ORIGINAL RESEARCH

New Factor Structure of the Tampa Scale for Kinesiophobia in Older Japanese Adults After Lumbar Surgery

This article was published in the following Dove Press journal: Journal of Pain Research

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Purpose: The Tampa Scale for Kinesiophobia (TSK) has been used worldwide as a measure of kinesiophobia, but its factor structure in older Japanese adults after lumbar surgery is unknown. The purpose of this study was to fill this research gap by identifying the factors that comprise TSK in older Japanese adults after lumbar surgery.

Patients and Methods: Participants were older Japanese adults who had undergone surgery for lumbar spinal stenosis. Clinicodemographic data, TSK, intensity of low back pain and leg pain, dysesthesia (using an 11-point numerical rating scale), and HRQOL (using the EQ-5D-5L) were collected. After supplementing the missing values by the multiple assignment method, the hypothetical model of TSK was developed by categorical exploratory factor analysis (weighted least squares method, promax rotation). Confirmatory factor analysis (WLSMV method, promax rotation) was used to compare the hypothetical model and the traditional one-factor and two-factor models. Furthermore, we confirmed the relationship between factors extracted from the hypothetical model and HRQOL, pain, and dysesthesia.

Results: Questionnaires were mailed to 302 individuals, and responses were obtained from 211 (72.4 \pm 4.2 years [range: 65–88]; 115 men and 96 women; 804 \pm 343.1 [380–1531] days after surgery; 137 who had undergone decompression and fixation surgery, 74 who had undergone decompression surgery) (response rate: 69.9%). The hypothesized model consisted of "somatic focus," "activity avoidance," and "efficacy of physical activities," all of which were highly consistent. The fit of the hypothetical model was slightly inferior to that of the traditional two-factor model, but the hypothetical model met the criteria for fit. Somatic focus in the hypothetical model was slightly associated with HRQOL, pain, and dysesthesia.

Conclusion: In older Japanese adults after lumbar surgery, the goodness of fit of the TSK model was maintained by adding efficacy of physical activities as a third factor to the traditional two factors.

Keywords: kinesiophobia, postoperative pain, confirmatory factor analysis, older Japanese adults

Introduction

Pain is known to be a symptom strongly affecting adults' cognition and emotions and consequently has a notable impact on physical activity. In Japan, where the aging population is the highest in the world, it is especially important to understand the relationship between pain and physical activity in older adults. In a longitudinal study among older Japanese patients with chronic pain, shorter durations of physical activity were found to be associated with a higher risk of functional disability.¹ It has also been reported that low physical activity is associated with poor subjective health in older

Journal of Pain Research 2021:14 601-612

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601

Japanese women with chronic pain.² However, although it has been reported that it is the perception of pain,³ rather than pain itself, that determines physical activity, there has been insufficient evidence on the perception of pain in the older Japanese population.

The fear-avoidance model was proposed by Lethem et al in 1983⁴ as a framework to explain the relationship among pain, pain perception (fear), and physical activity. The model poses that when the experience of pain causes fear of pain, adults avoid physical activity. In fact, it has been confirmed that older adults with chronic pain are more inactive and have poorer subjective health than those without chronic pain, and that physical activity in older adults with chronic pain was associated with a fear of movement, also called kinesiophobia.^{3,5} It has further been reported that fear of movement is significantly associated with health-related quality of life (HRQOL) in patients diagnosed with musculoskeletal pain.⁶ Thus, fear of movement is one of the important pain perceptions associated with both pain and physical activity.

Questionnaires such as the Fear-Avoidance Beliefs Questionnaire (FABQ) and the Tampa Scale for Kinesiophobia (TSK) have been developed as measures to assess kinesiophobia.⁷ The FABQ is a scale specific to low back pain, whereas the TSK can be applied to pain in general. For this reason, the TSK is more flexible for clarifying future kinesiophobia toward various pain conditions in older Japanese. A questionnaire related to the TSK is the pain catastrophizing scale,⁸ which assesses irrational thoughts about pain, but the TSK is more likely to be able to assess physical activity more directly, as the model suggests that catastrophizing affects disability through the mediation of fear of movement.⁹

The factor structure, which is one aspect of psychometric properties, of the TSK was originally assumed to comprise one factor, but two-factor¹⁰ and four-factor^{11,12} structures have now been proposed. Multiple structural models have been compared simultaneously in people with chronic neck or back pain^{13,14} and osteoarthritis,¹⁵ and the two-factor structure has been found to have the best model fit. However, the difference in fit between the one- and two-factor structures is said to be negligible,¹⁶ and no firm conclusions have been drawn about the factor structure of the TSK. As such, it has been pointed out that the TSK's construct validity has not yet been fully tested.⁷ In addition, it has been argued that there are cultural differences between Asian and Western populations with regard to kinesiophobia,¹⁷ and a fuller examination of the Japanese population is needed to

better understand the factor structure of the TSK. Clarifying the factor structure of TSK in specific populations of Japanese people will help us understand the process by which pain experiences specific to those populations lead to the avoidance of physical activity and further declines in HRQOL, following the so-called fear-avoidance model, which in turn will lead to the practice of effective interventions.

