Variants in Exon 11 of MEF2A Gene and Coronary Artery Disease: Evidence from a Case-Control Study, Systematic Review, and Meta-Analysis

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Abstract

Background: Coronary artery disease (CAD) is the most common heart disease worldwide. Association of CAD with variants in the myocyte enhancer factor 2A (*MEF2A*) gene, the first identified CAD-causing gene, has attracted special attention but the results are controversial. We aimed to evaluate this genetic association via a case-control study and meta-analysis.

Methodology/Principal Findings: We performed a case-control association study to investigate the relationship between variations in exon 11 of *MEF2A* gene and CAD in 1045 sporadic patients and 1008 controls enrolled angiographically among southern Chinese population, and then the data from this study were compared and discussed in a systematic review and meta-analysis with all available published studies on *MEF2A* gene and CAD. In total, eight variants were identified (21-bp deletion, CAG repeats, CCG repeats, a CCA deletion and four SNPs). No significant link was observed between the common (CAG)_n polymorphism and CAD, whereas the rare 21-bp deletion was detected only in five affected individuals. The meta-analysis of (CAG)_n polymorphism and CAD risk, including nine studies with 3801 CAD patients and 4020 controls, also provided no convincing evidence for the genetic association, even upon stratification by race (mainly Whites and Chinese). However, the 21-bp deletion was regarded as a potentially logical, albeit undetermined, candidate for CAD in the following systematic review.

Conclusions/Significance: Our findings failed to demonstrate a correlation between (CAG)_n polymorphism with CAD, however, we concluded that the rare 21-bp deletion might have a more compelling effect on CAD than the common (CAG)_n polymorphism, and *MEF2A* genetic variant might be a rare but specific cause of CAD/MI.

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Introduction

Coronary artery disease (CAD) is a common complex disorder resulting from both genetic and environmental influences [1,2], and it has become a major cause of death and disability in China. The role of genetic alterations and their impact on CAD susceptibility remains unclear and has attracted more attention. In the past three decades, genetic association studies and genomewide linkage scans have revealed a considerable number of candidate loci and genes for CAD and myocardial infarction (MI) [3–7], but results are not often reproducible [8–10].

In 2003, a 7-amino acid deletion, caused by a 21-base pair (bp) coding sequence deletion in exon 11 of the myocyte enhancer factor 2A (*MEF2A*) gene, was reported as a causative mutation in a single large CAD/MI family of Scandinavia ancestry [11]. *In vitro* functional analysis indicated that the 21-bp deletion disrupted the nuclear localization of mature protein and decreased *MEF2A*-induced transcriptional activation. Thus this genetic imperfection might lead to a defective or abnormal vascular endothelium, which

could promote the genesis of atherosclerotic plaque or thrombosis and influence the whole process of atherogenesis [11]. Subsequently, the same researchers discovered three functional variants (Asn263Ser, Pro279Leu and Gły283Asp) in exon 7 in approximately 2% of the affected population, but none in unaffected individuals [12]. Thence, *MEF2A* gene has been considered as the first CAD-causing gene to be identified.

The genomic sequence of MEF2A gene is highly polymorphic. It is thus of added interest to detect which or how many MEF2Agenetic variants might have functional potential to affect the final bioavailability of MEF2A, and further the development of CAD. In fact, many case-control studies have attempted to investigate the unequivocal effects of MEF2A gene on CAD, especially its exon 11, claimed as the most polymorphic locus harboring various substitution and insertion/deletion (indel) polymorphisms such as a common variant (CAG)_n polymorphism. However, the results have been inconsistent [13–20].

With the improved genotyping technologies and the completion of the human HapMap project, Genome-Wide Association Studies (GWASs) have been developed as an important approach in genetic research. Thus far, a large number of candidate loci conferring risk of or protection from common complex diseases such as CAD have been proposed [21–25]. Nonetheless, neither the *MEF2A* locus on chromosome 15q26 nor its adjacent region has been identified in any of the previous GWASs, thus generating debate over the nature of *MEF2A* genetic contribution to individual susceptibility to CAD.

To elucidate the relationship between MEF2A gene and its effect on CAD risk, we focused on its exon 11, the highly polymorphic and controversial region, and established a wellcharacterized case-control study of 1045 sporadic CAD patients and 1008 controls with normal coronary arteries. In addition, we reviewed all available studies reported in the literature to examine the association of the common (CAG)_n polymorphism and the rare 21-bp deletion with CAD, and to assess whether variations in study design and study population ethnicity could lead to potential biases and be the sources of between-study heterogeneity.

Materials and Methods

Case-control study

Ethics Statement. Approval to undertake this study was granted by the Ethics Review Committee of Ruijin Hospital, Shanghai Jiao Tong University School of Medicine and was conducted according to the Declaration of Helsinki Principles. Written informed consents were obtained from each participant at enrollment.

Study population. This was a hospital-based case-control study including a total of 2053 unrelated Han Chinese admitted to Ruijin Hospital, Shanghai Jiao Tong University School of Medicine when they were experiencing various symptoms or for a medical checkup from January 2006 to September 2009. All participants underwent coronary angiography and were divided into CAD group and control group according to their angiographic results.

The CAD group contained 1045 sporadic patients aged 65.49 ± 9.83 years, and the diagnosis of CAD was determined angiographically based on the presence of more than 70% stenosis in at least one of the three major coronary arteries or major branches. Patients with simple spasm of coronary arteries, myocardial bridge or other non-coronary atherosclerotic lesions were excluded. The remaining participants (n = 1008), aged 60.23\pm10.49 years, had normal coronary arteries (NCA) on angiography, formed the control group.

All patients with the 21-bp deletion were followed up every year in a special CAD clinic. At each visit, clinical manifestations and echocardiography were recorded. Adverse events (e.g. hospitalization, cardiac dysfunction, percutaneous coronary intervention, coronary artery bypass grafting, or death) were reported during the visit or through telephone conversation with the patients or their family members. Two trained physicians independently reviewed all medical notes, including emergency department visit forms and hospital medical records.

