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Prediagnosis Leisure-Time Physical Activity and Lung Cancer Survival: A Pooled Analysis of 11 Cohorts

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Abstract

Background: Little is known about the association between physical activity before cancer diagnosis and survival among lung cancer patients. In this pooled analysis of 11 prospective cohorts, we investigated associations of prediagnosis leisure-time physical activity (LTPA) with all-cause and lung cancer–specific mortality among incident lung cancer patients. **Methods:** Using self-reported data on regular engagement in exercise and sports activities collected at study enrollment, we assessed metabolic equivalent hours (MET-h) of prediagnosis LTPA per week. According to the Physical Activity Guidelines for Americans, prediagnosis LTPA was classified into inactivity, less than 8.3 and at least 8.3 MET-h per week (the minimum recommended range). Cox regression was used to estimate hazard ratios (HRs) and 95% confidence interval (CIs) for all-cause and lung cancer–specific mortality after adjustment for major prognostic factors and lifetime smoking history. **Results:** Of 20 494 incident lung cancer patients, 16 864 died, including 13 596 deaths from lung cancer (overall 5-year relative survival rate = 20.9%, 95% CI = 20.3% to 21.5%). Compared with inactivity, prediagnosis LTPA of more than 8.3 MET-h per week was associated with a lower hazard of all-cause mortality (multivariable-adjusted HR = 0.93, 95% CI = 0.88 to 0.99), but not with lung cancer–specific mortality (multivariable-adjusted HR = 0.95, 055 to 1.04), among the overall population. Additive interaction was found by tumor stage (P_{interaction} = .008 for all-cause mortality and .003 for lung cancer–specific mortality).

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When restricted to localized cancer, prediagnosis LTPA of at least 8.3 MET-h per week linked to 20% lower mortality: multivariableadjusted HRs were 0.80 (95% CI = 0.67 to 0.97) for all-cause mortality and 0.80 (95% CI = 0.65 to 0.99) for lung cancer–specific mortality. **Conclusions:** Regular participation in LTPA that met or exceeded the minimum Physical Activity Guidelines was associated with reduced hazards of mortality among lung cancer patients, especially those with early stage cancer.

Lung cancer is the most common cancer in the world and accounts for approximately 2.09 million new cases and 1.76 million deaths in 2018 (1). Despite the recent advances in lung cancer screening and treatments, more than half of newly diagnosed patients die within a year of diagnosis; the overall 5year survival rate of lung cancer remains under 20% worldwide (2,3). To reduce the global burdens of lung cancer, it is crucial to identify potential risk and prognostic factors apart from smoking cessation.

Physical activity (PA) has attracted great attention in cancer research because of its benefits in reducing inflammation, regulating hormones (eg, insulin), and improving immune function and energy balance (4-6). Epidemiological evidence to date supports a link of PA to cancer prevention and survival (7-11). The Physical Activity Guidelines Advisory Committee and the American College of Sports Medicine Roundtable have recently concluded that PA prevents at least 7 cancers (eg, breast and colon) and might confer survival benefits to patients with breast, colon, and prostate cancers (8,9,11,12). Yet, evidence on PA and lung cancer remains moderate or limited, especially for survival outcomes (ie, all-cause and lung cancer-specific mortality) (8,9,11,12). Although a recent meta-analysis and some prospective studies have shown a statistically significant reduction in lung cancer mortality attributed to prediagnosis PA (7,13-15), most of these studies had a small sample size. They are also limited by residual confounding because of smoking and lack of consideration of major prognostic factors and are unable to address subgroup variations. Large-scale, population-based prospective investigations that overcome previous limitations are needed to fill a research gap and provide convincing evidence.

To this end, this pooled analysis of 11 cohorts from the United States (US), Europe, and Asia aims to investigate associations of prediagnosis leisure-time PA (LTPA) with all-cause and lung cancer-specific mortality among more than 20000 incident primary lung cancer patients. Given the substantial difference in survivorship across lung cancer stage (2,3), analyses were conducted in the overall study population and subgroups defined by tumor stage. Furthermore, we assessed effect modification by established prognostic factors (ie, stage, histology, and grade), time interval from LTPA assessment to cancer diagnosis (given possible measurement errors or biologically relevant time windows), lifetime smoking history, and other risk factors.

Methods

Study Populations

We harmonized de-identified, individual participant data from 11 cohorts (16,17), including 7 US cohorts (National Institute of Health–American Association of Retired Persons Diet and Health Study; Health Professionals Follow-up Study; Nurses' Health Study; Iowa Women's Health Study; Prostate, Lung, Colorectal, and Ovarian Cancer Screening Trial; Southern Community Cohort Study; and VITamins And Lifestyle Study), 2 European cohorts (European Prospective Investigation into Cancer and Nutrition and Trøndelag Health Study), and 2 Asian cohorts (Shanghai Men's Health Study and Shanghai Women's Health Study). All studies were approved by the institutional review boards and ethics committees of the hosting institutes.

Among 1588 378 initial participants, we identified 22762 incident lung cancer patients diagnosed after study enrollment. Of those, we excluded individuals who had no data on LTPA (n = 2031), smoking history (n = 183), and survival time (n = 45). Cancer in situ was also excluded (n = 9), thus leaving 20494 patients. Characteristics of each participating cohort and our analytic sample are summarized in Table 1.

Assessment and Parameterization of LTPA Exposure

This study used LTPA data only, not incorporating other domains of PA. Prediagnosis LTPA was assessed at baseline using validated cohort-specific questionnaires asking about regular engagement in exercise and sports activities (18-26). Details of LTPA assessment in each cohort, including original questions, intensity, and exposure windows, are shown in Supplementary Table 1 (available online). The level of LTPA was quantified in metabolic equivalent hours per week (MET-h/ week), using the Compendium of Physical Activities (27). Based on the Physical Activity Guidelines for Americans (12,28,29), prediagnosis LTPA was classified into inactive, low active (>0 to <8.3 MET-h/week), moderately active (the recommended level for health benefits: 8.3-16.0 MET-h/week, equivalent to 150-300 minutes of moderate or 75-150 minutes of vigorous intensity activity per week), and highly active (>16.0 MET-h/week). Because of the limited number of study participants in the highly active group, this group was finally combined with the moderately active group, referred to as met or exceeded the minimum recommendation (\geq 8.3 MET-h/week).

