificity within the | e set of |

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Given their common sources, these factors were evaluated together based on studies of fish/seafood, dietary long-chain omega-3 fats, and fish oil supplements. In secondary stratified analyses, +++ for women, and—for men (main effect: null)

aboratory studies

curve.

cardiometabolic outcomes (coronary heart disease (CHD), stroke, diabetes mellitus). Analogy: based on the effects of similar factors on the disease outcome; see detailed footnotes

below. Plausibility: association supported by one or more credible biological mechanisms. Biological gradient: exposure and outcome are related by a monotonic dose-response Experiment: association is also supported by evidence from randomized controlled trials on intermediate risk factors (or, less commonly, disease outcomes) plus supportive

Effect size does not correspond to relative risk, but comparison with effect sizes on body mass index or blood pressure (BP) for other lifestyle-based interventions. For BP, the overall scored strength reflects the average of +++ for older adults, blacks, and hypertensives; and + for healthy, white, younger adults.

Based on analogies with other minimally processed, higher fiber, phytochemical rich foods.

Based on analogies with other less-processed foods, dietary fiber, and glycemic load.

Based on analogies to processed meats (or unprocessed red meats), blood ferritin levels, and hemochromatosis.

Based on analogies to sodium.

Based on analogies to probiotics in relation to weight gain.

¹⁰ Based on analogies to other poor-quality carbohydrates in relation to both CHD and weight gain.

¹¹ Based on analogies to other poor-quality carbohydrates in relation to both diabetes mellitus and weight gain.

¹² Based on analogies to vegetable oils in relation to CHD and cardiovascular risk factors.

¹³ Based on analogies to fish.

¹⁴ Based on analogies to other dietary fats.

¹⁵ Based on analogies to other higher-quality carbohydrates and other fiber-rich foods such as nuts, fruits, and vegetables.

¹⁷ Based on analogies to other poor-quality carbohydrates in relation to diabetes mellitus, and diabetes as a risk factor for stroke. ¹⁶ Based on analogies to other higher-quality carbohydrates in relation to both diabetes mellitus and weight gain

¹⁸ Based on analogies to other lifestyle-related and nonlifestyle-related blood pressure interventions and to foods high in sodium (e.g., processed meats).

¹⁹ Based on analogies to potassium.

²⁰ Based on insulin resistance/diabetes mellitus pathways.

²¹ Several individual studies show dose response; no published dose-response meta-analyses.

²² Between 0 and 250 mg/d; meta-analyses suggest no major additional benefits for fatal CHD above 250 mg/d.

23 Because while strong and consistent evidence from trials of dietary patterns rich in fruits and vegetables, few trials separately evaluated only fruits or vegetables.

²⁴ Based on overall effects of carbohydrate quality, including studies of dietary fiber and glycemic load; much less evidence for benefits of whole grains independent of dietary fiber

²⁵ Based on findings for yogurt and weight gain (animal studies, human cohorts) and for probiotics and weight gain (animal and human experiments). and glycemic load.

| Table 3. Estimates | of etiologic effects of dietary | factors and risk of c | ardiovascular diseases | and type 2 dia | abetes mell | itus ¹ . | | |
|---------------------------|---------------------------------|---|--|--------------------|------------------|---|--------------------------|-------------------------------------|
| Dietary Factor | Outcome | Studies in Each Meta-analysis ² | Source | No. of Subjects | No. of Events | Unit of RR ³ | RR (95% CI) ³ | Statistical Heterogeneity |
| Foods | | | | | | | | |
| Fruits ⁴ | ¢CHD | 16 cohorts (22 estimates) | Gan, 2015 [71] | 817,977 | 13,786 | per 1 serving/d (100 g/d) | 0.94 (0.91, 0.98) | l ² = 31.7% p = 0.08 |
| | ↓ Ischemic stroke | 9 cohorts (10 estimates) | <i>De novo</i> meta- analysis ⁵ [72–79] | 329,204 | 5,517 | per 1 serving/d (100 g/d) | 0.88 (0.83, 0.93) | l ² = 77.1% p<0.001 |
| | ↓ Hemorrhagic stroke | 5 cohorts (7 estimates) | <i>De novo</i> meta- analysis ⁵ [74–77] | 175,035 | 1,535 | per 1 serving/d (100 g/d) | 0.73 (0.62, 0.87) | l ² = 81.4% p<0.001 |
| Vegetables ⁶ | (CHD | 9 cohorts | Gan, 2015 [71] | 761,612 | 13,135 | per 1 serving/d (100 g/d) | 0.95 (0.92, 0.98) | l ² = 35.6% p = 0.07 |
| | ↓ Ischemic stroke | 9 cohorts (10 estimates) | <i>De novo</i> meta- analysis ⁵ [72–75, 77– 79] | 329,204 | 5,515 | per 1 serving/d (100 g/d) | 0.83 (0.75, 0.93) | l ² = 89.9% p<0.001 |
| | ↓ Hemorrhagic stroke | 5 cohorts (7 estimates) | <i>De novo</i> meta- analysis ⁵ [74–77, 79] | 175,035 | 1,535 | per 1 serving/d (100 g/d) | 0.83 (0.72, 0.96) | l ² = 30% p = 0.20 |
| Beans/legumes | ¢ CHD | 5 cohorts | Afshin, 2014 [80] | 198,904 | 6,514 | per 1 serving/d (100 g/d) | 0.77 (0.65, 0.90) | l ² = 0.2% p = 0.41 |
| Nuts/seeds | ↓ CHD (fatal) | 1 RCT and 5 cohorts | Afshin, 2014 [80] | 206,114 | 6,749 | per 4 servings/wk (4 oz/wk) | 0.76 (0.69, 0.84) | l ² = 27.2% p = 0.23 |
| | ↓ CHD (non-fatal) | 1 RCT and 3 cohorts | Afshin, 2014 [80] | 141,390 | 2,101 | per 4 servings/wk (4 oz/wk) | 0.78 (0.67, 0.92) | l ² = 0.0% p = 0.46 |
| | ↓ Diabetes | 1 RCT and 5 cohorts | Afshin, 2014 [80] | 230,216 | 13,308 | per 4 servings/wk (4 oz/wk) | 0.87 (0.81, 0.94) | l ² = 21.6% p = 0.27 |
| Whole grains | ¢CVD | 7 cohorts (9 estimates) 7 | New GLST ⁸ of Mellen 2008 [54] | 285,217 | 7,005 | per 1 serving/d (50 g/d) | 0.91 (0.86, 0.97) | l ² = 84.0% p<0.001 |
| | ¢ CHD | 6 cohorts | New GLST ⁸ of Mellen 2008 [54] | 281,633 | 4,593 | per 1 serving/d (50 g/d) | 0.97 (0.94, 0.99) | l ² = 75.5% p = 0.001 |
| | ↓ Diabetes | 10 cohorts | Aune 2013 [59] | 385,868 | 19,791 | per 1 serving/d (50 g/d) | 0.88 (0.83, 0.93) | l ² = 82% p = <0.0001 |
| Red meats, unprocessed | ↑ Diabetes | 9 cohorts (10 estimates) | Pan, 2011 [81] | 442,101 | 28,228 | per 1 serving/d (100 g/d) | 1.19 (1.04, 1.37) | l ² = 93% p<0.001 |
| Processed meats | ↑ CHD | 5 cohorts (6 estimates) | Micha, 2010 [82] | 614,062 | 21,308 | per 1 serving/d (50 g/d) | 1.37 (1.11, 1.68) | l ² = 76.2% p = 0.001 |
| | ↑ Diabetes | 8 cohorts (9 estimates) | Pan, 2011 [81] | 371,492 | 26,256 | per 1 serving/d (50 g/d) | 1.51 (1.25, 1.83) | l ² = 94.3% p<0.001 |
| Fish/Seafood ⁹ | ↓ CHD (fatal) | 16 cohorts (17 estimates) | Zheng, 2012 [83] | 315,812 | 4,472 | per 15 g/d (~1– 100 g- serving/ wk) | 0.94 (0.90– 0.98) | l ² = 63.1 p<0.005 |
| Yoghurt | ↓ Diabetes | 9 cohorts | Chen, 2014 [<u>84</u>] | 408,096 | 32,995 | per 1 serving/d (8 oz/d, 244 g/d) | 0.82 (0.70, 0.96) | $l^2 = 65.3 \text{ p} = 0.003$ |

