

- 6 O'Halloran HM, Kwong K, Veldhoen RA, Maslove DM. Characterizing the patients, hospitals, and data quality of the eICU Collaborative Research Database. *Crit Care Med* 2020;48:1737–1743.
- 7 Bell BA, Ferron JM, Kromrey JD. Cluster size in multilevel models: the impact of sparse data structures on point and interval estimates in two-level models. American Statistical Association; 2008.
- 8 Austin PC, Leckie G. The effect of number of clusters and cluster size on statistical power and Type I error rates when testing random effects variance components in multilevel linear and logistic regression models. *J Stat Comput Simul* 2018;88:3151–3163.
- 9 Maas CJM, Hox JJ. Sufficient sample sizes for multilevel modeling. *Methodology* 2005;1:86–92.
- 10 Clarke P. When can group level clustering be ignored? Multilevel models versus single-level models with sparse data. *J Epidemiol Community Health* 2008;62:752–758.
- 11 Merlo J, Chaix B, Yang M, Lynch J, Råstam L. A brief conceptual tutorial of multilevel analysis in social epidemiology: linking the statistical concept of clustering to the idea of contextual phenomenon. *J Epidemiol Community Health* 2005;59:443–449.
- 12 Sanagou M, Wolfe R, Forbes A, Reid CM. Hospital-level associations with 30-day patient mortality after cardiac surgery: a tutorial on the application and interpretation of marginal and multilevel logistic regression. *BMC Med Res Methodol* 2012;12:28.
- 13 Wayne MT, Seelye S, Molling D, Wang XQ, Donnelly JP, Hogan CK, *et al.* Temporal trends and hospital variation in time-to-antibiotics among veterans hospitalized with sepsis. *JAMA Netw Open* 2021;4:e2123950.
- 14 Shinn JR, Kimura KS, Campbell BR, Sun Lowery A, Wootten CT, Garrett CG, *et al.* Incidence and outcomes of acute laryngeal injury after prolonged mechanical ventilation. *Crit Care Med* 2019;47:1699–1706.
- 15 Lyon SM, Benson NM, Cooke CR, Iwashyna TJ, Ratcliffe SJ, Kahn JM. The effect of insurance status on mortality and procedural use in critically ill patients. *Am J Respir Crit Care Med* 2011;184:809–815.
- 16 Gershengorn HB, Garland A, Kramer A, Scales DC, Rubenfeld G, Wunsch H. Variation of arterial and central venous catheter use in United States intensive care units. *Anesthesiology* 2014;120:650–664.
- 17 Wunsch H, Hill AD, Bosch N, Adhikari NKJ, Rubenfeld G, Walkey A, *et al.* Comparison of 2 triage scoring guidelines for allocation of mechanical ventilators. *JAMA Netw Open* 2020;3:e2029250.
- 18 Althouse AD, Raffa GM, Kormos RL. Your results, explained: clarity provided by row percentages versus column percentages. *Ann Thorac Surg* 2016;101:15–17.

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Smoking Prevalence during the COVID-19 Pandemic in the United States

To the Editor:

External shocks can yield insight into the socioeconomic determinants of health. In the wake of some major previous social (1) and economic disruptions (2), cigarette smoking increased. The coronavirus disease (COVID-19) pandemic may have affected smoking behavior in myriad ways. Financial strain from job loss (3, 4) and pandemic-related anxiety might have increased unhealthy coping behaviors, including tobacco use. But financial strain was cushioned by trillions of dollars of federal government pandemic-relief spending, and worries about a dangerous respiratory virus might have deterred cigarette use.

We examined changes in smoking prevalence in the United States during the COVID-19 pandemic.

Methods

We analyzed data from the 2016–2020 Behavioral Risk Factor Surveillance System (BRFSS) ($N = 2,193,981$), a nationally representative telephone survey of U.S. adults conducted by the Centers for Disease Control and Prevention with U.S. states/territories.

