## Research article

# Association between sleep duration and serum neurofilament light chain levels among adults in the United States 

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#### Abstract

Background: Neurofilaments are neuron specific skeleton proteins maintaining axon transduction speed, leaked into cerebrospinal fluid and serum after axonal injury or neuron death. Sleep duration change has long related to many health issues but lack laboratory examination. Methods: This study enrolled total 10,175 participants from 2013 to 2014 National Health and Nutrition Examination Survey and used a multi-variable linear model to analyze the relationship between sleep duration and serum neurofilament light chain ( sNfL ) level. Results: There was a fixed relationship between sleep duration and sNfL level ( $\beta=0.65, p=$ 0.0280 ). After adjusted for covariates, this relationship still ( $\beta=0.82, p=0.0052$ ). Segmented regression showed that the turning point of sleep duration was 7 h 1 h decrease in sleep duration was significantly associated with -1.26 higher sNfL level (95 \% CI: $2.25,-0.28 ; p=0.0115$ ) when sleep duration $<7 \mathrm{~h}$; however, 1 h increase in sleep duration was significantly associated with 3.20 higher sNfL level ( 95 \% CI: 2.13, 4.27; $p<0.0001$ ) when sleep duration $>7 \mathrm{~h}$. Furthermore, the stratified analysis indicated that the associations between sleep duration and sNfL level were stronger among those normal body mass index and trouble sleeping ( $p$-interaction $<0.0001$ and 0.0003). Conclusion: In summary, there was a J-shaped relationship between sleep duration and sNfL level in the United States of America representative group, these may suggest that extreme sleep duration can be deleterious judged by sNfL level. And still need large cohort study to determine the accurate relationship, and cluster analysis to infer the nervous disease connected with extreme sleep duration.


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## 1. Background

Neurofilaments (NFs) are neuron specific intermediate filaments composed by three subunits with apparent molecular weights of $68,000 \mathrm{Da}(\mathrm{NfL}), 145,000 \mathrm{Da}(\mathrm{NF}-\mathrm{M})$, and $200,000 \mathrm{Da}(\mathrm{NF}-\mathrm{H})$ (determined by SDS-gel electrophoresis) [1-3], one of the components of the neuron cytoskeleton, confirmed help maintaining axon structure, axon and dendritic branching and growth and axonal transport speed [2-5], by assembled into 10 nm filamentous structures [2,3,6]. Neurofilament light chain ( NfL ) is a member of filament proteins family [7,8], most abundant and most soluble subunit [3], widely thought could be a potential biomarker in neuro-damage events, cause robust and early levitated in both cerebrospinal fluid (CSF) and serum after neuronal cell damage such as axonal injury and axonal death [4,8-11], besides, the serum (or plasma) NfL level is correlated with CSF NfL level, implies that two measurements reflect similar physiological process [12,13], this elevation effect always before obvious neurological disease onset, and a higher level of serum NfL generally comes with worse outcome [4,9,10,12,14].

In 2005 Abdo and coworkers found out that elevated CSF NfL level can be used to differentia multiple system atrophy (MSA) from Parkinson's disease [15]; then it was confirmed a promising biomarker in nervous diseases like sporadic Creutzfeldt-Jakob disease (sCJD) [16]; amyotrophic lateral sclerosis (ALS) [17]; 2016, Bacioglu and his colleges claimed that both serum and CSF NfL level can be used in detecting neurodegenerative disease [4]; next year, Fyfe alleged that high plasma NfL level is associated with cognitive decline in Alzheimer disease (AD) [18]. The level of NfL is not only related to the outcome, but also to the responses to treatment [6,12, 19,20]. Beyond these highly sensitivity, third-generation (electrochemiluminescence) and fourth generation (single molecule array) assay also enabled reliable serum NfL measurement make it easier to determine NfL level $[6,12]$.

Sleep is an essential step for optimal health and under genetic regulation [21-23], The American Academy of Sleep Medicine and the Sleep Research Society determined that adults require $\geq 7 \mathrm{~h}$ sleep per day [24]. Healthy sleep requires adequate duration, quality, appropriate timing and regularity [25,26], sleep loss confirmed slowing information processing speed and other neurocognitive performances [27,28], short sleep duration has long been considered as risk of behavioral or problems such as anxiety, aggressive behavior, daytime sleepiness and lowered cognitive function [28,29], even some disease states like cardiovascular disease, obesity, depression, type-2 diabetes, cancer even suicidal ideation [23,25,27,30-35]. Liu confirmed that lack of non-rapid eye movement sleep increased serum and CSF NfL level in C57BL/6J mice [36], but the detailed relationship between neuro-damage and sleep duration still need further investigation. In this study we enrolled 2013-2014 residents from National Health and Nutrition Examination Survey


Fig. 1. Flow chart of study design and sample disposition.
(NHANES), analyzed the connection between serum NfL level and sleep time duration, tried to figure out the detailed neuro-damage progression with differential sleep time.

## 2. Methods

### 2.1. Study design and population

This study enrolled 2013-2014 respondents from NHANES (which also is where can trace the original data), a complex, multi-stage representative national survey leaded by National Center for Health Statistics (NCHS), reports mainly about the health and nutrition characteristics of non-institutionalized populations lived in the United States. Participants not only underwent questionnaires assessing demographic, socioeconomic, nutritional, health, and other domains during face-to-face interviews in their homes, but also underwent physical examinations in medical facilities to collect medical and physiological data; additional laboratory tests were performed on blood and urine samples collected in the field, and surplus or pristine serum samples were stored for future studies, were eligible. All respondents who participated in the study provided informed consent. The NCHS Ethics Review Board (ERB) approved the ethical review of the project (www.cdc.gov/nchs/nhanes/irba98.htm). This study followed the Strengthening the Reporting of Observational Studies in Epidemiology (STROBE) reporting guidelines, thus the IRB/ethics committee approval is not applicable in this submission.

The downloaded data was analyzed visually using the statistical package R (R4.2.0). In this study, a total of 10,175 participants were investigated. Serum neurofilament light chain ( sNfL ) analysis measurements were received in surplus or pristine serum samples ( $\mathrm{n}=2071$, ages $25-75$ years). After excluding missing sleep data ( $\mathrm{n}=2$ ), baseline self-reported Stroke ( $\mathrm{n}=51$ ), and organic brain lesions diagnosed by ICD-10 codes [37]:Transient cerebral ischemic attack (G45.9), Epilepsy and recurrent seizures (G40), Alzheimer's disease (G30.9), Multiple sclerosis (G35), Mild cognitive impairment (G31.84), Parkinson's disease (G20), Degenerative disease of nervous system (G31.9), Spastic hemiplegia (G81.1), Trigeminal neuralgia (G50.0)) ( $\mathrm{n}=25$ ) and Pregnancy ( $\mathrm{n}=18$ ), a total of 1975 participants were included in the final analysis (Fig. 1).

### 2.2. Serum neurofilament light chain

Detailed specimen collection and handling instructions are discussed in NHANES Laboratory Methods (wwwn.cdc.gov/Nchs/ Nhanes/2013-2014/SSSNFL_H.htm). The method quantifies sNfL by a highly sensitive immunoassay. Using acridinium ester (AE) chemiluminescence and paramagnetic particles and may be run on an existing, high-throughput, automated platform (Attelica). Initially, the sample is incubated with AE labeled antibodies, which bind to the NfL antigen. Following this step, paramagnetic particles (PMP) coated with capture antibody are added to the sample, forming complexes of antigen bound to AE-labeled antibodies and PMP. Unbound AE-labeled antibodies are then separated and removed, following which acid and base are added to initiate chemiluminescence and light emission is measured. All steps are performed on the fully automated Attelica immunoassay system.

