

RESEARCH ARTICLE

Association and biomarker potential of elevated serum adiponectin with nephropathy among type 1 and type 2 diabetics: A meta-analysis

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Abstract

Background

Managing nephropathy associated with diabetes mellitus warrant investigation of relevant biomarkers in predicting this condition. Adiponectin (ADP) may hold promise as a biomarker for diabetic nephropathy (DN). In this study, we examine associations of ADP with DN by meta-analyzing relevant literature. We also examined the predictive potential of ADP and estimate progression of DN.

Methods

Multi-database literature searches and serial omissions of articles yielded 13 studies for inclusion in the meta-analysis. We compared ADP levels between controls/ normoalbuminuria and cases with micro- and macroalbuminuria (MI and MA, respectively) as well as MI versus MA using standardized mean differences (SMD). Associations of ADP with DN were indicated with the P-value considered significant at ≤ 0.05 . Subgrouping was based on diabetes type (1 and 2). Predictive potential of ADP was explored with AUC (area under the curve) derived from Receiver Operating Characteristic curve analysis.

Results and conclusion

At high P-values of $<10^{-5}$, overall and subgroup outcomes indicated ADP associations with DN (up to SMD = 1.89–2.26, respectively). However, heterogeneity of the initial SMD effects (up to $I^2 = 99\%$) warranted examination of their sources which with the Galbraith plot method, either eliminated or reduced their heterogeneity, signifying combinability of the

Abbreviations: A, Metric for Area Under the Curve; ACEI, Angiotensin-converting enzyme inhibitors; ACR, Albumin-to-creatinine ratio; ADP, Adiponectin; AER, Albumin excretion rate; AJR, Rodriguez et al. meta-analysis; AM, Analysis model; ARBs, Angiotensin receptor blockers; CC, Case-control; CI, Confidence interval; CN, Control; CS, Cross-sectional; CVD, Cardiovascular disease; df, degree of freedom; DN, Diabetic nephropathy; DT, Diabetes type; EH, Eliminated heterogeneity; ESRD, End-stage renal disease; F, Fixed-effects; GS, Gain in significance; HMS, High to moderate significance; MA, Macroalbuminuria; MHS, Moderate to high significance; MI, Microalbuminuria; n, Number of studies / indicators; NCH, No change in heterogeneity; NM, No mention; NO, Normoalbuminuria; NOS, Newcastle–Ottawa Scale; NSO, Number of studies omitted; P_a , P-value for association; P_b , P-value for heterogeneity; PO, Prospective observational; Post, Post-outlier; Pre, Pre-outlier; PRISMA, Preferred Reporting Items for Systematic Reviews and Meta-Analyses; [R], Reference number; R, Random-effects; RH, Reduced heterogeneity; RHS, Retained high significance; ROC, Receiver Operating Characteristic; SC, Short communication; SD, Standard deviation; SE, Standard error; SMD, Standardized mean difference; T1DM, Type 1 diabetes mellitus; T2DM, Type 2 diabetes mellitus; TZD, Thiazolidinediones; UTI, Urinary tract infection; V, Versus.

studies. This feature along with consistency of significant associations, robust outcomes and significant AUC values provide good evidence of the associative and predictive roles of ADP in DN.

Introduction

Diabetes mellitus (DM) is a global health problem affecting an estimated 422 million people. This estimate may increase to about 48% by the year 2045 [1]. A salient feature of DM is the increasing prevalence among children, adolescents and younger adults warranting early and effective screening of this disease [2]. A serious microvascular complication of DM is diabetic nephropathy (DN), which occurs in 30–50% of patients with type 1 (T1) and type 2 (T2) DM. A widely used screening tool for DN is measurement of urine albumin in a 24-hour sample. A level of urine albumin at $> 30\text{mg/day}$ suggests DN [3]. Molecular mechanisms for DN are currently unclear, but adiponectin (ADP) may play a role in DN pathogenesis [4]. ADP is a small collagen-like protein encoded by the adiponectin gene, which is primarily expressed by adipocytes [5]. ADP is critical in glucose and fat metabolism where it regulates glucose uptake, gluconeogenesis and fatty acid oxidation in hepatocytes [6]. ADP has been shown to have anti-inflammatory, insulin-sensitizing, anti-atherogenic and cardioprotective activities [7–9] and has been linked to insulin resistance, obesity and cardiovascular diseases [10]. Patients at risk for these complications were found to have lower serum ADP levels [11]. In contrast, elevated ADP levels were found to be an independent predictor of the progression to end-stage renal disease (ESRD) [12]. Glomerular damage may be facilitated with passage of ADP through the glomerular basement membrane and excreted in the urine. Indeed, various ADP isoforms (low, medium and high molecular weights) can be measured in the urine, where ADP is considered to be a marker of vascular damage [13, 14]. Reports of ADP effects on DN have been variable, which prompted a meta-analysis to obtain more precise estimates. Here, we provide meta-analytical evidence to associate ADP with DN as well as examine biomarker potential of ADP.

Materials and methods

Search strategy

Literature search and article selection using the terms, “adiponectin”, “albuminuria” and “diabetic nephropathy” in English, we searched MEDLINE using PubMed, Science Direct and Google Scholar for publications as of March 07, 2018 (S1 Table). References cited in the retrieved publications were screened manually to identify additional eligible articles. Inclusion criteria included articles that presented ADP data with Mogensen’s design for albuminuria and DN [15].

Data extraction and eligibility criteria

Two investigators (NP and HJ) independently extracted data and reached consensus on all the items. For each eligible study, the following information was extracted: the first author’s name, publication year, country, ethnicity, patient sample size, source of samples, type of diabetes and whether the article was included in a previous meta-analysis [16]. Studies eligible for meta-analysis would have reported data for mean ADP (serum or plasma) concentrations both in patients with albuminuria (case) and without (controls, [CN]) or its levels below 30 mg/day

(normoalbuminuria, [NO]). Studies were excluded from our meta-analysis if they did not specifically report ADP concentrations in patients in terms of Mogensen's criteria [15]. In addition, we focused on studies with patients targeted for secondary prevention. Study authors were contacted to obtain additional information.

Methodological quality of the studies

The Newcastle–Ottawa Scale (NOS) assessment tool [17] was used to evaluate methodological quality of the included studies. These studies were judged based on three broad perspectives: selection, comparability and exposure (S2 Table). The star rating system has scores ranging from zero (worst) to 9 (best). Scores of 5–6 and ≥ 7 stars indicate moderate and high quality, respectively.

Meta-analysis protocol

The relationship between ADP and DN was expressed as standardized mean difference (SMD) and 95% confidence intervals (CI) using Review Manager 5.3 (Copenhagen: Nordic Cochrane Centre, Cochrane Collaboration, 2014). SMD estimates were considered as statistically significant if their CI did not cross zero and their P-values were < 0.05 . Most of the papers presented median and range, which were converted into approximate mean and standard deviation (SD) following Hozo et al. [18]. Where mean and 95% CI or standard error were reported, SD were derived as described by the Cochrane Collaboration [19]. From albuminuria data in DN patients, we compared NO with microalbuminuria (MI); NO with macroalbuminuria (MA) and MI with MA. Where CN were explicitly present, we also compared these with MI and with MA, respectively. Pooled estimates were obtained using either the fixed [20] (absence of heterogeneity) or random [21] (in its presence) effects models. We used random-effects model when results of the χ^2 -based Q test [22] yielded P-values of < 0.10 [23]. Sources of heterogeneity were identified with outlier analysis using the Galbraith plot method [24]. Heterogeneity was quantified with the I^2 statistic which measures the degree of inconsistency among studies [25]. Pooled estimates were subjected to sensitivity analysis which involved omitting one study at a time followed by recalculation to test for robustness of the summary effects. In subgroup analysis, we were confronted with ethnicity and diabetes type [DT], both with equal logistic potential for analysis (number of studies). We opted for DT (1 and 2) because it was less heterogeneous than ethnicity, which was beset with different geographies and admixture. Publication bias assessment focused only on comparisons with the number of studies > 10 because lower than this number, sensitivity of the qualitative and quantitative tests are low [26]. With SigmaPlot 11.0 (Systat Software, San Jose, CA), we used the ROC (Receiver Operating Characteristic) curve to estimate potential of ADP in predicting DN. This curve yields AUC (area under the curve) values (range: zero to one) that indicate degree of predictive accuracy and P-values to indicate presence or absence of significance. All P-values were two-sided with significance set at ≤ 0.05 except in heterogeneity estimation.

