

Bidirectional Relationship Between Diabetes and Acute Pancreatitis

A Population-Based Cohort Study in Taiwan

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Abstract: The proposed bidirectional relationship between acute pancreatitis (AP) and diabetes has never been examined with the same source of data. Furthermore, the effects of disease severity on this relationship have not been fully evaluated. The present study employed the findings from a single database to measure the strength of the association between AP and diabetes.

Findings from 1 million National Health Insurance beneficiaries were utilized. Two cohort studies with this database were selected to evaluate the linkage between diabetes and AP. The first cohort analysis addressed the risk of AP among diabetic patients and was comprised of 42,080 diabetic patients and 672,146 unexposed subjects. The second cohort analysis considered the risk of diabetes among patients with AP and enrolled 3187 patients with AP and 709259 unexposed subjects. All adult beneficiaries were followed from January 1, 2005 to December 31, 2012 to identify outcomes of interest. Cox regression models were applied to compare hazards adjusted for potential confounders.

For the first cohort, the adjusted hazard ratio (HR) of AP was significantly increased by the presence of diabetes (1.72; 95% confidence interval [CI], 1.52–1.96). In diabetic patients with a history of hyperglycemic crisis episodes (HCEs), the HR was even higher (6.32; 95% CI, 4.54–8.81). For the second cohort, the adjusted HR of diabetes in patients with AP was increased compared to the general population (2.15; 95% CI, 1.92–2.41). For patients with severe AP, the HR was also higher (2.22; 95% CI, 1.50–3.29) but did not differ significantly from that for patients with nonsevere AP.

The 2 cohort studies provided evidence for the bidirectional relationship between diabetes and AP. Moreover, diabetic patients with history of HCEs may be associated with higher risk of AP.

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Abbreviations: AP = acute pancreatitis, CCI = Charlson Comorbidity Index, CI = confidence interval, HCE = hyperglycemic crisis episode, HR = hazard ratio, ICD-9-CM = International Classification of Diseases (Ninth Revision) Clinical Modification, NHI = National Health Insurance, SES = socioeconomic status.

INTRODUCTION

Acute pancreatitis (AP), a condition characterized by pancreatic inflammation, is often associated with increased morbidity and mortality.^{1,2} The etiology of AP is multifactorial, with factors such as alcohol consumption, gallstones, certain drugs, renal insufficiency, and hypertriglyceridemia reported to increase the risk of this condition.^{3–5} Additionally, recent findings of several large cohort studies reveal that patients with diabetes mellitus are at increased risk of AP.^{6–8} Despite this well-supported linkage, no study has yet been performed to examine the possibility that the risk of AP increases with the severity of diabetes.

Although hyperglycemia is commonly observed in patients with AP,^{9,10} findings of studies regarding newly diagnosed diabetes in subjects with AP are conflicting. Whereas a positive relationship between pre-existing AP and development of diabetes has been observed in some studies,^{11,12} no association was found in others.¹³ Authors of a recently published meta-analysis concluded that patients treated for AP may be at higher risk of developing diabetes after discharge; however, the severity of AP had a minimal effect on the studied outcomes in this analysis.¹⁴

To the knowledge of the authors of this report, no study investigating the bidirectional relationship between AP and diabetes has been performed with the same population and with the same source of data. Furthermore, the effects of disease severity on this relationship have not been fully evaluated. The present study employed a large administrative database from the Taiwan National Health Insurance (NHI) program, which provides healthcare coverage to Taiwan residents of all ages. The overall goal was to assist clinicians in the identification of individuals at risk of these frequently encountered diseases. Given the increasing incidence of both diseases,^{15,16} efforts to prevent their occurrences should result in a decrease in the population-attributable risk percentage.

METHODS

Ethics Statement

This study was initiated after approval from the Institutional Review Board of Dalin Tzu Chi Hospital, Buddhist Tzu

Chi Medical Foundation, Taiwan. Since all personal identification was removed from the secondary files before any analysis was performed, the review board waived the requirement for written informed consent from the involved patients.

