# **ORIGINAL RESEARCH**

# Comparing Effect Estimates in Randomized Trials and Observational Studies From the Same Population: An Application to Percutaneous Coronary Intervention

Anthony A. Matthews , PhD; Karolina Szummer, MD, PhD; Issa J. Dahabreh , MD ScD; Bertil Lindahl, MD, PhD; David Erlinge , MD, PhD; Maria Feychting, PhD; Tomas Jernberg, MD, PhD; Anita Berglund, PhD; Miguel A. Hernán, MD, DrPH

**BACKGROUND:** To understand when results from observational studies and randomized trials are comparable, we performed an observational emulation of a target trial designed to ask similar questions as the VALIDATE (Bivalirudin Versus Heparin in ST-Segment and Non–ST-Segment Elevation Myocardial Infarction in Patients on Modern Antiplatelet Therapy) randomized trial. The VALIDATE trial compared the effect of bivalirudin and heparin during percutaneous coronary intervention on the risk of death, myocardial infarction, and bleeding across Sweden.

**METHODS AND RESULTS:** We specified the protocol of a target trial similar to the VALIDATE trial, then emulated the target trial in the period before the VALIDATE trial took place using data from the SWEDEHEART (Swedish Web System for Enhancement and Development of Evidence-Based Care in Heart Disease Evaluated According to Recommended Therapies) registry—the same registry in which the trial was undertaken. The target trial emulation and the VALIDATE trial both estimated little or no effect of bivalirudin versus heparin on the risk of death or myocardial infarction by 180 days (target trial emulation risk ratio for death, 1.21 [95% CI, 0.88 – 1.54]; VALIDATE trial hazard ratio for death, 1.05 [95% CI, 0.78 – 1.41]). The observational data, however, could not capture less severe cases of bleeding, resulting in an inability to define a bleeding outcome like the trial, and could not accurately estimate the comparative risk of death by 14 days, which may be the result of intractable confounding early in follow-up or the inability to precisely emulate the trial's eligibility criteria.

**CONCLUSIONS:** Using real-world data to emulate a target trial can deliver accurate effect estimates. Yet, even with rich observational data, it is not always possible to estimate the short-term effect of interventions or the effect on outcomes for which data are not routinely collected.

Key Words: methodology 
observational studies 
randomized controlled trials 
registry 
SWEDEHEART 
target trial emulation

bservational analyses of routinely collected healthcare data, a form of "real-world evidence," are often used to evaluate the benefits and risks of clinical interventions for acute coronary syndromes. Most criticisms of these observational analyses revolve around the lack of randomized assignment of the treatment strategies under comparison, which may result in confounded effect estimates.<sup>1–3</sup> Despite this limitation, routinely collected healthcare data can be used to explicitly emulate target trials, as recent applications in diverse clinical areas have shown.<sup>4–8</sup>

Correspondence to: Anthony A. Matthews, PhD, Unit of Epidemiology, Institute of Environmental Medicine, Karolinska Institutet, Stockholm, Sweden. E-mail: anthony.matthews@ki.se

Preprint posted on MedRxiv February 3, 2021. 10.1101/2021.02.01.21250739.

Supplementary Material for this article is available at https://www.ahajournals.org/doi/suppl/10.1161/JAHA.120.020357

For Sources of Funding and Disclosures, see page 11.

<sup>© 2021</sup> The Authors. Published on behalf of the American Heart Association, Inc., by Wiley. This is an open access article under the terms of the Creative Commons Attribution License, which permits use, distribution and reproduction in any medium, provided the original work is properly cited. *JAHA* is available at: www.ahajournals.org/journal/jaha

## **CLINICAL PERSPECTIVE**

## What is New?

- Using observational data to emulate a target trial can deliver accurate estimates for the comparative effect of bivalirudin and heparin during percutaneous coronary intervention on the risk of death and myocardial infarction by 180 days.
- Observational data cannot be used to answer all clinical questions as study-specific information is not always routinely collected. In this instance, it was not possible to accurately estimate the effect on the risk of bleeding or the short-term effect on the risk of death.

## What are the Clinical Implications?

• Synergy between randomized trials and observational analyses of real-world data has the potential to revolutionize the length of time patients can be followed up and the speed at which results from trials reach clinical practice.

## Nonstandard Abbreviations and Acronyms

| RIKS-HIA   | Register of Information and<br>Knowledge About Swedish Heart<br>Intensive Care Admissions   |
|------------|---|
| SCAAR      | Swedish Coronary Angiography and Angioplasty Registry   |
| SEPHIA     | Swedish Heart Surgery Register<br>and the National Registry of<br>Secondary Prevention  |
| SWEDEHEART | Swedish Web System for<br>Enhancement and Development<br>of Evidence-Based Care in Heart<br>Disease Evaluated According to<br>Recommended Therapies |
| VALIDATE   | Bivalirudin Versus Heparin in<br>ST-Segment and Non–ST-<br>Segment Elevation Myocardial<br>Infarction in Patients on Modern<br>Antiplatelet Therapy |

Routinely collected observational data can only be used to emulate pragmatic trials, not trials with placebo control and blind treatment assignment, or those that rely on detailed data collection to determine eligibility or ascertain outcomes.<sup>9</sup> Specifically, the lack of detailed information on eligibility criteria and outcomes restricts the type of causal questions that can be answered using real-world data, a point that is not always emphasized in discussions about the topic.<sup>4</sup>

To study the strengths and limitations of observational analyses that emulate a trial, a near-ideal scenario is to compare a registry-based trial versus the emulation of a target trial with the same protocol as the registry-based trial using observational data from the same registry.<sup>10</sup> This comparison ensures that the causal question is asked in the same population and healthcare setting.11 SWEDEHEART (Swedish Web System for Enhancement and Development of Evidence-Based Care in Heart Disease Evaluated According to Recommended Therapies) is a national quality registry of myocardial infarction (MI), coronary intervention, and heart surgery, which includes longitudinal information on demographic and clinical characteristics, use of therapeutic and preventive services, diagnostic procedures, and various measures of healthcare utilization for the whole of Sweden.<sup>12</sup> By using the registry to enroll participants, randomize interventions, and report outcomes, SWEDEHEART can be used to run trials nested within the registry.<sup>13</sup> An example of a registrybased trial within SWEDEHEART is the VALIDATE (Bivalirudin Versus Heparin in ST-Segment and Non-ST-Segment Elevation Myocardial Infarction in Patients on Modern Antiplatelet Therapy) trial, which compared 2 anticoagulation interventions, bivalirudin and heparin, over 180 days in patients with acute MI undergoing percutaneous coronary intervention (PCI) in Sweden.<sup>14</sup>

Here, we use the observational data from the SWEDEHEART registry to emulate a target trial designed to answer similar questions as the VALIDATE trial, and we outline conditions under which one can successfully design a target trial and emulate it using observational data. We describe the VALIDATE trial and the observational data from the SWEDEHEART registry, the specification of the target trial, and the emulation procedures. Our example illustrates the opportunities and limitations of healthcare registries to emulate a target trial.

## VALIDATE: THE INDEX TRIAL

## **Trial Design and Analysis**

The VALIDATE trial was a multicenter, randomized, controlled, open-label clinical trial performed between June 2014 and September 2016. Individuals in the SWEDEHEART registry were eligible to participate if they were admitted to the hospital with a diagnosis of ST-segment–elevation MI (STEMI) or non–ST-segment–elevation MI (NSTEMI), urgent PCI was planned in 1 of 25 of the 29 PCI centers in Sweden, and some other eligibility criteria were met (Table 1). Individuals who accepted the invitation to participate were randomly assigned to receive either

