- 1 A Claims-Based Machine Learning Classifier of Modified Rankin Scale 2 in Acute Ischemic Stroke 3 4 Mamoon Habib, MSc^a; Rafaella Cazé de Medeiros, MD^a; Syed Muhammad Ahsan, MBBS, MD 5 ^a; Aidan McDonald Wojciechowski, BS^a; Maria A. Donahue, MD^a; Deborah Blacker, MD, ScD^{b,} 6 ^d; Joseph P. Newhouse, PhD ^{c, e, f, i}; Lee H. Schwamm, MD ^g; M. Brandon Westover, MD, PhD ^h*; 7 Lidia MVR Moura, MD, PhD, MPH^a* 8 9 *co-senior authors 10 11 ^a Department of Neurology, Massachusetts General Hospital, Harvard Medical School, Boston, 12 Massachusetts 13 ^b Department of Psychiatry, Massachusetts General Hospital, Harvard Medical School, Boston, 14 Massachusetts 15 ^c Department of Health Care Policy, Harvard Medical School, Boston, Massachusetts 16 ^d Department of Epidemiology, Harvard T.H. Chan School of Public Health, Boston, 17 Massachusetts. 18 ^e Department of Health Policy and Management, Harvard T.H. Chan School of Public Health, 19 Boston, Massachusetts. 20 ^f Harvard Kennedy School, Cambridge, Massachusetts 21 ^g Yale School of Medicine, New Haven, Connecticut. 22 ⁿ Department of Neurology, Beth Israel Lahey Health Medical System, Boston, Massachusetts. 23 ¹National Bureau of Economic Research, Cambridge, Massachusetts. 24 25 Corresponding Author: 26 Lidia MVR Moura 27 lidia.moura@mgh.harvard.edu 28 617.726.3311 29 55 Fruit Street, Wang 739D
- 30 Boston, MA 02114
- 31
- 32
- 33
- 34 Word Count: 5,503

35 ABSTRACT

36 Background:

37 We developed a classifier to infer acute ischemic stroke (AIS) severity from Medicare claims

38 using the Modified Rankin Scale (mRS) at discharge. The classifier can be utilized to improve

39 stroke outcomes research and support the development of national surveillance tools.

40 Methods:

41 This was a multistate study included all participating centers in the Paul Coverdell National

42 Acute Stroke Program (PCNASP) database from nine U.S. states. PCNASP was linked to

43 Medicare data sets for patients hospitalized with AIS, employing demographics, admission

44 details, and diagnosis codes to create unique patient matches. We included Medicare

45 beneficiaries aged 65 and older who were hospitalized for an initial AIS from January 2018 to

46 December 2020. Using Lasso-penalized logistic regression, we developed and validated a

47 binary classifier for mRS outcomes and as a secondary analysis we used ordinal regression to

48 model the full mRS scale. Performance was evaluated on held-out test data using ROC AUC,

49 ROC Precision-Recall, sensitivity, and specificity.

50 **Results:** We analyzed data from 68,636 eligible patients. The mean age was 79.5 years old.

51 77.5% of beneficiaries were White, 14% were Black, 2.6% were Asian, and 2% were Hispanic.

52 The classifier achieved an ROC AUC score of 0.85 (95%CI: 0.85-0.86), sensitivity of 0.81

53 (95%CI: 0.80-0.81), specificity of 0.73 (0.72 - 0.74), and Precision-Recall AUC of 0.90 (95%CI:

54 0.90-0.91) on the test set.

55 **Conclusion:** Among Medicare beneficiaries hospitalized for AIS, the claims-based classifier

56 demonstrated excellent performance in ROC AUC, Precision-Recall AUC, sensitivity, and

57 acceptable specificity for mRS classification.

58 Key Words: Acute Ischemic Stroke (AIS), Classifier, Medicare, Modified Rankin Scale (mRS),

59 Paul Coverdell National Acute Stroke Program (PCNASP)

61

62 Clinical Perspective

- 63 What Is New?
- Developed a novel claims-based classifier to infer acute ischemic stroke (AIS) severity using
- 65 the Modified Rankin Scale (mRS) at discharge.
- Integrated Medicare claims with clinical data from the stroke registry, utilizing penalized
- 67 logistic regression for both binary and ordinal classification.
- 68 What Are the Clinical Implications?
- Provides a robust tool for assessing stroke severity, which can enhance stroke outcomes
- 70 research and quality improvement initiatives.
- Supports the development of national surveillance tools, potentially guiding clinical decision-
- 72 making and resource allocation in stroke care.

73 Research Perspective

- 74 What New Question Does This Study Raise?
- How can claims-based severity classifiers be effectively integrated into existing stroke
- 76 research and clinical practice to enhance outcome measurement?
- To what extent is the classifier generalizable to diverse populations beyond Medicare
- 78 beneficiaries?
- 79 What Question Should be Addressed Next?
- Future research should evaluate the impact of incorporating such classifiers into risk
- 81 adjustment processes and their effect on long-term stroke outcomes.
- 82 Investigate whether similar modeling approaches can be adapted for other patient groups
- 83 and healthcare settings to improve surveillance and treatment strategies.

