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# Original article

# Predicting critical illness on initial diagnosis of COVID-19 based on easily obtained clinical variables: development and validation of the PRIORITY model

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#### ABSTRACT

*Objectives*: We aimed to develop and validate a prediction model, based on clinical history and examination findings on initial diagnosis of coronavirus disease 2019 (COVID-19), to identify patients at risk of critical outcomes.

*Methods:* We used data from the SEMI-COVID-19 Registry, a cohort of consecutive patients hospitalized for COVID-19 from 132 centres in Spain (23rd March to 21st May 2020). For the development cohort, tertiary referral hospitals were selected, while the validation cohort included smaller hospitals. The primary outcome was a composite of in-hospital death, mechanical ventilation, or admission to intensive

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Keywords: COVID-19 Critical illness Easily obtained clinical variables Initial assessment Prognostic models care unit. Clinical signs and symptoms, demographics, and medical history ascertained at presentation were screened using least absolute shrinkage and selection operator, and logistic regression was used to construct the predictive model.

Results: There were 10 433 patients, 7850 in the development cohort (primary outcome 25.1%, 1967/7850) and 2583 in the validation cohort (outcome 27.0%, 698/2583). The PRIORITY model included: age, dependency, cardiovascular disease, chronic kidney disease, dyspnoea, tachypnoea, confusion, systolic blood pressure, and  $SpO_2 \leq 93\%$  or oxygen requirement. The model showed high discrimination for critical illness in both the development (C-statistic 0.823; 95% confidence interval (CI) 0.813, 0.834) and validation (C-statistic 0.794; 95%CI 0.775, 0.813) cohorts. A freely available web-based calculator was developed based on this model (https://www.evidencio.com/models/show/2344).

Conclusions: The PRIORITY model, based on easily obtained clinical information, had good discrimination and generalizability for identifying COVID-19 patients at risk of critical outcomes. **Miguel Martínez-Lacalzada**, **Clin Microbiol Infect 2021;27:1838** 

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#### Introduction

Coronavirus disease 2019 (COVID-19) has spread globally, with a clinical spectrum ranging from an asymptomatic state to critical illness [1–3]. Notably, Spain was one of the countries with the highest incidence of COVID-19 during the first pandemic peak [4]. To optimize the use of limited healthcare resources, it is essential to identify, as early as possible, those patients who are at high risk of progressing to critical illness.

To date, studies of COVID-19 prognostic factors have focused on laboratory measurements and radiological examinations obtained following admission [5–15], which are not available in outpatient or resource-limited settings. Recently published well-developed models tend not to include clinical variables obtained from history and examination carried out on initial assessment [9–13]. Where one machine-learning model has addressed basic clinical features, it has narrowed down the prediction to the mortality outcome only and lacks wider generalizability [16]. Furthermore, a critical appraisal of the COVID-19 models has shown poor reporting and high risk of bias [14].

Prediction models based on easy-to-collect data have previously been developed for other infectious diseases, e.g. meningitis and pneumonia [17–19]. As a global health emergency, management of COVID-19 would benefit from a prediction model that could be readily applied for initial diagnosis. Therefore, we developed and externally validated a prediction model—based on easily obtainable clinical measures at presentation with confirmed COVID-19 diagnosis—to identify patients at risk of developing critical outcomes.

# Methods

Study design and data source

This study was based on the SEMI (Sociedad Española de Medicina Interna) COVID-19 Registry [20]. It is an ongoing multicentre nationwide cohort of consecutive patients hospitalized for COVID-19 across Spain. Eligibility criteria were age ≥18 years, confirmed diagnosis of COVID-19 (defined as a positive result on real-time reverse-transcription-PCR (RT-PCR) for the presence of severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2) in nasopharyngeal swab specimens or sputum samples), first hospital admission for COVID-19, and hospital discharge or in-hospital death [20]. The SEMI-COVID-19 Registry was approved by the Provincial Research Ethics Committee of Málaga (Spain) and the

Institutional Research Ethics Committees of each participating hospital.

For the study, we retrieved from the Registry clinical baseline data, history of previous medication, and comorbidities collected on admission, as well as complications during hospitalization and status at discharge. We used data from patients admitted to 132 hospitals between 23rd March and 21st May 2020. We chose hospital complexity as the criterion to assess the transportability of the prognostic model in a setting other than the one in which it was derived [21,22]. Patients admitted to tertiary referral hospitals (≥300 beds, according to the Ministry of Health of Spain [23]) were selected for the development cohort, while patients from smaller hospitals (<300 beds) were included in a separate validation cohort.

