

Methylation Clocks Do Not Predict Age or Alzheimer's Disease Risk Across Genetically Admixed Individuals

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Abstract

Epigenetic clocks that quantify rates of aging from DNA methylation patterns across the genome have emerged as a potential biomarker for risk of age-related diseases, like Alzheimer's disease (AD), and environmental and social stressors. However, methylation clocks have not been validated in genetically diverse cohorts. Here we evaluate a set of methylation clocks in 621 AD patients and matched controls from African American, Hispanic, and white cohorts. The clocks are less accurate at predicting age in genetically admixed individuals, especially those with substantial African ancestry, than in the white cohort. The clocks also do not consistently identify age acceleration in admixed AD cases compared to controls. Methylation QTL (meQTL) commonly influence CpGs in clocks, and these meQTL have significantly higher frequencies in African genetic ancestries. Our results demonstrate that methylation clocks often fail to predict age and AD risk beyond their training populations and suggest avenues for improving their portability.

33 1 Introduction

34 Biological aging is the progressive accumulation of cellular damage leading to degeneration
35 and organismal death (Aunan et al., 2016). DNA methylation patterns at CpG sites across
36 the genome correlate strongly with the aging process, an effect that has been quantified using
37 statistical models called “methylation clocks” (Jones, Goodman, and Kobor, 2015). The first-
38 generation of methylation clocks were trained to predict chronological age from methylation
39 levels at selected CpGs from across the genome (Hannum et al., 2013; Horvath, 2013; Zhang
40 et al., 2019). A second-generation of clocks were trained to use methylation levels to predict
41 mortality risk as proxied by a combination of biomarkers of frailty and physiological decline
42 (Morgan E. Levine et al., 2018; Lu et al., 2019). Finally, a third-generation of clocks have
43 been trained to predict the rate of aging based cohorts with longitudinal data on biomarkers of
44 frailty (Belsky et al., 2022).

45 Greater predicted DNA methylation age compared to an individual’s chronological age,
46 known as epigenetic age acceleration, has been associated with an increased risk of many
47 age-related diseases, including coronary heart disease, white matter hyperintensities, Type 2
48 diabetes mellitus, Parkinson’s disease, and Alzheimer’s disease (AD) (Hodgson et al., 2017;
49 Horvath, Gurven, et al., 2016; Horvath and Ritz, 2015; Morgan E Levine et al., 2015, 2018;
50 Lu et al., 2019; Raina et al., 2017). As such, methylation clocks show potential as predictive
51 biomarkers of the aging process and age-related health outcomes, and may capture relevant
52 biological signals associated with aging. The clocks are also increasingly being used in social
53 epidemiology research to quantify associations of epigenetic aging with exposure to adverse so-
54 cial and environmental factors that often differ across groups (Aiello et al., 2024; Chiu et al.,
55 2024; Krieger et al., 2024; Non, 2021).

56 While methylation is shaped by the environment of an individual, it is also strongly influ-
57 enced by genetic variation (Kader and Ghai, 2017). Millions of methylation quantitative trait
58 loci (meQTLs)—genetic variants that associate with the methylation level of a CpG site across
59 individuals—have been identified (Smith et al., 2014). MeQTL influence methylation levels via
60 many mechanisms, including disruption of CpGs and effects on transcription factor binding,
61 gene expression, and other gene regulatory processes (Banovich et al., 2014; Oliva et al., 2023).
62 Methylation patterns also vary between human groups, and approximately 75% of variance in
63 methylation between human groups associates with genetic ancestry (Galanter et al., 2017).
64 This suggests that meQTLs often vary in frequency in different genetic ancestries.

65 Despite these factors that lead to differential methylation levels in different genetic ances-
66 tries, methylation clocks have been developed and evaluated primarily individuals of European
67 genetic ancestry due to biases in available genomic data. We hypothesized that lack of genetic
68 diversity in the training data of methylation clocks could limit their generalizability across global
69 and admixed populations. Similar factors have posed challenges for the application of polygenic
70 risk scores (PRS) across human groups; PRS models often rapidly decrease in accuracy when
71 applied to individuals not represented in the training set (Martin et al., 2019; Novembre et al.,
72 2022; Privé et al., 2022).

73 To quantify whether current methylation clocks are generalizable across global populations,
74 we analyzed data from MAGENTA, a diverse AD study which has generated blood methylation
75 and genotyping data for 621 individuals from the Americas, including genetically admixed
76 individuals from African American, Puerto Rican, Cuban, and Peruvian cohorts. We evaluate

77 the accuracy of first-, second-, and third-generation methylation clocks at predicting age in these
78 individuals, and evaluate whether age acceleration metrics from these clocks associate with AD
79 risk, as they do in individuals of European ancestry. We also investigate the impact of genetic
80 diversity on clock CpGs by intersecting clock CpG sites with variants from different human
81 groups from gnomAD, and by comparing the frequencies of meQTL that influence clock CpG
82 sites from three sets of independent meQTLs across different genetic ancestries. Our results
83 highlight obstacles to the application of methylation clocks as biomarkers for precision medicine
84 and epidemiology, but they also identify promising avenues for considering genetic diversity in
85 the development, application, and interpretation of methylation clocks.

86 2 Results

87 Our analyses are based on genotyping and blood DNA methylation data collected by the MA-
 88 GENTA study from 621 individuals from the Americas with AD and non-demented controls.
 89 The MAGENTA study is focused on late-onset Alzheimer’s and thus the average age of partici-
 90 pants is 76 years old. Reflecting AD prevalence, the study is 68% female. The individuals come
 91 from five cohorts, collected from the United States (white, African American, Cuban), Peru,
 92 and Puerto Rico (**Table 1**). As described in detail in the Methods, to facilitate comparisons
 93 relevant to understand global differences in our study, we opted to use a combination of geo-
 94 graphic and race-based identifiers that are likely to best reflect underlying differences in genetic
 95 ancestry and admixture components.

96 We apply a range of first-, second-, and third-generation methylation-based epigenetic aging
 97 clocks to these individuals. We then evaluate their accuracy in predicting chronological age,
 98 quantify whether they identify accelerated aging in individuals with AD, and explore genetic
 99 factors that may influence clock performance (**Figure 1**).