| Target Trial Emulation Using<br>SWEDEHEART | Same as target trial  | Same as target trial. Individuals were classified according to the strategy that their data were compatible with at date of PCI.   | Assignment was treated as if randomized<br>within levels of the following baseline<br>covariates: severity of MI (STEMI/NSTEMI,<br>Killip class), center where PCI took<br>place (university hospital/nonuniversity<br>hospital), demographics (sex, age),<br>lifestyle characteristics (weight), laboratory<br>measurements (GFP), diagnoses (bleeding,<br>anemia), and medications (warfarin, novel<br>oral anticoagulants, P2Y12 inhibitors) |
|--|---|--|---|
| Target Trial                               | Same as VALIDATE apart from:<br>Study period between January 2012 and May 2014<br>Without asking for informed consent<br>Known terminal disease defined as palliative care, dialysis,<br>dementia, or aggressive cancer (pancreatic, mesothelioma,<br>lung, liver, stomach, or brain) in 3 y before PCI<br>Unable to identify ongoing bleeding<br>No information on dose of heparin before PCI, so excluded all<br>individuals given any prior heparin<br>individuals who died on the day of PCI were excluded and<br>identification of outcomes start from day after PCI   | (1) Individual administered bivalrudin under PCI<br>(2) Individual administered heparin under PCI<br>(2) Individual administered heparin under PCI<br>If the patient was given both bivalirudin and heparin, then<br>it was assumed that heparin was low-dose intra-arterial<br>administration of up to 3000 U, and the treatment strategy<br>was therefore bivalirudin, as no information was available on<br>dosage. | Individuals randomized to a treatment strategy under PCI and were aware of the assigned strategy  |
| VALIDATE Trial                             | Age ≥18 years between June 2014 and September 2016<br>Diagnosis of NSTEMI as defined by guidelines (positive troponin) or<br>STEMI as defined by chest pain for at least 30 minutes and an ECG<br>with new ST-segment elevation in ≥2 contiguous leads of ≥0.2 mV in<br>leads V2–V3 and/or ≥0.1 mV in other leads or a probable new-onset<br>left bundle branch block<br>Therapeutic PCI (not primarily diagnostic PCI) after angiography in 1<br>of 25 (of 29) selected clinics<br>Ability to provide informed consent: In patients with STEMI,<br>witnessed oral consent was obtained after angiography and before<br>randomization. Within the following 24 h, after written information<br>about the trial had been provided, the patients confirmed further<br>participation by providing written informed consent. In patients with<br>NSTEMI, written consent was obtained before angiography.<br>Treated with a bolus dose of ticagrefor, prasugrel, or cangrelor<br>Paticipation by providing dose of ticagrefor, prasugrel, or cangrelor<br>Patients with known ongoing bleeding<br>No heparin 5000 U before arriving to PCI laboratory. Up to 3000 U<br>of heparin allowed during angiography.<br>No uncontrolled hypertension, subacute bacterial endocarditis,<br>severe renal (GFR <30 mL/min) and/or liver dysfunction,<br>thrombocytopenia or thrombocyte function defects<br>No ocortraindication for the study medications<br>No gycoprotein IIb/III inhibitors given or are preplanned to be given<br>during the procedure | <ol> <li>Bivalirudin: intravenous bolus of 0.75 mg/kg of body weight<br/>followed by an infusion of 1.75 mg/kg per hour. Treatment<br/>was started as soon as PCI of the culprit lesion was planned.<br/>Continuation of the bivalirudin infusion after PCI until completion of<br/>the last vial is strongly recommended.</li> <li>(2) Heparin: 70 to 100 U/kg</li> </ol>   | Before PCI, patients were randomly assigned to receive in an open-label fashion either intravenous bivalirudin or intra-arterial unfractionated heparin. Randomization was performed in a 1:1 ratio with stratification according to type of MI (STEMI or NSTEMI) and hospital.   |
| Protocol Component                         | Eligibility criteria  | Treatment strategies   | Treatment assignment  |

Table 1. Description of the VALIDATE Randomized Trial, Target Trial, and Target Trial Emulation Using SWEDEHEART Register

(Continued)

| Target Trial Emulation Using<br>SWEDEHEART | Same as target trial   | Same as target trial   | Observational analogue of the per-protocol effect | Same as target trial  | ntion; STEMI, ST-segment-elevation myocardi                            |
|--|--|--|---|---|--|
| Target Trial                               | Same as VALIDATE but outcomes identified as follows:<br>Composite death from any cause identified from the Swedish<br>Cause of Death register, MI identified from the SWEDEHEART<br>register, and major bleeding events identified from the Swedish<br>Inpatient and Outpatient registers or the SWEDEHEART<br>register by 180 d | Same as VALIDATE apart from:<br>Unable to identify migration date  | Intention-to treat-effect<br>Per-protocol effect  | Same intention-to-treat analysis as VALIDATE Per-protocol<br>analysis:<br>Kaplan-Meier plots adjusted for baseline covariates via inverse<br>probability weighting<br>Treatment differences were estimated with use of pooled<br>logistic regression models with a flexible time-varying intercept<br>and product terms between the treatment and time. Baseline<br>covariates adjusted for using 2 atternative approaches:<br>inverse probability weighting (the baseline covariates were<br>not included in the outcome model) and standardization (the<br>baseline covariates were included in the outcome model). | vation myocardial infarction; PCI, percutaneous coronary interver      |
| VALIDATE Trial                             | Composite of death from any cause, MI, or major bleeding events<br>by 180 d<br>Each separate component of the composite  | Starts at treatment assignment and ends at date of first outcome<br>(separately for analysis of each outcome), migration, or 180 d after<br>baseline | Intention-to-treat effect                         | Intention-to-treat analysis<br>Kaplan-Meier plots<br>Treatment differences were estimated with the use of the log-rank<br>test and Cox regression   | filtration rate; MI, myocardial infarction; NSTEMI, non-ST-segment-ele |
| Protocol Component                         | Outcomes   | Follow-up  | Causal contrast                                   | Statistical analysis  | GFR indicates glomerular   |

Table 1. Continued

infarction; SWEDEHEART, Swedish Web System for Enhancement and Development of Evidence-Based Care in Heart Disease Evaluated According to Recommended Therapies; and VALIDATE, Bivalirudin Versus Heparin in ST-Segment and Non-ST-Segment Elevation Myocardial Infarction in Patients on Modern Antiplatelet Therapy. yocardial

Comparing Trials and Observational Studies

intravenous bivalirudin (0.75 mg/kg) or intra-arterial unfractionated heparin (5000 U/mL) under PCI intervention. The primary end point was the composite of death from any cause, MI, or major bleeding by 180 days. The individual components of the composite event were also assessed. Research nurses screened for clinical end point events by contacting the patients or first-degree relatives by telephone 7 days and 180 days after PCI. If the patient or relatives could not be contacted after repeated telephone calls and a mailed letter, information was collected through review of hospital records. The intention-to-treat analysis, described in detail elsewhere, relied on the comparison of 180-day risk differences and hazard ratios (HRs).<sup>14</sup>

## **Trial Results**

A total of 6006 patients underwent randomization in the VALIDATE trial, with 3004 assigned to bivalirudin and 3002 assigned to heparin. The risk of the composite end point did not differ between the treatment groups 30 days after PCI. By 180 days, the composite end point occurred in 12.3% of patients (369 of 3004) in the bivalirudin group and in 12.8% (383 of 3002) in the heparin group, with an HR of 0.96 (95% Cl, 0.83 - 1.10). Death occurred in 2.9% of patients in the bivalirudin group and in 2.8% in the heparin group (without accounting for censoring), with an HR of 1.05 (95% CI, 0.78 - 1.41). MI occurred in 2.0% of patients in the bivalirudin group and in 2.4% in the heparin group, with an HR of 0.84 (95% Cl, 0.60 – 1.19). Major bleeding occurred in 8.6% of patients in both the bivalirudin and heparin groups, with an HR of 1.00 (95% CI, 0.84 - 1.19).

## SPECIFYING AND EMULATING A TARGET TRIAL IN THE SWEDEHEART REGISTRY

## **Data Sources**

SWEDEHEART collects data from all patients hospitalized for acute coronary syndrome or undergoing coronary or valvular intervention for any indication in all relevant hospitals across Sweden.<sup>13</sup> The registry was officially launched in 2009 when 4 existing cardiovascular healthcare quality registries were merged: RIKS-HIA (Register of Information and Knowledge About Swedish Heart Intensive Care Admissions), SCAAR (Swedish Coronary Angiography and Angioplasty Registry), SEPHIA (Swedish Heart Surgery Register and the National Registry of Secondary Prevention), and the Swedish Heart Surgery Registry. SWEDEHEART was used to collect information for patients when they were randomized in the VALIDATE trial; hence, the target trial was emulated in the same population as the original trial and the data collection process was broadly similar between the 2 studies. SWEDEHEART is also linked to the Inpatient and Outpatient Register, which records all primary and secondary diagnoses and procedures from inpatient hospitalizations and outpatient specialist care visits across Sweden; the Swedish Cause of Death register, which records all deaths and causes of death; and the Prescribed Drug register, which collects information on all dispensed medications.<sup>15–17</sup>

The approach for emulating a target trial similar to the VALIDATE trial had 2 steps: (1) specifying the protocol of the target trial, and (2) emulating the target trial using the observational data from the SWEDEHEART registry. The first step is straightforward because our target trial has the same protocol as the VALIDATE trial, with exceptions only when the observational data were not adequate to identify the information collected in the trial. Table 1 summarizes the target trial protocol and outlines the emulation procedure described below.

## **Eligibility Criteria**

We identified individuals in the SWEDEHEART registry who met the target trial's eligibility criteria. The eligibility criteria for the target trial were the same as the VALIDATE trial with 6 exceptions. First, the study period was between January 2012 to May 2014, which immediately precedes the period of the VALIDATE trial, as all eligible patients during that period were considered in the VALIDATE trial. Second, no informed consent was requested and therefore the target trial could not exclude individuals who would have not been asked or declined participation if asked. Third, the target trial used a proxy measure to indicate known terminal disease with life expectancy <1 year (palliative care; dialysis; dementia; or malignant disease, including pancreatic, lung, liver, stomach, and brain cancer, or mesothelioma). Fourth, some eligibility criteria could not be applied in the target trial because of unavailable data (ongoing bleeding, contraindications to study medication). Fifth, there was no information on the dosage of heparin before PCI, thus we excluded all individuals with any prior heparin use. Sixth, individuals who died on the day of PCI were excluded and identification of outcomes started from MI and bleeding occurred before or after PCI when the events occurred on the same day as the procedure.

## **Treatment Strategies**

The treatment strategies in the target trial closely mimicked those in the VALIDATE trial and were: (1) administration of bivalirudin, or (2) administration of



**Figure 1.** Flowchart of individuals eligible for the observational emulation of a target trial of bivalirudin vs heparin, SWEDEHEART (Swedish Web System for Enhancement and Development of Evidence-Based Care in Heart Disease Evaluated According to Recommended Therapies) register, 2012 to 2014.

NSTEMI indicates non-ST-segment-elevation myocardial infarction; PCI, percutaneous coronary intervention; and STEMI, ST-segment-elevation myocardial infarction.

heparin during PCI. The SWEDEHEART registry contains information on the anticoagulant given under PCI but does not contain information on dosage. In the target trial, we therefore assumed that if a patient was given both treatments, then the heparin was low-dose, and the treatment strategy was defined as bivalirudin.