Assessment of Outcome

Each cohort has followed-up cancer incidence and mortality via linkages to national and regional registries, follow-up surveys, medical record reviews, or a combination of these methods. Incident lung cancer patients were ascertained using the *International Classification of Diseases* 9th or 10th revision (162 and C34, respectively). All patients were subclassified by histological type (adenocarcinoma, squamous cell carcinoma, other nonsmall cell carcinoma, small cell carcinoma, or unspecified or unknown), stage (localized [stage I/II], regional [stage III], distant [stage IV], and unknown), and grade (well, moderately, and poorly differentiated; undifferentiated; and unknown). For deceased ones, we obtained information on the underlying cause and date of death. Survival time was calculated as total years from the date of lung cancer diagnosis to the date of death, loss-to-follow-up, or end of follow-up, whichever occurred first.

Covariates

Potential confounders were selected a priori based on literature review and risk factors found in our study populations (16,17). Included were age at diagnosis (continuous), sex, race and

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^ancluding primary lung cancer patients who were eligible for the current pooled analysis. AARP = National Institute of Health-American Association of Retired Persons Diet and Health Study; CI = confidence interval; EPIC = European Prospective Investigation into Cancer & Nutrition; HPFS = Health Professionals Follow-up Study; HUNT = Trøndelag Health Study; IQR = interquartile range, IWHS = Iowa Women's Health Study; LTPA = leisure-time physical activity; MET-h/week = metabolic equivalent hours per week; NHS = Nurses' Health Study; NSCC = non-small cell carcinoma; PLCO = Prostate, Lung, Colorectal and Ovarian Cancer Screening Trial; SCCS = Southern Community Cohort Study; VITAL = VITamins And Lifestyle Study; SMHS = Shanghai Men's Health Study; SWHS = Shanghai Women's Health Study.

 $^{\mathrm{b}}\mathrm{Number}$ of deaths from all causes (deaths from lung cancer).

^{cy} ears from leisure-time physical activity assessment to lung cancer diagnosis. ^d percentage of adherence to the minimum recommended range of the Physical Activity Guidelines at least 500 MET-min (8.3 MET-h) per week. ^eIncluding current and former smokers.

ethnicity (Asian, Black, Other [Hispanic and Latino, American Indian, and other racial or ethnic group] and White), smoking status (never, former, current), smoking pack-years (continuous), education (less than high school, high school graduation, vocational education, college, university or higher), alcohol consumption (none, moderate drinking up to 1 and 2 drinks per day, heavy drinking >1 and >2 drinks per day for women and men, respectively; 1 drink = 14 grams of ethanol), history of diabetes (yes, no), body mass index (BMI; <18.5, 18.5-24.9, 25.0-29.9, and \geq 30.0 kg/m²), and hormone therapy in women (yes, no)—all of which were assessed at LTPA assessment. Missing covariates were independently imputed by cohort (16). Established prognostic factors were further included: tumor stage (localized, regional, distant, unknown), histological type (adenocarcinoma, squamous cell carcinoma, other non-small cell carcinoma, small cell carcinoma, unspecified or unknown), and grade (well, moderately, and poorly differentiated; undifferentiated; unknown).

Statistical Analysis

Using the life-table method and log-rank test, we assessed 5year relative survival rates by baseline characteristics and clinical features of cancer. Cox proportional hazard regression was used to estimate multivariable-adjusted hazard ratios (HRs) and 95% confidence intervals (CIs) for all-cause and lung cancerspecific mortality, using inactivity as the reference. The global goodness-of-fit test with Schoenfeld residuals found no violation of the proportional hazard assumption. For lung cancerspecific mortality, death from other causes was treated as a competing risk. Given inter- and intrastudy variability, Cox models were stratified by cohort, calendar years of lung cancer diagnosis, and time interval between LTPA assessment and diagnosis. A random-effects meta-analysis was complimented with I^2 and $P_{heterogeneity}$ to offset potential concerns of residual heterogeneity.

Stratified analyses were conducted by major prognostic and risk factors; to avoid reverse causality due to high fatality and short survival time of distant stage cancer, these analyses were restricted to early stage lung cancer. Additive interactions were evaluated by the relative excess risk due to interaction (30,31), referring to the excess risk from interaction between prediagnosis LTPA and stratification variables as compared with baseline risk without exposure. P values were corrected for multiple comparisons by controlling the false-discovery rate. A series of sensitivity analyses were conducted using another LTPA categorization (cohort- and sex-specific quartiles and common quartiles across total participants), excluding participants with long time intervals between LTPA assessment and cancer diagnosis (over median years), further adjusting for dietary calcium intake statistically significantly associated with lung cancer survival in our populations (16), and excluding 1 cohort at a time from analyses. All procedures were performed using SAS 9.4 (SAS Institute, Inc, Cary, NC), and 2-sided P-values less than .05 were considered statistically significant.

Results

Of the 20 494 incident lung cancer patients, 16 864 patients died, including 13 596 deaths from lung cancer (Table 1); the median survival time was 0.9 years (interquartile range = 0.3-2.7). The mean age at lung cancer diagnosis was 68.5 years. Most patients were ever-smokers (ranging from 82.3% to 94.8% across studies), except for those from the Shanghai Women's Health Study. About 43% of patients met the minimum recommendation before lung cancer diagnosis. The overall 5-year survival rate was 20.9% (95% CI = 20.3% to 21.5%).

Five-year relative survival rates were higher among patients with high educational attainment, noncurrent smokers, smokers with less than 30 pack-years, nondiabetic patients, and women taking hormone therapy (all P < .05); those patients also showed higher proportions of meeting the guidelines. When stratified by histological type, stage, and grade, we observed much lower 5-year survival rates for small cell carcinoma (10.0%), distant stage carcinoma (5.9%), and undifferentiated lung carcinoma (11.8%) (Table 2).

