Fish Consumption and Incidence of Diabetes

Meta-analysis of data from 438,000 individuals in 12 independent prospective cohorts with an average 11-year follow-up

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OBJECTIVE—Ecological data suggest an inverse correlation between fish consumption and diabetes prevalence. However, epidemiological data on fish intake and diabetes incidence are controversial and inconclusive. Therefore, we aimed to assess the literature and determine the association between fish consumption and diabetes risk quantitatively.

RESEARCH DESIGN AND METHODS—Prospective cohort studies published through August 2011 in peer-reviewed journals indexed in PubMed were selected. Additional information was retrieved through Google or a hand search of the references from relevant articles. The weighted relative risk (RR) and corresponding 95% CI for incident diabetes was estimated using random-effects models.

RESULTS—A database was derived from nine eligible studies (12 independent cohorts), including 438,214 individuals with an average 11.4-year follow-up. Compared with those who never consumed fish or ate fish less than once per month, the pooled RR of incident diabetes was 0.99 (95% CI 0.85–1.16) for individuals who ate fish five or more times per week ($P_{trend} = 0.80$). Similar results were found for long-chain n-3 polyunsaturated fatty acid intake. Study location was an effect modifier. An inverse association between fish intake and diabetes incidence was found by combining studies conducted in Eastern but not Western countries.

CONCLUSIONS—Accumulated evidence generated from this meta-analysis does not support an overall inverse association of fish or fish oil intake with incidence of diabetes. The null association was modified by study location (Eastern vs. Western countries), which may reflect the possible difference between Eastern and Western dietary patterns. Further studies are warranted.

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Ithough cardioprotective effects of fish or long-chain n-3 polyunsaturated fatty acid (LCn3PUFA) intake have been documented (1), the role of these dietary factors in the development of diabetes remains uncertain. Since ecological data indicate low prevalence of diabetes in populations with a large amount of fish consumption (2) and laboratory studies suggest that insulin sensitivity is directly associated with the content of LCn3PUFA in cell membranes (3), it has been hypothesized that fish or LCn3PUFA

intake may provide beneficial effects on diabetes risk. In the past decade, a number of longitudinal studies have examined the association between fish intake and incidence of diabetes, but the findings are conflicting. Although some studies reported a significant inverse association (4–6), other cohort studies (7–9) found a positive relation between fish consumption and incidence of diabetes after adjustment for various potential confounders. To provide a reliable quantitative assessment of the association of fish and LCn3PUFA

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See accompanying commentary, p. 666, and meta-analysis, p. 918.

intake with diabetes risk and to explore major sources of heterogeneity among studies, we conducted a meta-analysis of all eligible prospective cohort studies with the existing data.

RESEARCH DESIGN AND

METHODS—We undertook this metaanalysis according to MOOSE (Meta-analysis of Observational Studies in Epidemiology) guidelines during all stages of design, implementation, and reporting (10).

Study selection

All prospective cohort studies published in English-language journals from 1966 to August 2011 that reported the association between fish consumption and incidence of diabetes were identified by searching PubMed using the terms "fish," "seafood," "animal products," "omega-3 fatty acids," "n-3 fatty acids," "diabetes," "glucose tolerance," and "insulin resistance." Additional information was retrieved through Google and a hand search of the references from relevant articles. Both authors independently reviewed all relevant papers and identified eligible studies. Discrepancies were resolved by discussion. An article would be included if the study was a prospective cohort design and the relative risks (RRs) and their corresponding 95% CIs of diabetes relating to fish and/or LCn3PUFA intake were presented or such information could be recalculated.