#### **Treatment Assignment**

We classified eligible individuals in the SWEDEHEART registry into 2 groups according to the strategy their data were compatible with at baseline. That is, our emulation presumed that all individuals assigned to a treatment strategy ended up receiving it. Our emulation treated individuals as if they were randomly

# Table 2.Baseline Characteristics of Eligible Individualsfor the Observational Emulation of a Target Trial ofBivalirudin Versus Heparin, SWEDEHEART Register, 2012to 2014

|  | Bivalirudin   | Heparin       |
|--|---------------|---------------|
| No.  | 2634          | 2306          |
| Women, n (%)                                     | 731 (27.8)    | 652 (28.3)    |
| Age, median (IQR), y                             | 68 (59–76)    | 69 (61–76)    |
| University hospital, n (%)                       | 1074 (40.8)   | 995 (43.1)    |
| STEMI, n (%)                                     | 2131 (80.9)   | 398 (17.3)    |
| Killip class, n (%)                              | I             |               |
| 1  | 2413 (91.6)   | 1894 (82.1)   |
| 2  | 96 (3.6)      | 34 (1.5)      |
| 3  | 21 (0.8)      | 8 (0.3)       |
| 4  | 26 (1.0)      | 5 (0.2)       |
| Missing  | 78 (3.0)      | 365 (15.8)    |
| Angiography finding, n (%)                       |               |               |
| Normal   | 3 (0.1)       | 50 (2.2)      |
| 1 Vessel + no left main                          | 1365 (51.8)   | 1135 (49.2)   |
| 2 Vessels + no left main                         | 720 (27.3)    | 701 (30.4)    |
| 3 Vessels + no left main                         | 426 (16.2)    | 330 (14.3)    |
| Left main  | 120 (4.6)     | 90 (3.9)      |
| Heart rate, median (IQR), beats per min          | 75 (63–89)    | 76 (66–89)    |
| Missing, n (%)                                   | 98 (3.7)      | 67 (2.9)      |
| Systolic blood pressure, median<br>(IQR), mm Hg  | 144 (125–165) | 155 (140–174) |
| Missing, n (%)                                   | 99 (3.8)      | 67 (2.9)      |
| Diastolic blood pressure,<br>median (IQR), mm Hg | 85 (75–97)    | 88 (79–99)    |
| Missing, n (%)                                   | 279 (10.6)    | 144 (6.2)     |
| Anemia severity category, n (%)                  |               |               |
| Severe   | 2 (0.1)       | 3 (0.1)       |
| Moderate   | 91 (3.5)      | 42 (1.8)      |
| Mild   | 318 (12.1)    | 211 (9.2)     |
| No anemia  | 2039 (77.4)   | 1836 (79.6)   |
| Missing  | 184 (7.0)     | 214 (9.3)     |
| Hemoglobin, median (IQR), g/L                    | 142 (131–152) | 144 (133–153) |
| Weight, median (IQR), kg                         | 80 (70–90)    | 81 (72–92)    |
| Smoking status, n (%)                            |               |               |
| Never  | 959 (36.4)    | 895 (38.8)    |
| Ex-smoker (>1 mo)                                | 776 (29.5)    | 843 (36.6)    |
| Current  | 736 (27.9)    | 522 (22.6)    |
| Missing  | 163 (6.2)     | 46 (2.0)      |
| Prior MI, n (%)                                  | 406 (15.4)    | 519 (22.5)    |
| Prior PCI, n (%)                                 | 323 (12.3)    | 426 (18.5)    |
| Prior coronary artery bypass grafting, n (%)     | 115 (4.4)     | 185 (8.0)     |
| Prior serious bleeding, n (%)                    | 87 (3.3)      | 112 (4.9)     |
| Diabetes mellitus, n (%)                         | 429 (16.3)    | 459 (19.9)    |
| Kidney function (GFR), median<br>(IQR)           | 87 (65–112)   | 85 (64–107)   |

(Continued)

#### Table 2. Continued

|                                     | Bivalirudin | Heparin     |
|-------------------------------------|-------------|-------------|
| Hypertension treatment, n (%)       | 1256 (47.7) | 1343 (58.2) |
| Lipid-lowering treatment, n (%)     | 733 (27.8)  | 1035 (44.9) |
| Warfarin, n (%)                     | 27 (1.0)    | 40 (1.7)    |
| Outcomes                            |             |             |
| Composite outcome (30 d),<br>n (%)  | 140 (5.3)   | 83 (3.6)    |
| Composite outcome (180<br>d), n (%) | 253 (9.6)   | 211 (9.2)   |
| Death (30 d), n (%)                 | 86 (3.3)    | 22 (1.0)    |
| Death (180 d), n (%)                | 130 (4.9)   | 62 (2.7)    |
| MI (30 d), n (%)                    | 27 (1.0)    | 27 (1.2)    |
| MI (180 d), n (%)                   | 59 (2.2)    | 75 (3.3)    |
| Bleeding (30 d), n (%)              | 32 (1.2)    | 37 (1.6)    |
| Bleeding (180 d), n (%)             | 84 (3.2)    | 91 (3.9)    |

GFR indicates glomerular filtration rate; IQR, interquartile range; MI, myocardial infarction; PCI, percutaneous coronary intervention; STEMI, ST-segment–elevation myocardial infarction; and SWEDEHEART, Swedish Web System for Enhancement and Development of Evidence-Based Care in Heart Disease Evaluated According to Recommended Therapies.

assigned to a treatment strategy conditional on the baseline covariates: severity of MI (STEMI/NSTEMI, Killip class, angiography finding), center where the PCI took place (university hospital/nonuniversity hospital), demographics (sex, age), lifestyle characteristics (weight, smoking), laboratory measurements (glomerular filtration rate, heart rate, systolic blood pressure, diastolic blood pressure), diagnoses (diabetes mellitus, severe bleeding, anemia), medications (warfarin, antihypertensives, lipid-lowering treatment), and prior cardiovascular disease or cardiovascular procedures (MI, PCI, coronary artery bypass grafting). See Table S1 for further details on covariates and their definitions.

#### Outcomes

The outcomes in the target trial were the same as the VALIDATE trial and included death from any cause, MI, major bleeding events, and a composite of all of the above outcomes. Death was identified from the Cause of Death register, MI from the SWEDEHEART registry, and bleeding from the Inpatient and Outpatient register or SWEDEHEART registry (full outcome definitions in are shown in Table S2).

### Follow-Up

Each individual was followed from the day after PCI until the outcome of interest or 180 days, whichever occurred first. In the MI and bleeding analyses, individuals were not censored if they died during follow-up, which is valid when estimating the total effect of treatment, but we performed sensitivity analysis to explore

|                              | Risk, % (95% Cl)  |                    |                             |                     |
|------------------------------|-------------------|--------------------|-----------------------------|---------------------|
| Outcome                      | Bivalirudin       | Heparin            | Risk Difference, % (95% CI) | Risk Ratio (95% CI) |
| Inverse-probability weighted |                   | -<br>-             | `<br>                       |                     |
| Composite                    | 9.3 (8.2 to 10.4) | 10.0 (8.7 to 11.3) | -0.7 (-2.5 to 1.1)          | 0.93 (0.77 to 1.12) |
| Death                        | 4.1 (3.4 to 4.8)  | 3.4 (2.5 to 4.2)   | 0.7 (-0.4 to 1.8)           | 1.21 (0.88 to 1.68) |
| MI                           | 3.0 (2.3 to 3.8)  | 2.8 (2.2 to 3.4)   | 0.2 (-0.8 to 1.2)           | 1.08 (0.76 to 1.54) |
| Bleeding                     | 3.2 (2.5 to 3.9)  | 4.6 (3.7 to 5.6)   | -1.4 (-2.6 to -0.3)         | 0.69 (0.50 to 0.95) |
| Standardized                 |                   |                    |                             |                     |
| Composite                    | 9.2 (7.5 to 10.8) | 9.9 (8.0 to 11.9)  | -0.8 (-3.6 to 2.0)          | 0.92 (0.74 to 1.14) |
| Death                        | 4.3 (3.2 to 5.4)  | 3.4 (2.1 to 4.6)   | 1.0 (-0.8 to 2.7)           | 1.28 (0.92 to 1.80) |
| MI                           | 2.8 (1.6 to 3.9)  | 2.7 (1.7 to 3.7)   | 0.1 (-1.7 to 1.8)           | 1.03 (0.66 to 1.59) |
| Bleeding                     | 3.0 (1.9 to 4.1)  | 4.3 (2.6 to 5.9)   | -1.3 (-3.6 to 1.1)          | 0.70 (0.46 to 1.09) |

# Table 3. Estimated 180-d Risk, Risk Difference, and RRs from the Observational Emulation\* of a Target Trial of Bivalirudin Versus Heparin, SWEDEHEART Register, 2012–2014

MI indicates myocardial infarction; and SWEDEHEART, Swedish Web System for Enhancement and Development of Evidence-Based Care in Heart Disease Evaluated According to Recommended Therapies.

\*Adjusted at baseline for severity of MI (STEMI/NSTEMI, Killip class, angiography finding), center where percutaneous coronary intervention (PCI) took place (university hospital/nonuniversity hospital), demographics (sex, age), lifestyle characteristics (weight, smoking), laboratory measurements (glomerular filtration rate, heart rate, systolic blood pressure, diastolic blood pressure), diagnoses (diabetes mellitus, severe bleeding, anemia), medications (warfarin, antihypertensives, lipid lowering treatment), and prior cardiovascular disease or cardiovascular procedures (MI, PCI, coronary artery bypass grafting).

censoring at death (Sensitivity 10).<sup>18</sup> Outcome data on individuals who migrated out of Sweden were unavailable, but the probability of migration during such a short period is low.