#### Outcome description

The primary outcome, critical illness during hospitalization, was defined as the composite of in-hospital death, mechanical ventilation or admission to the intensive care unit (ICU), according to previously published studies [10,24,25].

# Potential predictors

To develop a predictive model based only on easily measurable variables registered at admission, we considered clinical signs and symptoms, demographic variables, and medical history. An initial list of 29 candidate variables was selected based on review of the existing evidence [5–16], clinical plausibility, and relevance to clinical care. To improve consensus on model applicability, a oneround online questionnaire was conducted among a multidisciplinary panel of 24 physicians involved in the clinical management of COVID-19 at nursing homes, emergency departments, primary care centres and hospitalization wards (six per each setting). The panellists were asked to rate (on a nine-point Likert scale) the availability/reliability of each predictor, its ability to predict the outcome, the best way to merge predictors of rare occurrence, and the maximum number of variables the model should contain. Agreement was considered when seven or fewer panellists rated outside the three-point region containing the median [26].

# Statistical analysis

The predictive model, called PRIORITY, was presented as the formula for estimating the probability of COVID-19 critical illness

outcome, as well as an associated web-based calculator. Patients' characteristics were summarized as frequencies and percentages or means and standard deviations. Statistical analysis was performed using R v4.0.0, with mice, mfp, glmnet, pROC, rms and rmda packages.

#### Model development

Missing values in the potential predictors were imputed using single imputation, a reasonable alternative to multiple imputation when dealing with relatively few missing values [27]. A stochastic single imputation dataset was created for both cohorts (development and validation) through multiple imputation by chained equations. Quantitative variables were kept as continuous to avoid loss of prognostic information, and non-linear relationships were modelled by multivariate fractional polynomials with a maximum of two degrees of freedom [28]. The least absolute shrinkage and selection operator (LASSO) was the feature selection method used to reduce the number of predictors down to the maximum agreed by the expert panel [29]. Briefly, the potential predictors were entered into the LASSO regularization process, which penalizes the coefficients by gradually shrinking them to zero. We selected the penalty parameter  $(\lambda)$  that minimized the deviance within the given maximum number of predictors. Those variables with nonzero coefficients were retained for risk estimation using a logistic regression model. Coefficients were presented as odds ratios (ORs) and 95% confidence intervals (95%CIs).

#### Model performance

Nagelkerke's R<sup>2</sup> and Brier score were used as overall performance measures. We assessed the discriminative ability of the model using the C-statistic, calculated as the area under the receiver operating characteristic curve, with 95%CI. Calibration of the model was visually assessed by plotting deciles of predicted versus observed probabilities, and the calibration slope with 95%CI was calculated [22].

#### Model validation

Internal validation was carried out to assess optimism-corrected performance by repeating the entire model development over 1000 bootstrap samples drawn from the development cohort [27]. We externally validated the model in a separate cohort of patients admitted at smaller hospitals to evaluate model transportability [21]. Within this validation cohort, we reassessed model performance and compared its discrimination ability with models based on oxygen saturation and/or age, the most discriminating univariate predictors for in-hospital mortality previously reported [15]. We also undertook a decision curve analysis, a method to ascertain the adequacy of prediction models based on the relative value of benefits (true positives) and harms (false positives) [30]. We plotted the net benefit of the models for the full range of critical illness probability thresholds.

#### Sensitivity analysis

To assess the impact of assumptions adopted in the model development, we carried out a complete-case analysis using only those patients with complete data in the potential predictors. We also developed models without restricting the maximum number of predictors ( $\lambda$  at one-standard-error of the minimum) or using linear continuous predictors instead of non-linear terms.

#### Results

We considered data from 10 433 patients included in the SEMI-COVID-19 Registry. The development cohort included 7850 patients, of whom 1967 (25.1%) presented critical illness: 650 (8.3%) were admitted to the ICU and 1598 (20.4%) died. The mean age

was 65.8  $\pm$  16.4 years and 57.2% (4483/7834) were male. Demographics and clinical characteristics for the development cohort are shown on Table 1.