Of 283 identified abstracts from the PubMed, we excluded 269 by screening because they were reviews, editorials, or letters to the editor; were not prospective cohort studies; studied diabetes patients; were animal studies; did not report RR of fish or fish oil and incident diabetes; or were not in English (Fig. 1). Two investigators independently examined the full text of the remaining 14 studies (4-9,11-18) to confirm eligibility for inclusion. Of the 14 studies, 1 study (11) was excluded because it only categorized fish consumption to two groups (yes vs. no). Two studies (14,15) were excluded because more detailed results were published

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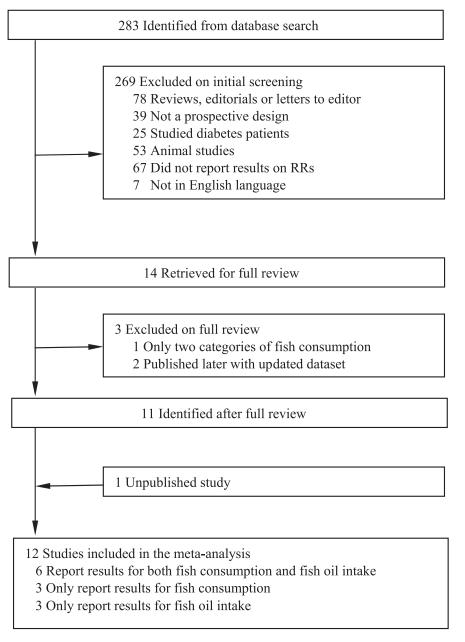


Figure 1—Process of study selection.

(7) later with the same dataset. In addition, 1 study (7) reports findings from 3 independent cohorts and another study (5) presents results from 2 independent cohorts so that these 2 studies were counted as 5 separate cohorts in the meta-analysis. Moreover, we generated unpublished de novo results for this meta-analysis by analyzing data from the Coronary Artery Risk Development in Young Adults (CARDIA) Study. Thus, 12 studies (15 independent cohorts) (4–9,12,13,16–18) were included for the meta-analysis. Of them, 9 studies (12 cohorts) (4–9,12,13) reported results on fish consumption and 9 studies (11 cohorts) (5,7–9,13,16–18) reported results on LCn3PUFA intake.

Quality assessment and data extraction

To evaluate the quality of the studies included, we designed a scoring system with several a priori selected important characteristics that may affect study quality based on literature (19). The system was created to account for study eligibility (1 point if appropriate inclusion and exclusion criteria was clearly stated), exposure (1 point if fish intake was assessed by a validated tool; and 1 point if fish intake was standardized [1]), outcome (1 point if diabetes cases were diagnosed at clinical examination or self-reported cases verified by clinicians on the basis of established clinical criteria [20]), and statistical analysis (1 point if adjustment included age, BMI, energy intake, and family history of diabetes). The scoring system was designed with reference to MOOSE (10) and allowed a total score from 0 to 5 points, with 5 reflecting the highest quality. Two authors independently assessed the quality for all identified studies.

The data included the first author's name, year of publication, study population, country of origin, sample size, proportion of men, duration of follow-up, range or mean of participants' age, number of events, case identification methods, categories of fish or LCn3PUFA intake, the amount of fish or LCn3PUFA intake for each category, methods for diet measurement, adjusted covariates, as well as RRs and 95% CIs of incident diabetes in the corresponding categories. RRs transformed to their natural logarithms (ln) and the 95% CIs were used to compute the corresponding SEs.

To standardize fish intake, we first converted frequency into grams per day. The amount of fish consumption (g/day) was estimated by multiplying the frequency of consumption (serving/day) by the corresponding portion size (g/serving). For example, the derived average portion size in Kaushik et al. (7) was 105 g/serving. The range of fish consumption for one to three times per month (shown as, for example, 1-3/month hereafter) was converted to 3.5 g/day (105 \times 1/30) to 10.5 g/day (105 \times 3/30). The median or mean value between two adjacent categories was used as the cut-off point. When the portion size of fish intake in an individual study was not available from the published paper, the information was acquired from authors of the primary study or the value was determined based on data from Kaushik et al. because this study includes three large cohort studies (the Nurses' Health Study [NHS], the Nurses' Health Study 2 [NHS2] and the Health Professionals Follow-up Study [HPFS]) and taking into consideration that the food frequency questionnaire (FFQ) used in these three studies has been validated (21). If the upper-bound of the highest fish intake category was uncertain, for instance, fish intake \geq 5/week, we assigned 1 serving/day of fish as the upper limit. If the information from the primary