Results

Search outcomes

Fig 1 outlines the study selection process in a flowchart following PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-Analyses) guidelines [27]. A total of 567 citations during the initial search were followed by a series of omissions (S1 List) that eventually yielded 13 studies for inclusion in the meta-analysis [28–40].

Characteristics of the studies

Table 1 features characteristics of the included publications, the years of which ranged from 2004 to 2013. Overall, the total sample size included in the meta-analysis was 5,632 with a wide range among patients across all studies (38 to 2,090). In terms of DT, seven examined T2DM

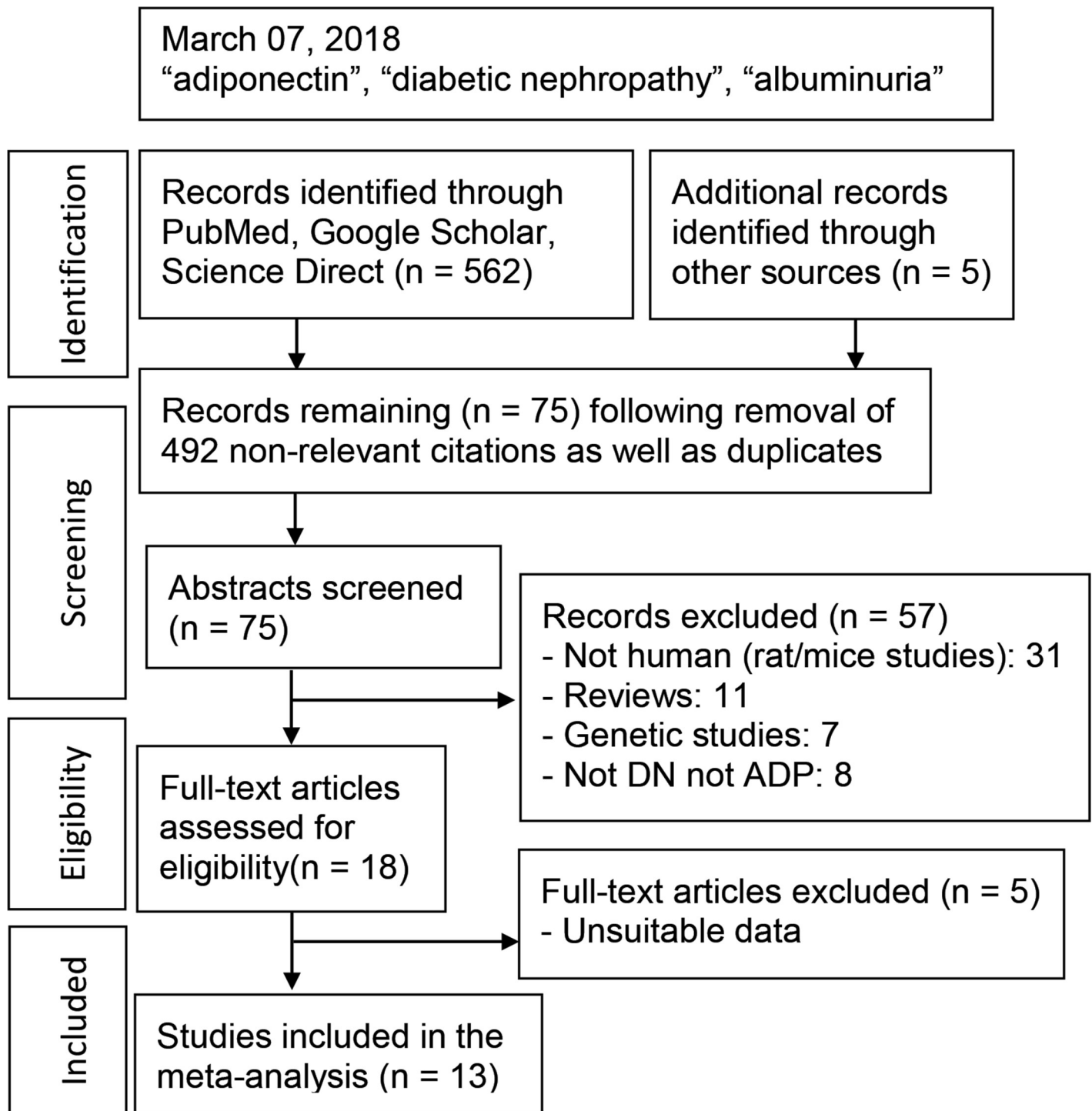


Fig 1. Summary of literature search: DN: diabetic nephropathy; ADP: adiponectin.

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Table 1. Characteristics of the included studies.

First author	[R]	Year	Country	Ethnicity	Sample size	Source	DT	Comparisons (n)	AJR	NOS
Ran	[28]	2010	China	Asian	50	Plasma	2	CN/NO* (18) MI (17) MA (15)	Yes	7
Fujita	[29]	2006	Japan	Asian	53	Serum	2	CN (20) NO (19) MI (18)	Yes	6
Kato	[30]	2008	Japan	Asian	198	Serum	2	NO (116) MI (47) MA (24)	Yes	4
Komaba	[31]	2006	Japan	Asian	179	Serum	2	NO (86) MI (44) MA (23)	Yes	5
Koshimura	[32]	2004	Japan	Asian	38	Serum	2	NO (18) MI (7) MA (13)	Yes	7
Saito	[33]	2007	Japan	Asian	280	Serum	2	CN (49) MA (22)	Yes	6
Jorsal	[34]	2013	Denmark	Western	145	Serum	1	CN (55) NO (58) MI (43) MA (44)	No	5
Schalkwijk	[35]	2006	Europe	Western	543	Serum	1	NO (328) MI (82) MA (128)	No	5
Panduru	[36]	2015	Finland	Western	2,090	Serum	1	CN (111) NO (1451) MI 319) MA (320)	No	7
Saraheimo	[37]	2008	Finland	Western	1,330	Serum	1	CN (204) NO (818) MI (216) MA (296)	No	6
Saraheimo	[38]	2005	Finland	Western	189	Serum	1	NO (66) MI (63) MA (60)	No	7
Hadjadj	[39]	2005	France	Western	126	Plasma	1	MI (18) MA (108)	No	8
Yilmaz	[40]	2008	Turkey	Western	411	Serum	2	CN (38) MI (40) MA (45)	Yes	8

[R] Reference number; DT: diabetes type; n: number of studies (n = 13); CN: control; NO: normoalbuminuria; MI: microalbuminuria; MA: macroalbuminuria; CN/NO*: NO is CN; AJR: Rodriguez et al. meta-analysis (yes and no indicate that the article was either in AJR or not, respectively); NOS: Newcastle-Ottawa Scale

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[28–33, 40] with remaining six focused on T1DM [34–39]. This meta-analysis included eight studies [29, 33–39] not found in the Rodriguez et al. meta-analysis [16]. S3 Table summarizes additional qualitative features of the component articles that include clinical and study design parameters. In clinical terms, patient diagnoses and restrictions to the cases were compared between studies. NOS scoring showed the mean and SD to be 6.2 ± 1.2 and a median of 6 indicating that the included studies were of moderate quality. S2 Table details the scoring outcomes from each included study. The PRISMA checklist was generated to provide detailed description of this meta-analysis (S4 Table).