These assays were developed by Siemens Healthineers and use Quanterix proprietary NfL antibody for the quantitative determination of NfL concentrations in samples. Analytical measurements are performed in accordance with strict quality control/quality assurance procedures. Researchers calculated the coefficient of variation (CV) and other relevant statistics to describe the quality control (QC) samples across the spectrum of sNfL measures. The lower limit of quantification (LLOQ) was defined as the concentration at which the CV was less than or equal to $20 \%$. The assays range from 3.9 to $500 \mathrm{pg} / \mathrm{ml}$. In addition to study samples, low, medium, and high concentration QC samples were run each 8-h shift as well as additional replicate samples to ensure accuracy and reliability of the derived data.

### 2.3. Sleep duration

Baseline sleep duration is self-reported as "How much sleep do you get (hours)?" Responses are coded as integers. There were no given categories in the face-to-face interviews. Participants were divided into 6 groups for analysis based on sleep duration $(<6,6,7,8$, 9 , or $\geq 10 \mathrm{~h}$ per night).

### 2.4. Assessment of covariates

Sex, age, race/ethnicity (non-Hispanic white, Black, Mexican American, or other race), poverty level ( $<2.0, \geq 2.0$ ), Education level, waist circumference and body mass index (BMI), measured objectively as part of the NHANES physical examination. The NHANES interview also assessed the following: Diabetes defined as self-reported physician diagnosis of diabetes or HbA1c $\geq 6.5 \%$. Hypertension was defined as systolic blood pressure (SBP) $\geq 140 \mathrm{mmHg}$ and diastolic blood pressure (DBP) $\geq 90 \mathrm{mmHg}$, or taking medication to control blood pressure. In addition, participants' renal function was assessed by measuring eGFR using the executive summary of the KDIGO 2021 Clinical Practice Guideline [38]. The urinary albumin/creatinine ratio (ACR) was also calculated. According to the ex-tended definition of chronic kidney disease (CKD), participants with ACR $\geq 30 \mathrm{mg} / \mathrm{g}$ or eGFR $\leq 60 \mathrm{ml} / \mathrm{min} / 1.73 \mathrm{~m}^{2}$ had CKD. According to the NCHS classification, individuals are classified as never smokers, former smokers, and current smokers. Drinking status was categorized as never (had $<12$ drinks in life time), former (had $\geq 12$ drinks in 1 year and did not drink last year, or did not drink last year but drank $\geq 12$ drinks in lifetime), mild drinker ( $\leq 1$ drink per day for women or $\leq 2$ drinks per day for men on average over the past year), moderate drinker ( $\leq 3$ drink per day for women or $\leq 4$ drinks per day for men on aver-age over the past year), or heavier
drinker ( $\geq 4$ drink per day for women or $\geq 5$ drinks per day for men on average over the past year [39]. These covariates were chosen because they were previously used to study the relationship between sleep duration.

### 2.5. Statistical analyses

Data were statistically analyzed according to CDC guidelines [40], and all models also incorporated the minimization weights. The weighted chi-square test was carried out for the classified variables, and the $P$ value of the continuous variables was calculated with the weighted linear regression model. Data are expressed as numbers of percentages of categorical variables and as means of standard deviations of continuous variables. Three multivariable regression analysis were developed to investigate the relationship between sNFL levels and sleep duration. Crude Model (unadjusted); Model I adjusted for Sex, Age, and Race/Ethnicity; Model II adjusted for Sex, Age, Race/Ethnicity, Education level, BMI, Smoking status, Drinking status, Hypertension, Diabetes, CKD, and Self-reported sleep quality. The subgroup analysis was performed by stratified multivariate regression analysis. Generalized additive model and curve fitting (Restricted cubic spline, RCS) were generalized to explore the potential nonlinear correlation, six different plot were generated in which RCS package default set.

We first used smooth curve fitting to examine whether the independent variable is partitioned into intervals. We applied segmented regression (also known as piece-wise regression) that used a separate line segment to fit each interval. Log-likelihood ratio test comparing one-line (non-segmented) model to segmented regression model was used to determine whether threshold exists. The inflection point that connecting the segments was based on the model gives maximum likelihood, and it was determined using two steps recursive method.

The step 1 was to narrow down the inflection point to a 10-percentile range of the independent variable. From $5 \%$ to $95 \%$ incremented by $5 \%$, we tested 19 segmented regression models using these 19 percentile points of independent variable as the inflection point respectively to find out which percentile points gives the model with highest likelihood. The precise inflection point was narrowed down to $\pm 4 \%$ percentile of the percentile points which gives highest likelihood among the 19 models, called $\mathrm{K}_{\min }$ and $\mathrm{K}_{\max }$ respectively. The step 2 was to determine precise inflection point between $K_{\min }$ and $K_{\max }$ using the recursive method. The specific method was to first run 3 models with inflection point equals Q1 ( $25 \%$ percentile), Q2 ( $50 \%$ percentile) and Q3 (75 \% percentile)

Table 1
Baseline Characteristics of participants.

| Characteristic | Neurofilament light chain (pg/ml) |  |  |  |  | $P$-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | $\begin{aligned} & \text { Q1 } \\ & <8.10 \end{aligned}$ | $\begin{aligned} & \text { Q2 } \\ & \geq 8.10 \text { to }<12.10 \end{aligned}$ | $\begin{aligned} & \text { Q3 } \\ & \geq 12.10 \text { to }<18.70 \end{aligned}$ | $\begin{aligned} & \text { Q4 } \\ & \geq 18.70 \end{aligned}$ |  |
| Age (years) | $44.81 \pm 15.05$ | $34.02 \pm 10.01$ | $42.67 \pm 13.12$ | $49.13 \pm 14.41$ | $54.35 \pm 14.09$ | <0.0001 |
| Sex (\%) |  |  |  |  |  | 0.0049 |
| Men | 49.49 | 43.70 | 49.64 | 50.04 | 55.07 |  |
| Woman | 50.51 | 56.30 | 50.36 | 49.96 | 44.93 |  |
| Race/ethnicity (\%) |  |  |  |  |  | <0.0001 |
| Mexican American | 9.71 | 16.18 | 9.21 | 5.96 | 7.06 |  |
| Non-Hispanic White | 64.66 | 53.60 | 62.23 | 71.59 | 72.13 |  |
| Non-Hispanic Black | 11.92 | 13.94 | 14.36 | 7.90 | 11.28 |  |
| Other race/ethnicity | 13.71 | 16.28 | 14.20 | 14.55 | 9.53 |  |
| Education (\%) |  |  |  |  |  | 0.2927 |
| Less than high school | 15.37 | 18.21 | 13.85 | 14.86 | 14.39 |  |
| High school or equivalent | 20.30 | 18.27 | 23.06 | 19.38 | 20.56 |  |
| College or above | 64.33 | 63.52 | 63.09 | 65.76 | 65.04 |  |
| Poverty Income Ratio (\%) |  |  |  |  |  | 0.2134 |
| $<2$ | 36.88 | 40.85 | 35.46 | 35.38 | 35.60 |  |
| $\geq 2$ | 63.12 | 59.15 | 64.54 | 64.62 | 64.40 |  |
| Smoke (\%) |  |  |  |  |  | 0.0003 |
| Now | 20.85 | 20.27 | 19.25 | 22.28 | 21.72 |  |
| Former | 22.35 | 16.28 | 21.25 | 25.82 | 26.57 |  |
| Never | 56.79 | 63.45 | 59.51 | 51.90 | 51.71 |  |
| Alcohol user (\%) |  |  |  |  |  | 0.0001 |
| Heavy | 22.47 | 26.40 | 24.32 | 18.78 | 20.20 |  |
| Mild | 34.30 | 30.48 | 36.70 | 35.44 | 34.69 |  |
| Moderate | 19.79 | 21.99 | 20.02 | 20.20 | 16.86 |  |
| Former | 12.11 | 7.72 | 9.71 | 14.12 | 17.12 |  |
| Never | 11.32 | 13.41 | 9.24 | 11.46 | 11.12 |  |
| BMI ( $\mathrm{kg} / \mathrm{m}^{2}$ ) | $29.41 \pm 7.38$ | $30.01 \pm 7.45$ | $28.83 \pm 7.29$ | $28.76 \pm 7.04$ | $30.04 \pm 7.62$ | 0.0036 |
| Waist circumference (cm) | $99.69 \pm 16.96$ | $99.44 \pm 17.13$ | $97.33 \pm 16.64$ | $99.13 \pm 16.97$ | $103.03 \pm 16.55$ | <0.0001 |
| Hypertension (\%) | 36.36 | 18.29 | 33.46 | 40.84 | 54.48 | <0.0001 |
| Diabetes (\%) | 13.86 | 5.97 | 9.11 | 15.64 | 25.60 | <0.0001 |
| CKD (\%) | 8.98 | 8.98 | 8.56 | 10.11 | 21.51 | <0.0001 |
| Trouble sleeping (\%) | 12.14 | 20.55 | 27.36 | 25.65 | 31.24 | 0.0017 |
| Sleep Durations (hours) | $6.89 \pm 1.32$ | $6.95 \pm 1.30$ | $6.74 \pm 1.25$ | $6.97 \pm 1.25$ | $6.90 \pm 1.48$ | 0.0331 |

