ORIGINAL ARTICLE

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Different therapeutic associations of renin-angiotensin system inhibitors with coronavirus disease 2019 compared with usual pneumonia

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¹Korean Society of Hypertension; ²Department of Internal Medicine, Seoul National University College of Medicine, Seoul; ³Department of Internal Medicine, Seoul National University Hospital, Seoul; ⁴Department of Public Health Sciences, ⁵Institute of Health and Environment, Seoul National University, Seoul; ⁶National Health Insurance Service, Wonju; ⁷Department of Benefits Strategy, National Health Insurance Service, Wonju; ⁸National Medical Center, Seoul; ⁹National Committee for Clinical Management of Emerging Infectious Diseases, Seoul; ¹⁰The Central Infectious Disease Hospital, Seoul; ¹¹Division of Cardiology, Department of Internal Medicine, Ewha Womans University School of Medicine, Seoul, Korea

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Korean Society of Hypertension and Division of Cardiology, Department of Internal Medicine, Ewha Womans University School of Medicine, 1071 Anyangcheonro, Yangcheon-gu, Seoul 07985, Korea TEL: +82-2-6986-1627 FAX: +82-2-2650-6166 E-mail: pwb423@ewha.ac.kr https://orcid.org/0000-0002-6377-0411 **Background/Aims:** Although it is near concluded that renin-angiotensin system inhibitors do not have a harmful effect on coronavirus disease 2019 (COVID-19), there is no report about whether angiotensin receptor blockers (ARBs) and angiotensin-converting enzyme inhibitors (ACEIs) offer any protective role. This study aimed to compare the association of ARBs and ACEIs with COVID-19-related mortality.

Methods: All patients with COVID-19 in Korea between January 19 and April 16, 2020 were enrolled. The association of ARBs and ACEIs with mortality within 60 days were evaluated. A comparison of hazard ratio (HR) was performed between COVID-19 patients and a retrospective cohort of pneumonia patients hospitalized in 2019 in Korea.

Results: Among 10,448 COVID-19 patients, ARBs and ACEIs were prescribed in 1,231 (11.7%) and 57 (0.6%) patients, respectively. After adjusting for age, sex, and history of comorbidities, the ARB group showed neutral association (HR, 1.034; 95% CI, 0.765 to 1.399; p = 0.8270) and the ACEI groups showed no significant associations likely owing to the small population size (HR, 0.736; 95% CI, 0.314 to 1.726; p = 0.4810). When comparing HR between COVID-19 patients and a retrospective cohort of patients hospitalized with pneumonia in 2019, the trend of ACEIs showed similar benefits, whereas the protective effect of ARBs observed in the retrospective cohort was absent in COVID-19 patients. Meta-analyses showed significant positive correlation with survival of ACEIs, whereas a neutral association between ARBs and mortality.

Conclusions: Although ARBs or ACEIs were not associated with fatal outcomes, potential beneficial effects of ARBs observed in pneumonia were attenuated in COVID-19.

Keywords: COVID-19; Pneumonia; Mortality; Angiotensin receptor antagonists; Angiotensin-converting enzyme inhibitors

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INTRODUCTION

When coronavirus disease 2019 (COVID-19, severe acute respiratory syndrome coronavirus 2 [SARS-CoV-2]) emerged as a global health problem, initial epidemiologic studies reported high prevalence of hypertension, diabetes mellitus and coronary artery disease [1,2]. These chronic diseases are highly prevalent among COVID-19 patients in Korea [3]. Given that SARS-CoV-2 enters target cells via angiotensin-converting enzyme 2 (ACE2) cell surface receptors [4], the use of angiotensin receptor blockers (ARBs) and angiotensin-converting enzyme inhibitors (ACEIs), which reportedly augment ACE2 levels [5], raise concern regarding antihypertensive agents, especially renin-angiotensin-aldosterone system (RAAS) inhibitors [6]. Although it is near concluded that ACEIs and ARBs do not increase mortality of COVID-19 patients [7-9], there are no reports as to whether ARBs or ACEIs have any protective role, which have been repetitively reported in pneumonia-related outcomes [10-12]. Given that beneficial effects have been reproducibly reported in pneumonia patients, neutral results suggest that there might be possible offset effects.

In the present study, we compared the association between the use of ARBs or ACEIs and COVID-19-related mortality in Korea. The hazard ratio (HR) was compared between COVID-19 patients and a retrospective cohort of patients hospitalized with pneumonia between January and June 2019 in Korea. Lastly, a meta-analysis was performed to compare the results of this study to those of other reports.

METHODS

Study design

This was a population-based cohort study supported by the Korea Disease Control and Prevention Agency (KDCA), the National Health Insurance Service (NHIS), and the Korean Society of Hypertension, and was approved by the Institutional Review Board of Seoul National University Hospital (No. 2003-102-1109). Informed consent was waived because the study was based on routinely collected administrative and claims data. All authors reviewed the manuscript for accuracy and completeness of the data.

Data collection

Data related to all 10,448 patients, who were infected with laboratory-confirmed SARS-CoV-2 until April 16, 2020, in Korea [13,14], were retrieved from compiled information available from the NHIS. Information related to medical utilization was extracted from the National Health Information Database (NHID), a public database formed by the NHIS [15].

For indirect comparative analyses, the association between the use of ARBs and/or ACEIs and the mortality of pneumonia in hospitalized Korean patients between January and June 2019 was investigated using the information retrieved from the NHID [15].

Study outcomes and definitions

The association of ARBs and ACEIs with COVID-19-related mortality within 60 days was investigated. Disease-related mortality was defined as mortality during hospitalization with the indicated disease (COVID-19 or pneumonia) as the main disease code of hospitalization. All commercially available ARBs and ACEIs in Korea were included in the analysis. Medications were classified under ARBs (candesartan, irbesartan, valsartan, losartan, telmisartan, eprosartan, fimasartan, azilsartan, and olmesartan) and ACEIs (captopril, enalapril, lisinopril, perindopril, ramipril, and zofenopril). Dichotomous variables were created to identify the use of ARBs and ACEIs defined as a filled prescription for the medication of interest with sufficient supply to overlap the date of admission, assuming 80% compliance [16].

