

# 1 Association between plausible genetic factors and weight loss from GLP1-RA and 2 bariatric surgery: a multi-ancestry study in 10 960 individuals from 9 biobanks

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## 68 Abstract

69 Obesity is a significant public health concern. GLP-1 receptor agonists (GLP1-RA), predominantly  
70 in use as a type 2 diabetes treatment, are a promising pharmacological approach for weight loss,  
71 while bariatric surgery (BS) remains a durable, but invasive, intervention. Despite observed  
72 heterogeneity in weight loss effects, the genetic effects on weight loss from GLP1-RA and BS  
73 have not been extensively explored in large sample sizes, and most studies have focused on  
74 differences in race and ethnicity, rather than genetic ancestry. We studied whether genetic  
75 factors, previously shown to affect body weight, impact weight loss due to GLP1-RA therapy or  
76 BS in 10,960 individuals from 9 multi-ancestry biobank studies in 6 countries. The average weight  
77 change between 6 and 12 months from therapy initiation was -3.93% for GLP1-RA users, with  
78 marginal differences across genetic ancestries. For BS patients the weight change between 6  
79 and 48 months from the operation was -21.17%. There were no significant associations between  
80 weight loss due to GLP1-RA and polygenic scores for BMI or type 2 diabetes or specific missense  
81 variants in the *GLP1R*, *PCSK1* and *APOE* genes, after multiple-testing correction. However, a  
82 higher polygenic score for BMI was significantly linked to lower weight loss after BS (+0.7% for 1  
83 standard deviation change in the polygenic score,  $P = 1.24 \times 10^{-4}$ ). In contrast, higher weight at  
84 baseline was associated with greater weight loss. Our findings suggest that existing polygenic  
85 scores related to weight and type 2 diabetes and missense variants in the drug target gene do  
86 not have a large impact on GLP1-RA effectiveness. Our results also confirm the effectiveness of  
87 these treatments across all major continental ancestry groups considered.

## 88 Introduction

89 The obesity epidemic presents a significant public health burden, driving the need for effective  
90 treatment options.<sup>1,2</sup> Despite 75% of adults with obesity having attempted to lose weight, most  
91 have not achieved lasting success.<sup>3</sup> Amid these challenges, glucagon-like peptide-1 receptor  
92 agonists (GLP1-RA) have emerged as a promising solution. As first discovered in the late 1980s,  
93 glucagon-like peptide-1 (GLP-1) is linked to the incretin effect, which promotes insulin release  
94 from pancreatic beta cells, thus playing a crucial role in regulating blood sugar levels.<sup>4,5</sup> This  
95 evidence led to the approval of the first GLP1-RA for treatment of type 2 diabetes in 2005.<sup>6-8</sup> Nine  
96 years later, an association between GLP-1 and reduced food intake paved the way for the  
97 development and approval of the first GLP1-RA for obesity treatment.<sup>5,9</sup>

98  
99 Clinical trials conducted in the following years demonstrated weight loss effects ranging from 6%  
100 to 16% for liraglutide and semaglutide, leading to a substantial increase in GLP1-RA usage  
101 worldwide<sup>10-12</sup>. Additional trials have further revealed cardiovascular<sup>13,14</sup> and renal benefits<sup>15</sup>, with  
102 ongoing trials currently underway.<sup>16</sup> Overall, anticipated trends suggest that by 2030, the US will  
103 witness approximately 30 million patients using GLP1-RA drugs, representing roughly 9% of its  
104 population, underscoring their potential in combating the obesity crisis and exploring new  
105 therapeutic applications.<sup>17-19</sup>

106  
107 Meanwhile, bariatric surgery (BS) stands as a durable and effective treatment for individuals with  
108 severe obesity (BMI  $\geq 40$  kg/m<sup>2</sup>) and metabolic comorbidities, and consistent with its  
109 effectiveness, the threshold has been reduced to a BMI  $\geq 30$  kg/m<sup>2</sup>.<sup>20,21</sup> Meta-analyses combining  
110 results from randomized controlled trials and observational studies found that, on average,  
111 individuals undergoing BS experienced a weight loss of 26 kg.<sup>22,23</sup> However, such an invasive  
112 procedure inherently carries increased risks of severe adverse events.<sup>24,25</sup>

113  
114 Despite the robust weight loss effects observed with both GLP1-RA treatment and BS, there  
115 exists heterogeneity in their effectiveness. One out of every five patients undergoing BS may not  
116 achieve the desired weight loss within the first year or may experience weight regain within two  
117 years.<sup>26-29</sup> The variability of GLP1-RA treatment effects on glycemic outcomes has been widely  
118 investigated and can be partially explained by clinical indicators like low  $\beta$ -cell function or C-  
119 peptide levels as well as genetic variants.<sup>30-35</sup> However, with the exception of type 2 diabetes,

120 sex, and baseline weight, the factors influencing the diversity in weight loss effects of these drugs  
121 are still poorly understood.<sup>36–38</sup> Additionally, differences in patient adherence to treatment  
122 regimens may further contribute to this variability.<sup>39</sup>

123

124 Genetics has been considered a potential factor for heterogeneity in weight loss treatments, but  
125 its role remains largely unexplored.<sup>36</sup> A previous study has examined the role of a common  
126 missense variant in the *GLP1R* genes in 57 women and found a weak link with GLP1-RA  
127 treatment response.<sup>32,40</sup> A greater number and better-powered studies have been conducted in  
128 individuals with BS<sup>41,42</sup>, however, a consistent genome-wide association for weight loss has yet  
129 to be identified.

130

131 Finally, there has been limited exploration into the effectiveness of these obesity treatments  
132 across different genetic ancestries and across various healthcare settings.<sup>10,43</sup>

133

134 In our study we seek to i) characterize the observed GLP1-RA and BS body weight-lowering  
135 effects using real-world data from 10 960 individuals across 9 biobanks in 6 countries ,with 6  
136 major continental ancestry groups; ii) identify whether plausible genetic factors (i.e. polygenic  
137 scores for BMI and T2D and coding variants in *GLP1R*, *PCSK1* and *APOE* gene) associate with  
138 heterogeneity in GLP1-RA and BS weight loss effects, and thus identify patient groups that will  
139 benefit most from such treatments; iii) compare the effects of GLP1-RA treatment and BS on  
140 weight loss across major continental ancestry groups.

