

Clinic continuity of care, clinical outcomes and direct costs for COPD in Sweden: a population based cohort study

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ABSTRACT

Introduction: In this study we investigate whether clinic level continuity of care (COC) for individuals with chronic obstructive pulmonary disease (COPD) is associated with better health care outcomes and lower costs in a Swedish setting.

Methods: Individuals with COPD (N = 20,187) were identified through ICD-10 codes in all Stockholm County health care registries in 2007–2011 (59% female, 40% in the age group 65–74 years). We followed the individuals prospectively for 365 days after their first outpatient visit in 2012. Individual associations between COC and incidence of any hospitalization or emergency department visit and total costs for health care and pharmaceuticals were quantified by regression analysis, controlling for age, sex, comorbidity and number of visits. Clinic level COC was measured through the Bice–Boxerman COC index, grouped into quintiles.

Results: At baseline, 26% of the individuals had been hospitalized at least once and 73% had dispensed at least seven prescription drugs (23% at least 16) in the last year. Patients in the lowest COC quintile (Q1) had higher probabilities of any hospitalization and any emergency department visit compared to those in Q5 (odds ratio 2.17 [95% CI 1.95–2.43] and 2.06 [1.86–2.28], respectively). Patients in Q1 also on average had 58% [95% CI: 52–64] higher costs.

Conclusion: The findings show robust associations between clinic level COC and outcomes. These results verify the importance of COC, and suggest that clinic level COC is of relevance to both better outcomes for COPD patients and more efficient use of resources.

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KEYWORDS

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Introduction


Continuity of care (COC) is a complex concept. It entails information exchange, disease management and interpersonal relationships, and correlates with a variety of positive health care outcomes.[1–4] For patients with chronic obstructive pulmonary disease (COPD), COC correlates with lower rates of avoidable hospitalization and emergency department visits, fewer complications, lower costs and lower all-cause mortality.[5–7]

In Sweden, COPD is underdiagnosed and undertreated.[8,9] It has been estimated that 8–10% of the adult population have chronic airflow limitation, fulfilling the spirometric criteria of COPD, but less than one in five of those are aware of their diagnosis.[8,10] The mean age at the time of diagnosis is 66 years. A large share of COPD patients suffers from one or several comorbidities, such as hypertension, heart failure, ischemic heart disease, diabetes or depression.[11]

Studies of COC have used different types of data; either patient surveys, which enable specific information gathering, but are expensive and bias prone, or routinely collected data, which are cheap and readily available, but not specifically designed to measure COC.[12,13]

Despite a comprehensive structure of registries in Sweden, studies of COC using routinely collected data are scarce. This is possibly because Swedish register data do not distinguish between individual caregivers, so COC has to be measured in relation to a certain *clinic*, rather than to a certain *clinician*.

A few decades ago, measurement of COC at clinic level was considered suboptimal.[14] This is starting to change, not least in the light of organizational advances such as clinical management protocols, increased emphasis on timely access to care and facilitated information exchange through electronic medical records.[15] Some previous studies have even found clinic-level COC to be more associated with desirable outcomes

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 Supplemental data for this article can be accessed [here](#).

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than clinician-level COC for multi-chronic patients, indicating that a single clinician could not meet the patients' entire health care needs.[16]

There are some further features which potentially make clinic-level COC relevant in Sweden. The structure of Swedish primary health care has typically comprised multidisciplinary teams, and the Swedish guidelines on COPD management emphasize inter professional cooperation. This includes continuity of care with a number of professional categories, such as physicians, nurses and physiotherapists.[17] In this context it is relevant to investigate whether the association between high COC and positive clinical outcomes found in clinician-level COC studies also hold when COC is measured at the clinic level.

In the present study we therefore investigate whether clinic-level COC for individuals with COPD is associated with better health care outcomes and lower costs in a Swedish setting.

Material and methods

Study subjects

We identified the study subjects from a population of 2.1 million inhabitants in Stockholm County, using a comprehensive anonymized individual level database (the VAL database). The VAL database includes complete data on all types of publicly funded public and private care as well as dispensed prescription drugs. It is held by the County Council, the regional body responsible for health care financing and delivery in Stockholm.

We identified 40,381 individuals with at least one main- or secondary diagnosis of COPD (ICD-10 J44, including subcategories) in any type of care (in and outpatient to any caregiver) during 2007–2011. Non-eligible subjects were excluded due to specific exclusion criteria (Figure 1), resulting in a final study population of 20,187 individuals.