Statistical analysis

The association of clinical characteristics with mortality within 60 days was tested using bivariate statistics. Categorical variables were analyzed with a chi-square test. Continuous variables were analyzed using an analysis of variance test. Time to occurrence of the COVID-19-related death in patients prescribed with ARBs or ACEIs was analyzed using Kaplan-Meier graphs displaying failure functions. Statistical significance was assessed using a log-rank test. Cox proportional hazard models calculated estimates of unadjusted HRs and 95% confidence intervals (CIs) for ARBs or ACEIs. The additional factors included in the models as covariates were age group, sex, and comorbidities including hypertension, type 2 diabetes mellitus, coronary heart diseases, heart failure,



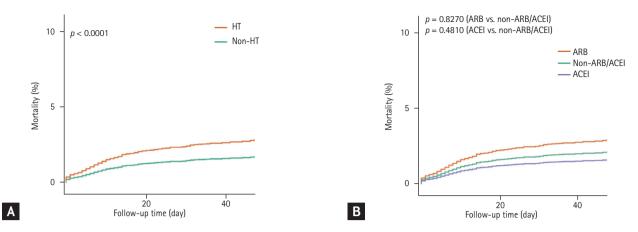


Figure 1. Survival within 60 days with prior statins or angiotensin receptor blockers (ARBs)/angiotensin-converting enzyme inhibitors (ACEIs) received. (A) Comparison with hypertension (HT) vs. non-HT (adjusted for age, sex, and history of comorbidities before the laboratory-confirmed diagnosis of coronavirus disease 2019). (B) Comparison with the patients treated with ARBs vs. ACEIs vs. the patients with non-ARBs/ACEIs (adjusted).

stroke, chronic obstructive pulmonary disorder (COPD), cancer, and chronic kidney diseases. We diagnosed the proportional hazard assumption using the Schoenfeld residuals plot. Comparisons were considered statistically significant if two-tailed *p* values were less than 0.05. All analyses were performed using SAS version 7.15 (SAS Institute Inc., Cary, NC, USA) and R version 4.0.0 (The R Development Core Team, Vienna, Austria).

Meta-analysis

Using PubMed searches of the MEDLINE database, we identified original papers published from 2019 to September 2020 in order to evaluate the association of ARBs or ACEIs with COVID-19 related death. The search strategy was based on the search terms "ARBs" or "ARB" or "angiotensin receptor blockers," or "ACEIs" or "ACEI" or "angiotensin converting enzyme inhibitors" or "angiotensin-converting enzyme inhibitors," and COVID-19 or COVID19 or coronavirus, and death or mortality. All available English abstracts were reviewed, and the full text was consulted as necessary to clarify eligibility status. We excluded the review articles, editorials and original papers that did not evaluate mortality (infection or hospitalization). Also, we performed an online search under the following terms: "COVID-19 and hypertension and ARBs or ACEIs." The initial search identified 259 articles. Of these, 25 articles were included for meta-analysis [17-33]. Details of the search strategy are summarized in Supplementary Fig. 1. The authors (H.Y.L. and J.A.) reviewed all abstracts independently to evaluate the eligibility criteria and appropriateness of the research topics. If the inclusion criteria were met, the article was retrieved and reviewed thoroughly. There were no discrepancies in this process.

Statistical calculations and graphs were made using Rex software version 3.3.1 (RexSoft Inc., Seoul, Korea) [34]. Two-tailed statistical significance was set at the 5% level, except for the Cochran's chi-square test for heterogeneity, which used a 10% level of significance. The pooled results for each outcome are presented as odds ratios (OR) with 95% CIs. Before applying approximate chi-square tests for heterogeneity, we clinically assessed studies for heterogeneity. Statistical heterogeneity was also examined with the I^2 statistics, where I^2 values \geq 50% were considered to be indicators of a substantial level of heterogeneity. Forest plots were used for visual inspection. Funnel plots of effect estimates against its standard error and Copas selection model analysis were conducted to assess publication bias. We combined p values for each OR and HR from meta-analysis using Fisher and Liptak's methods [35].

RESULTS

The present study includes data related to a total of 10,448 COVID-19 patients who were hospitalized or isolated in Korea from January 19, 2020, through April 16,



Characteristic	Total (n = 10,448)	Non-ARB/ACEI (n = 9,170)	ARB (n = 1,221)	ACEI (n = 57)	p value
Age, yr	44.87 ± 19.81	41.97 ± 18.83	65.50 ± 13.01	70.03 ± 12.30	< 0.0001
0–19	503 (4.81)	503 (5.49)	0	0	< 0.0001
20-59	7,293 (69.80)	6,879 (75.02)	403 (33.01)	11 (19.30)	
60–79	2,124 (20.33)	1,476 (16.10)	616 (50.45)	32 (56.14)	
≥80	528 (5.05)	312 (3.40)	202 (16.54)	14 (24.56)	
Female sex	6,264 (59.95)	5,547 (60.49)	697 (57.08)	20 (35.09)	< 0.0001
Comorbidity					
Any \geq comorbidity	3,979 (38.08)	2,720 (29.66)	1,202 (98.44)	57 (100.0)	< 0.0001
Hypertension	2,149 (20.57)	899 (9.80)	1,195 (97.87)	55 (96.49)	< 0.0001
Diabetes mellitus	1,874 (17.94)	1,219 (13.29)	613 (50.20)	42 (73.68)	< 0.0001
Coronary artery disease	633 (6.06)	358 (3.90)	240 (19.66)	35 (61.40)	< 0.0001
Heart failure	345 (3.30)	188 (2.05)	133 (10.89)	24 (42.11)	< 0.0001
Stroke	393 (3.76)	209 (2.28)	169 (13.84)	15 (26.32)	< 0.0001
COPD	1,487 (14.23)	1,161 (12.66)	306 (25.06)	20 (35.09)	< 0.0001
Cancer (any)	470 (4.50)	348 (3.79)	110 (9.01)	12 (21.05)	< 0.0001
Chronic kidney disease	112 (1.07)	50 (0.55)	59 (4.83)	3 (5.26)	< 0.0001
Medications					
ARBs	1,221 (11.69)	-	1,221 (100)	-	-
ACEIs	57 (0.55)	-	-	57 (100)	< 0.0001
Death within 60 days	228 (2.18)	127 (1.38)	95 (7.78)	6 (10.53)	< 0.0001