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142

## 143 Results

### 144 GLP1-RA and bariatric surgery associated weight loss in 10 960 145 individuals from 9 cohorts and 6 major continental ancestry groups

146 We investigated changes in body weight associated with GLP1-RA treatment and BS across 9  
147 biobank studies: Helsinki University Hospital (HUS, Finland; N = 633), Estonian Biobank (ESTBB;  
148 N = 464), UK Biobank (UKBB, United Kingdom; N = 810), All of Us (AoU, USA; N = 559), BioMe  
149 Biobank (BioMe, USA; N = 2,170), Mass General Brigham Biobank (MGBB, USA; N = 2,141),  
150 Atlas Biobank (UCLA-ATLAS, USA; N = 1,445), Qatar Biobank (QBB, Qatar; N = 2,383), Bialystok  
151 Bariatric Surgery Study (BBSS, Poland, N = 355). We included individuals that were at least 18  
152 years of age when starting a GLP1-RA treatment or undergoing BS. Individuals undergo GLP1-  
153 RA treatment for different medical reasons, not only for weight loss. We defined GLP1-RA usage  
154 based on prescription or purchase data (ATC codes starting with A10B\*; including exenatide,  
155 liraglutide, lixisenatide, albiglutide, dulaglutide, semaglutide, and beinaglutide) covering at least  
156 12 months, and excluding individuals with indication of treatment discontinuation within this time  
157 window. For BS, we considered individuals undergoing the following procedures: Roux-en-Y  
158 gastric bypass, sleeve gastrectomy, adjustable gastric band, vertical banded gastroplasty,  
159 biliopancreatic diversion with duodenal switch. More information can be found in the **Methods**  
160 section.

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162 We included only individuals who had an initial body weight measurement recorded at most 12  
163 months prior to the baseline (defined as treatment initiation or surgery) and a follow-up  
164 measurement taken at least 6 but no more than 12 months after the baseline for GLP1-RA  
165 treatments, or within 48 months after baseline for BS. In case of multiple measurements within  
166 this time frame, we considered the median of such measurements (**Supplementary Figure 1** for  
167 a schematic of the approach used). We defined our main outcome of interest as the percentage  
168 change in body weight, calculated as the difference between the second body weight  
169 measurement (or median of multiple measurements) and the initial one, divided by the initial  
170 measurement .

171  
172 Our study population included a total of 6,750 GLP1-RA users and 4,210 individuals undergoing  
173 BS (baseline characteristics are reported in **Table 1**). Overall, GLP1-RA users 1) had a lower

174 proportion of women (61.0% weighted average across studies) when compared to BS patients  
 175 (73.1% weighted average across studies), 2) were on average older (57.30 vs 46.50), and 3) had  
 176 a lower initial body weight (99.80 kg vs 116.96 kg). The proportion of individuals diagnosed with  
 177 type 2 diabetes at time of GLP1-RA initiation was at least 48% across all studies and major  
 178 continental ancestry groups.  
 179

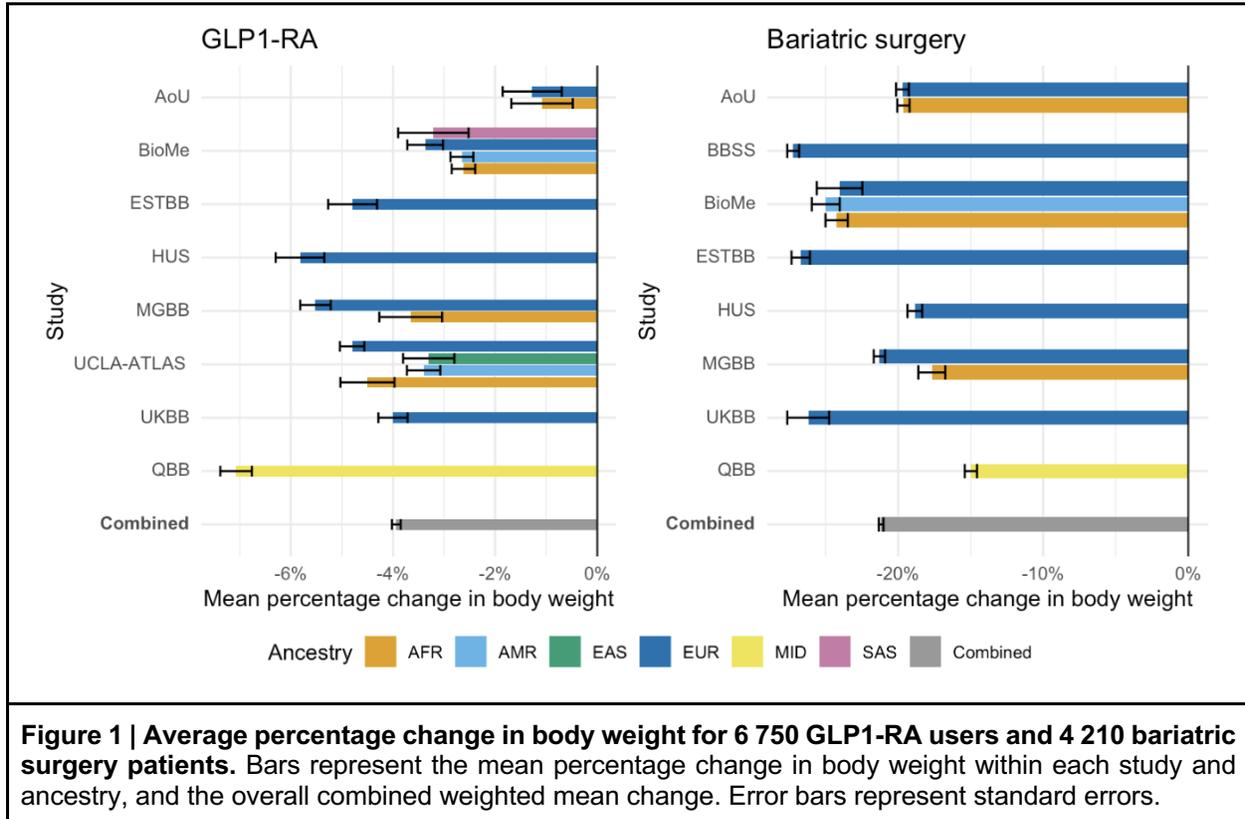
Study	Ancestry	Number of individuals	Proportion of women (%)	Mean body weight (in Kg) at baseline (SD)	Mean age at baseline (SD)	T2D prevalence (%)	Most common GLP1-RA type	Proportion of semaglutide (%)
<b>GLP1-RA</b>								
HUS	EUR	255	68.00	113.70 (22.90)	53.79 (13.52)	65.49	semaglutide	60.40
ESTBB	EUR	191	62.30	109.18 (20.3)	56.95 (10.96)	89.01	semaglutide	72.77
UKBB	EUR	614	44.30	105.00 (19.90)	60.50 (7.15)	99.80	liraglutide	0.00
AoU	EUR	117	64.96	106.66 (24.18)	57.01 (11.03)	82.91	liraglutide	11.97
	AFR	91	83.52	104.8 (25.01)	54.57 (11.93)	92.31	liraglutide	18.68
MGBB	EUR	857	57.01	99.99 (21.48)	59.88 (12.17)	56.33	semaglutide	40.72
	AFR	156	54.32	104.65 (23.80)	54.31 (11.91)	60.24	dulaglutide	36.14
BioMe	AMR	614	66.45	91.19 (22.45)	59.07 (11.29)	67.75	dulaglutide	29.32
	AFR	728	69.78	101.49 (25.42)	57.00 (12.24)	69.50	dulaglutide	29.12
	EUR	369	44.44	98.26 (21.93)	60.32 (11.57)	47.97	semaglutide	37.13
	SAS	77	50.65	83.79 (20.10)	56.48 (12.59)	70.13	semaglutide	41.56
UCLA-ATLAS	AFR	133	64.00	101.6 (23.20)	61.00 (13.00)	91.00	semaglutide	57.10
	AMR	336	66.00	94.40 (22.60)	56.00 (13.00)	84.00	liraglutide	26.50
	EAS	120	47.00	81.50 (16.30)	60.00 (13.00)	92.00	semaglutide	61.70
	EUR	856	53.00	101.60 (22.00)	63.00 (13.00)	80.00	semaglutide	67.80
QBB	MID	1236	72.17	97.68 (24.54)	49.52 (10.75)	75.97	liraglutide	3.64
<b>Bariatric surgery</b>								
HUS	EUR	378	81.00	113.63 (19.5)	49.90 (9.39)	37.04		
ESTBB	EUR	273	76.56	121.24 (20.57)	44.72 (10.07)	28.21		
UKBB	EUR	196	69.40	121.00 (29.70)	55.50 (8.21)	50.00		
AoU	EUR	168	81.55	121.32 (29.88)	50.74 (12.62)	18.45		
	AFR	183	89.62	127.23 (32.55)	46.09 (11.67)	48.63		
MGBB	EUR	978	71.06	121.21 (26.52)	49.37 (12.16)	39.98		
	AFR	150	86.67	122.89 (23.92)	43.68 (11.67)	42.67		
BioMe	AMR	143	79.72	112.36 (20.81)	44.62 (11.48)	30.07		
	AFR	188	82.98	124.88 (25.63)	45.35 (12.16)	29.26		
	EUR	51	70.59	120.23 (25.29)	49.51 (13.34)	17.65		
QBB	MID	1147	69.49	102.14 (23.18)	41.66 (10.66)	40.02		
BBSS	EUR	355	56.06	138.48 (27.07)	47.10 (7.11)	35.21		