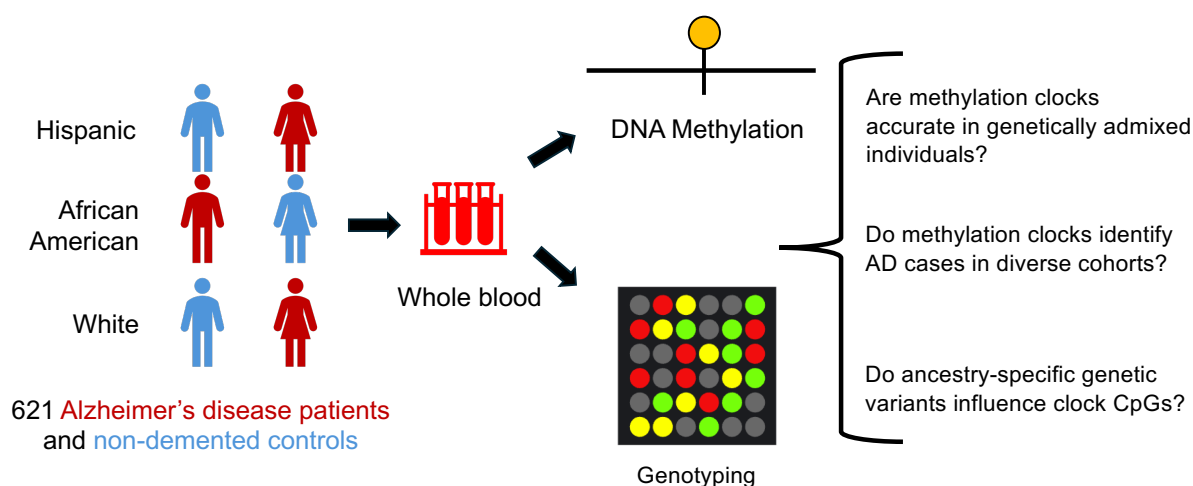


Figure 1: Schematic of the workflow of the study. We analyzed genome-wide methylation and genotyping data from blood samples from 621 AD and non-demented control individuals from the MAGENTA study. We applied a set of first-, second-, and third-generation methylation clocks to the individuals and estimated their genetic ancestry. This enabled us to explore the performance of methylation clocks in individuals with different genetic ancestries.

	Alzheimer’s (N=313)	Control (N=308)	Overall (N=621)
COHORT			
African American	98 (31.3%)	107 (34.7%)	205 (33.0%)
Cuban	22 (7.0%)	21 (6.8%)	43 (6.9%)
White	68 (21.7%)	65 (21.1%)	133 (21.4%)
Peruvian	41 (13.1%)	41 (13.3%)	82 (13.2%)
Puerto Rican	84 (26.8%)	74 (24.0%)	158 (25.4%)
SEX			
Female	206 (65.8%)	213 (69.2%)	419 (67.5%)
Male	107 (34.2%)	95 (30.8%)	202 (32.5%)

Table 1: Demographics of the MAGENTA study cohorts.

100 **2.1 Methylation clock accuracy is lower in cohorts with substantial African** 101 **ancestry**

102 To test whether current methylation clocks are able to predict age accurately in diverse, geneti-
103 cally admixed groups, we evaluated age predictions for the control individuals in the MAGENTA
104 study. We first analyzed the widely used Horvath clock, which was trained on data from several
105 tissues and cell types to accurately predict age across the lifespan using methylation levels at
106 353 CpG sites.

107 Age predicted from DNA methylation (DNAm age) using the Horvath clock was strongly
108 correlated with chronological age (Pearson $r = 0.72$) (**Figure 2A**). While this correlation is
109 lower than reported in the original study (>0.9), it is consistent with previous studies of older
110 individuals (Horvath, 2013; Marioni et al., 2015).

111 In contrast to the white cohort, the correlation between DNAm age and chronological age
112 were significantly lower for Puerto Ricans ($r = 0.45$, $p = 0.007$) and African Americans ($r =$
113 0.51 , $p = 0.016$) (**Figure 2B**). The correlations for Cubans ($r = 0.68$, $p = 0.385$) and Peruvians
114 ($r = 0.72$, $p = 0.52$) were similar in magnitude to the white cohort.

115 We noticed that the two cohorts with low correlation come from regions where individuals
116 often have substantial amounts of African ancestry. To explore how admixture levels associated
117 with the accuracy of the Horvath clock in predicting age, we estimated the global proportions
118 of African (YRI), European (CEU), and American (PEL) ancestries in each individual from the
119 MAGENTA cohort using reference groups from the 1000 Genomes Project (The 1000 Genomes
120 Project Consortium et al., 2015).

121 Methylation clock accuracy was lowest for the cohorts with substantial African ancestry:
122 African Americans (median 85% African) and Puerto Ricans (median 15% African). In contrast,
123 the clocks performed similarly to the white cohort in groups lacking substantial African ancestry:
124 Cubans (6% African) and Peruvians (2% African).

125 **2.2 Accuracy of age prediction on admixed individuals varies across methy-** 126 **lation clocks**

127 To investigate the performance of other methylation clocks at predicting chronological age in
128 admixed individuals, we selected several additional publicly available open-source clocks. We
129 considered two other “first-generation” clocks that were trained to predict chronological age:
130 the Hannum clock (Hannum et al., 2013) and a model developed by Zhang et al., 2019 that
131 used large datasets for training and achieved substantially higher performance than previous
132 age predictors. We hereafter refer to this elastic net model as “Zhang2019_EN”.

133 Both models achieved higher correlations with chronological age than the Horvath clock
134 across the cohorts in the MAGENTA study. For example, Hannum has a correlation of 0.74 in
135 the white cohort, and consistent with previous evaluations, Zhang2019_EN has a correlation of
136 0.88. These relative performance trends held across cohorts, but again, the African Americans
137 and Puerto Ricans, the cohorts with substantial African ancestry, had the lowest age correlations
138 for each clock (**Figure 2C**).

139 Next, we evaluated the PhenoAge clock, a “second-generation” clock that is trained on
140 biomarkers of frailty and physiological deterioration (Morgan E. Levine et al., 2018). The
141 correlation between DNAm age and chronological age was lower for this clock in comparison to
142 the other methylation clocks ($r = 0.53$ in the white cohort). This is likely due to the fact that

