# Hyperglycemia and Adverse Pregnancy Outcome Follow-up Study (HAPO FUS): Maternal Gestational Diabetes Mellitus and Childhood Glucose Metabolism 

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## OBJECTIVE

Whether hyperglycemia in utero less than overt diabetes is associated with altered childhood glucose metabolism is unknown. We examined associations of gestational diabetes mellitus (GDM) not confounded by treatment with childhood glycemia in the Hyperglycemia and Adverse Pregnancy Outcome (HAPO) cohort.

## RESEARCH DESIGN AND METHODS

HAPO Follow-up Study (FUS) included 4,160 children ages $10-14$ years who completed all or part of an oral glucose tolerance test (OGTT) and whose mothers had a 75-g OGTT at ~28 weeks of gestation with blinded glucose values. The primary predictor was GDM by World Health Organization criteria. Child outcomes were impaired fasting glucose (IFG), impaired glucose tolerance (IGT), and type 2 diabetes. Additional measures included insulin sensitivity and secretion and oral disposition index.

## RESULTS

For mothers with GDM, $10.6 \%$ of children had IGT compared with $5.0 \%$ of children of mothers without GDM; IFG frequencies were $9.2 \%$ and $7.4 \%$, respectively. Type 2 diabetes cases were too few for analysis. Odds ratios ( $95 \% \mathrm{Cl}$ ) adjusted for family history of diabetes, maternal BMI, and child BMI z score were 1.09 (0.78-1.52) for IFG and 1.96 (1.41-2.73) for IGT. GDM was positively associated with child's $30-\mathrm{min}$, 1-h, and 2-h but not fasting glucose and inversely associated with insulin sensitivity and oral disposition index (adjusted mean difference -76.3 [95\% CI -130.3 to -22.4$]$ and $-0.12[-0.17$ to -0.064$]$ ), respectively, but not insulinogenic index.

## CONCLUSIONS

Offspring exposed to untreated GDM in utero are insulin resistant with limited $\beta$-cell compensation compared with offspring of mothers without GDM. GDM is significantly and independently associated with childhood IGT.

The incidence of type 2 diabetes among youth is rising, due, in part, to increasing prevalence of childhood obesity (1-3). Worldwide, it is estimated that type 2 diabetes in children will continue to increase, posing a significant public health and financial burden (4). In addition to childhood obesity, intrauterine exposure to maternal preexisting diabetes or gestational diabetes mellitus (GDM) is associated with a higher

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[^0]risk for offspring abnormal glucose metabolism (5). Early evidence for these associations came from two long-term studies: a longitudinal study of Pima Indians (6) and cohort study of a mixed ethnic population in Chicago (7). However, these studies focused on women with a high prevalence of diabetes, making it unclear whether these results could be extrapolated to other populations. More recently, maternal treatment during pregnancy has confounded studies examining the impact of the intrauterine milieu on offspring risk of hyperglycemia (8). Studies have also not adequately addressed whether hyperglycemia in utero less than overt diabetes is associated with altered glucose metabolism in childhood.

The Hyperglycemia and Adverse Pregnancy Outcome (HAPO) Follow-up Study (FUS) offered a unique opportunity to examine associations of maternal glycemia during pregnancy, less than overt diabetes and not confounded by maternal treatment, with childhood glucose metabolism. The HAPO study, an observational epidemiological study that recruited a large, multinational, racially and ethnically diverse cohort, demonstrated that glucose levels below those diagnostic of diabetes were associated with adverse pregnancy outcomes (9). Based on these results and others, new diagnostic criteria for GDM were proposed by the International Association of the Diabetes and Pregnancy Study Groups (IADPSG) (10) and adopted by the World Health Organization (WHO) and others (11). These IADPSG/WHO criteria, based on newborn outcomes, reduced both the glucose levels and required number of abnormal values (from two to one) to diagnose GDM compared with CarpenterCoustan criteria (12). This report from the HAPO FUS examines whether in utero exposure to untreated GDM, defined post hoc by IADPSG/WHO criteria, is associated with prespecified glucose outcomes (impaired fasting glucose [IFG], impaired glucose tolerance [IGT], or type 2 diabetes) in children ages $10-$ 14 years.

## RESEARCH DESIGN AND METHODS

HAPO was a population-based study in which women underwent a 75 -g oral glucose tolerance test (OGTT) at $\sim 28$ weeks of gestation (9). Fasting plasma glucose (FPG) and 1-h and 2-h plasma glucose (PG) were measured at a central laboratory (9). OGTT results remained blinded to caregivers and participants unless FPG $>5.8 \mathrm{mmol} / \mathrm{L}$ and/or 2-h PG $>11.1 \mathrm{mmol} / \mathrm{L}$, either measure was $<2.5 \mathrm{mmol} / \mathrm{L}$, or random PG at 34-37 weeks of gestation was $>8.9 \mathrm{mmol} / \mathrm{L}$ (9). Using these criteria, 427 (1.8\%) participants were unblinded based on FPG and/or 2-h PG. Blinded participants were untreated. Height, weight, and blood pressure were measured using standardized procedures. Demographic and lifestyle characteristics, including age, self-reported race/ ethnicity, and any smoking or alcohol use during pregnancy, were collected via questionnaire and parity by medical record abstraction.