We tabulated the prevalence of current smoking, defined as having smoked 100 or more lifetime cigarettes and to be currently smoking, each survey year and interview quarter. We considered quarter 1 (Q1)-2016 to Q4-2019 the prepandemic period and Q2-2020 to Q4-2020 the pandemic period. We excluded 95,396 subjects with missing data on smoking, and (for

quarterly estimates) 11,288 subjects whose 2020 BRFSS interview occurred in early 2021.

We first examined trends according to interview quarter, both for the overall population and for age, sex, racial/ethnicity, income, and (for 2018–2020 only) urban/rural county subgroups.

To assess changes associated with the pandemic's onset, we performed an interrupted time series analysis using multivariable linear probability regression with the pre/post pandemic indicator as the predictor of interest, and with interview quarter treated as a continuous variable. To account for demographic shifts, we also included covariates for age, region, sex, education, race/ethnicity, and marital status. To evaluate whether changes in prevalence differed among aforementioned subgroups, we performed regressions that additionally included interaction terms for subgroup indicator \times pre/post indicator and subgroup indicator \times interview quarter.

For multivariable regressions, we excluded an additional 105,646 individuals interviewed Q1-2020, the quarter during which the pandemic unfolded, and then 82,559 individuals with missing data for one or more covariates.

Analyses were conducted with Stata 16 using BRFSS sampling weights and appropriate procedures for complex survey data. The institutional review board at the authors' institutions do not consider study of deidentified, publicly available data to be Human Subjects Research.

Results

Our study population consisted of 2,087,297 adults. Mean age was 47.5 years prepandemic and 47.6 years postpandemic; 51% of respondents were female in both periods.

Between survey years 2016 and 2019, annual smoking prevalence fell from 16.3% to 15.3%, and then more abruptly to 14.2% in 2020 (data not shown). Figure 1 provides quarterly trends in smoking prevalence overall and by subgroup. Smoking prevalence fell rapidly starting in Q1-2020 among those ages 18–54 years, while remaining

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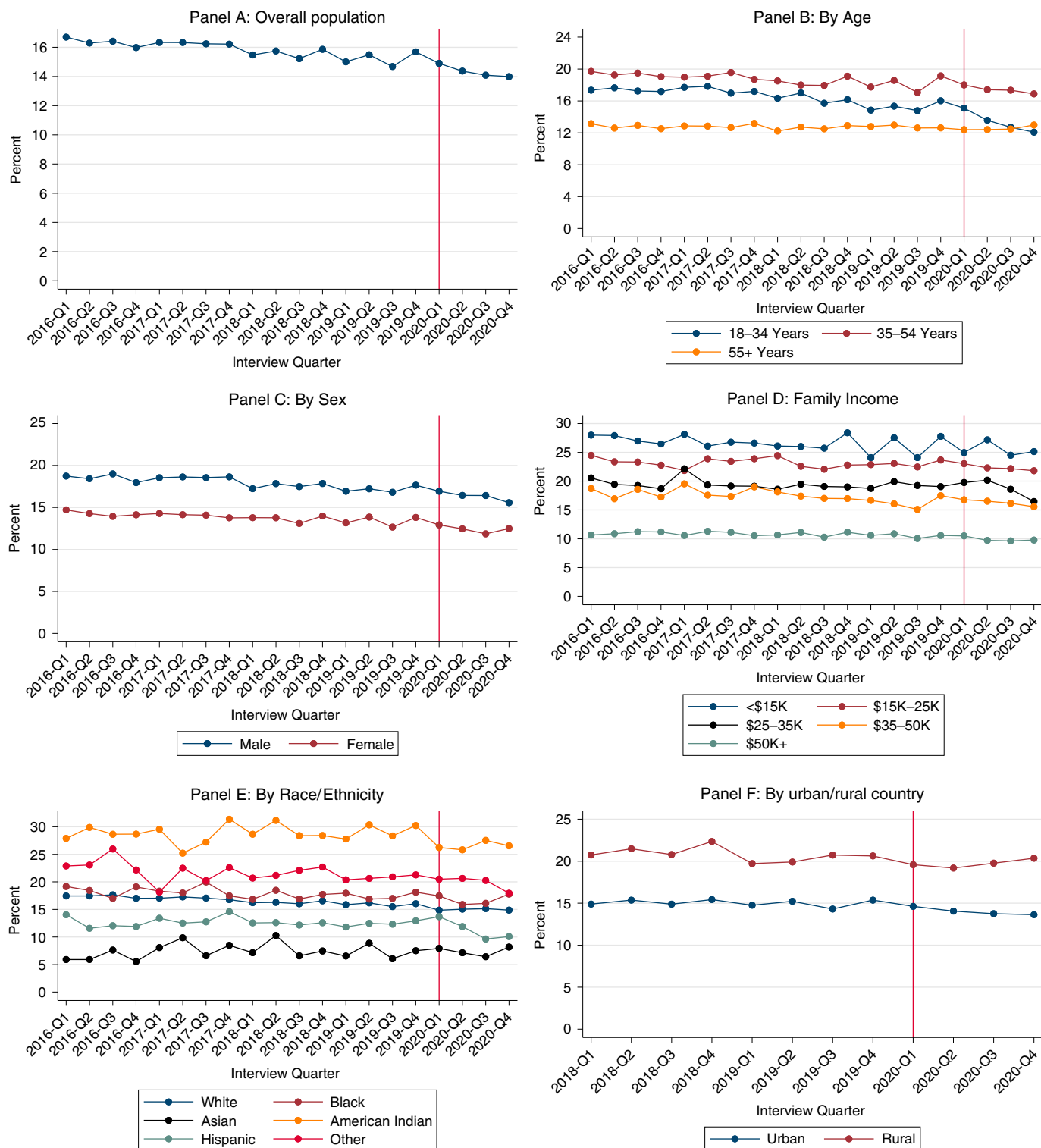