**Table 1 | Baseline characteristics for 6 750 GLP1-RA users and 4 210 patients undergoing bariatric surgery included in our study population.** For each cohort and major continental ancestry group (AFR: African, AMR: admixed American, EAS: East Asian, EUR: European, MID: Middle Eastern, SAS: South Asian), we report the total number of individuals eligible for analysis, the proportion of females, the mean (SD) body weight (in Kg) at baseline, the mean (SD) age (in years) at baseline, the prevalence of T2D, and (for GLP1-RA users only) the most commonly used type of medication and the proportion of semaglutide users.

HUS, Helsinki University Hospital; ESTBB, Estonian Biobank; UKBB, UK Biobank; AoU, All of US; MGBB, Mass General Brigham Biobank; BioMe, BioMe; UCLA-ATLAS, Atlas Biobank; QBB, Qatar Biobank; BBSS, Bialystok Bariatric Surgery Study

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Among GLP1-RA users, we observed an average body weight change (**Figure 1**) across studies of -3.93% (ranging from -1.08% to -7.07%), in line with results from observational studies.<sup>44,45</sup> Among BS patients, we observed an average change of -21.17% (-15.00% to -27.72%).



## 186 Demographic factors associated with GLP1-RA and bariatric 187 surgery weight change

188 We explored the effect of baseline characteristics on body weight changes associated with GLP1-  
189 RA treatment and BS. We fitted a multivariable linear model with percentage change in body  
190 weight as outcome and baseline body weight, sex, age at baseline, the first 20 principal genetic  
191 components and medication type (for GLP1-RA users only) as predictors. Each cohort was  
192 analyzed separately, stratifying by major continental ancestry group, and the effects were then  
193 meta-analyzed (**Table 2**). Baseline body weight was significantly associated with greater weight  
194 loss for both GLP1-RA and BS treatments ( $\beta_{GLP1-RA} = -0.05$  % weight change compared to  
195 baseline,  $P = 1.63 \times 10^{-35}$ ;  $\beta_{BS} = -0.14$  % weight change compared to baseline,  $P = 5.49 \times 10^{-69}$ )  
196 and women had a significantly greater weight loss for both treatments ( $\beta_{GLP1-RA} = -1.54$  % weight

197 change compared to baseline,  $P = 2.47 \times 10^{-16}$ ;  $\beta_{BS} = -2.66$  % weight change compared to  
198 baseline,  $P = 6.83 \times 10^{-10}$ ). Being older at treatment initiation was only associated with less weight  
199 loss for bariatric surgery ( $\beta_{BS} = 0.15$  % weight change compared to baseline,  $P = 2.54 \times 10^{-15}$ ).  
200 The effects of baseline weight and sex on weight loss were highly heterogeneous across  
201 ancestries and studies ( $I^2$  for GLP1-RA= 0.99 and 0.83 for baseline weight and sex, respectively)  
202 (Supplementary Figures 2-4 and Supplementary Tables 1-2).

203  
204

Variable	GLP1-RA			Bariatric surgery		
	Effect on percentage weight change	Standard error	P value	Effect on percentage weight change	Standard error	P value
Baseline weight (kg)	-0.05	0.004	$1.63 \times 10^{-35}$	-0.14	0.008	$5.49 \times 10^{-69}$
Sex (female)	-1.54	0.188	$2.74 \times 10^{-16}$	-2.66	0.431	$6.83 \times 10^{-10}$
Age at baseline (years)	-0.02	0.008	0.052	0.15	0.016	$9.85 \times 10^{-19}$

**Table 2 | Effect of baseline weight, sex and age at baseline of percentage weight changes associated with GLP1-RA and bariatric surgery.** Multivariable model adjusting for the three variables described in the table and the first 20 principal genetic components and medication type (for GLP1-RA users only). Negative effects indicate larger weight loss.

