# Coffee Consumption and Lung Cancer Risk: The Japan Public Health Center-Based Prospective Study 

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#### Abstract

Background: Many epidemiological studies have indicated a positive association between coffee intake and lung cancer risk, but such findings were suggested to be confounded by smoking. Furthermore, only a few of these studies have been conducted in Asia. Here, we investigated the association between coffee intake and lung cancer risk in one of the largest prospective cohort studies in Japan. Methods: We investigated the association of coffee drinking and subsequent incidence of lung cancer among 41,727 men and 45,352 women in the Japan Public Health Center-based Prospective Study using Cox proportional hazards regression, with adjustment for potential confounders and by strata of smoking status. Coffee and other dietary intakes were assessed once at baseline with a food frequency questionnaire ( FFQ ). Results: During 1,481,887 person-years of follow-up between 1990 and 2011, a total of 1,668 lung cancer cases were identified. In a multivariate regression model, coffee consumption was not associated with risk of lung cancer (HR 1.16; 95\% CI, $0.82-1.63 ; P_{\text {trend }}=0.285$ for men and HR 1.49;95\% CI, $0.79-2.83 ; P_{\text {trend }}=0.942$ for women). However, there was a significant increase in the risk for small cell carcinoma (HR 3.52; 95\% CI, 1.49-8.28; $P_{\text {trend }}<0.001$ ). Conclusion: Our prospective study suggests that habitual consumption of coffee is not associated with an increased risk of lung cancer incidence, despite observing a significant increase in the risk for small cell carcinoma.


Key words: coffee; lung cancer; Japan Public Health Center-based Prospective (JPHC) Study
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## INTRODUCTION

Coffee is one of the most widely consumed beverages around the world, including Japan. ${ }^{1,2}$ Since coffee drinking is often associated with smoking, it is sometimes considered part of an unhealthy lifestyle. However, coffee is a rich source of biochemically active compounds with anti-oxidative, ${ }^{3}$ antiinflammatory, ${ }^{4,5}$ and insulin-sensitizing effects. ${ }^{6,7}$ A number of studies have reported inverse associations between coffee intake and some cancers, as well as reduced mortality. ${ }^{4,8-15}$ Lung cancer, the world's leading cause of cancer death, ${ }^{16}$ however, appears to be an exception. More than $60 \%$ of the world's current smokers live in Asia, and the burden of lung cancer is expected to further increase in low- and middle-income countries. ${ }^{17}$

Epidemiological investigation into the association of coffee drinking and lung carcinogenesis has been extensive. The majority of studies have indicated positive associations. Nevertheless, adjustment for tobacco smoking remains a major issue in
these studies, as it may confound the associations. ${ }^{18-22}$ While recent findings from a meta-analysis of 17 studies ${ }^{18}$ and a large prospective cohort study ${ }^{19}$ reported an elevated risk of lung cancer, no significant association was observed in their stratified analyses among non-smokers. In addition, very few studies have explored the association by histological subtype, despite the fact that smoking is closely linked to squamous cell and small cell carcinomas, ${ }^{23}$ whereas the majority of diagnoses in never smokers are adenocarcinoma. ${ }^{24}$ Moreover, only a single hospitalbased case-control study ${ }^{25}$ and a few small-scale prospective cohort studies, ${ }^{15,26}$ with either limited information on exposure or case numbers, have been conducted in Japan.

Here, we conducted a population-based prospective cohort study to examine the association between coffee consumption and risk of lung cancer. We also investigated the association by histological subtype and comprehensive stratification of cigarette smoking.

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## MATERIALS AND METHODS

## Study population

The Japan Public Health Center-based Prospective Study (JPHC Study) is an ongoing cohort study designed to explore the associations between lifestyle factors and the incidence of diseases. ${ }^{27}$ A total of 140,420 residents ( 68,722 men and 71,698 women) in 11 public health center (PHC) areas nationwide (Iwate-Ninohe, Akita-Yokote, Nagano-Saku, TokyoKatsushika, Okinawa-Chubu, Niigata-Nagaoka, Ibaraki-Mito, Osaka-Suita, Kochi-Chuohigashi, Nagasaki-Kamigoto, and Okinawa-Miyako) aged 40-69 years at baseline were registered between 1990 and 1994. Details of the JPHC Study have been provided elsewhere. ${ }^{27}$ Participants from the Tokyo-Katsushika ( $n=7,097$ ) and Osaka-Suita ( $n=16,427$ ) PHC areas were excluded due to a lack of access to data on cancer incidence. We further excluded 229 participants with non-Japanese nationality, incorrect birth date, duplicate registration, or precommencement emigration, and a total of 455 subjects who had late report of emigration before the start of the follow-up period or were lost to follow-up.