Database

The NHI program was implemented in 1995 and provides compulsory universal health insurance, which enrolls up to 99% of the Taiwanese population and has contracts with 97% of all medical providers.¹⁷ The database analyzed in this study included 1 million beneficiaries randomly selected from all beneficiaries insured in 2005. Statistically significant differences were not observed for this group as compared to the larger cohort with respect to age, sex, or healthcare costs according to the Taiwan National Health Research Institute.^{18,19} The database contains comprehensive information on all insured individuals, including gender, date of birth, residential location, dates of clinical visits, the International Classification of Diseases (Ninth Revision) Clinical Modification (ICD-9-CM) diagnostic codes, details regarding prescribed medications, and outcomes at hospital discharge.^{20,21}

Analysis of the Association Between Diabetes and Risk of Acute Pancreatitis

The sampled population was followed from January 1, 2003 to December 31, 2012. Individuals were initially identified for the study cohort who were still alive in 2005 and who were older than 18 years. Patients diagnosed with diabetes (ICD-9-CM code: 250) were then identified from records of their ambulatory care claims during the follow-up period.²² To avoid coding error, an individual could be classified as a diabetic patient only if she/he received a diagnosis of diabetes and then experienced another one or more diagnoses within the subsequent 12 months. Moreover, the first and last visits during the 10-year period had to be separated by at least 30 days to avoid

accidental inclusion of patients with miscoded diagnoses. AP was defined by the presence of ICD-9-CM code 577.0 in any position of the diagnoses. In order to maximize case ascertainment, only patients hospitalized for AP were included. These selection processes have been well-validated with high positive predictive values.^{4,23}

After excluding patients with diabetes and AP before January 1, 2005, 42,080 patients were included in the diabetic group and 672,146 in the unexposed group. Diabetes with poor compliance was defined as a hospitalization with the diagnosis of a hyperglycemic crisis episode (HCE) (ICD-9-CM codes: diabetic ketoacidosis, 250.1 or hyperosmolar hyperglycemic state, 250.2).²⁴ The index date for each diabetic patient was the date of his or her first diabetes diagnosis. The index date for subjects in the unexposed group was set as January 1, 2005. Subjects in the diabetic and unexposed groups were then followed through December 31, 2012 for possible episodes of AP. Cases were censored for patients who were no longer beneficiaries of the NHI Program (ie, death or transfer out) or who were still robust at the end of the follow-up period (Figure 1A).

Analysis of the Association Between Acute Pancreatitis and Risk of Diabetes

The second cohort analysis assessed the association of AP with an increased risk of diabetes. The diagnostic criteria for AP and diabetes were the same as those for the first cohort. In total, 3187 subjects with AP were enrolled in the AP group and 709,259 were enrolled in the unexposed group after excluding patients with diabetes and AP before January 1, 2005. Severe AP was defined as a hospitalization in an intensive care unit setting.² The index date for each AP patient was the date of his or her first admission. The index date for subjects in the unexposed group was set as January 1, 2005. Subjects in the AP and unexposed groups were then followed until December 31, 2012 for possible diagnosis of diabetes. Cases were censored for patients who were