85 INTRODUCTION

Every 40 seconds, someone in the United States (U.S.) has a stroke.¹ Stroke is one of 86 87 the leading causes of long-term disability, affecting about 795,000 people in the U.S. annually.² 88 Acute ischemic stroke (AIS) severity can be variable, with a significant portion of discharged 89 patients presenting with declining functionality, leading to increased needs for rehabilitation and admission to nursing facilities.³ Both modifiable (i.e., obesity, diabetes, cardiovascular disease, 90 91 certain medications, physical inactivity, etc.) and non-modifiable stroke risk factors (i.e., age, 92 sex, race/ethnicity, genetics) can help determine prognosis, which is crucial for early tailored intervention.4 93

Functional outcome prediction in AIS impacts the quality of patient care decisions.^{5,6} 94 95 Recent advances in computational and software technologies have greatly impacted the rise of Machine learning (ML) studies, offering more precise outcome measures.^{7–9} ML models have 96 97 identified several crucial factors to predict and classify functional outcomes, such as an initial 98 National Institutes of Health Stroke Scale (NIHSS) score, age, fasting blood glucose, creatinine levels, and the modified Rankin Scale (mRS).^{10,11} mRS has been widely used to assess AIS 99 severity and clinical prognosis in electronic health records (EHRs) and registries.¹² The creation 100 101 of models and classifiers can be personalized to assess outcomes in AIS patients, including the 102 classification of mRS.^{8,9,13} However, limited valid measures of stroke severity have hindered 103 national, large-scale, claims-based studies.¹⁴

Despite this limitation, claims data may offer indirect clues about a patient's level of disability based on the types of claims filed. Leveraging a dataset that links claims to mRS scores, we explored whether supervised ML could develop a classifier to infer mRS from claims information. Such a model could enable the personalization of outcome assessments for AIS patients and the classification of mRS in large, claims-based studies, thereby configuring a tool for national surveillance of stroke severity.

We linked the Paul Coverdell National Acute Stroke Program (PCNASP) and Medicare
claims-based inpatient data of older adults presenting with AIS to develop and validate the mRs
classifier of stroke severity at discharge.

113

114 METHODS

The Medicare data supporting this study's findings are collected routinely by The Centers for Medicare & Medicaid Services (CMS) for billing purposes and were made available by CMS with no direct identifiers. All results were aggregated following CMS Cell Suppression Policies. Restrictions apply to the availability of these data, which were used under license for this study. Medicare data are available through CMS with their permission. PCNASP data are available through the CDC with their permission.

121 This study was approved by the Mass General Brigham Institutional Review Board's 122 (IRB) ethical guidelines and followed the Strengthening the Reporting of Observational Studies 123 in Epidemiology (STROBE) guidelines for observational studies¹⁵ and the transparent reporting 124 of multivariable prediction models developed or validated using clustered data (TRIPOD)¹⁶ and 125 the updated guidance for reporting clinical prediction models that use regression or machine 126 learning methods (TRIPOD-AI).¹⁷

127

128 Study Design

We conducted a retrospective analysis of claims data from AIS patients using a sample from nine large U.S states. We aimed to develop and validate a classifier based on claims data that infers mRS at discharge.

132

133 Data Source

We accessed data from the PCNASP registry and Medicare Claims data. PCNASP
 collects data on stroke cases and captures discharge mRS scores reported by clinicians or

hospital staff.¹⁸ The PCNASP registry includes information from 2008 to 2020from the following.
states: California; Georgia; Massachusetts; Michigan; Minnesota; New York; Ohio; Washington;
and Wisconsin.

We then matched the PCNASP data on individuals aged 65 or older with data from feefor-service Medicare, a national health insurance program administered by the Centers for
Medicare & Medicaid Services (CMS).¹⁹ The Medicare Provider Analysis and Review
(MEDPAR) files contain extensive information about these beneficiaries, including patient
demographics, admission and discharge dates, diagnosis, procedure codes, provider identifiers,
and comorbidities.²⁰

145

146 **Study Population**

147 We analyzed Medicare claims data for beneficiaries aged 65 and older hospitalized for 148 AIS from January 2018 to December 2020. We included beneficiaries who were enrolled in 149 traditional Medicare Part A (inpatient hospital insurance: care in a skilled nursing facility. 150 hospice care, and some home health care) and Part B (physician and other medical provider 151 services; outpatient care, medical supplies, and preventive services) who had mRS values 152 documented in the PCNASP clinical database (based on ICD-10 code information). 153 We used a multi-step exclusion and inclusion process to refine our patient population. 154 First, we excluded patients with missing mRS scores and deceased patients in the PCNASP

data and then linked the remaining data with Medicare claims data. We found patients with a
diagnosis of AIS in the Medicare claims data during 2018-2020 and used only their first stroke
encounter. We next created two groups based on the availability of an mRS score for any stroke
(Supplemental Figure 1). The first group included patients admitted to the hospital with a ≥90%
or more completion rate of mRS, while the second group included patients admitted to hospitals
with less than <90% of mRS completion. We used 20% of the first group and all of the second

161 group as a training sample; the remaining 80% of the first group was set aside as an

- 162 independent test sample.
- 163

164 Linking Databases

Because there were no unique patient identifiers common to both databases, we applied a matching strategy to link individuals in the PCNASP and Medicare datasets.²¹ For this linkage we used variables such as age, gender, admission and discharge dates, diagnosis code, hospitals, and state. After linkage, we retained patients with unique matches, excluding cases where PCNASP IDs corresponded to multiple Medicare Beneficiary IDs and vice versa. Due to limited access to baseline institutionalized (non-outpatient) data, we excluded patients transferred from another hospital, skilled nursing facility (SNF), or other healthcare facilities.