Of 94,917 eligible participants who returned a completed selfadministered baseline questionnaire (response rate: $81.7 \%$ ), those who reported a past history of cancer at any site were also not included ( $n=1,915$ ). After excluding subjects with extreme caloric intake ( $\pm 2.5 \%$ of the distribution, $n=4,407$ ) and missing data on covariates, including coffee intake and cigarette smoking ( $n=1,516$ ), 41,727 men and 45,352 women were included in the analytic cohort. This study was approved by the institutional review boards of the National Cancer Center (approval number: 13-021) and the University of Tokyo (approval number: 10508).

## Exposure measurement

Dietary intake, including coffee, was measured through a validated self-administered food frequency questionnaire (FFQ). In the FFQ, coffee intake was assessed using six categories based on the frequency and the amount of consumption as follows: almost never, 1-2, 3-4 times/week, and almost everyday, which was further divided into $1-2,3-4$, and $\geq 5$ cups/day. Coffee consumption at baseline was used as an exposure in this study. Due to the small number of participants, occasional drinkers (1-2 and 3-4 times/week) were merged into one group ( $<1$ cup/day). The validity of the questionnaire was assessed using dietary records for 28 days (1-week dietary records measured quarterly) or 14 days. Spearman rank correlation coefficients between the dietary records and coffee intake estimated from the FFQ were 0.59 for men and 0.51 for women. ${ }^{28}$

## Follow-up and case ascertainment

Participants were followed from the date of baseline survey (1990-1994) until the date of diagnosis of cancer, date of death, date of migration out of the study area, or the end of follow-up (December 31, 2011), whichever occurred first. Death certificates were used with the permission of the Ministry of Health, Labour and Welfare to confirm the cause of death with the International Statistical Classification of Diseases and Related Health Problems, Tenth Revision. ${ }^{29}$ Among the study participants, 375 ( $0.4 \%$ ) were lost to follow-up during the follow-up period.

Identification of lung cancer cases was conducted through active patient notification from major local hospitals in each PHC area and record linkage with data from population-based cancer
registries. Information on diagnosis was supplemented with information from death certificates, and $7 \%$ of cases were ascertained from death certificates only. Lung cancer cases were confirmed microscopically in $87 \%$ of cases and coded using the International Classification of Diseases for Oncology, Third Edition (C34.0-34.9). ${ }^{30}$ Histological subtypes included adenocarcinoma, squamous cell carcinoma, small cell carcinoma, and other types of lung carcinoma. For participants who had multiple incidences of lung cancer, only the first incidence was used.

## Statistical analysis

Cox proportional hazards regression was used to estimate hazard ratios (HRs) and $95 \%$ confidence intervals (CIs) for lung cancer incidence with coffee consumption. Participants were divided into five groups according to coffee intake, with non-drinkers as the referent group. $P$-values for trends were calculated by assigning a continuous variable from the median value for coffee intake in the regression models. In primary analyses, risk estimates for overall lung cancer incidence and histological subtypes (adenocarcinoma, squamous cell carcinoma, small cell carcinoma, and other types of lung carcinoma) were determined in separate models. To further examine the effect of residual confounding by smoking, secondary analyses were stratified by gender and detailed smoking categories, namely smoking status (never, former, and current) and smoking intensity by cumulative smoking exposure in current and former smokers by pack-years ( $\leq 19.9$ or $\geq 20.0$ ). Pack-years were calculated by multiplying the average number of packs smoked per day with the number of years smoked. In order to maintain sample size in scarcely populated categories, HRs were estimated by merging coffee intake into three groups as follows: almost never, $<3$ cups/day, or $\geq 3$ cups/day. All covariates were assessed at baseline, and dietary intake was adjusted for total energy intake using the residual method (see Table 2 and Table 3 for details). Isoflavone was included as a confounding factor, since it was shown to have a protective effect against lung cancer risk in a previous JPHC Study. ${ }^{31} P$-values for interaction were calculated using likelihood-ratio tests comparing Cox proportional-hazards models with and without cross-product terms for the combination of lung cancer subtypes and coffee intake or smoking status and coffee intake, with coffee as a continuous term. The proportional hazards assumption was tested with Schoenfeld residuals and shown to be insignificant. To avoid potential bias due to undiagnosed lung cancer, sensitivity analyses were conducted after excluding lung cancer cases that were diagnosed within 5 years after baseline, and these analyses produced a similar direction of associations as the main analyses. All statistical analyses were performed with Stata MP version 13.0 (StataCorp, College Station, TX, USA).