## Participants

HAPO FUS participants were recruited during 2013-2016 from 10 of 15 HAPO field centers based on feasibility to recruit participants. Eligibility criteria for HAPO FUS included caregivers and participants being blinded to HAPO OGTT results, gestational age at delivery $\geq 37$ weeks, and no major neonatal malformations or fetal/neonatal death. This yielded 15,812 eligible mother-child pairs. The recruitment target was 7,000 pairs, based on the primary childhood outcome of overweight/obesity (13). Multiple attempts were made to contact all eligible participants through institutional review board (IRB)-approved means. Of the 15,812 eligible motherchild pairs, 6,490 could not be contacted and 4,488 declined participation (Supplementary Fig. 1). A total of 4,834 children completed all or part of the HAPO FUS visit. OGTT completion was not required for participation. One child was excluded due to inadequate fasting and a second for lack of cooperation. Of the remaining 4,832 children, data were analyzed from

4,160 who had an FPG and at least one other timed OGTT measurement or who reported diabetes on treatment and were not excluded as having type 1 diabetes by antibody testing (see below).

The protocol was approved by each center's IRB. All mothers gave written informed consent for their child, and children assented where required by the local IRB. There was an external Observational Study Monitoring Board.

## Study Visit

Height was measured twice without shoes to the nearest 0.5 cm with a stadiometer and again if results differed by $>1.0 \mathrm{~cm}$. Weight was measured twice to the nearest 0.1 kg and again if results differed by $>0.5 \mathrm{~kg}$. Participants underwent a 2-h OGTT with a glucose load of $1.75 \mathrm{~g} / \mathrm{kg}$ body wt (maximum 75 g ) after an 8-h overnight fast with samples drawn for glucose and C -peptide at fasting and $30 \mathrm{~min}, 1 \mathrm{~h}$, and 2 h . If the child had self-reported diabetes on drug treatment, only a nonfasting blood sample was collected. All samples were processed at the field center laboratory and stored at $-80^{\circ} \mathrm{C}$ until shipment to the Central Laboratory.

Skinfolds (triceps, subscapular, and suprailiac) were measured twice with calibrated calipers (Harpenden, London, U.K.) to the nearest 0.1 mm and again if results differed by $>1.0 \mathrm{~mm}$. Percent fat was measured by air displacement plethysmography (BOD POD; COSMED, Rome, Italy). Tanner staging was performed by trained individuals using breast/areolar development for girls and testicular volume (Prader orchidometer) for boys. Child's age, first-degree family history of diabetes, and menstrual history for girls were collected from the mother via questionnaire.

## Laboratory Measurements

Glucose was measured by hexokinase in the Clinical Chemistry Laboratory of Northwestern Memorial Hospital on a Beckman Coulter SYNCHRON LX analyzer. Blinded duplicate samples were assayed several weeks apart and coefficients of

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variation (CVs) calculated within pairs for a random $10 \%$ subset; mean CV was $1.5 \%$ for FPG and 1-h and 2-h PG and $1.3 \%$ for $30-\mathrm{min}$ PG. C-peptide was measured in the Comprehensive Metabolic Core at Northwestern using the electrochemiluminescence immunoassay method on a Roche cobas e 411 immunoassay analyzer (14). Mean CVs were 2.8\% for fasting, $2.9 \%$ for $30-\mathrm{min}, 3.0 \%$ for 1-h, and $3.2 \%$ for 2 -h C-peptide. To diagnose type 1 diabetes, serum antiGAD65, -insulin, -ZnT8, and -IA-2 antibodies were measured at the Barbara Davis Center for Diabetes (Aurora, CO) for children reported to have diabetes on treatment ( $n=9$ ) and for children with OGTT values indicative of incident diabetes ( $n=5$ ). Of these 14 children, 4 had positive antibody results and were excluded from analyses.

## Outcomes and Predictors

Dichotomous childhood outcomes were as follows: IGT (2-h glucose 7.8-11.0 $\mathrm{mmol} / \mathrm{L}$ ); IFG by American Diabetes Association (ADA) and International Society for Pediatric and Adolescent Diabetes criteria (IFG-ADA, FPG 5.6-6.9 $\mathrm{mmol} / \mathrm{L}$ ); IFG by WHO criteria (IFGWHO, FPG 6.1-6.9 mmol $/ \mathrm{L}$ ); and type 2 diabetes ( $F P G \geq 7.0 \mathrm{mmol} / \mathrm{L}$ and/or 2-h $P G \geq 11.1 \mathrm{mmol} / \mathrm{L})$. Continuous outcomes included FPG and 30-min, 1-h, and 2-h PG from the OGTT at individual time points. The sum of individual glucose $z$ scores, an integrated measure that gives equal weight to each of the four glucose values during the OGTT, was also examined. This involved calculating $z$ scores at each OGTT time point by subtracting the mean glucose level from all observed values at that time point, dividing by the SD of the glucose values at that time point, and summing the four individual z scores.