### **Causal Contrast**

In the target trial, the intention-to-treat effect was the effect of being assigned to bivalirudin versus heparin at baseline and the per-protocol effect was the effect of receiving the assigned treatment. Our emulation focused on an observational analog of the per-protocol effect. In the VALIDATE trial there was 0.5% crossover between treatment groups, so the intention-to-treat effect was approximately equal to the per-protocol effect.

## **Statistical Analysis**

We estimated the 14-, 30-, and 180-day risk (cumulative incidence), risk difference (RD), and risk ratio (RR) for each outcome using pooled logistic regression models with a flexible time-varying intercept and product terms between the treatment and time.<sup>19</sup> We adjusted for baseline covariates using 2 alternative approaches: inverse probability weighting and outcome regression followed by standardization. We truncated the stabilized inverse probability weights at their 99th percentile to prevent outliers with extreme weights from influencing effect estimates. We plotted unweighted and inverse probability-weighted Kaplan-Meier survival curves for all outcomes. We used nonparametric bootstrapping with 500 samples to calculate all 95% Cls. Further details on our modelling approach are in Data S1.

We repeated the 180-day analysis using inverse probability-weighted models stratified by MI type (STEMI/NSTEMI) to explore effect modification. Finally, we performed several sensitivity analyses using inverse probability-weighted models to assess the sensitivity of the eligibility criteria, treatment strategies, outcomes, confounders, follow-up, and analysis. Full details on sensitivity analyses are in Data S2.

## Data, Materials, and Code Disclosure

All analysis codes are available at: https://github.com/ tonymatthews/validate. Pseudonymized personal data were obtained from national Swedish Registry holders after ethical approval and secrecy assessment. According to Swedish laws and regulations, personal sensitive data can only be made available for researchers who fulfill legal requirements for access to personal sensitive data.

## RESULTS

Figure 1 shows a flowchart of patient selection for the target trial emulation. Table 2 shows the baseline characteristics of the 4940 eligible patients, of whom 2634 were given bivalirudin and 2306 heparin. Compared with those in the heparin group, patients in the bivalirudin group were more likely to be diagnosed with STEMI over NSTEMI before PCI (81% versus 17%).

Table 3 shows the estimated 180-day risks, RD, and RRs obtained via inverse probability weighting and standardization. The inverse probability–weighted risk of the composite outcome was 9.3% (95% CI, 8.2%



Figure 2. Inverse-probability weighted survival curves from an observational emulation of a target trial of bivalirudin vs heparin, SWEDEHEART (Swedish Web System for Enhancement and Development of Evidence-Based Care in Heart Disease Evaluated According to Recommended Therapies) register, 2012–2014.

– 10.4%) in the bivalirudin group and 10.0% (95% Cl, 8.7% – 11.3%) in the heparin group, which results in a RD of –0.7% (95% Cl, –2.5% to 1.1%) and an RR of 0.93 (95% Cl, 0.77 – 1.12). The risk of death was 4.1% (95% Cl, 3.4% – 4.8%) in the bivalirudin group and 3.4% (95% Cl, 2.5% – 4.2%) in the heparin group, which results in a risk difference of 0.7% (95% Cl, –0.4% to

1.8%) and RR of 1.21 (95% Cl, 0.88 - 1.68). The risk of MI was 3.0% (95% Cl, 2.3% - 3.8%) in the bivalirudin group and 2.8% (95% Cl, 2.2% - 3.4%) in the heparin group, which results in a RD of 0.2% (95% Cl, -0.8% to 1.2%) and RR of 1.08 (95% Cl, 0.76 - 1.54). The risk of bleeding was 3.2% (95% Cl, 2.5% - 3.9%) in the bivalirudin group and 4.6% (95% Cl, 3.7% - 5.6%) in

Comparing Trials and Observational Studies

the heparin group, which results in a RD of -1.4% (95% CI, -2.6% to -0.3%) and an RR of 0.69 (95% CI, 0.50 - 0.95). Standardized estimates were similar.

Figure 2 shows the inverse probability–weighted survival curves (unweighted survival curves in Figure S2). The MI-free survival almost overlapped throughout the 180 days of follow-up. However, there was an elevated risk of death in the bivalirudin group compared with the heparin group in the 14 days after PCI. The short-term difference in risk of death can also be seen in the estimated risk difference and RR by 14 and 30 days (Tables S3 and S4).

Table S5 shows the 180-day risks, RDs, and RRs obtained via inverse probability weighting, stratified by MI type, and there were no apparent differences between strata. Tables S6 to S15 and Figure S2 show results from sensitivity analyses; all results were broadly similar to those estimated in the main analyses.

## DISCUSSION

We emulated a target trial similar to the VALIDATE trial using real-world observational data from SWEDEHEART, the same register in which the trial was undertaken. There was broad agreement in the estimates for death and MI by 180 days from the trial and the observational emulation of the target trial; both found little differences in risk between patients given heparin and bivalirudin. However, there were some important differences between the estimates from the VALIDATE trial and its observational emulation.

First, the target trial could not define a bleeding outcome similar to that in the VALIDATE trial, which was ascertained using a combination of phone calls to patients at 7 and 180 days after PCI and a review of hospital records from registers. Exclusively using the registers, as was done in the target trial, only allowed us to identify the most severe cases of bleeding. We also attempted to expand the bleeding outcome in a sensitivity analysis (outcome definition in Table S2 and results in Table S12), but we were still unable to fully capture less severe cases of bleeding. This difference in the definition of bleeding is likely responsible for the lower risk estimates for the bleeding and the composite outcomes in the VALIDATE trial and target trial emulation.

Second, the observational estimates showed a RD of 1.0% (95% Cl, 0.1% – 1.8%) for death after 14 days of follow-up (Figure 2, Table S3), whereas the trial estimates showed no discernible difference in mortality throughout the entire follow-up. This almost instantaneous difference in mortality may be attributable to prognostic factors that were imbalanced between treatment groups, and unadjusted for, but that had less impact over the entire follow-up period. It is possible,

for example, that patients perceived to be at high risk of bleeding, and hence at higher risk of early death, may have been more likely to be administered bivalirudin (because bivalirudin was marketed as having a lower risk of bleeding than heparin). Another explanation is that the data available on severity of MI were not granular enough to fully account for the fact that individuals administered bivalirudin were more likely to be severely ill or have a poor early prognosis, as reflected, for example, by the higher proportion of patients diagnosed with STEMI in the bivalirudin compared with the heparin group. Even with rich data from the SWEDEHEART register, we were not able to accurately identify these potentially important characteristics. One way to aid future researchers in identifying significant potential confounders would be if registers included more detailed information on reasons for making treatment decisions. Another possible, but perhaps less likely, explanation for this difference in the short-term risk of death is the inability to emulate a target trial that would have precisely the same eligibility criteria as the index trial; we discuss this issue in more detail below.

The number of individuals randomized in the VALIDATE trial was 6006 (2971 individuals either declined, were not asked, or could not be asked to participate), whereas the number of individuals included in the target trial emulation over a recruitment period of similar length was 4940. This discrepancy is partly because of our inability to emulate the eligibility criteria for the VALIDATE trial. Individuals given >5000 U heparin before PCI in the VALIDATE trial were excluded, but the SWEDEHEART data do not include details on the dosage of heparin administered. We, therefore, excluded all 3070 individuals given any dose of heparin before PCI, of whom some would have received low doses. A sensitivity analysis that did not exclude any patient given heparin before PCI (Table S9) included 7652 individuals, and results for death and MI by 180 days were broadly similar to the results from primary analyses. The RD for death between patients given bivalirudin and heparin early in follow-up was, however, reduced in this sensitivity analysis, as shown in the inverse probability-weighted Kaplan-Meier survival curves (Figure S2). This suggests that by excluding patients given any prior heparin in the main analysis, rather than only those given high-dose heparin as in the trial, we may be selecting a group of individuals given heparin during PCI who are at lower risk of death soon after start of follow-up compared with those in the trial. Results from this sensitivity analysis demonstrate the importance of designing a target trail that closely mimics the eligibility criteria of the index trial if aiming for comparable results, and, as mentioned, is another possible explanation for the differences in short-term risk of death discussed above.

The SWEDEHEART register contains some of the most complete and rich routinely collected data available worldwide when a patient undergoes PCI. Even with high-quality data, however, close harmonization of protocols, adjustment for important confounders, and analytic methods appropriate for estimating causal quantities analogous to those estimated in the trial, we have shown that it is not always possible for the target trial emulation to obtain the same results as an index trial attributable to trials collecting more detailed, study-specific information that are not routinely collected. Nevertheless, we have shown that using real-world data to emulate a target trial can deliver accurate effect estimates for this specific clinical question. It is, therefore, possible to foresee a system in which data from trials with short-term follow-up are combined with real-world data to estimate the effects of certain interventions over longer periods of follow-up.<sup>20</sup> Estimating the effect from the short trial would overcome the potential problem of confounding soon after the intervention. We could then use observational data to estimate the effect during the remainder of follow-up and create synthetic survival curves using both data sources. This would enable prompt estimation of effects of interventions over longer periods of follow-up and allow adaptive clinical or regulatory changes to be put guickly into practice. Synergy between trials and real-world data has the potential to revolutionize both the length of time patients can be followed up and the speed at which results from trials reach clinical practice.