Screening for variations in *MEF2A* exon 11. Blood samples (5 ml) were drawn and genomic DNA was extracted from peripheral blood leukocytes by standard phenol-chloroform extraction. To assess the distribution patterns of the structural variations of *MEF2A* exon 11 in this cohort study, we sequenced the entire exon 11 using the direct DNA sequencing method in all 2053 subjects. Primers were designed by the Primer3 software (http://frodo.wi.mit.edu/cgi-bin/primer3/primer3) according to reference sequence (NM_005920.2). In detail, the sequence of the forward primer was 5'-gca gag gta ctt gca agc cat ctg-3' and the reverse was 5'-ggt cgg cca agc aca att gga gaa-3'. The sequencing

primer was 5'-caa gca caa ttg gag aat gga-3'. Sequences were analyzed using an ABI Prism BigDye Terminator Cycle Sequencing Kit on an ABI Prism 3700 sequencer, version 3.1 (Applied Biosystems, Foster City, CA, USA), as described in detail in Text S1.

Systematic review and meta-analysis

Data sources and search strategies. We collected information via two international searching engines, viz. PubMed and Excerpta Medica Database (EMBASE), and two Chinese searching engines, viz. Wanfang database (http://www. wanfangdata.com.cn) and China Biological Medicine (CBM) (http://sinomed.imicams.ac.cn/index.jsp) with the last update on July 31, 2011. We restricted search results to papers published in English or Chinese. We combined the subject terms of 'coronary artery disease or coronary disease or arteriosclerosis or atherosclerosis or myocardial infarction or angina pectoris' and 'myocyte enhancer factor 2A' with either 'gene', 'variation', 'variant', 'mutation', 'polymorphism' or 'allele', which were all MeSH (Medical Subject Headings in the US National Library of Medicine) terms. The "related articles" in the MEDLINE option as well as reference lists of all retrieved studies were also checked for citations of other relevant publications that were not identified initially. All studies were considered potentially eligible if they aimed to investigate the relationship between MEF2A genetic polymorphisms and CAD risk. If there were multiple publications from the same study group, the most complete and recent results were extracted. Search results were limited to studies performed in human subjects without country restrictions and ethnic restrictions.

Inclusion/exclusion criteria. We enrolled all prevalent case-control or nested case-control or cross-sectional studies in this meta-analysis regardless of sample size, if 1) they explored the association of *MEF2A* genetic polymorphisms with CAD/MI, 2) genotyping had been performed by using validated methods, and 3) they provided the sufficient information on genotype/allele counts or frequencies for estimating odds ratio (OR) and 95% confidence interval (95% CI). We calculated the effect estimate against healthy subjects/NCA controls.

Data extraction. Data were extracted independently and entered into separate databases by two authors (Y. Liu and W. Niu) from each qualified study: first author's last name, publication date, population ethnicity, study design, diagnostic criteria, genotyping methods, baseline characteristics of the study population, such as age, gender, history of hypertension and diabetes mellitus, if available, and the number of persons with different alleles in cases and controls and available subgroups. Discrepancies between the two databases were identified by comparison. A third author (W. Jin) checked for them and a consensus was reached after discussion. For consistency, continuous variables such as age were uniformly expressed as mean \pm standard deviation (S.D.)

Statistical analysis. For our case-control study, database management and statistical calculation were conducted using SPSS version 13.0 (SPSS Inc., Chicago, Illinois, USA). The Student's *t*-test for continuous variables and the χ^2 -test for categorical ones were used to test differences between cases and controls, OR of CAD risk and their 95% CI were calculated as well.

Hardy-Weinberg equilibrium calculations were performed with the Arlequin program (http://anthro.unige.ch/software/arlequin). The Haplo.stats package (version 1.4.0) in the R statistical computing software (http://www.r-project.org) was used to analyze haplotype-based association study. Two-tailed P<0.05 was accepted as statistically significant.

In the following meta-analysis, pooled association relating $(CAG)_n$ polymorphism to CAD risk was performed by the Review Manager software (version 5.0.19; http://www.cc-ims. net/revman/download). Using the most common type $(CAG)_9$ allele as a reference, comparisons of other $(CAG)_n$ alleles between cases and controls were expressed in the form of OR and 95% CI. The allele effects were estimated using the model-free approach, where no assumption about genetic models was required. In addition, stratification analyses were conducted to seek more narrowly drawn subsets of the studies such as different genotyping methods, population origins and study designs. We implemented the random-effects model using the method of DerSimonian and Laird, instead of fixed-effects model, to bring the individual effect-size estimates together, and the estimate of heterogeneity was analyzed by the Mantel-Haenszel method [26–28].

The presence of between-study heterogeneity across all eligible comparisons was calculated using the χ^2 -based Cochrane's Q statistic with statistical significance at the level of 0.10 as this statistic has proven to have poor power if there are few studies [28,29]. Besides, the I^2 statistic was documented for the percentage of the observed between-study variability due to heterogeneity rather than chance with the ranges of 0–100% ($I^2 = 0-25\%$, no heterogeneity; $I^2 = 25-50\%$, moderate heterogeneity; $I^2 = 50-75\%$, large heterogeneity; $I^2 = 75-100\%$, extreme heterogeneity) [28].

Finally, publication bias was assessed by the fail-safe number (\mathcal{N}_{fs}) of each meta-analysis [30]. If the \mathcal{N}_{fs} was smaller than the number of observed studies for a polymorphism, it is believed that the meta-result might have a significant publication bias. In this study, the \mathcal{N}_{fs} significance was established at $P < 0.05 \ (\mathcal{N}_{fs0.05} = (\Sigma Z / 1.64)^2 - k$; where k is the number of articles included in each meta-analysis).