Matsuda, insulinogenic, and disposition indices were examined as continuous outcomes. The Matsuda index estimates insulin sensitivity using glucose and insulin levels from an OGTT (15). A modified Matsuda index was calculated using OGTT glucose and C-peptide levels (15). The insulinogenic index was calculated using C-peptide levels and defined as the $\Delta \mathrm{C}$-peptide ( $0-30 \mathrm{~min}, \mathrm{nmol} / \mathrm{L}$ )/Dglucose ( $0-30 \mathrm{~min}$, $\mathrm{mmol} / \mathrm{L}$ ) (16). The disposition index was calculated as the product of
the Matsuda index and insulinogenic index and then log transformed (17).

The primary predictor for child glucose outcomes was mother's GDM status during the HAPO pregnancy using IADPSG/WHO criteria where one or more glucose values from a 75-g OGTT equaled or exceeded the following thresholds: FPG $5.1 \mathrm{mmol} / \mathrm{L}, 1-\mathrm{h}$ PG $10.0 \mathrm{mmol} / \mathrm{L}$, and 2-h PG $8.5 \mathrm{mmol} / \mathrm{L}$.

Exploratory analyses were conducted by removing women from the data set who met Carpenter-Coustan GDM criteria, defined in this study as two abnormal glucose values from a 2-h 75-g OGTT that equaled or exceeded the following thresholds: FPG $5.3 \mathrm{mmol} / \mathrm{L}, 1-\mathrm{h}$ PG $10.0 \mathrm{mmol} / \mathrm{L}$, and $2-\mathrm{h}$ PG $8.6 \mathrm{mmol} / \mathrm{L}$. A predictor for mothers who only met IADPSG/WHO GDM criteria and not Car-penter-Coustan criteria, versus mothers without GDM, was then evaluated.

## Statistical Analyses

Data were summarized using frequencies and counts for categorical variables and means and SDs for continuous variables. Histograms and box plots were examined to assess distributions and identify potential outliers. Multiple logistic regression was used to estimate odds ratios (ORs) and $95 \%$ Cls for dichotomous outcomes, and modified least squares regression with Huber-White robust SE was used to estimate risk differences with $95 \%$ Cls (18). Multiple linear regression was used for continuous outcomes to estimate adjusted mean differences with $95 \%$ Cls. Two-sided $P<0.05$ was used for evaluating statistical significance. All statistical analyses were conducted in $R$ (3.4.1) (19).

Multiple models were considered for all outcomes, with variables identified according to study design, known potential confounders, and adjustments used in HAPO analyses (9). Covariate adjustments were as follows: model 1: field center (proxy for race/ancestry since most centers were predominantly one race/ancestry group), child age, sex, and pubertal status (Tanner stage 1, 2/3, or $4 / 5$ ) with sex by Tanner stage interaction term, maternal variables at pregnancy OGTT (age, height, mean arterial pressure, parity [0 or 1+], smoking [yes/no], drinking alcohol [yes/no], and gestational age), and child's family history of diabetes in first-degree relatives; model 2: model $1+$ maternal BMI at
pregnancy OGTT; model 3: model $1+$ child's BMI $z$ score; and model 4: model 1 + maternal BMI at pregnancy OGTT + child's BMI z score. Child BMI z scores were calculated according to skewness-median-coefficient of variation curves used by the International Obesity Task Force (20). Adjustments for other child adiposity variables were also explored, substituting child's percent body fat in models 5 and 6 and child's sum of skinfolds in models 7 and 8 for child's BMI z score. Although the study was not powered to evaluate Tanner stage-specific associations, an interaction term between GDM status and Tanner stage was evaluated to explore potential effect modification by Tanner stage. Exploratory analyses were also performed within Tanner stage $1,2 / 3$, and $4 / 5$ groups, using model covariates just listed but removing main effects for Tanner stage and sex $\times$ Tanner stage interactions. Multiple imputation was used to account for missing data using a "missing at random" assumption after confirming findings varied little under "missing not at random" (13). Logistic regression model fit was measured using C-statistics and confirmed by HosmerLemeshow goodness-of-fit tests (21). Colinearity of model predictors was evaluated using pairwise correlations.

## RESULTS

## Participants

Characteristics of participating children (mean age 11.4 years) and their mothers during the HAPO study, overall and by GDM status, are shown in Table 1. At the HAPO OGTT, mothers with GDM were on average older with higher weight, BMI, and mean arterial pressure. Among mothers of children who did and did not participate, mean age and frequency of GDM were 30.0 years and $14.9 \%$ and 29.1 years and $16.9 \%$, respectively (weighted summaries) (Supplementary Table 1). Mean maternal BMI, FPG, and 1-h and 2-h PG during the HAPO OGTT and race/ethnicity were similar.