#### Model development and performance

From an initial list of 29 candidate variables, the expert panel forged an agreement on 21 potential predictors for further evaluation in the predictive model. Chronic liver disease, previous medication with angiotensin converting enzyme inhibitors or angiotensin receptor blockers, cough, arthralgia/myalgia, ageusia/ anosmia, asthenia/anorexia, headache, gastrointestinal symptoms were excluded. Consensus was achieved for including a range between five and nine variables in the final model. For transparency, univariate analysis is shown in Supplementary Material Table S1, even though it was not part of the model development process. The 21 potential predictors were included in the LASSO selection process, retaining a subset of nine variables as the best predictors of critical illness (Supplementary Material Fig. S1). A multivariable logistic regression model was then fitted with these nine variables. All of them, except for moderate or severe dependency, were statistically significant (Table 2).

Based on the logistic regression model, the probability of critical COVID-19 illness could be calculated as:

Probability (%) =  $100/(1+\exp(-z))$ , where  $z = -4.665 + 2.663 \cdot ((age/100)^2) + 0.164 \cdot (dependency) + 0.316 \cdot (cardiovascular disease) + <math>0.586 \cdot (chronic kidney disease) + 0.504 \cdot (dyspnoea) + 0.844 \cdot (1/(SBP/100)^2) + 0.911 \cdot (tachypnoea) + 1.200 \cdot (SpO_2 < 93\% or oxygen requirement) + 0.681 \cdot (confusion).$ 

All predictors were coded as binary variables (1 when present and 0 when absent) except for age (years) and systolic blood pressure (SBP, mmHg). We also developed an online calculator based on this model (Supplementary Material Fig. S2), accessible at https://www.evidencio.com/models/show/2344.

In the development cohort, the PRIORITY model had an R<sup>2</sup> of 0.347 and a Brier score of 0.138. The apparent C-statistic was 0.823 (95%CI 0.813, 0.834) (Fig. 1a). After bootstrap internal validation, optimism-corrected C-statistic was 0.821 (95%CI 0.810, 0.832). The model showed good calibration across the range of predicted probabilities within the development cohort (calibration slope 0.996, 95%CI 0.989, 0.999; Supplementary Material Fig. S3a).

#### External validation

The validation cohort included 2583 patients, of whom 698 (27.0%) presented critical illness: 200 (7.7%) were admitted to the ICU and 594 (23.0%) died. The mean age was  $69.5 \pm 16.0$  years, 54.8% (1415/2580) were male (Table 1). The PRIORITY model showed good discrimination for critical illness within the validation cohort (C-statistic 0.794, 95%CI 0.775, 0.813) (Fig. 1b), and a calibration slope of 0.875, 95%CI 0.790, 0.960 (Supplementary Material Fig. S3b).

Our model compared well against the risk stratification based on univariate models including age (C-statistic 0.707, 95%CI 0.686, 0.729) or SpO $_2 \le 93\%$ /oxygen requirement at admission (C-statistic 0.652, 95%CI 0.635, 0.670) as sole predictors. Likewise, the PRIORITY model had better discriminative ability than the model including both age and SpO $_2 \le 93\%$ /oxygen supply (C-statistic 0.751, 95%CI 0.731, 0.771). Additionally, decision curve analysis showed that the PRIORITY model had higher net benefit across a wide range of threshold probabilities for developing critical illness compared to risk stratification using age and/or SpO $_2 \le 93\%$ /oxygen supply (Fig. 2).