Meta-analysis findings

Overall analysis

Table 2 summarizes the initial overall findings which indicate significantly higher ADP levels in MI and MA when compared to CN or NO (SMD = 0.74–1.89, P < 10⁻⁵–0.03). However,

Table 2. Overall summary associations of adiponectin with diabetic nephropathy.

	Test of association				Test of heterogeneity				Test of association				Test of heterogeneity			Effects of outlier treatment		
	n	SMD	95% CI	P _a	P _b	I ² (%)	AM	NSO (%)	n	SMD	95% CI	P _a	P _b	I ² (%)	AM			
Pre-outlier									Post-outlier									
CN v MI	6	0.73	-0.29, 1.74	0.16	10 ⁻⁵	98	R	3 (50)	3	1.19	0.41, 1.97	0.003	10 ⁻⁵	96	R	NCH, GS		
CN v MA	7	0.74	0.07, 1.41	0.03	10 ⁻⁵	97	R	3 (43)	4	0.88	0.75, 1.02	10⁻⁵	0.76	0	F	EH, MHS		
NO v MI	10	0.82	0.21, 1.43	0.009	10 ⁻⁵	98	R	4 (40)	6	0.15	0.03, 0.27	0.01	0.22	28	F	RH, HMS		
NO v MA	10	1.89	1.04, 2.73	10⁻⁵	10 ⁻⁵	99	R	3 (30)	7	0.82	0.70, 0.93	10⁻⁵	0.14	38	F	RH, RHS		
MI v MA	12	0.68	0.24, 1.13	0.003	10 ⁻⁵	94	R	6 (50)	6	0.55	0.42, 0.69	10⁻⁵	0.44	0	F	EH, MHS		

CN: control; NO: normoalbuminuria; MI: microalbuminuria; MA: macroalbuminuria; v: versus; n: number of studies in the comparison; SMD: standardized mean difference; CI: confidence interval; P_a: P-value for association; P_b: P-value for heterogeneity; AM: analysis model; R: random-effects; F: fixed-effects; NSO: number of studies omitted; NCH: no change in heterogeneity; EH: eliminated heterogeneity; RH: reduced heterogeneity; GS: gain in significance; MHS: moderate to high significance; HMS: high to moderate significance; RHS: retained high significance

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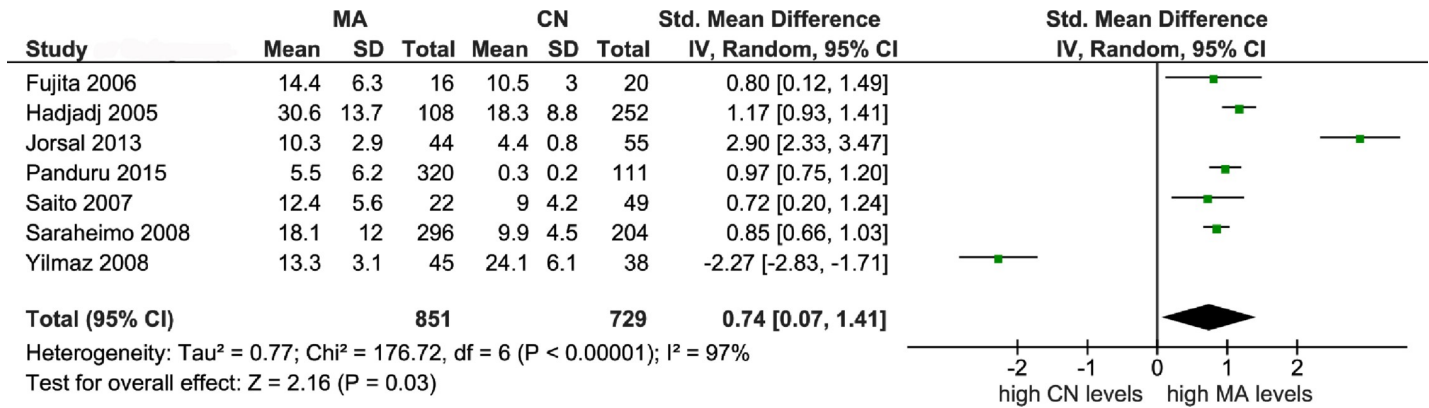


Fig 2. Forest plot of control versus macroalbuminuria comparison. LEGEND: MA: macroalbuminuria; CN: control; SD: standard deviation; CI: confidence interval; df: degree of freedom.

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these pre-outlier outcomes were all highly heterogeneous ($P_b < 10^{-5}$, $I^2 = 94-99\%$) with values obtained under the random-effects model. This heterogeneity warranted examining their sources, so we applied the outlier treatment which impacted on significance (reduced or elevated) and heterogeneity (eliminated or diminished) of the findings. The post-outlier values were thus, predominantly fixed-effects.

Figs 2–4 illustrate the mechanism of outlier treatment. Fig 2 details how the overall pooled SMD was derived from the CN versus MA comparison. In this forest plot, the random pooled effect was modestly significant (SMD = 0.74, $P_a = 0.03$) and highly heterogeneous ($P_b < 10^{-5}$, $I^2 = 97\%$). Sources of this heterogeneity were examined with the Galbraith plot (Fig 3) which identified three outlier studies [34, 39, 40]. Outlier treatment outcomes were (i) eliminated heterogeneity ($P_b = 0.76$, $I^2 = 0\%$), (ii) high statistical significance (SMD = 0.88, $P_a < 10^{-5}$) and (iii) shift to fixed-effects model (Fig 4).

Subgroup analysis

Table 3 shows all pre-outlier outcomes in the DT subgroup. Significant T1DM subgroup effects indicated higher ADP levels in MI and MA compared to either CN or NO (SMD = 0.75–2.26, $P_a = 10^{-5}$ –0.03). T2DM SMD outcomes ranged from -1.44 to 1.40 ($P_a < 10^{-5}$) indicating variability in the direction of effects. This variability disappeared following outlier treatment (SMD = 0.56–1.15, $P_a = 10^{-5}$ –0.001). Where subgroup effects were heterogeneous, their sources were examined with outlier treatment, outcomes of which are summarized in Table 4. All post-outlier outcomes in the subgroups were homogeneous ($I^2 = 0\%$) regardless of magnitude of associative significance ($P_a < 10^{-5}$ –0.006).

Table 5 presents the trends from pre- to post-outlier outcomes using three indicators. In the overall analysis, proportion of high significance ($P \leq 10^{-5}$) increased (20% to 60%), heterogeneity (random-effects) declined (100% to 20%) and homogeneity ($I^2 = 0\%$) was generated (0% to 40%). In the subgroup comparisons, high significance and homogeneity frequencies increased (40% to 57% and 20% to 100%, respectively). Most salient here was disappearance of heterogeneity (80% to 0%).