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study was insufficient for standardizing fish consumption, especially if the reference level of fish intake was substantially higher than the standardized level, de novo results based on the standardized fish consumption cut-off points were requested from authors of the primary study (4–6).

Statistical analysis

Fish consumption was standardized into five categories: never or < 1/month, 1-3/month, 1/week, 2-4/week, and \geq 5/week. We created a dataset by assigning each RR extracted from each primary study into its corresponding standardized category according to the range or median amount of fish intake in the category. If the median amount of fish consumption from more than one group in a single study fell into the same standardized category of fish intake in our meta-analysis, we pooled these RRs and used the combined estimate for that category. Also, if the range of fish intake covered more than one standardized category, we allocated RR based on the median fish intake. We estimated the pooled RRs and 95% CIs of diabetes for each standardized category of fish consumption compared with the lowest category using random-effects models. In addition, the Cochran test was used to test for heterogeneity among studies, and I^2 was computed to determine the degree of inconsistency across studies. A variance-weighted least squares regression was used to model the In RR for diabetes as a linear function of fish intake. The median intake of fish for each category was derived. Publication bias was assessed by visualizing the funnel plot and determined by the Egger asymmetry test (22).

We further carried out subgroup analysis to explore potential sources of heterogeneity based on sex, study location (Eastern vs. Western countries), duration of follow-up (<11.4 vs. ≥ 11.4 years), and outcome identification method (clinical examination vs. self-report) by using randomeffects models.

Two sensitivity analyses were conducted to test the robustness of our findings: *1*) we evaluated whether the results could have been affected markedly by a single study; and 2) we restricted our meta-analysis to studies that specified type 2 diabetes in the outcome.

 $P \le 0.05$ was considered statistically significant for all tests. All analyses were performed using STATA statistical software (Version 11.0. STATA Corp., College Station, TX).

RESULTS

Characteristics of included studies

The final dataset for our meta-analysis of fish consumption and incident diabetes included 12 independent cohorts from nine studies comprising 438,214 individuals (18,711 incident diabetes cases) aged between 18 and 98 years. Of the 12 cohorts, 9 were from Western countries and 3 from Asia. The sample sizes varied across studies from 3,088 (the Cardiovascular Health Study) (13) to 91,669 (NHS2) (7). The average duration of follow-up was 11.4 years (range 4-23 years), calculated based on the person-times of primary studies. Diet including fish consumption was collected by either self-administered or interview-based FFQ. Fish intake in primary studies was classified into four to five categories. Incident diabetes cases were self-reported (8 cohorts) or diagnosed at clinical examinations (5 cohorts). Two studies did not specify type of diabetes in the outcome. The quality scores for all included studies are ≥ 3 points based on our 5-point scoring system. All reported RRs (95% CI) of diabetes in each study were adjusted for multiple covariates (Table 1).

Association of fish consumption and LCn3PUFA intake with incidence of diabetes

After pooling data from the identified studies, neither fish consumption nor LCn3PUFA intake was significantly associated with incidence of diabetes. Compared with those who never consumed fish or ate fish less than once per month, the pooled RR of incident diabetes was 0.99 (95% CI 0.85-1.16) for 5 or more times per week ($P_{\text{trend}} = 0.80$) (Table 2). We tested the heterogeneity among the included cohorts by calculating I^2 for each categories of fish intake with the lowest group as the reference. The range of I^2 values was from 4.5 to 83.7% for fish consumption. As shown in Fig. 2, there was significant heterogeneity between studies $(I^2 =$ 83.7%, P < 0.01) when comparing participants who consumed fish 2-4/week with those in the reference group. However, a sensitivity analysis revealed that none of the studies significantly influenced the pooled estimate, with the pooled RR ranging from 0.96 (0.82–1.12) after excluding Djoussé et al. (8) to 1.06 (0.91–1.23) after excluding the study on men by Villegas et al. (5).