Study design

In this population based cohort study, modified after Hussey et al. [6], we followed the identified individuals through registries for 365 days after their first outpatient



Figure 1. Inclusion of individuals in the study population.

visit (any type) in 2012. Clinically relevant baseline characteristics (age, sex, comorbidity (measured as described below), previous hospitalizations and drug use) were identified using information from the registries 365 days prior to this index date (Figure 2). Using data from the follow up period 365 days after the index date, we then quantified associations between clinic level COC and three outcome measures: incidence of any hospitalization or emergency department visit, and total costs for health care and pharmaceuticals. In doing this we controlled for age, sex, number of visits and three different comorbidity measures. We refrained from using disease specific outcomes, for example exacerbations defined by ICD-10 codes, due to incomplete diagnosis coding.

Methods

Continuity of care measurement

COC was measured through the Bice–Boxerman COC index. The index was computed using the standard formula

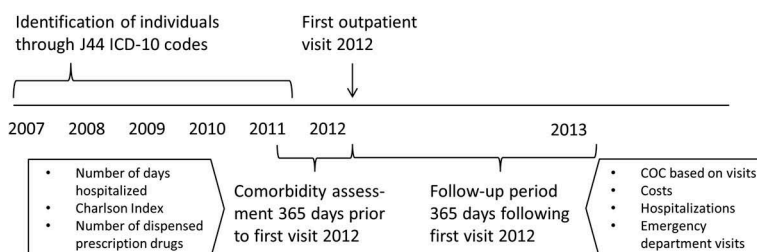


Figure 2. Overview of study design.

($COC = ((\sum(n_j^2)) - n) / (n(n-1))$) from $j = 1$ to $j = s$) as described by Bice and Boxerman, where n is the total number of visits, n_j is the number of visits to caregiver j and s is the total number of caregivers.[18] An index of 0 indicates that the individual had different caregivers at all visits (low continuity) and an index of 1 indicates the same caregiver at all visits (high continuity).

We included all visits, including home visits, during the 365-day follow-up period, regardless of diagnosis or provider type. As the role of nurses has been increasingly highlighted in COPD care in recent years, we included not only doctor visits but visits to all kinds of health care professionals. We excluded emergency department visits (as they were already used as outcome), and multiple visits to the same clinic the same day. The index was calculated for 19,704 individuals, using a total of 526,082 individual/day/clinic combinations. A clinic was defined as a center in primary care and as a specific department in hospital care. Individuals with only one visit were not assigned an index.

We finally grouped the individuals into quintiles based on index values. This was motivated by the non-normal distribution of the index, which complicated the interpretation of actual index values.[12]

Measurement of costs and comorbidity

Costs were calculated based on all outpatient visits, inpatient stays and dispensed prescription drugs for the 365-day follow-up period. For specialized outpatient care and inpatient stays, registry data on diagnosis related group (DRG) costs from the VAL-database was used. DRG is a resource classification system for specialized out-patient and inpatient care. For primary care, home care, emergency care and psychiatry, unit costs were used.¹ Costs were converted to Euros using the average 2012 exchange rate (1 Euro = 8.65 SEK). As the distribution of costs was clearly left skewed, the variable was log transformed prior to regression analysis.

Three comorbidity measures were computed: (1) the number of inpatient days, (2) Charlson index based on all registered diagnoses, and (3) previous drug use, defined as the number of dispensed prescription drugs (unique fifth level, i.e. seven digit or chemical substance, ATC codes). These have previously been found valid indicators

of comorbidity [19–21] and were computed based on data during a period 365 days prior to the first visit 2012.

Statistical analysis

We used regression analysis (SAS software version 9.4 for Windows²) to quantify associations between COC and the three different outcome measures; logistic regression for hospitalizations and emergency department visits, and multivariate linear regression with log of costs for costs. We controlled for age (discretized into four age groups), sex, number of visits (discretized into nine groups) and comorbidity. The grouping of number of visits was based on correlations with costs – for lower visit volumes the relative cost increase per incremental visit is higher than for higher volumes. To capture this nonlinearity, we created increasingly wide-ranging groups (for example 2–4, 5–9 and 55–99 visits).