Results

Clinical characteristics of our study population

The clinical characteristics of our study population are shown in Table S1. Compared with NCA controls, CAD patients were older (P<0.001) and more often of the male gender (P<0.001). As expected, the CAD group had a higher prevalence of conventional cardiovascular risk factors, including diabetes and dyslipidemia (P < 0.05). They had higher serum levels of fasting glucose, total cholesterol, triglycerides, and low density lipoprotein-cholesterol, and lower levels of high density lipoprotein-cholesterol. However, the morbidity of hypertension was similar between the two groups.

Genetic information on our case-control study

Eight variants were identified by sequencing the entire exon 11 in 2053 unrelated Chinese individuals (Table 1). No significant deviation from the Hardy-Weinberg equilibrium was detected for each polymorphism in both CAD patients and NCA controls.

The number of the CAG triplet repeats (polyglutamine tandem repeats, $(Q)_n$) spanned from 4 to 15, and the majority of individuals had 9–11 repeats. Shown in Table 2 are the allele distributions of (CAG)_n polymorphism. No statistical significance was observed (P=0.347) for the allelic association of this polymorphism with CAD, and the distribution of genotypes was also similar in two groups (data not shown).

Closely following the (CAG)_n polymorphism was the CCG triplet repeats varying between 4 and 5 proline tandem repeats (P_4 or P_5). More than 95% of individuals in both CAD patients and NCA controls contained five prolines, and the frequencies of (CCG)_n allele and genotypes yielded no significant differences between two groups (data not shown). Additionally, a CCA deletion resulting in lack of one proline amino acid, located adjacent to (CCG)_n site, was detected only in one unaffected subject.

Interestingly, the 21-bp deletion was found only in five independent CAD patients, and none in NCA subjects. In this cohort, they all had some traditional CAD risk factors, including dyslipidemia, hypertension and family history of cardiovascular diseases; three showed severe lesion in the left main coronary artery and two were diagnosed with premature CAD. After a 5year follow up, one died of sudden cardiac death, one took stent treatment and three underwent coronary artery bypass grafting (Table S2).

Table 1. Genetic variations in *MEF2A* gene exon 11 discovered by sequencing in our study population.

Categories	Variants	AA code ¹	MAF (allele frequency,	%)	P value; OR [95% CI]
			CAD Cases(n = 1045)	Controls(n = 1008)	
STR	(CAG) _n (1257–1290)	$(Q)_n (n = 4 - 15)$	More details in Table 4.		
Deletion	CCG (1291–1293)	P deletion (n=4 or 5, 431/432)	111(5.3)	109 (5.4)	0.892; 1.019 [0.777, 1.337]
	CCA (1297–1299)	P deletion (433/434)	0	1(0.05)	
	21-bp deletion (1303–1337)	QPPQPQP deletion (434–446 AA)	5(0.2)	0	
SNP	A1299G	P433P	1(0.05)	1(0.05)	0.980; 1.036 [0.065, 16.571]
	C1303T	P435S	9 (0.4)	5 (0.2)	0.315; 0.574 [0.192, 1.717]
	G1305A	P435P	92 (4.4)	91 (4.5)	0.867; 0.975 [0.725, 1.312]
	G1353T	G443G	693 (33.2)	689 (34.2)	0.490; 1.047 [0.920, 1.191]

 $^{1}Q = Gln; P = Pro; S = Ser; G = Gly.$

STR: short tandem repeat polymorphism; SNP: single nucleotide polymorphism; AA: amino acid; MAF: minor allele frequency; OR: odds ratio; 95%CI: 95% confidence interval.

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Table 2. The baseline characteristics of all the studies relevant to (CAG)_n polymorphism in this meta-analysis.