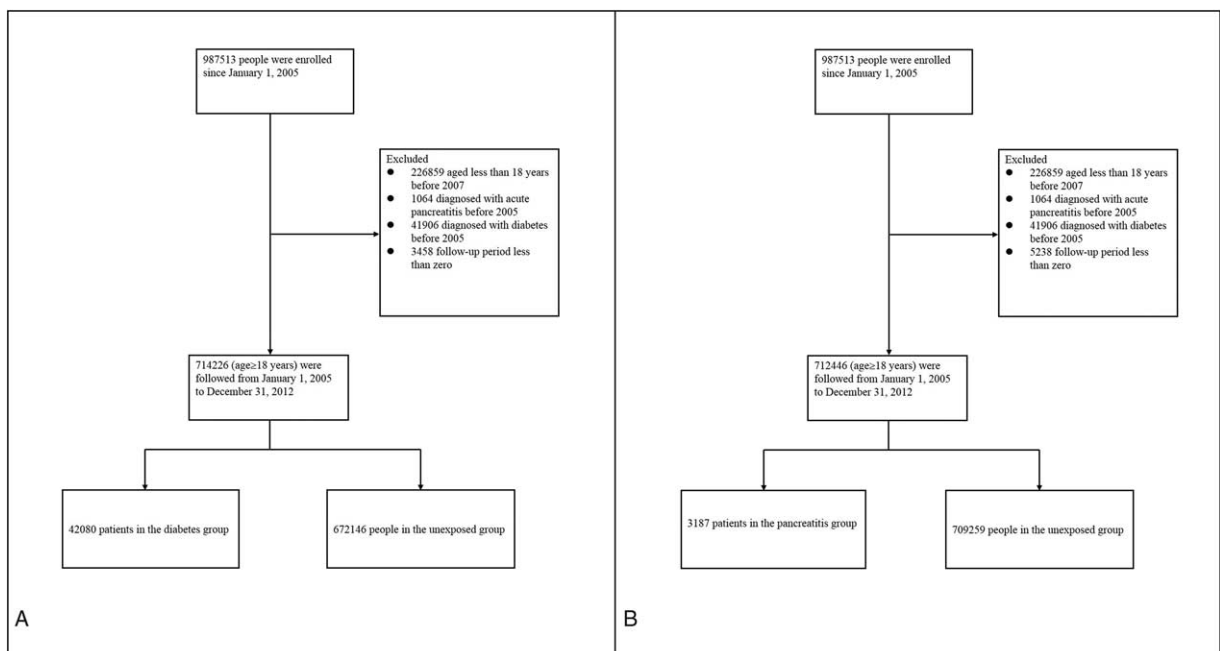


FIGURE 1. Flow diagram of the population-based study. (A) Diabetes and the risk of acute pancreatitis. (B) Pancreatitis and the risk of diabetes.

no longer beneficiaries of the NHI Program (ie, death or transfer out) or who were still robust at the end of the follow-up period (Figure 1B).

Covariates

To better characterize the relationships between diabetes and AP, several covariates were used, including patient demographics such as age, gender, urbanization level, and socioeconomic status (SES). The age of each patient was defined by the difference between the index date and the date of birth. Income-related insurance payment amounts were used as a proxy measure of individual SES at follow-up.²⁰

Additionally, the prevalence of specific comorbid conditions (chronic liver disease, hypertension, coronary artery disease, hypertriglyceridemia, malignancies, smoking, obesity, biliary tract disease, chronic renal insufficiency, and a history of alcohol intoxication) and the Charlson Comorbidity Index (CCI) score were determined according to discharge diagnoses following either outpatient clinic visits or hospitalizations before January 1, 2005. The detailed ICD-9-CM codes for comorbidities are described elsewhere, and the processes for selecting comorbidities are widely used and accepted.^{4,18,23,25}

Statistical Analysis

All covariates were taken as categorical variables excepting age, which was treated as a continuous variable. To reveal

the baseline heterogeneity in the 2 groups, categorical variables were compared with the Pearson Chi-square test and continuous variables were compared with the *t* test. The Nelson–Aalen cumulative hazard estimates were plotted initially to reveal the trends. Cox proportional hazard regression models were then used to calculate the hazard ratios (HRs) of AP for individuals with diabetes and of diabetes for individuals with AP; calculations were performed after adjustments for age, gender, urbanization level, SES, chronic liver disease, hypertension, coronary artery disease, hypertriglyceridemia, history of alcohol intoxication, malignancies, smoking, obesity, biliary tract disease, chronic renal insufficiency, and the CCI. The SAS statistical package, version 9.4 (SAS Institute, Inc., Cary, NC), and STATA version 11.2 (StataCorp, College Station, TX) were used for data analysis. A 2-tailed *P*-value of <0.05 was considered significant.

RESULTS

Cohort Analysis of the Association Between Diabetes and Risk of Acute Pancreatitis

The distributions of demographic characteristics and selected comorbidities for the study subjects are shown in Table 1. There were 42,080 patients in the diabetes group (982 with HCE) and 672,146 in the unexposed group. The average follow-up durations for the diabetes and exposed