Data are presented as mean $\pm$ SD or $n(\%)$.
within the range of $\mathrm{K}_{\min }$ and $\mathrm{K}_{\max }$ respectively to find out which quartile point gives the model with highest likelihood among the three models. Then we narrowed down the $K_{\min }$ and $K_{\max }$ to the range of $\pm 25 \%$ of the corresponding quartile point. By doing so, we narrowed the range of $\mathrm{K}_{\min }$ and $\mathrm{K}_{\max } 50 \%$ recursively each time until the specific value of the independent variable was identified, that if used as inflection point will give the segmented regression model highest likelihood.

Furthermore, the inflection point is calculated by using a two-stage linear regression model. All analyses were performed with package R (http://www.R-project.org) and EmpowerStats (http://www.empowerstats.com). $p<0.05$ was of statistical significance.

## 3. Results

### 3.1. Baseline characteristics of study participants

In this study, the analyzed data set consisted of data from 1975 participants without neurological disorders (mean age: $44.81 \pm$ 15.05 years; 49.49 \% males). The baseline characteristics of the study population according to sNFL levels are reported in Table 1. Participants with higher sNFL levels were more likely to be older, male, non-Hispanic white, and more likely to have poor lifestyle habits like smoking and drinking, worse health status (obesity, hypertension, diabetes, chronic kidney disease), and poor sleep quality ( $p<0.01$ ).

## 4. Sleep duration associated with sNfL

We designed three multivariable regression analysis to investigate the relationship be-tween sNfL and sleep duration. The results of the multivariate regression analyses are presented in Table 2. In the crude model, the shortest ( $<6 \mathrm{~h}$ ) and longest ( $>10 \mathrm{~h}$ ) sleep durations were associated with higher sNfL, with the highest sNfL levels in those who slept more than 10 h per night. And this pattern was maintained in Model II after correction for adjustments including age, gender, and race/ethnicity confounders. After multivariate adjustment in Model II, including further adjustment for education lev-el, BMI, smoking status, alcohol consumption status, sleep condition, disease condition, multivariate adjusted $\beta$ and $95 \%$ confidence intervals (CIs) from lowest to highest sleep duration category ( $<6,6,7,8,9$, or $\geq 10 \mathrm{~h}$ ) with sNfL of 0.00 (reference), $-2.68(-5.32,-0.04),-2.65(-5.26,-0.05),-2.79(-5.45,-0.13)$,

Table 2
Regression coefficients $(\beta)$ and $95 \%$ confidence interval for sNfL Levels across sleep duration categories.

| Outcome | Crude Model |  | Model I |  | Model II |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\beta$ (95\%CI) | P-value | $\beta$ (95\%CI) | P-value | $\beta$ (95\%CI) | P-value |
| Complete sample |  |  |  |  |  |  |
| Sleep duration | 0.65 (0.07, 1.24) | 0.0280 | 0.48 (-0.08, 1.03) | 0.0946 | 0.82 (0.25, 1.40) | 0.0052 |
| Sleep duration categories |  |  |  |  |  |  |
| $<6 \mathrm{~h}$ | Reference |  | Reference |  | Reference |  |
| 6 h | -3.24 (-5.93, -0.54) | 0.0216 | -3.19 (-5.75, -0.63) | 0.0148 | -2.68 (-5.32, -0.04) | 0.0464 |
| 7 h | -3.03 (-5.61, -0.45) | 0.0104 | -4.03 (-6.50, -1.57) | 0.0014 | -2.65 (-5.26, -0.05) | 0.0459 |
| 8 h | -3.50 (-6.18, -0.82) | 0.2608 | -4.06 (-6.61, -1.51) | 0.0018 | -2.79 (-5.45, -0.13) | 0.0398 |
| 9 h | 2.33 (-1.73, 6.39) | <0.0001 | 1.31 (-2.56, 5.17) | 0.5075 | 2.38 (-1.56, 6.32) | 0.2366 |
| $\geq 10 \mathrm{~h}$ | 11.60 (5.95, 17.25) | 0.0216 | 11.23 (5.86, 16.60) | <0.0001 | 12.31 (6.54, 18.07) | <0.0001 |
| P for trend | 0.110 |  | 0.319 |  | 0.125 |  |
| Men |  |  |  |  |  |  |
| Sleep duration | 1.02 (0.04, 2.00) | 0.0419 | 0.80 (-0.16, 1.75) | 0.1018 | 1.28 (0.32, 2.25) | 0.0094 |
| Sleep duration categories |  |  |  |  |  |  |
| $<6 \mathrm{~h}$ | Reference |  | Reference |  | Reference |  |
| 6 h | -4.06 (-8.69, 0.57) | 0.0861 | -4.33 (-8.82, 0.16) | 0.0589 | -2.73 (-7.28, 1.81) | 0.2392 |
| 7 h | -4.41 (-8.78, -0.05) | 0.0478 | -5.73 (-9.99, -1.47) | 0.0085 | -3.28 (-7.68, 1.12) | 0.1441 |
| 8 h | -5.13 (-9.73, -0.52) | 0.0295 | -5.88 (-10.36,-1.40) | 0.0103 | -3.40 (-7.95, 1.15) | 0.1433 |
| 9 h | 3.34 (-4.11, 10.80) | 0.3794 | 0.37 (-6.87, 7.60) | 0.9210 | 3.80 (-3.44, 11.05) | 0.3040 |
| $\geq 10 \mathrm{~h}$ | 16.27 (7.33, 25.21) | 0.0004 | 16.67 (8.00, 25.33) | 0.0002 | 17.32 (7.97, 26.66) | 0.0003 |
| P for trend | 0.194 |  | 0.389 |  | 0.109 |  |
| Woman |  |  |  |  |  |  |
| Sleep duration | 0.28 (-0.35, 0.90) | 0.3863 | 0.10 (-0.47, 0.68) | 0.7256 | 0.17 (-0.46, 0.80) | 0.5957 |
| Sleep duration categories |  |  |  |  |  |  |
| $<6 \mathrm{~h}$ | Reference |  | Reference |  | Reference |  |
| 6 h | -2.30 (-5.12, 0.53) | 0.1114 | -1.95 (-4.54, 0.64) | 0.1401 | -2.72 (-5.49, 0.05) | 0.0542 |
| 7 h | -1.75 (-4.50, 1.00) | 0.2126 | -2.25 (-4.78, 0.27) | 0.0806 | -2.30 (-5.09, 0.49) | 0.1062 |
| 8 h | -1.80 (-4.60, 1.00) | 0.2074 | -2.35 (-4.92, 0.21) | 0.0728 | -2.54 (-5.35, 0.27) | 0.0771 |
| 9 h | 2.32 (-1.72, 6.36) | 0.2602 | 2.19 (-1.52, 5.90) | 0.2481 | 1.45 (-2.51, 5.41) | 0.4743 |
| $\geq 10 \mathrm{~h}$ | 3.61 (-2.94, 10.15) | 0.2803 | 2.80 (-3.19, 8.80) | 0.3596 | 3.55 (-2.96, 10.07) | 0.2854 |
| P for trend | 0.320 |  | 0.687 |  | 0.501 |  |