Lumbar spinal canal stenosis (LSS) is one of the most common pain-causing diseases in the older Japanese population, with an estimated prevalence of 5.7% and 3.65 million patients between the ages of 40 and 79.18 If conservative treatment fails, and the pain and dysesthesia are judged to have a significant impact on daily life, surgery is indicated; however, pain and dysesthesia may be present even after surgery.¹⁹ Thus, it has been hypothesized that the experience of persistent chronic pain and dysesthesia after surgery may induce kinesiophobia, leading to inactivity and low HRQOL. The use of the TSK has already been analyzed with Western people with LSS; its relationship with pain and HRQOL²⁰ and the effect of participation in rehabilitation programs on reducing kinesiophobia²¹ have also been reported, but there is no report on the use of TSK in older patients with postoperative LSS, and so the nature of their kinesiophobia is not well understood. Moreover, the TSK has already been used in Japanese patients with diseases other than LSS such as knee osteoarthritis,¹⁵ anterior cruciate ligament injuries,¹⁷ and neck and back injuries due to traffic trauma,²² but, in contrast to when it has been used with a Western population, the factor structure of the TSK in Japanese has not been reported. Therefore, clarification of psychometric structures as well as detailed information on the TSK scores in older Japanese adults with LSS after surgery is expected to lead to the development of research on kinesiophobia in older Japanese adults with LSS after surgery.

To fill these research gaps in characteristics of the psychometric structure of the TSK in older Japanese adults with LSS after surgery, we aimed to identify the factors comprising the TSK in older Japanese adults who had undergone surgery for LSS at least 1 year prior, considering that the majority of people who had undergone instrumentation surgery had bone union by 1 year after surgery.²³

Patients and Methods

Study Design and Ethical Considerations

This was an observational study conducted between October 2019 and February 2020 in accordance with the Declaration of Helsinki. The ethical review committees of Sapporo Maruyama Orthopedic Hospital and Harunaso Hospital reviewed the research, and approval for the study was obtained from each institution in advance (approval numbers: 000025 and 190,105, respectively). A research description, a consent form, and a withdrawalof-consent form were sent to the subjects along with a questionnaire. We asked respondents to write their names on the questionnaire so that we could combine their responses with their medical records in the hospitals. The participants were asked to return the consent form with the completed questionnaire if they agreed to participate. Even for the participants who completed the questionnaire, we provided the option of deleting their research data in case they submitted a consent withdrawal form.

Participants

People aged 65 years and older with a diagnosis of LSS, lumbar disc herniation, or lumbar degenerative spondylolisthesis based on radiological and clinical examination, and who had undergone surgery at Sapporo Maruyama Orthopedic Hospital and Harunaso Hospital at least one year ago were included in the study. LSS in this study was identified using the definition of Arnordi et al²⁴ of "any type of narrowing of the spinal canal, nerve root canals (or tunnels), or intervertebral foramina"; therefore, lumbar disc herniation and lumbar degenerative spondylo-listhesis were considered a type of LSS.

There were 308 older adults who met the inclusion criteria between December 2015 and January 2019. Five had no known addresses and one was dead; as a result, 302 older adults were sent questionnaires.

Assessments

Clinicodemographic Data

Participants' date of birth, gender, date of surgery, and surgical procedure were extracted from the medical records. Current age was calculated from the date of birth. The number of days from the date of surgery to the date of postmark on the return envelope was taken as the number of days since surgery. Surgical procedures were categorized as decompression and fixation, and decompression.

Kinesiophobia

To evaluate kinesiophobia, the Japanese version of the TSK was used,²⁵ which consists of 17 items rated on a 4-point Likert scale: "strongly disagree (1 point)," "disagree (2 points)," "agree (3 points)," and "strongly agree

(4 points)." For questions 4, 8, 12, and 16, the scores are reversed. Total scores range from 17 to 68 points, with higher scores indicating a stronger degree of kinesiophobia. TSK's internal consistency, retest reliability, and validity have already been confirmed.^{22,26}

Pain and Dysesthesia

An 11-point numerical rating scale (NRS), where a score of 0 was defined as "no pain (dysesthesia)" and a score of 10 was defined as "unbearable pain (dysesthesia)," was used. Participants were asked to recall their persistent low back pain, leg pain, and dysesthesia related to lumbar surgery in the last month and rate the average intensity of their pain or dysesthesia. Alghadir et al²⁷ assessed the intensity of knee pain on several scales, including the NRS, and demonstrated the validity of the NRS.