Table 1. Clinical characteristics of coronavirus disease 2019 patients

Values are presented as mean ± standard deviation or number (%).

ARB, angiotensin receptor blocker; ACEI, angiotensin-converting enzyme inhibitor; COPD, chronic obstructive pulmonary disorder.

2020. As of April 24, 2020, 228 of these patients (2.18%) succumbed to death.

The demographic and clinical characteristics of patients are summarized in Table 1. Of the patients evaluated, 38.1% had at least one preexisting comorbid condition including hypertension (20.6%) and diabetes mellitus (17.9%). Patients prescribed with ARBs or ACEIs were more than 20 years older with a higher number of comorbidities than non-users and accounted for 11.7% and 0.6% of the overall cases, respectively. A small portion of the ACEI group corresponded to the prescription pattern reported in the Korean hypertension fact sheet [36]. Between the patients prescribed ARBs or ACEIs, the ACEI group showed a higher proportion of males (64.9% vs. 40% in overall COVID-19 patients) and more comorbidities, including diabetes mellitus and cardiovascular disease. The mean age of hypertensive patients $(65.7 \pm 13.0 \text{ years})$ was more than 20 years greater than

that of the non-hypertensive patients (42.0 \pm 18.8 years, *p* < 0.0001).

The crude HR of old age (> 65 years) and hypertension were 36.5 (95% CI, 24.6 to 54.2; p < 0.0001) and 14.1 (95% CI, 10.3 to 19.2; p < 0.0001), respectively (Table 2). In multivariable regression analysis, old age (> 65 years) was by far the most important predictor of COVID-19-related mortality (Table 3). After adjusting for age, sex, and history of comorbidities, the association of chronic comorbid conditions including hypertension and diabetes mellitus remained significant (Fig. 1A).

In Cox regression analysis, the ARB group showed neutral association (HR, 1.034; 95% CI, 0.765 to 1.399; p = 0.8270) and the ACEI group showed no significant associations likely owing to the small population size (HR, 0.736; 95% CI, 0.314 to 1.726; p = 0.4810) (Fig. 1B). As there is a possibility that specific ARBs have a superior impact beyond class effect [5], we evaluated the subgroup anal-



Table 2. Results of univariable regression analysis for mortality within 60 days

-	-				
Variable	No. patients	No. death	HR	95% CI	p value
Age, yr					
0–64	8,636	28	1		
≥ 65	1,812	200	36.524	24.590-54.249	< 0.0001
Sex					
Male	4,184	121	1		
Female	6,264	107	0.578	0.446–0.750	< 0.0001
Comorbidity (upper, no; lower, yes)					
Hypertension	8,299	51	1		
	2,149	177	14.093	10.321–19.245	< 0.0001
Diabetes mellitus	8,574	95	1		
	1,874	133	6.628	5.093-8.625	< 0.0001
Coronary artery disease	9,815	173	1		
	633	55	5.129	3.787-6.948	< 0.0001
Stroke	10,055	182	1		
	393	46	7.020	5.078-9.703	< 0.0001
COPD	8,961	136	1		
	1,487	92	4.142	3.179-5.397	< 0.0001
Cancer (any)	9,978	196	1		
	470	32	3.581	2.465–5.204	< 0.0001
Chronic kidney disease	10,336	210	1		
	112	18	8.447	5.220–13.671	< 0.0001
Heart failure	10,103	175	1		
	345	53	19.772	7.186–13.290	< 0.0001
Medication					
Non-ACEIs/ARBs	9,170	127	1		
ARBs	1,221	95	5.809	4-453-7-579	< 0.0001
ACEIs	57	6	8.174	3.606–18.529	< 0.0001
Non-ACEIs/ARBs	9,170	127	1		
ARBs or ACEIs	1,278	101	5.911	4.551-7.676	< 0.0001

HR, hazard ratio; CI, confidence interval; COPD, chronic obstructive pulmonary disorder. ACEI, angiotensin-converting enzyme inhibitor; ARB, angiotensin receptor blocker.

ysis according to individual drugs. Expectedly, no statistical significance was observed likely due to the small cohort sizes for each drug. No prevailing therapeutic effect was observed among the tested drugs (Supplementary Fig. 2).