(Continued)

| | | | , | | | c | C | , |
|---------------------------------------|---|---|----------------------------|--------------------|------------------|--|--|------------------------------|
| Dietary Factor | Outcome | Studies in Each Meta-analysis ² | Source | No. of Subjects | No. of Events | Unit of RR ⁵ | RR (95% CI) ³ | Statistical Heterogeneity |
| Sugar-sweetened beverages | ↑ BMI (when baseline BMI <25 kg/m ²) ¹⁰ | 3 cohorts | Mozaffarian, 2011 [55] | 120,877 | n/a | per 1 serving/d (8 oz/d) ¹¹¹ | 0.10 kg/m ² (0.05, 0.15) | not reported |
| | ↑ BMI (when baseline BMI ≥25 kg/m²) ¹⁰ | 3 cohorts | Mozaffarian, 2011 [55] | 120,877 | n/a | per 1 serving/d (8 oz/d) ¹¹ | 0.23 kg/m ² (0.14, 0.32) | not reported |
| | \uparrow Diabetes (BMI-adjusted) ¹⁰ | 17 cohorts | Imamura, 2015 [85] | 464,937 | 38,253 | per 1 serving/d (8 oz/d) ¹¹ | 1.27 (1.10, 1.46) | l ² = 73% |
| | \uparrow CHD (BMI-adjusted) ⁹ | 4 cohorts | Xi, 2015 [86] | 173,753 | 7,396 | per 1 serving/d (8 oz/d) ¹¹ | 1.17 (1.10, 1.24) | $l^2 = 0.0\% p = 0.79$ |
| Nutrients | | | | | | | | |
| PUFA replacing Carbs ¹² | ↓ CHD | 9 cohorts (12 estimates) | Farvid, 2014 [58] | 262,612 | 12,198 | per 5%E/d | 0.90 (0.85, 0.94) | $l^2 = 47.3\% p = 0.04$ |
| PUFA replacing SFA ¹² | ↓ CHD | 8 cohorts (11 estimates) | Farvid, 2014 [58] | 262,612 | 12,198 | per 5%E/d | 0.91 (0.87, 0.96) | $l^2 = 55.9\% p = 0.01$ |
| Seafood omega-3 fats ¹³ | ↓ CHD (fatal) | 4 RCTs and 15 cohorts | Mozaffarian 2006 [87] | 363,003 | 5,951 | per 100 mg/d | 0.85 (0.79, 0.92) | not reported |
| Trans-fats ¹⁴ | ↑ CHD | 4 cohorts | Mozaffarian, 2006 [88] | 139,836 | 4,965 | per 2% %E/d | 1.23 (1.11, 1.37) | not reported |
| Dietary fiber ¹⁵ | ↓ CVD | 10 cohorts | Threapleton, 2013 [89] | 1,279,690 | 19,869 | per 20 g/d | 0.76 (0.70, 0.84) | $l^2 = 45\%$ |
| | ¢ CHD | 12 cohorts | Threapleton, 2013 [89] | 1,039,572 | 11,282 | per 20 g/d | 0.76 (0.68, 0.85) | l ² = 33% |
| | ↓ Stroke | 7 cohorts | Threapleton, 2013 [90] | 324,640 | 9,257 | per 20 g/d | 0.81 (0.70, 0.95) | l ² = 59% |
| | ↓ Diabetes | 5 cohorts | Yao, 2014 [91] | 157,336 | 3,029 | per 30 g/d | 0.76 (0.65, 0.88) | not reported |
| Glycemic load ¹⁶ | ↑ CHD | 9 cohorts (13 estimates) | Mirrahimi, 2014 [92] | 262,891 | 11,319 | high vs. low | 1.23 (1.06, 1.42) | $l^2 = 52\% p = 0.02$ |
| | ↑ Stroke | 6 cohorts (9 estimates) | Cai, 2015 [93] | 222,308 | 2,951 | high vs. low | 1.19 (1.05, 1.36) | $l^2 = 5.0\% p = 0.39$ |
| | ↑ Diabetes | 17 cohorts (30 estimates) | Bhupathiraju, 2014 [94] | 698,589 | 46,115 | high vs. low | 1.13(1.08, 1.17) | $l^2 = 26.4\% p = 0.09$ |
| Sodium ^{17, 18} | ↑ CVD (fatal) | 11 cohorts (16 estimates) | Poggio, 2015[95] | 220,249 | 9,628 | high vs. low | 1.12 (1.06, 1.19) | $l^2 = 57.6\%$ p = 0.002 |
| | ↑ SBP, main effect, white, age 50, normotensives ¹⁸ | 103 RCTs (107 estimates) | Mozaffarian, 2014 [24] | 6,970 | NA | per 2,300 mg/d (100 mmol/d) | 3.74 mm Hg (5.18, 2.29) | not reported |
| | \uparrow SBP, additional effect per year of age < or > 50 ¹⁸ | | | | | | 0.105 mm Hg (0.164, 0.047) | |
| | ↑ SBP, additional effect among Blacks ¹⁸ | | | | | | 2.49 mm Hg (4.85, 0.13) | |
| | ↑ SBP, additional effect among hypertensives ¹⁸ | | | | | | 1.87 mm Hg (3.63, 0.12) | |
| | | | | | | | | (Continued) |