172

173 Variables

174 We included demographic variables, medical history, treatments, and discharge 175 outcomes. Most variables were extracted from the MEDPAR files. Those not included in 176 MEDPAR were extracted from hospital level data by linking MEDPAR data with provider-level 177 data and included variables such as bed size and hospital location, category and level. We 178 included two stroke-related variables for inpatient conditions and procedures such as tissue 179 plasminogen activator (tPA) and endovascular treatment. We used the value "1" if the condition 180 or procedure was present and the value "0" if not. For continuous variables such as age and 181 length of stay, we standardized their values. Categorical variables, such as race and admission 182 type, were converted into dummy variables for use in the model. We used the variables included 183 in the Chronic Conditions Warehouse (CCW) algorithms from Medicare to determine comorbidities and relevant patient medical history in our patient population.²² CCW flagged 27 184 185 chronic conditions for each beneficiary within the study period, which we used to determine if

the beneficiary had any comorbidities. We selected the first-ever criteria a beneficiary met forthe chronic condition.

188

189 Construct of Interest (Endpoint)

190 Our primary endpoint was the accurate classification of mRS at discharge. We

191 dichotomized the mRS scale into "favorable" if valued as equal or less than 2 (from no

symptoms to slight disabilities) and "unfavorable" if the mRS score was > 2 (interval from

193 moderate disability to death).^{12,23}

As a secondary analysis, we developed ordinal classifiers using the previous sampling approach to obtain more granularity among mRS categories. The two approaches of ordinal classification consist of a full mRS scale, one represented by 0: no symptoms; 1: no significant disabilities, despite symptoms; 2: slight disabilities; 3: moderate disability; 4: moderate to severe disability; 5: severe disability; 6: death.^{23, 24} The second ordinal model consists of the same full scale but excludes the death category.

200

201 Model Development

202 *Primary analysis - Binary Classifier:* The binary classifier outputs probabilities for each
203 class. A threshold of 0.5 was used to convert the probabilities into binary values. Predictions
204 with a probability greater than or equal to 0.5 were assigned to the unfavorable mRS category,
205 and those below 0.5 to the favorable class.

For development of our binary classifier, binary logistic regression with a lasso penalty
was trained to predict the binary mRS category (favorable vs unfavorable). The best
hyperparameters were determined through a grid-search hyperparameter tuning process. The
hyperparameters included a range of the inverse regularization strength C (10⁻⁴ to 100),
tolerance values (1e-4 to 1e-1), maximum iterations (5000 to 50000), solver methods ('liblinear'

and 'saga'), and class weight settings (None and Balanced). The hyperparameters that

generated the largest area under the receiver operator characteristic curve (ROC AUC) were
chosen. Stratified 5-fold cross-validation was used to evaluate the classifier's performance
within the training set. The model was separately evaluated on the test set, which was not used
in model development.

Secondary analysis - Ordinal Classifier: We also trained a classifier on the full-scale
 mRS values using ordinal regression. The ordinal regression model outputs probabilities for
 each class. To assign class labels, we selected the class with the maximum predicted
 probability.

We fitted the model as a parallel classifier with a logit link and Lasso L1 penalty using the ordinalNet R package. Grid-search hyperparameter tuning was performed on the training dataset to select the best model based on lambda and family values. We defined a sequence of lambda values (ranging from 0.001 to 0.01) and multiple family values (cumulative, acat, sratio, cratio).

For each family type in the classifier, models were fitted across a range of lambda values and log-likelihood was used to evaluate model performance. The optimal lambda for each family type was selected as the value that achieved the highest log-likelihood, once we selected the optimal family type and lambda value, we refitted the final classifier on the training data with the chosen parameters. We tested the refitted model on the test dataset to check for its generalizability.

231

232 **Performance Metrics**

For both primary and secondary analyses, we evaluated classifier's performance using ROC AUC and Area Under the Precision-Recall Curve (PR AUC) to assess the model's ability to distinguish between classes. Sensitivity and specificity, were included to evaluate the model's ability in identifying true positives and true negatives.

To calculate confidence intervals (CI) for our performance metrics, we performed 10,000 iterations of bootstrap random sampling with replacement in each iteration. We created a distribution for each metric and calculated 95% confidence intervals to show the classifier's performance variability.

241

242 RESULTS

243 Characteristics of the samples

We assessed 295,241 hospital admissions for AIS between January 2018 and December 2020 for eligibility. After applying our inclusion and exclusion criteria, our sample included 68,636 unique Medicare beneficiaries who were 65 years old or older with a first admission for AIS and available discharge mRS scores. We obtained distinctive patient hospital encounters with < or \ge 90% completion of the mRS (N= 33,654 and N= 34,982, respectively) (Supplemental Figure 1).