The HR was compared between COVID-19 patients and a retrospective cohort of patients hospitalized with pneumonia between January and June 2019. Details related to comorbid conditions and administered medication were similarly extracted (Supplementary Table 1). Because there was no viral pneumonia epidemic in Korea, like SARS or Middle East respiratory syndrome (MERS), the 2019 pneumonia cohort was composed of bacterial and viral pneumonia patients combined. This retrospective cohort included 179,265 patients with a mean age of 39 years, of which 49.4% were females. The most common comorbid condition in the cohort was hypertension (33.3%) followed by COPD (25.1%) and diabetes mellitus (23.4%). The mortality rate at 60 days was 7.4% and sharply increased among elderly patients (≥ 60



5	1		1	1	
Variable	No. patients	No. death	HR	95% CI	<i>p</i> value
Age, yr					
0–64	8,636	28	1		
≥ 65	1,812	200	16.400	10.487–25.646	< 0.0001
Sex					
Male	4,184	121	1		
Female	6,264	107	0.556	0.427-0.723	< 0.0001
Comorbidity (upper, no; lower, yes)					
Hypertension	8,299	51	1		
	2,149	177	2.423	1.621–3.623	< 0.0001
Diabetes mellitus	8,574	95	1		
	1,874	133	1.495	1.118–2.000	0.0128
Coronary artery disease	9,815	173	1		
	633	55	0.860	0.613–1.206	0.3815
Stroke	10,055	182	1		
	393	46	1.242	0.887–1.739	0.2080
COPD	8,961	136	1		
	1,487	92	1.509	1.141–1.996	0.0040
Cancer (any)	9,978	196	1		
	470	32	1.096	0.743–1.617	0.6438
Chronic kidney disease	10,336	210	1		
	112	18	1.709	1.038–2.814	0.0352
Heart failure	10,103	175	1		
	345	53	1.826	1.294-2.577	0.0006
Medication					
Non-ACEIs/ARBs	9,170	127	1		
ARBs	1,221	95	1.034	0.765–1.399	0.8270
ACEIs	57	6	0.736	0.314–1.726	0.4810
Non-ACEIs/ARBs	9,170	127	1		
ARBs or ACEIs	1,278	101	1.011	0.750–1.361	0.9446

Table 3. Results of multi	variable regression	analysis of ARBs	and ACEIs for the n	ortality within 60 days
J		1		1

Adjusted for age, sex, history of medication (statins in case of ARBs/ACEIs), and history of comorbidities (hypertension, diabetes mellitus, cancer, COPD, stroke, coronary artery disease, heart failure, and chronic kidney disease) before the diagnosis of coronavirus disease 2019.

ARB, angiotensin receptor blocker; ACEI, angiotensin-converting enzyme inhibitor; HR, hazard ratio; CI, confidence interval; COPD, chronic obstructive pulmonary disorder.

years) to a value higher than that seen with COVID-19 (Fig. 2A). Furthermore, the mortality was higher in patients with comorbidities, including hypertension (Supplementary Tables 2 and 3). Upon consideration of secondary factors, the results of the present study indicated a significantly lower risk of 60 days mortality with the use of ARBs (HR, 0.853; 95% CI, 0.820 to 0.887; *p* < 0.0001) (Supplementary Table 3). Although ACEI groups showed an increase in HR, there was no statistical significance (HR, 1.089; 95% CI, 0.820 to 1.203; p = 0.0987). When ARB and ACEI groups were merged, they showed a significant decrease in the mortality by about 14% (HR, 0.866; 95% CI, 0.834 to 0.900; p < 0.0001). These results are in agreement with previously published data [10-12,16,37].



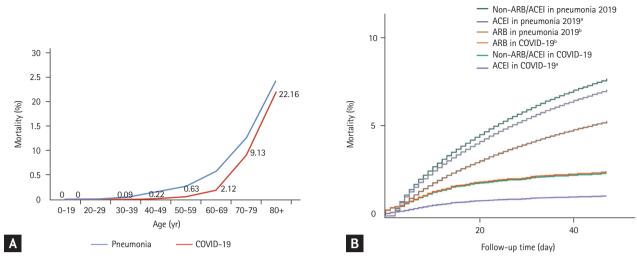


Figure 2. Survival within 60 days in coronavirus disease 2019 (COVID-19) patients and retrospective cohort of patients hospitalized with pneumonia between January and June 2019. (A) Mortality rate according to age groups. (B) Comparison between patients treated with angiotensin receptor blockers (ARBs) or angiotensin-converting enzyme inhibitors (ACEIs) vs. patients with non-ARBs/ACEIs (adjusted). ^ap = 0.1319, ^bp = 0.0040 (pneumonia 2019 vs. COVID-19).

We compared HR between COVID-19 patients and a retrospective cohort of patients hospitalized with pneumonia 2019 in Korea. The trend of ACEIs was similar to that in COVID-19 patients (p = 0.1319). However, the HRs of ARBs significantly differed between the retrospective cohort and the COVID-19 patients (p = 0.0040). The protective effect of ARBs observed in the retrospective cohort was lost in the COVID-19 patients (Table 3, Fig. 2B).

Lastly, we performed meta-analyses of ARB and ACEI groups using the data from previous studies. The 25 eligible studies and our study included 190,135 COVID-19 patients (Supplementary Fig. 1). Meta-analysis of ACEIs showed significant decrease in relative risk by 18% to 25% (p = 0.028 [Fisher], p = 0.017 [Liptak]) (Fig. 3A). However, the ARB group showed neutral association (relative risk, 0.96 to 0.99; p = 0.915 [Fisher], p = 0.886 [Liptak]) (Fig. 3B). Combined p values for ORs and HRs are robust through Fisher and Liptak's methods. Interestingly, combined p values in ARBs have the same trend with our results, whereas those for ACEIs show a significant protective effect for all-cause mortality (Table 4). We did not detect any significant publication bias by visual inspection of the funnel plots.

DISCUSSION

In the present study, we showed that ARBs and ACEIs are

not associated with COVID-19-related fatal outcomes. While the ARB group showed a neutral association, the ACEI group showed no significant results, likely due to the small cohort size.

The RAAS is activated in sepsis, and angiotensin II, the main effector of the RAAS, has been recognized as an important inflammatory agent associated with organ failure and mortality [38]. Therefore, there is a competing hypothesis that ARBs and ACEIs may provide protection against pneumonia-related fatality. Enhanced expression of ACE and angiotensin II was reported in patients with acute respiratory distress syndrome and mechanical ventilation [30-41]. Injurious ventilation enhances the expression of the angiotensin type 1 receptor, a potent promoter of inflammatory response [40]. Angiotensin II, a pro-inflammatory mediator, induces damage to the pulmonary and vascular endothelial cells [42,43]. ARBs and ACEIs are reported to protect the vascular endothelium [44-46] and to possess immunomodulatory effects [47,48]. In vivo, angiotensin II deletion and generation of anti-inflammatory Ang 1-7 revealed the potential benefits of ACE2 in H7N9 virus-induced acute lung injury [49].