To calculate estimates and odds ratios (OR) we used the following reference groups: male individuals, 55–64 years of age, with 2–4 outpatient visits, who had the lowest comorbidity for each of the three indicators, and belonged to the 20% of individuals with highest COC.

For the purpose of sensitivity analysis we fitted four additional models with further (mutually exclusive) exclusion criteria. These were: exclusion of (1) individuals >85 years of age, (2) individuals with <5 visits, (3) individuals with >50 visits, and (4) visits with diagnoses indicating respiratory complications (acute upper respiratory infections (ICD-10 J00–J06), influenza and pneumonia (J09–J18), other acute lower respiratory infections (J20–J22), bronchitis (J40–J42), emphysema (J43), COPD exacerbation (J44.0, J44.1), asthma (J45), bronchiectasis (J47), suppurative and necrotic conditions of lower respiratory tract (J85–J86), respiratory failure (J96.0, J96.1, J96.9), cough (R05) or dyspnea (R06.0)).

Results are presented as mean values \pm standard deviations or [95% confidence intervals].

Ethical considerations

The study was based on anonymized administrative data, which cannot be traced back to identifiable individuals. The analysis of this data is classified as continuous health care quality monitoring. The first author is employed by Stockholm County Council and has solely handled the

¹For individuals with at least one primary health care visit, as well as for individuals enrolled to basic home care, a yearly cost (189/73 Euro (age ≥ 65 / < 65) and 393 Euro, respectively) was added to the registered per visit cost. For advanced home care patients, a yearly cost of 811 Euro was added, and per visit costs were estimated to 140 Euro. Visits to local emergency wards were estimated to 116 Euro. Costs for psychiatric care were estimated to 116 Euro (outpatient visit), 231 Euro (emergency department visit) and 578 Euro (inpatient day).

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data processing and analyses in accordance with the county's regulation about patient confidentiality.

Results

Characteristics of the individuals

The study population comprised more women than men, and the most common age group was 65–74 years (Table 1). At baseline, 26% of the individuals had been hospitalized at least once and 73% had dispensed at least seven prescription drugs (23% at least 16) in the last year. During the 365-day follow-up period, almost 70% of the individuals had more than 10 outpatient visits, 40% had at least one emergency department visit and 31% were hospitalized at least once.

Approximately 2% ($N = 483$) of the individuals had <2 visits during the follow-up period, and were hence not assigned a COC index. A total of 1588 individuals (8%) had index values of 1, indicating maximum COC, or all visits to the same clinic. Due to the skewed distribution of the COC-index, the grouping of individuals into quintiles resulted in different width in terms of index values (Figure 3).

Costs were mostly driven by inpatient stays (data not shown), and mean per individual costs were augmented by very high costs for some individuals ($7977 \pm 15,033$ Euro). Mean exponentially back-transformed log of costs was 3647 ± 0.40 Euro per individual, i.e. approaching the median of 3478 Euro.

Median income of the individual's immediate residential area was initially included as proxy for socioeconomic position, but had no explanatory value in the final model and was hence excluded.

Descriptive differences between individuals with different COC

Individuals within the 20% of the highest COC (Q5) were on average older than individuals with lower continuity, but had lower incidence of hospitalizations and emergency department visits (Table 2). There was also a clear association between COC and the number of visits. More individuals in Q5 had few (2–4) or many (50–99, 100+) visits whereas none in Q2 had <5 visits. This is partly explained by the mathematical features of the index.³ The level of comorbidity, measured as number of inpatient days, Charlson index, and previous drug use, was more or less similar in all COC quintiles.

Table 1. Characteristics of the 20,187 individuals.