HRQOL

The five-level version of the five-dimensional EuroQol scale (EQ-5D-5L) was used to assess the HRQOL of the participants in the previous month.²⁸ This scale comprises five subscales, with one question each: mobility, self-care, usual activities, pain/discomfort, and anxiety/depression, with five levels of subjective health state for each subscale. Respondents rate each item on a Likert scale ranging from level 1 indicating no problem to level 5 indicating the inability to walk or perform self-care/activities, extreme pain, or extreme anxiety/depression, depending on the subscale being assessed. The EQ-5D-5L has been previously validated.^{29,30} The index value is calculated based on the answers to the five questions and ranges from a score of 0, indicating "a state as bad as being dead" to a score of 1, indicating "full health":³¹ the conversion values were available for the Japanese population.³²

Analytic Procedures

Missing data were complemented through multiple imputation using the chained equation method. All survey items were submitted to the multiple assignment method, and the number of assignments was set to 10. After reviewing the distribution of scores for the 17 questions comprising the TSK, a categorical exploratory factor analysis (C-EFA) (weighted least squares method, promax rotation) based on polychoric correlations was performed on the 17 questions of the TSK for the development of a hypothetical model of the factor structure of the TSK. Parallel analysis was used to determine the number of factors. If items had factor loadings of less than 0.40, the items were removed, and the factor analysis was performed again. The factor analysis was completed when the factor loadings for all items were above 0.40. Once the final factors were determined, the reliability coefficients (Cronbach's alpha coefficient) for each factor were checked. Cronbach's alpha coefficient was set to be between 0.70 and 0.95 as an acceptable range.³³

Subsequently, we conducted confirmatory factor analysis (CFA) (WLSMV method, promax rotation) for each of our hypothetical models as well as the traditional onefactor and two-factor models.¹⁰ Standardized path coefficients and the fit index of the model (Tucker Lewis index [TLI], root mean square error of approximation [RMSEA], standardized root mean residual [SRMR], expected crossvalidation index [ECVI]) were calculated. A TLI of 0.95 or higher, an RMSEA of 0.06 or lower, and an SRMR of 0.08 or lower were judged to be adequate.³⁴ The smaller the value of ECVI, the better the fit of the model.

Finally, Spearman's partial rank correlation coefficients between factors extracted from the TSK, the EQ-5D-5L index value, and scores of pain and dysesthesia were calculated. To calculate the correlation coefficients, age, sex (0: male, 1: female), and days after surgery were used as control variables, referring to the following studies: 1) postoperative course was poor when the patient was over 80 years,^{35,36} 2) being female was a predictor of poor subjective health after lumbar spine surgery,³⁶ and 3) HRQOL changed with postoperative course.³⁷

The surgical procedure (decompression and fixation surgery or decompression surgery) was not employed as a control variable because it has been found to make no difference in long-term outcomes.

R ver. 4.0 (Foundation for Statistical Computing, Vienna, Austria) was used for statistical analysis, and a risk rate of 5% was considered significant.

Results

Participant Characteristics and Data Completion

A set of questionnaires was mailed to 302 older adults (72.8 \pm 4.5 [range: 65–88] years; 164 men, 138 women; 806.0 \pm 341.3 [380–1537] days after surgery; 189 who had undergone decompression and fixation surgery, 113 who had undergone decompression surgery), and responses were obtained from 221 people (73.2%). After excluding four people with additional spinal surgery, three with surgery for osteoarthritis of the lower extremities, two with severe rheumatoid arthritis, and one with a recent cardiac pacemaker implantation, 211

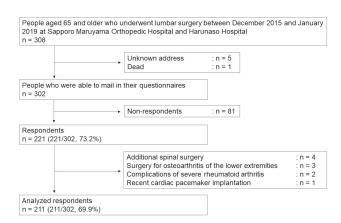


Figure I Flowchart of participants.

people (72.4 \pm 4.2 years [range: 65–88]; 115 men and 96 women; 804 \pm 343.1 [380–1531] days after surgery; 137 who had undergone decompression and fixation surgery, 74 who had undergone decompression surgery) were analyzed (69.9%) (Figure 1). There were no significant differences in age, sex ratio, postoperative days, or ratio of surgical procedures between respondents and non-respondents.

There were 25 items per respondent in this study (pain and dysesthesia: 3 items, EQ-5D-5L: 5 items, TSK: 17 items); hence, the total number of items answered by all 211 respondents was 5275. As 170 items (3.2%) were missing, they were supplemented by the multiple assignment method.

Distribution of Scores on the 17 Items of the TSK

The medians and frequency distributions of the 17 items comprising the TSK are summarized in Table 1. The medians of all items were 2 or 3 points, all items showed an unimodal distribution, and no deviance to minimum or maximum values was observed.