Figure 1. Current smoking prevalence by interview quarter, Behavioral Risk Factor Surveillance System (BRFSS) 2016–2020. The red line demarcates the beginning of the coronavirus disease (COVID-19) pandemic in the United States. Quarter 1 (Q1) = January–March; Q2 = April–June; Q3 = July–September; Q4 = October–December. Quarters refer to time when interview was conducted and do not necessarily correspond with survey year, as a small number of individuals in each BRFSS survey year are not interviewed until early in the following calendar year; here, they are presented at the time of interview, although the annual estimates reported in the main text refer to survey year. Study population: (A) Overall: $N = 2,087,297$; (B) Age: $N = 2,087,297$; (C) Sex: $N = 2,086,076$; (D) Income: $N = 1,747,808$; (E) Race: $N = 2,048,779$; and (F) Urban/rural: $N = 1,169,146$. The BRFSS imputed age for those with missing data for all years. The BRFSS also imputed race/ethnicity for those with missing data for 2017–2020 based on most common race/ethnicity in the region. All race/ethnicities except “Hispanic” are non-Hispanic; Hispanic individuals can be of any race. Other race includes Native Hawaiian or Pacific Islander, other, and those of multiple races. American Indian includes Alaskan Native.

Table 1. Adjusted change in smoking prevalence pre- (Q1-2016 to Q4-2019) to post (Q2-2020 to Q4-2020) COVID-19 pandemic among U.S. adults

Characteristic	Unweighted N for Analysis*	Percentage Point Change (95% CI)	P Value
Overall population	1,899,092	-0.59 (-0.99 to -0.19)	0.004
Age, yr	1,899,092		
18–34		-1.66 (-2.62 to -0.70)	0.001
35–54		-0.28 (-1.22 to 0.66)	0.56
≥55		Reference	
Sex	1,899,092		
Male		Reference	
Female		-0.44 (-1.24 to 0.36)	0.28
Race/ethnicity [†]	1,899,092		
White		Reference	
Black		0.07 (-1.34 to 1.49)	0.92
Asian		0.15 (-1.81 to 2.10)	0.88
American Indian/Alaskan Native		-1.79 (-5.82 to 2.25)	0.39
Hispanic		-1.45 (-2.79 to -0.12)	0.03
Other		-0.54 (-3.00 to 1.92)	0.67
Family income [‡]	1,603,982		
<\$15K		0.78 (-1.14 to 2.71)	0.43
\$15–25K		-0.27 (-1.67 to 1.13)	0.70
\$25–35K		0.21 (-1.46 to 1.87)	0.81
\$35–50K		0.96 (-0.36 to 2.27)	0.15
≥\$50K		Reference	
Urban/rural county [§]	1,036,995		
Urban		Reference	
Rural		0.31 (-1.51 to 2.13)	0.74