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| Meta- | Studies in Each Meta-analysis ² | Source | No. of Subjects | No. of Events | Unit of RR ³ | RR (95% CI) ³ | Statistical Heterogeneity |
|-------------------------|---|---------------------------|--------------------|------------------|---------------------------------|--------------------------|------------------------------|
| Potassium ¹⁷ | 9 cohorts (11 estimates) | D'Elia, 2011 [<u>96]</u> | 233,606 | 7,066 | per 1,000 mg/d (25.7 mmol/d) | 0.87 (0.79, 0.94) | $l^2 = 55\% p = 0.01$ |

Dietary factors with probable or convincing evidence, based on the Bradford-Hill criteria for assessing causality [11], for etiologic effects on cardiometabolic outcomes including ² Number of estimates can be higher than the number of studies if more than one arm in a randomized controlled trial or if estimates were separately reported by sex or age in coronary heart disease (CHD), stroke, cardiovascular disease (CVD), type 2 diabetes, body mass index (BMI), or systolic blood pressure (SBP)

prospective cohort studies

Based on published or de novo dose-response meta-analyses of prospective cohorts or randomized trials. Meta-analyses were evaluated based on design, number of studies and actors on cardiometabolic outcomes vary by age, with an inverse log-linear age association [22]. We derived age specific RRs for diet-cardiometabolic disease relationships based available data by means of generalized least squares (GLST in STATA) for trend estimation. Effect sizes are relative risks (RRs) (95% confidence intervals (CIs)) except for sugarindividual studies per uniform servings of intake. When necessary, original data were extracted from individual studies to perform de novo dose-response meta-analyses using all consumption of each dietary target per unit of RR and respective change in cardiometabolic risk (directionality in risk: 1 increased). Proportional effects of major risk on the age patterns of RRs for metabolic risk factors and incident cardiometabolic disease events (see Figure B in S1 File) [22]. Except as indicated (SSBs, sodium), we did not events, definition of dietary exposure and disease outcomes, statistical methods, evidence of bias, and control for confounders. Relative risks (RRs) were standardized across sweetened beverage (SSB) effects on BMI (absolute in kg/m²) and sodium effects on SBP (absolute in mm Hg). Effect sizes correspond to the relationship between increased dentify sufficient evidence for effect modification by other factors beyond age, e.g. race, obesity, or overall diet quality. Excluding 100% juices.

All four of these de novo meta-analyses were performed for consumption of fruits and vegetables and stroke subtypes due to absence of recent published meta-analyses; all details are provided in S1 File.

commonly included as vegetables in many of the identified studies, the etiologic effects identified for vegetables should be considered as representing the effects of vegetables Excluding vegetable juices, starchy vegetables such as potatoes or corn, and salted or pickled vegetables. Because certain beans/legumes (e.g., black beans, lentils) were ncluding beans/legumes. We also evaluated etiologic effects of beans/legumes separately When a trial did not report an effect for total CVD separately (n = 3 cohorts), CHD and stroke estimates from each trial were first pooled using fixed effects (n = 2 cohorts), or the CHD estimate was used in place of CVD when that was the only reported outcome (n = 1 cohort).

Data were re-extracted from all original investigations identified in the meta-analysis to assess dose-response using two-step generalized least squares for trend estimation [18, 19] Etiologic effects are limited to fatal CHD only due to absence of probable or convincing evidence for benefits on nonfatal CHD events. Benefits for were identified up to 3.5 servings/ week of fish/seafood and 250 mg/d of eicosapentaenoic acid (EPA) + docosahexaenoic acid (DHA), with little evidence for additional benefits at higher intakes.

¹⁰ Available evidence suggests that SSBs increase risk through effects on both BMI and additional BMI-independent effects on type 2 diabetes and CHD [22, 23]

¹¹ Depending on study-specific assumptions, use of UK or US conversion factors, and study weighting, the serving size is in this analysis could also be 8.7–9.1 oz.

¹² Reported effects are nearly identical for polyunsaturated fats (PUFA) replacing carbohydrates (Carbs) or saturated fats (SFA).