**Table 1**Demographic and clinical characteristics among patients included in the development and validation cohorts

		Development cohort		Validation cohort	
		No of patients (%) or mean ± SD	Total No (%)	No of patients (%) or mean ± SD	Total No (%)
Characteristics of the	population				
Critical illness	•	1967 (25.1%)	7850 (100%)	698 (27.0%)	2583 (100%)
Age [years]		$65.8 \pm 16.4$	7816 (99.6%)	$69.5 \pm 16.0$	2575 (97.3%)
Male		4483 (57.2%)	7834 (99.8%)	1415 (54.8%)	2580 (99.9%)
Ethnicity	Caucasian	6836 (89.1%)	7677 (98.8%)	2340 (91.0%)	2572 (99.6%)
	Latino	693 (9.0%)	, ,	193 (7.5%)	, ,
	Other	148 (1.9%)		39 (1.5%)	
Smoking history	Never	5270 (70.9%)	7433 (94.7%)	1625 (65.7%)	2475 (95.8%)
	Former smoker	1764 (23.7%)	, ,	718 (29.0%)	, ,
	Active Smoker	399 (5.4%)		139 (5.3%)	
Medical history		, ,		, ,	
Obesity <sup>a</sup>		1665 (23.7%)	7012 (89.3%)	584 (24.3%)	2401 (93.0%)
Hypertension		3803 (48.6%)	7833 (99.8%)	1444 (56.1%)	2576 (99.7%)
Diabetes mellitus		1440 (18.4%)	7820 (99.6%)	509 (19.8%)	2570 (99.5%)
Cardiovascular disease <sup>b</sup>		1974 (25.3%)	7800 (99.4%)	806 (31.7%)	2545 (98.5%)
Pulmonary diseases <sup>c</sup>		1625 (20.9%)	7776 (99.1%)	576 (22.6%)	2583 (98.9%)
Severe chronic kidney disease <sup>d</sup>		488 (6.2%)	7825 (99.7%)	163 (6.3%)	2583 (99.7%)
Malignancy <sup>e</sup>		793 (10.2%)	7803 (99.4%)	259 (10.1%)	2571 (99.5%)
Immunocompromised status <sup>f</sup>		650 (8.6%)	7549 (96.2%)	187 (7.6%)	2473 (95.7%)
Dependency (moderate/severe) <sup>g</sup>		1129 (14.7%)	7701 (98.1%)	605 (23.7%)	2555 (98.9%)
Symptoms at admission	on .				
Fever <sup>h</sup>		5138 (67.0%)	7663 (97.6%)	1670 (65.6%)	2544 (98.5%)
Dyspnoea		4427 (56.7%)	7805 (99.4%)	1523 (59.4%)	2562 (99.2%)
Clinical signs and phys	sical exploration at admission	, ,	, ,	• •	, ,
SBP (mmHg)		$129.0 \pm 21.5$	7430 (94.6%)	$127.6 \pm 21.0$	2451 (94.9%)
HR (beats/minute)		$88.6 \pm 17.4$	7500 (95.5%)	87.5 ± 17.5	2504 (96.9%)
Tachypnoea (>20 breaths/min)		2271 (29.9%)	7604 (96.9%)	879 (35.1%)	2504 (96.9%)
SpO <sub>2</sub> ≤93% or oxygen requirement at presentation		4152 (52.9%)	7842 (99.9%)	1605 (62.1%)	2583 (100%)
Pulmonary rales		4630 (60.7%)	7626 (97.1%)	1588 (63.6%)	2495 (96.6%)
Confusion		849 (11.0%)	7736 (98.5%)	384 (15.1%)	2546 (98.6%)

SD, standard deviation; HR, heart rate; SBP, systolic blood pressure; SpO<sub>2</sub>, peripheral oxygen saturation.

- <sup>a</sup> Obesity is defined as medical history or body mass index  $\geq 30 \text{ kg/m}^2$ .
- b History of cerebrovascular disease, peripheral arterial disease, myocardial infarction, angina pectoris, heart failure or atrial fibrillation.
- <sup>c</sup> Chronic obstructive pulmonary disease, obstructive sleep apnoea/hypopnoea syndrome and asthma.
- d History of serum creatinine level >3 mg/dL or history of dialysis.
- e History of solid tumour, leukaemia or lymphoma.
- f History of autoimmune diseases, solid-organ transplant recipients, HIV infection or previous immunosuppressive treatment including systemic steroids.
- g Moderate or severe dependency for activities of daily living (Barthel index score <60).
- <sup>h</sup> Temperature  $\ge$ 38°C or history of fever.

### Sensitivity analysis

We carried out a complete-case analysis selecting as development cohort the 5513 patients with complete data on the 21 potential predictors and the outcome. The resulting model had the same predictors as the PRIORITY model with apparent C-statistic 0.813 (95%CI 0.800, 0.826) and calibration slope 0.993 (95%CI 0.986, 0.997). Next, we fitted a new model with no restriction in the maximum number of variables, resulting in a model which added sex, diabetes mellitus, malignancy, immunocompromised status, pulmonary rales, and heart rate cubed to the predictors in the PRIORITY model. C-statistic was 0.831 (95%CI 0.821, 0.842) and slope 0.990 (95%CI 0.986, 0.996). Likewise, we fitted an alternative model using linear continuous predictors instead of non-linear terms, which included sex but excluded the systolic blood pressure. C-statistic was 0.823 (95%CI 0.812, 0.833) and slope 0.994 (95%CI 0.988, 0.999).