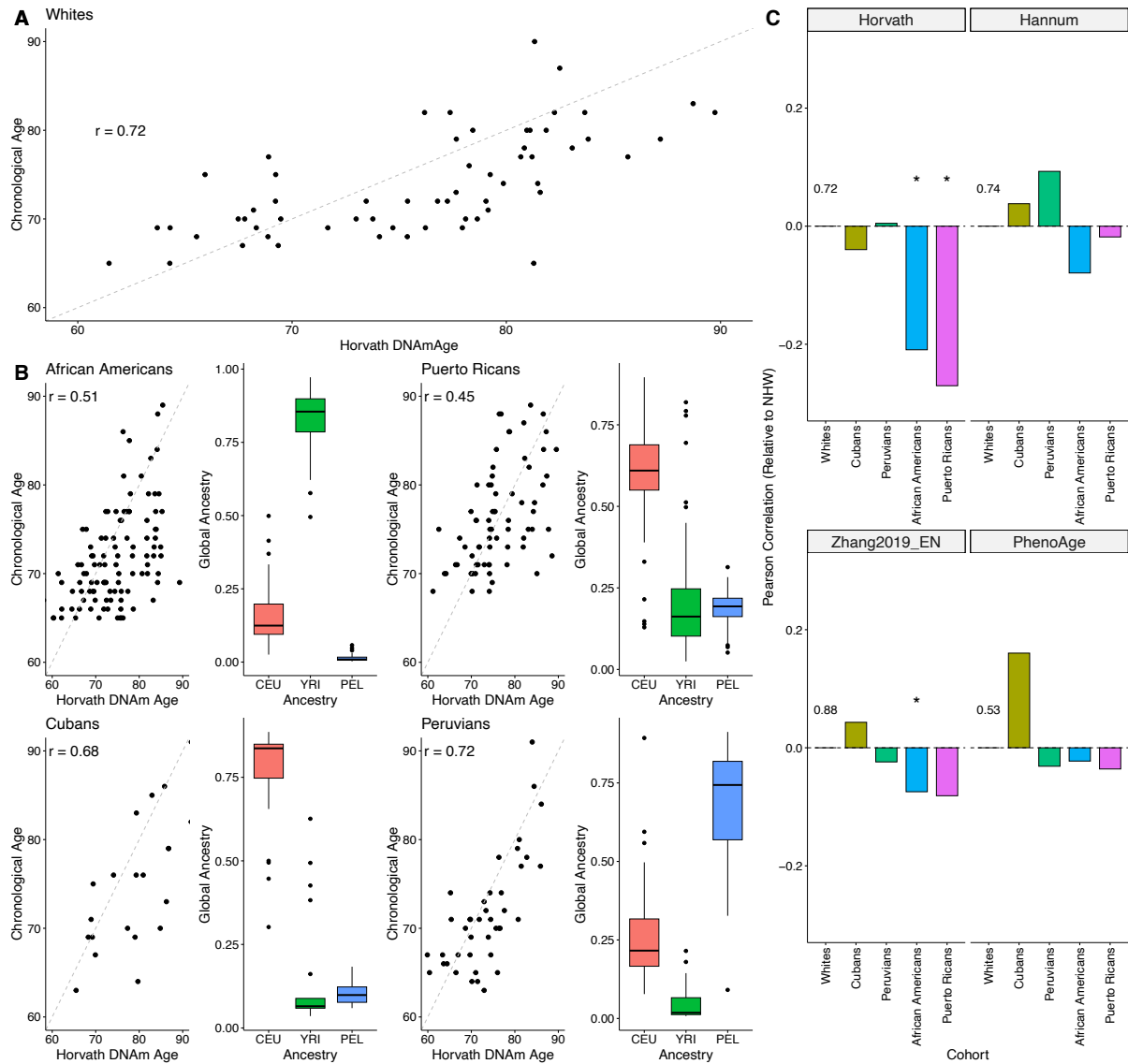


Figure 2: Methylation clock accuracy is lower in cohorts with substantial African genetic ancestry. **A:** Pearson correlation between chronological age and DNAm age predicted by the Horvath clock for controls in the white MAGENTA cohort. The correlation of 0.72 is similar to previous studies of older cohorts. **B:** Pearson correlation between chronological age and DNAm age predicted by the Horvath clock for the genetically admixed cohorts in MAGENTA. The right plot in each pair shows the proportion of European (CEU), African (YRI), and American (PEL) global ancestry for each individual in each cohort. The two cohorts with substantial African ancestry—African Americans and Puerto Ricans—have significantly lower correlations than the other cohorts. **C:** Difference in correlation between chronological and predicted DNAm age for controls in each cohort compared to the white cohort controls for four methylation clocks. The baseline correlation for the white cohort controls is given in each panel; asterisks indicate a statistically significant difference from the baseline. * $p < 0.05$.

143 this clock was not trained to predict age directly, but rather markers of aging. This clock did
144 not show as substantial a difference in performance between cohorts as seen for the Horvath
145 clock, but the African American and Puerto Rican populations again had the lowest correlation
146 of all cohorts.

147 Overall, these results demonstrate that current methylation clocks vary in the correlation of
148 their predicted DNAm age with chronological age in genetically admixed cohorts. The clocks
149 are also consistently the least accurate in predicting age cohorts with substantial proportions
150 of African ancestry.

151 **2.3 Most methylation clocks do not identify accelerated aging in admixed** 152 **Alzheimer’s cohorts**

153 DNAm age has been proposed as a promising biomarker and predictive tool for age-related
154 disease risk, particularly because of associations between accelerated DNAm age (compared to
155 chronological age) and the presence of diseases such as coronary heart disease, Parkinson’s dis-
156 ease, and AD. However, these results have largely been observed in European-ancestry cohorts.

157 To evaluate the ability of methylation clocks to identify accelerated aging and risk for age-
158 related disease in diverse, genetically admixed individuals, we quantified the association of
159 methylation age acceleration with AD status in cohorts from the MAGENTA study. In addi-
160 tion to the clocks tested in the previous section, we also included a “third generation” clock,
161 DunedinPACE, that aims to predict the pace of aging as measured by change in biomarkers
162 over time from methylation data, rather than age itself (Belsky et al., 2022).

163 The cell type composition in blood is known to change with age, which if not accounted
164 for, can confound age acceleration estimates (Jaiswal and Ebert, 2019). Thus, we focused on
165 intrinsic age acceleration estimates computed using established algorithms to correct for cell
166 type composition.

167 To establish a baseline for this analysis, we tested whether individuals with AD in the white
168 cohort show significantly greater age acceleration than non-demented controls. As expected from
169 previous studies (Morgan E Levine et al., 2015, 2018), AD cases have modest but significantly
170 greater age acceleration as measured by the Horvath clock than controls (**Supplementary**
171 **Figure 1**; median 1.7 vs. 1.5 years, $p = 0.041$). For each of the other clocks, AD cases had
172 higher median age acceleration than controls (**Figure 3B**), though due to the relatively small
173 sample size, the differences only reached statistical significance for the DunedinPACE clock
174 (median 1.09 vs. 1.07, $p = 0.044$).

175 Having established that previously reported age acceleration in AD was detectable in our
176 framework, we evaluated whether the clocks found accelerated aging in the admixed AD co-
177 horts. Focusing first on the Horvath clock, we observed inconsistent relationships between
178 age acceleration and AD status. None of the admixed cohorts showed a significant difference,
179 and controls even had higher median age acceleration in Peruvians and Cubans (**Figure 3A**).
180 Across the other clocks, none consistently identified greater age acceleration in AD cases across
181 all populations (**Figure 3B**). Among the first- and second-generation clocks, only PhenoAge
182 demonstrated a significant ability to differentiate AD cases from controls in any of the non-
183 European ancestry groups, specifically in African American individuals ($p = 0.008$), which were
184 included in its training set. Cubans consistently showed greater age acceleration in controls
185 rather than cases, while none of the other admixed cohorts even had consistent directions of