In the initial reports of COVID-19, it was hypothesized that the use of ARBs or ACEIs may alter ACE2, which is a cell receptor of SARS-CoV-2, hence enhancing the infectivity of SARS-CoV-2. However, recent studies commonly reported that renin-angiotensin sys-

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ACEIs						Weight	Weight
Study	TE s	seTE	Odds Ratio	OR	95%-CI	(fixed)	(random)
Bauer et al (2020)	0.20 0.	1784			[0.86; 1.73]	12.4%	10.5%
Bean et al (2020)	-1.24 0.				[0.10; 0.84]	1.3%	3.3%
Bravi et al (2020)	-0.20 0.1				[0.49; 1.37]	5.7%	8.0%
Dalan et al (2020)	-0.69 0.				[0.08; 3.12]	0.5%	1.3%
de Abajo et al (2020)	-0.08 0.	1772			[0.65; 1.30]	12.6%	10.5%
Felice et al (2020)	-0.39 0.4				[0.28; 1.68]	1.9%	4.2%
Feng et al (2020)	-0.92 1.	0712 -			[0.05; 3.27]	0.3%	1.0%
Li et al (2020)	-0.08 0.4				[0.38; 2.23]	1.9%	4.3%
Liu et al (2020)	-0.56 0.				[0.14; 2.35]	0.8%	2.1%
Mancia et al (2020)	-0.09 0.		*		[0.69; 1.20]	19.8%	11.6%
Mehra et al (2020, Retracted			- -		[0.20; 0.54]	6.0%	8.2%
Mehta et al (2020)	0.30 0.				[0.74; 2.46]	4.2%	6.9%
Meng et al (2020)	-1.11 0.				[0.09; 1.21]	0.9%	2.4%
Peng et al (2020)	-0.06 0.				[0.24; 3.68]	0.8%	2.2%
Rentsch et al (2020)	0.52 0.		<u>š</u> — — —		[1.01; 2.83]	5.7%	8.0%
Reynolds (2020)	-0.11 0.				[0.70; 1.16]		12.0%
Yang et el (2020)	-1.14 0.				[0.07; 1.46]	0.7%	1.8%
Zeng et al (2020)	-0.43 0.	8620		0.65	[0.12; 3.52]	0.5%	1.5%
Fixed effect model				0.89	[0.78; 1.00]	100.0%	
Random effects model			4		[0.66; 1.02]		100.0%
Prediction interval					[0.42; 1.60]		
Heterogeneity: $I^2 = 53\%$, $\tau^2 = 0$	0.0882, p <	0.01					
			0.1 0.5 1 2 10			Weight	Weight
Study	TE	seTE	Hazard Ratio	HR	95%-CI		(random)
Fosbol et al (2020)	-0.02 0	1644	i <u></u>	0.98	[0.71; 1.35]	6.0%	17.4%
Khera et al (2020)	-0.03 0				[0.81; 1.16]		
Mehra et al (2020, Retractre					[0.51; 0.62]		
Rossi et al (2020)	-0.03 0				[0.69; 1.36]		
Zhang et al (2020)	-0.87 0				[0.19; 0.93]		
Zhou et al (2020)	-0.71 0				[0.20; 1.20]		
Lee et al (2020)	-0.31 0				[0.31; 1.73]		
200 01 01 (2020)	0.01 0			0.74	[0.01, 1.10]	0.070	1.070
Fixed effect model			6	0.67	[0.62; 0.72]	100.0%	
Random effects model			\sim		[0.56; 1.01]		100.0%
Prediction interval			-		[0.30; 1.90]		
Heterogeneity: $I^2 = 85\%$, $\tau^2 = 0$	0.1070. p <	0.01	гтт	٦			
			0.2 0.5 1 2	5			
	Combi	ning J	p = 0.028 (Fisher), 0.017 (I	iptak)		
ARBs							
						Maight	