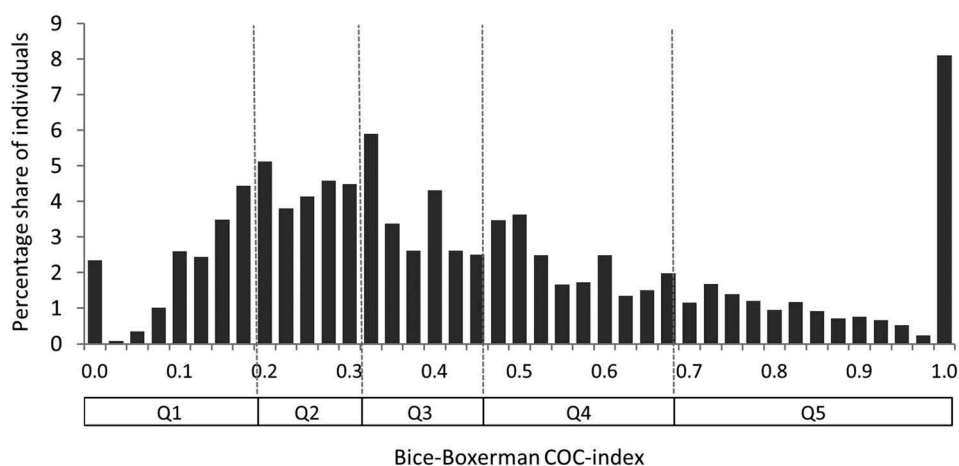
| | N | % |
|---|--------|----|
| Baseline characteristics | | |
| Age group | | |
| 55–64 years | 4665 | 23 |
| 65–74 years | 8030 | 40 |
| 75–84 years | 5634 | 28 |
| 85+ years | 1858 | 9 |
| Sex | | |
| Male | 8341 | 41 |
| Female | 11846 | 59 |
| Comorbidity indicators based on information within 365 days prior to first visit 2012 | | |
| Number of days hospitalized | | |
| 0 | 14,978 | 74 |
| 1–3 | 2008 | 10 |
| 4–7 | 1191 | 6 |
| 8–14 | 835 | 4 |
| 15–29 | 745 | 4 |
| 30–59 | 326 | 2 |
| 60+ | 104 | 1 |
| Charlson index | | |
| 0 | 7042 | 35 |
| 1 | 7779 | 39 |
| 2 | 2827 | 14 |
| 3 | 1447 | 7 |
| 4 | 578 | 3 |
| 5+ | 514 | 3 |
| Number of dispensed prescription drugs | | |
| 0–3 | 2245 | 11 |
| 4–6 | 3319 | 16 |
| 7–10 | 5047 | 25 |
| 11–15 | 4910 | 24 |
| 16–20 | 2675 | 13 |
| 21+ | 1991 | 10 |
| Outcomes measured during the follow-up period 365 days from first outpatient visit 2012 | | |
| Number of outpatient visits | | |
| 2–4 | 2173 | 11 |
| 5–9 | 3988 | 20 |
| 10–14 | 3167 | 16 |
| 15–19 | 2240 | 11 |
| 20–29 | 3071 | 15 |
| 30–49 | 2857 | 14 |
| 50–99 | 1679 | 8 |
| 100+ | 529 | 3 |
| 1 (thus excluded from COC measurement) | 483 | 2 |
| Any emergency department visit | | |
| No | 12,193 | 60 |
| Yes | 7994 | 40 |
| Any hospitalization | | |
| No | 13,987 | 69 |
| Yes | 6200 | 31 |

Age group was defined by age at 31 December 2012. COC = continuity of care. Charlson index was calculated based on all registered diagnoses within 365 days prior to first visit 2012. Number of dispensed prescription drugs = unique fifth level ATC codes.

Associations between COC and hospitalizations, emergency department visits and costs

In the main model (where previous drug use was used as comorbidity indicator), the odds ratio for any hospitalization for individuals with low COC (Q1) as compared to high (Q5) was 2.17 [1.95–2.43], and for any emergency department visit 2.06 [1.86–2.28], when controlled for

³The Bice–Boxerman COC formula does not result in values in the range 0.2–0.32 (the width of Q2) for any combination of clinics if the number of visits is <5. The only possible values are 0.0, 0.17, 0.33, 0.5 and 1.0 (online Appendix A in Supplementary Material).



N=19,704. The width of the horizontal Q1-Q5 bars indicate which values of the Bice-Boxerman continuity of care (COC) index belong to which quintile (Q1-Q5=quintile 1-quintile 5).

Figure 3. Histogram for Bice-Boxerman continuity of care (COC) index, including quintile grouping in horizontal bars.

differences in age, sex, number of visits and comorbidity (Table 3).⁴ Relative increases in total costs amounted to 58%. [52–64] The direction and significance of the results were consistent in the models where we applied further exclusion criteria (data not shown).

We chose the model using previous drug use as comorbidity indicator as the main model for two reasons. First, it had somewhat higher explanatory power for costs ($R^2 = 0.51$) than had number of days hospitalized ($R^2 = 0.48$) and Charlson index ($R^2 = 0.49$), respectively. Second and more importantly, these models provided the most conservative associations between COC and outcomes, as reflected in parameter estimates and odds ratios closer to one (Figure 4).