¹³ Linear reduction in risk observed until 250 mg/day, with little evidence for additional benefits at higher intakes.

¹⁴ The overall causal effect was based on 4 cohorts; the final RR (95% CI) used herein was very similar but based on the isocaloric replacement of trans-fats with an equal distribution of SFA, monounsaturated fats (MUFA), and PUFA based on a meta-analysis of 2 cohorts.

¹⁵ Possible evidence for larger effects at intakes above 20 g/d.

¹⁶ Glycemic load is calculated as the glycemic index of a food multiplied by its carbohydrate content. Higher values reflect both higher glycemic index and higher quantities of refined grains, starches, and sugars. We also identified evidence for causal effects of dietary fiber. Glycemic load and dietary fiber each overlap with foods in this Table including fruits, vegetables, beans/legumes, nuts/seeds, and whole grains.

 17 Assessed by 24h dietary recall, food frequency questionnaire, or 24h urine excretion.

above or below age 50, there was 0.105 mm Hg (95% CI: 0.047, 0.164) larger or smaller BP reduction, respectively. Effects on CVD vs. SBP were separately identified and are not ¹⁸ Available evidence suggests that sodium increases mortality from CHD, stroke, and other BP-related cardiovascular diseases through effects on SBP [22, 24]. For every year independent (i.e., effects on CVD are at least partly mediated by SBP effects). fiber, and harms of unprocessed red meats, processed meats, SSBs, and glycemic load. SSBs and glycemic load were most frequently studied (17 cohorts each) (Table 3). Processed meats had the strongest estimated effect, with 1.51 RR per daily serving; other foods had more modest effects, such as 0.82 and 0.88 RR per daily serving of yogurt and whole grains, respectively. SSBs had a small but statistically significant etiologic effect on body weight, with smaller effects in normal weight (per daily serving, 0.10 kg/m² increase in BMI) vs. overweight or obese (0.23 kg/m²) individuals.

Heterogeneity in etiologic effects

Proportional effects of most cardiovascular risk factors decline with age (inverse age association), likely related to competing risks, while absolute risk differences increase with age due to increased baseline risk [22]. We evaluated and found similar log-linear inverse associations by age for etiologic effects of major diet-related cardiometabolic risk factors [22]. We therefore applied the mean proportional differences in RRs across 6 age groups (25–34, 35–44, 45–54, 55–64, 65–74, 75+ y) for these risk factors to the dietary RRs (S6 Text & Figure B in S1 File). For both major cardiometabolic risk factors and most diet-disease relationships with sufficient evidence, [22, 54–59] we identified similar RRs by sex. One exception was glycemic load and CHD, for which stronger effects were suggested in women. For effects of SSBs on BMI, we identified and incorporated effect modification by baseline BMI, with larger effects among overweight compared with normal weight individuals [55]. For effects of sodium on SBP, we identified and incorporated joint effect modification by age, race, and hypertensive status [24].

Evidence for optimal intakes

Based on risk as well as feasibility and consistency, we characterized optimal intakes for each dietary factor (Table 4). Potential choices of optimal ranges were broadest for dietary sodium, with differing observed optimal intakes from different CVD outcomes ranging from 614 to 2,391 mg/d [46, 60, 61] and from different major dietary guidelines ranging from 1,200 to 2,400 mg/d [33, 62–65]. Based on all available evidence, we identified a conservative optimal intake level of 2,000 mg/d as previously described [24], consistent with WHO guidelines [64].

Assessment of validity and bias

Because our risk estimates were mostly derived from observational studies of individual dietary components, we performed several validity analyses to compare our estimated etiologic effects to other lines of evidence (Table 5, Tables D-F in S1 File). Evaluating cohort studies of dietary patterns and incident CHD, the estimated vs. observed risks were generally similar. Our estimated etiologic effects did not appreciably overestimate benefits in any study; largest differences were seen in studies of Western dietary patterns (in which our estimated etiologic effects underestimated the observed harms) and in one Greek dietary pattern study (in which our estimated etiologic effects underestimated the observed benefit). Based on changes in BP and LDL-C in dietary pattern feeding trials, the observed effects of individual dietary components were similar to our estimated etiologic effect for that dietary component, except for whole grains for which our estimated etiologic effect was smaller than that predicted by BP and LDL-C changes; and for fish for which our estimated etiologic effect on CHD death was larger than that predicted by BP and LDL-C changes alone. Finally, based on clinical events in a large randomized primary prevention trial, the observed vs. estimated relative risk reductions were similar, except for modest overestimation of benefits based on our etiologic effects in the mixed nuts group.