					Mainha	Mainha
Study	TE seTI	E Odds Ratio	OR	95%-CI	Weight (fixed)	Weight (random)
Bauer et al (2020)	0.00 0.385	2	1.00	[0.47; 2.13]	2.8%	4.6%
Bravi et al (2020)	-0.19 0.258	6 -	0.83	[0.50; 1.38]	6.1%	8.2%
Dalan et al (2020)	1.05 0.992	3	- 2.87	[0.41; 20.09]	0.4%	0.8%
de Abajo et al (2020)	0.22 0.173	3 1		[0.89; 1.76]		12.8%
eng et al (2020)	-0.73 0.560	5		[0.16; 1.44]		2.4%
_i et al (2020)	-0.26 0.321	5	0.77	[0.41; 1.45]	4.0%	6.0%
_iu et al (2020)	-1.39 0.695	2	0.25	[0.06; 0.98]	0.9%	1.6%
Mancia et al (2020)	-0.19 0.140	7 🚽	0.83	[0.63; 1.09]	20.8%	15.2%
Mehra et al (2020, Retracted)	0.21 0.176	7 1=	1.23	[0.87; 1.74]	13.2%	12.6%
Mehta et al (2020)	0.11 0.327) —	1.12	[0.59; 2.13]	3.8%	5.9%
Meng et al (2020)	-1.11 0.662			[0.09; 1.21]		1.8%
Peng et al (2020)	-0.06 0.696	6		[0.24; 3.68]		1.6%
Rentsch et al (2020)	0.52 0.262	6 -	1.69	[1.01; 2.83]	6.0%	8.0%
Reynolds (2020)	-0.07 0.130	6 –	0.93	[0.72; 1.20]	24.1%	16.0%
rang et al (2020)	-1.14 0.775	4	0.32	[0.07; 1.46]	0.7%	1.3%
Zeng et al (2020)	-0.43 0.862		0.65	[0.12; 3.52]	0.6%	1.1%
Fixed effect model		Į.	0.00	[0.86; 1.11]	100.0%	
			0.90		100.070	
Random effects model		4		[0.80; 1.15]		100.0%
Prediction interval						100.0%
	0352, <i>p</i> = 0.10			[0.80; 1.15]		100.0%
Prediction interval	0352, <i>p</i> = 0.10	0.1 0.5 1 2 10		[0.80; 1.15]		100.0%
Prediction interval leterogeneity: $l^2 = 33\%$, $\tau^2 = 0.0$		0.1 0.5 1 2 10	0.96	[0.80; 1.15] [0.61; 1.50]	 Weight	Weight
Prediction interval	0352, <i>p</i> = 0.10 TE seT	0.1 0.5 1 2 10		[0.80; 1.15] [0.61; 1.50]	 Weight	Weight
Prediction interval Heterogeneity: $I^2 = 33\%$, $\tau^2 = 0.0$		0.1 0.5 1 2 10 E Hazard Ratio	0.96 HR	[0.80; 1.15] [0.61; 1.50]	 Weight	Weight
Prediction interval Heterogeneity: /² = 33%, τ² = 0. Study	TE seT	0.1 0.5 1 2 10 E Hazard Ratio 8	0.96 HR 0.80	[0.80; 1.15] [0.61; 1.50] 95%-CI	 Weight (fixed) 5.4%	Weight (random) 14.5%
Prediction interval Heterogeneity: / ² = 33%, τ ² = 0.1 Study Fosbol et al (2020)	TE seT -0.22 0.146 0.14 0.097	0.1 0.5 1 2 10 E Hazard Ratio 8	0.96 HR 0.80 1.15	[0.80; 1.15] [0.61; 1.50] 95%-CI	 Weight (fixed) 5.4% 12.4%	Weight (random) 14.5%
Prediction interval Heterogeneity: /² = 33%, τ² = 0. Study Fosbol et al (2020) Khera et al (2020)	TE seT -0.22 0.146 0.14 0.097	0.1 0.5 1 2 10 E Hazard Ratio 8	0.96 HR 0.80 1.15 0.99	[0.80; 1.15] [0.61; 1.50] 95%-CI [0.60; 1.07] [0.95; 1.39]	 Weight (fixed) 5.4% 12.4% 72.5%	Weight (random) 14.5% 22.4%
Prediction interval leterogeneity: $l^2 = 33\%$, $\tau^2 = 0.1$ Study Fosbol et al (2020) Khera et al (2020) Mehra et al (2020)	TE seT -0.22 0.146 0.14 0.097) -0.01 0.040	0.1 0.5 1 2 10 E Hazard Ratio 8	0.96 HR 0.80 1.15 0.99 1.16	[0.80; 1.15] [0.61; 1.50] 95%-Cl [0.60; 1.07] [0.95; 1.39] [0.91; 1.07]	 Weight (fixed) 5.4% 12.4% 72.5% 4.0%	Weight (random) 14.5% 22.4% 34.8%
Prediction interval Heterogeneity: / ² = 33%, τ ² = 0.1 Study Fosbol et al (2020) Khera et al (2020) Mehra et al (2020, Retracted Rossi et al (2020)	TE seT -0.22 0.146 0.14 0.097) -0.01 0.040 0.15 0.170	0.1 0.5 1 2 10 E Hazard Ratio 8	0.96 HR 0.80 1.15 0.99 1.16 0.42	[0.80; 1.15] [0.61; 1.50] 95%-Cl [0.95; 1.39] [0.91; 1.07] [0.93; 1.39] [0.91; 1.07]	 Weight (fixed) 5.4% 12.4% 72.5% 4.0% 0.7%	Weight (random) 14.5% 22.4% 34.8% 11.8% 2.8%
Prediction interval Heterogeneity: $l^2 = 33\%$, $\tau^2 = 0.1$ Study Fosbol et al (2020) Khera et al (2020) Mehra et al (2020) Rossi et al (2020) Zhang et al (2020)	TE seT -0.22 0.146 0.14 0.097) -0.01 0.040 0.15 0.170 -0.87 0.404	0.1 0.5 1 2 10 E Hazard Ratio 8	0.96 HR 0.80 1.15 0.99 1.16 0.42 1.03	[0.80; 1.15] [0.61; 1.50] 95%-CI [0.60; 1.07] [0.95; 1.39] [0.91; 1.07] [0.83; 1.62] [0.19; 0.93]	 Weight (fixed) 5.4% 12.4% 72.5% 4.0% 0.7% 5.0%	Weight (random) 14.5% 22.4% 34.8% 11.8% 2.8%
Prediction interval Heterogeneity: $l^2 = 33\%$, $\tau^2 = 0.1$ Study Fosbol et al (2020) Khera et al (2020) Mehra et al (2020, Retracted Rossi et al (2020) Zhang et al (2020) Lee et al (2020)	TE seT -0.22 0.146 0.14 0.097) -0.01 0.040 0.15 0.170 -0.87 0.404	0.1 0.5 1 2 10 E Hazard Ratio 8	0.96 HR 0.80 1.15 0.99 1.16 0.42 1.03 1.00	[0.80; 1.15] [0.61; 1.50] (0.61; 1.50] (0.95; 1.39] (0.91; 1.07] (0.83; 1.62] (0.19; 0.93] (0.77; 1.40]	 Weight (fixed) 5.4% 12.4% 72.5% 4.0% 0.7% 5.0% 100.0%	Weight (random) 14.5% 22.4% 34.8% 11.8%
Prediction interval Heterogeneity: $l^2 = 33\%$, $\tau^2 = 0.1$ Study Study Khera et al (2020) Mehra et al (2020) Mehra et al (2020) Zhang et al (2020) Lee et al (2020) Fixed effect model	TE seT -0.22 0.146 0.14 0.097) -0.01 0.040 0.15 0.170 -0.87 0.404	0.1 0.5 1 2 10 E Hazard Ratio 8	0.96 HR 0.80 1.15 0.99 1.16 0.42 1.03 1.00	[0.80; 1.15] [0.61; 1.50] 95%-Cl [0.60; 1.07] (0.95; 1.39] (0.91; 1.07] [0.83; 1.62] [0.19; 0.93] [0.77; 1.40] [0.93; 1.07]	 Weight (fixed) 5.4% 12.4% 72.5% 4.0% 0.7% 5.0% 100.0%	Weight (random) 14.5% 22.4% 34.8% 11.8% 2.8% 13.6%