lef.	Study	Years	Ethnicity	Status	Enrollment criteria		Genotyping methods	Study siz	e	Cases (alle	ele freque	ncies, %)		Controls ((allele frec	quencies, 9	4 (%	value
					Cases	Controls		Cases (n)	Controls (n)	6	10	۲	others 9		10	11	others	
17	Weng et al.	2005	White (Canadian)	premature CAD/MI	Angiography confirmed CAD (>50% stenosis), MI or CABG	Symptom investigation	Sequence	287	296	197(32.8)	76(12.7)	292(48.7)	9(1.5) ¹	217(36.2)	80(13.3)	288(48.0)	7(1.2) 0	.787
31	Yuan et al.	2006	Asian (southern Chinese)	CAD/MI	Angiography confirmed CAD (>50% stenosis), MI or CABG	Normal angiography	PCR-SSCP	175	228	85(24.3)	66(18.9)	185(52.9)	14(4.0) ² 1	06(23.2)	85(18.6)	246(53.9)	19(4.2) 0	.985
13	González et al.	2006	White (Spanish)	W	Angiography confirmed CAD (>30% stenosis), MI	Symptom investigation	Sequence	211	301	136(32.2)	68(16.1)	216(51.2)	2(1.0) ³ 1	84(30.6)	97(16.1)	316(52.5)	5 (0.8)	× 0.05
4	Han et al.	2007	Asian (Northern Chinese)	CAD/MI	Angiography confirmed CAD (>50% stenosis), CABG or MI	Normal angiography	PCR-SSCP	378	348	276(36.5)	158(20.9)	306(40.5)	16(2.1) ⁴ 1	58(22.7)	146(21.0)	362(52.0)	30(4.3) C	.001
32	Gulec et al.	2008	Turk	premature MI	W	Symptom investigation	PCR-SSCP	69	87	70(51.0)	I	62(44.5)	6(4.5) ⁵ 5	8(33.5)	I	98 (56.0)	18(10.5) >	-0.05
33	Lieb et al.	2008	White (German)	premature MI	W	Symptom investigation	PCR-SSCP	543	1190	393(36.2)	164(15.1)	510(47.0)	19(1.7) 8	865(36.3)	350(14.7)	1132(47.6)	33(1.4) 0	.800
18	Hsu et al.	2009	Asian (Taiwanese)	CAD/MI	Angiography confirmed CAD (>50% stenosis), MI	Symptom investigation	Maldi- Tof MS	258	258	209(40.5)	108(20.9)	166(32.2)	33(6.4) ⁶ 1	80(34.9)	119(23.1)	189(36.6)	28(5.4) C	.179
19	Dai et al.	2010	Asian (Northern Chinese)	CAD	Angiography confirmed CAD (≥75% stenosis)	Normal angiography	Sequence	835	304	675(40.4)	319(19.1)	551 (33.0)	125(7.5) ⁷ 2	254(41.8)	110(18.1)	213(35.0)	31(5.1) 0	.052
	This study		Asian (Southern Chinese)	CAD/MI	Angiography confirmed CAD (≥70% stenosis)	Normal angiography	Sequence	1045	1008	812(38.9)	458(21.9)	670(32.1)	150(7.2) ⁸ 8	809(40.1)	434(21.5)	646(32.0)	127(6.3) 0	.374
CAD: MALI Num Num 2 Num 4 Num 5 Num 7 Num 8 Num Comp	coronary ar coronary ar bler OF MS: 1, bler of (CAK hber	tery disea matrix-ass matrix-ass a), repeat a), repeat b), repeat b), repeat b), repeat b), repeat veen (CAC veen (CAC	se, MI: myoca isted laser de: s: 9, 10 or 11 s: 9, 10 or 11 j) repeats an 031406.t002	dial infarction sorption/foniz repeats; othe repeats; othe repeats; othe repeats; othe repeats; othe repeats; othe repeats; othe repeats; othe dother types	ty PTCA: percutaneous \cdot action time-of-flight m ⁶ r, alleles with 5–8, 12, \cdot , alleles with 4, 7 or 8, \cdot , alleles with 6, 7, 8, 1 \cdot , alleles with 4, 5, 7, 8, \cdot , alleles with 4, 5, 7, 8 \cdot , alleles with 4, 5, 7, 8 \cdot , alleles with 4–8, 12– \cdot , alleles with 4–8, 12– \cdot , alleles with 4–8, 12– \cdot , alleles with 4–8, 12–	coronary angiop ass spectrometry or 13 repeats. 12, or 13 repeats. 12, or 13 repeats. 1, 12-15 repeats. 1, 12, 14, or 15 ru 15 repeats. 15 repeats.	lasty; CABG: con epeats. inager software;	onary arten	/ bypass graff	ing; PCR: p 95% confic	olymerase o dence inter-	chain reacti val.	on; SSCP: s	ingle stran	id conform.	ational poly	norphism	analysis;

Besides the aforementioned four variants, we identified three synonymous SNPs (A1299G, G1305A, and G1353T) and one non-synonymous SNP (C1303T) in exon 11. The A1299G and C1303T polymorphisms were rare SNPs with minor allele frequencies being 0.1% and 0.4% among NCA controls, respectively. As a result, none of these four SNPs were associated with CAD (data not shown).

After dropping four rare variations (21-bp deletion, CCA deletion, A1299G and C1303T polymorphisms), we evaluated the remaining four common variants for further haplotype analysis. Using the most common haplotype $(Q)_9$ - P_5 -G-T (in order of $(CAG)_n$, $(CCG)_n$, G1305A and G1353T) (26.0% in NCA controls) as the baseline for the comparison of the rest, the differences in haplotype distributions between the two groups did not achieve nominal significance (data not shown).

Meta-analysis results of (CAG)_n polymorphism

The initial search strategy retrieved forty-three relevant articles in English (n = 24) and Chinese (n = 19), in which the effect of MEF2A gene variations on CAD was evaluated. A total of fifteen studies met selection criteria, whereas only eight studies [13,14,17-19,31-33] were tailored to the inclusion criteria in the metaanalysis because six [15,16,20,34-36] lacked the necessary information on (CAG)_n genotypes/alleles, and one shared the same population [19,37]. Twenty-five studies, including fourteen review papers, comments and editorials, and fourteen relating to other diseases or polymorphisms in MEF2A gene, were excluded for the final analysis. The flow chart of study selection was summarized in Figure 1. Therefore, data from nine studies, including the present study, totaling 3801 CAD patients and 4020 controls were finally identified in the meta-analysis. Of these, five studies were carried out on Chinese (including this study, 61.85%) [14,18,19,31], three on Whites (36.16%) [13,17,33] and one on Turks (1.99%) [32].

The demographics and clinical features of all eligible studies are summarized in Table 2. The sample sizes ranged from 156 to 2061. The percentage of males ranged from 72.4% to 87.4% in CAD patients and 49.4% to 79.8% in controls. The mean age was greater than 56 years old in CAD patients and 51 in controls. All studies had allele data of $(CAG)_n$ polymorphism except for two with only genotype counts [14,32]. Seven studies provided information on this polymorphism associated with CAD/MI, and two with MI only [32,33]. The $(CAG)_9$ allele frequency differed widely in diverse ethnic groups. In Whites, the frequencies were in the ranges of 32.2% to 36.2% for CAD cases and 30.6% to 36.3% for controls, which were lower than that in Chinese ranging from 24.3% to 40.5% for cases and 22.7% to 41.8% for controls. In contrast, the Turks had a higher frequency of 51% for cases and 33.5% for controls (Table 2).

As shown in Figure 2, compared with other tandem repeats carriers, those with the (CAG)₉ allele yielded a non-significant 15% increased risk for CAD (95% CI = 0.97–1.37, P=0.1) under a random-effects model. Whereas, of nine studies, only two individual OR estimates showed a higher risk of CAD that was statistically significant for (CAG)₉ allele compared with other alleles (Han et al.: OR = 1.96, 95% CI = 1.55–2.47, P<0.001; Gulec et al.: OR = 2.06, 95% CI: 1.30–3.26, P= 0.002, respectively).