Similar results were found for LCn3PUFA intake. The pooled RR (95%

CI) of incident diabetes across quintiles of LCn3PUFA intake were 1.00, 1.00 (0.96–1.05), 1.00 (0.93–1.08), 1.04 (0.93–1.17), and 1.04 (0.92–1.18; P_{trend}=0.38).

Stratified analysis

We conducted stratified analyses based on the available data (Table 2). The observed null associations were modified by study location (Eastern vs. Western countries) and follow-up period (<11.4 vs. \geq 11.4 years). An inverse association between fish consumption and diabetes risk was documented by combining studies conducted in Asia and studies with a follow-up period <11.4 years. However, the associations of fish consumption with risk of diabetes were not appreciably modified by sex and outcome identification methods (clinical examination vs. self-report). In addition, two studies did not specify type of diabetes in the outcome. Nevertheless, the pooled results were consistent when excluding these two studies (data not shown).

Publication bias

Publication bias was assessed by Egger test, and no evidence was found (P = 0.14).

CONCLUSIONS—In this quantitative meta-analysis of prospective cohort studies, we did not find an overall pooled association of fish consumption or LCn3PUFA intake with diabetes incidence. However, evidence from studies conducted in Asia suggests an inverse association between fish consumption and diabetes risk. The follow-up period appeared to also modify the association, which may reflect the fact that all studies with a follow-up period longer than the average follow-up were conducted in Western countries and most studies with a follow-up period shorter than the average were conducted in Eastern countries. No other effect modifiers were identified.

This meta-analysis comprises more than 435,000 male and female adults with a wide age range. Most cohorts or studies included in this meta-analysis had a large sample size and long-term follow-up periods that increased the statistical power to examine the overall associations of fish intake and incidence of diabetes. The prospective cohort study design also reduced the likelihood of selection bias and recall bias. Although randomized placebo controlled trials are the best approach to establish causal inference, no randomized data on fish consumption and diabetes incidence are available yet, presumably, in part because of the feasibility considerations such as participant's long-term adherence. Thus, evidence generated from meta-analysis of cohort studies is very valuable, especially for reliably estimating the long-term association between fish consumption and diabetes risk. Another major strength to be highlighted is that we standardized fish consumption in the meta-analysis. Specifically, for those primary studies (4-6) with insufficient information for fish consumption standardization, de novo results were obtained from the investigators of the primary studies, which substantially reduced the misclassification of fish intake.

Because this meta-analysis is based on observational studies, the inherent limitations of primary studies may have affected our findings. For example, the individual RR estimate included in the meta-analysis was adjusted for different covariates in the different studies. Thus, the possibility of residual confounding due to inclusion of different factors or bias due to measurement errors or unmeasured factors cannot be ruled out. Alternatively, a meta-analysis based on primary data with subsequent adjustment for identified covariates may provide more solid evidence. In addition, differences in method of case identification (i.e., clinical examination vs. self-report) might lead to difficulties in estimating true effects of fish intake on diabetes risk. For instance, individuals with high fish intake are generally more health conscious, and therefore they may be less likely to have undiagnosed cases of diabetes. By contrast, those in the low fish intake groups may be more likely to have undiagnosed diabetes, thus underestimation of diabetes cases could occur particularly in the studies with self-reported cases. This limitation may bias the results toward the null. Nevertheless, our stratified analyses did not support the presence of substantial effect modifications by case identification methods.