Discussion

In all models, individuals belonging to lower COC quintiles had higher probabilities of hospitalizations and emergency department visits, and higher costs. These statistically significant associations followed a proportional pattern; individuals belonging to lower quintiles had worse outcomes than individuals in the next, higher, group. The models explained roughly half of the variation in costs between patients. The findings were stable in the main model as well as in four models with further exclusion criteria. In conclusion, associations

between high COC and better outcomes were consistent, relatively strong and followed a logical pattern.

One strength of the study is the full use of a comprehensive dataset of routinely collected data, which enabled individual tracking of patients through all care levels. The inclusion of visits not only to doctors but also to other professions (nurses and physiotherapists, etc.) extended the scope in relation to previous studies, and is in line with modern guidelines for COPD management. [17] Further, the inclusion of all visits, regardless of diagnosis, enabled a more complete picture of the individuals' health care utilization, regardless of whether they were correctly diagnosed at each time point.

Also, in this study we had access to almost complete DRG cost data for hospitalizations. Although we had to approximate unit costs for other types of care, as costs are largely driven by hospitalizations, we can anticipate our cost estimate to be a fairly valid measure of actual total individual costs. In most previous studies, costs are either not included, or calculated only based on prices or billing data (which discloses cost variability).

Another strength of the study is the use of three different comorbidity measures, yielding consistent results for COC. Model consistency reduces the likelihood that associations are driven by some feature of the comorbidity measure as such, and hence increases the reliability of the study. Among the comorbidity measures, we found previous drug use resulting in somewhat

⁴Specifications using the other two comorbidity indicators are found in the online appendices (online Appendix B, C in Supplementary Material). The results were consistent also in models where all three comorbidity indicators were included (not shown).

Table 2. Characteristics of individuals per continuity of care (COC) quintile.