Combining p = 0.915 (Fisher), 0.886 (Liptak)

Figure 3. Meta-analysis of survival with prior use of angiotensin-converting enzyme inhibitors (ACEIs) and angiotensin receptor blockers (ARBs). Gray squares represent treatment-to-control ratios in the reports; their size is proportional to the number of events. The 95% confidence intervals (CIs) for individual reports are denoted by lines, while those for pooled odds ratios are denoted by diamonds. (A) ACEIs. (B) ARBs. TE, estimated treatment effect; seTE, standard error of pooled estimated treatment effect; OR, odds ratio; HR, hazard ratio.

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	Df	Wald chi-square	p value
Contrasts statement (COVID-19 or retrospective cohort)			
$H_o: \beta_{ARBs COVID-19} = \beta_{ARBs pneumonia}$	1	8.2740	0.0040
$H_o: \beta_{ACEI COVID-19} = \beta_{ACEi pneumonia}$	1	2.2706	0.1319

Table 4. Comparison of hazard ratio between COVID-19 patients and a retrospective cohort of patients hospitalized with pneumonia between January and June 2019 in Korea

COVID-19, coronavirus disease 2019; ARB, angiotensin receptor blocker; ACEI, angiotensin-converting enzyme inhibitor.

tem inhibitors did not increase SARS-CoV-2 infectivity [23,24]. Another hypothesis is that ARBs are less effective in attenuating inflammatory response in COVID-19 than in typical pneumonia. Interestingly, ARBs showed a loss of beneficial effect on COVID-19 conflicting with previous meta-analysis results in pneumonia patients [11,16], and with data from pneumonia patients in Korea in 2019. Although several studies, which urgently reported no harm with ARBs and ACEIs in COVID-19 patients (see references for the meta-analysis), given that beneficial effects have been reproducibly reported in pneumonia patients, neutral results suggest that there might be possible offset effects. Although ARBs showed non-inferior cardiovascular protective effect to ACEIs in a large scaled randomized controlled trial [50], several meta-analyses suggested superior cardiovascular protecting effects of ACEIs over ARBs in high risk hypertensive patients [51,52]. A replenished effect of bradykinin by ACEIs, which is reduced in cardiovascular disease status, might exert a superior anti-inflammatory effect over ARBs [53]. The lack of benefit of ARBs observed in this study might be significant, because multiple clinical trials using ARBs are recruiting patients based on the initial clinical attention associated with ACE2-COVID-19 issues (14 studies, with the search terms of "COVID-19" and "ARBs," identified in ClinicalTrials.gov, accessed January 25, 2021).

The current study has several limitations. Firstly, data related to drug exposure were measured based on claim data. Hence, detailed information about drug exposure, such as adherence to the medication or discontinuation during COVID-19 hospitalization, remains unknown. However, the Korean Society of Hypertension and the Korean Society of Cardiology formally recommended that doctors continue administering anti-hypertensive medication without change during COVID-19 treat-

ment [54-56]. Therefore, we rationally assumed that the infected hypertensive patients continued the use of anti-hypertensive medications during the period of hospitalization. However, there might be a difference in the compliance of antihypertensive medications. Indeed, it was reported that about 26% of COVID-19 patients with hypertension might exhibited drug compliance below 80% [8], which potentially influenced the association between the use of ARBs or ACEIs and mortality in COVID-19 patients either directly or indirectly through the healthy adherer effect or selective prescribing [57]. Secondly, although we evaluated a total of 10,448 COVID-19 patients who were hospitalized in Korea from January 19, 2020, through April 15, 2020, the interpretation of our findings is influenced by the limited sample size. Notably, the prescription rate of ACEIs was barely 3% in hypertensive patients likely because of dry cough, a common side effect among Asians [36]. To overcome this limitation, we compared the results with a retrospective cohort of more than 170,000 pneumonia patients hospitalized from January to June in 2019 to include more than 1,000 patients with prescribed use of ACEIs, assuming that 3% of prescription rate had 40% hypertension prevalence among all pneumonia cases. The results indicated similar trends to those followed by COVID-19 and pneumonia cases. Also, we performed meta-analyses of ARBs and ACEIs using the data from previous studies merged with our study, thus evaluating 145,124 patients. The results of the meta-analysis supported the findings of our study.

In conclusion, while ARBs or ACEIs were not associated with fatal outcomes, the possible beneficial effect of ARBs observed in pneumonia patients was not observed in COVID-19 patients. The findings of this paper implicate that ARBs might have no role as a COVID-19 therapeutic.

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KEY MESSAGE

- In coronavirus disease 2019 (COVID-19) patients in Korea, angiotensin receptor blockers (ARBs) showed neutral association, whereas angiotensin converting enzyme (ACE) inhibitors showed no significant associations likely owing to the small population size.
- 2. When comparing a retrospective cohort of patients hospitalized with pneumonia in 2019, ACE inhibitors showed similar benefits. In contrast, the protective effect of ARBs was not reproduced in the COVID-19 cohort.
- 3. Meta-analysis merging the data from our and previously published studies showed significant benefit of ACE inhibitors, but a neutral association between ARBs and the mortality.