## RESULTS

During an average follow-up of 17.0 years ( $1,481,887$ personyears), a total of 1,668 participants ( 1,227 men and 441 women) were newly diagnosed with lung cancer. Of histologically confirmed cases, $44 \%$ were adenocarcinoma ( $n=729$ ). More than $67 \%$ of the study participants reported drinking at least 1 cup/day of coffee (Table 1). Compared with non-drinkers, coffee drinkers were much more likely to smoke cigarettes. Male coffee drinkers tended to consume less alcohol, while female coffee drinkers were likely to report higher alcohol consumption. Coffee drinkers also tended to consume lower levels of fruits, vegetables,

Table 1. Baseline characteristics of the study participants by daily coffee consumption

| Characteristics | Men ( $n=41,727$ ) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | No Coffee | <1 Cup/day | 1-2 Cups/day | 3-4 Cups/day | $\geq 5$ Cups/day |  |
| Number of participants | 13,191 | 12,767 | 10,414 | 3,921 | 1,434 |  |
| Number of lung cancer cases | 403 | 370 | 296 | 109 | 49 |  |
| Age, years, mean (SD) | 53.7 (7.7) | 52.0 (7.7) | 50.6 (7.8) | 48.4 (7.3) | 48.9 (7.4) | <0.001 |
| Body mass index, $\mathrm{kg} / \mathrm{m}^{2}$, mean (SD) | 23.5 (2.9) | 23.6 (2.8) | 23.5 (2.8) | 23.4 (2.9) | 23.3 (3.1) | <0.001 |
| Diabetes, \% | 8.6 | 6.3 | 4.7 | 4.1 | 3.9 | <0.001 |
| Current smoker, \% | 43.4 | 50.0 | 56.2 | 69.6 | 76.9 | <0.001 |
| Physical activity almost daily, \% | 6.1 | 4.9 | 4.8 | 4.3 | 5.5 | <0.001 |
| Dietary intake |  |  |  |  |  |  |
| Total energy intake, kcal/day, mean (SD) | 1,952 (552) | 1,953 (551) | 1,901 (516) | 1,881 (512) | 1,919 (527) | <0.001 |
| Fruits, g/day, mean (SD) | 27 (30) | 27 (26) | 27 (28) | 26 (27) | 28 (37) | <0.001 |
| Vegetables, g/day, mean (SD) | 47 (45) | 45 (46) | 44 (42) | 40 (39) | 44 (53) | <0.001 |
| Isoflavone, mg/day, mean (SD) | 18 (10) | 17 (9) | 16 (9) | 15 (9) | 15 (9) | <0.001 |
| Alcohol consumption, g/week, mean (SD) | 205 (218) | 186 (209) | 174 (201) | 157 (197) | 154 (220) | <0.001 |
| Green tea $>1$ time/day, \% | 75.6 | 75.2 | 74.5 | 70.1 | 65.1 | <0.001 |
| Chinese tea $>1$ time/day, \% | 10.4 | 8.7 | 11.8 | 9.3 | 10.9 | <0.001 |
| Black tea $>1$ time/day, \% | 1.7 | 1.6 | 3.2 | 2.6 | 3.8 | <0.001 |
| Characteristics | Women ( $n=45,352$ ) |  |  |  |  | $P$ |
|  | No Coffee | <1 Cup/day | 1-2 Cups/day | 3-4 Cups/day | $\geq 5$ Cups/day |  |
| Number of participants | 15,202 | 13,554 | 12,573 | 3,155 | 868 |  |
| Number of lung cancer cases | 163 | 139 | 103 | 23 | 13 |  |
| Age, years, mean (SD) | 55.1 (7.7) | 52.4 (7.8) | 50.0 (7.5) | 47.3 (6.6) | 48.2 (7.3) | <0.001 |
| Body mass index, $\mathrm{kg} / \mathrm{m}^{2}$, mean (SD) | 23.7 (3.3) | 23.6 (3.1) | 23.4 (3.1) | 23.2 (3.0) | 23.4 (3.2) | <0.001 |
| Diabetes, \% | 4.9 | 2.5 | 2.0 | 1.3 | 1.3 | <0.001 |
| Current smoker, \% | 4.2 | 4.0 | 6.3 | 13.1 | 23.6 | <0.001 |
| Physical activity almost daily, \% | 5.8 | 4.4 | 4.4 | 3.2 | 5.6 | <0.001 |
| Dietary intake |  |  |  |  |  |  |
| Total energy intake, kcal/day, mean (SD) | 1,221 (305) | 1,248 (304) | 1,247 (296) | 1,230 (297) | 1,295 (322) | <0.001 |
| Fruits, g/day, mean (SD) | 102 (75) | 100 (66) | 95 (67) | 89 (75) | 90 (90) | <0.001 |
| Vegetables, g/day, mean (SD) | 112 (92) | 106 (82) | 103 (84) | 92 (81) | 93 (92) | <0.001 |
| Isoflavone, mg/day, mean (SD) | 26 (13) | 24 (12) | 23 (11) | 22 (12) | 20 (12) | <0.001 |
| Alcohol consumption, g/week, mean (SD) | 10 (54) | 11 (54) | 11 (50) | 17 (82) | 19 (58) | <0.001 |
| Green tea $>1$ time/day, \% | 77.1 | 77.8 | 74.7 | 66.4 | 66.6 | <0.001 |
| Chinese tea $>1$ time/day, \% | 12.1 | 11.0 | 14.0 | 13.7 | 20.3 | <0.001 |
| Black tea $>1$ time/day, \% | 1.8 | 1.8 | 4.5 | 3.4 | 5.0 | <0.001 |