Compared with inactivity (Table 3), LTPA of 8.3 MET-h/week or more before cancer diagnosis was associated with a lower hazard of all-cause mortality among the overall population (HR = 0.93, 95% CI = 0.88 to 0.99), but not with lung cancer–specific mortality (HR = 0.99, 95% CI = 0.95 to 1.04), after adjustment for all potential covariates. No heterogeneity was observed across cohorts (P_{heterogeneity}=.41 for all-cause and .28 for lung cancer–specific mortality), with comparable HRs from random-effects meta-analyses of 0.95 (95% CI = 0.90 to 1.00) and 1.00 (95% CI = 0.95 to 1.05), respectively (Supplementary Figures 1

Table 2. Leisure-time physical activity and 5-year survival rates among lung cancer patients by baseline characteristics

| | | | Moot the | Modian LTPA | 5-year surviv | val |
|---------------------|--------------|---------------|-----------------------------|------------------|---------------------|------------------|
| Characteristics | No. of cases | No. of deaths | guideline, (%) ^a | (IQR) MET-h/week | Rate (95% CI), % | P^{b} |
| Age at diagnosis, y | | | | | | |
| <70 | 10718 | 8659 | 41.4 | 4.5 (0.3-15.0) | 22.3 (21.5 to 23.1) | <.001 |
| ≥70 | 9776 | 8205 | 44.0 | 4.9 (0.9-15.0) | 18.9 (18.2 to 19.7) | |
| Sex | | | | . , | . , | |
| Men | 11 010 | 9384 | 45.2 | 4.5 (1.0-15.0) | 17.8 (17.1 to 18.6) | <.001 |
| Women | 9484 | 7480 | 39.7 | 4.5 (0.4-15.0) | 24.1 (23.2 to 25.0) | |
| Race and ethnicity | | | | . , | . , | |
| Asian | 1852 | 1356 | 29.7 | 0.0 (0.0-10.5) | 21.3 (19.4 to 23.3) | .02 |
| Black | 954 | 716 | 25.3 | 0.3 (0.0-9.2) | 21.8 (19.1 to 24.5) | |
| Other | 219 | 186 | 32.0 | 2.2 (0.3-10.5) | 16.9 (12.2 to 22.2) | |
| White | 17469 | 14606 | 45.1 | 5.2 (1.5-15.0) | 20.7 (20.1 to 21.3) | |
| Education | | | | . , | . , | |
| \leq High school | 8552 | 7091 | 39.6 | 4.5 (0.3-15.0) | 18.4 (17.5 to 19.2) | <.001 |

(continued)

Table 2. (continued)

| | | | Moot the | Modian LTPA | 5-year surviv | val |
|---|--------------|---------------|-----------------------------|------------------------------------|--|----------------|
| Characteristics | No. of cases | No. of deaths | guideline, (%) ^a | (IQR) MET-h/week | Rate (95% CI), % | P ^b |
| Vocational school/some college | 6662 | 5441 | 41.5 | 4.5 (1.5-12.2) | 21.7 (20.7 to 22.7) | |
| ≥University graduation | 5280 | 4332 | 49.0 | 7.7 (1.5-15.0) | 23.2 (22.1 to 24.4) | |
| Smoking status | | | | | | |
| Never | 2457 | 1760 | 46.0 | 4.8 (0.2-15.0) | 29.4 (27.6 to 31.3) | <.001 |
| Former | 8233 | 6875 | 48.0 | 6.8 (1.5-15.0) | 21.2 (20.3 to 22.1) | |
| Current | 9804 | 8229 | 37.8 | 4.5 (0.3-13.5) | 18.2 (17.5 to 19.0) | |
| Smoking pack-years in smokers | | | | | | |
| <30 | 5647 | 4536 | 48.4 | 7.5 (1.5-15.0) | 21.1 (20.1 to 22.3) | <.001 |
| 30-49 | 6324 | 5324 | 42.4 | 4.5 (0.8-15.0) | 19.5 (18.5 to 20.5) | |
| ≥50 | 6066 | 5244 | 37.0 | 4.5 (0.3-10.5) | 18.3 (17.3 to 19.3) | |
| Alcohol consumption ^c | | | | | | |
| None | 6216 | 5051 | 34.8 | 4.0 (0.0-10.5) | 20.0 (19.0 to 21.1) | .004 |
| Moderate | 10 031 | 8255 | 46.6 | 6.0 (1.5-15.0) | 21.7 (20.9 to 22.5) | |
| Heavy | 4247 | 3558 | 44.9 | 4.5 (0.9-15.0) | 19.4 (18.2 to 20.6) | |
| Body mass index, kg/m ² | | | | · , | · · · · | |
| <18.5 | 432 | 357 | 32.2 | 2.3 (0.0-10.5) | 17.9 (14.4 to 21.8) | .29 |
| 18.5-24.99 | 8836 | 7227 | 44.5 | 5.1 (0.7-15.0) | 21.2 (20.3 to 22.0) | |
| 25.0-29.99 | 8048 | 6664 | 44.1 | 4.5 (1.3-15.0) | 20.3 (19.5 to 21.2) | |
| >30.0 | 3178 | 2616 | 35.2 | 4.5 (0.3-10.5) | 20.8 (19.4 to 22.2) | |
| History of diabetes | | | | · · · · | x y | |
| No | 18931 | 15 493 | 43.2 | 4.5 (0.7-15.0) | 21.2 (20.6 to 21.8) | <.001 |
| Yes | 1563 | 1371 | 35.5 | 4.5 (0.3-10.5) | , 14.8 (13.1 to 16.6) | |
| Hormone therapy in women | | | | | | |
| No | 5548 | 4420 | 38.2 | 4.5 (0.3-14.9) | 22.8 (21.7 to 24.0) | <.001 |
| Yes | 3936 | 3060 | 41.8 | 4.5 (0.9-15.0) | 26.0 (24.6 to 27.4) | |
| Histological type | | | | (| | |
| Adenocarcinoma | 7543 | 5753 | 43.5 | 4.5 (0.7-15.0) | 27.0 (26.0 to 28.0) | <.001 |
| Squamous cell carcinoma | 3591 | 2874 | 42.4 | 4 5 (0 6-14 9) | 24 8 (23 4 to 26 3) | |
| Other non-small cell carcinoma | 3224 | 2673 | 42.5 | 4.5 (1.1-14.8) | 20.5 (19.1 to 22.0) | |
| Small cell carcinoma | 2723 | 2500 | 40.7 | 4 5 (0 5-12 3) | 10.0 (8.9 to 11.2) | |
| Unspecified | 3413 | 3064 | 42.8 | 4 5 (0 3-15 0) | 11 8 (10 7 to 12 9) | |
| Tumor stage ^d | 5115 | 5001 | 12.0 | 1.5 (0.5 15.0) | 11.0 (10.7 to 12.3) | |
| Localized | 2776 | 1482 | 40.4 | 4 5 (0 3-12 5) | 54 3 (52 3 to 56 2) | < 001 |
| Regional | 3347 | 2776 | 42.4 | 4 5 (0 3-14 5) | 21.0 (19.6 to 22.4) | 2.001 |
| Distant | 5566 | 5206 | 38.8 | 4.5 (0.3-10.5) | 59(53to 65) | |
| Unknown | 8805 | 7400 | 45.9 | 6.1 (1.5-15.0) | 20 2 (19 3 to 21 0) | |
| Tumor grade | 0005 | 7 100 | 15.5 | 0.1 (1.5 15.0) | 20.2 (19.5 to 21.0) | |
| Well differentiated | 710 | 102 | 40.1 | 1 5 (1 5 10 5) | 16 1 (12 2 to 10 9) | < 001 |
| Moderately differentiated | 2183 | 1521 | 41.6 | 4.5 (1.5-10.5) | 40.1 (42.3 to 49.8) | <.001 |
| Roorly differentiated | 2105 | 2075 | 40.2 | 4.5 (0.7 10.5) | $22.5(20.4 \pm 0.220)$ | |
| Indifferentiated | 1206 | 1200 | 40.5 | 4.5 (0.7 - 10.5) | 22.3 (21.2 to 23.9) | |
| Unknown | 12 500 | 10645 | 33.0 AA 0 | 4.2 (0.0-10.2) | 16 7 (16 0 to 17 2) | |
| From baseline to diagnosis w ^e | 17 272 | 10 040 | '1'1 .0 | 4.0 (0.4-15.0) | 10.7 (10.0 to 17.3) | |
| | 7615 | 6706 | 40.1 | | 10 0 /10 0 +~ 20 7) | 02 |
| < <u>5</u> | 7010 | 6040 | 40.1 12 2 | 4.5 (0.3-10.5) | 13.0 (10.3 LU 2U./) 20.6 (10.9 to 21.7) | .02 |
| S-5 ∖10 | 4220 | 0742 | 45.0 | +.3 (0.7 - 14.0) 7 0 (1 E 10 1) | 20.0 (12.0 LU 21.3) | |
| <u></u> | 4320 | 5210 | 43.9 | /.2 (1.3-18.1) | 22.0 (21.3 (0 23.9) | |