Conflict of interest

No potential conflict of interest relevant to this article was reported.

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Supplementary Table 1. Clinical characteristics of pneumonia patients in Korea (2019)

Values are presented as mean \pm standard deviation or number (%).

ARB, angiotensin receptor blocker; ACEI, angiotensin-converting enzyme inhibitor; COPD, chronic obstructive pulmonary disorder.



Variable	No. patients	No. death	HR	95% CI	p value
Age, yr					
0-64	111,602	634	1		
≥ 65	67,663	12,650	36.223	33.446–39.230	< 0.000
Sex					
Male	90,705	7,429	1		
Female	88,560	5,855	0.803	0.776–0.831	< 0.000
Comorbidity (upper, no; lower, yes)					
Hypertension	119,536	2,884	1		
	59,729	10,400	7.814	7.498-8.143	< 0.000
Diabetes mellitus	137,397	6,344	1		
	41,868	6,940	3.822	3.694–3.955	< 0.000
Coronary artery disease	156,504	9,310	1		
	22,761	3,974	3.120	3.006-3.238	< 0.000
Stroke	159,465	8,983	1		
	19,800	4,301	4.199	4.049-4.354	< 0.000
COPD	134,212	7,522	1		
	45,053	5,762	2.366	2.286–2.449	< 0.000
Cancer (any)	166,875	10,448	1		
	12,390	2,836	4.000	3.837–4.169	< 0.000
Chronic kidney disease	172,876	11,850	1		
	6,389	1,434	3.553	3.364-3.753	< 0.000
Heart failure	161,697	9,656	1		
	17,568	3,628	3.737	3.597-3.882	< 0.000
Medication					
Non-ACEIs/ARBs	143,583	7,612	1		
ARBs	33,626	5,243	3.100	2.992-3.211	< 0.000
ACEIs	2,056	429	4.266	3.871–4.702	< 0.000
Non-ACEIs/ARBs	143,583	7,612	1		
ARBs or ACEIs	35,682	5,672	3.165	3.058-3.276	< 0.000

Supplementary Table 2	. Results of univariable re	gression analysis	for the mortality	within 60 days
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HR, hazard ratio; CI, confidence interval; COPD, chronic obstructive pulmonary disorder. ACEI, angiotensin-converting enzyme inhibitor; ARB, angiotensin receptor blocker.



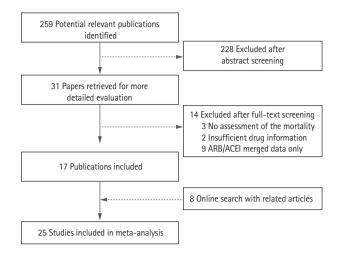
Variable	No. patients	No. death	HR	95% CI	p value
Age, yr					
0–64	111,602	634	1		
≥ 65	67,663	12,650	27.036	24.765–29.516	< 0.0001
Sex					
Male	90,705	7,429	1		
Female	88,560	5,855	0.862	0.832-0.893	< 0.0001
Comorbidity (upper, no; lower, yes)					
Hypertension	119,536	2,884	1		
	59,729	10,400	1.433	1.363–1.507	< 0.0001
Diabetes mellitus	137,397	6,344	1		
	41,868	6,940	1.061	1.022-1.100	0.0017
Coronary artery disease	156,504	9,310	1		
	22,761	3,974	0.982	0.943–1.024	0.3988
Stroke	159,465	8,983	1		
	19,800	4,301	1.356	1.306–1.408	< 0.0001
COPD	134,212	7,522	1		
	45,053	5,762	0.864	0.833-0.895	< 0.0001
Cancer (any)	166,875	10,448	1		
	12,390	2,836	1.631-	1.563–1.702	< 0.0001
Chronic kidney disease	172,876	11,850	1		
	6,389	1,434	1.264	1.193–1.338	< 0.0001
Heart failure	161,697	9,656	1		
	17,568	3,628	1.281	1.228–1.335	< 0.0001
Medication					
Non-ACEIs/ARBs	143,583	7,612	1		
ARBs	33,626	5,243	0.853	0.820-0.887	< 0.0001
ACEIs	2,056	429	1.089	0.986–1.203	0.0987
Non-ACEIs/ARBs	143,583	7,612	1		
ARBs or ACEIs	35,682	5,672	0.866	0.834–0.900	< 0.0001

Supplementary Table	3. Results of multivariable re	egression analysis of AF	RBs and ACEIs for the mortal	ity within 60 days

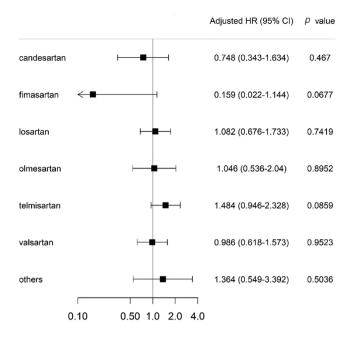
Adjusted for age, sex, the history of medication (statins in case of ARBs/ACEIs), the history of comorbidities (hypertension, diabetes mellitus, cancer, COPD, stroke, coronary artery disease, heart failure, and chronic kidney disease, statin use) before the diagnosis of coronavirus disease 2019.

ARB, angiotensin receptor blocker; ACEI, angiotensin-converting enzyme inhibitor; HR, hazard ratio; CI, confidence interval; COPD, chronic obstructive pulmonary disorder.





Supplementary Figure 1. Flow diagram for meta-analysis. ARB, angiotensin receptor blocker; ACEI, angiotensin-converting enzyme inhibitor.



Supplementary Figure 2. Subgroup analysis for survival within 60 days by each angiotensin receptor blocker. Only individual drugs having sufficient prescription numbers for analysis were analyzed. HR, hazard ratio; CI, confidence interval.