250 The mean age for the full sample was 79.53 (SD 8.7), and 77.5% of beneficiaries were 251 White, 14% were Black or African American, 2.7% were Asian, and 2% were Hispanic (Table 1). 252 The mean age for our test data was 79.76 (SD 8.7). Approximately 91% of our patient sample 253 was admitted through emergency care. Regarding discharge disposition, the test set data was 254 more evenly distributed between home, SNFs, and inpatient rehabilitation facilities with 28%, 255 23%, and 19%, respectively, followed by interventions, such as receipt of tissue plasminogen 256 activator and endovascular intervention. The remaining percentage was distributed between 257 approximately 100 other discharge disposition variables. Concerning comorbidities, 71% of 258 beneficiaries had hypertension, 39% diabetes, and 29% congestive heart failure. A further 259 breakdown of the full sample, training, and test set demographics can be found in Table 1. We 260 used 63 covariates to predict a scale score, such as demographics, medical history, treatments, 261 and discharge outcomes (a list can be found in Figure 1 and Supplement Table 5).

262

263 Binary Classifier

264 On the held-out test data, our binary classifier achieved an ROC AUC score of 0.85 265 (95%CI: 0.85 – 0.86, Figure 2), sensitivity of 0.81 (95%CI: 0.80 – 0.81), specificity of 0.73 (0.72) 266 - 0.74), and Precision-Recall AUC of 0.90 (95%CI: 0.90 – 0.91, Figure 3). Figure 1 shows the 267 model's feature coefficients sorted/ranked by their contribution to its predictions. Palliative care 268 was the strongest predictor (2.02) of unfavorable mRS outcomes. Similarly, coded hemiplegia 269 (0.71), and the use of ventilator during the AIS hospitalization (0.61) were strong predictors of 270 unfavorable outcomes. Several features were also associated with a lower likelihood of 271 unfavorable outcomes. For instance, binary discharge disposition (home vs others) had the 272 strongest negative coefficient (-1.95), suggesting that favorable discharge outcomes strongly 273 predict better recovery. Transesophageal echocardiogram (-0.31), and tPA administration (-274 0.25), were associated with favorable outcomes.

275

276 Ordinal Classifier

For our secondary analysis, the ordinal model's overall performance on the test data is presented in Table 2. The model demonstrates a stronger ability to distinguish between mRS scores 0 (No Symptoms) and 5/6 (Severe Disability/Death) compared to its performance in differentiating intermediate outcomes (1–4) [see Supplementary Figure 2].

Classes 2 (Slight Disability) and 3 (Moderate Disability) showed the lowest ROC AUC and PR AUC scores. Supplementary Figure 4 presents a box plot of grouped probabilities, highlighting how the model conflates mRS scores 2 and 3 with mRS score 4. The model's ability to distinguish between mRS scores 0 (No Symptoms) and 5/6 (Severe Disability/Death) is higher compared to its performance in differentiating intermediate outcomes (1–4) [see Supplementary Figure 2].

Additionally, we excluded death to evaluate whether the model's performance improves in predicting intermediate outcomes 2 and 3, however, no significant changes in performance

were observed. The model's performance is presented in the supplementary section. The coefficients from both ordinal models (see Supplementary Tables 6 and 7) were consistent with those observed in the binary model. For instance, in the full-scale mRS ordinal model, discharge disposition [i.e., discharged home] (coefficient = 1.99) increased the odds of falling into a lower (better) mRS category, whereas palliative care (coefficient = -2.72) increased the odds of a higher (worse) category.

295

296 DISCUSSION

297 Considering the clinical burden of AIS and its influence on patient mortality, rate of 298 disability, medical complications, and healthcare expenditures, it is fundamental to monitor the 299 impact, severity, and prognosis of this condition.^{1,31,34} Our interpretation of the identified factors 300 driving the classification highlights their strong face validity and consistency with existing 301 literature as they align with clinical expectations and prior studies. Palliative care, hemiplegia, 302 endotracheal intubation, and feeding device usage were strong predictors of unfavorable mRS 303 outcomes, which is consistent with established knowledge on poor prognostic factors in acute 304 ischemic stroke. Similarly, favorable discharge disposition (e.g., discharged home), tPA 305 administration and brain imaging (CT or MRI) were associated with better outcomes, reinforcing 306 the importance of early and effective stroke management.

We developed and validated a claims-based classifier to accurately identify stroke
severity measured by mRS at discharge in patients aged 65 or older who experienced AIS. By
leveraging administrative claims data, our classifier demonstrates strong predictive performance,
achieving excellent accuracy for categorizing stroke severity. This tool holds significant potential
for facilitating large-scale research on stroke outcomes and improving national surveillance
efforts, enabling more effective monitoring of stroke care quality and recovery outcomes.
Validated claims-based classifiers for AIS surveillance are also important for observing

314 geographic trends and are essential for population health research, which in turn can inform

315 public health policy and national guidelines to improve clinical practice.³

316 Previous studies have utilized ML methods for stroke functional outcome 317 assessment.^{5,13,28} Joon Nyung Heo et al. measured mRS 90 days after hospital discharge using 318 three learning algorithm models: deep neural network, random forest, and logistic regression. 319 The study had similar results with the logistic regression model (AUC 0.85), while the best performance was by the deep neural network model (AUC 0.88)²⁸ In our study, logistic 320 321 regression for mRS classification at discharge yielded positive results with the ROC AUC score of 0.85, reiterating the results seen in other models.^{5,13,28,31} 322 323 Most importantly, the previous studies were limited by selection bias due to their 324 sampling from single regions of the US.^{5,13,28,31} Our study overcomes this challenge by including 325 a national, large-scale sample with representation of patients and practices from nine U.S. 326 states spanning all regions of the US. Therefore, our cohort provides a more robust, inclusive,

327 and representative claims-based classifier for beneficiaries with AIS than has been heretofore328 available.