^aPercentage of adherence to the recommended physical activity guidelines, \geq at least 500 MET-minutes (8.3 MET-hours) per week. CI = confidence interval; IQR = interquartile range; LTPA = leisure-time physical activity; MET-h/week = metabolic-equivalent hours per week.

^bStatistical differences across survival rates (2-sided P values) were estimated by the log-rank test and corrected for multiple comparisons by controlling the false-discovery rate.

 $^{\circ}$ Moderate defined as >0 to \leq 1 (women) or >0 to \leq 2 (men) drinks per day and heavy defined as >1 (women) or >2 (men) drinks per day.

^dLocalized, regional, and distant stages included stage I and II, stage III, and stage IV, respectively.

^eTime interval from physical activity assessment to lung cancer diagnosis.

and 2, available online). Clear evidence of additive interaction was found by stage ($P_{\rm interaction} = .008$ for all-cause mortality and .003 for lung cancer–specific mortality). For localized lung cancer, prediagnosis LTPA of at least 8.3 MET-h/week was

associated with 20% lower mortality (HR = 0.80, 95% CI = 0.67 to 0.97, for all-causes, and HR = 0.80, 95% CI = 0.65 to 0.99, for lung cancer). Meanwhile, overall associations were weaker for regional stage lung cancer (HR = 0.89, 95% CI = 0.77 to 1.02, and