SD denotes standard deviation. The body mass index is the weight divided by the square of the height in meters.
and fish, but not meat. Both male and female coffee drinkers were less likely to have diabetes.

In models adjusted for age and PHC area, coffee consumption was associated with an increased risk of lung cancer incidence (HR 1.88; 95\% CI, 1.39-2.54; $P_{\text {trend }}<0.001$ for men and HR 2.06; $95 \%$ CI, 1.16-3.67; $P_{\text {trend }}=0.253$ for women; comparing $\geq 5$ cups/day with non-drinkers) (Table 2). However, participants who routinely consumed coffee also tended to be smokers, and after adjusting for smoking, the association was notably attenuated (HR 1.14; 95\% CI, 0.83-1.55; $P_{\text {trend }}=0.405$ for men and HR 1.55; $95 \% \mathrm{CI}, 0.86-2.79 ; P_{\text {trend }}=0.839$ for women; comparing $\geq 5$ cups/day with non-drinkers). Similar associations were also observed for histological subtypes. The highest hazard was seen for small cell carcinoma, with HR for the highest level of coffee intake being reduced by three-quarters after adjustment for tobacco smoking and other confounders, without significant effect modification by histological subtypes ( $P_{\text {interaction }}=0.331$ for men; $P_{\text {interaction }}=0.169$ for women). In models separately analyzed by sex, hazard ratios comparing $\geq 5$ cups/day with nondrinkers for small cell carcinoma were higher for female coffee drinkers than male drinkers (HR 4.07; 95\% CI, 0.74-22.49; $P_{\text {trend }}=0.033$ for women and HR 2.82; 95\% CI, 1.00-7.94; $P_{\text {trend }}=0.005$ for men) (eTable 1). No other clear gender differences were observed in the histological analyses.