The Development of a Hypothetical Model of the TSK by C-EFA

Initial parallel analysis suggested that TSK is a three-factor structure (Figure 2). Based on the results of the parallel analysis, the initial C-EFA was conducted and, since the factor loading of item 8 was 0.233 at maximum, item 8 was deleted (left column of Table 2). The three-factor model was still maintained even after item 8 was deleted (Figure 3). In the second C-EFA, the factor loadings for item 1 did not exceed 0.40; therefore, item 1 was also deleted. Again, the three-factor model was maintained (Figure 4). Since the factor loadings for all items exceeded 0.40 in the third

| ltem | Median (QD) | Proportion (%) | Proportion (%) | | | | | |
|------|-------------|----------------|----------------|------------|-----------|--|--|--|
| | | I | 2 | 3 | 4 | | | |
| I | 2 (0.5) | 40 (19.0) | 96 (45.5) | 70 (33.2) | 5 (2.4) | | | |
| 2 | 2 (0) | 52 (24.6) | 128 (60.7) | 28 (13.3) | 3 (1.4) | | | |
| 3 | 2 (0) | 46 (21.8) | 117 (55.5) | 44 (20.9) | 4 (1.9) | | | |
| 4 | 2 (0.5) | 12 (5.7) | 133 (63.0) | 50 (23.7) | 16 (7.6) | | | |
| 5 | 2 (0) | 47 (22.3) | 115 (54.5) | 46 (21.8) | 3 (1.4) | | | |
| 6 | 2 (0.5) | 39 (18.5) | 105 (49.8) | 59 (28.0) | 8 (3.8) | | | |
| 7 | 3 (0.5) | 28 (13.3) | 66 (31.3) | 98 (46.4) | 19 (9.0) | | | |
| 8 | 3 (0.5) | 2 (0.9) | 70 (33.2) | (52.6) | 28 (13.3) | | | |
| 9 | 3 (0.5) | 22 (10.4) | 68 (32.2) | 3 (53.6) | 8 (3.8) | | | |
| 10 | 3 (0.5) | 19 (9.0) | 66 (31.3) | 106 (50.2) | 20 (9.5) | | | |
| П | 2 (0.5) | 42 (19.9) | 114 (54.0) | 49 (23.2) | 6 (2.8) | | | |
| 12 | 2 (0.5) | 11 (5.2) | 132 (62.6) | 56 (26.5) | 12 (5.7) | | | |
| 13 | 3 (0.5) | 11 (5.3) | 78 (37.0) | 4 (54.0) | 8 (3.8) | | | |
| 14 | 2 (0.5) | 21 (10.0) | 109 (51.7) | 70 (33.2) | 11 (5.2) | | | |
| 15 | 2 (0.5) | 32 (15.2) | 121 (57.3) | 53 (25.1) | 5 (2.4) | | | |
| 16 | 3 (0.5) | 5 (2.4) | 88 (41.7) | 104 (49.3) | 14 (6.6) | | | |
| 17 | 3 (0) | 6 (2.8) | 45 (21.3) | 137 (64.9) | 23 (10.9) | | | |

| Table I Medians and Frequency Distributions of the 17 Items in the Tampa Scale for Kinesiophob |
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|--|

Note: n = 211.

Abbreviation: QD, quartile deviation.

C-EFA, this model was adopted as our hypothetical model (right column in Table 2). Following previous studies,^{10,15,38} Factor 1 was designated as "somatic focus" and Factor 2 as "activity avoidance." We named Factor 3 "efficacy of physical activities." The Cronbach's alpha coefficients for factors 1, 2, and 3 increased through the C-EFA iterations, eventually reaching 0.829, 0.770, and 0.705, respectively. These values met the criteria for internal consistency.

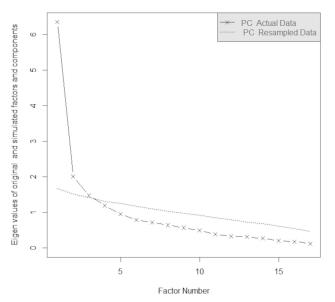
Fitting and Comparison of the Three Models of the TSK by CFA

The hypothetical model, the traditional one-factor model, and the two-factor model were evaluated by CFA in older Japanese adults after lumbar spine surgery (Table 3). No model had a TLI greater than 0.95. RMSEA was below 0.06 and SRMR was below 0.09 in the hypothetical and two-factor models. The one-factor model did not meet the criteria for any of the fit indices. The ECVI of the two-factor model, the hypothetical model, and the one-factor model were evaluated in decreasing order.

The minimum path coefficient to the observed variable in the one-factor model was 0.047, while the minimum path coefficients in the hypothetical and two-factor models were -0.402 and 0.390, respectively, indicating that the observed variables were moderately related to the latent variables. The path coefficients for item 16 belonging to Factor 2 in the hypothetical model and items 16 and 8 in the one-factor model were negative, while the coefficients for the other items were positive.