205

## 206 Marginal effect of ancestry on GLP1-RA and bariatric surgery 207 associated weight loss

208 We evaluate whether the effect of GLP1-RA and BS associated weight loss was significantly  
209 different across genetically-defined continental ancestry groups after accounting for differences  
210 in weight at baseline, age, sex and, only for GLP1-RA, medication type (Table 3). Only multi-  
211 ancestry biobanks (AoU, BioMe, MGBB, UCLA-ATLAS) were included in this analysis (N = 6 315).  
212 Compared to individuals of European ancestry, all other ancestry groups had lower change in  
213 body weight. However, this effect was statistically significant only among GLP1-RA users of  
214 African ( $\beta = 0.76$  % weight change from baseline weight,  $P = 0.02$ ) and admixed American ( $\beta =$   
215  $0.87$  % weight change from baseline weight,  $P = 3 \times 10^{-3}$ ) ancestries.

<b>GLP1-RA</b>			
<b>Major continental ancestry group</b>	<b>Effect on percentage weight change</b>	<b>Standard error</b>	<b>P value</b>
AFR	0.758	0.327	0.021
AMR	0.874	0.295	0.003
EAS	1.093	0.664	0.100
SAS	0.118	0.759	0.876
<b>Bariatric surgery</b>			
<b>Major continental ancestry group</b>	<b>Effect on percentage weight change</b>	<b>Standard error</b>	<b>P value</b>
AFR	1.688	1.308	0.197
AMR	-0.542	1.428	0.704

**Table 3 | Associations between weight change and major continental ancestry groups.** Meta-analysis effect sizes, standard errors and P values for the association between weight change and major continental ancestry groups (AFR: African, AMR: admixed American, EAS: East Asian, EUR: European, SAS: South Asian). Coefficients are derived from a multivariable model including percentage weight change as outcome and ancestry as categorical variable (with EUR as reference category), and adjusting for sex, age and weight at baseline and, for GLP1-RA only, medication type. Negative effects indicate larger weight loss

216

## 217 Plausible genetic factors do not associate with weight change due 218 to GLP1-RA

219 We investigated the effect of genetic exposures with plausible effects on body weight changes.  
220 Genetic exposures were selected upon discussion between authors and before initiating  
221 analyses. First, we considered polygenic scores for BMI based on previous observation that a  
222 polygenic score was associated with BMI trajectories.<sup>46,47</sup> We also consider the only replicated  
223 genome-wide variant associated with BMI change in the largest study of BMI trajectories  
224 (*rs429358* in *APOE*).<sup>46</sup> Second, we considered a polygenic score for T2D based on the  
225 observation that biomarkers of type 2 diabetes were associated with effectiveness of GLP1-  
226 RA.<sup>31,33</sup> We confirmed that both polygenic scores for BMI and T2Ds were significantly associated  
227 with their respective traits across all major continental ancestry groups (**Supplementary Table**  
228 **3**) although the effects were larger in individuals of European ancestry.

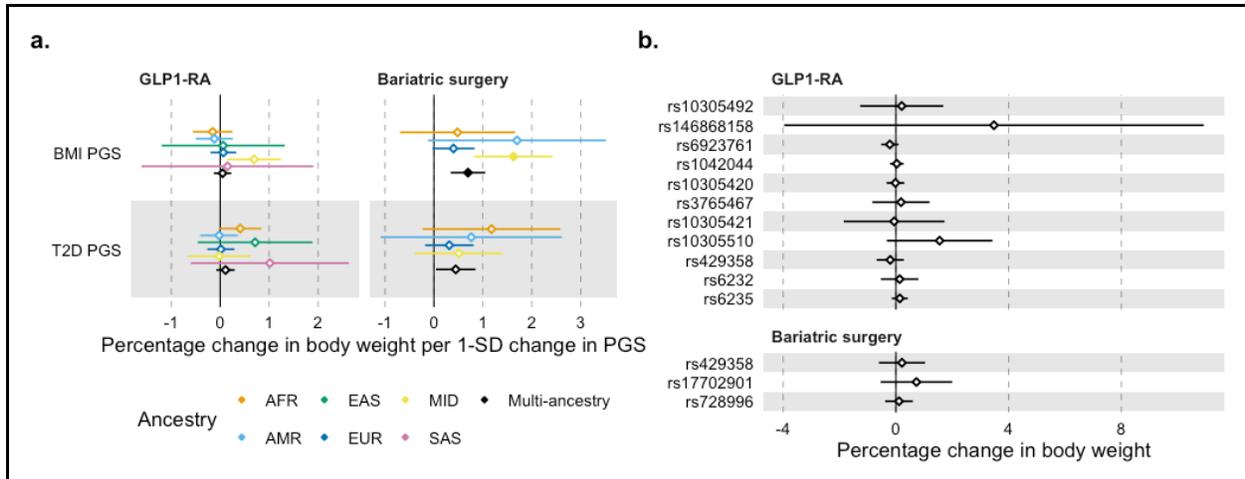
229 Third, we considered 13 SNPs in the *GLP1R* and *PCSK1* genes. We considered all missense  
230 variants in the *GLP1R* gene with at least 1% minor allele frequency in at least one of major  
231 continental ancestry group, as per gnomAD (version 4.1)<sup>48</sup>, three of which have been functional  
232 characterized as gain or loss of function (**Supplementary Table 4**).<sup>49</sup> We also considered two  
233 non-synonymous variants in the *PCSK1* gene that have been associated with BMI variation in the  
234 population.<sup>50</sup> Finally, we sought to replicate the effect of two variants associated with weight loss  
235 in one of the largest studies of BS<sup>42</sup>(**Methods**). After analyses, two missense variants in the  
236 *GLP1R* genes, *rs2295006* and *rs201672448* were too rare to be tested across studies and were  
237 not considered further.

238  
239 We estimated the effect of each genetic exposure on weight change in each cohort and major  
240 continental ancestry groups separately, fitting a linear model adjusted for baseline body weight,  
241 sex, age at baseline, first 20 principal genetic components, and medication type (for GLP1-RA  
242 users only), and we subsequently combined the effects through ancestry-specific and multi-  
243 ancestry meta-analyses (**Figure 2, Supplementary Tables 5-6**). As the approach of adjusting  
244 longitudinal change phenotypes for the baseline trait has been criticized for potentially inducing  
245 biases for genetic variants associated with the baseline trait<sup>51</sup>, we also considered a second model  
246 using post-treatment weight as outcome, including the same predictors as in the model above  
247 and an additional term modeling the interaction between baseline weight and each genetic  
248 exposure (**Supplementary Tables 7-8**).