Offspring of mothers with and without GDM had similar age and height. Offspring of mothers with GDM were heavier and had a higher sum of skinfolds, more frequent family history of diabetes in first-degree relatives, as well as lower insulin sensitivity (Matsuda index), insulin secretion (insulinogenic index), and $\beta$-cell compensation for insulin resistance

Table 1-Characteristics of mothers during HAPO pregnancy OGTT and their children at follow-up according to mother's GDM status

| During HAPO pregnancy | Overall | GDM* | No GDM |
| :---: | :---: | :---: | :---: |
|  | $n=4,160$ | $n=589$ | $n=3,571$ |
| Characteristics (mothers) | Mean (SD) | Mean (SD) | Mean (SD) |
| Age at OGTT (years) | 29.9 (5.7) | 31.8 (5.3) | 29.6 (5.7) |
| Gestational age at OGTT (weeks) | 27.6 (1.7) | 27.9 (1.7) | 27.6 (1.7) |
| Height (cm) | 161.7 (6.8) | 160.9 (7.0) | 161.8 (6.8) |
| Weight (kg) | 71.8 (14.1) | 77.2 (15.9) | 70.9 (13.6) |
| BMI ( $\mathrm{kg} / \mathrm{m}^{2}$ ) | 27.4 (4.9) | 29.7 (5.4) | 27.0 (4.7) |
| Mean arterial pressure ( mmHg ) | 80.2 (7.9) | 83.3 (7.7) | 79.7 (7.8) |
| FPG ( $\mathrm{mmol} / \mathrm{L}$ ) | 4.5 (0.4) | 4.9 (0.4) | 4.4 (0.3) |
| 1-h PG (mmol/L) | 7.4 (1.7) | 9.6 (1.6) | 7.0 (1.4) |
| 2-h PG (mmol/L) | 6.1 (1.3) | 7.7 (1.5) | 5.9 (1.1) |
|  | $n$ (\%) | $n$ (\%) | $n$ (\%) |
| Race/ethnicity |  |  |  |
| White, non-Hispanic | 1,778 (42.7) | 215 (36.5) | 1,563 (43.8) |
| Hispanic | 465 (11.2) | 104 (17.7) | 361 (10.1)) |
| Black, non-Hispanic | 719 (17.3) | 71 (12.1) | 648 (18.1) |
| Asian | 1,125 (27.0) | 184 (31.2) | 941 (26.4) |
| Other | 73 (1.8) | 15 (2.5) | 58 (1.6) |
| Any prenatal smoking | 177 (4.3) | 33 (5.6) | 144 (4.0) |
| Any prenatal alcohol use | 283 (6.8) | 39 (6.6) | 244 (6.8) |
| Parity (any prior delivery $\geq 20$ weeks) | 2,146 (51.6) | 334 (56.7) | 1,812 (50.7) |
| Characteristics (children), at follow-up | Mean (SD) | Mean (SD) | Mean (SD) |
| Age (years) | 11.4 (1.2) | 11.4 (1.3) | 11.3 (1.2) |
| Height (cm) | 148.5 (10.4) | 149.4 (9.9) | 148.4 (10.4) |
| Weight (kg) | 43.2 (13.6) | 45.9 (14.5) | 42.7 (13.4) |
| BMI z score | 0.47 (1.25) | 0.73 (1.31) | 0.43 (1.24) |
| Sum of skinfolds (mm) | 38.8 (21.3) | 43.6 (23.8) | 38.1 (20.8) |
|  | $n$ (\%) | $n$ (\%) | $n$ (\%) |
| Sex (female) | 2,036 (48.9) | 272 (46.2) | 1,764 (49.4) |
| Tanner stage (girls) |  |  |  |
| 1 | 340 (19.0) | 39 (16.0) | 301 (19.5) |
| 2/3 | 761 (42.6) | 108 (44.4) | 653 (42.3) |
| 4/5 | 686 (38.4) | 96 (39.5) | 590 (38.2) |
| Tanner stage (boys) |  |  |  |
| 1 | 523 (35.8) | 66 (31.4) | 457 (36.6) |
| 2/3 | 673 (46.1) | 100 (47.6) | 573 (45.9) |
| 4/5 | 263 (18.0) | 44 (21.0) | 219 (17.5) |
| Family history of diabetes | 1,899 (45.7) | 358 (60.8) | 1,541 (43.2) |
| Outcomes (children) | Mean (SD) | Mean (SD) | Mean (SD) |
| Fasting glucose ( $\mathrm{mmol} / \mathrm{L}$ ) | 5.0 (0.4) | 5.1 (0.4) | 5.0 (0.4) |
| $30-\mathrm{min}$ glucose ( $\mathrm{mmol} / \mathrm{L}$ ) | 7.8 (1.3) | 8.2 (1.4) | 7.7 (1.3) |
| 1-h glucose ( $\mathrm{mmol} / \mathrm{L}$ ) | 6.8 (1.7) | 7.3 (1.8) | 6.8 (1.7) |
| 2-h glucose ( $\mathrm{mmol} / \mathrm{L}$ ) | 6.0 (1.2) | 6.2 (1.2) | 5.9 (1.1) |
| Sum of glucose z scores | -0.01 (2.8) | 0.9 (2.9) | -0.2 (2.8) |
| Matsuda index | 1,810.9 (747.3) | 1,620.0 (681.8) | 1,838.4 (753.3) |
| Insulinogenic index | 0.82 (0.69) | 0.78 (0.64) | 0.82 (0.70) |
| Disposition index | 7.02 (0.58) | 6.88 (0.59) | 7.05 (0.57) |
|  | $n$ (\%) | $n$ (\%) | $n$ (\%) |
| IFG-ADA | 316 (7.6) | 54 (9.2) | 262 (7.4) |
| IFG-WHO | 23 (0.6) | 3 (0.5) | 20 (0.6) |
| IGT | 238 (5.8) | 61 (10.6) | 177 (5.0) |
| Type 2 diabetes | 10 (0.2) | 2 (0.3) | 8 (0.2) |