Definition of abbreviations: BRFSS = Behavioral Risk Factor Surveillance System; CI = confidence interval; COVID-19 = coronavirus disease; Q = quarter. All models adjusted for age (18–34, 35–54, ≥55 yr); region (Northeast, Midwest, South, West); sex (male/female); education (less than high school, high school graduate, some college/technical school, college/technical school graduate); marital status (married vs. unmarried, which includes: divorced, widowed, separated, never married, unmarried couple); race/ethnicity (non-Hispanic White, non-Hispanic Black, non-Hispanic Asian, non-Hispanic American Indian/Alaskan Native, Hispanic, non-Hispanic other); an indicator of interview quarter (treated continuously); and a pandemic onset indicator (pre- [Q1-2016 to Q4-2019] vs. post- [Q2-2020 to Q4-2020]). Models examining differences among subgroups were additionally adjusted for subgroup indicator × pandemic onset indicator and subgroup indicator × interview quarter interaction terms. Percentage point change estimates for the overall population are the coefficient for the pandemic onset indicator multiplied by 100; for subgroups, estimates are the coefficient for the pandemic onset indicator × subgroup interaction term multiplied by 100.

*All analyses exclude individuals with missing smoking data ($n = 95,396$ of total BRFSS population of $N = 2,193,981$), followed by those interviewed Q1-2021 ($n = 11,288$), followed by those interviewed in Q1-2020 ($n = 105,646$), followed by those with missing data on one or more covariate (total $n = 82,559$, including those in U.S. territories who could not be assigned a census region), leaving a final population of 1,899,092 for adjusted analyses.

[†]All race/ethnicities except Hispanic are non-Hispanic; Hispanic individuals can be of any race. Other race includes Native Hawaiian or Pacific Islander, other, and those of multiple races. The BRFSS imputed race for individuals with missing data for all years except 2016.

[‡]Family income missing for 295,110 of 1,899,092 individuals.

[§]Includes survey years 2018–2020 only.

stable among persons ages 55 years or older. In all years, smoking rates were higher among men than women and among lower-income versus higher-income individuals. Smoking prevalence was highest among American Indians/Alaskan Natives and lowest among Asians, with a substantial decrease among Hispanics after Q1-2020. After pandemic onset, smoking prevalence appeared to decrease in urban counties (which began the period with lower smoking rates) but changed little in rural counties.

Table 1 provides results adjusted for demographic factors and linear smoking trends, both for the overall population and among subgroups. Smoking prevalence in the overall population fell by an adjusted 0.59 percentage points (95% confidence interval [CI], -0.99 to -0.19) after pandemic onset. After pandemic onset, individuals ages 18–34 years experienced an adjusted 1.66 percentage point reduction (95% CI, -2.62 to -0.70) in smoking prevalence relative to individuals ages 55 years or older. Relative to non-Hispanic White individuals,

Hispanic individuals had a 1.45 percentage point decrease (-2.79 to -0.12) in smoking prevalence after onset of the pandemic. Differential changes between other subgroups did not reach statistical significance.

Discussion

Smoking prevalence fell during the COVID-19 pandemic in the United States, particularly among younger and Hispanic individuals, a trend that runs counter to those observed in some past social upheavals.