Crude Model: no covariates were adjusted.
Model I: age, sex, and race/ethnicity were adjusted.
Model II: age, sex, race/ethnicity, education level, BMI, smoking status, alcohol consumption status, trouble sleeping, history of hypertension, diabetes and chronic kidney disease; In the subgroup analysis stratified by sex, the model is not adjusted for sex.
$2.38(-1.56,6.32)$, and $12.31(6.54,18.07)$, ( $p$ for trend $=0.125$ ). This relationship persisted. We also found that when considering men separately, for those with $<6 \mathrm{~h}$ of sleep, elevated sNFL was observed only in crude model and model I; long sleep duration was significantly associated with elevated sNFL in all 3 models. For women, however, no significant difference in sNFL with sleep duration was observed in all models.

### 4.1. The detection of nonlinear relationships

Weighted restricted cubic splines were used to visually assess nonlinear relationships. We found a or J-shaped association between sNFL and sleep duration (Fig. 2 Total). Then, we used a recursive algorithm to calculated the inflection point between sNFL and sleep duration and performed two linear regression models on both sides of the inflection point to investigate the analysis confirming the non-linear relationship association between sNFL and sleep duration. (Log-likelihood p $<0.05$ ) (Table 3). We found that the inflection point of sNFL in the fully adjusted model was 7 h per night. When sleep duration was less than 7 h per night, each 1-h decrease in sleep duration was associated with an adjusted increase in sNFL of $1.26(\mathrm{pg} / \mathrm{ml})(\beta=-1.26 ; 95 \% \mathrm{CI}: 2.24,-0.28)$. More than 7 h of sleep per night and each additional hour of sleep was associated with a $3.20(\mathrm{pg} / \mathrm{ml})$ adjusted increase in $\mathrm{sNFL}(\beta=3.20 ; 95 \% \mathrm{CI}: 2.13,4.27)$.


Fig. 2. Relationship between sleep duration and serum NfL level. A: Relationship between sleep duration and serum NfL level in all groups. B: Relationship between sleep duration and serum NfL level in different ages. C: Relationship between sleep duration and serum NfL level in different sexes. D: Relationship between sleep duration and serum NfL level in people with different BMIs. E: Relationship between sleep duration and serum NfL level in if suffered trouble sleeping. F: Relationship between sleep duration and serum NfL level in if suffered CKD. G: Relationship between sleep duration and serum NfL level in if suffered diabetes. H: Relationship between sleep duration and serum NfL level in if suffered hypertensions.

Table 3
Threshold Effect Analysis of sleep duration and sNfL Levels using Piece-wise Linear Regression.

|  | $\beta(95 \% \mathrm{CI}) P$-value |
| :--- | :--- |
| Crude Model <br> Fitting by the standard linear model <br> Fitting by the two-piecewise linear model |  |
| Inflection point | $0.65(0.07,1.24) 0.0280$ |
| $\quad$ sleep duration $<8 \mathrm{~h}$ |  |
| sleep duration $>8 \mathrm{~h}$ | 8 |
| Log likelihood ratio | $-0.87(-1.56,-0.18) 0.0141$ |
| Model I | $7.29(5.53,9.06)<0.0001$ |
| Fitting by the standard linear model | $<0.001$ |
| Fitting by the two-piecewise linear model |  |
| Inflection point | $0.48(-0.08,1.03) 0.0946$ |
| $\quad$ sleep duration $<8 \mathrm{~h}$ | 8 |
| $\quad$ sleep duration $>8 \mathrm{~h}$ | $-1.05(-1.71,-0.39) 0.0017$ |
| Log likelihood ratio | $7.12(5.44,8.80)<0.0001$ |
| Model II | $<0.001$ |
| Fitting by the standard linear model | $0.82(0.25,1.40) 0.0052$ |
| Fitting by the two-piecewise linear model | 7 |
| Inflection point | $-1.26(-2.24,-0.28) 0.0115$ |
| sleep duration $<7 \mathrm{~h}$ | $3.20(2.13,4.27)<0.0001$ |
| sleep duration $>7 \mathrm{~h}$ | $<0.001$ |
| Log likelihood ratio |  |

Crude Model: no covariates were adjusted.
Model I: age, sex, and race/ethnicity were adjusted.
Model II: age, sex, race/ethnicity, education level, BMI, smoking status, alcohol consumption status, trouble sleeping, history of hypertension, diabetes and chronic kidney disease.

### 4.2. Stratified analyses

As shown in Fig. 3, the subgroup analysis revealed a highly consistent pattern. In the subgroup analysis, we further explored the role of other covariates on the association between sleep duration and sNFL. The positive correlation between sleep duration and sNFL was more significant at normal body mass index ( $\beta=2.24,95 \% \mathrm{CI}$ : $1.40,3.07$; P-interaction $<0.0001$ ). Subjects who reported trouble sleeping had higher sNFL levels than those who were not ( $\beta=2.32,95 \% \mathrm{CI}$ : $1.33,3.30 p$ for interaction $=0.0003$ ). However, the association between sleep duration and sNFL was consistent across the following subgroups: sex, age, race, hypertension, diabetes, and chronic kidney disease ( $p$ for covariates interaction $>0.05$ ).

We further used multivariate regression models to investigate the relationship between sNfL levels and extreme sleep durations in different subgroups (Tables 4 and 5). After controlling for other variables (Table 4), remarkable elevated sNFL was observed in $>10-\mathrm{h}$ sleepers at normal body mass index ( $\beta=22.01$; 95\%CI15.34, 28.68). Similarly, after controlling for other variables (Table 5), sNfL levels were significantly highest in the group who self-reported trouble sleeping ( $\beta=41.78$; 95\%CI: 29.34, 54.22). However, this trend was not present in those who not self-reported trouble sleeping ( $\beta=1.48,95 \% \mathrm{CI}:-5.00,7.95$ ).

## 5. Discussion

During the past century human sleep duration was significantly reduced, 9 h in 1910 to 7.5 h in 1975 and 6.8 h in2005 [22,41], which raised lots of diseases incidence. Sleep restriction confirmed reduce plasma leptin and ghrelin concentration [35,42]and other negative effects such as anxiety [24], depression [43], obesity [42], diabetes [44] even cancer [23,33]. The relationship between extreme sleep duration and neurological dysfunction has long been identified [24,43,45], however, the detail measurement of neurological function or structure impairment linked with extreme sleep duration left blank remain.