Type 2 diabetes ( $F P G \geq 7.0$ and/or 2-h PG $\geq 11.0 \mathrm{mmol} / \mathrm{L}$ ). Disposition index is reported on a log scale. *GDM is defined by IADPSG/WHO criteria (one or moreglucose values from a $75-\mathrm{g}$ OGTT equals or exceeds FPG $5.1 \mathrm{mmol} / \mathrm{L}$, 1-h $10.0 \mathrm{mmol} / \mathrm{L}$, 2-h $8.5 \mathrm{mmol} / \mathrm{L})$.
(disposition index), compared with offspring of mothers without GDM.

IFG-ADA was present in 54 (9.2\%) of offspring of mothers with GDM compared with 262 (7.4\%) offspring of mothers without GDM, and 61 (10.6\%) offspring of mothers with GDM had IGT compared with 177 (5.0\%) offspring of mothers without GDM (Table 1). IFGWHO was present in 3 ( $0.5 \%$ ) offspring of mothers with GDM and 20 ( $0.6 \%$ ) offspring of mothers without GDM. Type 2 diabetes was present in 2 ( $0.3 \%$ ) offspring of mothers with GDM and 8 ( $0.2 \%$ ) offspring of mothers without GDM.

## Model Diagnostics

Hosmer-Lemeshow $P$ values for logistic regression models ranged 0.59-0.97 for all outcomes, indicating reasonable model fit. C-statistics for logistic regression models ranged $0.70-0.78$ and changed very little for each outcome for models 1-4. $R^{2}$ values ranged $0.071-$ 0.23 for continuous outcomes and changed very little across models 1-4. Colinearity was not a concern with pairwise correlations ranging from 0 to 0.20 for model covariates. Visual inspection of residual plots confirmed linear modeling assumptions, and DFBETA statistics indicated no observations of undue influence.

## Association of Maternal GDM Status With Childhood Glucose Levels

Offspring of mothers with GDM had higher 30-min, 1-h, and 2-h PG during their OGTT and sum of glucose $z$ scores compared with offspring of mothers without GDM. GDM was positively associated with child $30-\mathrm{min}, 1-\mathrm{h}$, and 2-h PG and sum of glucose $z$ scores but not child FPG (Table 2). For the above associations, adjusting for maternal BMI at the time of the pregnancy OGTT (model 2), child's BMI $z$ score (model 3), or both maternal BMI and child's BMI z score (model 4) had little effect. Results were similar after adjusting for either child percent body fat or sum of skinfolds at follow-up instead of child's BMI z score (Supplementary Table 2), demonstrating independence of the associations from maternal and child BMI and adiposity.

## Association of Maternal GDM Status With Childhood Dichotomous Glucose Outcomes

GDM was associated with offspring IGT. After model 1 adjustments, the OR (CI)

Table 2-Association of maternal GDM status during pregnancy with child glucose outcomes