329 Prior studies creating mRS stroke-severity classifiers used a random assignment approach within hospitals to create training and test sets.^{9,13} This approach is potentially biased 330 331 because random sampling does not account for hospital-level patterns in patient intake and 332 reporting. We addressed this by categorizing the training and test data sets depending on 333 whether hospitals reported $< \text{ or } \ge 90\%$ mRS completion. We only used data from those with 334 \geq 90% mRS completion as the test set, with a random 20% allocated to the training set for 335 representativeness, allowing the classifier to be trained and tested with higher-quality data and 336 partially accounting for potential bias in random sampling.

Furthermore, our study used binary and ordinal regression methods to classify the mRS
score in AIS patients. Binary analyses yield results that are more easily interpreted by
examining the absolute risk reduction between the two groups but do not exploit the within-

group variation .²³ We therefore also implemented an ordinal approach to achieve better use of
 the dataset.^{23,32} The use of the ordinal method increased statistical power and decreased loss of
 information when compared with previous studies.^{5,33}

343 Other research groups have focused on validating admission stroke severity, such as electronic health record (EHR)-based classifiers of NIHSS at admission.²⁵ This is important work, 344 345 as classifiers of stroke severity at admission can inform resource allocation while patients are 346 admitted and guide other care measures. However, we focused on leveraging claims data to 347 classify stroke functional outcomes at discharge using the mRS. The mRS is important because 348 it provides information on patient functional outcomes, which can inform the prioritization of 349 post-discharge stroke care allocation and predictions of long-term outcomes, among other 350 applications.^{26,27} The score's ability to predict the level of functionality makes it an essential tool 351 for national-level surveillance using administrative databases.⁵

352

353 Limitations

354 While we used a nationally representative stroke registry covering nine U.S. states and 355 its major stroke centers linked to administrative claims data, results may not be generalized to 356 states not included in our data set or smaller community healthcare centers. In addition, our 357 selection of older adults ≥65 covered by fee-for-service Medicare may not represent other 358 patient populations. Slightly over half of eligible Medicare beneficiaries are now enrolled in 359 Medicare Advantage "Part C" instead of traditional Medicare. Beneficiaries must also be 360 enrolled in Parts A and B, as well as Part B premium. Recent studies have shown that 361 enrollment in lower-cost Medicare Advantage plans has increased among low-income and 362 racial/ethnic minorities.³⁵ Future studies assessing these groups would benefit these 363 populations.

We excluded also 12,894 patients transferred from another hospital, skilled nursing
facility (SNF), or other healthcare facilities from the analytical sample, which may have omitted a

subset of the AIS population with a higher burden of baseline comorbidities. We selected this
approach due to limited access to predictor data from these groups. Including these patients
could have enhanced classifier representativeness and performance by increasing the sample
size and introducing greater variability. Nevertheless, our classifier demonstrated high
performance while capturing a broad and still nationally representative segment of the AIS
population.

372 We were limited by data availability for the Medicare and PCNASP datasets. While 373 utilization of administrative claims linked to data registries represents a vast source of information for research purposes,³⁶ some inherent limitations (e.g., human-type errors of 374 375 scores and clinical scales and missing data e.g., missing mRS scores and other stroke-related 376 variables) are surely present. Despite these limitations, national administrative claims data 377 remains valuable in representing large-sized populations and their reflections.^{37,38} 378 Lastly, the replicability of our classifier can present some challenges, for example, 379 requiring at least two databases to perform linkage of common unique identifiers and extract

380 multiple variables. Users looking to replicate should have experience in Python and R

381 Programming and can refer to the GitHub link in the Supplementary Material for replication.

382

383 Conclusion

We developed a claims-based classifier to identify stroke severity in AIS patients using discharge mRS. Importantly, we partially addressed potential bias by accounting for hospitallevel patterns in sampling using mRS completion rates. Our classifier has expanded on previous research by using PCNASP and Medicare-linked data from several states to assess stroke severity.