249  
250 Among GLP1-RA users, we did not observe any significant genetic exposure associated with  
251 weight loss after multiple testing correction ( $P < 0.004$ , Bonferroni correction for 13 exposures  
252 tested). Using the alternative model, we identified one missense variant in the *GLP1R* genes,  
253 *rs3765467*, for which we observed a statistically significant interaction with baseline weight  
254 ( $\beta_{w0*rs3765467} = 0.09$  kg weight change from baseline weight,  $P = 2.7 \times 10^{-4}$ ) on the weight measured  
255 post GLP1-RA treatment after multiple testing correction. When considering nominal significance,  
256 another missense variant in the *GLP1R* genes, *rs6923761*, had a statistically significant  
257 interaction with baseline weight ( $\beta_{w0*rs6923761} = 0.02$  kg weight change from baseline weight,  $P =$   
258  $0.01$ ) on the weight measured post GLP1-RA treatment.

259  
260 We further tested whether the lack of associations observed was possibly due to high effect-size  
261 heterogeneity across ancestries, or due to lack of statistical power. We did not observe any  
262 significant heterogeneity across ancestries for any of the genetic exposure (average  $I^2$  across

263 genetic exposures for GLP1-RA was 0.06 - **Supplementary Tables 9-10**). Moreover, given the  
264 current sample size for GLP1-RA users (N = 6 750) our power calculation showed we would have  
265 at least 80% power to detect a statistically significant effect (at  $P < 0.05$ ) of at least a 0.3% change  
266 in body weight for 1 standard deviation (SD) change in the polygenic score, or for SNPs with a  
267 minor allele frequency of at least 1% (**Methods**).  
268  
269



**Figure 2 | Effect of 15 genetic exposures on body weight changes associated with GLP1-RA treatment and bariatric surgery.**

a. Ancestry-specific and multi-ancestry meta-analysis effect sizes for association between percentage change in body weight and PGS for BMI and type 2 diabetes. Dots represent the percentage change in body weight per one standard deviation change in PGS, error bars represent the 95% confidence interval. Full dots represent statistical significance at  $P < 0.004$ .

b. Multi-ancestry meta-analysis effect sizes for the association between percentage change in body weight and genotype at each locus. Dots represent the effect size, error bars represent the 95% confidence interval. PGS, polygenic score.

270  
271

## 272 A polygenic score for BMI associates with weight change due to 273 BS

274  
275 Among BS patients, a higher polygenic score for BMI was significantly associated with lower  
276 weight loss ( $\beta_{BMI\ PGS} = 0.70\%$  weight change compared to baseline for 1 SD change in the  
277 polygenic score,  $P = 1.24 \times 10^{-4}$ ) after multiple testing correction ( $P < 0.01$ , Bonferroni correction  
278 for 5 exposures tested), in the opposite direction than observed for weight at baseline (**Figure 2**).

279 A higher polygenic score for T2D was nominally associated with lower weight loss ( $\beta_{T2D\ PGS} = 0.45$   
280 % weight change compared to baseline for 1 SD change in the polygenic score,  $P = 0.03$ ).  
281 We observed significant heterogeneity in the association between BMI PGS and weight loss  
282 across studies ( $I^2 = 0.64$ , **Supplementary Table 10**).

## 283 Discussion

284 In this study we provide a comprehensive assessment of major demographic factors and plausible  
285 genetic factors on body weight changes associated with GLP1-RA use and BS.

286 In line with previous work,<sup>36,52-55</sup> our main outcome of interest was the percentage change in body  
287 weight between 6 and 12 months after GLP1-RA treatment initiation as well as between 6 and 48  
288 months after BS.

289  
290 By considering studies from 6 different countries, we were able to highlight the heterogeneity in  
291 use of GLP1-RA across different healthcare systems. GLP1-RA users consisted mostly of  
292 individuals with type 2 diabetes, reflecting their original use for treatment of this condition. Biobank  
293 studies covering Boston (MGBB), New York (BioMe), Los Angeles (UCLA-ATLAS) and Helsinki  
294 (HUS) had overall lower rates of type 2 diabetes individuals compared to the other studies and  
295 lower weight at baseline, which, together with the availability of more updated electronic health  
296 records data, suggest a larger number of individuals being prescribed GLP1-RA for weight loss.  
297 However, care should be taken in interpreting variation in rate of type 2 diabetes across biobanks  
298 as this might reflect different approaches used for capturing this diagnosis from the electronic  
299 health records as well as different sampling strategies of biobanks.

300  
301 The comparison between GLP1-RA and BS is informative as it highlights the younger age, higher  
302 weight and lower percentage of men undergoing BS compared to GLP1-RA users. The weight  
303 reduction was also approximately 5 times higher among BS patients compared to GLP1-RA  
304 users. The weight reduction among GLP1-RA users was also more heterogeneous across  
305 biobanks compared to BS, but in line with what reported previously in the literature using real  
306 world data<sup>34</sup> and lower than results from clinical trials.<sup>11,56</sup> Our findings also highlight a consistent  
307 effect of baseline body weight and sex on weight loss associated with both GLP1-RA treatment  
308 and BS, suggesting that individuals with higher weight at baseline, as well as women benefit more  
309 from both interventions. These findings underscore the importance of considering individual  
310 characteristics when designing weight loss interventions.

311  
312 The main aim of this study was to evaluate the impact of genetic exposures on body weight  
313 changes induced by GLP1-RA or BS. At the current sample size, a genome-wide scan would  
314 have limited power to identify potential new signals. With 6,750 individuals, we would have at  
315 least 80% statistical power to identify, at a genome-wide significant threshold, a genetic variant  
316 with an allele frequency of at least 1% and an effect larger than 0.3% weight change per change  
317 in one SD, which is unlikely given previous effects of common variants on treatment response.<sup>51</sup>  
318 Instead, we formulated a set of hypotheses based on the results from the largest genome-wide  
319 association studies of BMI and type 2 diabetes, and, explored biologically plausible mutations in  
320 several candidate genes. Polygenic scores represent the most powerful genetic predictors for  
321 BMI and type 2 diabetes and we confirm their significant association with the corresponding trait  
322 across all major continental ancestry groups. Thus, lack of associations with these polygenic  
323 scores would suggest that GLP1-RA and BS-related weight loss is independent of one's genetic  
324 predisposition to BMI or type 2 diabetes, as explained by common genetic variants. Moreover,  
325 because the biological underpinning of weight change is partially distinct from weight at baseline,  
326 we also consider the few replicated genetic signals for BMI variability and BMI trajectories.