Sensitivity analysis and publication bias

Outcomes from all levels of comparisons were found to be robust indicating the stability of our findings (data not shown). Table 6 shows that all comparisons with studies >10 showed

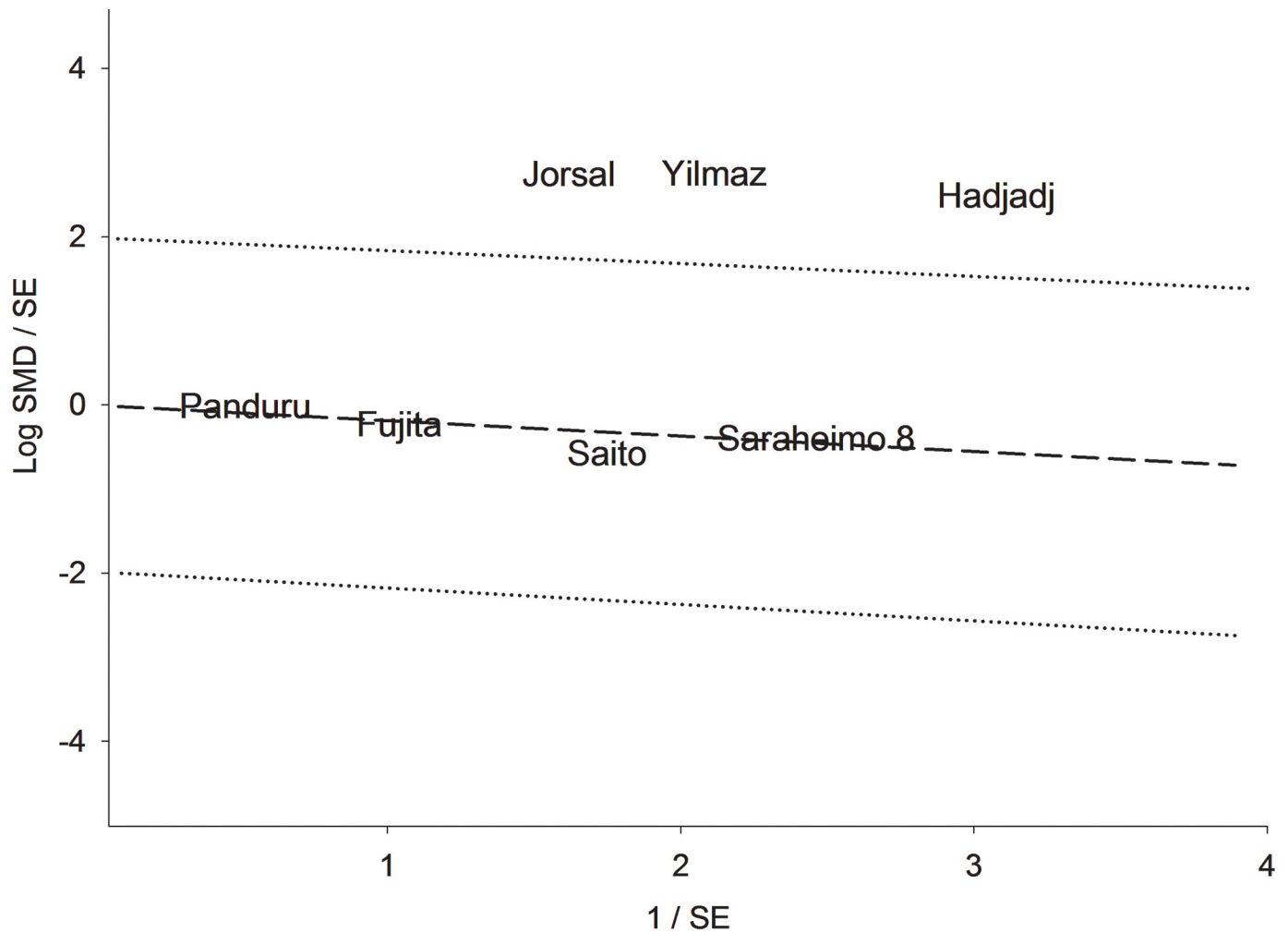


Fig 3. Galbraith plot analysis of outliers from the control versus macroalbuminuria comparison. LEGEND: SMD: standardized mean difference; SE: standard error.

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no evidence of publication bias, except the NO versus MA comparison in the correlation test ($P < 0.001$).

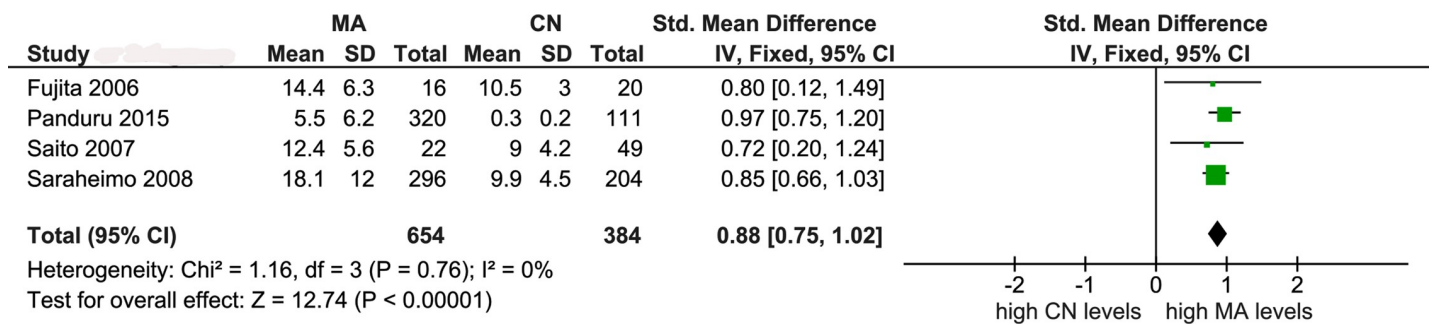


Fig 4. Forest plot of post-outlier control versus macroalbuminuria comparison. LEGEND: MA: macroalbuminuria; CN: control; SD: standard deviation; CI: confidence interval; df: degree of freedom.

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Table 3. Subgroup analysis of adiponectin associations with diabetic nephropathy.

	n	Test of association			Test of heterogeneity		
		SMD	95% CI	P _a	P _b	I ² (%)	AM
T1DM							
CN v MI	4	1.74	0.78, 2.69	0.0004	10 ⁻⁵	97	R
CN v MA	4	1.39	0.88, 1.89	10 ⁻⁵	10 ⁻⁵	93	R
NO v MI	5	1.10	0.13, 2.07	0.03	10 ⁻⁵	99	R
NO v MA	5	2.26	0.97, 3.56	0.0006	10 ⁻⁵	99	R
MI v MA	6	0.75	0.22, 1.29	0.006	10 ⁻⁵	95	R
T2DM							
CN v MI	2	-1.44	-1.85, -1.03	10 ⁻⁵	0.48	0	F
CN v MA	3	-0.25	-2.28, 1.78	0.81	10 ⁻⁵	97	R
NO v MI	5	0.46	0.25, 0.68	10 ⁻⁵	0.43	0	F
NO v MA	5	1.40	0.93, 1.87	10 ⁻⁵	0.03	62	R
MI v MA	6	0.64	-0.29, 1.57	0.18	10 ⁻⁵	92	R

T1DM: Type 1 diabetes mellitus; T2DM: Type 2 diabetes mellitus; CN: control; NO: normoalbuminuria; MI: microalbuminuria; MA: macroalbuminuria; v: versus; n: number of studies; SMD: standardized mean difference; CI: confidence interval; P_a: P-value for association; P_b: P-value for heterogeneity; AM: analysis model; F: fixed-effects, R: random-effects

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Evaluating the potential of ADP in predicting DN

We generated ROC curves for the CN, NO, MI and MA comparisons, details of which are summarized in S5 Table. Those with significant AUC outcomes are shown in Fig 5. ROC curve for ADP (NO versus MA and MI versus MA) showed respective post-outlier (non-heterogeneous) AUC values of 0.88 (95% CI 0.68–1.07, P = 0.02) and 0.83 (95% CI 0.59–1.10, P = 0.05) indicating good potential in discriminating patients with and without DN.