The causes of this salutary behavioral change are uncertain. In the past, increasing unemployment has sometimes been associated with increases in unhealthy coping behaviors, including cigarette smoking (3, 4) (e.g., a 600,000-person increase in the number of smokers in the United States [2] [although not in Iceland (5)] during the Great Recession). Dislocation, social isolation, and anxiety (6) might also be expected to increase psychosocial strain

and promote tobacco use, an effect that may, in part, explain increased smoking among those exposed to Hurricane Katrina (1) and terrorist attacks (7).

However, during COVID-19, the untoward effect of unemployment may have been blunted by the infusion of government relief funds, causing the poverty rate to paradoxically fall in 2020 despite economic contraction. Moreover, specific worries about lung health occasioned by the epidemic of severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2) respiratory illness may have deterred smoking uptake. School closures could also have reduced smoking behaviors among parents seeking to avoid smoking around children; in addition, for some, work- or commuting-related stress was potentially alleviated.

A few studies have examined the impact of COVID-19 on smoking. For instance, a survey conducted in Italy found that smoking prevalence decreased, although cigarette consumption increased among those who continued to smoke (8). However, previous analyses have had major limitations, including a lack of pre-pandemic sampling or limited generalizability (9, 10).

Although our study avoids these limitations, selection bias caused by the disruption of data collection because of the pandemic could raise concern about the comparability of the samples. However, the BRFSS was conducted by telephone both before and during the pandemic, and the 2020 response rate (47.9%) was, reassuringly, similar to that of the 2019 survey (49.4%). Unfortunately, we had no information on the quantity of cigarettes consumed by smokers.

Our findings suggest that increases in tobacco uptake may be avoidable during a period of pandemic-induced economic and social disruption. Public health agencies and clinicians should seek to consolidate these gains, and policymakers should assist such efforts by assuring adequate financial supports for disadvantaged Americans. ■

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References

- Alexander AC, Ward KD, Forde DR, Stockton M. Are posttraumatic stress and depressive symptoms pathways to smoking relapse after a natural disaster? *Drug Alcohol Depend* 2019;195:178–185.
- Gallus S, Ghislandi S, Muttarak R. Effects of the economic crisis on smoking prevalence and number of smokers in the USA. *Tob Control* 2015;24:82–88.
- Macy JT, Chassin L, Presson CC. Predictors of health behaviors after the economic downturn: a longitudinal study. *Soc Sci Med* 2013; 89:8–15.
- Falba T, Teng HM, Sindelar JL, Gallo WT. The effect of involuntary job loss on smoking intensity and relapse. *Addiction* 2005;100: 1330–1339.
- McClure CB, Valdimarsdóttir UA, Hauksdóttir A, Kawachi I. Economic crisis and smoking behaviour: prospective cohort study in Iceland. *BMJ Open* 2012;2:e001386.
- Cai C, Woolhandler S, Himmelstein DU, Gaffney A. Trends in anxiety and depression symptoms during the COVID-19 pandemic: results from the US Census Bureau's Household Pulse Survey. *J Gen Intern Med* 2021;36:1841–1843.
- DiMaggio C, Galea S, Li G. Substance use and misuse in the aftermath of terrorism: a Bayesian meta-analysis. *Addiction* 2009; 104:894–904.
- Carreras G, Lugo A, Stival C, Amerio A, Odone A, Pacifici R, et al. Impact of COVID-19 lockdown on smoking consumption in a large representative sample of Italian adults. *Tob Control* [online ahead of print] 29 Mar 2021; DOI: 0.1136/tobaccocontrol-2020-056440.
- Zhang X, Oluyomi A, Woodard L, Raza SA, Adel Fahmideh M, El-Mubasher O, et al. Individual-level determinants of lifestyle behavioral changes during COVID-19 lockdown in the United States: results of an online survey. *Int J Environ Res Public Health* 2021;18:4364.
- Vanderbruggen N, Matthys F, Van Laere S, Zeeuws D, Santermans L, Van den Ameel S, et al. Self-reported alcohol, tobacco, and cannabis use during COVID-19 lockdown measures: results from a web-based survey. *Eur Addict Res* 2020;26:309–315.

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