327  
328 For bariatric surgery we observed that genetically higher BMI was associated with weight gain, in  
329 the opposite direction for the phenotypic association with initial body weight. This might suggest  
330 that individuals with a higher genetic predisposition for elevated BMI might regain weight, after  
331 the operation, faster than those with lower genetic predisposition. To maintain weight loss, they  
332 may need additional lifestyle modifications, behavioral health therapy or adjuvant pharmacologic  
333 therapies as a GLP1-RA treatment.<sup>57,58</sup>

334  
335 On the contrary, none of the genetic exposures were associated with weight loss among GLP1-  
336 RA users, after multiple-testing correction. These results were not attributable to high  
337 heterogeneity across studies or across major continental ancestry groups. Whether we directly  
338 tested the impact of genetic exposure on percentage body weight change or used post-treatment  
339 weight as outcome and tested the interaction between baseline weight and the genetic exposure,  
340 results did not significantly change. In light of the associations observed for BS and following the  
341 hypothesis that individuals with genetically higher BMI tend to re-gain weight faster, it would not  
342 be surprising to find, at larger sample size, a significant association between BMI polygenic score  
343 and weight loss due to GLP1-RA. However, our analyses exclude that this effect, should it exist,  
344 is unlikely to guide clinical treatment assignment. Our results showed that the considered genetic

345 influence on weight loss in response to GLP1-RA treatment may be too small to impact clinical  
346 decision-making or patient outcomes. This statement should be revisited either by using more  
347 powerful polygenic scores or by considering individuals undergoing GLP1-RA treatment only for  
348 weight loss.

349  
350 Another aim of the study was to compare the effects of GLP1-RA treatment and BS on weight  
351 loss across major continental ancestry groups. Ancestry can provide more useful information  
352 about population health than racial and ethnic categories because it is not directly influenced by  
353 geographic, cultural, and sociopolitical forces.<sup>59</sup> Nonetheless, ancestry is also partially correlated  
354 with self-reported race and ethnicity and socio-economic characteristics do differ across racial  
355 and ethnic groups. Overall, we did not find large differences in GLP1-RA and BS associated  
356 weight loss between major continental ancestry groups. We do observe that GLP1-RA users of  
357 African and admixed American ancestry had statistically significantly lower change in body weight  
358 compared to Europeans, after accounting for differences in sex, age, weight at baseline and  
359 medication type. The effects were consistent in BioMe and UCLA-ATLAS biobanks for admixed  
360 American ancestry, and additionally in AoU and MGBB for African ancestry. This might reflect  
361 different socio-economic characteristics between individuals of admixed American and European  
362 ancestry resulting, for example, in differences in access to treatment.

363  
364 The main strength of our study is its size and diversity across countries, healthcare systems and  
365 ancestries. Our study presents various limitations. First, we cannot be sure that all GLP1-RA  
366 users fulfilled their prescription. This is a potential issue only in studies with prescription  
367 information but not in studies with purchase information. Nonetheless weight loss was comparable  
368 across both study types. Moreover, we ensured that observed changes in body weight are due to  
369 the treatment by considering only measurements within one month from the last purchase or  
370 prescription. Second, we could not accurately estimate treatment adherence, but overcome this  
371 limitation by considering treatments of maximum 12 months, a period for which adherence has  
372 been observed to be around 65%.<sup>52</sup> Third, we used ancestry labels based on major continental  
373 ancestry groups, while ancestry can be better characterized as a continuous quantity.<sup>59</sup> This  
374 discretization reflects pragmatic considerations to enable comparable analyses across the  
375 different studies.

376  
377 In conclusion, our study suggests that GLP1-RA treatments work equally well in individuals  
378 carrying common non-synonymous mutations in the GLP1R gene and in individuals at high

379 genetic risk for BMI and type 2 diabetes. It also suggests that BS is less effective among  
380 individuals with higher polygenic score for BMI, but the effect is not large enough to be clinically  
381 relevant. In fact, for each additional SD in the polygenic score for BMI, individuals undergoing BS  
382 have on average 0.7% lower weight loss from their initial body weight. Finally, these two major  
383 weight-loss interventions achieved sustained weight reduction regardless of an individual's  
384 genetic ancestry. However, socio-economic characteristics correlating with ancestry should be  
385 further studied to better understand some of the remaining differences observed in this study.

## 386 Methods

### 387 Study population

388 In the current study, we included samples from 10 960 individuals from the following 9 biobanks:  
389 Helsinki University Hospital (HUS, Finland; N = 633), Estonian Biobank (ESTBB, Estonia; N =  
390 464), UK Biobank (UKBB, United Kingdom; N = 810), All of Us (AoU, USA; N = 559), BioMe  
391 (BioMe, USA; N = 2,170), Mass General Brigham Biobank (MGBB, USA; N = 2,141), Atlas  
392 Biobank (UCLA-ATLAS, USA; N = 1,445), Qatar Biobank (QBB, Qatar; N = 2,383), Bialystok  
393 Bariatric Surgery Study (BBSS, Poland, N = 355) . The biobank studies include samples from  
394 (hospital) biobanks, prospective epidemiological and disease-based cohorts. Follow-up covers a  
395 total of 20 years with the earliest study starting follow-up in 2004 (AoU) and the latest study ending  
396 follow-up in 2024 (BioMe). In the **Supplementary Material** we provide a detailed description of  
397 the population selected in each study.

### 398 Inclusion and exclusion criteria for GLP1-RA analysis

399 For our GLP1-RA analysis we implemented the following inclusion and exclusion criteria. We  
400 required an initiation of a GLP1-RA treatment (ATC codes: A10BJ\*, A10BX04, A10BX10,  
401 A10BX13, A10BX14) defined as the date of first medication purchase or medication prescription  
402 of GLP1-RA. Individuals were required to be on a GLP1-RA treatment for at least twelve months.  
403 This was defined through prescription length, regular drug prescription refills or drug purchases,  
404 depending on the dataset and healthcare setting.  
405 Individuals aged under 18 were excluded from the analysis. Furthermore, individuals were  
406 required to have at least one body weight measurement maximum one year prior to treatment  
407 initiation or within 14 days after initiation, if no measure before initiation was available. If we  
408 observed multiple weight measurements within this time frame, only the closest one to initiation  
409 was considered. This measurement defined our baseline weight variable. At least one body  
410 weight measurement between 26 weeks and 52 weeks was necessary for each individual to be  
411 included in the analysis. Individuals were excluded if they underwent BS prior to or during the first  
412 year of GLP1-RA treatment.