Discussion

Summary interpretations

The meta-analysis findings in this study provide good evidence of associative and predictive potential of ADP, for two reasons. First, consistent outcomes of significance were supported

Table 4. Outlier outcomes in the subgroups.

	n	Test of association			Test of heterogeneity		
		SMD	95% CI	P _a	P _b	I ² (%)	AM
T1 DM							
CN v MA	2	0.90	0.75, 1.04	10 ⁻⁵	0.40	0	F
NO v MI	2	0.37	0.11, 0.63	0.006	0.47	0	F
NO v MA	3	0.76	0.63, 0.88	10 ⁻⁵	0.51	0	F
MI v MA	2	0.61	0.45, 0.77	10 ⁻⁵	0.46	0	F
T2 DM							
CN v MA	2	0.75	0.34, 1.16	0.0004	0.85	0	F
NO v MA	4	1.15	0.86, 1.43	10 ⁻⁵	0.54	0	F
MI v MA	3	0.56	0.23, 0.90	0.001	0.85	0	F

T1DM: Type 1 diabetes mellitus; T2DM: Type 2 diabetes mellitus; CN: control; NO: normoalbuminuria; MI: microalbuminuria; MA: macroalbuminuria; v: versus; n: number of studies; SMD: standardized mean difference; CI: confidence interval; P_a: P-value for association; P_b: P-value for heterogeneity; AM: analysis model; F: fixed-effects; CN versus MI in T1DM did not change with outlier treatment.

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Table 5. Summary of outlier outcomes by indicator.

Indicator	Feature	Overall ^a			Subgroup		
		n	Pre (%)	Post (%)	n	Pre ^b (%)	Post ^c (%)
$P \leq 10^{-5}$	high significance	5	1 (20)	3 (60)	10	4 (40)	7 (57)
Random-effects	heterogeneity	5	5 (100)	1 (20)	10	8 (80)	0 (0)
$I^2 = 0\%$	homogeneity	5	0 (0)	2 (40)	10	2 (20)	7 (100)

n: number of indicators

^a: data from Table 2

^b: data from Table 3

^c: data from Table 4; Pre; pre-outlier; Post: post-outlier

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by homogeneity of post-outlier results indicating combinability of the studies. Furthermore, high magnitude (multiple P_a -values of 10^{-5}), precision of effects (indicated by the narrow CIs), robustness (from sensitivity treatment) and lack of bias (outcomes of Egger’ tests) helped improve the outcomes. Key indicators used in this meta-analysis rendered outlier treatment as instrumental in generating high significance and homogeneity as well as eliminating heterogeneity. In addition, outlier treatment generated associative similarities between T1DM and T2DM by virtue of common high significance ($P \leq 10^{-5}$) and homogeneity ($I^2 = 0\%$). Second, the difference in AUC (NO versus MA and MI versus MA) between the non-significant pre-outlier outcomes ($P = 0.32$ and 0.13) and significant post-outlier results ($P = 0.02$ and 0.05) indicates the contributive utility of outlier treatment to the discriminative potential of ADP in predicting DN (NO versus MI) and estimating its progression (MI versus MA). However, marginal significance of the MI versus MA outcome warrant caution in its interpretation.

Primary study findings on the role of ADP in DN have been variable. Several methodological problems may explain the discrepancies, including limited statistical power, unrecognized confounding factors and stratification of populations [41]. Reporting study-specific effects of ADP has ranged from association to prediction and progression. In presence of association, risk effects may indicate protection or susceptibility which may or may not be significant. Meta-analysis, however, informs more in reporting effects for the role of ADP in DN. These involve exploration of magnitude, precision, consistency, stability and heterogeneity of outcomes. These features, along with significant AUC outcomes, raise the levels of evidence to support conclusions of ADP-DN associations and potential of ADP in helping predict DN progression.

Comparison with previous studies

Significant AUC outcomes in our post-outlier analysis likely render a role for ADP as susceptibility screening tool for DN. Of the five component articles that addressed the role of ADP in

Table 6. Publication bias assessment of adiponectin association outcomes with diabetic nephropathy.

	n	Egger regression asymmetry test		Begg-Mazumdar correlation test	
		Intercept	P-value	Kendall’s τ	P-value
NO v MI	10	1.69	0.68	0.20	0.42
NO v MA	10	3.35	0.49	0.49	< 0.001
MI v MA	12	-1.07	0.68	-0.03	0.89

NO: normoalbuminuria; MI: microalbuminuria; MA: macroalbuminuria; v: versus; n: number of studies

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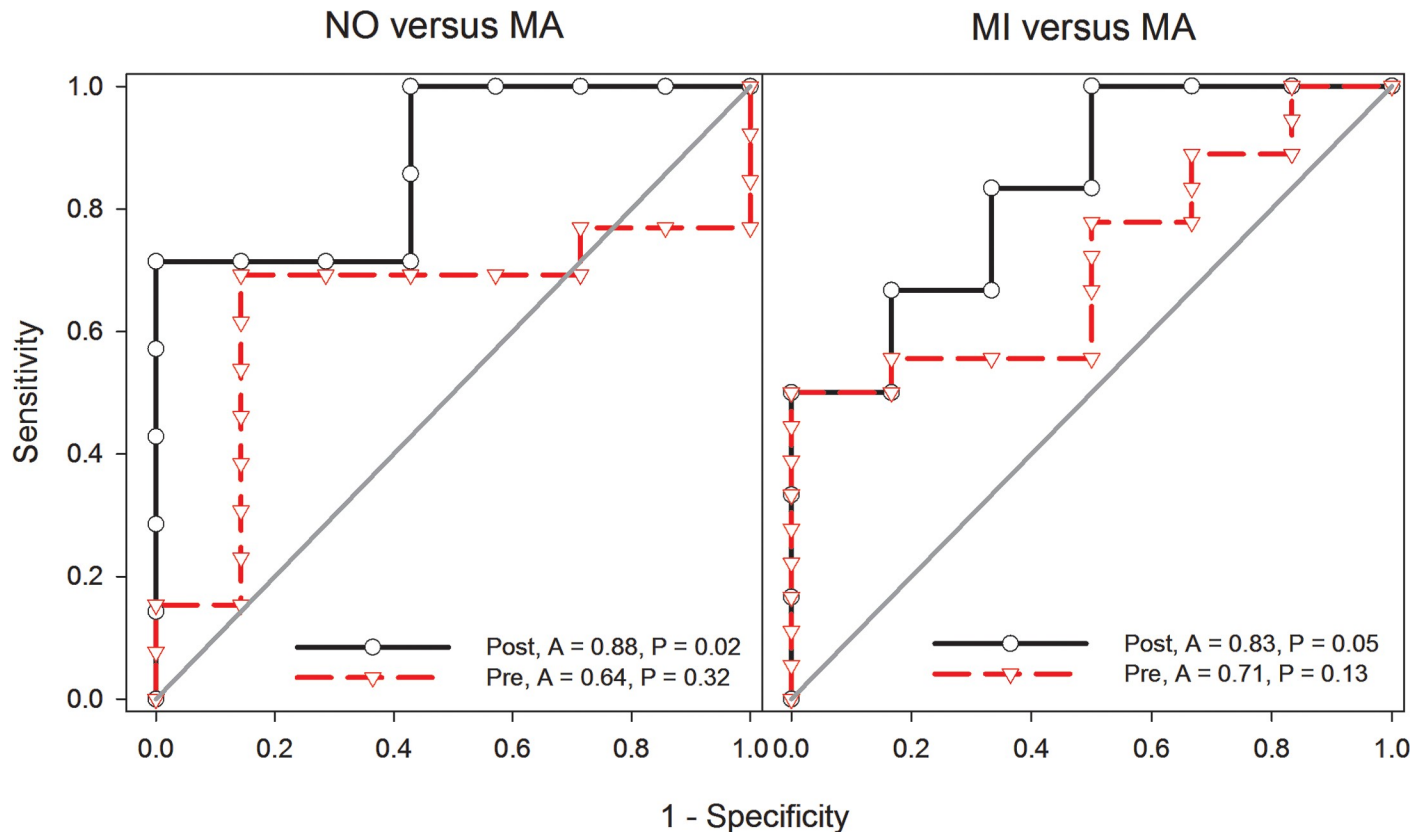