#### **ARTICLE INFORMATION**

Received November 30, 2020; accepted March 22, 2021.

#### Affiliations

Unit of Epidemiology, Institute of Environmental Medicine, Karolinska Institutet, Stockholm, Sweden (A.A.M., M.F., A.B., M.A.H.); Department of Cardiology, Karolinska University Hospital, Stockholm, Sweden (K.S.); Department of Medicine, Karolinska Institutet, Huddinge, Sweden (K.S.); Department of Biostatistics (I.J.D., M.A.H.); and Department of Epidemiology (I.J.D., M.A.H.), Harvard T.H. Chan School of Public Health, Boston, MA; Department of Medical Sciences, Cardiology and Uppsala Clinical Research Center, Uppsala University, Uppsala, Sweden (B.L.); Department of Cardiology, Clinical Sciences, Lund University, Skåne University Hospital, Lund, Sweden (D.E.); Department of Clinical Sciences, Danderyd University Hospital-Karolinska Institute, Danderyd, Sweden (T.J.); and Harvard-MIT Division of Health Sciences and Technology, Boston, MA (M.A.H.).

#### Sources of Funding

This work was supported by a grant from the Swedish Research Council (registration number 2018-03028).

#### Disclosures

Dr Matthews reports grants from FORTE during the conduct of the study; Dr Dahabreh reports grants from PCORI (award ME-1502-27794) during the conduct of the study; Dr Erlinge reports personal fees from AstraZeneca, outside the submitted work; Dr Feychting, Dr Berglund, and Dr Hernan report grants from the Swedish Research Council during the conduct of the study; and Dr Hernán reports personal fees from Cytel and ProPublica during the conduct of the study. The remaining authors have no disclosures to report.

#### Supplementary Material

Data S1–S2 Tables S1–S15 Figures S1–S2

#### REFERENCES

- Gerstein HC, McMurray J, Holman RR. Real-world studies no substitute for RCTs in establishing efficacy. *Lancet*. 2019;393:210–211. DOI: 10.1016/S0140-6736(18)32840-X
- Collins R, Bowman L, Landray M, Peto R. The magic of randomization versus the myth of real-world evidence. N Engl J Med. 2020;382:674– 678. DOI: 10.1056/NEJMsb1901642
- Bowman L, Baras A, Bombien R, Califf RM, Chen Z, Gale CP, Gaziano JM, Grobbee DE, Maggioni AP, Muse ED, et al. Understanding the use of observational and randomized data in cardiovascular medicine. *Eur Heart J*. 2020;41:2571–2578. DOI: 10.1093/eurheartj/ehaa020
- Booth CM, Karim S, Mackillop WJ. Real-world data: towards achieving the achievable in cancer care. *Nat Rev Clin Oncol.* 2019;16:312–325. DOI: 10.1038/s41571-019-0167-7
- Petito LC, Garcia-Albeniz X, Logan RW, Howlader N, Mariotto AB, Dahabreh IJ, Hernan MA. Estimates of overall survival in patients with cancer receiving different treatment regimens: emulating hypothetical target trials in the surveillance, epidemiology, and end results (SEER)medicare linked database. *JAMA Netw Open.* 2020;3:e200452. DOI: 10.1001/jamanetworkopen.2020.0452
- Dickerman BA, Garcia-Albeniz X, Logan RW, Denaxas S, Hernan MA. Avoidable flaws in observational analyses: an application to statins and cancer. Nat Med. 2019;25:1601–1606. DOI: 10.1038/s41591-019-0597-x
- Lodi S, Phillips A, Lundgren J, Logan R, Sharma S, Cole SR, Babiker A, Law M, Chu H, Byrne D, et al. Effect estimates in randomized trials and observational studies: comparing apples with apples. *Am J Epidemiol.* 2019;188:1569–1577. DOI: 10.1093/aje/kwz100
- Hernan MA, Alonso A, Logan R, Grodstein F, Michels KB, Willett WC, Manson JE, Robins JM. Observational studies analyzed like randomized experiments: an application to postmenopausal hormone therapy and coronary heart disease. *Epidemiology*. 2008;19:766–779. DOI: 10.1097/EDE.0b013e3181875e61
- Hernán MA, Robins JM. Using big data to emulate a target trial when a randomized trial is not available. *Am J Epidemiol.* 2016;183:758–764. DOI: 10.1093/aje/kwv254
- Forbes SP, Dahabreh IJ. Benchmarking observational analyses against randomized trials: a review of studies assessing propensity score methods. J Gen Intern Med. 2020;35:1396–1404. DOI: 10.1007/s11606-020-05713-5
- James S, Rao SV, Granger CB. Registry-based randomized clinical trials—a new clinical trial paradigm. *Nat Rev Cardiol.* 2015;12:312–316. DOI: 10.1038/nrcardio.2015.33
- Thygesen LC, Ersboll AK. When the entire population is the sample: strengths and limitations in register-based epidemiology. *Eur J Epidemiol.* 2014;29:551–558. DOI: 10.1007/s10654-013-9873-0
- Jernberg T, Attebring MF, Hambraeus K, Ivert T, James S, Jeppsson A, Lagerqvist B, Lindahl B, Stenestrand U, Wallentin L. The Swedish Web-system for enhancement and development of evidence-based care in heart disease evaluated according to recommended therapies (SWEDEHEART). *Heart.* 2010;96:1617–1621. DOI: 10.1136/ hrt.2010.198804
- Erlinge D, Omerovic E, Fröbert O, Linder R, Danielewicz M, Hamid M, Swahn E, Henareh L, Wagner H, Hårdhammar P, et al. Bivalirudin versus heparin monotherapy in myocardial infarction. *N Engl J Med.* 2017;377:1132–1142. DOI: 10.1056/NEJMoa1706443
- Ludvigsson JF, Andersson E, Ekbom A, Feychting M, Kim JL, Reuterwall C, Heurgren M, Olausson PO. External review and validation of the Swedish national inpatient register. *BMC Public Health*. 2011;11:450. DOI: 10.1186/1471-2458-11-450
- Brooke HL, Talback M, Hornblad J, Johansson LA, Ludvigsson JF, Druid H, Feychting M, Ljung R. The Swedish cause of death register. *Eur J Epidemiol.* 2017;32:765–773. DOI: 10.1007/s10654-017-0316-1
- Wettermark B, Hammar N, Fored CM, Leimanis A, Otterblad Olausson P, Bergman U, Persson I, Sundstrom A, Westerholm B, Rosen M. The new Swedish Prescribed Drug Register–opportunities for pharmacoepidemiological research and experience from the first six months.

Pharmacoepidemiol Drug Saf. 2007;16:726–735. DOI: 10.1002/ pds.1294

- Young JG, Stensrud MJ, Tchetgen Tchetgen EJ, Hernan MA. A causal framework for classical statistical estimands in failure-time settings with competing events. *Stat Med.* 2020;39:1199–1236. DOI: 10.1002/ sim.8471
- 19. Hernán MA, Robins JM. Causal Inference: What If. Boca Raton, FL: Chapman & Hall/CRC. 2020.
- Dahabreh IJ, Robins JM, Hernan MA. Benchmarking observational methods by comparing randomized trials and their emulations. *Epidemiology.* 2020;31:614–619. DOI: 10.1097/EDE.000000000 001231

# SUPPLEMENTAL MATERIAL

Data S1.

#### Inverse probability weighting and standardization

In our target trial, the intention to treat effect is the effect of being assigned to bivalirudin versus heparin under percutaneous coronary intervention on the risk of death, myocardial infarction, bleeding, or a composite of all three. Estimating its observational analog requires adjustment for baseline confounders, which we did in two ways. First, we used inverse probability weighting of marginal structural models, and second, we used standardization.

#### Inverse probability weighed models

First, we estimate the stabilized inverse probability weight,  $SW^A$ , truncated to the 99<sup>th</sup> percentile, for each individual in the study population. Informally, the denominator of this weight is the probability that an individual received her observed treatment, A, given her measured confounders, L, and the numerator is the probability of receiving her observed treatment, A.

$$SW^A = f(A)/f(A|L)$$

Second, we fit pooled logistic regression models where individuals are weighted by their estimated  $SW^A$ .

$$logit(Pr[Y_{m+1} = 0|A, Y_m = 0]) = \alpha_{0,m} + \alpha_1 A + \alpha_2 Am + \alpha_3 Am^2$$

where:

 $Y_{m+1}$  = Indicator for the outcome of interest at time m+1.

A = Indicator for the treatment group.

 $\alpha_{0,m}$  = Time-varying intercept estimated as a constant plus linear and quadratic terms at time *m*.

The predicted probabilities from the inverse probability weighted hazards model can then be multiplied over time, in the absence of censoring, to obtain an estimate of survival under each treatment strategy at each time, m.

$$\widehat{IPW_k^a} = \prod_{m=1}^k expit(\widehat{\alpha}_{0,m} + \widehat{\alpha}_1 a + \widehat{\alpha}_2 am + \widehat{\alpha}_3 am^2)$$

The risk at time, k, can then be estimated by taking one minus  $\widehat{IPW}_k^a$ , and risk differences and risk ratios can be calculated. Finally, we use nonparametric bootstrapping with 500 samples to calculate all 95% confidence intervals.