In this meta-analysis, we found that study location modified the association between fish consumption and diabetes risk. A combined inverse association was observed in 3 studies conducted in Eastern countries (Japan and China). The reasons for this modification are unclear. Presumably, fish consumption may be a marker of healthy dietary pattern in the observational studies. Thus, the effect modification of study location we observed may reflect the differences in dietary pattern in relation to diabetes risk between Eastern and Western population. For example, fast

food and sweetened beverages-possible risk factors of diabetes-are more popular in Western countries (23). Also, food preparation methods may partially explain this effect modification. Raw, boiled, and steamed fish are major preparation methods in Eastern cuisine. In contrast, fried fish may be more popular in Western food. Studies suggest that frying may modify the lipid profile through a decrease in LCn3PUFA content (24). Deep frying (e.g., fried fast food) may also cause the presence of *trans*-fatty acids and lipid oxidation products that may consequently increase the risk of diabetes (18) by promoting systemic inflammation (25) and oxidant stress (26). In addition, we noted that the amounts of fish consumption in studies conducted in Asia were much higher than the amounts of fish intake in Western countries. The possibility cannot be completely ruled out that the potential benefit of fish consumption on diabetes development would only be seen at a high LCn3PUFA levels. In other words, a relatively low LCn3PUFA intake may not be able to outweigh the risk of contaminants in fish. Further studies are clearly warranted.

In addition, we observed effect modification by study follow-up period in this meta-analysis. Of note, this modification may simply be a surrogate of effect modification of study location because all studies with a follow-up period (mean \sim 16 years) above the average study duration (~11 years) were conducted in Western countries, whereas most studies with a follow-up period (mean \sim 8 years) below the average study duration were conducted in Asia. Another possible explanation of this effect modification may be that the incident diabetes cases were delayed rather than completely prevented by fish consumption so that an inverse association was more likely to be observed in studies with a relatively short follow-up period.

Our results are robust in a couple of sensitivity analyses. First, none of the primary studies significantly influenced the pooled estimate when we excluded one study at a time in the meta-analysis. Second, the results remained when we excluded two studies without specified type 2 diabetes.

The possibility of publication bias is always a concern in a meta-analysis. By visualizing the funnel plot and conducting statistical tests, we found no strong evidence of publication bias in our results. Nevertheless, a potential bias resulting from excluding studies published in other languages (if any) is possible.

Because the prevalence of diabetes is very rare in populations with high fish consumption (27) and animal studies indicate a beneficial effect of fish oil intake on glucose metabolism (28), it was hypothesized that dietary enrichment with LCn3PUFAs may reduce risk of diabetes (2). However, the present meta-analysis does not generate strong evidence supporting this hypothesis. In addition to the heterogeneity of individual studies, some other important issues that may explain the overall null association between fish intake and diabetes incidence merit consideration. Although LCn3PUFAs are the key nutrients in fish responsible for the potential cardioprotective effect of fish consumption, fish is not only LCn3PUFAs but a package of nutrients and contaminants (29). Notably, fish is a major source of exposure to methylmercury and polychlorinated biphenyls (PCBs) (30,31). Studies suggest that PCBs may be a risk factor of type 2 diabetes (32). Also, laboratory studies indicate that even low dose mercury-induced oxidative stress and phosphoinositide 3-kinase activation can cause pancreatic B-cell dysfunction and increase risk of diabetes (33). Hence, assuming LCn3PUFAs are beneficial to risk of diabetes development, contaminants in fishparticularly PCBs and methylmercurymay substantially attenuate or even cancel the benefit of LCn3PUFAs. In addition. cooking methods may alter the nutrient content in fish and generate unexpected chemicals (24), which may confound the effects of fish intake. In this meta-analysis, we also combined studies that reported results on LCn3PUFA intake and incidence of diabetes. We found no association between LCn3PUFAs and diabetes, which was generally in concordance with what we observed in fish. In fact, consistent associations of fish and LCn3PUFA intake with diabetes were found in all studies with data available on both fish and LCn3PUFAs, which may reflect the fact that LCn3PUFA intake assessed by diet measurement instrument is more likely to be a surrogate marker of fish consumption because total LCn3PUFA intake is derived from the nutrient table based on the frequency of fish intake. The effects of dietary LCn3PUFA intake cannot be isolated from fish intake. Data on objective biomarkers of LCn3PUFAs and diabetes are sparse. Three studies found no significant correlations between serum or plasma LCn3PUFAs and diabetes (34-36), and