However, statistically significant heterogeneity was evident in most subgroups, according to covariates identified by our qualitative assessment (Table 3). In view of genotyping methods, we classified the nine studies into sequence (including this study) [13,17,19]/other (matrix-assisted laser desorption/ionization time-of-flight mass spectrometry, MALDI-TOF MS) [18] and PCR-SSCP (single-strand conformational polymorphism analysis) [14,31–33] groups. There was no significant heterogeneity in the sequence/other group (OR = 1.00, 95% CI: 0.90–1.11; $P_{heterogeneity} = 0.25$). In comparison, studies in PCR-SSCP group were heterogeneous (OR = 1.41, 95% CI: 0.94–2.12; $P_{heterogeneity}$



Figure 1. Flow chart of studies identified through the systematic literature search. doi:10.1371/journal.pone.0031406.g001

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	Cases	5	Contro	ols		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H. Random, 95% CI Year	M-H. Random, 95% Cl
Weng	197	574	217	592	11.3%	0.90 [0.71, 1.15] 2005	
González	136	422	184	602	10.7%	1.08 [0.83, 1.41] 2006	
Yuan	85	350	106	456	9.5%	1.06 [0.76, 1.47] 2006	
Han	276	756	158	696	11.5%	1.96 [1.55, 2.47] 2007	· · · · · · · · · · · · · · · · · · ·
Lieb	393	1086	865	2380	13.1%	0.99 [0.86, 1.15] 2008	· •
Gulec	70	138	58	174	7.1%	2.06 [1.30, 3.26] 2008	· · · · · · · · · · · · · · · · · · ·
Hs⊔	209	516	180	516	11.0%	1.27 [0.99, 1.64] 2009	
Dai	675	1670	254	608	12.3%	0.95 [0.78, 1.14] 2010	
This study	812	2090	809	2016	13.5%	0.95 [0.84, 1.07] 2011	4
Total (95% Cl)		7602		8040	100.0%	1.15 [0.97, 1.37]	•
Total events	2853		2831				
Heterogeneity: Tau ² =	0.05; Chi ²	= 44.0	0, df = 8 ((P < 0.0	00001); l ² :	= 82%	
Test for overall effect:	Z = 1.63 (F	P = 0.1	D)	-			U.Z U.S 1 2 5
	•						Decreased risk increased risk

Figure 2. The comparison of the MEF2A (CAG)₉ allele versus other alleles (with 4–8, 10–15 repeats) under a random effects model. doi:10.1371/journal.pone.0031406.g002

<0.00001). However, differences in *MEF2A* genotyping methods did not affect the overall results materially.

After stratification by control selection criteria based on clinical symptoms or coronary angiographic data, significant heterogeneity was observed in both the NCA group and the symptom investigation group ($P_{heterogeneity} = 0.01$ and < 0.00001, respectively). Moreover, negative associations persisted across all comparisons.

We divided the population into three groups by ethnicity, Chinese (including this study) [14,18,19,31], White (Spanish, German and Canadian) [13,17,33] and Turk [32]. Although, there was no evidence of heterogeneity in White population (OR = 0.99, 95% CI: 0.88–1.11; $P_{heterogeneity} = 0.62$), it was significant in Chinese population possibly due to the wide spectrum of (CAG)₉ allele (OR = 1.18, 95% CI: 0.90–1.54; $P_{heterogeneity} < 0.00001$). Since there was only one study performed in Turks with a relatively small sample size (n = 156), the risk estimate showed a significant higher risk of (CAG)₉ allele with CAD (P=0.002), and there was no difference in the pooled risk estimates.

To assess publication bias, we calculated the fail safe number (N_{f_0}) at the level of 0.05 for each comparison. The $N_{f_0,0.05}$ values for all the comparisons were greater (62.63) than the number of studies (n = 9) included in this meta-analysis. Therefore, no

evidence showed publication bias for association between MEF2A gene (CAG)_n polymorphism and CAD susceptibility.

Systematic review of the 21-bp deletion

Of the forty-three potentially relevant studies and the present study, fourteen were eligible for a systematic review of the 21-bp deletion and CAD risk, and thirty studies were excluded (Figure 1). Three of these were family-based studies and the remaining eleven used a hospital-based case-control design (Table 4). Of the latter, four studies (including this study) [14,19,38] had used coronary angiography as critical criteria for classification the enrollments, and four studies (including this study) involved more than 1000 subjects in controls [13,17,20]. Seven studies were conducted on Whites (53.85%) [11,13,17,20,33,34,38], six on East Asian populations (including this study, 38.46%) [14,16,18,19,39] and one on Turks (7.69%) [32]. The frequency of the 21-bp deletion in sporadic patients differed substantially, from 0.09% to 1.92%, mainly 0.16% in Whites [20,38] and 0.65% in Asian (including this study) [16,19,39], all were less than 5%. Two studies on Whites [17,20] and one on Japanese [16] confirmed the 21-bp deletion in controls (not angiographically tested), and the frequency was 0.12% and 0.51%, respectively. The overall

Table 3. Meta-analysis of the effect of (CAG)₉ allele on CAD risk according to potential sources of heterogeneity.