| Table 4. Data sources and ident | tified optimal intake levels of specific diet | tary factors for reducing cardiometaboli | c diseases ¹ . | |
|--|---|--|---|--|
| Dietary Factor (standardized serving size) | Observed intake levels associated with lowest disease risk in meta-analyses (health outcome) | Observed mean national intakes in 2010 ² | Recommended intakes by major dietary guidelines ³ | Optimal mean population intake ⁴ |
| FOODS | | | | |
| Fruits ⁵ (100 g/serving) | 2.4 servings/d (CHD) 3.4 servings/d (ischemic stroke) 2.1 servings/d (hemorrhagic stroke) | Top 3 countries: • Barbados: 410 g/d (4.1 servings/d) • Jamaica: 315 g/d (3.2 servings/d) • Malaysia: 299 g/d (3.0 servings/d) | DGA 2015: 2 cups/d ● AHA 2020: ≥4.5 cups/d (fruits and vegetables) | 3 (100 g) servings/d |
| Vegetables (including beans/legumes) ⁶ (100 g/serving) | 3.7 servings/d (CHD) 3.4 servings/d (ischemic stroke) 1.5 servings/d (hemorrhagic stroke) | Top 3 countries: • Lao PDR: 364 g/d (3.6 servings/d) • Bhutan: 302 g/d (3.0 servings/d) • Taiwan: 293 g/d (2.9 servings/d) | DGA 2015: 1.8 cups/d (excluding starchy vegetables) AHA 2020: ≥4.5 cups/d (fruits and vegetables) | 4 (100 g) servings/d |
| Beans/legumes (100 g/serving) | • 4.2 servings/wk (CHD) (0.6 servings/d) | Top 3 countries: • Brazii: 194 g/d (1.9 servings/d) • Colombia: 137 g/d (1.4 servings/d) • Mexico: 103 g/d (1.0 serving/d) | DGA 2015: 1 ½ cups/wk (0.2 servings/d) | 1 (100 g) serving/d |
| Nuts/seeds (1 oz [28.35 g)/serving) | 5.0 servings/wk (fatal CHD) 5.2 servings/wk (non-fatal CHD) 4.9 servings/wk (diabetes) | Top 3 countries: • Malaysia: 74 g/d (18.3 servings/wk) • Lebanon: 24 g/d (6.0 servings/wk) • UK: 15 g/d (3.7 servings/wk) | DGA 2015: 5 oz/wk (including soy products) AHA 2020: ≥4 servings/wk⁷ | 5 (1 oz) servings/wk |
| Whole grains (50 g/serving) | 2.5 servings/d (CVD) 2.5 servings/d (CHD) 3.0 servings/d (diabetes) | Top 3 countries: • Germany: 128 g/d (2.6 servings/d) • Barbados: 118 g/d (2.4 servings/d) • Australia: 33 g/d (1, 9 servings/d) | DGA 2015: ≥3 (1 oz) servings/d (≥1.7 servings/d) • AHA 2020: ≥3 (1 oz) servings/d (≥1.7 servings/d) | 2.5 (50 g) servings/d |
| Red meats, unprocessed (100 g/serving) | 0.19 servings/d (diabetes) (1.3 servings/wk) | Bottom 3 countries: • Indonesia: 12 g/d (0.8 servings/wk) • Armenia: 15 g/d (1.0 servings/wk) • Georgia: 15 g/d (1.0 servings/wk) | DGA 2015: 26 oz/wk (lean meat, poultry and eggs) (7.4 servings/wk) | 1 (100 g) serving/wk |
| Processed meats (50 g/serving) | 0.07 serving/d (CHD) (0.5 servings/wk) 0.11 serving/d (diabetes) (0.8 servings/wk) | Bottom 3 countries: Iran: 2.5 g/d (0.4 servings/wk) Korea: 3.1 g/d (0.4 servings/wk) China: 3.3 g/d (0.5 servings/wk) | DGA 2015: Choose fresh lean meat, rather than processed meat AHA 2020: none or ≤2 servings/wk | o |
| Fish/Seatood (100 g/serving) | 3 servings/wk (fatal CHD) 6.5 servings/wk (total stroke) | Top 3 countries: • Japan: 105 g/d (7,4 servings/wk) • Korea: 73 g/d (5.1 servings/wk) • Iceland: 66 g/d (4,6 servings/wk) | DGA 2015: 8 oz/wk (2 servings/wk) AHA 2020: 22 (3.5 oz) servings/wk (preferably oily fish) (>2 servings/wk) | 3.5 (100 g) servings/wk |
| Yoghurt (8 oz (244 g)/serving) | 2.6 servings/wk (diabetes) | Not available | DGA 2015: 3 cups/d of dairy products (21 cups/wk) | 2.5 (8 oz) servings/wk |
| Sugar-sweetened beverages (8 oz (236.5 g)/ serving) | 0 servings/d (body mass index) 0.017 serving/d (diabetes) | Bottom 3 countries: • China: 14 g/d (0.06 servings/d) • Iran: 24 g/d (0.1 servings/d) • Poland: 24 g/d (0.1 servings/d) | DGA 2015: <10%E from added sugars AHA 2020: ≤450 kcal (36 oz)/wk (≤0.6 servings/d) | o |
| NUTRIENTS | | | | |
| Polyumsaturated fats replacing carbohydrates or saturated fats (%E) | 5.1%E (CHD, cohorts) 14.9%E (CHD, randomized trials) | Top 3 countries, PUFA: • Bulgaria: 11.3%E • Lebanon: 9.3%E • Hungary: 8.9%E Bottom 3 countries, SFA: • Japan: 7.0%E • Mexico: 7.4%E • Barbados: 7.6%E | DGA 2015: <10%E from saturated fat replaced with unsaturated especially polyunsaturated fats FAO 2010: 6-11%E from PUFA | 11%E ⁹ |
| Seafood omega-3 fats (mg/d) | • 250 mg/d (tatal CHD) | Top 3 countries: • Barbados: 1,191 mg/d • Iceland: 1,120 mg/d • Japan: 993 mg/d | • DGA 2015: 250 mg/d ⁸ | 250 mg/d |
| | | | | (Continued) |

| Dietary Factor (standardized serving size) | Observed intake levels associated with lowest disease risk in meta-analyses (health outcome) | Observed mean national intakes in 2010 ² | Recommended intakes by major dietary guidelines ³ | Optimal mean population intake ⁴ |
|--|---|--|--|--|
| <i>Trans</i> -fats (% <i>E</i>) | • 0%E (CHD) | Bottom 3 countriles: • Barbados: 0.2%E • Finland: 0.4%E • Italy: 0.5%E | DGA 2015: As low as possible • WHO: <1%E[97] | 0.5%E ¹⁰ |
| Dietary fiber (g/d) | 23.8 g/d (CVD) 22.8 g/d (CHD) 19.4 g/d (stroke) 30.0 g/d (diabetes) | Top 3 countries: • Barbados: 28 g/d • Mexico: 26 g/d • Bulgaria: 25 g/d | DGA 2015: 22.4–33.6 g/d (depending on gender & age group) | 30 g/d |
| Sodium (mg/d) | 614 mg/d (lower SBP)¹¹ [60] 1,500 mg/d (reduced BP in randomized trials) [46] 1,787 mg/d (CHD mortality) [61] 2,245 mg/d (stroke mortality) [61] 2,391 mg/d (stroke) | 5 nations with mean intakes at or below 2,000 mg/d: Jamaica, Colombia, South Africa, Mexico, Iran | UK NICE: <1,200 mg/d [62] AHA 2020: <1,500 mg/d [63] WHO: 2,000 mg/d [64] DGA 2015: <2,300 mg/d UK FSA: <2,400 mg/d [65] | 2,000 mg/d ¹² |
| Potassium (mg/d) | 4,136 mg/d (stroke) | Not available | DGA 2015: 4,700 mg/d WHO: 23,500 mg/d [38] | 4,500 mg/d |