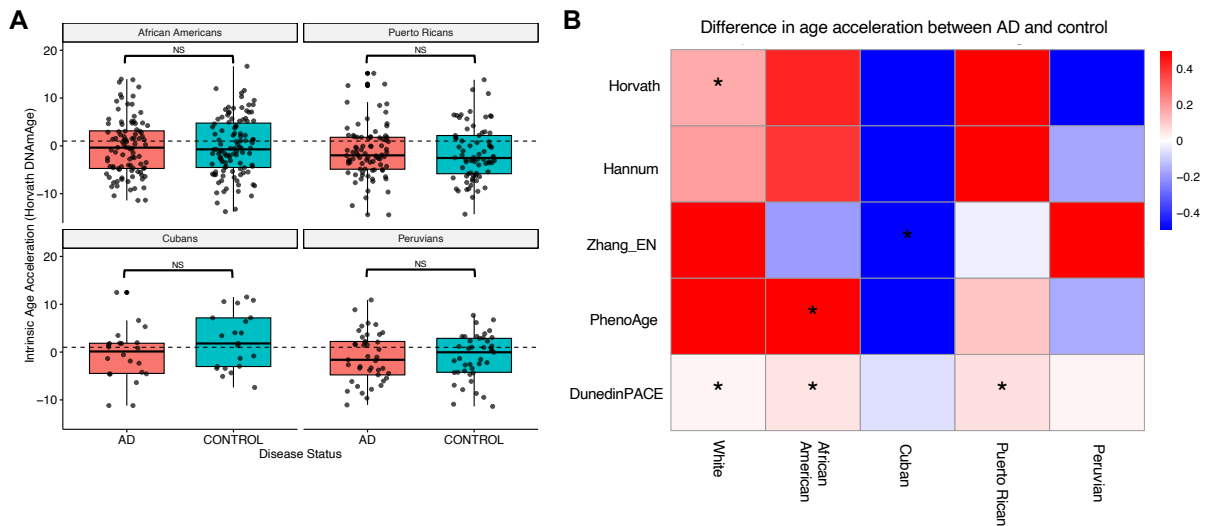


Figure 3: Methylation clocks rarely identify accelerated aging in admixed Alzheimer’s cohorts. **A:** Comparison of the distributions of Horvath intrinsic age acceleration for AD patients and non-demented controls for each of the admixed cohorts in MAGENTA. AD patients do not show significantly higher age acceleration in any of the admixed cohorts. In contrast, the AD cases had significantly greater acceleration than controls in the white cohort (**Supplementary Figure 1**). NS = Not significant. **B:** Median differences in intrinsic age acceleration between AD patients and non-demented controls for five methylation clocks for each cohort in MAGENTA. The clocks do not consistently identify accelerated aging in AD across cohorts, and the results also vary within cohorts. * $p < 0.05$.

186 effect across methods.

187 DunedinPACE stood out in the evaluation, as it identified significantly greater aging in AD
 188 cases compared to controls in the white ($p = 0.044$), African American ($p = 0.0019$), and Puerto
 189 Rican ($p = 0.0090$) cohorts using its “pace of aging” metric. However, no significant differences
 190 were found for Cubans ($p = 0.26$) or Peruvians ($p = 0.81$).

191 2.4 Combining predictions across methylation clocks does not improve their 192 performance

193 Inspired by recent work on the ensembling of PRS to better predict disease risk from genetic vari-
 194 ation across populations (Monti et al., 2024), we evaluated whether combining age predictions
 195 would lead to greater accuracy in age prediction and AD risk prediction in the admixed cohorts.
 196 To investigate this, we averaged the age predictions for each individual in the MAGENTA study
 197 across five methylation clocks: Horvath, Hannum, Zhang2019_EN, Zhang2019_BLUP, and Phe-
 198 noAge clocks. (The Zheng2019_BLUP is a variation of the Zhang2019_EN clock that does not
 199 use strong regularization.)

200 The ensemble method’s DNAm age prediction is more strongly correlated with chronological
 201 age in comparison to the Horvath and PhenoAge clocks, but it could not improve upon the best
 202 predictors (Zhang2019 clocks) across populations (**Supplementary Figure 2A**).

203 We next evaluated whether the ensemble intrinsic age acceleration estimates would associate
 204 more strongly with AD disease status relative to the standalone methylation clocks. Following
 205 the same evaluation framework as for the individual clocks, we found only one significant differ-
 206 ence in age acceleration. Cuban control individuals had significantly *lower* age acceleration than
 207 AD cases (**Supplementary Figure 2B**). Thus, a simple ensemble does not lead to stronger
 208 performance at either task in admixed cohorts.

209 **2.5 Many methylation clock CpGs are disrupted by genetic variants, but the** 210 **variants are extremely low frequency**

211 Our results so far demonstrate that existing methylation clocks do not perform consistently
212 across genetically admixed individuals with ancestries under-represented in clock training data.
213 Thus, we sought to investigate potential mechanisms underlying the decreased performance of
214 some clocks. We first quantified how often a genetic variant disrupted a CpG site included in
215 a clock. This scenario could lead to inaccuracies across cohorts given the loss of potential for
216 methylation and the ability of the site to contribute to the age prediction.

217 Of the 353 CpG sites considered in the Horvath clock, 245 (69%) have at least one disruptive
218 genetic variant observed in at least one individual in the gnomAD database of variants identified
219 in a cohort of 76,156, including thousands of individuals of non-European ancestry. However,
220 these variants are extremely rare (**Figure 4**); the average frequency is 0.0001, with the most
221 common case being a variant observed in just one individual. Only one clock CpG disrupting
222 variant had a frequency greater than 1%. Thus, genetic variation in clock CpG sites themselves
223 is unlikely to be the main cause of lack of generalization of the Horvath clock.

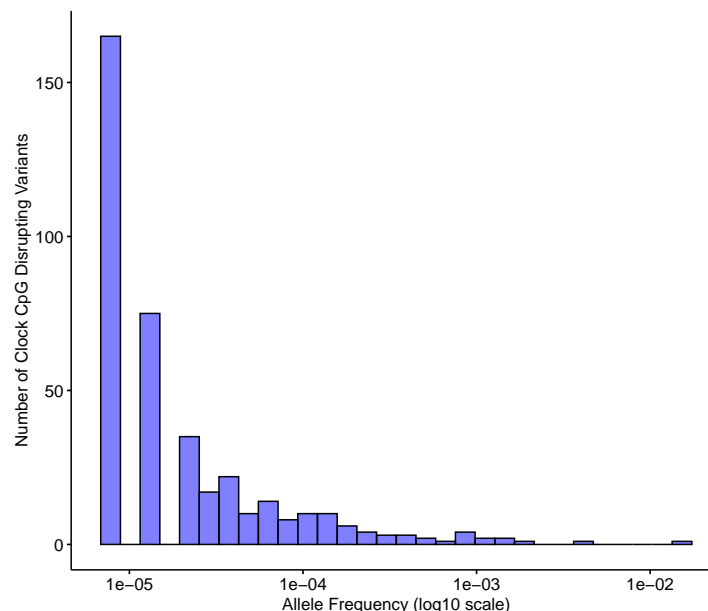


Figure 4: Genetic variants that disrupt methylation clock CpG sites appear at extremely low frequencies. The allele frequency distribution for single nucleotide variants that disrupt one of the 353 CpG sites considered by the Horvath clock. Allele frequencies were computed across 76,156 individuals from large-scale sequencing studies harmonized in gnomAD (version 3.0).