	Studies (cases/controls), N (n/n)	(CAG) ₉ repeats carriers	
		Overall effect (Z, OR [95%CI], P value)	Heterogeneity (<i>I</i> ² , <i>P value</i>)
Genotyping methods			
Sequence/other	5 (2636/2167)	0.07, 1.00 [0.90, 1.11], 0.94	26%, 0.25
PCR-SSCP	4 (1165/1853)	1.66, 1.41 [0.94, 2.12], 0.10	90%, <0.00001
Control selection			
Symptom investigation	5 (1368/2132)	1.24, 1.13 [0.93, 1.38], 0.21	68%, 0.01
Normal angiography	4 (2433/1888)	0.90, 1.16 [0.84, 1.61], 0.37	90%, <0.00001
Ethnic			
White	3 (1041/1787)	0.22, 0.99 [0.88, 1.11], 0.82	0%, 0.62
Chinese	5 (2691/2146)	1.22, 1.18 [0.90, 1.54], 0.22	88%, <0.00001
Turk	1 (69/87)	3.08, 2.06 [1.30, 3.26], 0.002	N/A

CAD: coronary artery disease; PCR: polymerase chain reaction; SSCP: single strand conformational polymorphism analysis. doi:10.1371/journal.pone.0031406.t003

Method Suby size Static strain Subscription 1 Ware et al. (200) Mine (wence) Subvisite Subvisite <td< th=""><th></th><th></th><th></th><th></th><th></th><th></th><th></th><th></th><th></th></td<>									
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Image: control in the second study in the s	Ref. no	Study	Ethnicity	Study design	Genotyping methods	Study size		21-bp deletion	
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17 Weng et al. (2005) Wine (canadium) cae-control study Sequence 300 permature 1821 none name indexiet, a sevent of attack 13 Gonzáleze et al. (2006) Wine (riph) cae-control study Sequence 43 Mi 1189 none in any individuals 14 Holandia (riph) Mine (riph) cae-control study Sequence 43 Mi 1189 none in any individuals 15 Hoan et al. (2006) Wine (riph) Secontrol study Sequence 43 Mi 1189 none in any individuals 16 Hoan et al. (2006) Wine (riph) Secontrol study Sequence 45 Ho Mindials 45 Ho Mindials 17 Hoan et al. (2006) Wine (riph) Secontrol study FCR-SCP 156 C.D/Mi 39 64 Paeraod make tim or Application 18 Hoat et al. (2007) Asian Northern Chinesi Cae-control study FCR-SCP 37 C.D/Mi 39 64 Paeraod make tim or Application 18 Hoat et al. (2008) Mine Chinesi Cae-control study FCR-SCP 37 C.D/Mi 36 FCR-SCP 37 C.D/Mi 19 Hoat et al. (2008) Mine Chinesi Cae-control study FCR-SCP 37 C.D/Mi 36 FCR-SCP 37 C.D/Mi 10 Edie e	16	Kajimoto et al. (2005)	Asian (Japanese)	case-control study	Sequence	379 MI	589	3, no detailed data	3, not angiographically tested
13 González et al (2006) Wite (Spanish) Gase control study Sequence 43 Mit 1189 none in any individuals 39 Hoarn et al (2006) Wite (irith) family-based study Sequence 43 Mit prioriduals from 573 families none in any individuals 39 Li et al. (2006) Wite (irith) family-based study FG+S5CP 156 CuDMin 93 a-0-year-old male; and none 30 Li et al. (2006) Main (Northern Chinese) Gase control study FC+S5CP 378 CuD 340 none in any individuals 31 Han et al. (2009) Main (Northern Chinese) Gase control study FC+S5CP 378 CuD 348 none in any individuals 32 Gue et al. (2009) Wite (Garman) Gase control study FC+S5CP 378 CuD 349 none in any individuals 33 Li be et al. (2009) Wite (Garman) Gase control study FC+S5CP 378 CuD 349 none in any individuals 34 Li be et al. (2009) Wite (Garman) Gase control study FC+S5CP 378 CuD 378	17	Weng et al. (2005)	White (Canadian)	case-control study	Sequence	300 premature CAD	1821	none	3 elderly subjects in a 71-year-old female kindred; a 45-year-old obese male with diabetes; a 45-year-old normal-weight male; all of them not angiographically tested
34 Hoan et al. (2006) White (rish) Emily-based study Sequence (LO) Mith premature (LO) One in any individuals 39 Li et al. (2006) Asian (Northern Chinese) ase-control study PC-SSC 155 G.CD/MI 93 AP-ase-old males with opsiloptement and individuals 14 Han et al. (2007) Asian (Northern Chinese) ase-control study PC-SSC 373 C.DD 348 none in any individuals 15 Han et al. (2007) Asian (Northern Chinese) ase-control study PC-SSC 373 C.DD 348 none in any individuals 16 Han et al. (2003) White (German) Rescortrol study PC-SSC 373 C.DD/M 348 none in any individuals 16 Hau et al. (2003) White (German) Rescortrol study PC-SSC 373 C.DD/M 87 None in any individuals 16 Hau et al. (2003) White (German) Rescortrol study PC-SSC 372 C.DD/M 87 None in any individuals 16 Hau et al. (2003) White (German) Rescortrol study PC-SSC 372 C.DD/M 87 None 175 C.DD/M 100 16 Hau et al. (2003)<	13	González et al. (2006)	White (Spanish)	case-control study	Sequence	483 MI	1189	none in any individuals	
39 Li et al. (2006) Aian (Northern Chinese) Gae-control study FC-SICH 15 6 CAD/MI 9. year-old male with yeigoldermia, antibly history of dipolermia, antibly history of dipolermia, antibly history of dipolermia, antibly history of dipolermia, and disberses. SO-year-old male with yeigoldermia and with yeigoldermia and disberses. SO-year-old male with yeigoldermia and disberses. SO-year-old male with yeigoldermia and disberses. SO-year-old male with yeigoldermia and with yeigoldermia and disberses. SO-year-old male with yeigoldermia and disberses. SO-year-old male with yeigoldermia and with yeigold with antibula with final with with yeigold with antibula	34	Horan et al. (2006)	White (Irish)	family-based study	Sequence	1481 individuals from 573 f with premature CAD	families	none in any individuals	
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32 Gulec et al. (2008) Turk case-control study PCR-S5CP 69 prenature MI 87 none in any individuals 33 Lieb et al. (2008) White (German) family-based study PCR-S5CP 23 representative individuals with familial MI none 18 Hsu et al. (2009) Aian (Taian) case-control study MLD-I-TOF MS 258 CAD/MI 258 none none 20 Guella et al. (2009) White (Italian) case-control study Sequence 3127 CAD/MI 3083 5 obsee males with angle and	14	Han et al. (2007)	Asian (Northern Chinese)	case-control study	PCR-SSCP	378 CAD	348	none in any individuals	
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20 Guella et al. (2009) White (Italian) case-control study Sequence 3127 CAD/MI 303 5 obese males with a 55-year-old woman, in angiographically tested 10 Dai et al. (2010) Asian (Northern Chinese) case-control study PCR-SSCP 257 CAD 154 1, no detailed data none 38 Maiolino et al. White (Italian) case-control study FRET and HRMA 1079 CAD 301 a 52-year-old male with early none 38 Maiolino et al. White (Italian) case-control study FRET and HRMA 1079 CAD 301 a 52-year-old male with early none 10 Dai at 30 on set three vessels CAD, hypertension, smoking, family history of Mi and sudden death none 10 This study Asian (Southern Chinese) case-control study Squence 1045 CAD 301 a 52-year-old male with early none 10 This study Asian (Southern Chinese) case-control study Southern Chinese none 500 57-year-old male with early none 10 This study Asian (Southern Chinese) case-control study Southern Chinese None None	18	Hsu et al. (2009)	Asian (Taiwanese)	case-control study	MALDI-TOF MS	258 CAD/MI	258	none	
19 Dai et al. (2010) Asian (Northern Chinese) case-control study PCR-SSCP 257 CAD 154 1, no detailed data none 38 Maiolino et al. White (Italian) case-control study FRET and HRMA 1079 CAD 301 a 52-year-old male with early none 38 (2011) Control study FRET and HRMA 1079 CAD 301 a 52-year-old male with early none 4 (2011) Case-control study FRET and HRMA 1079 CAD 301 a 52-year-old male with early none 7 Naiset three vessels CAD, hypertension, smoking, family hypertension, smoking, family hypertension, smoking, family none 7 This study Asian (Southern Chinese) case-control study Sequence 1045 CAD 1008 5, more details in table 3 none	20	Guella et al. (2009)	White (Italian)	case-control study	Sequence	3127 CAD/MI	3083	5 obese males with premature MI (4 with smoking, 2 with dyslipidemia)	a 55-year-old woman, not angiographically tested
38 Maiolino et al. White (Italian) case-control study FRET and HRMA 1079 CAD 301 a 52-year-old male with early none (2011) (2011) onset three vessels CAD, hypertension, smoking, family hypertension, smoking, family none This study Asian (Southern Chinese) case-control study Sequence 1045 CAD 1008 5, more details in table 3 none	19	Dai et al. (2010)	Asian (Northern Chinese)	case-control study	PCR-SSCP	257 CAD	154	1, no detailed data	none
This study Asian (Southern Chinese) case-control study Sequence 1045 CAD 1008 5, more details in table 3 none	38	Maiolino et al. (2011)	White (Italian)	case-control study	FRET and HRMA	1079 CAD	301	a 52-year-old male with early onset three vessels CAD, hypertension, smoking, family history of MI and sudden death	none
		This study	Asian (Southern Chinese)	case-control study	Sequence	1045 CAD	1008	5, more details in table 3	none