| | | Deaths from all c | causes | | | Deaths from lung o | cancer ^c | |
|---|--|---|---|-------------------------------------|---|---|--|---------------------------|
| Tumor stage | None | >0 to <8.3 | >8.3 | $P_{\rm trend}^{\rm d}$ | None | >0 to <8.3 | >8.3 | $P_{\rm trend}^{\rm d}$ |
| Total cases | | | | | | | | |
| Deaths/cases, No. | 2380/3008 | 7297/8749 | 7187/8737 | | 1965/2822 | 5764/8267 | 5867/8328 | |
| 5-year survival rate (95% CI), % | 18.1 (16.7 to 19.6) | 21.1 (20.2 to 22.0) | 21.2 (20.3 to 22.1) | | 23.7 (22.0 to 25.4) | 28.5 (27.5 to 29.5) | 27.4 (26.4 to 28.4) | |
| Hazard ratio (95% CI) ^e | 1 (Referent) | 0.96 (0.91 to 1.01) | 0.92 (0.87 to 0.97) | .002 | 1 (Referent) | 0.99 (0.94 to 1.04) | 0.98 (0.93 to 1.02) | .29 |
| Hazard ratio (95% CI) ^f | 1 (Referent) | 0.97 (0.91 to 1.02) | 0.93 (0.88 to 0.99) | .01 | 1 (Referent) | 1.00 (0.96 to 1.05) | 0.99 (0.95 to 1.04) | .63 |
| Localized lung cancer cases | | | | | | | | |
| Deaths/cases, No. | 226/485 | 654/1170 | 602/1121 | | 165/473 | 391/1152 | 405/1097 | |
| 5-year survival rate (95% CI), % | 52.9 (47.9 to 57.6) | 54.3 (51.3 to 57.3) | 54.8 (51.7 to 57.8) | | 60.1 (55.0 to 64.9) | 66.2 (63.2 to 69.1) | 64.3 (61.2 to 67.2) | |
| Hazard ratio (95% CI) ^e | 1 (Referent) | 0.92 (0.77 to 1.10) | 0.79 (0.66 to 0.95) | 600. | 1 (Referent) | 0.84 (0.68 to 1.03) | 0.80 (0.65 to 0.99) | .42 |
| Hazard ratio (95% CI) ^f | 1 (Referent) | 0.93 (0.78 to 1.12) | 0.80 (0.67 to 0.97) | .02 | 1 (Referent) | 0.84 (0.68 to 1.04) | 0.80 (0.65 to 0.99) | .46 |
| Regional lung cancer cases | | | | | | | | |
| Deaths/cases, No. | 397/510 | 1176/1419 | 1203/1418 | | 332/480 | 945/1377 | 996/1353 | |
| 5-year survival rate (95% CI), % | 20.4 (16.9 to 24.1) | 22.3 (20.1 to 24.5) | 19.8 (17.8 to 22.0) | | 25.6 (21.5 to 29.8) | 29.5 (27.1 to 32.1) | 25.4 (23.0 to 27.8) | |
| Hazard ratio (95% CI) ^e | 1 (Referent) | 0.94 (0.82 to 1.08) | 0.87 (0.76 to 1.00) | .06 | 1 (Referent) | 0.98 (0.86 to 1.11) | 0.93 (0.83 to 1.05) | .42 |
| Hazard ratio (95% CI) ^f | 1 (Referent) | 0.94 (0.82 to 1.09) | 0.89 (0.77 to 1.02) | .17 | 1 (Referent) | 0.98 (0.86 to 1.11) | 0.95 (0.84 to 1.07) | .51 |
| Distant lung cancer cases | | | | | | | | |
| Deaths/cases, No. | 995/1093 | 2174/2315 | 2037/2158 | | 885/1048 | 1900/2252 | 1815/2113 | |
| 5-year survival rate (95% CI), % | 6.3 (4.9 to 7.8) | 5.6 (4.7 to 6.6) | 6.0 (5.1 to 7.1) | | 10.0 (8.2 to 12.0) | 10.7 (9.4 to 12.1) | 10.4 (9.1 to 11.8) | |
| Hazard ratio (95% CI) ^e | 1 (Referent) | 0.98 (0.89 to 1.07) | 0.97 (0.89 to 1.06) | .61 | 1 (Referent) | 1.03 (0.97 to 1.10) | 1.04 (0.98 to 1.10) | .50 |
| Hazard ratio (95% CI) ^f | 1 (Referent) | 0.98 (0.90 to 1.08) | 0.98 (0.89 to 1.07) | .73 | 1 (Referent) | 1.03 (0.97 to 1.10) | 1.04 (0.98 to 1.11) | .53 |
| ^a Localized, regional, and distant stages in tality. CI = confidence interval; MET-h/we b_≤00 MET-min/wk (≥8.3 MET-h/week) w | cluded stage I and II, stage ek = metabolic-equivalent ras the level recommender | III, and stage IV, respectivel hours per week. d for substantial health be: | y. Additive interactions we nefits based on the physic | re statistically al active guide | significant: P interaction = . lines such as the World H | .008 for all-cause mortality e ealth Organization Global I | and .003 for lung cancer-sp Recommendations and 201 | ecific mor- 8 Physical |
| ACUVITY GUIDEN. | | | | | | | | |

Table 3. Association of prediagnosis leisure-time physical activity with all-cause and lung cancer-specific mortality among lung cancer patients by tumor stage^a

^c Patients missing cause of death were excluded from the analysis; death from other causes was treated as a competing risk.

^dCorrected for multiple comparisons by controlling the false-discovery rate.

Adjusted for age at diagnosis, sex, smoking status, smoking pack-years, race and ethnicity, education, alcohol consumption, history of diabetes, body mass index levels, hormone therapy in women, and histological type and *Adjusted for age at diagnosis, sex, smoking status, and smoking pack-years and stratified by cohort, year of lung cancer diagnosis, and time interval from leisure-time physical activity assessment to lung cancer diagnosis. grade of lung cancer and stratified by cohort, year of lung cancer diagnosis, and time interval from leisure-time physical activity assessment to lung cancer diagnosis.

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HR = 0.95, 95% CI = 0.84 to 1.07, respectively), and null associations were found for distant stage (Table 3). Results from random-effects meta-analyses yielded similar results as those presented, with about 20% lower mortality among localized cases who adhered to prediagnosis LTPA of at least 8.3 MET-h/ week (HR = 0.80, 95% CI = 0.66 to 0.98, for all-cause mortality; $P_{heterogeneity} = .48$; and HR = 0.76, 95% CI = 0.61 to 0.94, for lung cancer-specific mortality; $P_{heterogeneity} = .50$; Supplementary Figures 3 and 4, available online). Exclusion of any cohort from the main analysis one at a time had limited impacts on the above-reported associations (Supplementary Table 2, available online).

In stratified analyses of localized and regional-stage lung cancer (Figures 1 and 2), we observed additive interactions when jointly considering prediagnosis LTPA and histological type ($P_{interaction} = .04$ for all-cause mortality and .003 for lung cancer-specific mortality). Prediagnosis LTPA of at least 8.3 MET-h/week was associated with 21% lower all-cause mortality for adenocarcinoma (HR = 0.79, 95% CI = 0.66 to 0.95). The overall association pattern remained consistent across the potential prognostic and risk factors (eg, race and ethnicity and BMI; all $P_{\text{interaction}} > .05$), but the magnitude of the associations attenuated when LTPA assessment was far from cancer diagnosis. Notably, never-smokers also appeared to have survival benefits from prediagnosis LTPA of at least 8.3 MET-h/week (HR = 0.76, 95% CI = 0.55 to 1.06, for all-cause mortality, and HR = 0.79, 95% CI=0.58 to 1.07, for lung cancer-specific mortality), but the point estimates failed to reach statistical significance because of the small sample size.