Relationship Between TSK Factors, and EQ-5D-5L and Pain/Dysesthesia

The median EQ-5D-5L was 0.740 (quartile deviation: 0.097, range: 0.010–1.000). The median NRS scores for low back pain, leg pain, and dysesthesia were 1 (quartile deviation: 1.5, range: 0–10), 1 (1.5, 0–8), and 1 (1.5, 0–8), respectively.



Eigen values of tetrachoric/polychoric matrix

Figure 2 Parallel analysis of the initial categorical exploratory factor analysis.

Since the path coefficient of item 16 belonging to Factor 2 was negative, the total score when item 16 was reversed and the total score when item 16 was deleted were calculated. Spearman's rank correlation coefficients between the total scores of the three TSK factors in the hypothetical model and the index value of EQ-5D-5L as well as the scores of low back pain and leg pain and dysesthesia are summarized in Table 4. The correlation coefficients were adjusted for age, sex, and days after surgery. Factor 1 and Factor 2 showed a significant positive association with the index value of EQ-5D-5L. Factor 3 showed a weak positive effect on EQ-5D-5L, although it was not statistically significant. Factor 1 showed a significant positive association with NRS scores for low back and leg pain and dysesthesia, and Factor 2 showed a significant positive association with NRS scores for low back pain. The correlation coefficients were almost the same when item 16 of Factor 2 was reversed as well as when item 16 was deleted.

Discussion

Psychometric Structure of the Hypothetical Model of TSK in Older Japanese Adults After Lumbar Surgery

Items 1 and 8 were deleted during the process of three rounds of C-EFA, resulting in the development of the hypothetical model with a three-factor structure that ensured internal consistency.

Both items 1 and 9 refer to anxiety about physical injury, which belongs to activity avoidance in the traditional two-factor model, while in the hypothetical model, item 1 was deleted and item 9 was assigned to the somatic focus instead of the activity avoidance. This indicates that the activity avoidance in the hypothetical model reduces the presence of anxiety as a reason for activity avoidance compared to that in the two-factor model. Since anxiety is known to have little mediating effect on pain intensity and disability,³⁹ it may not be necessary to emphasize anxiety in activity avoidance as it is in the hypothetical model. In addition, item 8 was one of the items that was difficult to translate into Japanese.²⁶ Although the appropriateness of the Japanese sentence was assessed by adults and the item eventually passed, it may have been confusing for the older adults to understand. Item 8 did not load strongly with any of the three factors of the hypothesized model; thus, it did not assess any specific construct.

One of the features of the hypothesized model was that the third factor was composed of items 4 and 12, which were deleted in the two-factor model, and we named this factor "efficacy of physical activities." Exercise has been shown to be effective in reducing chronic pain, so-called exercise-induced hypoalgesia (EIH),^{40,41} so it is possible to consider increasing the amount of physical activity in people with chronic pain as a strategy for treating pain. However, EIH is impaired in people with chronic pain;⁴² therefore, encouraging people in pain to continue exercising is a challenge. As the expression of bodyweight exercise behavior has been reported to be influenced by outcome expectations,⁴³ it is important for physical therapists in Japan, whose duties include the management and promotion of physical activity and the prescription of therapeutic exercise, to evaluate the expectation that physical activity or exercise will contribute to pain reduction or health improvement.

Validation of the TSK Hypothesis Model and Comparison to Other Traditional Models

The validity of our hypothetical model as well as the traditional one-factor and two-factor models was tested with CFA. The hypothesized three-factor model and the two-factor model met the criteria for all fit indices except TLI. In contrast, the one-factor model did not meet the criteria for any of the fit indices. The ECVI, where smaller