389 ACKNOWLEDGEMENTS

- 390 FUNDING
- 391 This study was supported by the NIH (1R01AG073410-01)
- 392 DISCLOSURES OF CONFLICT OF INTEREST
- 393 M.H., R.C.M., S.M.A., A.M.W., and M.A.D. have no conflict of interest to disclose.
- 394 D.B. receives support from the National Institute of Neurological Disorders and Stroke and the
- 395 National Institute on Aging and reports no conflict of interest.
- 396 J.P.N. reports no conflict of interest.
- 397 M.B.W. is a co-founder, scientific advisor, and consultant to Beacon Biosignals and has a
- 398 personal equity interest in the company.
- 399 L.M.S. receives research support from the National Institute of Neurological Disorders and
- 400 Stroke and the National Institute on Aging and reports no conflict of interest.
- 401 L.M.V.R.M. receives research support from the Epilepsy Foundation of America, the National
- 402 Institute of Neurological Disorders and Stroke, and the National Institute on Aging and reports
- 403 no conflict of interest.
- 404 Supplemental Materials
- 405 Link to GitHub Code to replicability
- 406 Figures S1-S4
- 407 Tables S1-9

409 **REFERENCES**

- Tsao CW, Aday AW, Almarzooq ZI, et al. Heart Disease and Stroke Statistics—2023
 Update: A Report From the American Heart Association. *Circulation*.
 2023;147(8):e93-e621. doi:10.1161/CIR.00000000001123
- 413 2. CDC. Stroke Facts. Stroke. October 24, 2024. Accessed November 22, 2024.
 414 https://www.cdc.gov/stroke/data-research/facts-stats/index.html
- Ziaeian B, Xu H, Matsouaka RA, et al. US Surveillance of Acute Ischemic Stroke
 Patient Characteristics, Care Quality, and Outcomes for 2019. *Stroke*.
 2022;53(11):3386-3393. doi:10.1161/STROKEAHA.122.039098
- 418
 4. Boehme AK, Esenwa C, Elkind MSV. Stroke Risk Factors, Genetics, and Prevention.
 419
 419 *Circ Res.* 2017;120(3):472-495. doi:10.1161/CIRCRESAHA.116.308398
- 420 5. Zhang MY, Mlynash M, Sainani KL, Albers GW, Lansberg MG. Ordinal Prediction
 421 Model of 90-Day Modified Rankin Scale in Ischemic Stroke. *Front Neurol.* 2021;12.
 422 doi:10.3389/fneur.2021.727171
- 423 6. Fonarow GC, Kapral MK, Schwamm LH. Future of Quality and Outcomes Research
 424 in Stroke. *Circ Cardiovasc Qual Outcomes*. 2015;8(6_suppl_3):S66-S68.
 425 doi:10.1161/CIRCOUTCOMES.115.002309
- 426 7. Fralick M, Colak E, Mamdani M. Machine Learning in Medicine. *N Engl J Med.*427 2019;380(26):2588-2590. doi:10.1056/NEJMc1906060
- Esra Zihni MSC, Bryony McGarry PHD, John Kelleher PHD. Moving Toward
 Explainable Decisions of Artificial Intelligence Models for the Prediction of Functional
 Outcomes of Ischemic Stroke Patients. *Exon Publ.* Published online April 29,
 2022:73-90. doi:10.36255/exon-publications-digital-health-explainable-decisions
- 432 9. Su PY, Wei YC, Luo H, et al. Machine Learning Models for Predicting Influential
 433 Factors of Early Outcomes in Acute Ischemic Stroke: Registry-Based Study. *JMIR*434 *Med Inform.* 2022;10(3):e32508. doi:10.2196/32508
- 435 10. Li X, Pan X, Jiang C, et al. Predicting 6-Month Unfavorable Outcome of Acute
 436 Ischemic Stroke Using Machine Learning. *Front Neurol.* 2020;11.
 437 doi:10.3389/fneur.2020.539509
- 438 11. Liu Y, Yu Y, Ouyang J, et al. Functional Outcome Prediction in Acute Ischemic
 439 Stroke Using a Fused Imaging and Clinical Deep Learning Model. *Stroke*.
 440 2023;54(9):2316-2327. doi:10.1161/STROKEAHA.123.044072
- 441 12. Banks JL, Marotta CA. Outcomes Validity and Reliability of the Modified Rankin
 442 Scale: Implications for Stroke Clinical Trials. *Stroke*. 2007;38(3):1091-1096.
 443 doi:10.1161/01.STR.0000258355.23810.c6