Approximately half of male participants were current smokers, whereas $93 \%$ of female participants were never-smokers (Table 3). In stratified analyses by smoking strata, current and former male smokers were examined by pack-years. A higher level of coffee intake among male heavy smokers who had $\geq 20.0$ pack-years of tobacco exposure was associated with an increased risk of lung cancer (HR 1.30; $95 \% \mathrm{CI}, 1.04-1.63 ; P_{\text {trend }}=$ 0.014 , comparing $\geq 3$ cups/day with non-drinkers), whereas no substantial effect modification by smoking status was observed in both men and women ( $P_{\text {interaction }}=0.266$ for men and $P_{\text {interaction }}=0.243$ for women). While tea is also known to have free radical scavenging activity and chemopreventive effects on lung cancer, ${ }^{32,33}$ a separate analysis on tea, including green tea, Chinese (oolong) tea, and black tea, and lung cancer did not find a significant association (eTable 2).

## DISCUSSION

In this prospective Japanese cohort study, our overall results suggested that coffee consumption is not significantly associated with an increased risk of lung cancer, after adjustment for smoking and other potential confounders. Residual confounding by smoking is a major concern in lung cancer research; the majority of our present coffee drinkers tended to be smokers, and

Table 2. Hazard ratios (95\% confidence intervals) for lung cancer incidence by daily coffee consumption

|  | Coffee intake |  |  |  |  | $P$-trend | $P$-interaction |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | No Coffee | <1 Cup/day | 1-2 Cups/day | 3-4 Cups/day | $\geq 5$ Cups/day |  |  |
| Lung cancer (all) |  |  |  |  |  |  |  |
| Men |  |  |  |  |  |  | 0.331 |
| Number of cases | 403 | 370 | 296 | 109 | 49 | - | - |
| Age, area | 1.00 | 1.07 (0.93-1.23) | 1.27 (1.09-1.48) | 1.54 (1.24-1.91) | 1.88 (1.39-2.54) | $<0.001$ | - |
| Age, area, smoking | 1.00 | 0.95 (0.82-1.10) | 1.05 (0.89-1.23) | 1.02 (0.82-1.28) | 1.14 (0.83-1.55) | 0.405 | - |
| Multivariate-adjusted | 1.00 | 1.03 (0.88-1.22) | 1.11 (0.93-1.22) | 1.05 (0.82-1.34) | 1.16 (0.82-1.63) | 0.285 | - |
| Women |  |  |  |  |  |  | 0.169 |
| Number of cases | 163 | 139 | 103 | 23 | 13 | - | - |
| Age, area | 1.00 | 1.08 (0.86-1.36) | 1.00 (0.77-1.30) | 1.08 (0.69-1.70) | 2.06 (1.16-3.67) | 0.253 | - |
| Age, area, smoking | 1.00 | 1.08 (0.86-1.36) | 0.95 (0.73-1.24) | 0.90 (0.56-1.43) | 1.55 (0.86-2.79) | 0.839 | - |
| Multivariate-adjusted | 1.00 | 1.06 (0.82-1.36) | 0.94 (0.70-1.25) | 0.90 (0.55-1.48) | 1.49 (0.79-2.83) | 0.942 | - |
| Adenocarcinoma |  |  |  |  |  |  |  |
| Number of cases | 266 | 216 | 176 | 46 | 25 | - | - |
| Age, gender, area | 1.00 | 0.97 (0.81-1.16) | 1.10 (0.90-1.33) | 1.06 (0.77-1.46) | 1.66 (1.09-2.51) | 0.087 | - |
| Age, gender, area, smoking | 1.00 | 0.91 (0.76-1.09) | 0.97 (0.80-1.19) | 0.83 (0.59-1.16) | 1.21 (0.79-1.86) | 0.829 | - |
| Multivariate-adjusted | 1.00 | 0.95 (0.78-1.16) | 1.01 (0.81-1.26) | 0.80 (0.56-1.16) | 1.13 (0.70-1.83) | 0.778 | - |
| Squamous cell carcinoma |  |  |  |  |  |  |  |
| Number of cases | 107 | 106 | 76 | 30 | 16 | - | - |
| Age, gender, area | 1.00 | 1.20 (0.91-1.57) | 1.31 (0.97-1.77) | 1.80 (1.19-2.73) | 2.55 (1.50-4.35) | <0.001 | - |
| Age, gender, area, smoking | 1.00 | 1.06 (0.80-1.40) | 1.06 (0.78-1.45) | 1.10 (0.71-1.70) | 1.45 (0.84-2.49) | 0.303 | - |
| Multivariate-adjusted | 1.00 | 1.14 (0.83-1.55) | 1.10 (0.78-1.55) | 1.04 (0.63-1.69) | 1.24 (0.66-2.32) | 0.578 | - |
| Small cell carcinoma |  |  |  |  |  |  |  |
| Number of cases | 28 | 61 | 37 | 21 | 9 | - | - |
| Age, gender, area | 1.00 | 2.60 (1.66-4.08) | 2.28 (1.38-3.77) | 4.50 (2.51-8.06) | 5.39 (2.51-11.58) | $<0.001$ | - |
| Age, gender, area, smoking | 1.00 | 2.47 (1.56-3.93) | 1.88 (1.12-3.16) | 2.83 (1.54-5.17) | 2.62 (1.16-5.90) | 0.003 | - |
| Multivariate-adjusted | 1.00 | 2.51 (1.47-4.30) | 2.35 (1.33-4.17) | 3.48 (1.79-6.73) | 3.52 (1.49-8.28) | $<0.001$ | - |
| Other |  |  |  |  |  |  |  |
| Number of cases | 165 | 126 | 110 | 35 | 12 | - | - |
| Age, gender, area | 1.00 | 0.93 (0.74-1.17) | 1.17 (0.92-1.51) | 1.33 (0.92-1.94) | 1.27 (0.70-2.29) | 0.072 | - |
| Age, gender, area, smoking | 1.00 | 0.87 (0.68-1.10) | 1.00 (0.77-1.29) | 0.94 (0.64-1.38) | 0.83 (0.46-1.38) | 0.707 | - |
| Multivariate-adjusted | 1.00 | 0.94 (0.72-1.23) | 1.00 (0.75-1.35) | 0.99 (0.65-1.50) | 1.02 (0.55-1.88) | 0.958 | - |