| | Initial Model Constructed with the First C-EFA | | | Modified Model Constructed with the Third C-EFA | | | | |
|---------------------------------|--|----------|---------------|--|-------|----------|---------------|---------------|
| | Items | Factor I | Factor 2 | Factor 3 | Items | Factor I | Factor 2 | Factor 3 |
| Factor loadings | 3 | 1.176 | -0.442 | 0.027 | 3 | 1.033 | -0.348 | -0.004 |
| | 11 | 0.771 | -0.068 | -0.004 | 11 | 0.811 | -0.108 | 0.003 |
| | 6 | 0.701 | 0.026 | -0.044 | 6 | 0.781 | -0.027 | -0.063 |
| | 5 | 0.486 | 0.072 | -0.130 | 7 | 0.552 | 0.279 | -0.011 |
| | 7 | 0.483 | 0.296 | -0.079 | 5 | 0.527 | 0.098 | -0.018 |
| | 2 | 0.463 | 0.343 | 0.258 | 2 | 0.489 | 0.202 | 0.282 |
| | 9 | 0.450 | 0.314 | 0.101 | 9 | 0.486 | 0.243 | 0.080 |
| | 13 | -0.179 | 0.813 | 0.232 | 17 | -0.238 | 0.831 | -0.011 |
| | 17 | -0.23 I | 0.726 | -0.054 | 16 | 0.034 | -0.647 | 0.246 |
| | 14 | 0.135 | 0.648 | 0.162 | 13 | -0.096 | 0.644 | 0.302 |
| | 10 | 0.168 | 0.618 | 0.080 | 14 | 0.196 | 0.550 | 0.196 |
| | 15 | 0.113 | 0.566 | 0.020 | 10 | 0.245 | 0.517 | 0.104 |
| | 16 | 0.056 | -0.550 | 0.225 | 15 | 0.170 | 0.495 | 0.057 |
| | I | 0.226 | 0.426 | 0.129 | | | | |
| | 4 | 0.088 | 0.042 | 0.785 | 12 | 0.031 | -0.054 | 0.836 |
| | 12 | 0.074 | 0.137 | 0.765 | 4 | -0.163 | -0.085 | 0.771 |
| | 8 | -0.108 | -0.198 | 0.233 | | | | |
| Contribution ratio (cumulative) | | 0.209 | 0.199 (0.408) | 0.092 (0.499) | | 0.240 | 0.179 (0.420) | 0.110 (0.529) |
| Cronbach's alpha | | 0.829 | 0.780 | 0.489 | | 0.829 | 0.770 | 0.705 |
| Factor correlations | Factor 1 <> Factor 2 0.676 Factor 2 <> Factor 3 0.096 Factor 1 <> Factor 3 0.0.25 | | | Factor 1 <> Factor 2 0.649 Factor 2 <> Factor 3 0.347 Factor 1 <> Factor 3 | | | | |

| Table 2 Factor Loadings, Alpha Coefficients, and Correlation Coefficients for Each Factor of the Tampa Scale for Kinesiophobi | Table 2 Factor Loadings, | Alpha Coefficients | , and Correlation Coefficient | s for Each Factor of the | Tampa Scale for Kinesiophobia |
|---|--------------------------|--------------------|-------------------------------|--------------------------|-------------------------------|
|---|--------------------------|--------------------|-------------------------------|--------------------------|-------------------------------|

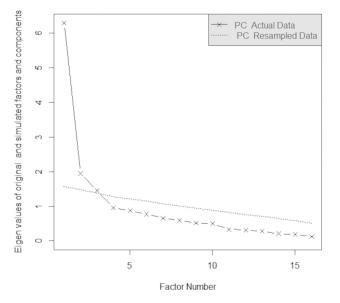
Notes: n = 211. Values above 0.40 are in bold.

Abbreviation: C-EFA, categorical explanatory factor analysis.

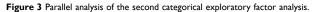
values indicate a better fit, decreased for the one-factor model, hypothetical model, and two-factor model, in that order. The fact that the one-factor model was not a good fit is in accordance with the transactional model of stress,⁴⁴ which shows that the cognitive process of avoiding physical activity and exercise is not simple.

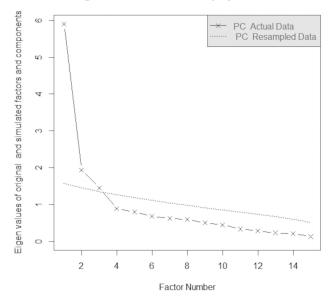
The hypothetical three-factor model was also statistically acceptable, although the two-factor model is a versatile model that has been shown to be valid not only for the older Japanese adults after lumbar surgery in this study but also for people with chronic low back pain, fibromyalgia, and osteoarthritis. The strengths of the hypothetical model in older Japanese adults after lumbar surgery are as described above, allowing for a positive outcome expectation of physical activity and exercise to be captured along with negative factors such as a sense of inadequacy about their own body and avoidance beliefs about physical activity and exercise. However, there are only two items belonging to the third factor, efficacy of physical activities, and we should point out that uncertainty remains. Therefore, additional validation of the factors of efficacy of physical activities is essential.

The validity of the hypothetical model is discussed by focusing on the differences between the items belonging to somatic focus and activity avoidance in the hypothetical model and the two-factor model. Basically, the items belonging to somatic focus and activity avoidance were similar, but items 2 (low motivation to move due to worsening pain) and 9 (anxiety about accidentally injuring the body) belonged to activity avoidance in the two-factor model, whereas they belonged to somatic focus in the hypothetical model. In our cohort, item 2 was significantly



Eigen values of tetrachoric/polychoric matrix





Eigen values of tetrachoric/polychoric matrix

Figure 4 Parallel analysis of the third categorical exploratory factor analysis.

correlated with both other items belonging to somatic focus and activity avoidance (results not shown). Based on this finding, we speculated that item 2 is an ambivalent item, as low motivation to move is likely to be associated with activity avoidance, while the expectation that moving will worsen pain is likely to be associated with somatic focus. Whether the belonging of item 2 to somatic focus is a unique tendency limited to older Japanese adults after lumbar surgery or whether it is due to differences in cultural backgrounds, including language, needs to be further investigated. In addition, the fact that the anxiety asked about in item 9 belonged to somatic focus was in line with research showing that the presence of type D personality, including anxiety tendency, was associated with a risk of low subjective health among older Japanese adults.⁴⁵ Since the report that Japanese university students were more anxious than American students⁴⁶ suggests that there are cultural differences in anxiety tendencies, it is possible that the finding that item 9 belongs to somatic focus is specific to Japanese people.