A series of sensitivity analyses showed a similar pattern of the associations (data not shown). Further analyses separating the moderately active (8.3-16.0 MET-h/week) and the highly active (\geq 16.0 MET-h/week) groups showed little evidence that the latter was more strongly related to lung cancer survival; however, the risk estimate was unstable because of its insufficient sample size (data not shown).

Discussion

In this pooled analysis of 20494 incident lung cancer patients from 11 prospective cohorts, regular participation in LTPA prior to cancer diagnosis, particularly when meeting or exceeding the minimum Physical Activity Guidelines, was associated with reduced hazards of mortality among lung cancer patients. Notably, patients diagnosed with localized lung cancer showed approximately 20% lower all-cause and lung cancer-specific mortality when engaging in LTPA of at least 8.3 MET-h/week before diagnosis compared with inactivity. A statistically significant additive interaction by stage was suggested. Our findings support a possible long-term benefit of habitual LTPA adhering to the Physical Activity Guidelines. If confirmed, pretreatment LTPA could be proposed as a possible stratification factor for future therapeutic trials, at least for early stage lung cancer patients.

Currently, epidemiological evidence on survival benefits attributed to prediagnosis LTPA remains limited among lung cancer patients (8,9,11,12). In a recent systematic review and metaanalysis of all existing epidemiologic studies and trials (7), the summary HR for lung cancer–specific mortality associated with higher levels of prediagnosis PA was 0.81 (95% CI = 0.75 to 0.87); no data was available for all-cause mortality. Despite adding an important piece of evidence, this finding was derived as a part of 5 studies accessing multiple cancer sites, not focusing only on lung cancer. Thus, overall numbers of lung cancer patients and deaths were limited, and lung cancer-specific prognostic and risk factors could not be properly considered. Several observational studies also showed the beneficial impact of prediagnosis PA on lung cancer mortality (13-15), in line with our findings. For example, the β -Carotene and Retinol Efficacy Trial, including 231 lung cancer cases and 141 deaths, found an inverse association of prediagnosis total PA with mortality only among women who smoked heavily (13). The Women's Health Initiative study, analyzing 2148 lung cancer patients and 1365 lung cancer deaths, reported 20%-32% lower lung-cancer mortality associated with at least 100 MET-minutes/week of exercise among postmenopausal women (14), with a dose-response association for adenocarcinoma only. Recently, a hospitalbased, case-control study (579 cases with 560 total deaths and 481 lung cancer-specific deaths) reported that lifetime recreational physical inactivity was associated with 30%-40% increased mortality (HR = 1.31, 95% CI = 1.09 to 1.58, for all-cause mortality, and HR = 1.40, 95% CI = 1.14 to 1.71, for lung cancerspecific mortality) (15). Consistent with our findings, this study observed a stronger inactivity-mortality association among early stage lung cancer cases; no association was found for distant stage. Similarly, some cohort studies showed that higher cardiorespiratory fitness levels prior to cancer diagnosis were linked to a lower risk of death among lung cancer patients (32-35). Nonetheless, most previous studies were limited to insufficient sample size, restricted to predominantly White populations, and lacked consideration of major prognostic factors, subgroup variations, and potential competing risks of death. Our large sample size, including 20494 incident lung cancer patients with different stage and histology and 16864 deaths from racially and ethnically diverse populations, detailed data on clinical features of lung cancer, and enhanced scientific rigor would overcome the previous limitations. Findings from this pooled analysis suggest that the association of prediagnosis LTPA with lung cancer survival could be modified by tumor stage and histological type. Black patients appeared to have lower HRs than others, although the test for interaction was not statistically significant. Also, the overall associations remained consistent across sex, smoking, BMI, and other factors potentially related to survival, adding a piece of epidemiologic evidence to a possible causal association.

PA has attracted much attention as a potential protective factor against cancer-related deaths, recurrence, or metastasis (4-6,36). During exercise, the body activates biological mechanisms which inhibit tumor growth, including modulation of carcinogenic factors (ie, inflammatory cytokines, insulin-like growth factors, and other hormones) and enhancement in immune function and metabolic health (6). Long-term habitual PA can lead to intratumoral adaptations, including improvements in blood perfusion, immunogenicity, and immune cell infiltration (4,5), which help inhibit cancer progression. Given these biological benefits, it is possible that PA before initiation of carcinogenesis may result in developing less aggressive tumors. Furthermore, PA plays a crucial role in enhancing drug tolerance and efficacy, alleviating treatment-related adverse effects, and reducing the likelihood of relapse and metastasis (4,36). Regarding lung-specific benefits from PA, evidence indicates that greater amounts of PA are associated with less lung function decline, better pulmonary functional capacity, and a lower risk of chronic obstructive pulmonary disease (37-39). Indeed, a recent systematic review of randomized-controlled trials has reported that presurgery exercise interventions could substantially improve physical and pulmonary functions among lung