TABLE 1. Baseline Characteristics of the Diabetes Group and the Unexposed Group

Variables	Diabetes (n = 41,098)		Diabetics With Poor Control (n = 982)		Unexposed Group (n = 672,146)		P-Value
	No.	%	No.	%	No.	%	
Admitted for acute pancreatitis	275	0.7	36	3.7	2838	0.4	<0.001
Male	21,780	53.0	591	60.2	327,721	48.8	<0.001
Mean age (SD)	57.3	13.5	58.2	18.3	41.1	15.9	<0.001
Socioeconomic status							<0.001
Low	20,421	49.7	625	63.6	274,127	40.8	
Moderate	14,950	36.4	303	30.9	268,739	40.0	
High	5727	13.9	54	5.5	129,280	19.2	
Urbanization level							<0.001
Urban	11,420	27.8	218	22.2	202,702	30.2	
Suburban	18,521	45.1	441	44.9	311,118	46.3	
Rural	11,157	27.1	323	32.9	158,326	23.5	
Charlson Comorbidity Index score							<0.001
0	23,337	56.8	503	51.2	481,485	71.6	
1	10,367	25.2	243	24.8	121,505	18.1	
≥2	7394	18.0	236	24.0	69,156	10.3	
Chronic liver disease	5688	13.8	109	11.1	50,317	7.5	<0.001
Hypertension	13,892	33.8	331	33.7	74,254	11.0	<0.001
Coronary artery disease	4934	12.0	112	11.4	30,049	4.5	<0.001
Hypertriglycemia	469	1.1	9	0.9	2170	0.3	<0.001
Malignancies	908	2.2	35	3.6	9588	1.4	<0.001
Smoking	19	0.1	2	0.2	280	0.04	0.043
Obesity	118	0.3	1	0.1	767	0.1	<0.001
History of alcohol intoxication	446	1.1	22	2.2	4328	0.6	<0.001
Chronic renal insufficiency	424	1.0	21	2.1	4102	0.6	<0.001
Biliary tract disease	249	0.6	0	0	2465	0.4	<0.001

SD = standard deviation.

p-Values less than 0.05 are labeled as bold types.

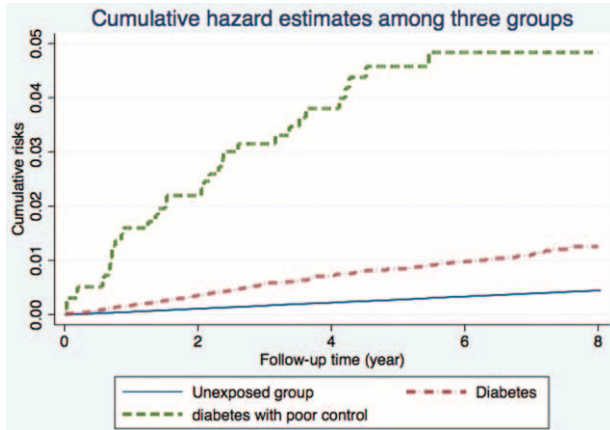


FIGURE 2. Nelson–Aalen curves showing a higher cumulative risk of acute pancreatitis in the diabetes group.

groups were 3.98 and 7.60 years, respectively. Patients with diabetes were predominantly male and significantly older. These patients were also more likely to have lower SES values, higher CCI scores, chronic liver disease, hypertension, coronary artery disease, hypertriglyceridemia, malignancies, chronic renal insufficiency, and biliary tract disease, and were more likely to be smokers, to be obese, to reside in nonrural areas, and to have a history of alcohol intoxication.

At the end of follow-up, 3149 patients had been admitted for AP; this group included 311 diabetic patients (36 with HCE) and 2838 nondiabetics. The average time from diabetes to AP was 3.80 years. The incidence rate of AP per 1000 person-years was 1.7 for patients with diabetes, 8.2 for diabetic patients with HCE, and 0.6 for nondiabetics. As compared to the general population, the crude HRs of AP for diabetic patients and for diabetics with HCE were 3.03 (95% confidence interval [CI],

2.67–3.43) and 14.75 (95% CI, 10.62–20.50), respectively. The Nelson–Aalen plot showed higher cumulative risk of AP for both diabetes groups (Figure 2).

The multivariate Cox regression model was then employed to determine the adjusted HRs of AP. After controlling for the above-mentioned covariates, an increased HR was still observed for diabetic patients (1.72; 95% CI, 1.52–1.92). For the subgroup of diabetic patients with HCE, the HR was also significantly higher and greater than that for diabetics with no history of HCE (6.32; 95% CI, 4.54–8.81). Other independent risk factors for AP included male gender, older age, lower SES, living outside of an urban area, higher CCI, chronic liver disease, hypertension, history of alcohol intoxication, chronic renal insufficiency, and biliary tract disease (Figure 3). Findings with relevant statistics are summarized in Table 2.