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coronary heart disease (CHD), stroke, type 2 diabetes, body mass index (BMI), and systolic blood pressure (SBP) [10]. For both studies of clinical outcomes and national surveys of Building up on our work in the 2010 Nutrition and Chronic Diseases Expert Group (NutriCoDE) in which we evaluated optimal intake levels for cardiometabolic outcomes including Jietary intakes, we used standardized servings to account for any variability in serving sizes across studies or countries. Thus, the characterized RR's are accurate for the listed serving size. For populations with smaller or larger serving sizes, the RR's should be appropriately adjusted.

Based on nationally representative, individual-level dietary surveys using optimal dietary metrics among both adult men and women in our 2010 NutriCoDE Global Dietary Database 27-31] and other sources [32]; adjusted to 2000 kcal/d.

³ For an average intake of 2,000 kcal/d.

consumption (SD) around the optimal population mean was determined to be ±10% of the mean, based on the average SD for diet-related metabolic risk factors [3, 36–39]. We could not comparably identify optimal intake levels of glycemic load due to absence of global data on mean intakes in most nations and of recommended levels in major dietary guidelines. easibility (observed national consumption levels globally), and consistency (with other assessments in major dietary guidelines) [33, 34]. The plausible population distribution of For an average energy intake of 2,000 kcal/d. The optimal mean levels for the population were determined based on risk (observed levels at which lowest disease risk occurs). Excluding 100% juices

Excluding vegetable juices, starchy vegetables such as potatoes or corn, and salted or pickled vegetables.

⁷ Including beans/legumes

Based on eicosapentaenoic acid (EPA) + docosahexaenoic acid (DHA) in common fish varieties, the DGA 2015 calculates that 1,750 mg/wk (250 mg/d) would be in concordance with the recommended fish intake of 8 oz/wk

³ 12%E in earlier analyses [3, 25, 29]. Lowered to 11% based on the present updated review of available evidence on optimal levels and considering observed national intakes and major dietary guidelines

¹⁰ Non-zero value to account for natural ruminant sources, for which probable or convincing evidence of causal effects on cardiometabolic outcomes was not identified.

¹¹ Based on ecologic evidence; these values are the mean 24-hour urine sodium excretion across the 4 populations with lowest levels (Brazil, Yanomano and Xingu, Papua New Guinea, Kenya) in the Intersalt Study [60]

effect of BP on cardiovascular disease (CVD), the absence of plausible biologic rationale for increased risk with sodium reduction, at least to 2,000 mg/d, and the plausible source of ²² As previously described in detail, we did not incorporate a potential U-shaped relationship with risk due to the linear dose-response effect of dietary sodium on BP, the log-linear bias that could explain the U-shaped relationships observed in some, but not all, prior observational studies [24]

AHA 2020, American Heart Association 2020 Strategic Impact Goals [35]; DGA, Dietary Guidelines for Americans [33]; FAO, United Nations Food and Agricultural Organization [34]; PUFA, Polyunsaturated fats; SFA, Saturated fats; UK FSA, Food Standards Agency; UK NICE, UK National Institute for Health and Clinical Excellence; WHO, World Health Organization

Table 5. Validity analyses comparing the observed relative risks for CHD based on evidence from prospective observational studies and randomized trials of dietary patterns versus the estimated relative risks for CHD based on the present analysis of individual dietary components.

| | Observed Relative Risk for CHD ¹ | Estimated Relative Risk for CHD ² |
|--|--|--|
| Prospective cohort studies evaluating associations of overall dietary patterns with i | ncident CHD ³ | · |
| Health Professionals Study—Prudent diet (average of all quintiles, with lowest quintile as the reference) [40] | 0.82 | 0.78 |
| Health Professionals Study—Western diet (average of all quintiles, with lowest quintile as the reference) [40] | 1.29 | 1.17 |
| Nurse's Health Study—Prudent diet (average of all quintiles, with lowest quintile as the reference) [41] | 0.82 | 0.80 |
| Nurse's Health Study—Western diet (average of all quintiles, with lowest quintile as the reference) [41] | 1.20 | 1.10 |
| Nurse's Health Study—Mediterranean diet (average of all quintiles, with lowest quintile as the reference) [42] | 0.84 | 0.81 |
| EPIC-Greek (per 2 units diet score increase) [43] | 0.78 | 0.90 |
| SUN-Spain (per 2 units diet score increase) [44] | 0.74 | 0.75 |
| Randomized controlled feeding trials evaluating effects of overall dietary patterns o | n BP and LDL-C 4 | |
| Fruits, serving/d (100 g/d) | 0.93 | 0.94 |
| Vegetables, serving/d (100 g/d) | 0.93 | 0.95 |
| Nuts and seeds, serving/wk (1 oz [28.35 g]/wk) | 0.93 | 0.93 |
| Whole grains, serving/d (50 g/d) | 0.88 | 0.97 |
| Fish, serving/d (100 g/d) | 0.87 | 0.66 |
| Red meat, serving/d (100 g/d) | 1.17 | 1.17 |
| Dietary fiber, 20 g/d | 0.77 | 0.76 |
| Randomized clinical trial evaluating effects of an overall dietary pattern on incident | CHD [<u>53, 99]</u> ⁵ | |
| Mediterranean diet + extra-virgin olive oil vs. placebo | 0.80 | 0.77 |
| Mediterranean diet + mixed nuts vs. placebo | 0.74 | 0.62 |
| Combined groups | 0.77 | 0.69 |

¹ Values are the observed relative risks (RRs) in these long-term prospective observational studies or randomized trials of dietary patterns.