#### Discussion

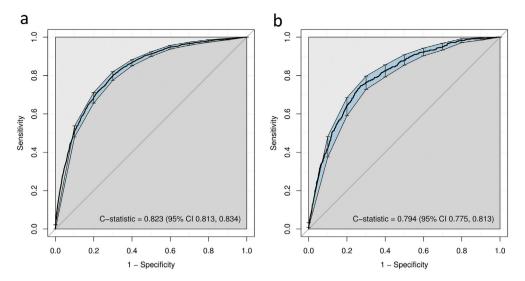
We developed and validated a new clinical risk model to predict COVID-19 critical illness based on nine simple clinical features easily available on initial assessment, which would be useful in resource-limited or out-of-hospital settings without access to other complementary tests. The model was well calibrated, had good discrimination, and performed robustly in an external validation cohort. Moreover, it showed a potential clinical benefit in a variety

**Table 2**Multivariate logistic regression of critical illness prediction in coronavirus 2019 (COVID-19)

Predictors	Odds ratio	95%CI
(Age/100) <sup>2</sup> Age in years <sup>a</sup>	14.339	10.054, 20.532
Cardiovascular disease	1.372	1.195, 1.573
Severe chronic kidney disease	1.797	1.433, 2.252
Dyspnoea	1.655	1.451, 1.891
1/(SBP/100) <sup>2</sup> SBP in mmHg <sup>a</sup>	2.326	1.837, 2.951
Tachypnoea (>20 breaths/min)	2.487	2.192, 2.824
SpO <sub>2</sub> ≤93% or oxygen requirement	3.320	2.889, 3.819
Confusion	1.976	1.642, 2.380
Dependency (moderate or severe)	1.178	0.989, 1.404

Predictors in the PRIORITY model retained after LASSO feature selection. Model coefficients were derived from a multivariate logistic regression, and presented as odds ratios (ORs) and 95% confidence intervals (95%Cls). Variables entered into the LASSO feature selection process were: age as a squared term, sex, ethnicity, smoking history, obesity, hypertension, diabetes mellitus, cardiovascular disease, pulmonary diseases, severe chronic kidney disease, malignancy, immunocompromised status, dependency, fever, dyspnoea, systolic blood pressure (SBP) as the inverse of a quadratic term, heart rate (HR) as a cubic term, tachypnoea, peripheral oxygen saturation (SpO<sub>2</sub>)  $\leq$ 93% on room air or oxygen requirement at presentation, pulmonary rales, and confusion. All predictors were coded as binary variables (1 when present and 0 when absent) except for age (years), SBP (mmHg) and HR (bpm).

<sup>a</sup> Continuous predictors modelled as fractional polynomial terms, including rescaling when the range of values of the predictor was reasonably large. As interpretability of the effect of non-linear continuous predictors can be difficult, linear local approximations of ORs for 10-unit variations are provided at selected values. ORs for age (10-year increments): OR (50/40) = 1.271; OR (70/60) = 1.414; OR (90/80) = 1.573. ORs for SBP (10-mmHg decreases): OR (110/120) = 1.118; OR (90/100) = 1.219; OR (70/80) = 1.497. Approximated ORs are provided for illustrative purposes only and were not used for making predictions.



**Fig. 1.** Discriminatory ability of the PRIORITY model in (a) the development and (b) the validation cohorts. Discriminative ability was assessed using the C-statistic, as the area under the receiver operating characteristic curve, with 95% confidence intervals (CIs) computed with 1000 bootstrap replicates.

of scenarios covering different healthcare situations over a range of threshold probabilities for critical illness. The web-based calculator can facilitate its immediate application for frontline clinicians.

Previously, an external validation of 22 prognostic models showed that none of the multivariate models offered incremental value for patient stratification compared to oxygen saturation or age [15]. In this regard, the PRIORITY model showed higher discriminative ability and net benefit than age and/or oxygen saturation. Additionally, despite its simplicity, our model had a performance similar to previously published prognostic tools including laboratory and imaging tests [9–16].

It is worth noting that the PRIORITY model could be applied in triage, using easily measurable variables available in settings without access to laboratory or radiology tests, identifying highrisk patients for referral to hospital. This model could be useful

for supporting clinical management decisions over a range of risk thresholds for critical illness which could be considered as relevant in clinical practice. The choice of thresholds will vary across different regions, according to changing epidemiological situations and availability of health resources. For example, under pandemic peak pressure or low-resource healthcare systems, policy-makers may consider a cut-off point up to 20%, a threshold that will be associated with higher reduction in unnecessary critical care admissions. However, at low risk of overwhelming the critical care capacity, a lower threshold may be considered at the expense of unnecessary referrals. We recommend objectively defining specific cut-off points considering the circumstances and the availability of health resources.