#### **Standardization models**

First, we fit the pooled logistic regression model.

$$logit(Pr[Y_{m+1} = 0|A, L, Y_m = 0]) = \beta_{0,m} + \beta_1 A + \beta_2 Am + \beta_3 Am^2 + \beta_4^T L$$

where:

 $Y_{m+1}$  = Indicator for the outcome of interest at time m+1.

A = Indicator for the treatment group.

L = Vector of potential confounders at baseline

 $\beta_{0,m}$  = Time-varying intercept estimated as a constant plus linear and quadratic terms at time *m*.

We then use the predicted probabilities from these models, multiplied over time, in the absence of censoring, to obtain an estimate of survival of each individual, *i*, under each treatment strategy at each time, *m*, conditional on the individual's baseline confounders,  $L_i$ .

$$\widehat{OM_{i,k}^{a}} = \prod_{m=1}^{k} expit(\widehat{\beta}_{0,m} + \widehat{\beta}_{1}a + \widehat{\beta}_{2}am + \widehat{\beta}_{3}am^{2} + \widehat{\beta}_{4}^{T}L_{i})$$

We then standardized the survival probabilities at each time point to the empirical distribution of the confounders at baseline.

$$\widehat{OM_k^a} = \frac{1}{n} \sum_{i=1}^n \widehat{OM_{i,k}^a}$$

Where *n* is the number of individuals in the study population. The risk at time, *k*, can then be estimated by taking one minus  $\widehat{OM}_k^a$ , and risk differences and risk ratios can be calculated. Finally, we used nonparametric bootstrapping with 500 samples to calculate all 95% confidence intervals.

## Data S2.

## Description of sensitivity analyses

## Sensitivity analysis 1: eligibility criteria 1

We expanded the definition of diagnoses that could indicate a life expectancy of less than a year (one of the exclusion criteria) to include all cancer diagnoses (ICD-10 code: C).

## Sensitivity analysis 2: eligibility criteria 2

We excluded all patients over the age of 75 years at the time of PCI.

## Sensitivity analysis 3: eligibility criteria 3

We only included those who were recorded as given P21Y2 before percutaneous coronary intervention, not under percutaneous intervention

## Sensitivity analysis 4: eligibility criteria 4

We included everyone given any dose of heparin before coronary intervention

## Sensitivity analysis 5: eligibility criteria 5

We included everyone with a missing GFR and imputed missing GFR and weight variables with the median GFR and weight in the analysis.

## Sensitivity analysis 6: treatment strategy

We restricted treatment to those who were only given bivalirudin or heparin (original definition is exposure to bivalirudin if given both treatments as assumed that heparin is low dose if have been given both). 1,466 of the 1,951 (75%) of the patients originally assigned to bivalirudin were also given low dose heparin at the same time. These patients were excluded from this analysis.

## Sensitivity analysis 7: outcome

We expanded the definition of bleeding to include less severe bleeding outcomes. The additional ICD-10 codes we identified in the inpatient and outpatient registers were:

- anal K62.5
- anovulatory N97.0
- atonic, following delivery O72.1
- capillary I78.8
  - o puerperal O72.2
- contact N93.0
- due to uterine subinvolution N85.3
- ear H92.2

- excessive, associated with menopausal onset N92.4
- following intercourse N93.0
- gastrointestinal K92.2
- hemorrhoids K64
- intermenstrual (regular) N92.3
  - o irregular N92.1
- irregular N92.6
- menopausal N92.4
- nipple N64.59
- nose R04.0
- ovulation N92.3
- perimenopausal N92.4
- postclimacteric N95.0
- postcoital N93.0
- postmenopausal N95.0
- pre-pubertal vaginal N93.1
- preclimacteric N92.4
- puberty N92.2 (excessive, with onset of menstrual periods)
  - rectum, rectal K62.5
  - o newborn P54.2
- throat R04.1

•

- tooth socket K91.840 (post-extraction)
- umbilical stump P51.9
- uterus, uterine NEC N93.9
  - o climacteric N92.4
  - o dysfunctional or functional N93.8
  - o menopausal N92.4
  - o preclimacteric or premenopausal N92.4
  - unrelated to menstrual cycle N93.9
- vagina, vaginal (abnormal) N93.9
  - o dysfunctional or functional N93.8
  - o newborn P54.6
  - o pre-pubertal N93.1
- vicarious N94.89

### Sensitivity analysis 8: confounders

We changed the definition to the confounders that were defined in the inpatient, outpatient, and prescribed drug registers (bleeding, warfarin, and NOAC), so they were identified from a 5-year look back period rather than 3-year.

## Sensitivity analysis 9: missing data

We carried out a complete case analysis, excluding patients that had missing data for the Killip class and anemia variables (18%).

## Sensitivity analysis 10: censoring at death

In the analyses for the myocardial infarction and bleeding outcomes, we censored individuals at date of death.

| Covariate Register Definition (if further detail required)<br>used |   | Form   | Categories        |   |
|--|---|--|-------------------|---|
| MI type  | SCAAR                                     |  | Indicator         | STEMI/NSTEMI                                      |
| Killip class   | SCAAR                                     |  | 5 categories      | 1, 2, 3, 4, missing                               |
| Angiography finding  | SCAAR                                     | $\begin{array}{l} 0 - normal \\ 1 - 1 \ vessel + no \ left \ main \\ 2 - 2 \ vessels + no \ left \ main \\ 3 - 3 \ vessels + no \ left \ main \\ 4 - left \ main \end{array}$  | 5 categories      | 0, 1, 2, 3, 4                                     |
| Age  | SCAAR                                     |  | Linear, quadratic | N/A   |
| Gender   | SCAAR                                     |  | Indicator         | Male/female                                       |
| Weight   | SCAAR                                     | Kilograms  | Linear, quadratic | N/A   |
|  |   | No missing as to calculate GFR for eligibility, a non-<br>missing weight is needed.  |                   |   |
| Prior bleeding   | Inpatient and<br>outpatient<br>registries | The patient registries were used to identify primary<br>diagnoses of intracranial bleeding; major gastrointestinal<br>bleeding; respiratory, renal/urinary tract, ocular,<br>retroperitoneal or pericardial bleeding; and bleeding<br>attributed to anemia within 3 years of PCI.  | Indicator         | Yes/No  |
|  |   | These outcomes were identified using the following<br>ICD-10 codes: D62, I60, I61, I62, R31, R04, D500,<br>H313, H356, H431, H450, I312, I850, K250, K252,<br>K254, K256, K260, K262, K264, K266, K270, K272,<br>K274, K276, K280, K282, K284, K286, K625, K661,<br>K920, K921, K922, S064, S065, S066, J942, K228F,<br>K298A, K638B, K638C, K838F, K868G, I864A,<br>H052A, S368D, G951A |                   |   |
| Warfarin   | SCAAR                                     | Warfarin before or under PCI, if missing assumed no  | Indicator         | Yes/No  |
| Anemia   | RIKSHIA                                   | Hemoglobin (g/L) was identified from the RIKSHIA<br>register, then categorized to indicate severity of anemia<br>Based on WHO anemia guidelines.   | 5 categories      | 1, 2, 3, 4, missing                               |
|  |   | 1 - Severe - lower than 70<br>2 - Moderate - 80-109<br>3 - Mild - 110-119 (women), 110-129 (men)<br>4 - No anemia - over 120 (women), over 130 (men)<br>9 - Missing  |                   |   |
| GFR  | SCAAR and<br>RIKSHIA                      | First used creatinine recording from SCAAR register. If<br>not available from SCAAR, then checked RIKSHIA<br>register one month prior to PCI.  | Linear, quadratic | N/A   |
|  |   | GFR then calculated using the Cockcroft-Gault formula:<br>GFR = $(((140-age)*weight)/(72*SCr))$ (*0.85 if female)  |                   |   |
|  |   | No missing as eligibility criteria required to have GFR reading (excluded if severe renal disease or missing GFR)  |                   |   |
| PCI Centre   | SCAAR                                     | Categorized into those that are part of a University<br>Hospital: Karolinska Solna, Linköping, Örebro,<br>Sahlgrenska, Lund, Malmö, Umeå, Uppsala.   | Indicator         | University<br>hospital/Not<br>university hospital |
| Prior myocardial infarction  | SCAAR                                     | If missing assumed no  | Indicator         | Yes/No  |
| Prior percutaneous coronary intervention                           | SCAAR                                     | If missing assumed no  | Indicator         | Yes/No  |
| Prior coronary artery bypass grafting                              | SCAAR                                     | If missing assumed no  | Indicator         | Yes/No  |

# Table S1. Covariate definitions in the target trial.

| Diabetes                         | SCAAR   | If missing assumed no  | Indicator         | Yes/No           |
|----------------------------------|---------|--|-------------------|------------------|
| Prior treatment for hypertension | SCAAR   | If missing assumed no  | Indicator         | Yes/No           |
| Prior lipid lowering treatment   | SCAAR   | If missing assumed no  | Indicator         | Yes/No           |
| Smoking                          |         | 0 – never<br>1 – ex-smoker (> 1 month)<br>2 – current<br>9 - missing                         | 4 categories      | 0, 1, 2, missing |
| Heart rate                       | RIKSHIA | Some missing as captured from RIKSHIA and not all<br>individuals have related RIKSHIA record | Linear, quadratic | N/A              |
|                                  |         | In main analysis missing imputed with median   |                   |                  |
| Systolic blood pressure          | RIKSHIA | Some missing as captured from RIKSHIA and not all individuals have related RIKSHIA record    | Linear, quadratic | N/A              |
|                                  |         | In main analysis missing imputed with median   |                   |                  |
| Diastolic blood pressure         | RIKSHIA | Some missing as captured from RIKSHIA and not all individuals have related RIKSHIA record    | Linear, quadratic | N/A              |
|                                  |         | In main analysis missing imputed with median   |                   |                  |