| | COC quintile (Q1 = 20% with lowest COC, Q5 = 20% with highest COC) | | | | | | | | | |
|---|--|----|------|----|------|----|------|----|------|----|
| | Q1 | | Q2 | | Q3 | | Q4 | | Q5 | |
| | N | % | N | % | N | % | N | % | N | % |
| Baseline characteristics | | | | | | | | | | |
| Age group | | | | | | | | | | |
| 55–64 years | 1025 | 25 | 846 | 22 | 917 | 23 | 881 | 22 | 810 | 21 |
| 65–74 years | 1769 | 44 | 1570 | 41 | 1583 | 40 | 1504 | 38 | 1404 | 36 |
| 75–84 years | 1026 | 25 | 1144 | 30 | 1134 | 29 | 1136 | 29 | 1117 | 28 |
| 85+ years | 208 | 5 | 285 | 7 | 311 | 8 | 424 | 11 | 610 | 15 |
| Sex | | | | | | | | | | |
| Male | 1632 | 41 | 1477 | 38 | 1639 | 42 | 1660 | 42 | 1698 | 43 |
| Female | 2396 | 59 | 2368 | 62 | 2306 | 58 | 2285 | 58 | 2243 | 57 |
| Comorbidity indicators based on information within 365 days prior to first visit 2012 | | | | | | | | | | |
| Number of days hospitalized | | | | | | | | | | |
| 0 | 3038 | 75 | 2786 | 72 | 2940 | 75 | 2875 | 73 | 2900 | 74 |
| 1–3 | 445 | 11 | 418 | 11 | 408 | 10 | 364 | 9 | 346 | 9 |
| 4–7 | 232 | 6 | 257 | 7 | 227 | 6 | 233 | 6 | 235 | 6 |
| 8–14 | 154 | 4 | 173 | 4 | 155 | 4 | 185 | 5 | 162 | 4 |
| 15–29 | 115 | 3 | 137 | 4 | 133 | 3 | 183 | 5 | 176 | 4 |
| 30–59 | 38 | 1 | 57 | 1 | 58 | 1 | 89 | 2 | 82 | 2 |
| 60+ | 6 | 0 | 17 | 0 | 24 | 1 | 16 | 0 | 40 | 1 |
| Charlson index | | | | | | | | | | |
| 0 | 1385 | 34 | 1214 | 32 | 1353 | 34 | 1318 | 33 | 1441 | 37 |
| 1 | 1607 | 40 | 1498 | 39 | 1505 | 38 | 1568 | 40 | 1473 | 37 |
| 2 | 552 | 14 | 585 | 15 | 559 | 14 | 541 | 14 | 573 | 15 |
| 3 | 284 | 7 | 309 | 8 | 292 | 7 | 299 | 8 | 259 | 7 |
| 4 | 92 | 2 | 117 | 3 | 135 | 3 | 119 | 3 | 113 | 3 |
| 5+ | 108 | 3 | 122 | 3 | 101 | 3 | 100 | 3 | 82 | 2 |
| Number of dispensed prescription drugs | | | | | | | | | | |
| 0–3 | 445 | 11 | 302 | 8 | 392 | 10 | 402 | 10 | 517 | 13 |
| 4–6 | 674 | 17 | 513 | 13 | 597 | 15 | 631 | 16 | 764 | 19 |
| 7–10 | 1032 | 26 | 898 | 23 | 987 | 25 | 1011 | 26 | 1021 | 26 |
| 11–15 | 906 | 22 | 1010 | 26 | 1040 | 26 | 961 | 24 | 958 | 24 |
| 16–20 | 539 | 13 | 608 | 16 | 528 | 13 | 571 | 14 | 411 | 10 |
| 21+ | 432 | 11 | 514 | 13 | 401 | 10 | 369 | 9 | 270 | 7 |
| Outcomes measured during the follow-up period 365 days from first outpatient visit 2012 | | | | | | | | | | |
| Number of outpatient visits | | | | | | | | | | |
| 2–4 | 610 | 15 | 0 | 0 | 442 | 11 | 282 | 7 | 839 | 21 |
| 5–9 | 925 | 23 | 685 | 18 | 773 | 20 | 916 | 23 | 689 | 17 |
| 10–14 | 728 | 18 | 810 | 21 | 614 | 16 | 608 | 15 | 407 | 10 |
| 15–19 | 507 | 13 | 606 | 16 | 453 | 11 | 385 | 10 | 289 | 7 |
| 20–29 | 640 | 16 | 782 | 20 | 610 | 15 | 556 | 14 | 483 | 12 |
| 30–49 | 418 | 10 | 615 | 16 | 623 | 16 | 624 | 16 | 577 | 15 |
| 50–99 | 186 | 5 | 305 | 8 | 366 | 9 | 412 | 10 | 410 | 10 |
| 100+ | 14 | 0 | 42 | 1 | 64 | 2 | 162 | 4 | 247 | 6 |
| Any emergency department visit | | | | | | | | | | |
| No | 3439 | 85 | 3325 | 86 | 3506 | 89 | 3485 | 88 | 3624 | 92 |
| Yes | 589 | 15 | 520 | 14 | 439 | 11 | 460 | 12 | 317 | 8 |
| Any hospitalization | | | | | | | | | | |
| No | 2691 | 67 | 2506 | 65 | 2698 | 68 | 2708 | 69 | 2952 | 75 |
| Yes | 1337 | 33 | 1339 | 35 | 1247 | 32 | 1237 | 31 | 989 | 25 |

higher explanatory power than other outcomes, which is in accordance with previous studies.[21]

Other authors have suggested that while individuals with higher comorbidity are more likely to have higher costs and worse health care outcomes, they are by definition also more likely to have low continuity as they require care from several different specialists.[7] This potential problem is not supported by our descriptive data, where the level of comorbidity is approximately the same in all COC quintiles.