224 **2.6 Common methylation QTL influence clock CpGs**

225 We next assessed the prevalence of meQTLs that affect clock CpG sites, another modifier of
226 methylation levels that could lead to spurious DNAm age predictions across individuals. We
227 gathered three sets of meQTLs from Europeans, South Asians, and African Americans. We
228 intersected the affected CpGs for the meQTLs with the Horvath clock CpG sites.

229 Out of the 353 CpGs included in the Horvath clock, 271 (77%) had at least one meQTL.
230 Overall, a total of 29,033 unique variants associated with methylation levels at clock CpGs. The
231 meQTL had an average allele frequency of 0.19, and 26,500 were common ($\geq 1\%$; **Figure 5A**).

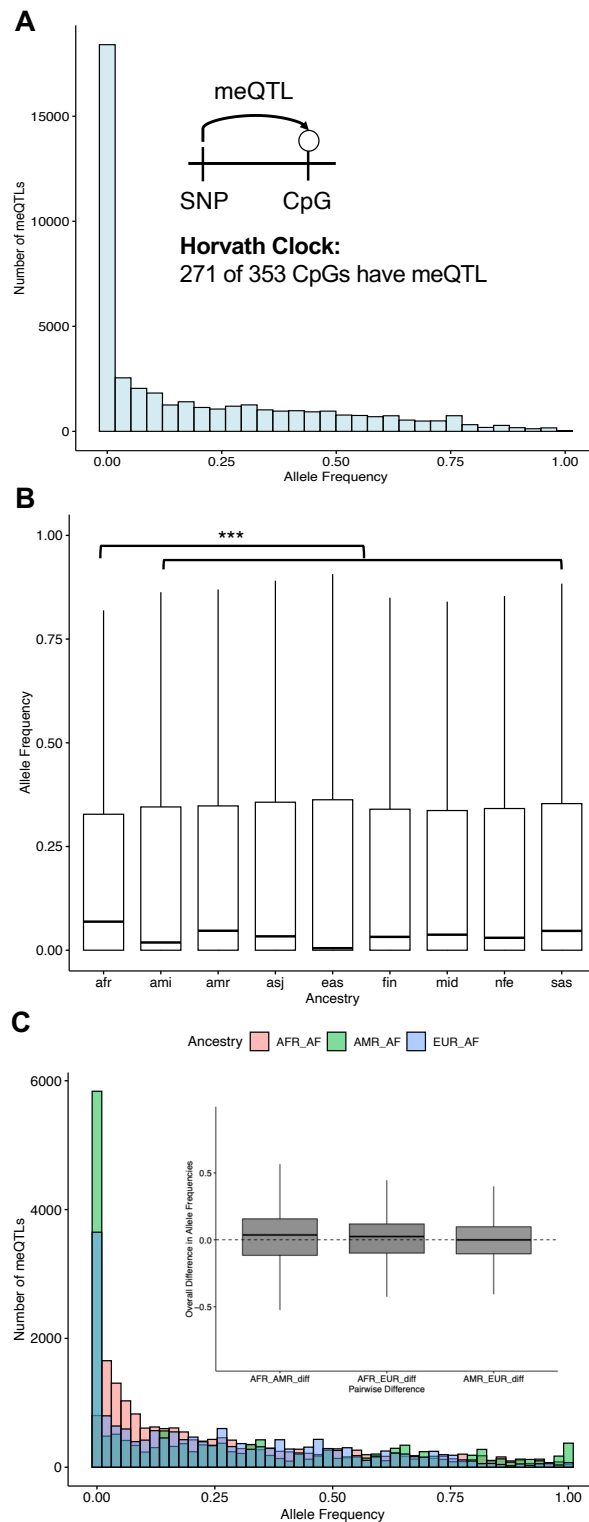


Figure 5: Common meQTLs affect most Horvath clock CpGs and vary in frequency across ancestries.

A: The allele frequency distribution of the 29,033 unique variants associated with methylation levels at Horvath clock CpGs. Allele frequencies were computed over the 76,215 genomes in gnomAD version 4.1. Inset: Out of the 353 CpGs in the Horvath clock, 271 (77%) have at least one meQTL, i.e., a genetic variant that is associated with methylation level. **B:** Clock meQTL have significantly higher allele frequency in individuals with African genetic ancestry from gnomAD than all other ancestry groups (median 0.068 for African vs. 0.004–0.046; $p < 3.85 \times 10^{-25}$). **C:** Clock meQTL have significantly higher allele frequency on African local ancestry genomic segments in 7,612 Latino admixed individuals with varying proportions of European, American, and African ancestry from gnomAD. Inset: The distribution of difference in frequencies for each meQTL for each pair of populations.

232 2.7 Clock CpG methylation QTLs vary in frequency across ancestries

233 Differences in the presence or frequency of meQTL that influence clock CpGs between genetic
234 ancestries could lead to decreased DNAm age prediction accuracy (and therefore weaker asso-
235 ciations with disease) in diverse and admixed cohorts. For example, if a clock is trained on a
236 cohort without an meQTL, the learned weights for the CpG will not have accounted for the
237 effects of the meQTL.

238 To quantify whether differences in meQTL across genetic ancestries could potentially in-
239 fluence methylation clocks, we analyzed the gnomAD allele frequencies for the 29,033 Horvath
240 clock meQTL tag variants in multiple global populations (**Figure 5B**). The meQTLs are at
241 significantly higher frequencies in African ancestry populations (median 0.068) than in each of
242 the eight other population groups considered ($0.004\text{--}0.046$; $p < 3.85 \times 10^{-25}$). There were also
243 2,328 meQTL that were only observed in African populations.

244 To connect these results to individuals with recent admixture, like many in the MAGENTA
245 study, we also tested whether the Horvath clock meQTL differed in frequency in local an-
246 cestry blocks of different origins in genetically admixed individuals from gnomAD. We used
247 pre-computed local ancestry calls for 7,612 Latino/Admixed American individuals to compare
248 allele frequencies for each meQTL in three ancestral backgrounds: African, Amerindigenous,
249 and European. The clock CpG-affecting meQTLs were at higher frequencies in African local
250 ancestry backgrounds (**Figure 5C**), consistent with our finding that these meQTLs are most
251 frequent in African ancestry individuals at the global population level.

252 2.8 Susceptibility to meQTL varies across methylation clocks

253 Given the strong potential for meQTL to influence the Horvath clock (76% of its CpGs have
254 an meQTL), we expanded this analysis and quantified the number and proportion of clock
255 CpGs that are affected by meQTL for all clocks considered here (**Figure 6**). The other first-
256 generation clocks have lower proportions of their CpG sites affected by meQTLs (Hannum: 8%,
257 Zhang2019_EN: 1%). PhenoAge is similarly low, with 7% of its CpGs affected by at least one
258 meQTL. Finally, DunedinPACE had no meQTLs affecting its 173 clock CpG sites.

259 Thus, the clock with the largest decrease in performance in admixed cohorts (in terms of
260 predicting chronological age and identifying age acceleration in AD) has the most and largest
261 fraction of meQTLs influencing its CpGs. On the opposite side of the spectrum, DunedinPACE,
262 the best performing clock at identifying AD cases in the MAGENTA study, had no meQTLs.
263 The three other clocks with intermediate performance in the admixed cohorts, all have meQTL
264 for some CpGs, but much lower fraction than the Horvath clock.