| | 10.01 | ueatiis | Aujusteu | | |
|---------------------------------|-------|---------|------------------|--|---------------|
| ≥ 8.3 MET-hr/wk vs. none | ≥ 8.3 | None | HR (95% CI) | | P Interaction |
| Age at lung cancer diagnosis | | | | | |
| <70 years | 966 | 398 | 0.77 (0.66-0.89) | H III H | .71 |
| ≥70 years | 839 | 225 | 0.99 (0.83-1.18) | | |
| Sex | | | . , | I I | |
| Men | 1139 | 313 | 0.86 (0.74-1.00) | F E | .34 |
| Women | 666 | 310 | 0.87 (0.74-1.03) | | |
| Race and ethnicity | | | . , | | |
| Asian | 94 | 171 | 0.91 (0.67-1.22) | ⊢ | .97 |
| Black | 43 | 94 | 0.68 (0.44-1.06) | ►_ | |
| White | 1654 | 347 | 0.86 (0.75-0.98) | H E H | |
| Education | | | · · · · · · | | |
| ≤High school graduation | 847 | 431 | 0.94 (0.81-1.08) | H II -1 | .97 |
| >High school graduation | 958 | 192 | 0.76 (0.64-0.91) | H E H | |
| Smoking status | | | | | |
| Never | 180 | 98 | 0.76 (0.55-1.06) | | .71 |
| Former | 832 | 166 | 0.80 (0.66-0.97) | H H | |
| Current | 793 | 359 | 0.92 (0.79-1.07) | H I H | |
| Alcohol consumption | | | | | |
| None | 455 | 317 | 0.91 (0.77-1.09) | ► ■ -4 | .97 |
| Moderate | 937 | 181 | 0.86 (0.71-1.04) | F | |
| Heavy | 413 | 125 | 0.77 (0.60-0.98) | ⊢ ∎ | |
| Obesity status | | | | | |
| ≤Normal weight, <i>BMI</i> <25 | 769 | 306 | 0.83 (0.70-0.98) | H | .97 |
| Overweight, 25≤BMI<30 | 780 | 198 | 0.93 (0.77-1.12) | H - | |
| Obese, <i>BMI≥30</i> | 256 | 119 | 0.78 (0.60-1.02) | ⊢_ | |
| History of diabetes | | | | | |
| No | 1672 | 562 | 0.86 (0.77-0.97) | • •• • | .75 |
| Yes | 133 | 61 | 0.95 (0.60-1.50) | ⊢−− | |
| Hormone therapy in women | | | | | |
| No | 363 | 225 | 0.88 (0.72-1.08) | ⊢ , | .79 |
| Yes | 303 | 85 | 0.94 (0.70-1.26) | F | |
| Histological type | | | | | |
| Adenocarcinoma | 685 | 230 | 0.79 (0.66-0.95) | Hand a state of the state of th | .04 |
| Squamous cell carcinoma | 434 | 164 | 0.84 (0.67-1.06) | ⊢ - ■ | |
| Other non-small cell carcinoma | 272 | 98 | 0.90 (0.68-1.19) | ⊢ | |
| Small cell carcinoma | 220 | 72 | 0.93 (0.65-1.32) | ⊢ | |
| Tumor grade | | | | | |
| Well/moderately differentiated | 378 | 95 | 0.82 (0.62-1.07) | | .71 |
| Poorly & undifferentiated | 585 | 156 | 0.92 (0.74-1.13) | ⊢ | |
| Time from baseline to diagnosis | | | | | |
| < 5 years | 827 | 352 | 0.87 (0.74-1.01) | H | .90 |
| 5-9 years | 785 | 223 | 0.80 (0.67-0.96) | H H | |
| ≥ 10 vears | 193 | 48 | 1.10 (0.75-1.60) | P | |

Figure 1. Prediagnosis leisure-time physical activity and all-cause mortality among lung cancer patients: stratified analyses of localized and regional stage cases. HRs (95% CIs) for \geq 8.3 MET-h/week vs none were shown after adjusting for age at diagnosis, sex, smoking status, smoking pack-years, race and ethnicity, education, alcohol consumption, history of diabetes, BMI levels, hormone therapy in women, histological type, tumor stage, and grade of lung cancer and stratifying by cohort, year of lung cancer diagnosis, and time interval from leisure-time physical activity assessment to lung cancer diagnosis. Interaction (additive) refers to global P value for relative excess risk due to interaction between prediagnosis leisure-time physical activity and each stratification variable. All P values were corrected for multiple comparisons by controlling the false-discovery rate. All statistical tests were 2-sided. Error bars represent the 95% CIs. BMI = body mass index; CI = confidence interval; HR = hazard ratio; MET-h/week = metabolic-equivalent hours per week.

cancer patients and reduce postsurgery complications (40). In our study, the association of prediagnosis LTPA appeared to be stronger when exposure was measured closer to diagnosis, lending some support to the mechanisms mentioned above. Existing biological evidence and our epidemiological observations suggest that habitual LTPA may improve lung cancer survivorship, especially for early stage lung cancer. However, we did not find any statistically significant association with distant stage cancer. We speculate that the high fatality rate and short survival time for late stage lung cancer made it difficult for us to detect a moderate association with prediagnosis LTPA, if one exists.

Our study has several strengths. This is the largest prospective investigation on the association of prediagnosis LTPA with