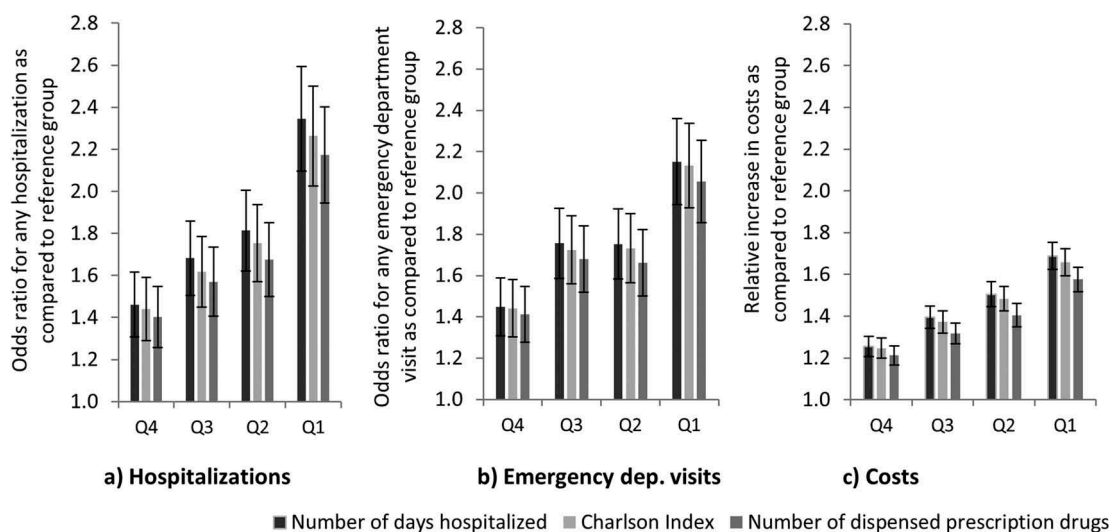
When interpreting the findings it is important to note that the study population only includes those individuals

who remained alive at least 365 days from their first visit in 2012. This inclusion criterion was necessary to make measurement of COC comparable across individuals, but it also has the implication that the results cannot be generalized to the patients with the most severe level of disease.

A limitation of the study is that, although we included several comorbidity measures, we were not able to specifically capture the severity of COPD based on disease specific measures. This was due to data constraints: disease severity is not included in registry data. Potentially, severity could have been indicated through use of specific medications. In a previous study, Make and colleagues [22] showed that

Table 3. Parameter estimates and odds ratios [95% confidence intervals] for the three outcomes; any hospitalization, any emergency department visit and costs during the follow-up period.

| | Odds ratio [95% CI] for any hospitalization as compared to reference group | <i>p</i> | Odds ratio [95% CI] for any emergency department visit as compared to reference group | <i>p</i> | Relative increase [95% CI] in costs as compared to reference group | <i>p</i> |
|--|--|----------|---|----------|--|----------|
| Estimated average for the reference group | 0.07 [0.06–0.08] | | 0.14 [0.12–0.17] | | 606 [574–641] | |
| Age group | | | | | | |
| 55–64 years (ref) | | | | | | |
| 65–74 years | 1.13 [1.04–1.24] | 0.01 | 0.89 [0.82–0.97] | 0.01 | 1.10 [1.06–1.13] | <0.01 |
| 75–84 years | 1.36 [1.23–1.49] | <0.01 | 0.99 [0.91–1.08] | 0.87 | 1.08 [1.04–1.11] | <0.01 |
| 85+ years | 1.88 [1.65–2.13] | <0.01 | 1.38 [1.22–1.55] | <0.01 | 1.11 [1.06–1.16] | <0.01 |
| Sex | | | | | | |
| Male (ref) | | | | | | |
| Female | 0.73 [0.68–0.78] | <0.01 | 0.82 [0.77–0.87] | <0.01 | 0.89 [0.87–0.91] | <0.01 |
| Number of outpatient visits during the follow-up period | | | | | | |
| 2–4 (ref) | | | | | | |
| 5–9 | 1.76 [1.49–2.09] | <0.01 | 1.76 [1.53–2.02] | <0.01 | 1.56 [1.49–1.64] | <0.01 |
| 10–14 | 2.46 [2.07–2.92] | <0.01 | 2.34 [2.03–2.69] | <0.01 | 2.29 [2.18–2.41] | <0.01 |
| 15–19 | 3.61 [3.03–4.30] | <0.01 | 3.40 [2.93–3.94] | <0.01 | 3.02 [2.87–3.19] | <0.01 |
| 20–29 | 4.33 [3.66–5.13] | <0.01 | 4.01 [3.48–4.63] | <0.01 | 3.91 [3.72–4.11] | <0.01 |
| 30–49 | 6.67 [5.63–7.92] | <0.01 | 5.34 [4.61–6.18] | <0.01 | 5.50 [5.23–5.80] | <0.01 |
| 50–99 | 11.43 [9.50–13.75] | <0.01 | 8.51 [7.23–10.03] | <0.01 | 8.82 [8.31–9.35] | <0.01 |
| 100+ | 19.21 [15.02–24.56] | <0.01 | 11.75 [9.33–14.80] | <0.01 | 17.71 [16.27–19.27] | <0.01 |
| Number of dispensed prescription drugs within 365 days prior to first visit 2012 | | | | | | |
| 0–3 (ref) | | | | | | |
| 4–6 | 0.93 [0.80–1.07] | 0.30 | 0.97 [0.85–1.10] | 0.62 | 1.26 [1.20–1.32] | <0.01 |
| 7–10 | 0.98 [0.86–1.12] | 0.76 | 0.97 [0.86–1.09] | 0.61 | 1.57 [1.50–1.64] | <0.01 |
| 11–15 | 1.15 [1.01–1.32] | 0.04 | 1.12 [1.00–1.27] | 0.06 | 1.87 [1.79–1.96] | <0.01 |
| 16–20 | 1.36 [1.17–1.57] | <0.01 | 1.25 [1.10–1.43] | <0.01 | 2.23 [2.12–2.35] | <0.01 |
| 21+ | 1.75 [1.50–2.05] | <0.01 | 1.73 [1.50–2.01] | <0.01 | 2.69 [2.54–2.84] | <0.01 |
| Continuity of care (COC) quintile (Q1 = 20% with lowest COC, Q5 = 20% with highest COC) | | | | | | |
| Q5 (ref) | | | | | | |
| Q4 | 1.40 [1.26–1.56] | <0.01 | 1.41 [1.28–1.56] | <0.01 | 1.21 [1.17–1.26] | <0.01 |
| Q3 | 1.57 [1.41–1.75] | <0.01 | 1.68 [1.52–1.86] | <0.01 | 1.32 [1.27–1.37] | <0.01 |
| Q2 | 1.68 [1.50–1.87] | <0.01 | 1.66 [1.50–1.84] | <0.01 | 1.41 [1.35–1.46] | <0.01 |
| Q1 | 2.17 [1.95–2.43] | <0.01 | 2.06 [1.86–2.28] | <0.01 | 1.58 [1.52–1.64] | <0.01 |