|  | Model 1 | Model 2 | Model 3 | Model 4 |
| :---: | :---: | :---: | :---: | :---: |
| ontinuous | $\beta$ (95\% CI) | $\beta$ (95\% CI) | $\beta$ (95\% CI) | $\beta$ (95\% CI) |
| metabolic outcomes | $P$ value, $R^{2}$ | $P$ value, $R^{2}$ | $P$ value, $R^{2}$ | $P$ value, $R^{2}$ |
| Fasting glucose (mmol/L) | $\begin{gathered} 0.026(-0.0056 \text { to } 0.056) \\ P=0.11,0.23 \end{gathered}$ | $\begin{gathered} 0.024(-0.0072 \text { to } 0.055) \\ P=0.13,0.23 \end{gathered}$ | $\begin{gathered} 0.021(-0.0094 \text { to } 0.052) \\ P=0.18,0.23 \end{gathered}$ | $\begin{gathered} 0.022(-0.0083 \text { to } 0.053) \\ P=0.16,0.23 \end{gathered}$ |
| $30-\mathrm{min}$ glucose ( $\mathrm{mmol} / \mathrm{L}$ ) | $\begin{aligned} & 0.27 \text { (0.16-0.39) } \\ & P<0.0001,0.12 \end{aligned}$ | $\begin{aligned} & 0.28(0.17-0.40) \\ & P<0.0001,0.12 \end{aligned}$ | $\begin{gathered} 0.27(0.15-0.38) \\ P<0.0001,0.12 \end{gathered}$ | $\begin{aligned} & 0.28(0.16-0.40) \\ & P<0.0001,0.12 \end{aligned}$ |
| 1-h glucose (mmol/L) | $\begin{gathered} 0.34(0.19-0.49) \\ P<0.0001,0.071 \end{gathered}$ | $\begin{gathered} 0.35(0.20-0.50) \\ P<0.0001,0.072 \end{gathered}$ | $\begin{gathered} 0.33(0.18-0.48) \\ P<0.0001,0.072 \end{gathered}$ | $\begin{gathered} 0.35(0.20-0.50) \\ P<0.0001,0.072 \end{gathered}$ |
| 2-h glucose (mmol/L) | $\begin{gathered} 0.20(0.10-0.30) \\ P<0.0001,0.079 \end{gathered}$ | $\begin{gathered} 0.20(0.10-0.30) \\ P<0.0001,0.079 \end{gathered}$ | $\begin{gathered} 0.18(0.08-0.28) \\ P=0.0004,0.095 \end{gathered}$ | $\begin{gathered} 0.20(0.10-0.30) \\ P=0.0001,0.096 \end{gathered}$ |
| Sum of glucose z scores | $\begin{gathered} 0.65(0.41-0.88) \\ P<0.0001,0.17 \end{gathered}$ | $\begin{gathered} 0.66(0.42-0.90) \\ P<0.0001,0.17 \end{gathered}$ | $\begin{gathered} 0.61(0.37-0.84) \\ P<0.0001,0.17 \end{gathered}$ | $\begin{aligned} & 0.65(0.41-0.88) \\ & P<0.0001,0.17 \end{aligned}$ |
| Matsuda index | $\begin{gathered} -114.5(-174.5 \text { to }-54.5) \\ P=0.0002,0.22 \end{gathered}$ | $\begin{gathered} -91.6(-151.6 \text { to }-31.6) \\ P=0.0031,0.23 \end{gathered}$ | $\begin{gathered} -70.9(-124.4 \text { to }-17.5) \\ P=0.010,0.38 \end{gathered}$ | $\begin{gathered} -76.3(-130.3 \text { to }-22.4) \\ P=0.0063,0.38 \end{gathered}$ |
| Insulinogenic index | $\begin{gathered} -0.034(-0.10 \text { to } 0.029) \\ P=0.28,0.052 \end{gathered}$ | $\begin{gathered} -0.053(-0.12 \text { to } 0.010) \\ P=0.10,0.058 \end{gathered}$ | $\begin{gathered} -0.052(-0.11 \text { to } 0.010) \\ P=0.10,0.080 \end{gathered}$ | $\begin{gathered} -0.060(-0.120 \text { to } 0.0030) \\ P=0.061,0.081 \end{gathered}$ |
| Disposition index | $\begin{gathered} -0.12(-0.17 \text { to }-0.06) \\ P<0.0001,0.071 \end{gathered}$ | $\begin{gathered} -0.12(-0.18 \text { to }-0.065) \\ P<0.0001,0.071 \end{gathered}$ | $\begin{gathered} -0.11(-0.17 \text { to }-0.057) \\ P<0.0001,0.073 \end{gathered}$ | $\begin{gathered} -0.12(-0.17 \text { to }-0.064) \\ P<0.0001,0.074 \end{gathered}$ |

Model 1 was adjusted for field center + child age, sex, and pubertal status (Tanner stage 1, 2/3, 4/5, and sex $\times$ Tanner stage interaction) + maternal variables at pregnancy OGTT (age, height, mean arterial pressure, parity [0 or 1+], smoking [yes/no], drinking [yes/no], and gestational age), and child's family history of diabetes in first-degree relatives. Model 2 was model $1+$ maternal BMI at pregnancy OGTT. Model 3 was model 1 + child's BMI z score. Model 4 was model 1 + maternal BMI at pregnancy OGTT + child's BMI z score.
and risk difference (CI) for IGT in offspring were 1.91 (1.38-2.64) and 0.046 (0.020-0.071), respectively. Adjusting for maternal BMI and/or child's BMI z score (models 2-4) did not attenuate associations (ORs 1.87-1.96 and risk
differences 0.044-0.047) (Fig. 1 and Table 3). In contrast, GDM was not associated with a higher risk for IFG-ADA (ORs 1.091.11 and risk differences $0.0057-0.0066$ ). The ORs and risk differences for IGT and IFG-ADA were similar after adjusting for
child percent body fat or sum of skinfolds instead of child BMI z score (models 5-8) (Supplementary Table 3). Due to small sample sizes for IFG-WHO and type 2 diabetes, associations of GDM with these outcomes were not analyzed.

A
IGT

$$
\begin{gathered}
1.91(1.38-2.64) \\
\mathrm{p}<0.001, \mathrm{C}=0.70 \\
1.96(1.41-2.72) \\
\mathrm{p}<0.001, \mathrm{C}=0.70 \\
1.87(1.35-2.59) \\
\mathrm{p}<0.001, \mathrm{C}=0.71 \\
1.96(1.41-2.73) \\
\mathrm{p}<0.001, \mathrm{C}=0.71
\end{gathered}
$$

Odds Ratio (95\% CI) p-value, C-statistic

## C

IFG-ADA


B
IGT

$$
\begin{gathered}
0.046(0.020-0.071) \\
\mathrm{p}<0.001 \\
0.047(0.022-0.073) \\
\mathrm{p}<0.001 \\
0.044(0.019-0.070) \\
\mathrm{p}<0.001 \\
0.047(0.021-0.072) \\
\mathrm{p}<0.001
\end{gathered}
$$

D
IFG-ADA


Figure 1-ORs ( $95 \% \mathrm{Cls}$ ), $P$ values, and C-statistics from logistic regression analyses and risk differences ( $95 \% \mathrm{Cls}$ ) and $P$ values for the association of maternal GDM during pregnancy with outcomes in HAPO FUS children. A and B: IGT. C and D: IFG-ADA. Results are presented for each outcome for model 1 (purple), model 2 (blue), model 3 (green), and model 4 (orange) covariates.