Source	Participants (n)	Age (years)	Men (%)	Duration of follow-up, (years)*	Exposure assessment	Exposure categories	No. of incident diabetes and case identification methods	Adjusted variables
Montonen et al. 2005 (12). The Finnish Mobile Clinic Health Examination Survey, Finland	4,304	40–69	53.1	23	Self-administered questionnaire	<5 g/day; 5−16 g/day; 17–36 g/day; ≥37 g/day.	383, diagnosed by physician	Age, sex, BMI, energy intake, smoking, family history of diabetes, and geographic area.
Kaushik et al. 2009 (7). Health Professionals Follow-up Study (HPFS), U.S.	42,504	39–78	100	15	Self-administered questionnaire	→ 1.03/month;I/week;2-4/week;=5/week.	2,493, self-report	Age, smoking, alcohol consumption, physical activity, family history of diabetes, BMI; intakes of saturated fat, <i>trans</i> fats, linolenic acid, linoleic acid, caffeine, and cereal fiber; glycemic index, calorie intake.
Kaushik et al. 2009 (7). Nurses' Health Study (NHS), U.S.	61,031	30-55	0	16	Self-administered questionnaire		4,159, self-report	Age, smoking, alcohol consumption, physical activity, family history of diabetes, BMI; intakes of saturated fat, <i>trans</i> fats, linolenic acid, linoleic acid, caffeine, and cereal fiber; glycemic index, calorie intake, menopausal status, and postmenopausal hormone use.
Kaushik et al. 2009 (7). Nurses' Health Study 2 (NHS2), U.S.	91,669	26-46	0	15	Self-administered questionnaire		2,728, self-report	Age, smolking, alcohol consumption, physical activity, family history of diabetes, BMI; intakes of saturated fat, <i>trans</i> fats, linolenic acid, inolecic acid, caffeine, and cereal fiber; glycemic index, calorie intake, use of hormone replacement therapy, and oral contraceptive use.
Patel et al. 2009 (4). EPIC-Norfolk Study, U.K.	21,984	40-79	45	10.2	Self-administered questionnaire	−2/month; 1–3/month;1–2/week; 3–4/week;≥5/week.	725, self-report	Age, sex, family history of diabetes, smoking, education level, physical activity, total energy intake, alcohol intake, plasma vitamin C, BMI, and waist circumference.
van Woudenbergh et al. 2009 (9). Rotterdam Study, the Netherlands	4,472	67.2 ± 7.7	41	12.4	Self-administered questionnaire	no; low; moderate; high.	463, diagnosed by clinical examination according to ADA 1997 criteria and WHO 1999 criteria	Age, sex, smoking, education level; intakes of energy, alcohol, <i>trans</i> fatty acids, and fiber.

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Table 1—Characteristics of 12 included cohorts (from nine studies) on fish consumption and incidence of diabetes