Multivariate analyses were adjusted for the following factors at baseline: age (continuous); gender; PHC area; body-mass index (<18.5, 18.5-24.9, 25.0-29.9, and $\geq 30 \mathrm{~kg} / \mathrm{m}^{2}$ ); smoking status (never, former: $<10,10-19$, and $\geq 20$ years of smoking cessation, current: $1-19,20-29,30-39,40-49,50-59, \geq 60$ pack-years); physical activity (almost never, <3 days/month, 1-2 days/week, 3-4 days/week, and almost everyday); alcohol consumption (never/former, <1 time/week, $<23,23-45,46-68,69-91$, and $>92 \mathrm{~g}$ of ethanol/day); consumption of green tea, Chinese tea, and black tea (almost never, $<1$ time/week, and $>1$ cup/day); energy-adjusted intake of fruit, vegetables, and isoflavone (continuous).
the adjustment for cigarette smoking might have been incomplete, which would have attenuated our results. The significant positive association among heavy smokers might be explained via the residual confounding by smoking. An updated meta-analysis of five cohort and 12 case-control studies indicated that recently published studies consistently reported a positive association of coffee intake with lung cancer incidence, despite a reduction in the risk of other types of cancers, including prostate, bladder, colorectal, and liver cancers. ${ }^{18}$ While the positive associations were consistent not only in case-control but also prospective cohort studies, many were detected among smokers, so cautious interpretation was recommended. ${ }^{19,34,35}$ Furthermore, the possibility of passive inhalation of cigarette smoke by non-smokers should also be considered. Presence or absence of passive smoking at home and work was included in a separate model in an attempt to assess environmental exposure, although the period and intensity of exposure could not be measured due to data unavailability. While this analysis may have only partially captured the influence of passive smoking, the risk estimates were not substantially altered (data not shown). Null associations between coffee intake and lung cancer risk observed among nonsmokers in the present study were consistent with the majority of recent studies. ${ }^{18,19}$