## Association of Maternal GDM Status With Childhood Insulin Sensitivity and Secretion

GDM was inversely associated with child insulin sensitivity after adjusting for covariates in model 1 . This association was attenuated by adjusting for maternal BMI (model 2) and/or child BMI z score (models 3 and 4) but remained significant, demonstrating independence from these two measures. Adjusting for either child percent body fat or sum of skinfolds at follow-up also attenuated the associations, but they remained significant (Supplementary Table 2). GDM was not associated with the insulinogenic index except for a borderline significant inverse association after adjusting for both maternal BMI and child $\mathrm{BMI} z$ score ( $P=0.06$ ). Results were comparable after adjusting for child percent body fat or sum of skinfolds (Supplementary Table 2). GDM was significantly inversely associated with the disposition index in all models, including after adjusting for child percent body fat and sum of skinfolds (Table 2 and Supplementary Table 2).

Puberty affects insulin sensitivity and glucose metabolism (22). Thus, additional exploratory analyses were conducted separately for Tanner stages $1,2 / 3$, and $4 / 5$. These analyses suggested that some of the associations described above may be strongest after the onset of puberty (model 4 results in Supplementary Table 4). However, for all childhood outcomes, statistical interaction terms between GDM and Tanner stage were not significant.

## Association of Maternal GDM Status According to IADPSG/WHO but Not Carpenter-Coustan Criteria With Childhood Glucose Outcomes

The diagnosis of GDM using CarpenterCoustan criteria requires two abnormal glucose values: $\mathrm{FPG}>5.3 \mathrm{mmol} / \mathrm{L}, 1-\mathrm{h}$ PG $>10.0 \mathrm{mmol} / \mathrm{L}, 2-\mathrm{h} \mathrm{PG}>8.6 \mathrm{mmol} / \mathrm{L}$, or $3-\mathrm{h}$ PG $>7.8 \mathrm{mmol} / \mathrm{L}$. The current study was not powered to assess differences in childhood glucose outcomes for women with Carpenter-Coustan GDM, GDM by IADPSG/WHO criteria only, and no GDM. However, exploratory analyses were conducted by examining childhood dichotomous outcome frequencies and childhood continuous outcome means across the three groups. In addition, fully adjusted model 4 logistic and linear
regression results were examined after removing women meeting CarpenterCoustan GDM criteria to evaluate associations with childhood outcomes for mothers with GDM by IADPSG/WHO criteria only compared with those without GDM. In general, childhood dichotomous outcome frequencies and continuous outcome means for mothers with GDM by IADPSG/WHO criteria only were midway between those for mothers with GDM by CarpenterCoustan criteria and mothers without GDM, except for FPG and insulinogenic index (Supplementary Table 5). As observed for analyses including mothers with Carpenter-Coustan GDM, GDM according to IADPSG/WHO criteria only was not associated with IFG or FPG in the fully adjusted model 4. However, GDM according to IADPSG/WHO criteria only was associated with IGT, 30-min, 1-h, and 2-h PG, glucose sum of $z$ scores, and disposition index. ORs and adjusted mean differences were attenuated compared with analyses including mothers with Carpenter-Coustan GDM, but results still indicate statistically significant higher risks of adverse metabolic outcomes for children of mothers with GDM according to IADPSG/WHO criteria alone.

## CONCLUSIONS

This study demonstrated that offspring of mothers with untreated GDM by IADPSG/WHO criteria are at high risk for IGT 10-14 years postpartum. In the U.S., over a time frame similar to HAPO, GDM prevalence using CarpenterCoustan criteria was $7.4 \%$ (23). In contrast, IADPSG/WHO criteria, which require only one abnormal glucose value during a 2-h OGTT, identify a larger group of mothers, e.g., $17.8 \%$ of women in the HAPO cohort (24). Thus, offspring of women meeting IADPSG/WHO criteria for GDM represent a sizable group at risk for IGT during childhood, which is important from a public health perspective, given the risk for progression to type 2 diabetes in children with IGT (25). The frequency of offspring type 2 diabetes in the current study was low, which precluded analyses of association of offspring type 2 diabetes with GDM.
Previous studies have examined the association between GDM or preexisting diabetes in utero and later risk
of altered offspring glucose metabolism ( $5,26-28$ ). Associations of preexisting maternal type 2 diabetes or GDM with offspring glycemic outcomes have been inconsistent, relating, in part, to the size of the studies and failure to adjust for maternal BMI (29). A recent metaanalysis of studies performed prior to the development of IADPSG/WHO criteria examined associations of GDM with glucose outcomes in children from prepuberty through young adulthood (30). Among 5,850 children, GDM was not associated with child FPG, whereas association of GDM with child 2-h glucose was demonstrated in a group of 890 children. The current study also demonstrated no association of GDM with child FPG but did demonstrate that after adjustment for both maternal and child BMI, GDM was associated with not only child 2-h glucose but also with $30-\mathrm{min}$ and 1-h PG.