Another point to be noted in the hypothetical model is that the path coefficient of item 16 (the belief that even if pain is felt, it will not hurt the body), which belongs to activity avoidance, was negative. This means that the strength of activity avoidance belief was paradoxically related to the confidence to cope with pain. It has been found that the reduction of kinesiophobia before and after lumbar surgery is linked to an increase in self-efficacy,⁴⁷ and that there is a positive relationship between selfefficacy and disability in patients with chronic pain and postoperative cardiac patients.^{48,49} Du et al⁵⁰ argue that fear-avoidance belief is a risk factor for QOL and selfefficacy is a protective factor. As far as I could find, no research has shown that activity avoidance beliefs coexist with self-efficacy. The fact that the path coefficient of item 16 in activity avoidance was negative may reflect the psychological state specific to older adults who have long since undergone surgery for LSS, but the Japanese text of item 16 may have been difficult for older adults to understand. Even when item 16 was deleted, the correlation between activity avoidance and HRQOL, pain, and abnormal sensation remained almost the same, so it is recommended that item 16 be deleted to make the interpretation of activity avoidance easier.

Among the three factors in the hypothetical model extracted by CFA, somatic focus was the most frequently associated with HRQOL, low back pain, and dysesthesia in our cohort although there are reports that older adults with chronic pain had lower kinesiophobia.⁵¹ As it has been reported that illness perception was related to HRQOL in patients with fibrous dysplasia,⁵² subjective health is determined by referring to the perception of one's own health status: the stronger the perception of poor health, the stronger the pain and abnormal sensation, and the lower the HRQOL. On the other hand, activity avoidance and efficacy of physical activities were considered to HRQOL, although they

| Model | Hypothetical Model | | Traditional One-Factor Model | | Traditional Two-Factor Model | | |
|---------------------------------|----------------------|-------------|------------------------------|--------|------------------------------|-----------------------|--|
| Latent Variables | ltem | | ltem | | Item | | |
| Factor I | 7 | 0.703 | 2 | 0.706 | 7 | 0.739 | |
| | 2 | 0.701 | 10 | 0.672 | 6 | 0.715 | |
| | 6 | 0.674 | 14 | 0.665 | 3 | 0.663 | |
| | 9 | 0.662 | 7 | 0.663 | 11 | 0.656 | |
| | 11 | 0.633 | 9 | 0.658 | 5 | 0.513 | |
| | 3 | 0.625 | 6 | 0.620 | | | |
| | 5 | 0.483 | 11 | 0.576 | | | |
| | | | 3 | 0.570 | | | |
| | | | 1 | 0.565 | | | |
| | | | 15 | 0.562 | | | |
| | | | 13 | 0.532 | | | |
| | | | 5 | 0.441 | | | |
| | | | 17 | 0.382 | | | |
| | | | 16 | -0.353 | | | |
| | | | 12 | 0.240 | | | |
| | | | 8 | -0.206 | | | |
| | | | 4 | 0.047 | | | |
| Factor 2 | 14 | 0.746 | N.A. | | 2 | 0.732 | |
| | 10 | 0.738 | | | 10 | 0.699 | |
| | 15 | 0.611 | | | 9 | 0.690 | |
| | 13 | 0.595 | | | 14 | 0.685 | |
| | 17 | 0.431 | | | 1 | 0.600 | |
| | 16 | -0.402 | | | 15 | 0.572 | |
| | | | | | 13 | 0.550 | |
| | | | | | 17 | 0.390 | |
| Factor 3 | 12 | 1.000 | N.A. | | N.A. | | |
| | 4 | 0.468 | | | | | |
| Covariances | Factor I <> Factor 2 | | N.A. | N.A. | | Factor I < > Factor 2 | |
| | 0.735 | | | | 0.728 | | |
| | Factor 2 | <> Factor 3 | | | | | |
| | 0.253 | | | | | | |
| | Factor I | <> Factor 3 | | | | | |
| | 0.145 | | | | | | |
| Fit index | | | | | | | |
| TLI | 0.793 | | 0.693 | | 0.905 | | |
| RMSEA (90% confidence interval) | - | 940–0.072) | 0.071 (0.059–0.084) | | 0.050 (0.028–0.069) | | |
| SRMR | 0.072 | | 0.092 | 0.092 | | 0.063 | |
| ECVI | 0.701 | | 1.128 | | 0.498 | | |

Table 3 Estimates and Fit Index of the Three Models According to Confirmatory Factor Analysis

Note: n = 211.