Hurst and his colleges managed to locate human neurofilament gene (NEFL) at short arm of chromosome 8 in 1987 [1], which encoded a 543 amino acids primary peptide chain. Neurofilaments light chain is a member of neurofilaments family, widely expressed in central and peripheral nervous system [46], are structural scaffolding proteins in neurons and can cross-bridging and interconnecting with other components of cytoskeleton [46,47], were synthesized in cell plasma then transported by slow axonal transport into distal axon by $0.1-0.3 \mathrm{~mm}$ per day [2]. Usually released into CSF and serum after axonal damage or neuron death, folds over normal level [47,48], have long been identified as a promising neuro-disease biomarker [2,4,15,16,49].

In our study, we found out that there was a J shape relationship between USA adult sleep duration and serum neurofilaments level. This $J$ shape relationship between sleep and other health state has long been discovered [23,25,31,34,42,50-53], however, there still lack nervous system data support this conception. Our effort in this study may further explain connection of extreme sleep duration and neurological dysfunction. Interestingly, our results showed that overlong sleep duration in male caused higher level sNfL than female might out of the poorer sleep in male [54]. Also, our data showed that people with normal BMI may get worsen neuro-stage, might traced back that obesity is chronical inflammatory state caused elevated sNfL level same as diabetes group [55,56].

Though in this article mainly about the relationship between sleep duration and NfL, the sleep quality has long been an important


Fig. 3. Forest plots of stratified analyses of sleep duration and serum NfL level.
part [26,29], normal hormone secretion adhered a time related circadian rhythms [24,35,57], however due to the origin questionnaire design, this study lack of these parts that shall be long concerned in the following randomized designed study. In these collected data, only 1 case reported apnea-hypopnea event, which is obvious couldn't match other database [34], so this index was removed in this study.

What should be noticed is that in some brain-blood barrier (BBB) dysfunction diseases CSF NfL were also reported elevated [3], we couldn't straight jump to the conclusion that elevated serum NfL level is definitely connected with potential neuron damage, detailed study still needs further explore.

There has long been considered that sleep duration has linked to many neurodegeneration diseases [29,58,59]. Apart from some extreme status such as ABCC9 (ATP-binding cassette, sub-family C, member9) mutations [60], our result showed that sleep duration do impair neurons, but the precise pathway and link still need further exploration. Qualification in the scales always led to recall bias, also may led us to a biased result, which is only can be partially narrowed by detailed curve optimization, which is undeniable. And samples

Table 4
Regression coefficients ( $\beta$ ) and 95 \% confidence interval for sNfL Levels across sleep duration categories.

| Outcome | Crude Model |  | Model I |  | Model II |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\beta$ (95\%CI) | P-value | $\beta$ (95\%CI) | P-value | $\beta$ (95\%CI) | P-value |
| BMI < 28 ( $\mathrm{kg} / \mathrm{m}^{2}$ ) |  |  |  |  |  |  |
| Sleep duration | 2.24 (1.51, 2.97) | <0.0001 | 1.95 (1.25, 2.65) | $<0.0001$ | 2.29 (1.55, 3.04) | <0.0001 |
| Sleep duration categories |  |  |  |  |  |  |
| < 6 h | Reference |  | Reference |  | Reference |  |
| 6 h | 0.54 (-2.99, 4.08) | 0.7635 | -0.17 (-3.54, 3.20) | 0.9212 | 0.70 (-2.87, 4.27) | 0.6998 |
| 7 h | 0.65 (-2.73, 4.03) | 0.7063 | -1.22 (-4.46, 2.03) | 0.4625 | 0.77 (-2.80, 4.33) | 0.6726 |
| 8 h | 2.05 (-1.48, 5.59) | 0.2555 | 0.74 (-2.63, 4.12) | 0.6663 | 2.21 (-1.43, 5.85) | 0.2339 |
| 9 h | 7.80 (2.87, 12.74) | 0.0020 | 5.68 (0.98, 10.38) | 0.0181 | 6.29 (1.37, 11.21) | 0.0124 |
| $\geq 10 \mathrm{~h}$ | 19.75 (13.28, 26.21) | <0.0001 | 18.54 (12.39, 24.69) | <0.0001 | 22.01 (15.34, 28.68) | <0.0001 |
| P for trend | <0.001 |  | <0.001 |  | <0.001 |  |
| BMI $\geq 28\left(\mathrm{~kg} / \mathrm{m}^{2}\right)$ |  |  |  |  |  |  |
| Sleep duration | -0.46 (-1.36, 0.44) | 0.3181 | $-0.52(-1.38,0.34) 0$ | 0.2398 | -0.54 (-1.43, 0.35) | 0.2325 |
| Sleep duration categories |  |  |  |  |  |  |
| $<6 \mathrm{~h}$ | Reference |  | Reference |  | Reference |  |
| 6 h | -4.38 (-8.39, -0.37) | 0.0326 | -3.77 (-7.59, 0.05) | 0.0535 | -4.79 (-8.71, -0.88) | 0.0166 |
| 7 h | $-3.85(-7.74,0.03) 0$ | 0.0520 | -4.04 (-7.74, -0.34) | 0.0328 | -4.21 (-8.05, -0.36) | 0.0322 |
| 8 h | -6.35 (-10.32, -2.38) | 0.0018 | -6.28 (-10.07, -2.49) | 0.0012 | -6.18 (-10.08, -2.28) | 0.0020 |
| 9 h | $-0.46(-7.09,6.16) 0$ | 0.8911 | -0.61 (-6.93, 5.71) | 0.8507 | -1.41 (-7.78, 4.95) | 0.6637 |
| $\geq 10 \mathrm{~h}$ | $4.14(-6.08,14.35) 0$ | 0.4277 | 4.45 (-5.29, 14.20) | 0.3707 | -0.56 (-10.84, 9.72) | 0.9143 |
| P for trend | 0.039 |  | 0.146 |  | 0.116 |  |

Crude Model: no covariates were adjusted.
Model I: age, sex, and race/ethnicity were adjusted.
Model II: age, sex, race/ethnicity, education level, smoking status, alcohol consumption status, trouble sleeping, history of hypertension, diabetes and chronic kidney disease.

Table 5
Regression coefficients ( $\beta$ ) and 95 \% confidence interval for sNfL Levels across sleep duration categories.

| Outcome | Crude Model |  | Model I |  | Model II |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\beta$ (95\%CI) | P-value | $\beta$ (95\%CI) | P-value | $\beta$ (95\%CI) | P -value |
| History of trouble sleeping |  |  |  |  |  |  |
| Sleep duration | 2.32 (1.17, 3.46) | <0.0001 | 1.67 (0.55, 2.78) | 0.0036 | 1.91 (0.71, 3.11) | 0.0020 |
| Sleep duration categories |  |  |  |  |  |  |
| $<6 \mathrm{~h}$ | Reference |  | Reference |  | Reference |  |
| 6 h | -0.75 (-5.57, 4.07) | 0.7593 | -1.08 (-5.69, 3.53) | 0.6472 | -0.83 (-5.67, 4.01) | 0.7368 |
| 7 h | -2.35 (-7.37, 2.68) | 0.3606 | -4.98 (-9.85, -0.10) | 0.0461 | -2.99 (-8.18, 2.21) | 0.2604 |
| 8 h | -1.33 (-6.65, 3.98) | 0.6232 | -2.98 (-8.10, 2.14) | 0.2543 | -3.85 (-9.21, 1.51) | 0.1603 |
| 9 h | 7.57 (-0.40, 15.54) | 0.0633 | 2.47 (-5.27, 10.20) | 0.5326 | 3.45 (-4.70, 11.60) | 0.4076 |
| $\geq 10 \mathrm{~h}$ | 39.99 (27.86, 52.12) | <0.0001 | 36.91 (25.29, 48.54) | <0.0001 | 41.78 (29.34, 54.22) | <0.0001 |
| P for trend | 0.003 |  | 0.065 |  | 0.064 |  |
| No history of trouble sleeping |  |  |  |  |  |  |
| Sleep duration | $0.03(-0.65,0.72)$ | 0.9266 | -0.09 (-0.75, 0.56) | 0.7812 | 0.25 (-0.42, 0.92) | 0.4623 |
| Sleep duration categories |  |  |  |  |  |  |
| $<6 \mathrm{~h}$ | Reference |  | Reference |  | Reference |  |
| 6 h | -4.37 (-7.69, -1.06) | 0.0098 | -4.65 (-7.80, -1.50) | 0.0039 | -3.38 (-6.59, -0.18) | 0.0385 |
| 7 h | -3.23 (-6.37, -0.09) | 0.044 | -4.35 (-7.36, -1.35) | 0.0046 | -2.78 (-5.86, 0.30) | 0.0767 |
| 8 h | -4.07 (-7.29, -0.84) | 0.0137 | -4.95 (-8.03, -1.86) | 0.0017 | -2.85 (-5.99, 0.30) | 0.0761 |
| 9 h | 0.14 (-4.59, 4.87) 0 | 0.9531 | -0.01 (-4.53, 4.51) | 0.9972 | 1.83 (-2.73, 6.38) | 0.4314 |
| $\geq 10 \mathrm{~h}$ | 1.87 (-4.45, 8.19) 0 | 0.5623 | 1.73 (-4.29, 7.75) | 0.5732 | 1.48 (-5.00, 7.95) | 0.6554 |
| P for trend | 0.822 |  | 0.890 |  | 0.478 |  |