Fig 5. Receiver Operating Characteristic (ROC) curve analysis of ADP potential to predict DN. Area under the curve (A) values (range: zero-one) indicate predictive potential classified as poor (< 0.69), fair (> 0.70–0.79) and good (0.80–0.89). P-values associated with AUC are significant at ≤ 0.05 . ADP: adiponectin; DN: diabetic nephropathy; NO: normoalbuminuria; MA: macroalbuminuria; A: area under the curve; Pre: pre-outlier; Post: post-outlier.

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DN progression or its predictive capability [28, 29, 36, 37, 40], one used ROC curve analysis but focused on uADP [36]. An external study [42] that examined AUC in predicting renal decline in pediatric lupus nephritis also used uADP. In terms of direction of effect and significance of association, our DN findings agree with a previous meta-analysis [16]. Still, differences exist between the two meta-analyses (previous and ours). Rodriguez et al. [16] covered four studies examining ADP and DN in three milieus (NO, MI and MA) without the actual CN data. We covered 13 studies, using CN where data were available. In terms of levels of comparisons, the previous meta-analysis covered three (NO versus MI; NO versus MA and MI versus MA) while we covered up to five which included CN. Of the five levels of comparisons in the overall analysis and subgroups, NO versus MA had the strongest associations as indicated by SMDs of up 2.3-fold and the most number of 10^{-5} P_a -value ($n = 7$). This finding agrees with that of Rodriguez et al. [16] (SMD = 1.37, $P < 10^{-5}$). In contrast to the previous meta-analysis, however, our reason for exclusive focus on ADP was so that we could apply a number of meta-analytical treatments to our findings, endeavoring to raise the quality of evidence. The findings from both meta-analyses (theirs and ours) were not materially different (both were statistically significant). What differentiates our findings is that we explore the heterogeneous outcomes with outlier analysis, thus generating pooled SMDs of either non-heterogeneous or homogeneous nature. Thus, we address combinability of the studies in addition to consistency (significance across comparisons), stability (subgroup analysis) and robustness (sensitivity analysis) of the findings.

Physiological correlates

Although our meta-analysis focused on serum and plasma ADP, some of the papers [32, 34, 36] included urinary ADP (uADP), but the data was too sparse to warrant further analysis. uADP has been linked to renal tubular injury [29]. Thus, uADP may reflect both glomerular and tubular damage in DN. Shimotomai et al. [43] have demonstrated that uADP levels were significantly increased among patients with DN. Such increased levels might result from enhanced filtration of circulating ADP through the damaged kidney. Concordant with our results, Yamamoto et al [44] reported that serum ADP levels are increased in DN patients with MA. However, serum ADP level has been reported to be twice as high in patients with nephrotic syndrome compared to those without, irrespective of glomerular filtration rate [45].

Among type 1 diabetics, patients with MA on average had a higher level of serum ADP when compared to patients with NO and MI, and high ADP levels in urine predicted the progression of MA to ESRD [38]. On the other hand, patients with T1DM, uADP levels were reported to be an independent predictor of DN progression from MA to ESRD. Thus, increase in uADP was identified to be associated with ESRD risk in patients with T1DM [36]. These results indicate that in diabetic patients, an increase in uADP is associated with reduced renal function [46]. Given these physiological correlates, uADP levels maybe a good predictor of DN progression. Diseased states notwithstanding, we need to consider that circulating ADP levels have been positively and independently associated with albuminuria in non-diabetic subjects [47] and that there is a cause-effect relationship between the two variables.

Strengths and limitations

Interpreting these meta-analysis results warrant awareness of its strengths and limitations. Strengths include: (i) Post-outlier outcomes in zero heterogeneity highlight combinability of the studies; (ii) consistency of significant associations strengthens the evidence for increased ADP associations with DN; (iii) predictive potential of ADP is bolstered by significant AUC under non-heterogeneous conditions and (iv) all comparisons were deemed robust underpinning the stability of our findings. On the other hand, limitations of our study include: (i) the wide range of patient sample sizes in the studies elicited emphasized contributions from large studies while minimizing those from small studies. This renders caution when interpreting the results; (ii) resulting losses and reductions in heterogeneity from outlier analysis were obtained at the expense of statistical power; where the greater the percentage of omitted studies, the less power is accorded to the resulting post-outlier pooled effects. This is most evident in the outlier-treated subgroups; (iii) issues of multiple comparisons and inadequate statistical power [48] precluded examination of the ethnic subgroups (Asian and Western) as well as ADP isoforms; (iv) ADP may operate in conjunction with other adipocytokines [16, 49] which may likely confound our findings and (v) the relationship established in our study was, for the most part, provided at univariate level of analysis but only from one study at the multivariate level [40].

Conclusions

In conclusion, our findings indicate that ADP is significantly associated with DN. High significance, non-heterogeneity and stability of the pooled SMDs confirms its associative potential of ADP. In addition, its predictive potential is underpinned by significant AUC values under conditions of non-heterogeneity. Further studies regarding interaction of ADP with other markers and variables may help better understand its role in DN.

Supporting information

S1 Table. Database search algorithms for adiponectin associations with diabetic nephropathy.

(DOCX)

S2 Table. Assessment of methodological quality of the articles using the Newcastle-Ottawa Scale.

(DOCX)

S3 Table. Qualitative features of the studies based on methodological, biochemical and clinical characteristics.

(DOCX)

S4 Table. PRISMA checklist.

(DOCX)

S5 Table. Evaluating predictive potential of adiponectin in diabetic nephropathy.

(DOCX)

S1 List. Excluded citations.

(DOCX)

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References

1. World Health Organization Global Report on Diabetes (2016) http://www.who.int/about/licensing/5Cnhttp://apps.who.int/iris/bitstream/10665/204871/1/9789241565257_eng.pdf Accessed March 2018.
2. Smriti K, Pai KM, Ravindranath V, Gadicherla S, Pentapati KC. Salivary Glucose as a Diagnostic Marker for Diabetes Mellitus. *Journal of diabetes science and technology*. 2016; 10(4):991–2. <https://doi.org/10.1177/1932296816637619> PMID: 26961976; PubMed Central PMCID: PMC4928231.