² Values are the estimated RRs based on the reported differences in intakes of individual dietary factors across each category of the diet pattern study and our estimated quantitative effects for these individual dietary factors (Table 2), assuming a multiplicative relation of proportional effects of individual components. Not all dietary factors in Table 2 were included due to insufficient reporting of differences in these components across studies of dietary patterns. We focused on foods and excluded overlapping nutrients (e.g., we included whole grains, fruits, and vegetables; and excluded fiber, glycemic load). We also assumed no benefits from differences in other dietary factors (e.g., coffee) in the dietary pattern for which we had not determined a causal etiologic effect.

³ Because the observed relative risks in most of these cohorts were based on serial dietary measures with time-varying updating, the predicted relative risks for each dietary factor were adjusted for comparability for regression dilution bias due to the observed changes over time of each dietary factor in these cohorts. See Table D in <u>S1 File</u> for more details.

⁴ For randomized controlled feeding trials of dietary patterns and cardiovascular risk factors, we performed inverse-variance-weighted meta-regression across all of the treatment arms of three large, well-established dietary pattern trials [45–47] to estimate the independent effects of five different dietary components, when consumed as part of an overall dietary pattern, on systolic blood pressure (SBP) and LDL-cholesterol (LDL-C). We evaluated achieved dietary changes in fruits, vegetables, nuts, whole grains, and fish simultaneously as independent variables in the meta-regression, with changes in SBP or LDL-C as the dependent variable. For each dietary factor, we then calculated how the identified change in SBP and LDL-C from the meta-regression would alter cardiovascular risk, based on the established relationship between SBP and LDL-C and clinical events [48–52], assuming independent, multiplicative effects of SBP and LDL-C on risk. These observed effects, calculated based only on how each dietary factor altered SBP and LDL-C in randomized controlled feeding trials of diet patterns, were then compared to our estimated etiologic effect on coronary heart disease (CHD) events for that dietary factor (Table 2). See Table E in S1 File for more details.

⁵ We compared the observed vs. predicted risk in a large randomized clinical trial evaluating the effects of two overall dietary patterns on incidence of cardiovascular events [53]. A similar analysis was previously reported using 2010 NutriCode RR's [99]; the values here are based on the updated RR's in the current investigation (Table 2). The predicted risk reductions were calculated by combining the observed differences in individual dietary components achieved in the trial with our estimated quantitative effects, assuming multiplicative effects of each individual component. Because we had not identified sufficient studies to quantify etiologic effects of extra-virgin olive oil, to enable comparison we imputed potential effects of extra-virgin olive oil from our estimated relative risk for polyunsaturated vegetable fats. See Table F in S1 File for more details.

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Discussion

This systematic evaluation of the evidence for effects of dietary habits on CHD, stroke, and diabetes identified and quantified probable or convincing etiologic effects and optimal consumption levels for 10 foods and 7 nutrients. Generally, minimally processed, bioactive-rich foods like fruits, vegetables, nuts/seeds, beans/legumes, and whole grains had protective effects, whereas certain more highly processed foods such as processed meats and SSBs had harmful effects. Other identified protective dietary factors were characterized by relatively unique attributes, such as fish/seafood and long-chain omega-3s (linked to lower risk of fatal CHD) or yogurt (containing active probiotics; linked to lower risk of diabetes) [12]. Fewer etiologic relationships were identified for isolated nutrients, and these were generally consistent with the findings for foods: for example, lower risk with dietary fiber and higher risk with additives such as trans-fats and sodium. Notably, many identified findings were specific for particular cardiometabolic outcomes: e.g., low fruit intake was identified as an etiologic risk factor for cardiovascular diseases, but not diabetes; unprocessed red meat and low yogurt intakes, as etiologic risk factors for diabetes, but not CHD or stroke; and low long-chain omega-3 intake, as an etiologic risk factor for fatal CHD, but not nonfatal CHD, stroke, or diabetes. To our knowledge, these novel findings provide the most updated, comprehensive estimates of quantitative effects of specific dietary factors on cardiometabolic disease burdens.

Whereas our evaluation for causality was based on diverse types of evidence [10], our estimates of quantitative etiologic effects mostly relied on prospective observational studies. While such studies represent a reasonable study design for evaluating long-term effects of lifestyle (as compared with pharmaceutical drugs), the results could be biased by residual confounding, particularly from other correlated dietary habits. Yet, the results of separate validity analyses, each examining estimated effects of individual dietary components as compared to observational studies or RCTs of dietary patterns, suggested low likelihood of large magnitudes of bias in our quantified etiologic effects.

Interestingly, the majority of the identified causal factors represented food groups, rather than isolated nutrients. These results are consistent with advances in nutritional science that suggest a greater relevance of foods, rather than nutrient-based metrics, for risk of chronic diseases [12, 33, 66]. Exceptions included polyunsaturated fats, representing certain vegetables, nuts/seeds, and vegetable oils; long-chain omega-3 fats, representing seafood; and dietary fiber and potassium, representing intakes of whole foods such as fruits, vegetables, nuts/seeds, and whole grains. The other identified nutrients with evidence for etiologic effects were sodium and *trans*-fats—industrial additives that can be increased or decreased in any otherwise similar food—and glycemic load, representing higher intakes of refined starches and sugars. These new findings add to a growing evidence base that emphasizes the importance of food-based diet quality in general, and minimally processed, bioactive-rich foods in particular, as key priorities for reducing burdens of cardiometabolic diseases.