In the stratified analyses by subsite of lung cancer, an increased risk of small cell carcinoma was found among coffee drinkers. This finding was congruous with a recent large-scale prospective cohort study in the United States wherein coffee intake was positively associated with small cell carcinoma. ${ }^{19}$ Smoking has been closely linked with small cell and squamous cell carcinomas in previous studies, ${ }^{23,36}$ and we found a similar association in the present study, with approximately $92 \%$ ( $126 / 137$ cases) of small cell carcinoma cases in men observed among heavy smokers with $\geq 20.0$ pack-years (data not shown). The steep increase in the risk of these subtypes of lung carcinoma could be explained via residual confounding by smoking. The oncogenic mechanism of this association has yet to be elucidated, but one possible pathway might be through the activity of TP53 and RB1, the tumor suppressor genes that play a crucial role in the oncogenesis of small cell carcinoma. ${ }^{37}$ Contrary to its chemopreventive components, ${ }^{3-5,38,39}$ coffee has been recently discovered to contain gallic acid and pyrogallol, which are considered responsible for DNA-damaging activities involving TP53. ${ }^{40}$ Further analysis is required to examine the potential impact of coffee intake on the etiology of subtypes of lung carcinogenesis and to investigate the possibility of chance findings, as there were only modest numbers of cases in some of our subanalyses.

Table 3. Subgroup analysis by smoking status: hazard ratios ( $95 \%$ confidence intervals) for lung cancer incidence by daily coffee consumption

|  | All Participants | Coffee intake |  |  | $P$-trend | $P$-interaction |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | No Coffee | <3 Cups/day | $\geq 3$ Cups/day |  |  |
| Men |  |  |  |  |  | 0.266 |
| Never-smokers |  |  |  |  |  |  |
| Number of participants | 10,138 | 3,804 | 5,615 | 719 | - | - |
| Number of cases | 87 | 40 | 40 | 7 | - | - |
| Age, area |  | 1.00 | 0.75 (0.48-1.17) | 1.20 (0.53-2.71) | 0.587 | - |
| Multivariate-adjusted |  | 1.00 | 0.81 (0.49-1.33) | 1.41 (0.58-3.46) | 0.940 | - |
| Former smokers |  |  |  |  |  |  |
| Number of participants | 9,789 | 3,660 | 5,323 | 806 | - | - |
| Number of cases | 188 | 75 | 105 | 8 | - | - |
| Age, area |  | 1.00 | 1.11 (0.82-1.50) | 0.85 (0.41-1.78) | 0.802 | - |
| Multivariate-adjusted |  | 1.00 | 1.25 (0.90-1.74) | 0.88 (0.40-1.95) | 0.461 | - |
| Current smokers |  |  |  |  |  |  |
| Number of participants | 21,800 | 5,727 | 12,243 | 3,830 | - | - |
| Number of cases | 952 | 288 | 521 | 143 | - | - |
| Age, area |  | 1.00 | 1.04 (0.90-1.20) | 1.18 (0.96-1.46) | 0.155 | - |
| Multivariate-adjusted |  | 1.00 | 1.12 (0.95-1.32) | 1.24 (0.99-1.57) | 0.058 | - |
| Light smokers: pack-years $\leq 19.9^{\text {a }}$ |  |  |  |  |  |  |
| Number of participants | 7,846 | 2,654 | 4,446 | 746 | - | - |
| Number of cases | 100 | 38 | 58 | 4 | - | - |
| Age, area |  | 1.00 | 1.14 (0.75-1.73) | 0.61 (0.21-1.73) | 0.876 | - |
| Multivariate-adjusted |  | 1.00 | 1.35 (0.83-2.18) | 0.92 (0.31-2.68) | 0.500 | - |
| Heavy smokers: pack-years $\geq 20.0^{a}$ |  |  |  |  |  |  |
| Number of participants | 23,743 | 6,733 | 13,120 | 3,890 | - | - |
| Number of cases | 1,040 | 325 | 568 | 147 | - | - |
| Age, area |  | 1.00 | 1.08 (0.94-1.24) | 1.29 (1.05-1.58) | 0.023 | - |
| Multivariate-adjusted |  | 1.00 | 1.16 (0.99-1.35) | 1.30 (1.04-1.63) | 0.014 | - |
| Women |  |  |  |  |  | 0.243 |
| Never-smokers |  |  |  |  |  |  |
| Number of participants | 42,155 | 14,374 | 24,458 | 3,323 | - | - |
| Number of cases | 375 | 147 | 204 | 24 | - | - |
| Age, area |  | 1.00 | 1.01 (0.81-1.25) | 1.14 (0.72-1.79) | 0.716 | - |
| Multivariate-adjusted |  | 1.00 | 1.00 (0.79-1.28) | 1.13 (0.70-1.84) | 0.759 | - |
| Current smokers |  |  |  |  |  |  |
| Number of participants | 2,600 | 645 | 1,337 | 618 | - | - |
| Number of cases | 61 | 14 | 35 | 12 | - | - |
| Age, area |  | 1.00 | 1.42 (0.76-2.67) | 1.39 (0.61-3.14) | 0.380 | - |
| Multivariate-adjusted |  | 1.00 | 1.27 (0.66-2.45) | 1.38 (0.58-3.28) | 0.441 | - |