3. John S. Complication in diabetic nephropathy. *Diabetes & metabolic syndrome*. 2016; 10(4):247–9. <https://doi.org/10.1016/j.dsx.2016.06.005> PMID: 27389078.
4. Zha D, Wu X, Gao P. Adiponectin and Its Receptors in Diabetic Kidney Disease: Molecular Mechanisms and Clinical Potential. *Endocrinology*. 2017; 158(7):2022–34. Epub 2017/04/13. <https://doi.org/10.1210/en.2016-1765> PMID: 28402446.
5. Maeda K, Okubo K, Shimomura I, Funahashi T, Matsuzawa Y, Matsubara K. cDNA cloning and expression of a novel adipose specific collagen-like factor, apM1 (adipose most abundant gene transcript 1). 1996. *Biochem Biophys Res Commun*. 2012; 425(3):556–9. Epub 2012/08/29. <https://doi.org/10.1016/j.bbrc.2012.08.023> PMID: 22925673.
6. Sheng T, Yang K. Adiponectin and its association with insulin resistance and type 2 diabetes. *Journal of genetics and genomics = Yi chuan xue bao*. 2008; 35(6):321–6. [https://doi.org/10.1016/S1673-8527\(08\)60047-8](https://doi.org/10.1016/S1673-8527(08)60047-8) PMID: 18571119.
7. Gable DR, Hurel SJ, Humphries SE. Adiponectin and its gene variants as risk factors for insulin resistance, the metabolic syndrome and cardiovascular disease. *Atherosclerosis*. 2006; 188(2):231–44. <https://doi.org/10.1016/j.atherosclerosis.2006.02.010> PMID: 16581078.
8. Bjornstad P, Pyle L, Kinney GL, Rewers M, Johnson RJ, Maahs DM, et al. Adiponectin is associated with early diabetic kidney disease in adults with type 1 diabetes: A Coronary Artery Calcification in Type 1 Diabetes (CACTI) Study. *Journal of diabetes and its complications*. 2017; 31(2):369–74. <https://doi.org/10.1016/j.jdiacomp.2016.06.012> PMID: 27368123; PubMed Central PMCID: PMC5156602.
9. Menzaghi C, Trischitta V, Doria A. Genetic influences of adiponectin on insulin resistance, type 2 diabetes, and cardiovascular disease. *Diabetes*. 2007; 56(5):1198–209. <https://doi.org/10.2337/db06-0506> PMID: 17303804.
10. Kawano J, Arora R. The role of adiponectin in obesity, diabetes, and cardiovascular disease. *Journal of the cardiometabolic syndrome*. 2009; 4(1):44–9. <https://doi.org/10.1111/j.1559-4572.2008.00030.x> PMID: 19245516.
11. Kim Y, Lim JH, Kim MY, Kim EN, Yoon HE, Shin SJ, et al. The Adiponectin Receptor Agonist AdipoRon Ameliorates Diabetic Nephropathy in a Model of Type 2 Diabetes. *Journal of the American Society of Nephrology: JASN*. 2018; 29(4):1108–27. <https://doi.org/10.1681/ASN.2017060627> PMID: 29330340.
12. Stenvinkel P, Marchlewska A, Pecoits-Filho R, Heimbürger O, Zhang Z, Hoff C, et al. Adiponectin in renal disease: relationship to phenotype and genetic variation in the gene encoding adiponectin. *Kidney international*. 2004; 65(1):274–81. <https://doi.org/10.1111/j.1523-1755.2004.00370.x> PMID: 14675060.
13. Shen YY, Peake PW, Charlesworth JA. Review article: Adiponectin: its role in kidney disease. *Nephrology*. 2008; 13(6):528–34. <https://doi.org/10.1111/j.1440-1797.2008.01018.x> PMID: 19138207.
14. von Eynatten M, Liu D, Hock C, Oikonomou D, Baumann M, Allolio B, et al. Urinary adiponectin excretion: a novel marker for vascular damage in type 2 diabetes. *Diabetes*. 2009; 58(9):2093–9. <https://doi.org/10.2337/db09-0204> PMID: 19509019; PubMed Central PMCID: PMC2731516.
15. Mogensen CE. Practice guidelines for chronic kidney disease. *Annals of internal medicine*. 2004; 140(11):933–4; author reply 4–5. PMID: 15172914.
16. Rodriguez AJ, Nunes Vdos S, Mastrorandi CA, Neeman T, Paz-Filho GJ. Association between circulating adipocytokine concentrations and microvascular complications in patients with type 2 diabetes mellitus: A systematic review and meta-analysis of controlled cross-sectional studies. *Journal of diabetes and its complications*. 2016; 30(2):357–67. <https://doi.org/10.1016/j.jdiacomp.2015.11.004> PMID: 26684169.
17. Wells G, Shea B, O'Connell D, Peterson J, Welch V, Losos M, Tugwell P. The Newcastle-Ottawa Scale (NOS) for assessing the quality of nonrandomised studies in meta-analyses. Ottawa Hospital Research Institute. http://www.ohri.ca/programs/clinical_epidemiology/oxford.asp. Accessed March, 2018.
18. Hozo SP, Djulbegovic B, Hozo I. Estimating the mean and variance from the median, range, and the size of a sample. *BMC medical research methodology*. 2005; 5:13. <https://doi.org/10.1186/1471-2288-5-13> PMID: 15840177; PubMed Central PMCID: PMC1097734.
19. Imputing standard deviations for changes from baseline. Available: http://handbook-5-1.cochrane.org/chapter_16/16_1_3_2_imputing_standard_deviations_for_changes_from_baseline.htm Accessed March, 2018.
20. Mantel N, Haenszel W. Statistical aspects of the analysis of data from retrospective studies of disease. *J Natl Cancer Inst*. 1959; 22(4):719–48. PMID: 13655060.
21. DerSimonian R, Laird N. Meta-analysis in clinical trials. *Control Clin Trials*. 1986; 7(3):177–88. PMID: 3802833.
22. Lau J, Ioannidis JP, Schmid CH. Quantitative synthesis in systematic reviews. *Annals of internal medicine*. 1997; 127(9):820–6. PMID: 9382404.