We have demonstrated associations of GDM according to IADPSG/WHO criteria with offspring obesity and adiposity in the HAPO FUS cohort (13). However, despite the known impact of childhood obesity on glucose metabolism (25), a singular feature of the current report is that association of GDM with childhood IGT was not attenuated by adjusting for the child's BMI z score or other measures of adiposity or maternal BMI during pregnancy. Compared with mothers without GDM, mothers with GDM more frequently have a family history of type 2 diabetes and are at higher risk for developing type 2 diabetes $(31,32)$. Offspring of mothers with GDM in this study more frequently had a family history of diabetes in first-degree relatives; however, the association of maternal GDM with child IGT was independent of the child's family history of diabetes. Together, these data demonstrate that GDM is associated with an independent risk for offspring IGT, suggesting that fetal programming may contribute to the observed associations, although additional factors, e.g., shared environmental factors and shared genetics not captured by family history of diabetes, may also contribute. Importantly, this study also demonstrated that the risk of IGT in offspring of mothers with GDM occurred with lesser degrees of hyperglycemia during pregnancy compared with GDM diagnostic thresholds used in earlier studies. Consistent with that,
exploratory analyses that removed women with Carpenter-Coustan GDM showed that GDM according to IADPSG/ WHO criteria was significantly associated with child IGT and other measures of glucose metabolism compared with mothers without GDM.

Previous studies have demonstrated the association of maternal diabetes with offspring insulin resistance during childhood $(25,33,34)$. In the current study, insulin sensitivity was also lower in offspring of mothers with GDM compared with those without GDM. Adjusting for child BMI or adiposity attenuated the association of GDM with insulin sensitivity, but the association remained significant, suggesting that it was, in part, independent of child BMI and adiposity. Insulin resistance is an early abnormality in children with abnormal glucose metabolism and a major driving force behind dysglycemia in adolescents (35). Progression to type 2 diabetes in children occurs, in part, secondary to a rapid decline in $\beta$-cell function (36-38). We found that offspring of mothers with GDM had, on average, a lower insulinogenic index, but the association was borderline significant only after adjusting for both maternal BMI during pregnancy and child's BMI z score. However, the oral disposition index in offspring of mothers with GDM was consistently lower than in offspring of mothers without GDM. This is consistent with a $\beta$-cell defect in these children, manifest as limited $\beta$-cell compensation for the underlying insulin resistance. Greater insulin resistance together with lower disposition index is seen in individuals who tend to progress to type 2 diabetes $(39,40)$ and likely contributes to the higher risk for IGT in offspring of mothers with GDM.

Of particular note in this study was the positive association of GDM with the child's $30 \mathrm{~min}, 1-\mathrm{h}$, and 2-h PG but not FPG. Similarly, GDM was associated with a higher risk of IGT but not IFG-ADA. IFG and IGT are considered to be distinct metabolic conditions with differing pathophysiologies $(41,42)$. Isolated IFG is thought to reflect high hepatic insulin resistance with relatively normal insulin sensitivity in skeletal muscle, whereas those with isolated IGT typically exhibit insulin resistance in skeletal muscle (41). Thus, GDM may preferentially impact offspring skeletal muscle as opposed to liver.

Strengths of the study include that HAPO was a blinded observational study in which both caregivers and mothers were unaware of maternal glucose levels; thus, child outcomes were not confounded by treatment of maternal hyperglycemia. Also, HAPO included participants from multiple races and ethnicities from field centers around the world, making the results broadly applicable.

This study had some limitations. First, the proportion of participants who met IADPSG/WHO criteria and participated in the HAPO FUS (weighted estimate 14.9\%) is lower than in all eligible participants (16.2\%). Second, we were unable to diagnose IGT in all participants due to some missing 2-h PG measurements. Participants with a normal FPG but missing 2-h PG value ( $n=54$ ) were defined as having normal glucose metabolism; thus, the number of individuals with IGT may have been underestimated. Third, $1.8 \%$ of HAPO participants with OGTT values higher than predefined thresholds were unblinded and excluded from HAPO primary analyses and this study (24). This subgroup would likely have included children at highest risk for IGT. Fourth, the HAPO FUS was not powered to formally examine Tanner stage-specific associations. Fifth, paternal BMI data were not available. Finally, postnatal factors that might influence child glucose outcomes were not available for analyses.

In summary, against a background of limited data and conflicting results, the present large, racially and ethnically diverse study demonstrates that maternal hyperglycemia during pregnancy is associated with a higher risk of IGT in childhood. In addition, this association is evident even when maternal GDM is defined by the less stringent IADPSG/ WHO criteria. Importantly, this association persisted after adjustment for maternal BMI and childhood adiposity, which are also associated with GDM in this cohort (13). Given the risk for progression to type 2 diabetes in children with IGT (42), our findings have important public health implications. A review of trials examining the effect of GDM treatment on child metabolic outcomes reported no difference in offspring of treated versus nontreated mothers, but the overall strength of evidence was deemed insufficient (43). With the
increasing prevalence of GDM and potential transgenerational impact of in utero exposure to GDM, future, wellpowered interventional trials are needed to address the impact of prevention and treatment of GDM diagnosed using IADPSG/ WHO criteria on subsequent childhood glucose outcomes.

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