This study has several methodological strengths maximizing internal and external validity [23]. To the best of our knowledge,

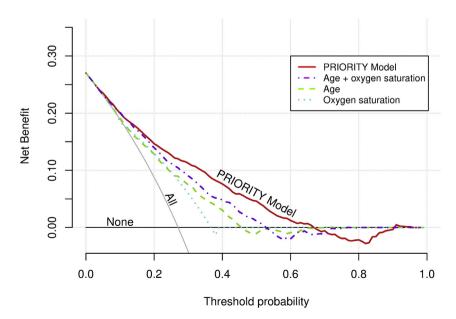


Fig. 2. Decision curve analysis within the validation cohort. Clinical usefulness of the PRIORITY model compared to risk stratification based on oxygen saturation (binary:  $SPO_2 \le 93\%$  or oxygen requirement) and/or age (quadratic term). The x-axis represents the whole range of decision threshold probabilities for critical illness  $(p_t)$  and the y-axis represents the net benefit (NB). NB calculated as: True positives/N - (false positives/N)\* $(p_t/(1-p_t))$ , with N total sample size.

this is the first generalizable COVID-19 predictive model built with simple clinical information (excluding imaging and laboratory data). We developed and validated the model in a large multicentre national cohort. The methodology was rigorous, avoiding data-driven predictor selection and biases that have affected previous studies [14]. The practical application of the model was maximized by forging an agreement among an expert panel on key issues. Moreover, the model was validated in a separate cohort of patients admitted in smaller hospitals, showing transportability to a setting with a different level of healthcare [21,22].

The strengths of our findings should be interpreted in light of some limitations. First, although we carefully selected easily available clinical and demographic variables, the data were collected at the time of hospital admission, which represents an important selection bias that would require further studies in an outpatient setting. Second, it could be suggested that, taking into account the situation of healthcare pressure, data quality may be affected. In this regard, it is notable that in this study there were relatively few missing data and we used imputation to reduce their impact. Third, since the COVID-19 pandemic has demonstrated significant differences between countries and time periods, it could affect the applicability of the model to other settings. However, we considered this early pandemic period in Spain to reflect a scenario with an overwhelmed healthcare system, where our predictive model could be particularly useful. Nevertheless, further studies introducing factors such as viral strains, healthcare system actions, new treatments, or vaccination could improve the applicability of the PRIORITY model. Finally, even though we compared the net benefit of using the model with discrimination based on oxygen saturation and/or age, its real clinical usefulness would require comparison with the best existing scores or the clinician's decision.

In summary, we have developed and validated a new prediction model, called PRIORITY, to estimate the risk of critical illness in patients with COVID-19 based on nine clinical variables easily measurable in resource-limited or out-of-hospital settings. The study could provide underpinning evidence to inform decision-making in health systems under pandemic pressure.

#### **Author contributions**

MML, AVN and MF planned and conceived the study and analysed and interpreted the data. MML, AVN, MF and LM wrote the original draft of the manuscript. MRR, SLG, FAF, JLBP, JAVN, ECM, ACEA, SJFC, JLA, PMPF, AP, AMAS, ASA, BGL, JLP, JSC, PCP, GMGG, JMNC, JMCR and RGH contributed to reading and approving the final version of the manuscript. MML and AVN are joint first authors. The corresponding author attests that all listed authors meet authorship criteria and that no others meeting the criteria have been omitted. LM is the guarantor.

## Transparency declaration

The corresponding author affirms that the manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned have been explained. The authors have no conflicts of interest to declare. No funding was received for this work.

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# Dissemination to participants and related patient and public communities

Results of this study have been made available to the public through an open access preprint posted to MedRxiv (https://doi.org/10.1101/2020.11.27.20237966). The Spanish Society of Internal Medicine (SEMI) shares the results of the studies derived from the SEMI-COVID-19 Registry through its public-facing website (https://www.fesemi.org/investigacion/proyectos/registro-semi-covid-19) and its twitter account (@Sociedad\_SEMI).

#### Appendix A. Supplementary data

Supplementary data to this article can be found online at https://doi.org/10.1016/j.cmi.2021.07.006.

# **Data sharing**

The data that support the findings of this study are available on request from the SEMI-COVID-19 Scientific Committee and the Registry Coordinating Center.

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