23. Higgins JP, Thompson SG. Quantifying heterogeneity in a meta-analysis. *Stat Med*. 2002; 21(11):1539–58. <https://doi.org/10.1002/sim.1186> PMID: 12111919.
24. Galbraith R. A note on graphical presentation of estimated odds ratios from several clinical trials. *Stat Med* 1988; 7((8)):889–94. PMID: 3413368
25. Higgins JP, Thompson SG, Deeks JJ, Altman DG. Measuring inconsistency in meta-analyses. *Bmj*. 2003; 327(7414):557–60. <https://doi.org/10.1136/bmj.327.7414.557> PMID: 12958120.
26. Ioannidis JP, Trikalinos TA. The appropriateness of asymmetry tests for publication bias in meta-analyses: a large survey. *CMAJ*. 2007; 176(8):1091–6. <https://doi.org/10.1503/cmaj.060410> PMID: 17420491; PubMed Central PMCID: PMC1839799.
27. Moher D, Shamseer L, Clarke M, Ghersi D, Liberati A, Petticrew M, et al. Preferred reporting items for systematic review and meta-analysis protocols (PRISMA-P) 2015 statement. *Systematic reviews*. 2015; 4:1. <https://doi.org/10.1186/2046-4053-4-1> PMID: 25554246; PubMed Central PMCID: PMC4320440.
28. Ran J, Xiong X, Liu W, Guo S, Li Q, Zhang R, et al. Increased plasma adiponectin closely associates with vascular endothelial dysfunction in type 2 diabetic patients with diabetic nephropathy. *Diabetes research and clinical practice*. 2010; 88(2):177–83. <https://doi.org/10.1016/j.diabres.2010.01.021> PMID: 20138682.
29. Fujita H, Morii T, Koshimura J, Ishikawa M, Kato M, Miura T, et al. Possible relationship between adiponectin and renal tubular injury in diabetic nephropathy. *Endocrine journal*. 2006; 53(6):745–52. PMID: 16966829.
30. Kato K, Osawa H, Ochi M, Kusunoki Y, Ebisui O, Ohno K, et al. Serum total and high molecular weight adiponectin levels are correlated with the severity of diabetic retinopathy and nephropathy. *Clinical endocrinology*. 2008; 68(3):442–9. <https://doi.org/10.1111/j.1365-2265.2007.03063.x> PMID: 17970779.
31. Komaba H, Igaki N, Goto S, Yokota K, Doi H, Takemoto T, et al. Increased serum high-molecular-weight complex of adiponectin in type 2 diabetic patients with impaired renal function. *American journal of nephrology*. 2006; 26(5):476–82. <https://doi.org/10.1159/000096870> PMID: 17095862.
32. Koshimura J, Fujita H, Narita T, Shimotomai T, Hosoba M, Yoshioka N, et al. Urinary adiponectin excretion is increased in patients with overt diabetic nephropathy. *Biochemical and biophysical research communications*. 2004; 316(1):165–9. <https://doi.org/10.1016/j.bbrc.2004.02.032> PMID: 15003525.
33. Saito T, Saito O, Kawano T, Tamemoto H, Kusano E, Kawakami M, et al. Elevation of serum adiponectin and CD146 levels in diabetic nephropathy. *Diabetes research and clinical practice*. 2007; 78(1):85–92. <https://doi.org/10.1016/j.diabres.2007.02.014> PMID: 17490776.
34. Jorsal A, Petersen EH, Tarnow L, Hess G, Zdunek D, Frystyk J, et al. Urinary adiponectin excretion rises with increasing albuminuria in type 1 diabetes. *Journal of diabetes and its complications*. 2013; 27(6):604–8. <https://doi.org/10.1016/j.jdiacomp.2013.06.008> PMID: 23969018.
35. Schalkwijk CG, Chaturvedi N, Schram MT, Fuller JH, Stehouwer CD, Group EPCS. Adiponectin is inversely associated with renal function in type 1 diabetic patients. *The Journal of clinical endocrinology and metabolism*. 2006; 91(1):129–35. <https://doi.org/10.1210/jc.2005-1117> PMID: 16219717.
36. Panduru NM, Saraheimo M, Forsblom C, Thorn LM, Gordin D, Waden J, et al. Urinary adiponectin is an independent predictor of progression to end-stage renal disease in patients with type 1 diabetes and diabetic nephropathy. *Diabetes care*. 2015; 38(5):883–90. <https://doi.org/10.2337/dc14-2276> PMID: 25720601.
37. Saraheimo M, Forsblom C, Thorn L, Waden J, Rosengard-Barlund M, Heikkila O, et al. Serum adiponectin and progression of diabetic nephropathy in patients with type 1 diabetes. *Diabetes care*. 2008; 31(6):1165–9. <https://doi.org/10.2337/dc07-2306> PMID: 18346990.
38. Saraheimo M, Forsblom C, Fagerudd J, Teppo AM, Petterson-Fernholm K, Frystyk J, et al. Serum adiponectin is increased in type 1 diabetic patients with nephropathy. *Diabetes care*. 2005; 28(6):1410–4. PMID: 15920060.
39. Hadjadj S, Aubert R, Fumeron F, Pean F, Tichet J, Roussel R, et al. Increased plasma adiponectin concentrations are associated with microangiopathy in type 1 diabetic subjects. *Diabetologia*. 2005; 48(6):1088–92. <https://doi.org/10.1007/s00125-005-1747-x> PMID: 15875155.
40. Yilmaz MI, Saglam M, Qureshi AR, Carrero JJ, Caglar K, Eyileten T, et al. Endothelial dysfunction in type-2 diabetics with early diabetic nephropathy is associated with low circulating adiponectin. *Nephrology, dialysis, transplantation: official publication of the European Dialysis and Transplant Association—European Renal Association*. 2008; 23(5):1621–7. <https://doi.org/10.1093/ndt/gfm828> PMID: 18175782.
41. Clark MF, Baudouin SV. A systematic review of the quality of genetic association studies in human sepsis. *Intensive Care Med*. 2006; 32(11):1706–12. <https://doi.org/10.1007/s00134-006-0327-y> PMID: 16957907.

42. Brunner HI, Gulati G, Klein-Gitelman MS, Rouster-Stevens KA, Tucker L, Ardoin SP, et al. Urine biomarkers of chronic kidney damage and renal functional decline in childhood-onset systemic lupus erythematosus. *Pediatric nephrology*. 2018. <https://doi.org/10.1007/s00467-018-4049-5> PMID: 30159624.
43. Shimotomai T, Kakei M, Narita T, Koshimura J, Hosoba M, Kato M, et al. Enhanced urinary adiponectin excretion in IgA-nephropathy patients with proteinuria. *Renal failure*. 2005; 27(3):323–8. PMID: 15957550.
44. Yamamoto M, Hujimoto Y, Hayashi S and Hashida S. A study of high, middle and low molecular weight adiponectin in urine as a surrogate marker for early diabetic nephropathy using ultra-sensitive ICT-EIA. *Annals of Clinical Biochemistry*. 2017; <https://doi.org/10.1177/0004563217748681> PMID: 29199440
45. Zoccali C, Mallamaci F, Panuccio V, Tripepi G, Cutrupi S, Parlongo S, et al. Adiponectin is markedly increased in patients with nephrotic syndrome and is related to metabolic risk factors. *Kidney international Supplement*. 2003;(84):S98–102. <https://doi.org/10.1046/j.1523-1755.63.s84.49.x> PMID: 12694320.
46. Kacso I, Lenghel A, Bondor CI, Moldovan D, Rusu C, Nita C, et al. Low plasma adiponectin levels predict increased urinary albumin/creatinine ratio in type 2 diabetes patients. *International urology and nephrology*. 2012; 44(4):1151–7. <https://doi.org/10.1007/s11255-011-0064-1> PMID: 21993769.
47. Menzaghi C, De Cosmo S, Copetti M, Salvemini L, De Bonis C, Mangiacotti D, et al. Relationship between ADIPOQ gene, circulating high molecular weight adiponectin and albuminuria in individuals with normal kidney function: evidence from a family-based study. *Diabetologia*. 2011; 54(4):812–8. <https://doi.org/10.1007/s00125-010-2037-9> PMID: 21229348.
48. Burke JF, Sussman JB, Kent DM, Hayward RA. Three simple rules to ensure reasonably credible subgroup analyses. *Bmj*. 2015; 351:h5651. <https://doi.org/10.1136/bmj.h5651> PMID: 26537915; PubMed Central PMCID: PMC4632208.
49. Goldberg RB. Cytokine and cytokine-like inflammation markers, endothelial dysfunction, and imbalanced coagulation in development of diabetes and its complications. *The Journal of clinical endocrinology and metabolism*. 2009; 94(9):3171–82. <https://doi.org/10.1210/jc.2008-2534> PMID: 19509100.