Table 4. Systematic review of the association between 21-bp deletion and CAD/MI.

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frequency of the 21-bp deletion was approximately 0.2% in the combined populations of studies published to date.

As shown in Table 4, more than half of patients bearing the deletion either suffered severe CAD who had undergone percutaneous coronary intervention or coronary artery bypass grafting, or had some traditional CAD risk factors, such as hypertension, diabetes, dyslipidemia, smoking, drinking, and/or family history of CAD/MI or sudden death. But, the results of different studies were inconsistent. Hsu et al. [18,33] failed to detect this rare variant in CAD patients. In contrast, Weng et al. [17] discovered this deletion in unaffected individuals rather than in CAD patients. However, Kajimoto et al. [16,20] reported this variant in both CAD patients and controls, and González et al. [13,14,32,34] did not reveal this deletion in any subject of the study population.

Discussion

In the present study, we verified eight variants in MEF2A exon 11 and found that the most conspicuously heterogeneous variant was the (CAG)_n polymorphism, while the other seven were all downstream of this polymorphism within 100 bp. Such intense variation in the context of a single exon led us to explore the link of MEF2A genetic polymorphisms to CAD/MI. A possible explanation might be the remarkable diversity embedded in (CAG)_n polymorphism. We therefore carried out a rigorously-designed case-control association study focusing on MEF2A exon 11 in southern Chinese and reviewed all available information regarding the relationship between this genetic hotspot and sporadic CAD/MI from the literature. To the best of our knowledge, this is the first meta-analysis seeking to clarify the association of MEF2A gene (CAG)_n polymorphism with CAD risk.

Data from our case-control study, which was in concordance with most previous observations, and in combination with all other eight studies involving a total of 3801 CAD patients and 4020 controls, failed to show a significant association between $(CAG)_n$ polymorphism and CAD susceptibility, even upon stratification by race (mainly Whites and Chinese). This meta-analysis had sufficient statistical power [40] to detect such genetic effect. Therefore, it is reasonable to surmise that either $(CAG)_n$ polymorphism itself exhibits null association with CAD, or its effect on CAD is small and depends on neighboring variants that compensate or dilute the variation under study.