- 444 13. Lee J, Park KM, Park S. Interpretable machine learning for prediction of clinical
 445 outcomes in acute ischemic stroke. *Front Neurol.* 2023;14:1234046.
 446 doi:10.3389/fneur.2023.1234046
- 447 14. ElHabr AK, Katz JM, Wang J, et al. Predicting 90-day modified Rankin Scale score
 448 with discharge information in acute ischaemic stroke patients following treatment.
 449 *BMJ Neurol Open.* 2021;3(1). doi:10.1136/bmjno-2021-000177
- 450 15. von Elm E, Altman DG, Egger M, et al. The Strengthening the Reporting of
 451 Observational Studies in Epidemiology (STROBE) statement: guidelines for
 452 reporting observational studies. *Lancet Lond Engl.* 2007;370(9596):1453-1457.
 453 doi:10.1016/S0140-6736(07)61602-X
- 454 16. Debray TPA, Collins GS, Riley RD, et al. Transparent reporting of multivariable
 455 prediction models developed or validated using clustered data (TRIPOD-Cluster):
 456 explanation and elaboration. *BMJ*. 2023;380:e071058. doi:10.1136/bmj-2022457 071058
- 458 17. Collins GS, Moons KGM, Dhiman P, et al. TRIPOD+AI statement: updated guidance
 459 for reporting clinical prediction models that use regression or machine learning
 460 methods. *BMJ*. 2024;385:e078378. doi:10.1136/bmj-2023-078378
- 18. Taha M, Habib M, Lomachinsky V, et al. Evaluating the concordance between
 International Classification of Diseases, Tenth Revision Code and stroke severity as
 measured by the National Institutes of Health Stroke Scale. *BMJ Neurol Open*.
 2024;6(2). doi:10.1136/bmjno-2024-000831
- 465 19. Centers for Medicare and Medicaid Services. Centers for Medicare and Medicaid
 466 Services. Accessed January 7, 2025. https://www.cms.gov
- 467 20. Medicare Provider Analysis and Review (MedPAR). ResDAC (Research Data
 468 Assistance Center). Accessed January 7, 2025. https://resdac.org/cms469 data/files/medpar
- 470 21. Patorno E, Schneeweiss S, George MG, et al. Linking the Paul Coverdell National
 471 Acute Stroke Program to commercial claims to establish a framework for real-world
 472 longitudinal stroke research. *Stroke Vasc Neurol*. 2022;7(2):114-123.
 473 doi:10.1136/svn-2021-001134
- 474 22. Chronic Conditions Data Warehouse. Chronic Conditions Data Warehouse. July 1,
 475 2025. Accessed January 7, 2025. https://www2.ccwdata.org/web/guest/home/
- 476 23. Ganesh A, Luengo-Fernandez R, Wharton RM, Rothwell PM, on behalf of the
 477 Oxford Vascular Study. Ordinal vs dichotomous analyses of modified Rankin Scale,
- 478 5-year outcome, and cost of stroke. *Neurology*. 2018;91(21):e1951-e1960.
- 479 doi:10.1212/WNL.00000000006554

- 480 24. Stinear CM, Smith MC, Byblow WD. Prediction Tools for Stroke Rehabilitation.
 481 *Stroke*. 2019;50(11):3314-3322. doi:10.1161/STROKEAHA.119.025696
- 482 25. Brott T, Adams HP, Olinger CP, et al. Measurements of acute cerebral infarction: a
 483 clinical examination scale. *Stroke*. 1989;20(7):864-870.
 484 doi:10.1161/01.STR.20.7.864
- 485 26. Rankin J. Cerebral Vascular Accidents in Patients over the Age of 60: II. Prognosis.
 486 Scott Med J. 1957;2(5):200-215. doi:10.1177/003693305700200504
- 487 27. de Haan R, Limburg M, Bossuyt P, van der Meulen J, Aaronson N. The Clinical
 488 Meaning of Rankin 'Handicap' Grades After Stroke. *Stroke*. 1995;26(11):2027-2030.
 489 doi:10.1161/01.STR.26.11.2027
- 490 28. Heo J, Yoon JG, Park H, Kim YD, Nam HS, Heo JH. Machine Learning–Based
 491 Model for Prediction of Outcomes in Acute Stroke. *Stroke*. 2019;50(5):1263-1265.
 492 doi:10.1161/STROKEAHA.118.024293
- 493 29. Lotan E. Emerging Artificial Intelligence Imaging Applications for Stroke
 494 Interventions. *Am J Neuroradiol.* 2021;42(2):255-256. doi:10.3174/ajnr.A6902
- 30. Mouridsen K, Thurner P, Zaharchuk G. Artificial Intelligence Applications in Stroke. *Stroke*. 2020;51(8):2573-2579. doi:10.1161/STROKEAHA.119.027479
- 497 31. Daidone M, Ferrantelli S, Tuttolomondo A. Machine learning applications in stroke
 498 medicine: advancements, challenges, and future prospectives. *Neural Regen Res.*499 2024;19(4):769. doi:10.4103/1673-5374.382228
- 32. Risselada R, Lingsma HF, Molyneux AJ, et al. Prediction of two month modified
 Rankin Scale with an ordinal prediction model in patients with aneurysmal
 subarachnoid haemorrhage. *BMC Med Res Methodol.* 2010;10(1):86.
 doi:10.1186/1471-2288-10-86
- 33. Roozenbeek B, Lingsma HF, Perel P, et al. The added value of ordinal analysis in
 clinical trials: an example in traumatic brain injury. *Crit Care*. 2011;15(3):R127.
 doi:10.1186/cc10240
- 34. Ovbiagele B, Goldstein LB, Higashida RT, et al. Forecasting the future of stroke in
 the United States: a policy statement from the American Heart Association and
 American Stroke Association. *Stroke*. 2013;44(8):2361-2375.
 doi:10.1161/STR.0b013e31829734f2
- 511 35. Meyers DJ, Mor V, Rahman M, Trivedi AN. Growth In Medicare Advantage Greatest
 512 Among Black And Hispanic Enrollees. *Health Aff (Millwood)*. 2021;40(6):945-950.
 513 doi:10.1377/hlthaff.2021.00118