Crude Model: no covariates were adjusted.
Model I: age, sex, and race/ethnicity were adjusted.
Model II: age, sex, race/ethnicity, education level, BMI, smoking status, alcohol consumption status, history of hypertension, diabetes and chronic kidney disease.
were extracted from NHANEs may not precisely represent population data.

## 6. Conclusion

The relationship between American sNfL level and sleep duration is J shaped, judged from a national representative database. Suggest extreme sleep duration could be deleterious to nervous system (overlong duration can be more obvious than too short). Also, there was an obvious elevate effect in people with normal BMI than higher BMI may indicated that people with higher BMI may
suffered from invisible nervous damage.

## Ethics declarations

Informed consent was not required for this study because this study just analysis second handed data, and the original data collection was approved by NCHS ERB.

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## Data availability

All the original data can be founded at NHANES website, and data generated during this manuscription would be available on reasonable request.

## CRediT authorship contribution statement

Jiaxing Liang: Writing - review \& editing, Writing - original draft, Conceptualization. Tengchi Ma: Visualization, Methodology, Data curation. Youlei Li: Writing - review \& editing. Ruixin Sun: Writing - review \& editing. Shuaishuai Zhao: Writing - review \& editing. Yuzhe Shen: Writing - review \& editing. Hui Gao: Writing - review \& editing. Yunhang Jing: Writing - review \& editing. Xinyue Bai: Writing - review \& editing. Mengze He: Writing - review \& editing. Qingyan Wang: Writing - review \& editing. Huilin Xi: Writing - review \& editing. Rui Shi: Writing - review \& editing. Yanling Yang: Writing - review \& editing, Supervision, Funding acquisition.

## Declaration of competing interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

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## List of abbreviations

ABCC9 ATP-binding cassette, sub-family C, member 9
ACR albumin/creatinine ratio
AD Alzheimer disease
AE acridinium ester
ALS amyotrophic lateral sclerosis
BBB brain-blood barrier
BMI body mass index
CIs confidence intervals
CKD chronic kidney disease
CSF cerebrospinal fluid
CV coefficient of variation
DBP diastolic blood pressure
ERB Ethics Review Board
LLOQ lower limit of quantification
MS multiple sclerosis
MSA multiple system atrophy
NCHS National Center for Health Statistics
NEFL neurofilament gene
NF neurofilament
NfL neurofilament light chain
NHANES National Health and Nutrition Examination Survey
PMP paramagnetic particles
QC quality control

## SBP systolic blood pressure

SCA3 spinocerebellar ataxia type 3
sCJD sporadic Creutzfeldt-Jakob disease
sNfL serum neurofilament light chain
STROBE Strengthening the Reporting of Observational Studies in Epidemiology