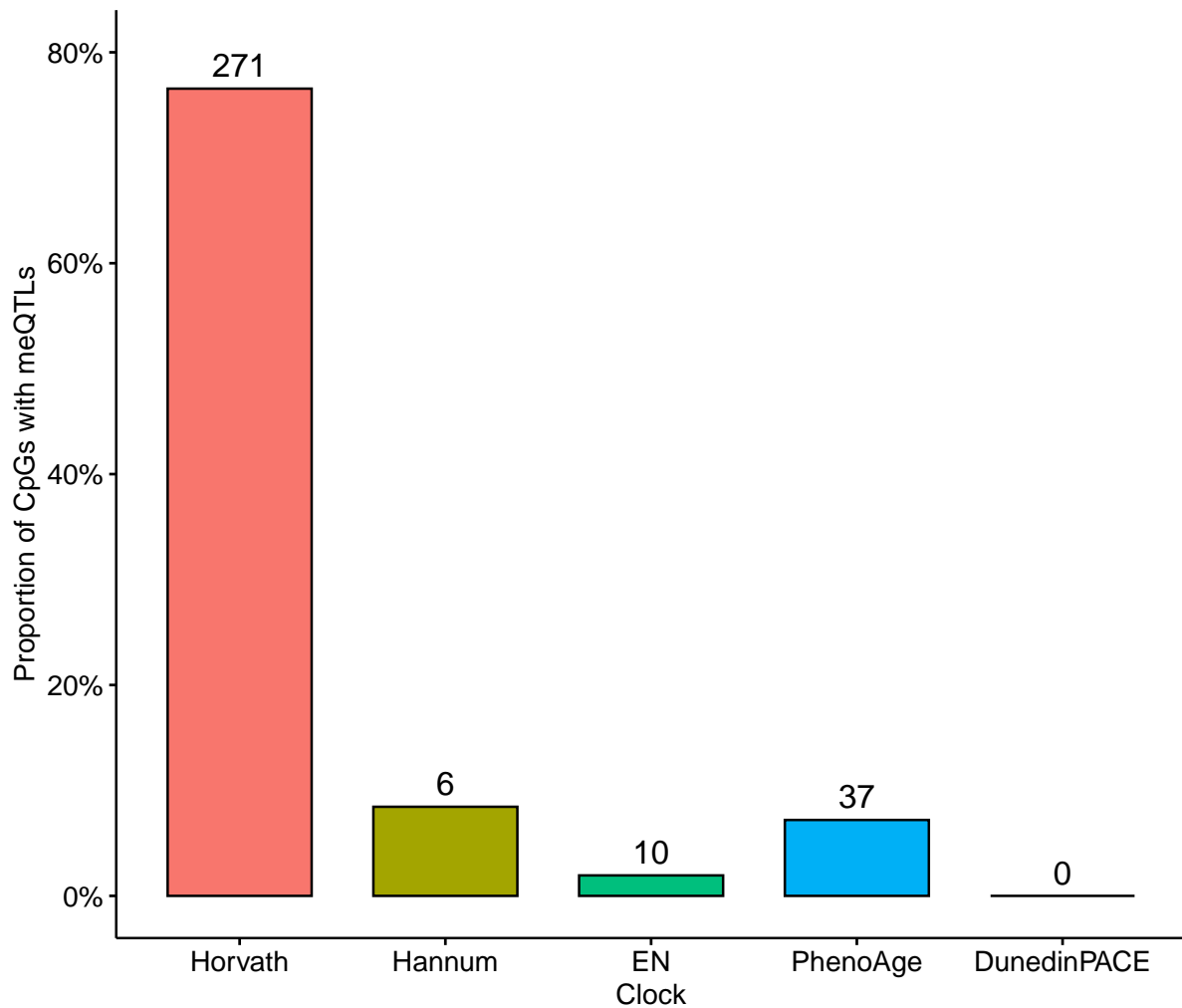


Figure 6: Methylation clocks vary in the number and proportion of CpGs affected by meQTLs. The proportion of clock CpGs for each clock that have at least one known meQTL. The number of unique clock CpGs affected by meQTLs for each clock is given on top of each bar. The meQTL were taken from three genome-wide studies in Europeans, South Asians, and African Americans.

265 3 Discussion

266 Methylation clocks are promising biomarkers of aging and social stress, and as tools for mech-
267 anistic studies of diseases related to the aging process. Despite their widespread use in these
268 applications, methylation clocks have not been comprehensively evaluated in diverse human
269 groups. These groups are underrepresented in genetic and genomic databases and underserved
270 in biomedical research in terms of access to healthcare and quality of healthcare.

271 In this study, we sought to fill this gap by evaluating commonly used methylation clocks in
272 genetically admixed individuals from the Americas. We found that most clocks did not predict
273 age as accurately in admixed control individuals as in the white cohort, especially for individuals
274 with substantial African genetic ancestry. We next found that most methylation clocks could
275 not discern between AD patients and non-demented controls based on age acceleration metrics
276 in cohorts with genetic ancestries distinct from their training populations.

277 To evaluate potential genetic factors that could contribute to this decrease in performance,
278 we hypothesized that two types of variants could reduce clock accuracy: 1) variants that disrupt
279 clock CpG sites prevent methylation, and 2) meQTLs that influence clock CpG site methylation.
280 Both scenarios could lead to over or under estimations of age, and therefore spurious associations
281 with age-related disease, if they differ in frequency across genetic ancestries. While we discovered
282 that 245 of the 353 CpG sites used by the Horvath clock are disrupted in at least one individual in
283 gnomAD, these variants are extremely rare and thus unlikely to be a major driver of differences
284 between the cohorts. In contrast, among meQTLs from multiple global populations, we found
285 29,033 unique variants that affected the Horvath clock CpG sites. Many of these variants
286 are common, and they are most frequent in African ancestry individuals from gnomAD. We
287 also showed that they are most frequent in African local ancestry blocks in admixed Latino
288 individuals from the Americas.

289 Our findings demonstrate that methylation clocks—a widespread tool in aging, genomics,
290 and social epidemiology research—perform inconsistently across individuals of different genetic
291 ancestries. These results further underline the need for more diversity in the development and
292 evaluation of genomic tools for precision medicine.

293 We hope that these results also encourage researchers using these tools to study diseases
294 of aging or social stressors in diverse groups to exercise caution when interpreting differences
295 in age acceleration. We have shown that many methylation clocks differ significantly in their
296 accuracy at predicting age between cohorts. Thus, what might appear to be a faster pace of
297 aging, could simply be the result of a difference in genetic ancestry from the training cohort.
298 Applying existing methylation clocks to diverse individuals could lead to grave consequences
299 and exacerbate existing disparities in access to quality healthcare, as well as provide spurious
300 conclusions about an individual's health. Due to the increased potential for false positives and
301 false negatives when applying the clocks as predictive biomarkers, individuals at risk might not
302 receive the medical attention they need, and additional stress could unnecessarily be placed on
303 individuals in good health. These challenges must be addressed before methylation clock are
304 adopted as biomarkers for precision medicine.