Cohort Analysis of the Association Between Acute Pancreatitis and Risk of Diabetes

The distributions of demographic characteristics and selected comorbidities for the study subjects are shown in Table 3. There were 3187 patients in the AP group and 709,259 in the unexposed group. Of those in the AP group, 255 were further defined as having severe AP. The average follow-up durations for the AP and unexposed groups were 3.21 and 7.39 years, respectively. Patients with AP were predominantly male and significantly older. They were also more likely to have lower SES values, higher CCI scores, chronic liver disease, hypertension, coronary artery disease, hypertriglyceridemia, malignancies, chronic renal insufficiency, and biliary tract disease, and were more likely to be smokers, to be obese, to reside in nonrural areas, and to have a history of alcohol intoxication.

At the end of follow-up, 41681 patients had been diagnosed with diabetes; this group included 324 patients with AP (25 with severe AP) and 41,357 without AP. The average time

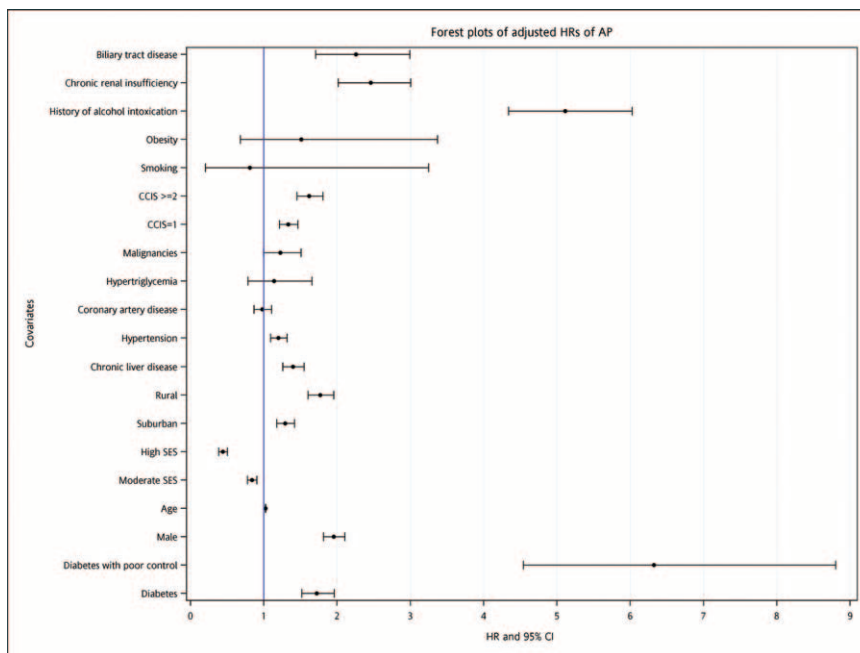


FIGURE 3. Forest plots of adjusted hazard ratios for acute pancreatitis.

TABLE 2. Adjusted HRs of AP in Patients With Diabetes

Variables	Hazard Ratio	95% Confidence Interval	P-Value
Diabetes	1.72	1.52–1.96	<0.001
Diabetes with poor control	6.32	4.54–8.81	<0.001
Male	1.96	1.82–2.11	<0.001
Patient age	1.02	1.02–1.03	<0.001
Socioeconomic status			<0.001
Low	1	—	—
Moderate	0.84	0.78–0.90	<0.001
High	0.44	0.38–0.50	<0.001
Urbanization level			
Urban	1	—	—
Suburban	1.29	1.18–1.42	<0.001
Rural	1.77	1.60–1.96	<0.001
Charlson Comorbidity Index score			
0	1	—	—
1	1.33	1.22–1.47	<0.001
≥2	1.62	1.45–1.81	<0.001
Chronic liver disease	1.40	1.26–1.55	<0.001
Hypertension	1.20	1.09–1.32	<0.001
Coronary artery disease	0.98	0.87–1.11	0.726
Hypertriglycemia	1.14	0.78–1.66	0.496
Malignancies	1.22	0.99–1.51	0.054
Smoking	0.81	0.20–3.25	0.768
Obesity	1.51	0.68–3.37	0.313
History of alcohol intoxication	5.11	4.34–6.03	<0.001
Chronic renal insufficiency	2.46	2.01–3.00	<0.001
Biliary tract disease	2.26	1.71–2.99	<0.001

p-Values less than 0.05 are labeled as bold types.