# Table S2. Outcome definitions in the target trial.

| VALIDATE trial  | Target trial definition  |
|---|--|
| Death from any cause  | 1. SCAAR<br>The SCAAR registry was used to identify date of death  |
|   | <b>2. Cause of death register</b><br>If there was no death record in the SCAAR register, but a record of death in the cause of death registry, then date of death was taken from cause of death register.  |
| Myocardial infarction - the third universal<br>myocardial definition (Thygesen et al Eur<br>Heart J 2012;33:2551- 67)                         | <b>RIKSHIA</b><br><b>a.</b> $\leq$ <b>2 days after PCI</b><br>If there was a record of a myocardial infarction in the RIKSHIA registry within 2 days<br>after PCI, then there also needed to be record of troponin levels greater than 40ng/L<br>(both high sensitivity and conventional troponin, and I and T troponin) or a reinfarction.<br>This was to ensure the record was not a repeat record of the original myocardial<br>infarction. |
|   | <b>b.</b> >2 days after PCI<br>After 2 days, through until the end of follow up, all myocardial infarction records in the<br>RIKSHIA registry were used (ICD 10 codes I21/I22).  |
| Major bleeding event - adjudicated according<br>to the BARC definition, type 2, 3, or 5<br>(Mehran R et al Circulation 2011; 123:2736-<br>47) | <b>1. Inpatient and outpatient registries</b><br>The patient registries were used to identify primary diagnoses of intracranial bleeding;<br>major gastrointestinal bleeding; respiratory, renal/urinary tract, ocular, retroperitoneal or<br>pericardial bleeding; and bleeding attributed to anemia.   |
|   | These outcomes were identified using the following ICD-10 codes: D62, I60, I61, I62, R31, R04, D500, H313, H356, H431, H450, I312, I850, K250, K252, K254, K256, K260, K262, K264, K266, K270, K272, K274, K276, K280, K282, K284, K286, K625, K661, K920, K921, K922, S064, S065, S066, J942, K228F, K298A, K638B, K638C, K838F, K868G, I864A, H052A, S368D, G951A  |
|   | <b>2. RIKSHIA</b><br>The RIKSHIA registry also records if a patient bled within a period of care. This was used to identify lethal bleeding, cerebral bleeding, or bleeding that required transfusion.   |
| Composite of all above  | As above   |

Table S3. Estimated 14-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, using inverse probability weighting models.

| Outcome               | <b>Risk (%, 95% CI)</b> |               | <b>Risk difference</b> | <b>Risk ratio</b> |
|-----------------------|-------------------------|---------------|------------------------|-------------------|
|                       | Bivalirudin             | Heparin       | (95% CI)               | (95% CI)          |
| Composite             | 3.7 (2.9,4.4)           | 3.3 (2.4,4.1) | 0.4 (-0.7,1.5)         | 1.12 (0.79,1.58)  |
| Death                 | 2.1 (1.5,2.6)           | 1.1 (0.5,1.7) | 1.0 (0.1,1.8)          | 1.85 (0.95,3.63)  |
| Myocardial infarction | 1.1 (0.6,1.6)           | 0.9 (0.6,1.3) | 0.2 (-0.5,0.8)         | 1.16 (0.58,2.32)  |
| Bleeding              | 0.6 (0.3,0.9)           | 1.3 (0.8,1.8) | -0.7 (-1.3,-0.1)       | 0.46 (0.22,0.98)  |

Table S4. Estimated 30-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, using inverse probability weighting models.

| Outcome               | Risk (%, 95% CI) |                | Risk difference | Risk ratio       |
|-----------------------|------------------|----------------|-----------------|------------------|
|                       | Bivalirudin      | Heparin        | (95% CI)        | (95% CI)         |
| Composite             | 5.1 (4.3,6.0)    | 4.6 (3.6,5.66) | 0.5 (-0.8,1.8)  | 1.10 (0.83,1.46) |
| Death                 | 2.7 (2.1,3.3)    | 1.7 (1.0,2.39) | 1.0 (0.1,1.9)   | 1.61 (0.99,2.63) |
| Myocardial infarction | 1.4 (0.8,2.0)    | 1.1 (0.7,1.56) | 0.3 (-0.4,1.0)  | 1.22 (0.68,2.19) |
| Bleeding              | 1.3 (0.8,1.7)    | 2.0 (1.4,2.69) | -0.8 (-1.6,0.1) | 0.63 (0.37,1.06) |

| Outcome               | Risk (%, 95% CI) |                 | Risk difference  | Risk ratio       |
|-----------------------|------------------|-----------------|------------------|------------------|
|                       | Bivalirudin      | Heparin         | (95% CI)         | (95% CI)         |
| STEMI                 |                  |                 |                  |                  |
| Composite             | 10.4 (9.3,11.4)  | 11.9 (9.3,14.6) | -1.6 (-4.4,1.2)  | 0.87 (0.68,1.11) |
| Death                 | 5.5 (4.8,6.3)    | 4.3 (2.6,6.1)   | 1.2 (-0.7,3.1)   | 1.28 (0.79,2.07) |
| Myocardial infarction | 2.1 (1.6,2.5)    | 2.5 (1.5,3.6)   | -0.5 (-1.6,0.7)  | 0.81 (0.48,1.37) |
| Bleeding              | 3.4 (2.8,4.0)    | 5.9 (3.9,7.9)   | -2.6 (-4.6,-0.5) | 0.57 (0.39,0.84) |
| NSTEMI                |                  |                 |                  |                  |
| Composite             | 8.4 (6.4,10.5)   | 8.5 (7.5,9.4)   | 0.0 (-2.3,2.3)   | 1.00 (0.75,1.33) |
| Death                 | 2.7 (1.6,3.7)    | 2.4 (1.9,2.9)   | 0.3 (-0.9,1.5)   | 1.11 (0.68,1.82) |
| Myocardial infarction | 4.1 (2.6,5.7)    | 3.0 (2.5,3.6)   | 1.1 (-0.6,2.7)   | 1.36 (0.86,2.15) |
| Bleeding              | 3.1 (1.9,4.4)    | 3.5 (2.9,4.2)   | -0.4 (-1.8,1.0)  | 0.89 (0.54,1.46) |

Table S5. Estimated 180-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, stratified by STEMI/NSTEMI, using inverse probability weighting models

\* Adjusted at baseline for severity of myocardial infarction (STEMI/NSTEMI, Killip class, angiography finding), center where percutaneous coronary intervention took place (university hospital/non-university hospital), demographics (sex, age), lifestyle characteristics (weight, smoking), laboratory measurements (GFR, heart rate, systolic blood pressure, diastolic blood pressure), diagnoses

(diabetes, severe bleeding, anemia), medications (warfarin, anti-hypertensives, lipid lowering treatment), and prior cardiovascular disease or cardiovascular procedures (myocardial infarction, percutaneous coronary intervention, coronary artery bypass grafting)

Table S6. Sensitivity 1 - Estimated 180-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, further excluding those with all cancer, using inverse probability weighting models (N=4,527).

| Outcome               | <b>Risk (%, 95% CI)</b> |                | <b>Risk difference</b> | Risk ratio       |
|-----------------------|-------------------------|----------------|------------------------|------------------|
|                       | Bivalirudin             | Heparin        | (95% CI)               | (95% CI)         |
| Composite             | 9.0 (7.9,10.1)          | 9.0 (7.7,10.3) | -0.1 (-1.8,1.7)        | 0.99 (0.81,1.21) |
| Death                 | 3.8 (3.1,4.5)           | 2.9 (2.0,3.8)  | 0.9 (-0.2,2.0)         | 1.30 (0.90,1.87) |
| Myocardial infarction | 3.1 (2.3,3.9)           | 2.6 (2.0,3.2)  | 0.5 (-0.5,1.6)         | 1.20 (0.81,1.76) |
| Bleeding              | 3.1 (2.4,3.7)           | 4.3 (3.4,5.3)  | -1.3 (-2.4,-0.1)       | 0.71 (0.51,0.98) |

Table S7. Sensitivity 2 - Estimated 180-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, excluding all patients over the age of 75 years, using inverse probability weighting models (N=3,624).

| Outcome               | <b>Risk (%, 95% CI)</b> |               | <b>Risk difference</b> | Risk ratio       |
|-----------------------|-------------------------|---------------|------------------------|------------------|
|                       | Bivalirudin             | Heparin       | (95% CI)               | (95% CI)         |
| Composite             | 6.7 (5.5,7.9)           | 6.7 (5.3,8.0) | 0.0 (-1.7,1.7)         | 1.00 (0.77,1.31) |
| Death                 | 2.4 (1.7,3.1)           | 1.5 (0.9,2.1) | 0.9 (0.0,1.8)          | 1.57 (0.91,2.69) |
| Myocardial infarction | 2.5 (1.6,3.4)           | 2.0 (1.4,2.5) | 0.5 (-0.5,1.6)         | 1.28 (0.78,2.08) |
| Bleeding              | 2.6 (2.0,3.3)           | 3.6 (2.6,4.6) | -1.0 (-2.1,0.2)        | 0.73 (0.50,1.07) |