413

## 414 Inclusion and exclusion criteria for BS analysis

415 For our BS analysis we implemented the following inclusion and exclusion criteria. Individuals  
416 were included if they underwent any type of BS. The respected code definitions for BS varied  
417 across countries and healthcare systems and can be found in **Supplementary Table 11**.  
418 Individuals aged under 18 at baseline were excluded from the analysis. Analogously to the GLP1-  
419 RA analysis, individuals were required to have at least one body weight measurement maximum  
420 one year prior to surgery or within 14 days after it, if no measure before baseline was available.  
421 If we observed multiple weight measurements within this time frame, only the closest one to  
422 initiation was considered. This measurement defined our baseline weight variable. Additionally,  
423 at least one body weight measurement between 26 weeks and 208 weeks was necessary for  
424 each individual to be included in the analysis.

## 425 Outcome definitions

426 In our primary model, for both GLP1-RA and BS we defined the outcome to be the percentage  
427 change in body weight from baseline. The baseline body weight measured closest to  $T_0$  and  
428 between -52 weeks and +2 weeks from  $T_0$  was defined as  $W_0$ .

429 The second body weight ( $mW_1$ ) was defined as the median of body weight measurements  
430 between 26 weeks and 52 weeks from  $T_0$  for the GLP1-RA analysis and between 180 days and  
431 1460 days for the BS analysis. This was to enhance robustness against outlier measurements  
432 occurring within these intervals.

433 Therefore, the percentage change in body weight from baseline was calculated using

$$434 \quad \% \text{ weight change} = \frac{mW_1 - W_0}{W_0} * 100 .$$

435 In our secondary model, we used  $mW_1$  as an outcome.

## 436 Genetic exposures

437 We identified 15 genetic variants (single nucleotide polymorphism, SNP) and two PGS of interest  
438 for our analysis. Among the included SNPs 3 are functionally characterized *GLP1R* variants  
439 (*rs10305492*, *rs146868158*, *rs6923761*) associated with random glucose levels that have shown  
440 decreased/increased response to different endogenous and exogenous GLP1-RA.<sup>49</sup> Seven SNPs

441 are *GLP1R* missense variants (*rs1042044*, *rs10305420*, *rs3765467*, *rs10305421*, *rs2295006*,  
442 *rs10305510*, *rs201672448*) with a frequency over 1% in at least one ancestry group in gnomAD  
443 (version 4.1).<sup>48</sup> One SNP is the only robustly replicated lead genome-wide significant variant  
444 (*rs429358*, *ApoE*) in a GWAS of BMI change over time.<sup>46,60</sup> Two SNPs are *PCSK1* missense  
445 variants (*rs6232*, *rs6235*) that have been found associated with BMI variations.<sup>50</sup> And two SNPs  
446 (*rs728996*, *rs17702901*) were associated with an effect on excess body mass index loss amongst  
447 BS patients.<sup>42</sup> A list of all genetic exposures and their frequencies across all major continental  
448 ancestry groups can be found in **Supplementary Table 4**. The included PGS are for body mass  
449 index (BMI) and type 2 diabetes (T2D).<sup>61</sup> The PGS weights were taken from Weissbrod et al.  
450 2022<sup>61</sup> and obtained from UK Biobank. The scores were computed in each study using PLINK  
451 2.0.<sup>62</sup> Because the PGS scores were derived from UK Biobank, we couldn't use the same score  
452 in UK Biobank, instead we used Thompson et al. 2022.<sup>63</sup>  
453

## 454 Statistical models

455 We employed two linear regression models (A and B) to investigate the effect of the chosen  
456 genetic exposures on weight change after initiation of GLP1-RA treatment and BS, respectively.  
457 In both models we adjusted for the baseline weight, sex, age at initiation of treatment or surgery  
458 in years, the first 20 genetic principal components and study-specific covariates, such as the  
459 genotyping batch. In the GLP1-RA analysis we additionally adjusted for the medication type within  
460 the drug class.

461 In model A % *weight change* was the outcome variable, while in model B the outcome was  $mW_1$ .  
462 In model B we additionally included the interaction term between the genetic exposure variable  
463 and the baseline weight.

464

### 465 **Model A (primary model):**

466 % *weight change* =  $\beta_0 + \beta_1 * \text{genetic exposure} + \beta_2 * W_0 + \beta_3 * \text{sex} + \beta_4 * \text{age at initiation} +$   
467  $\beta_5 * \text{medication (only for GLP1)} + \sum_{k=6}^{25} \beta_k * PC1:20 + \sum_{i=26}^n \beta_i * \text{study specific covariates}$

468

### 469 **Model B (secondary model):**

470  $mW_1 = \beta_0 + \beta_1 * \text{genetic exposure} + \beta_2 * W_0 + \beta_3 * \text{genetic exposure} * W_0 +$   
471  $\beta_4 * \text{sex} + \beta_5 * \text{age at initiation} + \beta_6 * \text{medication (only for GLP1)} + \sum_{k=7}^{26} \beta_k * PC1:20 +$

472

$$\sum_{i=27}^n \beta_i * \text{study specific covariates}$$

473

## 474 Meta-analyses

475 We conducted ancestry-specific and multi-ancestry fixed-effect meta-analyses for each of the  
476 genetic exposures. Effect sizes were combined using fixed-effect inverse-variance-weighting  
477 (IVW), as implemented in the R package “meta”.<sup>64</sup> For each exposure we reported the unadjusted  
478  $P$  values, but considered as threshold for statistical significance  $P < 0.004$  for GLP1-RA (i.e.  $P <$   
479  $0.05/13$ , Bonferroni correction for 13 exposures tested) and  $P < 0.01$  for BS (i.e.  $P < 0.05/5$ ,  
480 Bonferroni correction for 5 exposures tested). We tested for heterogeneity in effect sizes between  
481 ancestries using Cochran’s Q-test<sup>65</sup> and by inspecting the  $I^2$  statistics<sup>66</sup>, which is defined as:

482

483

$$I^2 = \left( \frac{Q - df}{Q} \right) * 100\%$$

484

485 where  $df$  is equal to *number of studies* – 1.

## 486 Power calculation

487 To calculate the statistical power for the PGS effect sizes, we used a t-test for linear regression  
488 coefficient as implemented in the R package `pwr`<sup>67</sup> assuming a standard deviation for the  
489 outcome variable  $SD_y = 7.8$  (weighted combine SD across studies), a sample size of  $n = 6\,750$ ,  
490 a total number of predictors  $k = 25$  and an adjusted  $r^2 = 0.05$ , to calculate power at  $P < 0.05$   
491 significance level.