Noteworthy, one study conducted by Han et al. [14] in a small northern Chinese cohort showed a positive and independent association of (CAG)₉ allele with an increased risk and severity of CAD, while data from Dai et al. [19] in another 1139 northern Chinese cohort displayed a marginal significance (P = 0.052). However, in our present study, we failed to replicate this association in southern Chinese, which was in line with two other Chinese populations [18,31] and in agreement with the pooled estimate of this meta-analysis. Except for the differences in diet and climate between northern and southern China, this discrepancy might be caused by misgenotyping as discussed by Hsu et al. [18]. All participants in Han's study were homozygous for (CAG)_n polymorphism; the phenomenon was not compatible with the situation expected from random mating. After separating analyses by genotyping methods, we found that heterogeneity between studies in the PCR-SSCP group was higher than the overall estimate; however, there was no indication attributable to the diversity between different experimental methods. Nevertheless, applying appropriate genotype techniques remain an open question.

Moreover, it is well known that CAD is frequently asymptomatic and the diagnosis relies on coronary angiography. However, the definition of controls was debatable in a number of the available studies. Some controls were enrolled according to their clinical symptoms and should have been properly defined as "uncertain phenotype" [41], whereas some were on the basis of explicit coronary angiographic results. Thus we cannot exclude the possibility that the apparently healthy elderly controls had underlying CAD, and so confuse and bias the study conclusions. Therefore, our meta-analysis pinpointed the different selective criteria of controls as a potentially significant source of betweenstudy heterogeneity. Nevertheless, deviation in the controls did not appear to be a significant source of between-study heterogeneity. Although this observation seems counterintuitive in terms of selective criteria, considering the relative small sample sizes even in the present meta-analysis and the possibility that (CAG)_n polymorphism might not be a major contributing locus or have limited values to assess an exact role of MEF2A in CAD/MI, we maintain that application of coronary angiographic criteria for controls is preferable, and the proper phenotype discrimination is critical in any genetic association study.

Meanwhile, we observed the wide divergence of $(CAG)_n$ repeats across different populations. Specifically, the high versus low frequency of $(CAG)_9$ allele was nearly double in both CAD patients and controls, suggesting a possible role of differences in genetic background and the environment in which the populations live. Of note, there was only one eligible Turkish population [32], and there was statistical evidence of heterogeneity between subgroups only in this population. It was likely that this positive association between the $(CAG)_9$ allele and CAD in the Turkish study might be due to chance or confounding, for its sample size was rather small (n = 156) and its deviation would have little or no effect on our null results. We agree that more studies in diverse ethnic/racial groups are required to draw a firm conclusion.

It was noteworthy that we identified the 21-bp deletion only in affected individuals, which was consistent with the results from Wang et al. [11,19,38,39], but which are contradicted by observations in other [13,14,16-18,20,32-34]. Considering that the 21-bp deletion was firstly identified in an exceptional CAD family displaying an autosomal-dominant pattern of inheritance and no families in the specific context were ever available for genetic linkage analysis thus far, the molecular case-control association studies of unrelated samples have become the alternative research strategy, but the results were inconsistent. Researchers have argued strongly against this deletion as a causal variant in the mechanisms of CAD pathopoiesis [41,42]. In view of possibly different genetic profiles and clinical features, we cannot jump to a conclusion regarding the cosegregation of the 21-bp deletion with CAD until validation in well-designed, large cohort studies. On the other hand, the susceptibility of patients with the 21-bp deletion to CAD supports the common-disease rare-variant (minor allele frequency less than 5%) hypothesis (CDRV) rather than the common-disease common-variant hypothesis (CDCV) [43,44]. There exists a 'common-variant, small-effect' model and also the possibility of a 'rare-variant, large-effect' model [34], the question is not which model is correct, but rather what is the relative contribution of each [43,45–48]. Although our findings add potent evidence favoring the association of the 21-bp deletion with CAD, the possibility of a founder effect from a common variant, such as (CAG)_n polymorphism, cannot be ruled out. The challenge here is to decide which observed variants in MEF2A gene could be considered as a susceptibility or causal mutation in CAD. Our data indicate that the rare 21-bp deletion might have a more compelling effect on CAD than the common (CAG)_n variant, and MEF2A genetic variants might, therefore, be a rare but specific cause of CAD/MI. Needless to say, a composite effect

encompassing all influential *MEF2A* genetic variants remains to be determined.

Finally, some limitations of this study should be acknowledged when interpreting the results. Firstly, it is recognized that differences in study design, genetic heterogeneity and statistical methods made it harder to estimate the exact underlying genetic contribution to disease susceptibility. Moreover, some large-scale studies [15,20,34] could not be included in our meta-analysis because of their incomplete raw data. These could have potentially introduced additional factors and influenced our results. Secondly, most of the enrolled study samples, including our affected population, were all survivors of CAD, as we could not evaluate those who did not survive. Thirdly, considering the complex interplay between the MEF2A gene and others that operate in the same pathway, the single-locus based case-control study and metaanalysis preclude the possibility of gene-gene and gene-environment interactions and may not reveal the full picture. Although there was no evidence showing publication bias in our overall meta-analysis, considering the above limitations, further studies with larger sample size and different ethnic compositions, which typically considered as small or moderate effects, are warranted to avoid study bias.

In conclusion, our case-control study and the following metaanalysis provide no convincing evidence for the genetic involvement of MEF2A gene (CAG)_n polymorphism in CAD. However,

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Supporting Information

Text S1 Sequence analysis.

(DOC)

Table S1The clinical characteristics of our study population.(DOC)

we suggested that the 21-bp deletion might be a rare but specific

cause of CAD. As few studies are available in this field and current

evidence remains limited, this conclusion requires further

confirmation by well-designed prospective studies with adequate

methodological quality and properly controlling for possible confounds, particularly different genetic approaches, homoge-

Table S2 Demographics of the CAD patients with 21-bpdeletion in our study population.(DOC)

Author Contributions

Conceived and designed the experiments: YL WJ. Performed the experiments: ZW XS QC. Analyzed the data: YL WN WJ. Contributed reagents/materials/analysis tools: WN ZW LL. Wrote the paper: YL WN WJ. Reference collection and data management: YL WN WJ.

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