TABLE 3. Baseline Characteristics of the Acute Pancreatitis Group and the Unexposed Group

Variables	Pancreatitis Group (n = 2932)		Severe Pancreatitis Group (n = 255)		Unexposed Group (n = 709,259)		P-Value
	No.	%	No.	%	No.	%	
Event	299	10.2	25	9.8	41,357	5.8	<0.001
Male	1942	66.2	171	67.1	347,178	48.9	<0.001
Mean age (SD)	53.8	17.7	63.6	18.9	41.9	16.1	<0.001
Socioeconomic status							<0.001
Low	1475	50.3	159	62.3	292,652	41.2	
Moderate	1185	40.4	92	36.1	282,078	39.8	
High	272	9.3	4	1.6	134,529	19.0	
Urbanization level							<0.001
Urban	624	21.3	44	17.3	213,062	30.1	
Suburban	1263	43.1	114	44.7	327,829	46.2	
Rural	1045	35.6	97	38.0	168,368	23.7	
Charlson Comorbidity Index score							<0.001
0	1469	50.1	111	43.5	502,402	70.8	
1	767	26.2	57	22.4	130,986	18.5	
≥2	696	23.7	87	34.1	75,871	10.7	
Chronic liver disease	531	18.1	48	18.8	55,410	7.8	<0.001
Hypertension	741	25.3	85	33.3	87,479	12.3	<0.001
Coronary artery disease	297	10.1	47	18.4	34,687	4.9	<0.001
Hypertriglycemia	24	0.8	0	0	2616	0.4	<0.001
Malignancies	93	3.2	14	5.5	10,414	1.5	<0.001
Smoking	2	0.1	0	0	299	0.1	0.749
Obesity	4	0.1	2	0.8	876	0.1	0.011
History of alcohol intoxication	169	5.8	15	5.9	4610	0.7	<0.001
Chronic renal insufficiency	91	3.1	17	6.7	4430	0.6	<0.001
Biliary tract disease	47	1.6	3	1.2	2656	0.4	<0.001

SD = standard deviation.

p-Values less than 0.05 are labeled as bold types.

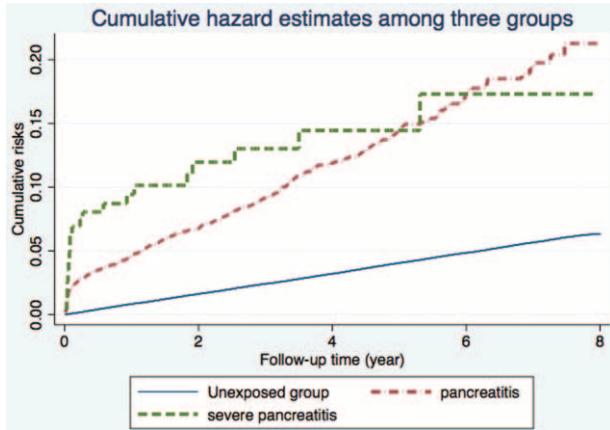


FIGURE 4. Nelson–Aalen curves showing a higher cumulative risk of diabetes in the acute pancreatitis group.

from AP to diabetes was 3.78 years. The incidence rates of diabetes per 100 person-years were 3.1 for patients with AP, 4.5 for patients with severe AP, and 0.8 for patients without AP. The crude HRs of diabetes for patients with AP and patients with severe AP were 3.81 (95% CI, 3.40–4.27) and 5.55 (95% CI, 3.75–8.21), respectively. The Nelson–Aalen plot also revealed higher cumulative risk of diabetes in both AP groups (Figure 4).

The multivariate Cox regression model was then applied to determine the adjusted HRs of diabetes. After controlling for the covariates, an increased HR was still observed for patients with AP (2.15; 95% CI, 1.92–2.41). For patients with severe AP, the HR was also higher but did not differ significantly from that for patients without severe disease (2.22; 95% CI, 1.50–3.29). Other independent risk factors for diabetes included male gender, older age, lower SES, higher CCI, chronic liver disease, hypertension, hypertriglyceridemia, obesity, and a history of

alcohol intoxication (Figure 5). Findings with relevant statistics are summarized in Table 4.

For subjects in the AP group who were diagnosed with diabetes, the diagnosis frequently occurred shortly after hospitalization for AP, with 45.1% of diagnoses occurring within 1 year of an AP episode. Further validation of an increased risk of diabetes to patients with AP was therefore sought. To this end, 324 patients diagnosed with diabetes after an episode of AP were followed for the presence of diabetes for a median period of 1.7 years. Findings revealed that the diagnosis of diabetes was not simply the result of a miscoded transient hyperglycemia; rather, a persistent diabetic state was confirmed for these AP patients.