Multivariate analyses were adjusted for the following factors at baseline: age (continuous); gender; PHC area; body-mass index (<18.5, 18.5-24.9, 25.0-29.9, and $\geq 30 \mathrm{~kg} / \mathrm{m}^{2}$ ); smoking status (never, former: $<10,10-19$, and $\geq 20$ years of smoking cessation, current: $1-19,20-29,30-39,40-49,50-59, \geq 60$ pack-years); physical activity (almost never, <3 days/month, 1-2 days/week, 3-4 days/week, and almost everyday); alcohol consumption (never/former, <1 time/week, $<23,23-45,46-68,69-91$, and $>92 \mathrm{~g}$ of ethanol/day); consumption of green tea, Chinese tea, and black tea (almost never, $<1$ time/week, and $>1$ cup/day); energy-adjusted intake of fruit, vegetables, and isoflavone (continuous).
${ }^{\text {a }}$ Model included current and former smokers. Pack-years is the number of cigarette packs (assuming 20 cigarettes per pack) smoked per day, multiplied by years of consumption.

Some limitations of the present study include the unavailability of information on types of coffee (caffeinated or decaffeinated), preparation methods (eg, drip, instant, or canned), duration of coffee consumption, and cumulative intake of coffee and caffeine, although consumption of decaffeinated coffee is infrequent in Japan. A recent prospective cohort study that investigated the effect of caffeine intake on cancer risk reported a null association with lung cancer. ${ }^{21}$ While coffee drinking is considered to be relatively stable over time, coffee intake was evaluated at baseline using a single self-reported measurement, and the possibility of misclassification cannot be ruled out. However, such misclassification due to fluctuation in habitual coffee consumption during the follow-up period would likely have occurred randomly and nondifferentially and might have attenuated the results towards null. We also excluded participants who reported a history of
cancer at any site and conducted a sensitivity analysis excluding participants with newly diagnosed lung cancer, which would refute a possibility of a bias from ongoing or undiagnosed illnesses. Further, cumulative exposure to cigarette smoking was taken into consideration as pack-years, but more detailed data on smoking, such as depth of inhalation or nicotine dose, was unavailable.

Despite these limitations, this is the first population-based prospective cohort study to investigate the association between coffee intake and lung cancer incidence in Japan. The main strengths of our study include its high response rate ( $81.7 \%$ ), long follow-up period and negligible number of lost participants $(0.4 \%)$, and minimal recall bias due to its prospective design. Misclassification of lung cancer cases was also unlikely because of the rate of microscopic diagnosis and limited reliance on death certificate notification.

In summary, we found that a higher level of coffee intake was not associated with an increased risk of lung cancer, although we did observe an elevated risk of small cell carcinoma. This positive association was substantially attenuated after adjusting for cigarette smoking. Further evidence is needed to determine the associations with subtypes of lung cancer. Future studies that capture more detailed quantitative information on tobacco exposure would help clarify the impact of residual confounding on lung carcinogenesis.

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Authors' contributions: M. Inoue and ST designed research; SN conducted research; SN and ES analyzed data; SN wrote the paper; M. Inoue had primary responsibility for final content; TY and MN provided with helpful comments regarding intellectual content; ES, NS, TS, M. Iwasaki, NS and SS made substantial contributions to strengthen the methods and discussion sections. All authors read and approved the final manuscript.

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## APPENDIX A. SUPPLEMENTARY DATA

Supplementary data related to this article can be found at https:// doi.org/10.2188/jea.JE20160191.

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