## References

[1] J. Hurst, D. Flavell, J.P. Julien, D. Meijer, W. Mushynski, F. Grosveld, The human neurofilament gene (NEFL) is located on the short arm of chromosome 8, Cytogenet. Cell Genet. 45 (1) (1987) 30-32.
[2] A. Al-Chalabi, C.C. Miller, Neurofilaments and neurological disease, Bioessays 25 (4) (2003) 346-355.
[3] J. Kuhle, K. Plattner, J.P. Bestwick, R.L. Lindberg, S.V. Ramagopalan, N. Norgren, A. Nissim, A. Malaspina, D. Leppert, G. Giovannoni, L. Kappos, A comparative study of CSF neurofilament light and heavy chain protein in MS, Mult. Scler. 19 (12) (2013) 1597-1603.
[4] M. Bacioglu, L.F. Maia, O. Preische, J. Schelle, A. Apel, S.A. Kaeser, M. Schweighauser, T. Eninger, M. Lambert, A. Pilotto, D.R. Shimshek, U. Neumann, P. J. Kahle, M. Staufenbiel, M. Neumann, W. Maetzler, J. Kuhle, M. Jucker, Neurofilament light chain in blood and CSF as marker of disease progression in mouse models and in neurodegenerative diseases, Neuron 91 (1) (2016) 56-66.
[5] A. Yuan, M.V. Rao, Nixon RA. Veeranna, Neurofilaments and neurofilament proteins in health and disease, Cold Spring Harbor Perspect. Biol. 9 (4) (2017).
[6] M. Carmona-Iragui, D. Alcolea, I. Barroeta, L. Videla, L. Munoz, K.L. Van Pelt, F.A. Schmitt, D.D. Lightner, L.M. Koehl, G. Jicha, S. Sacco, C. Mircher, S.E. Pape, R. Hithersay, I.C.H. Clare, A.J. Holland, G. Nubling, J. Levin, S.H. Zaman, A. Strydom, A.S. Rebillat, E. Head, R. Blesa, A. Lleo, J. Fortea, Diagnostic and prognostic performance and longitudinal changes in plasma neurofilament light chain concentrations in adults with Down syndrome: a cohort study, Lancet Neurol. 20 (8) (2021) 605-614.
[7] L. Gong, Q. Yu, H. Wang, C. Xu, Y. Dou, B. Mao, Y. Zhao, Neurofilament light chain (NF-L) stimulates lipid peroxidation to neuronal membrane through microglia-derived ferritin heavy chain (FTH) secretion, Oxid. Med. Cell. Longev. 2022 (2022) 3938940.
[8] W.W. Aamodt, T. Waligorska, J. Shen, T.F. Tropea, A. Siderowf, D. Weintraub, M. Grossman, D. Irwin, D.A. Wolk, S.X. Xie, J.Q. Trojanowski, L.M. Shaw, A. S. Chen-Plotkin, Neurofilament light chain as a biomarker for cognitive decline in Parkinson disease, Mov. Disord. 36 (12) (2021) $2945-2950$.
[9] J.M. Jarvis, E.L. Fink, Neurofilament light chain-it is not just about concussions, Pediatr. Crit. Care Med. 21 (7) (2020) 685-686.
[10] G. Giovannoni, Peripheral blood neurofilament light chain levels: the neurologist's C-reactive protein? Brain 141 (8) (2018) $2235-2237$.
[11] L. Parnetti, L. Gaetani, M. Di Filippo, Serum neurofilament light chain as a preclinical marker of neurodegeneration, Lancet Neurol. 18 (12) (2019) $1070-1071$.
[12] M. Khalil, C.E. Teunissen, M. Otto, F. Piehl, M.P. Sormani, T. Gattringer, C. Barro, L. Kappos, M. Comabella, F. Fazekas, A. Petzold, K. Blennow, H. Zetterberg, J. Kuhle, Neurofilaments as biomarkers in neurological disorders, Nat. Rev. Neurol. 14 (10) (2018) 577-589.
[13] Q.F. Li, Y. Dong, L. Yang, J.J. Xie, Y. Ma, Y.C. Du, H.L. Cheng, W. Ni, Z.Y. Wu, Neurofilament light chain is a promising serum biomarker in spinocerebellar ataxia type 3, Mol. Neurodegener. 14 (1) (2019) 39.
[14] A. Stokowska, L. Bunketorp Kall, C. Blomstrand, J. Simren, M. Nilsson, H. Zetterberg, K. Blennow, M. Pekny, M. Pekna, Plasma neurofilament light chain levels predict improvement in late phase after stroke, Eur. J. Neurol. 28 (7) (2021) 2218-2228.
[15] W.F. Abdo, B.R. Bloem, W.J. Van Geel, R.A. Esselink, M.M. Verbeek, CSF neurofilament light chain and tau differentiate multiple system atrophy from Parkinson's disease, Neurobiol. Aging 28 (5) (2007) 742-747.
[16] J.J. van Eijk, B. van Everbroeck, W.F. Abdo, B.P. Kremer, M.M. Verbeek, CSF neurofilament proteins levels are elevated in sporadic Creutzfeldt-Jakob disease, J. Alzheimers Dis. 21 (2) (2010) 569-576.
[17] R.A. Menke, E. Gray, C.H. Lu, J. Kuhle, K. Talbot, A. Malaspina, M.R. Turner, CSF neurofilament light chain reflects corticospinal tract degeneration in ALS, Ann. Clin. Transl. Neurol. 2 (7) (2015) 748-755.
[18] I. Fyfe, Alzheimer disease: neurofilament light in the blood marks Alzheimer degeneration, Nat. Rev. Neurol. 13 (5) (2017) 257.
[19] D. Leppert, J. Kuhle, Blood neurofilament light chain at the doorstep of clinical application, Neurol. Neuroimmunol. Neuroinflamm. 6 (5) (2019) e599.
[20] J. Piepgras, A. Muller, F. Steffen, J. Lotz, C. Loquai, F. Zipp, C. Dresel, S. Bittner, Neurofilament light chain levels reflect outcome in a patient with glutamic acid decarboxylase 65 antibody-positive autoimmune encephalitis under immune checkpoint inhibitor therapy, Eur. J. Neurol. 28 (3) (2021) 1086-1089.
[21] J.P. Chaput, A.W. McHill, R.C. Cox, J.L. Broussard, C. Dutil, B.G.G. da Costa, H. Sampasa-Kanyinga, K.P. Wright Jr., The role of insufficient sleep and circadian misalignment in obesity, Nat. Rev. Endocrinol. 19 (2) (2023) 82-97.
[22] R. Ren, N. Covassin, L. Yang, Y. Li, Y. Zhang, J. Zhou, L. Tan, T. Li, X. Li, Y. Wang, J. Zhang, Y.K. Wing, W. Li, V.K. Somers, X. Tang, Objective but not subjective short sleep duration is associated with hypertension in obstructive sleep apnea, Hypertension 72 (3) (2018) 610-617.
[23] Y.H. Chen, Z.Y. Lyu, G. Wang, X.S. Feng, S.H. Xie, S.H. Chen, J. Yin, J.S. Ren, Z.H. Mi, S. Wang, S.L. Wu, N. Li, M. Dai, Relationship of sleep duration and annual changes in sleep duration with the incidence of gastrointestinal cancers: a prospective cohort study, Chin. Med. J. 134 (24) (2021) $2976-2984$.
[24] T.M. Shockey, A.G. Wheaton, Short sleep duration by occupation group - 29 states, 2013-2014, MMWR Morb. Mortal. Wkly. Rep. 66 (8) (2017) 207-213.
[25] N.F. Watson, M.S. Badr, G. Belenky, D.L. Bliwise, O.M. Buxton, D. Buysse, D.F. Dinges, J. Gangwisch, M.A. Grandner, C. Kushida, R.K. Malhotra, J.L. Martin, S. R. Patel, S. Quan, E. Tasali, Recommended amount of sleep for a healthy adult: a joint consensus statement, of the American Academy of sleep medicine and sleep Research society, Sleep 38 (6) (2015).
[26] M.K. Lemke, Y. Apostolopoulos, A. Hege, S. Sonmez, L. Wideman, Understanding the role of sleep quality and sleep duration in commercial driving safety, Accid. Anal. Prev. 97 (2016) 79-86.
[27] J. Fernandez-Mendoza, S. Calhoun, E.O. Bixler, S. Pejovic, M. Karataraki, D. Liao, A. Vela-Bueno, M.J. Ramos-Platon, K.A. Sauder, A.N. Vgontzas, Insomnia with objective short sleep duration is associated with deficits in neuropsychological performance: a general population study, Sleep 33 (4) (2010) 459-465.
[28] F. Warsame, N.M. Chu, J. Hong, A. Mathur, D.C. Crews, G. Bayliss, D.L. Segev, M.A. McAdams-DeMarco, Sleep duration and cognitive function among older adults with chronic kidney disease: results from the National Health and Nutrition Examination Survey (2011-2014), Nephrol. Dial. Transplant. 38 (7) (2023) 1636-1644.
[29] X. Liu, H. Zhou, Sleep duration, insomnia and behavioral problems among Chinese adolescents, Psychiatr. Res. 111 (1) (2002) 75-85.
[30] C. Bornhorst, S. Hense, W. Ahrens, A. Hebestreit, L. Reisch, G. Barba, R. von Kries, O. Bayer, I. Consortium, From sleep duration to childhood obesity-what are the pathways? Eur. J. Pediatr. 171 (7) (2012) 1029-1038.