Abbreviations: TLI, Tucker Lewis index; RMSEA, root mean square error of approximation; SRMR, standardized root mean residual; ECVI, expected cross-validation index.

were weakly related to HRQOL and pain or dysesthesia. It can be inferred that avoidance beliefs are involved in the background of choosing maladaptive coping as represented by rest (avoidance of physical activity), which is treated by cognitive behavioral therapy and is negatively related to HRQOL.⁵³ Furthermore, positive expectation of

physical activity is necessary for performing physical activities that are considered to have a positive impact on HRQOL in people with various backgrounds of pain.^{54–57} Therefore, although uncertainty remains regarding the third factor of TSK (efficacy of physical activities), it is important to simultaneously observe the three factors in

| | Factor I | Factor 2 (R)* | Factor 2 (D)** | Factor 3 |
|-----------------|----------------------|----------------------|----------------------|-----------------------|
| EQ-5D | -0.506 (-0.6000.397) | -0.200 (-0.3270.066) | -0.210 (-0.3370.076) | 0.131 (-0.005-0.262) |
| Low back pain | 0.313 (0.185–0.431) | 0.146 (0.010–0.276) | 0.148 (0.012–0.278) | -0.114 (-0.247-0.022) |
| Leg pain | 0.304 (0.176–0.423) | 0.106 (-0.030-0.239) | 0.104 (-0.033-0.236) | -0.032 (-0.167-0.105) |
| Leg dysesthesia | 0.351 (0.226–0.465) | 0.113 (-0.023-0.245) | 0.117 (-0.019-0.249) | -0.015 (-0.151-0.121) |

Table 4 Spearman's Partial Rank Correlations Between Factors of the TSK and EQ-5D, Pain, and Dysesthesia

Notes: n = 211. Age, sex (0: male, 1: female), postoperative days, and surgery procedure (0: with fusion; 1: with fusion) were used as controlling variables. The values in parentheses represent the 95% confidence interval. *The total score was calculated by reversing the scores for item 16. **The total score was calculated by deleting item 16. **Abbreviations:** TSK, Tampa Scale for Kinesiophobia; EQ-5D, EuroQol-5 dimensions.

order to broadly understand kinesiophobia in older Japanese adults after lumbar surgery.

Study Limitations and Future Prospects

This study had two major methodological limitations. First, although the pain and dysesthesia after lumbar surgery reported by the older Japanese adults in this study were likely to be chronic in nature, the duration of the pain and dysesthesia was not collected. Therefore, we cannot conclude that the findings in this study are limited to chronic pain and dysesthesia. Second, as the survey was voluntary, including an optional written questionnaire, there may have been bias in the sample characteristics and responses (i.e., self-selection bias). Moreover, since the respondents were asked to recall their experiences in the most recent month when answering the questions, distortions in the respondents' memory, represented by the phenomenon of telescoping, is another bias that may have affected the results of this study. Thus, the results of this study need to be interpreted cautiously to account for the presence of such biases.

In spite of these limitations, the present study is the first to examine the psychometric structure of TSK in Japanese, especially in older adults after lumbar surgery, and confirmed that the goodness-of-fit of the model did not deteriorate considerably when a new three-factor model, in which efficacy of physical activities was added as a third factor to the traditional two factors of automatic focus and activity avoidance, was adopted for the TSK among older Japanese adults after lumbar surgery. However, there is uncertainty in the third factor, which consists of only two items, and verification is required to compensate for this uncertainty. Traditional somatic focus and activity avoidance were factors that were also valid for older Japanese adults after lumbar surgery; thus, the two factors were accepted internationally regardless of cultural background. Furthermore, among the three factors of the TSK, somatic focus and physical avoidance showed significant negative effects on HRQOL. In this regard, future studies are needed to test the effectiveness of interventions focusing on kinesiophobia to enhance HRQOL among older Japanese adults suffering from pain and dysesthesia after lumbar surgery.

Acknowledgments

We would like to express our deepest gratitude to Kentaro Tajima (physical therapist, Higashi-Maebashi Orthopaedic Hospital, formerly Harunaso Hospital) for his help with data collection, and Editage for English language editing.

Funding

This research was funded by JSPS KAKENHI Grant Number JP19K11201. The JSPS was not involved in any part of the study, except for the provision of funds.

Disclosure

Dr Daisuke Higuchi reports grants from Japan Society for the Promotion of Science (JSPS), during the conduct of the study. The authors declare that they have no other competing interests.

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