Our conservative approach did not identify sufficient accumulated evidence for probable or convincing causal effects on cardiometabolic endpoints of other promising dietary factors, e.g., plant omega-3 fats, coffee, tea, cocoa. The present findings represent an assessment of the current state of evidence, and undoubtedly continuing advances in science—e.g., better dietary assessment, biomarker measures, nutrigenomics, metabolomics, personalized nutrition, other technological advances—will lead to future identification and refinement of additional important etiologic dietary factors and mechanistic pathways, for instance including polyphenols, other trace bioactives, branched chain fatty acids, and the microbiome.

Our standardized assessment of feasible optimal intake levels, informed primarily by observed levels linked to lowest risk of clinical events, provide additional new evidence to

inform dietary guidelines, policy targets, and assessments of disease burdens. The identified optimal intake levels were generally similar to major dietary guidelines, supporting validity of our approach. These results do not imply assumptions about practicality or potential pace of achieving such optimal intakes for all countries, which will vary based on local cultural, economic, and political considerations. Yet, changes in national policies can induce brisk changes in dietary habits, cardiometabolic risk factors, and disease rates [67, 68], and systems approaches utilizing school, workplace, economic, built environment, and media/education strategies can effectively alter diets in populations [69, 70]. The present results on optimal intakes to quantify disease risk and inform policy priorities in different nations.

Using evidence published through 2007, Mente and colleagues reviewed evidence for etiologic effects of dietary factors on CHD; this work included only 4 Bradford-Hill criteria, compared with 9 in our investigation; and did not evaluate stroke or diabetes, evidence for optimal intake levels, or validity analysis to assess bias [8]. Others have reviewed, in narrative fashion [7], the published evidence on diet and cardiometabolic diseases, but without quantitative assessments of etiologic dose-responses, optimal consumption, or potential bias. This is the first study, to our knowledge, to systematically evaluate and quantify the current evidence for both etiologic effects and optimal levels of multiple dietary components for major cardiometabolic endpoints including CHD, stroke, and diabetes.

Our study has several strengths. We formally evaluated evidence for causality independently and in duplicate based on established Bradford-Hill criteria and assessed whether such evidence was at least probable or convincing. We quantified etiologic effects and optimal levels based on published or new meta-analyses of available evidence; including determination of dose-responses per standardized serving sizes where possible. We evaluated heterogeneity in etiologic effects by underlying individual characteristics such as age and sex. Importantly, the potential for over- or underestimation of identified effects was assessed in separate validity analyses incorporating data from long-term cohorts and randomized trials of dietary patterns.

Potential limitations should be considered. Dietary assessment in prospective cohort studies can be imperfect due to incomplete memory, questionnaire limitations, and changes in dietary habits over time; each of these factors would generally attenuate risk estimates, causing underestimation of etiologic effects. Conversely, some of the individual studies in these meta-analyses utilized serial measurements of diet, which would tend to reduce such misclassification. We did not identify sufficiently reliable data on temporal dietary changes to correct for regression dilution bias over time; our investigation identifies a need to generate such evidence across multiple cohorts and world regions. Outcome ascertainment (e.g. of CHD, diabetes) varied across studies and would be prone to error; because we focused on prospective studies, such misclassification would most often be random with respect to exposure and lead to smaller magnitudes of etiologic effects. We did not assess study quality, country, or year of publication for individual reports within each meta-analysis; we cannot exclude that differences in these factors might influence findings. Our validity analyses represented qualitative comparisons, not formal statistical tests. We limited our final estimates to dietary factors with strongest evidence, excluding many other dietary components which may influence cardiometabolic health. We did not grade strength of evidence for absence of health effects; a recent narrative review included some qualitative conclusions on this [12].

In sum, our novel findings provide a quantitative summary of the current evidence for causality, etiologic effects, and optimal intakes of individual dietary factors in relation to cardiometabolic disease. These findings facilitate assessment of diet-related disease burdens, investigation of comparative effectiveness and cost-effectiveness of individual and policy-level dietary interventions, and design of program priorities and prevention strategies to reduce diet-related cardiometabolic diseases.

Supporting information

S1 File. Grading the Evidence for Causality

Text A. Criteria for grading the evidence for etiologic effects of specific dietary factors on cardiometabolic outcomes.

Literature Searches for Published Meta-analyses

Text B. Searches for identifying meta-analyses of the effect of specified dietary risk factors on cardiometabolic diseases.

Table A. Search results, per each search strategy based on types of articles.

Figure A. Screening and selection process of meta-analyses evaluating etiologic effects of dietdisease relationships for dietary factors with probable or convincing evidence for effects on cardiometabolic diseases.

De Novo Meta-Analyses of Fruit and Vegetable Intake and Incident Stroke

Text C. Protocol for de novo meta-analyses of fruit and vegetable intake and incident stroke. **Text D.** Search terms used to identify published prospective cohort studies examining the fruit/vegetable and stroke relationship that were published after previous fruit and vegetable meta-analyses.

Text E. Search results of published prospective cohort studies examining the fruit/vegetable and stroke relationship.

Table B. Summary results of included cohort studies in de novo meta-analysis on fruit and vegetable intake and ischemic stroke.

Table C. Summary results of included cohort studies in de novo meta-analysis on fruit and vegetable intake and hemorrhagic stroke.

Etiologic Effects of Dietary Factors on Cardiometabolic Disease Risk

Text F. Heterogeneity in etiologic effects.

Figure B. Age-specific relative risks for fruit intake and coronary heart disease risk. **Validity Analyses**

Table D. Comparison of relative risks for CHD observed in prospective cohort studies of die-
tary patterns and estimated based on NutriCoDE relative risks for individual dietary factors.**Table E.** Comparison of relative risks for CHD calculated based on changes in systolic blood
pressure and LDL-cholesterol in randomized controlled feeding trials of dietary patterns vs.estimated relative risks based on NutriCoDE relative risks for individual dietary factors.**Table F.** Comparison of relative risks for CHD observed in a large randomized clinical trial of
dietary patterns vs. estimated relative risks based on NutriCoDE relative risks for individual

dietary factor.

(DOCX)

S2 File. PRISMA checklist. (DOC)

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