DISCUSSION

To the knowledge of the authors, the study described in the present report is the first to assess the strength of the bidirectional relationship between diabetes and AP using the same subject population. The database selected for this study was representative of the entire Taiwanese population; therefore, losses to follow-up and selection bias are not concerns.

In patients with diabetes, the risk of AP was significantly increased (HR, 1.72; 95% CI, 1.52–1.96); these findings are compatible with those of other reports.^{6–8} Moreover, diabetics with a history of HCE were found to have an even higher risk (HR, 6.32; 95% CI, 4.54–8.81) as compared to diabetics without HCE. The latter observation is consistent with a “severity-response” relationship between diabetes and risk of AP and represents a novel finding. In this regard, it is of interest that the increased production of reactive oxygen species and increased lipid peroxidation associated with chronic hyperglycemia may be key events in the pathogenesis of AP.^{26,27} Furthermore, diabetes is also associated with comorbidities, such as obesity, hyperlipidemia, or gallstones, with the ability to accelerate the development of AP. Additionally, ryanodine receptor-related alterations in cellular calcium homeostasis may be involved in

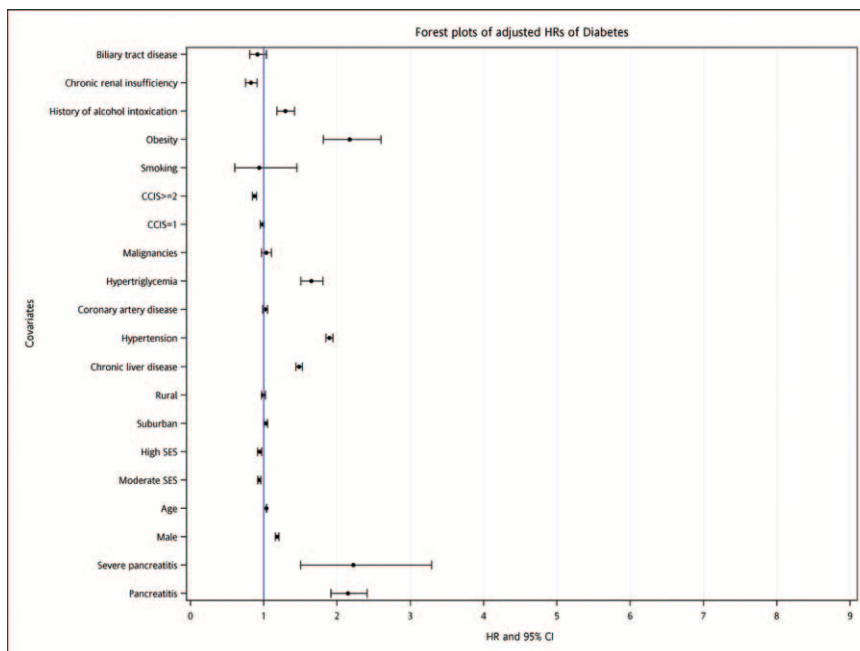


FIGURE 5. Forest plots of adjusted hazard ratios for diabetes.

TABLE 4. Adjusted HRs of Diabetes in Patients With AP

Variables	Hazard Ratio	95% Confidence Interval	P-Value
Pancreatitis	2.15	1.92–2.41	<0.001
Severe pancreatitis	2.22	1.50–3.29	<0.001
Male	1.18	1.16–1.21	<0.001
Patient age	1.04	1.03–1.04	<0.001
Socioeconomic status			<0.001
Low	1	—	—
Moderate	0.94	0.92–0.96	<0.001
High	0.94	0.91–0.97	<0.001
Urbanization level			
Urban	1	—	—
Suburban	1.03	1.00–1.05	0.031
Rural	0.99	0.97–1.02	0.621
Charlson Comorbidity Index score			
0	1	—	—
1	0.98	0.95–1.00	0.085
≥2	0.87	0.85–0.90	<0.001
Chronic liver disease	1.48	1.44–1.53	<0.001
Hypertension	1.89	1.85–1.94	<0.001
Coronary artery disease	1.02	0.99–1.05	0.303
Hypertriglyceridemia	1.65	1.51–1.81	<0.001
Malignancies	1.03	0.97–1.11	0.342
Smoking	0.94	0.60–1.45	0.769
Obesity	2.17	1.81–2.60	<0.001
History of alcohol intoxication	1.29	1.18–1.42	<0.001
Chronic renal insufficiency	0.83	0.75–0.91	<0.001
Biliary tract disease	0.91	0.81–1.04	0.162

p-Values less than 0.05 are labeled as bold types.