Q4-Q1=continuity of care (COC) quintiles, where 1=lowest COC. Reference group=Q5, the 20% of individuals with highest COC. Adjusted for age, sex, number of visits and one of the three comorbidity indicators: number of days hospitalized, Charlson Index and number of dispensed prescription drugs. All comorbidity indicators measured 365 days prior to follow-up. Error bars = 95% confidence intervals.

Figure 4. Relative effect of belonging to a lower continuity of care (COC) quintile in relation to the highest quintile (Q5) on (a) any hospitalization, (b) any emergency department visit, (c) costs.

COPD maintenance medications are important predictors of exacerbations. Future studies could refine the analysis by stratifying on COPD medication use.

Another limitation is the difficulty to assess the implication of the association between number of visits and COC – with some number of visits even not allowing for all COC levels. As the number of visits is included in the regression models, and the COC quintiles remain significantly associated to outcomes, it is clear that COC harbors an effect that is independent of visit volume, yet it is somewhat difficult to disentangle the two.

A third limitation is that, even though the dose-response-like relationship between COC and outcomes speaks in favor of a causal interpretation, we cannot ascertain that it is COC that *leads* to improved outcomes. This is of importance for the policy implications of the study, and calls for further investigation with other chronic conditions and over several years.

Swedish guidelines on COPD management emphasize inter-professional cooperation, and routinely collected data in Sweden do not distinguish between individual providers. We therefore measured COC at clinic instead of individual clinician level. There are two potential drawbacks to this. First, we cannot ascertain whether the measured associations between higher clinic-level COC and better outcomes is driven at the clinic level as such, or if it is simply an aggregation of a positive effect from high clinician-level COC. Second, although multidisciplinary teams are generally accepted as beneficial for COPD patients, it is not unlikely that some patients experience impaired continuity from visiting different care givers. In effect, although all modern systems for information exchange are in place, from the point of the patient, relational continuity with a particular care provider may be the most important. In order to create a comprehensive understanding of care continuity, a continuity measure such as the one used in this study must therefore be understood in relation to and in combination with the patients' own experiences. [12,23,24]

Conclusions

In conclusion, the findings show robust associations between high clinic-level COC, better clinical outcomes and lower health care costs. These results verify the importance of COC, and suggest that clinic-level COC is of relevance to both better outcomes for patients and more efficient use of resources.

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