[31] N.F. Watson, K.P. Harden, D. Buchwald, M.V. Vitiello, A.I. Pack, D.S. Weigle, J. Goldberg, Sleep duration and body mass index in twins: a gene-environment interaction, Sleep 35 (5) (2012) 597-603.
[32] C.A. Magee, J.K. Lee, S.A. Vella, Bidirectional relationships between sleep duration and screen time in early childhood, JAMA Pediatr. 168 (5) (2014) 465-470.
[33] Y. Qin, Y. Zhou, X. Zhang, X. Wei, J. He, Sleep duration and breast cancer risk: a meta-analysis of observational studies, Int. J. Cancer 134 (5) (2014) 1166-1173.
[34] R.N. Aurora, J.S. Kim, C. Crainiceanu, D. O'Hearn, N.M. Punjabi, Habitual sleep duration and all-cause mortality in a general community sample, Sleep 39 (11) (2016) 1903-1909.
[35] K.E. Minges, N.S. Redeker, Delayed school start times and adolescent sleep: a systematic review of the experimental evidence, Sleep Med. Rev. 28 (2016) 86-95.
[36] S. Liu, Z. Zhang, S. Shi, Y. Meng, X. Zhang, Q. Lei, Z. Li, NREM sleep loss increases neurofilament light chain levels in APP/PS1 and C57BL/6 J mice, Sleep Breath. 27 (4) (2023) 1495-1504.
[37] O. World Health, International Statistical Classification of Diseases and Related Health Problems, fifth ed., World Health Organization, Geneva, 2015.
[38] A.K. Cheung, T.I. Chang, W.C. Cushman, S.L. Furth, F.F. Hou, J.H. Ix, G.A. Knoll, P. Muntner, R. Pecoits-Filho, M.J. Sarnak, S.W. Tobe, C.R.V. Tomson, L. Lytvyn, J.C. Craig, D.J. Tunnicliffe, M. Howell, M. Tonelli, M. Cheung, A. Earley, J.F.E. Mann, Executive summary of the KDIGO 2021 clinical Practice guideline for the management of blood pressure in chronic kidney disease, Kidney Int. 99 (3) (2021) 559-569.
[39] Z. Zhang, Z. Li, X. Zhang, W. Ye, J. Chen, L. Wang, Z. Lin, J. Li, Z. Li, Association between secondhand smoke and cancers in adults in the US population, J. Cancer Res. Clin. Oncol. 149 (7) (2023) 3447-3455.
[40] NHANE Tutorial, August 02, 2023 [cited; NHANES Turials]. Available from: https://wwwn.cdc.gov/nchs/NHANES/tutorials/default.aspx, 2023.
[41] M. Kahn, O. Schnabel, M. Gradisar, G.S. Rozen, M. Slone, N. Atzaba-Poria, L. Tikotzky, A. Sadeh, Sleep, screen time and behaviour problems in preschool children: an actigraphy study, Eur. Child Adolesc. Psychiatr. 30 (11) (2021) 1793-1802.
[42] J. Theorell-Haglow, E. Lindberg, Sleep duration and obesity in adults: what are the connections? Curr. Obes. Rep. 5 (3) (2016) 333-343.
[43] L. Zhai, H. Zhang, D. Zhang, Sleep duration and depression among adults: a meta-analysis of prospective studies, Depress. Anxiety 32 (9) (2015) 664-670.
[44] W. Chen, J.P. Wang, Z.M. Wang, P.C. Hu, Y. Chen, Association between sleep duration and chest pain in US adults: a cross-sectional study, Front. Public Health 10 (2022) 952075.
[45] J.R. Winer, K.D. Deters, G. Kennedy, M. Jin, A. Goldstein-Piekarski, K.L. Poston, E.C. Mormino, Association of short and long sleep duration with amyloid-beta burden and cognition in aging, JAMA Neurol. 78 (10) (2021) 1187-1196.
[46] T. Geis, S. Brandstetter, A.A. Toncheva, O. Laub, G. Leipold, R. Wagner, M. Kabesch, S. Kasser, J. Kuhle, S. Wellmann, Study g CoKiBa, Serum neurofilament light chain (sNfL) values in a large cross-sectional population of children with asymptomatic to moderate COVID-19, J. Neurol. 268 (11) (2021) $3969-3974$.
[47] A.L. Fisse, K. Pitarokoili, D. Leppert, J. Motte, X. Pedreiturria, L. Kappos, R. Gold, J. Kuhle, M.S. Yoon, Serum neurofilament light chain as outcome marker for intensive care unit patients, J. Neurol. 268 (4) (2021) 1323-1329.
[48] Z. Alirezaei, M.H. Pourhanifeh, S. Borran, M. Nejati, H. Mirzaei, M.R. Hamblin, Neurofilament light chain as a biomarker, and correlation with magnetic resonance imaging in diagnosis of CNS-related disorders, Mol. Neurobiol. 57 (1) (2020) 469-491.
[49] S. Modvig, M. Degn, B. Sander, H. Horwitz, B. Wanscher, F. Sellebjerg, J.L. Frederiksen, Cerebrospinal fluid neurofilament light chain levels predict visual outcome after optic neuritis, Mult. Scler. 22 (5) (2016) 590-598.
[50] L. Stefan, G. Vrgoc, T. Rupcic, G. Sporis, D. Sekulic, Sleep duration and sleep quality are associated with physical activity in elderly people living in nursing homes, Int. J. Environ. Res. Publ. Health 15 (11) (2018).
[51] P. Zhang, C.W. Tan, G.H. Chen, Y.J. Ge, J. Xu, L. Xia, F. Wang, X.Y. Li, X.Y. Kong, Patients with chronic insomnia disorder have increased serum levels of neurofilaments, neuron-specific enolase and S100B: does organic brain damage exist? Sleep Med. 48 (2018) 163-171.
[52] Y. Li, F. Li, X. Liu, J. Zu, W. Zhang, S. Zhou, J. Zhu, T. Zhang, G. Cui, C. Xu, Association between serum neurofilament light chain levels and sleep disorders in patients with Parkinson's disease, Neurosci. Lett. 812 (2023) 137394.
[53] S. Ai, J. Zhang, G. Zhao, N. Wang, G. Li, H.C. So, Y. Liu, S.W. Chau, J. Chen, X. Tan, F. Jia, X. Tang, J. Shi, L. Lu, Y.K. Wing, Causal associations of short and long sleep durations with 12 cardiovascular diseases: linear and nonlinear Mendelian randomization analyses in UK Biobank, Eur. Heart J. 42 (34) (2021) 3349-3357.
[54] H. Lee, S. Kim, B.S. Kim, M. Kim, J. Yang, H. Bae, C.W. Won, Sexual difference in effect of long sleep duration on incident sarcopenia after two years in community-dwelling older adults, Annals of Geriatric Medicine and Research 26 (3) (2022) 264-274.
[55] A.J. Cox, N.P. West, A.W. Cripps, Obesity, inflammation, and the gut microbiota, Lancet Diabetes Endocrinol. 3 (3) (2015) 207-215.
[56] W. Ying, W. Fu, Y.S. Lee, J.M. Olefsky, The role of macrophages in obesity-associated islet inflammation and beta-cell abnormalities, Nat. Rev. Endocrinol. 16 (2) (2020) 81-90.
[57] F. Fusco, N. Longo, M. De Sio, D. Arcaniolo, G. Celentano, M. Capece, R. La Rocca, F. Mangiapia, G. Califano, S. Morra, C. Turco, G. Spena, L. Spirito, G.M. Fusco, L. Cirillo, L. De Luca, L. Napolitano, V. Mirone, M. Creta, Impact of circadian desynchrony on spermatogenesis: a mini review, Front. Endocrinol. 12 (2021) 800693.
[58] S. Cai, E.K.H. Tham, H.Y. Xu, X. Fu, R.S.M. Goh, P.D. Gluckman, Y.S. Chong, F. Yap, L.P. Shek, O. Hoe Teoh, J.J. Gooley, D. Yam-Thiam Goh, M.J. Meaney, N. Schneider, A. Rifkin-Graboi, B.F.P. Broekman, Trajectories of reported sleep duration associate with early childhood cognitive development, Sleep 46 (2) (2023).
[59] W. Liu, Q. Wu, M. Wang, P. Wang, N. Shen, Prospective association between sleep duration and cognitive impairment: findings from the China health and retirement longitudinal study (CHARLS), Front. Med. 9 (2022) 971510.
[60] K.V. Allebrandt, N. Amin, B. Müller-Myhsok, T. Esko, M. Teder-Laving, R.V. Azevedo, C. Hayward, J. van Mill, N. Vogelzangs, E.W. Green, S.A. Melville, P. Lichtner, H.E. Wichmann, B.A. Oostra, A.C. Janssens, H. Campbell, J.F. Wilson, A.A. Hicks, P.P. Pramstaller, Z. Dogas, I. Rudan, M. Merrow, B. Penninx, C. P. Kyriacou, A. Metspalu, C.M. van Duijn, T. Meitinger, T. Roenneberg, A K(ATP) channel gene effect on sleep duration: from genome-wide association studies to function in Drosophila, Mol. Psychiatr. 18 (1) (2013) 122-132.


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