492

493 For the coefficients of association with single variants, we estimated statistical power via the non-  
494 centrality parameter (NCP) of the chi-square distribution. We defined  $NCP = 2 f (1 - f) n \beta^2$   
495 where  $f$  is MAF,  $n$  is the effective sample size, and  $\beta$  is the expected effect size.<sup>68</sup>

496 We set a beta of 0.3 (% weight change from baseline) as the expected genetic effect on GLP1  
497 response, based on findings from a recent large-scale study on genetic influences on drug  
498 response.<sup>51</sup>

499

## 500 Code Availability

501 The study utilized previously published analysis tools as described in the **Methods** section.  
502 Additional code used for these analyses is available at [https://github.com/dsgelab/glp1-bs-](https://github.com/dsgelab/glp1-bs-genetics/)  
503 [genetics/](https://github.com/dsgelab/glp1-bs-genetics/)  
504

## 505 Data Availability

506 The Helsinki biobank can provide access for research projects within the scope regulated by the  
507 Finnish Biobank Act, which is research utilizing the biobank samples or data for the purposes of  
508 promoting health, understanding the mechanisms of disease or developing products and  
509 treatment practices used in health and medical care.

510 De-identified data of the MGBB that supports this study is available from the MGB Biobankportal  
511 at [https://www.massgeneralbrigham.org/en/research-and-innovation/participate-in-](https://www.massgeneralbrigham.org/en/research-and-innovation/participate-in-research/biobank/for-researchers)  
512 [research/biobank/for-researchers](https://www.massgeneralbrigham.org/en/research-and-innovation/participate-in-research/biobank/for-researchers). Restrictions apply to the availability of these data, which are  
513 available to MGB-affiliated researchers via a formal application.

514 UK Biobank data are available through a procedure described at  
515 <http://www.ukbiobank.ac.uk/using-the-resource/>.

516 Registered researchers whose institutions have Data Use and Registration Agreements in place  
517 with All of Us that include the Controlled Tier can access genomic data.

518 Clinical and genotype data from UCLA ATLAS Community Health Initiative patients, de-identified  
519 for research purposes, are accessible to UCLA-approved researchers through the Discovery Data  
520 Repository (DDR).

521 BioMe data is available through a process described at  
522 <https://icahn.mssm.edu/research/ipm/programs/biome-biobank/researcher-faqs>.

523 Pseudonymised data and/or biological samples can be accessed for research and development  
524 purposes in accordance with the Estonian Human Genome Research Act. To access data, the  
525 research proposal must be approved by the Scientific Advisory Committee of the Estonian  
526 Biobank as well as by the Estonian Committee on Bioethics and Human Research. For more  
527 details on data access and relevant documents, please see  
528 <https://genomics.ut.ee/en/content/estonian-biobank#dataaccess>.

529 For access to the Bialystok Bariatric Surgery Study please refer to the Medical University of  
530 Bialystok.

531 The data and biosamples collected or generated by QBB will be made available to researchers  
532 employed within or otherwise contractually bound to public and private institutions that conduct  
533 scientific research and that meet the requirements detailed in the Qatar Biobank Research Access  
534 policy.  
535 The PGS weights used in this study are available in Weissbrod et al. 2022 and Thompson et al.  
536 2022.<sup>61</sup>

## 537 Acknowledgements

538 We gratefully acknowledge the resources provided by the Institute for Precision Health and  
539 participating patients from the UCLA ATLAS Community Health Initiative  
540 ([https://www.uclahealth.org/international-services/consulting-services/clinical-research-](https://www.uclahealth.org/international-services/consulting-services/clinical-research-consultation/ucla-atlas-community-health-initiative)  
541 [consultation/ucla-atlas-community-health-initiative](https://www.uclahealth.org/international-services/consulting-services/clinical-research-consultation/ucla-atlas-community-health-initiative)). The UCLA ATLAS Community Health  
542 Initiative in collaboration with UCLA ATLAS Precision Health Biobank is a program of the Institute  
543 for Precision Health, which directs and supports the biobanking and genotyping of biospecimen  
544 samples from participating patients from UCLA in collaboration with the David Geffen School of  
545 Medicine, UCLA Clinical and Translational Science Institute and UCLA Health. The ATLAS  
546 Community Health Initiative is supported by UCLA Health, the David Geffen School of Medicine  
547 and a grant from the UCLA Clinical and Translational Science Institute (UL1TR001881). This work  
548 was financially supported in part by National Institutes of Health awards U01HG011715,  
549 R01HG009120 and R01MH115676. The content is solely the responsibility of the authors and  
550 does not necessarily represent the official views of the National Institutes of Health.

551 We would furthermore like to thank all participants of HUS  
552 (<https://www.helsinginbiopankki.fi/en/Helsinki-Biobank>; <https://www.hus.fi/en>), ESTBB  
553 (<https://genomics.ut.ee/en/content/estonian-biobank>), BBSS  
554 (<https://classic.clinicaltrials.gov/ct2/show/NCT04634591>), QBB (<https://www.qphi.org.qa/>), BioMe  
555 (<https://icahn.mssm.edu/research/ipm/programs/biome-biobank>), UKBB  
556 (<https://www.ukbiobank.ac.uk/>), MGBB ([https://www.massgeneralbrigham.org/en/research-and-](https://www.massgeneralbrigham.org/en/research-and-innovation/participate-in-research/biobank)  
557 [innovation/participate-in-research/biobank](https://www.massgeneralbrigham.org/en/research-and-innovation/participate-in-research/biobank)) and AOU (<https://allofus.nih.gov/>) for their generous  
558 contribution. This work was further supported in part by funding from the Eric and Wendy Schmidt  
559 Center at the Broad Institute of MIT and Harvard.

560 U.V. and K.A. have been funded by Estonian Research Council's personal research funding start-  
561 up grant PSG759 and European Research Council' starting grant under the grant agreement no  
562 101117251 (OBECAUSE). J.H.L. is supported by NIDDK K23 DK131345.

563 The activities of the EstBB are regulated by the Human Genes Research Act, which was adopted  
564 in 2000 specifically for the operations of the EstBB. Individual level data analysis in the EstBB  
565 was carried out under ethical approvals of 1.1-12/1409 and 1.1-12/2161 from the Estonian  
566 Committee on Bioethics and Human Research (Estonian Ministry of Social Affairs), using data  
567 according to release application 6-7/GI/18857 from the Estonian Biobank.  
568

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578 [attempted-to-lose-weight-in-the-past-year-but-most-have-been-unsuccessful/](https://easo.org/european-survey-finds-over-three-quarters-of-adults-with-obesity-have-attempted-to-lose-weight-in-the-past-year-but-most-have-been-unsuccessful/) (2022).
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590 [announce-fda-approval-byettatm-exenatide](https://investor.lilly.com/news-releases/news-release-details/amylin-and-lilly-announce-fda-approval-byettatm-exenatide).
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