Table 1—Continued

Source	Participants (n)	Age (years)	Men (%)	Duration of follow-up, (years)*	Exposure assessment	Exposure categories	No. of incident diabetes and case identification methods	Adjusted variables
Djoussé et al. 2011 (13). Cardiovascular Health Study (CHS), U.S.	3,088	65–98	38.9	9.6	Picture-sort questionnaire	<1/month; 1–3/month; 1–2/week; 3–4/week; ≥5/week.	204, diagnosed by clinical examination	Age, race, sex, clinic site, BMI, alcohol consumption, physical activity, current smoking, total energy intake, and LDL cholesterol.
Djoussé et al. 2011 (8). Women's Health Study (WHS), U.S.	36,328	54.6 ± 7.0	0	12.4	Self-administered questionnaire	Quintiles (median): 0.47/week; 0.93/week; 1.47/week; 1.93/week; 3.93/week.	2,370, self-report	Age, BMI, parental history of diabetes, smoking, exercise, alcohol intake, menopausal status; intakes of red-meat, energy, linoleic acid, α-linolenic acid, dietary magnesium, <i>trans</i> fat, saturated fat, and cereal fiber; and glycemic index.
Namri et al. 2011 (6). Japan Public Health Center- based Prospective Study (JPHC), Japan	52,680	45-75	43.5	Ŋ	Self-administered questionnaire	<1/month; 1–3/month; 1/week; 2–4/week; ≥5/week.	971, self-report	Age, study area, BMI, smoking, alcohol, family history of diabetes, total physical activity, history of hypertension, total energy intake, coffee consumption; and intakes of calcium, magnesium, dietary fiber, vegetables, fruit, meat, and rice.
Villegas et al. 2011 (5). Shanghai Men's Health Study (SMHS), China	52,054	40-74	100	4. l	Self-administered questionnaire	<1/month; 1–3/month; 1/week; 2–4/week; ≥5/week.	900, self-report	Age, energy intake, waist-to-hip ratio, BMI, smoking, alcohol consumption, physical activity, income level, education level, occupation, family history of diabetes, hypertension, and dietary pattern.
Villegas et al. 2011 (5). Shanghai Women's Health Study (SWHS), China	64,193	40-70	0	8.9	Self-administered questionnaire	<1/month; 1–3/month; 1/week; 2–4/week; ≥5/week.	3,034, self-report	Age, energy intake, waist-to-hip ratio, BMI, smoking, alcohol consumption, physical activity, income level, education level, occupation, family history of diabetes, hypertension, and dietary pattern.
He et al. Unpublished. CARDIA, U.S.	3,907	18–30	47.4	20	Interviewer- administered questionnaire	<1/month; 1–3/month; 1/week; 2–4/week; ≥5/week.	281, diagnosed by clinical examination according to ADA 2010 criteria	Age, race, sex, study center, education, smoking, alcohol consumption, physical activity, BMI, systolic blood pressure, triglycerides, family history of diabetes; and intakes of saturated fat, carbohydrates, fiber, and total energy.

	No. of	No. of participants			Fish consumpti	on	
	cohorts	(events)	<1/month	1–3/month	1/week	2–4/week	≥5/week
All studies	12	438,214 (18,711)	1.0	0.97 (0.90–1.04)	0.99 (0.89–1.10)	1.00 (0.85–1.18)	0.99 (0.85–1.16)
Sex							
Men	4	151,145 (4,645)	1.0	0.83 (0.60–1.15)	0.85 (0.60-1.21)	0.87 (0.55–1.37)	0.93 (0.64–1.36)
Women	6	273,480 (11,173)	1.0	0.99 (0.90-1.09)	1.06 (0.96–1.17)	1.11 (0.88–1.39)	1.08 (0.82-1.41)
Study location							
Eastern	3	168,927 (4,905)	1.0	0.81 (0.57–1.15)	0.76 (0.64–0.92)	0.71 (0.60–0.84)	0.79 (0.66-0.96)
Western	9	269,287 (13,806)	1.0	0.99 (0.92-1.06)	1.07 (1.00-1.15)	1.13 (0.98–1.29)	1.09 (0.94–1.27)
Follow-up							
<11.4 years	5	193,999 (5,834)	1.0	0.88 (0.71–1.10)	0.77 (0.66–0.90)	0.73 (0.63–0.85)	0.79 (0.67–0.94)
\geq 11.4 years	7	244,215 (12,877)	1.0	0.98 (0.91–1.06)	1.09 (1.03–1.15)	1.17 (1.02–1.34)	1.13 (0.96-1.33)
Case diagnosis methods							
Clinical examination	4	15,771 (1,331)	1.0	0.98 (0.80–1.21)	1.13 (0.90–1.42)	1.03 (0.77–1.37)	0.92 (0.78-1.09)
Self-report	8	422,443 (17,380)	1.0	0.97 (0.89–1.05)	0.97 (0.86–1.09)	0.99 (0.81–1.21)	1.00 (0.81-1.22)