| | No.o | f deaths | Adjusted | | |
|---|-------|----------|------------------|----------------------------|---------------|
| ≥ 8.3 MET-hr/wk vs. none | ≥ 8.3 | None | HR (95% CI) | | P Interaction |
| Age at lung cancer diagnosis | | | · · | | |
| <70 years | 793 | 331 | 0.81 (0.71-0.93) | F 2 -1 | .23 |
| ≥70 years | 608 | 166 | 1.03 (0.87-1.24) | | |
| Sex | | | | I | |
| Men | 893 | 253 | 0.91 (0.79-1.05) | F E H | .23 |
| Women | 508 | 244 | 0.93 (0.79-1.10) | F===-1 | |
| Race and ethnicity | | | . , | | |
| Asian | 89 | 162 | 0.91 (0.70-1.18) | | .66 |
| Black | 28 | 73 | 0.54 (0.35-0.85) | ⊢_ ∎1 | |
| White | 1272 | 257 | 0.91 (0.80-1.03) | H an t | |
| Education | | | | | |
| <high graduation<="" school="" td=""><td>666</td><td>348</td><td>0.97 (0.85-1.12)</td><td>HIIH</td><td>.89</td></high> | 666 | 348 | 0.97 (0.85-1.12) | H II H | .89 |
| >High school graduation | 735 | 149 | 0.76 (0.64-0.90) | H 20 -1 | |
| Smoking status | | | | | |
| Never | 143 | 87 | 0.79 (0.58-1.07) | ⊢−■ − <u>+</u> • | .66 |
| Former | 611 | 119 | 0.87 (0.71-1.06) | H | |
| Current | 647 | 291 | 0.95 (0.82-1.10) | H E H | |
| Alcohol consumption | | | | | |
| None | 352 | 252 | 0.99 (0.84-1.18) | | .89 |
| Moderate | 719 | 139 | 0.91 (0.76-1.10) | H E -1 | |
| Heavy | 330 | 106 | 0.74 (0.60-0.91) | H | |
| Obesity status | | | | | |
| ≤Normal weight, BMI<25 | 620 | 250 | 0.94 (0.79-1.11) | H R H | .66 |
| Overweight, 25≤BMI<30 | 592 | 162 | 0.88 (0.74-1.04) | H- | |
| Obese, <i>BMI≥30</i> | 189 | 85 | 0.86 (0.66-1.13) | | |
| History of diabetes | | | | | |
| No | 1297 | 455 | 0.89 (0.80-1.00) | P ara -i | .66 |
| Yes | 104 | 42 | 1.09 (0.73-1.63) | ⊢ ₩ 1 | |
| Hormone therapy in women | | | | | |
| No | 284 | 184 | 0.94 (0.77-1.14) | ⊢⊒ −1 | .66 |
| Yes | 224 | 60 | 0.97 (0.70-1.35) | | |
| Histological type | | | | | |
| Adenocarcinoma | 541 | 187 | 0.86 (0.72-1.04) | ⊢ ∎-+ | .003 |
| Squamous cell carcinoma | 320 | 131 | 0.84 (0.68-1.04) | F | |
| Other non-small cell carcinoma | 211 | 77 | 0.90 (0.70-1.17) | ₩ ;1 | |
| Small cell carcinoma | 190 | 61 | 0.94 (0.71-1.25) | ⊢ | |
| Tumor grade | | | | | |
| Well/moderately differentiated | 280 | 68 | 0.95 (0.71-1.26) | ⊨ | .42 |
| Poorly & undifferentiated | 458 | 125 | 0.86 (0.71-1.04) | H | |
| Time from baseline to diagnosis | | | | | |
| < 5 years | 658 | 278 | 0.97 (0.84-1.13) | F B | .65 |
| 5-9 years | 620 | 181 | 0.83 (0.70-0.98) | | |
| ≥ 10 years | 123 | 38 | 0.94 (0.65-1.36) | ⊢ −− ■ −−−−1 | |
| | | | | | |
| | | | | 0 0.5 1 1.5 | 2 |

Adjusted HR (95% CI)

Figure 2. Prediagnosis leisure-time physical activity and lung cancer-specific mortality among lung cancer patients: stratified analyses of localized and regional stage cases. HRs (95% CIs) for \geq 8.3 MET-h/week vs none were shown after adjusting for age at diagnosis, sex, smoking status, smoking pack-years, race and ethnicity, education, alcohol consumption, history of diabetes, BMI levels, hormone therapy in women, histological type, tumor stage, and grade of lung cancer and stratifying by cohort, year of lung cancer diagnosis, and time interval from leisure-time physical activity assessment to lung cancer diagnosis. For the lung-cancer mortality analyses, cases missing cause of death were excluded from the analysis, and death from other causes was treated as a competing risk. Interaction (additive) refers to global P value for relative excess risk because of interaction between prediagnosis leisure-time physical activity and each stratification variable. All P values were corrected for multiple comparisons by controlling the false-discovery rate. All statistical tests were 2-sided. Error bars represent the 95% CIs. BMI = body mass index; CI = confidence interval; HR = hazard ratio; MET-h/week = metabolic-equivalent hours per week.

lung cancer survival. We used individual participant data of more than 20 000 incident lung cancer patients from diverse populations. Our prospective design, large sample size, and extensive information on a wide range of clinical characteristics, smoking history, and other lifestyle factors enabled comprehensive analyses. All the analyses were controlled for or stratified by major prognostic and risk factors and considered potential competing risks of death, which enhanced the scientific rigor of our study. Nonetheless, several limitations should be acknowledged. First, because of a lack of information, we could not control for the influence of lung cancer treatment and care in the current study. To compensate for this limitation, we applied statistical models stratified by calendar year at lung cancer diagnosis with adjustment for treatment-related clinic factors (ie, histological type, stage, and grade). Second, we used a one-time measure of prediagnosis LTPA; thus, we could not consider changes in LTPA intensity or patterns over time, as well as other domains of PA (ie, occupation, household, and transportation). Measurement errors in self-reports (eg, overreporting) and MET estimations based on varying instruments across cohorts are another concern, despite using validated questionnaires (18-26). Our findings might be somewhat affected by these exposure misclassifications. Third, postdiagnosis information was unavailable in most participating studies. Given that the mean survival time after lung cancer diagnosis is relatively short, however, it may be challenging to find the actual impact of postdiagnosis factors, including LTPA, on lung cancer survival in a population-based setting. Furthermore, LTPA assessed after diagnosis would be more likely affected by diseases, resulting in bias because of reverse causation. Fourth, despite carefully adjusting for both smoking status and lifetime tobacco exposure in all analyses, residual confounding by smoking cannot be completely ruled out. For example, it is possible that heavy smokers are less likely to engage in LTPA and also more likely to underreport their smoking exposure. Finally, insufficient sample size of some subgroup analyses (eg, neversmokers, Black, and Asian participants) might contribute to the failure of reaching statistical significance in some of the observed associations.

This large-scale pooled analysis of 11 cohorts indicates that regular participation in LTPA, particularly when meeting or exceeding the minimum Physical Activity Guidelines, is associated with reduced mortality among lung cancer patients, especially those with early stage cancer. Our findings add the supporting evidence that adhering to the Physical Activity Guidelines before cancer diagnosis may have long-term benefits on lung cancer progression and/or survival. Future investigation incorporating various domains of PA objectively assessed at multiple time points throughout the life course and the lung cancer continuum is needed to confirm certain benefits of PA and develop PA promotion strategies for reducing the global burdens of lung cancer.

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Data Availability

The data underlying this article will be shared on reasonable request to the corresponding author after approval of principle investigators of participating cohorts.

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