305 The challenges we identify here for methylation clocks mirror the limitations of PRS, wherein
306 phenotype prediction models decrease in accuracy on individuals genetically distant from the
307 training population. While there are substantial biological differences in the processes modeled
308 by methylation clocks and PRS, we are optimistic that recent progress on building PRS that

309 are more portable across cohorts will provide strategies for improving methylation clocks.

310 Our results suggest two promising approaches for building more robust clocks. First, given
311 the large number of meQTL in the human genome and their differences in frequency across
312 human populations, we suggest training clocks only on CpG sites that do not have known
313 meQTL. Given that the biological signatures driving methylation clock performance appear
314 to influence large fractions of the genome, we do not anticipate that this will substantially
315 limit overall performance. The strong and relatively consistent performance across cohorts
316 of the DunedinPACE clock, which lacks CpGs with meQTL, supports this approach. It also
317 suggests that methylation clocks that predict the pace of aging (rather than age itself) may
318 be more robust, but further work is needed to validate this hypothesis. Second, we encourage
319 including individuals from multiple genetic ancestries in the training cohorts. The ability of the
320 PhenoAge clock, which included African Americans in the training cohort, to detect significant
321 age acceleration in the African American AD cases suggests this may improve generalizability.

322 There are caveats and limitations to our study that we hope future work will address. First,
323 the impact of environment on methylation clock accuracy and differences in environmental
324 factors for global populations that might lead to decreases in methylation clock accuracy are
325 difficult to study with the data at hand. Given this, we focused on genetic influences on CpG
326 sites that could lead to spurious associations in diverse populations. However, work is needed to
327 investigate other factors that might cause methylation clocks to not generalize across individuals,
328 such as the methods that account for cell type composition heterogeneity in blood that might
329 not be as accurate across individuals of different populations. Second, while we attempted to
330 evaluate a representative set of first-, second-, and third-generation clocks, we were not able to
331 evaluate all methylation clocks. In particular, some with closed source that were only available
332 as a web server could not be used due to data privacy restrictions for the MAGENTA samples.
333 Another limitation is the use of blood samples to generate methylation data for the study of
334 a neurodegenerative disease focused on the central nervous system. However, we note that
335 all methylation clocks tested in the present study were developed using blood samples, either
336 exclusively (Hannum, PhenoAge, Zheng2019_EN, Zheng2019_BLUP, and DunedinPACE) or as
337 part of the tissues used in training (Horvath). In addition, these clocks and their association
338 with age-related diseases such as AD have all been validated in multiple tissues, including blood.
339 Multiple studies in the AD literature point to changes in blood such as gene expression changes,
340 immune cell type composition, and disruption of the blood brain barrier in AD patients relative
341 to non-demented controls, such that signals related to AD pathology can be identified from
342 this peripheral tissue (Griswold et al., 2020; Shigemizu et al., 2022) and through epigenetic age
343 acceleration (Hodgson et al., 2017; Marioni et al., 2015; Raina et al., 2017). Finally, while the
344 MAGENTA study is an excellent resource for exploring methylation clocks and AD in admixed
345 individuals, it is not representative of all genetic ancestries and combinations. Moreover, the
346 sample sizes vary between the MAGENTA cohorts, with the Cuban and Peruvian cohorts
347 being particularly small relative to the others. This limits our ability to find differences in the
348 age accelerations between AD patients and controls as measured by the methylation clocks.
349 However, we are reassured by the complementary findings in African cohorts on the decreased
350 performance of methylation clocks at age prediction and the role of meQTL (Meeks et al., 2024).

351 In conclusion, our results show that many existing methylation clocks have inconsistent
352 performance and limited portability across genetically admixed cohorts. We encourage future
353 efforts in the development of methylation clocks and other genomics-based aging biomarkers

354 to be genetics- and ancestry-aware to ensure the accuracy of these tools for all individuals,
355 regardless of their genetic background.

356 4 Methods

357 4.1 MAGENTA study

358 Cohort Selection

359 All participants in the MAGENTA study were recruited through previous studies of AD, in-
360 cluding Feliciano-Astacio et al., 2019, Marca-Ysabel et al., 2021, and Griswold et al., 2020.
361 Blood samples were taken for all individuals ascertained and processed at the following sites:
362 the University of Miami Miller School of Medicine (Miami, FL, US), Wake Forest University
363 (Winston-Salem, NC, US), Case Western Reserve University (Cleveland, OH, US), Universi-
364 dad Central Del Caribe (Bayamón, PR), and the Instituto Nacional de Ciencias Neurologicas
365 (Lima, PE). Ascertainment protocols were consistent across sites and capture cognitive func-
366 tion, family history of AD/related dementias, sociodemographic factors, and dementia staging.
367 All diagnoses were assigned by clinical experts following criteria for diagnosis and staging from
368 the National Institute on Aging Alzheimer’s Association (NIA-AA).

369 The MAGENTA study is based on pre-existing sample collections which vary in terms of
370 the demographic information collected for each participant. Because the original ascertainment
371 of MAGENTA study participants was international, different population descriptors were used
372 across different ascertainment sites/cohorts. To facilitate comparisons relevant to understand
373 the global differences noted in our study, we use a combination of geographic and race-based
374 identifiers that are likely to best reflect underlying differences in genetic ancestry and admixture
375 components. The label “white” is applied to legacy samples from North Carolina, Tennessee,
376 and South Florida where participants either self-identified with this descriptor or were (in some
377 legacy instances) administratively assigned as White race. The label “African American” is ap-
378 plied to samples collected via ascertainment in North Carolina and South Florida using popula-
379 tion descriptors that specifically targeted enrollment of self-identified Black/African American
380 participants. “Puerto Rican”, “Cuban”, and “Peruvian” labels are applied to samples collected
381 as part of ascertainment efforts in these geographic areas. While more precise descriptors of
382 self-identity are preferred, the advanced age of study participants and the older dates of some
383 sample collections make recontact to collect these data impossible. All participants or their
384 consenting proxy provided written informed consent as part of the study protocols approved by
385 the site-specific Institutional Review Boards.

386 Genetic data and ancestry analysis

387 Genome-wide SNP genotyping was previously performed as previously described for the MA-
388 GENTA study cohorts. Briefly, samples were genotyped on the Illumina Infinium Global
389 Screening Array using standard quality control filters on call rate, quality, missingness, and
390 Hardy-Weinberg equilibrium.

391 For our analyses, local ancestry calls were generated using the *FLARE* software (S. R.
392 Browning, Waples, and B. L. Browning, 2023) with three reference panels from the 1000
393 Genomes Project: Utah residents with Northern and Western European ancestry (CEU), Pe-
394 ruvians in Lima (PEL), and Yoruba in Ibadan, Nigeria (YRI). To estimate global ancestry
395 proportions, we summed the haplotype lengths for each ancestry in each individual and divided
396 by the total number of sites considered.