one study reported a borderline significant inverse association (13). Because the amount of endogenous LCn3PUFA is small (37), the biomarker levels of LCn3PUFAs should be parallel with the total amount of fish consumption as well as the content of other constituents in fish. Thus, findings from studies using the objective biomarkers of LCn3PUFA are still subject to potential confounding from other components in fish such as methylmercury and PCBs.

In future studies, randomized placebocontrolled trials on fish oil supplements and glucose metabolism or diabetes risk are needed to elucidate the potential mechanism of action and to help us better understand the key nutrients—LCn3PUFAs—in fish in relation to incidence of diabetes. In cohort studies, investigators should pay attention to collecting information on types of fish, cooking methods, and other important constituents in fish. Besides methylmercury and PCBs, fish is rich in selenium and vitamin D (38). Recent studies suggest that high levels of selenium may be associated with elevated risk of

Study	Year						RR (95% CI)	Weight (%)
Montonen	2005						0.85 (0.69, 1.04)	10.38
Kaushik-HPFS	2009				+		1.10 (0.95, 1.28)	11.21
Kaushik-NHS	2009						1.22 (1.04, 1.43)	11.07
Kaushik-NHS2	2009						1.20 (1.01, 1.42)	10.92
Patel	2009*				•		0.77 (0.52, 1.13)	7.38
van Woudenbergh	2009				+	_	1.32 (1.02, 1.70)	9.51
Djoussé-CHS	2011				•		0.93 (0.40, 2.15)	2.93
Djoussé-WHS	2011				-		1.42 (1.29, 1.55)	11.88
Nanri	2011*						0.59 (0.18, 1.96)	1.64
Villegas-men	2011*				—		0.59 (0.43, 0.81)	8.37
Villegas-women	2011*			-			0.77 (0.63, 0.94)	10.44
He	Unpublished*						0.99 (0.52, 1.88)	4.26
Overall: <i>I</i> ² = 83.7%	, p = 0.000				\diamond		1.002 (0.85, 1.18)	100.00
* <i>de novo</i> results		125	ı .25	۱ .5	1 1	1 2		

Figure 2—Multivariable adjusted RR and 95% CI of incident diabetes. The pooled estimates were obtained using a random-effects model. The dots indicate the adjusted RRs by comparing fish consumption 2–4 times per week to never or less than once per month. The size of the shaded square is proportional to the percent weight of each study. The horizontal lines represent 95% CIs. The diamond data markers indicate the pooled RRs. (A high-quality color representation of this figure is available in the online issue.)

diabetes (39). Also, studies indicate that vitamin D may be a protective factor of diabetes development (40). Unless we clearly understand these both the nutritious and the possible harmful factors as well as their interactions in relation to diabetes risk, it may be too early for us to make any definitive recommendation on fish consumption with respect to primary prevention of diabetes.

In summary, this meta-analysis of prospective cohort studies does not find an overall association between fish consumption and diabetes risk. However, results from this meta-analysis suggest that fish consumption may be beneficial for diabetes development among Asians who had a relatively large amount of fish intake and different dietary pattern compared with Western populations. Until more studies are available, findings from this metaanalysis should not alter current advisories on fish intake or diminish the fact that fish is overall a healthy food, as a large body of evidence supports the cardioprotective effects of fish consumption.

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