- 514 36. Fernandes M, Cardall A, Jing J, et al. Identification of patients with epilepsy using
 515 automated electronic health records phenotyping. *Epilepsia*. 2023;64(6):1472-1481.
 516 doi:10.1111/epi.17589
- 517 37. Schneeweiss S, Avorn J. A review of uses of health care utilization databases for
 518 epidemiologic research on therapeutics. *J Clin Epidemiol*. 2005;58(4):323-337.
 519 doi:10.1016/j.jclinepi.2004.10.012
- 520 38. MacKay EJ, Stubna MD, Chivers C, et al. Application of machine learning
- 521 approaches to administrative claims data to predict clinical outcomes in medical
- 522 and surgical patient populations. *PLoS ONE*. 2021;16(6):e0252585.
- 523 doi:10.1371/journal.pone.0252585

525 TABLES AND FIGURES

526 Table 1. Demographic Characteristics

Characteristics	Full Sample (N = 68,636)	Training /	Tost	
		Validation	(n = 27,986)	
		(n = 40,650)		
Age, mean (SD)	79.53 (8.67)	79.38 (8.63)	79.76 (8.71)	
Gender (%)				
Female	37,439 (54.54)	22,045 (54.23)	15,394 (55.00)	
Male	31,197 (45.45)	18,605 (45.76)	12,592 (45.00)	
Race (%)				
White	53,192 (77.49)	31,794 (78.21)	21,398 (76.45)	
Black	9,629 (14.02)	5,394 (13.26)	4,235 (15.13)	
Asian	1,821 (2.65)	1,146 (2.81)	675 (2.41)	
Hispanic	1,361 (1.98)	753 (1.85)	608 (2.17)	
Other	1,483 (2.16)	876 (2.15)	607 (2.16)	
Unknown	997 (1.45)	593 (1.45)	404 (1.44)	
North American Native	153 (0.22)	94 (0.23)	59 (0.21)	
Admission Type (%)				
Emergency	62,639 (91.26)	36,657 (90.18)	25,982 (92.83)	
Urgently	4,911 (7.15)	3,375 (8.30)	1,536 (5.48)	
Trauma Center	559 (0.81)	326 (0.80)	233 (0.83)	
Intensive Care Unit (ICU) Type (%)				
Intermediate IOCU	13,325 (19.41)	8,379 (20.61)	4,946 (17.6)	
General	11,569 (16.85)	6,786 (16.69)	4,783 (17.09)	
Medical	3,033 (4.41)	1,599 (3.93)	1,434 (5.12)	

Surgical	1,501 (2.1	3) 1,073 (2.63)	428 (1.52)
Trauma	153 (0.22)	117 (0.28)	36 (0.12)
Other	144 (0.20)	63 (0.15)	81 (0.28)
Discharge Disposition (%)*		
Home/Self-care	18,931 (27	.58) 11,233 (27.6	3) 7,698 (27.50)
Skilled Nursing Facility	15,426 (22	47) 9,056 (22.27)) 6,370 (22.76)
Inpatient Rehabilitation	Facility 12,856 (18	.73) 7,213 (18.67)) 5,266 (18.81)
Interventions (%)			
Tissue Plasminogen Ac	tivator 9,001 (13.	11) 5,579 (13.72)) 3,422 (12.22)
Endovascular Intervent	ion 3,089 (4.50) 1,780 (4.37)	1,309 (4.67)
Comorbidities (%)			
Acute Myocardial Infarc	tion 4,290 (6.2	5) 2,544 (6.26)	1,746 (6.24)
Atrial Fibrillation	13,304 (19	.38) 7,700 (18.94)) 5,604 (20.02)
Diabetes	26,708 (38	.91) 15,581 (38.3	3) 11,127 (39.76)
Congestive Heart Failu	re 19,766 (28	.80) 11,555 (28.4	3) 8,211 (29.34)
Hypertension	48,451 (70	.59) 28,418 (69.9	1) 20,033 (71.58)

527 Legend: Baseline demographics, admission type, Intensive Care Unit (ICU) Type, and

528 comorbidities stratified by sample, training, and test groups.

- 529 *We did not include all discharge disposition variable in the table, as there are over 100 existing
- 530 items. We reported the most relevant ones to this table.

531 Table 2. Full-Scale Ordinal Model Performance

Metric	Score [CI]
ROC AUC	0.81 [0.80 – 0.81]
Precision-Recall AUC	0.39 [0.37 – 0.39]
Sensitivity	0.42 [0.41-0.2]
Specificity	0.89 [0.88 - 0.89]

532 Legend: Performance Metrics from Full-Scale Ordinal. We report micro-average ROC

533 AUC and Precision-Recall AUC.

534 Figure 1. Model Features



535

536 **Legend:** The full list of the classifier's features and their coefficient values. COPD: Chronic

537 obstructive pulmonary disease; ICU: Intensive Care Unit; tPA, tissue plasminogen activator; CT,

- 538 computed tomography; MRI, Magnetic resonance imaging.
- 539

540 Figure 2. ROC (Receiver Operating Characteristic) Curve



541

542 **Legend:** Comparison of the ROC in both the training and test sets of the classifier.

544 Figure 3. Precision-Recall Area Under the Curve for Binary Classifier



545

546 **Legend:** Comparison of Precision-Recall Curve of the classifier in the training and test sets.