the mechanisms through which diabetes facilitates the development of AP. In accord with this proposal, enhanced ryanodine receptor function, which has been reported for several disorders,^{28,29} is observed in both pancreatitis and diabetes.^{30–32} Further studies are required to delineate the biological mechanism(s) responsible for the development of AP in diabetic patients. However, studies that focus on methods to achieve better glycemic control and thereby avoid occurrence of AP in diabetics are particularly warranted.

A higher HR of development of diabetes was found in the present study for patients admitted with AP (HR, 2.15; 95% CI 1.92–2.41). This finding, which is in agreement with those of other studies,^{11,12,14} was confirmed by the observation that the “diabetes” diagnosed in these AP patients was not attributable to a transient hyperglycemia but rather to a persistent disease. In patients with severe as compared to mild AP, a slightly but not significantly increased HR of diabetes was observed (2.22; 95% CI, 1.50–3.29). This finding is compatible with findings from a recent comprehensive systemic review.¹⁴ Although loss of pancreatic β cells due to necrosis is generally considered responsible for development of diabetes in patients with AP,^{33,34} the minimal effect of AP severity on the risk of diabetes as described in the present report is consistent with existence of another mechanism. For example, it is possible that AP itself triggers an event in patients specifically at risk of developing diabetes^{35–37}; however, studies focusing on the mechanism(s) whereby diabetes develops in some patients with AP are needed to confirm this hypothesis.

LIMITATIONS

Five limitations of the present study are acknowledged. First, findings were derived from administrative data. The exposures and outcomes were collected using ICD-9-CM diagnosis codes, and the validities of diagnoses (ie, sensitivity, specificity, and accuracy) cannot be fully assessed. However, the definitions of exposures and outcomes in this study are well-accepted in the administrative database, and previous studies focusing on either AP or diabetes and using similar enrollment criteria from the same database revealed good validities.^{4,23,38,39}

Second, laboratory findings relevant to the severity of diabetes, including HbA1c and blood glucose values, were unavailable in the database. Alternatively, since uncontrolled diabetes is the most common factor precipitating HCE,²⁴ diabetes with poor compliance was defined as diabetes with at least one episode of HCE. Because diabetic ketoacidosis and the hyperosmolar hyperglycemic state are extremely severe states, they may not serve as perfect substitutes for poor control. However, if the definition of diabetes with poor compliance as “diabetes with at least one episode of HCE” failed to differentiate these patients from average diabetic patients, a higher HR would not have been observed.

Third, because of the limitation of administrative database, we believe the diagnostic coding could not perfectly differentiate type I diabetes from type II. As a result, we decide not to perform subgroup analyses regarding the different effects of AP among different types of diabetes. For the same reason, we did not perform analyses regarding the risks of AP on different types of HCE. Individualized studies are better options to solve the study question.

Fourth, the presumed etiology of AP was not considered in the analyses. For example, no effort was made to determine whether either biliary pancreatitis or alcoholic pancreatitis contributed to the higher risk of developing diabetes. Diagnostic coding in the administrative database was considered a limitation to differentiation among possible etiologies of AP. Although the lack of information regarding etiology did not bias the results of this study, it is acknowledged that availability of this information will promote further understanding of the pathophysiology whereby diabetes develops in patients with AP. Future individual-based cohort studies will be needed to address this limitation.

Finally, although extensive adjustments were made for possible comorbidities, unmeasured confounding remains an issue. Based on the nature of the dataset, certain important risk factors, such as body mass index, diet, and life style, could not be taken into account. However, the adjusted HRs were of high enough significance that residual confounding is not likely to fully explain the findings of a bidirectional relationship between diabetes and AP. Furthermore, the “severity-response” effect observed for development of AP in diabetic patients cannot be explained by unmeasured confounding.

CONCLUSIONS

The 2 cohort studies provided evidence for the bidirectional relationship between diabetes and AP. Moreover, diabetic patients with history of HCEs may be associated with higher risk of acute pancreatitis.

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