397 Methylation profiles

398 DNA methylation was quantified using the Illumina HumanMethylation EPICv2.0 according
399 to the manufacturer’s instructions. All quality control and data normalization were performed
400 using the the openSeSAmE pipeline from the SeSAmE (Wanding Zhou, 2018) tools for analyz-
401 ing Illumina Infinium DNA methylation arrays. Probes of poor design were removed from the
402 analysis as well as probes with signal detection P-value >0.05 in more than 5% of the sam-
403 ples. Non-CG probes and probes located on the X, Y, and mitochondrial chromosomes were
404 also removed. Samples with incomplete bisulfite conversion (GCT score >1.5) and principal
405 component analysis outliers were excluded. Noob normalization was performed with SeSAmE,
406 using a nonlinear dye-bias correction.

407 4.2 Estimating epigenetic age and its correlation with chronological age in 408 the MAGENTA study

409 We applied multiple commonly used first-, second-, and third-generation methylation clocks
410 to all individuals in the MAGENTA study with genome-wide methylation data. We used es-
411 tablished implementations of the Horvath, Hannum, Zhang2019_EN, Zhang2019_BLUP, and
412 PhenoAge clocks from the methylClock R library (Pelegí-Sisó et al., 2021). We also applied
413 DunedinPACE, a third-generation clock separately, because it was not included in the *methyl-*
414 *clock* library (Belsky et al., 2022). Because this clock does not explicitly predict age, it is not
415 included in the analyses of correlation with biological age. Unless otherwise specified, default
416 options were used for all clocks.

417 The methylation clocks considered analyze different numbers of CpG sites. For each clock,
418 the sites considered were taken from the *methylclock* library or the original publication. In the
419 case of missing data, the *methylClock* library imputes methylation status using the *mpute.knn*
420 function from the *impute* R library. The MAGENTA cohort had low proportions of missing
421 data for clock CpGs. Specifically, there were 3.7% missing for the Horvath clock, 12.7% for the
422 Hannum clock, 4.5% for the Zhang2019_EN clock, 3.7% for the PhenoAge clock, and 17.9% for
423 the DunedinPACE clock.

424 We computed the Pearson correlation of estimated epigenetic age and chronological age
425 for the controls in each cohort. To compare the strength of correlation between cohorts, we
426 computed p-values for the observed differences using Fisher’s z test and the Zou method for
427 computing confidence intervals as implemented in the *cocor* library (Diedenhofen and Musch,
428 2015).

429 4.3 Computing epigenetic age acceleration in Alzheimer’s disease patients 430 and controls

431 In order to quantify epigenetic age acceleration from blood methylation data, we estimated
432 raw, intrinsic, and extrinsic age acceleration for all clocks, except DunedinPACE, from the
433 methylClock library. Blood cell type composition differs between individuals and over the
434 lifespan; thus, we report results in the main text using intrinsic age acceleration estimates,
435 which capture age acceleration independently of blood cell proportions.

436 Using the age acceleration estimates, we compared epigenetic age association between AD
437 cases and matched controls using a Mann-Whitney U test, as implemented in the *stats* library

438 in R. We analyzed the study as a whole and stratified by cohort.

439 **4.4 Evaluating ensembles of age predictors**

440 We tested the performance of a simple ensemble of age prediction methods at both estimating
441 chronological age and distinguishing AD cases and controls. The ensemble was computed as
442 the average of the estimate of each method on each individual. The resulting predictions were
443 evaluated as described for each clock itself.

444 **4.5 Analysis of genetic variation at clock CpG sites**

445 To test whether the CpG sites included in the methylation clocks are variable between indi-
446 viduals, we intersected all the clock CpGs with variants identified in version 3.0 of gnomAD
447 (Karczewski et al., 2020). This database covers 76,156 individuals with whole genome sequenc-
448 ing harmonized from many large-scale sequencing studies. The intersection was performed using
449 bedtools in hg38 coordinates (Quinlan and Hall, 2010).

450 **4.6 Analysis of meQTL affecting clock CpGs**

451 **Identification of meQTL affecting clock CpGs**

452 We leveraged meQTL sets identified from blood samples by three independent studies. The first
453 study identified 11,165,559 meQTLs from 3,799 Europeans and 3,195 South Asians (Hawe et
454 al., 2022). The second study generated 4,565,687 meQTLs from 961 African Americans (Shang
455 et al., 2023). The final study (EPIGEN) identified 249,710 meQTLs from 2,358 UK individuals
456 (Villicaña et al., 2023). We filtered these sets separately on at a false discovery rate threshold of
457 0.05, correcting for multiple tests using the Benjamini-Hochberg method. These meQTL studies
458 published their results in hg19 coordinates. To integrate with genetic variation and clock CpG
459 data, we mapped the meQTL positions to hg38 using the UCSC liftOver tool (Hinrichs, 2006).
460 For each meQTL set, we intersected the target CpG site with the CpGs considered in each
461 clock and then combined across meQTL sets to generate a set of clock CpGs with evidence of
462 meQTL.

463 **Population-level allele frequencies of meQTL affecting clock CpGs**

464 We analyzed the frequency of clock CpG-affecting meQTLs within two different versions of
465 gnomAD. We used version 4.1 to quantify the allele frequencies of these meQTLs in the follow-
466 ing global populations: African, Middle Eastern, Admixed American, European (non-Finnish),
467 South Asian, Ashkenazi Jewish, East Asian, European (Finnish), Amish, and a “Remaining”
468 group defined by gnomAD as individuals that did not unambiguously cluster within these pre-
469 vious groups in a principal component analysis. We then used gnomAD version 3.1 to gather
470 allele frequencies for local ancestry blocks identified in 7,612 Latino admixed individuals with
471 varying proportions of European, Amerindigenous, and African ancestry.

472 **4.7 Data availability**

473 Raw and normalized beta matrices, along with genotyping data used in this study will be
474 made available at time of publication in the NIAGADS platform. Age predictions for all clocks

475 mentioned in this article will be made available in tab-delimited format in the same Github
476 repository in which all code used for these analyses is available.

477 **4.8 Code availability**

478 The publicly available code for analysis are available in the following repository: [https://](https://github.com/seba2550/methyl-clocks-admixture)
479 github.com/seba2550/methyl-clocks-admixture

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489 **4.10 Author Contributions**

490 Conceptualization: SCG, JAC; Methodology: SCG, JAC; Investigation: SCG, JAC; Resources:
491 EG, LG, MM, JMV, MLC, MRCO, BEFA, GSB, JLH, MAPV, AJG, WSB; Data Curation:
492 SCG, EG, LG, MM, AJG, WSB; Writing – Original Draft: SCG, JAC; Writing - Review and
493 Editing: SCG, AJG, WSB, JAC; Supervision: JAC; Project Administration: AJG, WSB, JAC;
494 Funding Acquisition: SCG, AJG, WSB, JAC.

495 **4.11 Competing interests**

496 The authors declare no competing interests.

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