Table S8. Sensitivity 3 - Estimated 180-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, including only patients given P21Y2 inhibitors before percutaneous coronary intervention, using inverse probability weighting models (N=4,113)

| Outcome               | Risk (%, 95% CI) |                | <b>Risk difference</b> | Risk ratio       |
|-----------------------|------------------|----------------|------------------------|------------------|
|                       | Bivalirudin      | Heparin        | (95% CI)               | (95% CI)         |
| Composite             | 9.0 (7.7,10.3)   | 9.5 (8.3,10.7) | -0.5 (-2.2,1.3)        | 0.95 (0.78,1.16) |
| Death                 | 3.5 (2.8,4.3)    | 3.1 (2.3,3.9)  | 0.5 (-0.6,1.5)         | 1.15 (0.82,1.63) |
| Myocardial infarction | 3.5 (2.5,4.5)    | 3.0 (2.4,3.6)  | 0.5 (-0.7,1.6)         | 1.16 (0.80,1.69) |
| Bleeding              | 3.1 (2.4,3.9)    | 4.3 (3.4,5.3)  | -1.2 (-2.4,0.0)        | 0.73 (0.52,1.02) |

Table S9. Sensitivity 4 - Estimated 180-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, including all individuals given heparin before percutaneous coronary intervention, using inverse probability weighting models (N=7,652).

| Outcome               | <b>Risk (%, 95% CI)</b> |                 | <b>Risk difference</b> | <b>Risk ratio</b> |
|-----------------------|-------------------------|-----------------|------------------------|-------------------|
|                       | Bivalirudin             | Heparin         | (95% CI)               | (95% CI)          |
| Composite             | 9.4 (8.6,10.2)          | 10.0 (8.7,11.2) | -0.5 (-2.0,0.9)        | 0.94 (0.81,1.10)  |
| Death                 | 3.9 (3.5,4.4)           | 3.6 (2.7,4.4)   | 0.4 (-0.5,1.3)         | 1.11 (0.86,1.44)  |
| Myocardial infarction | 2.6 (2.1,3.1)           | 2.6 (2.1,3.2)   | 0.0 (-0.8,0.7)         | 0.99 (0.73,1.34)  |
| Bleeding              | 3.7 (3.3,4.2)           | 4.4 (3.5,5.2)   | -0.6 (-1.6,0.3)        | 0.86 (0.67,1.09)  |

Table S10. Sensitivity 5 - Estimated 180-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, including individuals missing GFR, and imputing median GFR and weight, using inverse probability weighting models (N=6,172).

| Outcome               | <b>Risk (%, 95% CI)</b> |                | <b>Risk difference</b> | <b>Risk ratio</b> |
|-----------------------|-------------------------|----------------|------------------------|-------------------|
|                       | Bivalirudin             | Heparin        | (95% CI)               | (95% CI)          |
| Composite             | 9.3 (8.4,10.1)          | 9.4 (8.2,10.6) | -0.2 (-1.7,1.3)        | 0.98 (0.83,1.16)  |
| Death                 | 4.2 (3.6,4.8)           | 3.6 (2.7,4.4)  | 0.6 (-0.4,1.7)         | 1.18 (0.89,1.57)  |
| Myocardial infarction | 2.8 (2.2,3.5)           | 2.5 (2.0,3.0)  | 0.3 (-0.4,1.1)         | 1.14 (0.84,1.55)  |
| Bleeding              | 3.1 (2.6,3.6)           | 4.0 (3.2,4.8)  | -0.9 (-1.9,0.0)        | 0.77 (0.58,1.01)  |

Table S11. Sensitivity 6 - Estimated 180-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, restricting to those that were given only bivalirudin or heparin, using inverse probability weighting models (N=2,886).

| Outcome               | <b>Risk (%, 95% CI)</b> |                | <b>Risk difference</b> | <b>Risk ratio</b> |
|-----------------------|-------------------------|----------------|------------------------|-------------------|
|                       | Bivalirudin             | Heparin        | (95% CI)               | (95% CI)          |
| Composite             | 7.3 (4.7,9.9)           | 9.7 (8.7,10.8) | -2.4 (-5.2, 0.4)       | 0.75 (0.51,1.12)  |
| Death                 | 4.4 (2.1,6.7)           | 3.0 (2.4,3.7)  | 1.4 (-1.1, 3.8)        | 1.45 (0.76,2.77)  |
| Myocardial infarction | 1.6 (0.5,2.7)           | 3.1 (2.5,3.6)  | -1.5 (-2.6,-0.3)       | 0.52 (0.22,1.23)  |
| Bleeding              | 2.6 (0.9,4.3)           | 4.4 (3.6,5.1)  | -1.7 (-3.6, 0.1)       | 0.60 (0.27,1.34)  |

Table S12. Sensitivity 7 - Estimated 180-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, with expanded definition of the bleeding outcome, using inverse probability weighting models (N=4,940).

| Outcome   | <b>Risk (%, 95% CI)</b> |                 | <b>Risk difference</b> | Risk ratio       |
|-----------|-------------------------|-----------------|------------------------|------------------|
|           | Bivalirudin             | Heparin         | (95% CI)               | (95% CI)         |
| Composite | 9.5 (8.4,10.6)          | 10.2 (8.8,11.5) | -0.7 (-2.4,1.1)        | 0.94 (0.78,1.12) |
| Bleeding  | 3.5 (2.8,4.2)           | 4.8 (3.8,5.7)   | -1.3 (-2.5,-0.1)       | 0.74 (0.54,0.99) |

Table S13. Sensitivity 8 - Estimated 180-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, with 5 year look back period for confounder identified in the inpatient and outpatient register, using inverse probability weighting models (N=4,940).

| Outcome               | <b>Risk (%, 95% CI)</b> |                 | <b>Risk difference</b> | <b>Risk ratio</b> |
|-----------------------|-------------------------|-----------------|------------------------|-------------------|
|                       | Bivalirudin             | Heparin         | (95% CI)               | (95% CI)          |
| Composite             | 9.3 (8.2,10.4)          | 10.0 (8.7,11.3) | -0.7 (-2.5,1.0)        | 0.93 (0.77,1.12)  |
| Death                 | 4.1 (3.4,4.8)           | 3.4 (2.5,4.2)   | 0.7 (-0.4,1.8)         | 1.21 (0.89,1.66)  |
| Myocardial infarction | 3.0 (2.3,3.8)           | 2.8 (2.2,3.4)   | 0.2 (-0.7,1.2)         | 1.08 (0.77,1.52)  |
| Bleeding              | 3.2 (2.5,3.8)           | 4.6 (3.7,5.6)   | -1.4 (-2.6,-0.3)       | 0.69 (0.50,0.94)  |

Table S14. Sensitivity 9 - Estimated 180-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, using a complete case analysis, using inverse probability weighting models (N=3,704).

| Outcome               | <b>Risk (%, 95% CI)</b> |                 | <b>Risk difference</b> | <b>Risk ratio</b> |
|-----------------------|-------------------------|-----------------|------------------------|-------------------|
|                       | Bivalirudin             | Heparin         | (95% CI)               | (95% CI)          |
| Composite             | 9.2 (8.0, 10.4)         | 9.2 (7.9, 10.6) | 0.0 (-1.9, 1.8)        | 1.00 (0.81, 1.22) |
| Death                 | 3.6 (2.9, 4.3)          | 2.4 (1.6, 3.2)  | 1.2 (0.1, 2.2)         | 1.49 (1.00, 2.23) |
| Myocardial infarction | 3.1 (2.3, 4.0)          | 2.4 (1.8, 3.0)  | 0.7 (-0.4, 1.8)        | 1.29 (0.85, 1.96) |
| Bleeding              | 3.3 (2.6, 4.1)          | 4.8 (3.6, 6.0)  | -1.5 (-2.9, 0.0)       | 0.70 (0.49, 1.00) |

Table S15. Sensitivity 10 - Estimated 180-day risk, risk difference, and risk ratios from the observational emulation\* of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014, censoring at death, using inverse probability weighting models (N=4.940).

| Total additional and a control of the probability weighting models (11-4,940). |                         |                |                        |                   |  |  |
|--|-------------------------|----------------|------------------------|-------------------|--|--|
| Outcome  | <b>Risk (%, 95% CI)</b> |                | <b>Risk difference</b> | <b>Risk ratio</b> |  |  |
|  | Bivalirudin             | Heparin        | (95% CI)               | (95% CI)          |  |  |
| Myocardial infarction  | 3.1 (2.3, 3.9)          | 2.9 (2.2, 3.5) | 0.3 (-0.8, 1.3)        | 1.09 (0.76, 1.55) |  |  |
| Bleeding   | 3.3 (2.6, 4.0)          | 4.7 (3.7, 5.7) | -1.4 (-2.7, -0.2)      | 0.69 (0.51, 0.95) |  |  |



Figure S1. Unweighted survival curves from an observational emulation a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014.



Figure S2. Sensitivity 4 - Inverse-probability weighted survival curves from an observational emulation of a target trial of bivalirudin vs. heparin, SWEDEHEART Register, 2012-2014\*, including all individuals given heparin before percutaneous coronary intervention (N=7,652).

\*Adjusted at baseline for severity of myocardial infarction (STEMI/NSTEMI, Killip class, angiography finding), center where percutaneous coronary intervention took place (university hospital/non-university hospital), demographics (sex, age), lifestyle characteristics (weight, smoking), laboratory measurements (GFR, heart rate, systolic blood pressure, diastolic blood pressure), diagnoses (diabetes, severe bleeding, anemia), medications (warfarin, anti-hypertensives, lipid lowering treatment), and prior cardiovascular disease or cardiovascular procedures (myocardial infarction, percutaneous coronary intervention, coronary artery bypass grafting)