



Improvement of Sepsis Prognosis by Ulinastatin: A Systematic Review and Meta-Analysis of Randomized Controlled Trials

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Background: Ulinastatin has been prescribed to treat sepsis. However, there is doubt regarding the extent of any improvement in outcomes to guide future decision making.

Objectives: To evaluate the effects of ulinastatin on mortality and related outcomes in sepsis patients.

Methods: Thirteen randomized controlled trials and two prospective studies published before September 1, 2018, that included 1358 patients with sepsis, severe sepsis, or septic shock were evaluated. The electronic databases searched in this study were PubMed, Medline, Embase, and China National Knowledge Infrastructure (CNKI) for Chinese Technical Periodicals.

Results: Ulinastatin significantly decreased the all-cause mortality {odds ratio (OR) = 0.48, 95% confidence interval (CI) [0.35, 0.66], p < 0.00001, $I^2 = 13\%$ }, Acute Physiology, Age, Chronic Health Evaluation II (APACHE II) score {mean difference (MD) = -3.18, 95%CI [-4.01, -2.35], p < 0.00001, $I^2 = 33\%$, and reduced the incidence of multiple organ dysfunction syndrome (MODS) (OR = 0.3, 95% CI [0.18, 0.49], p < 0.00001, $I^2 = 0\%$). Ulinastatin also decreased the serum levels of IL-6 (MD = -53.00, 95% CI [-95.56, -10.05], p = 0.02), TNF-a MD = -53.05, 95%CI [-68.36, -37.73], p < 0.00001, and increased the serum levels of IL-10 (MD = 37.73, 95% CI [16.92, 58.54], p = 0.0004). Ulinastatin administration did not lead to any difference in the occurrence of adverse events.

Conclusions: Ulinastatin improved all-cause mortality and other related outcomes in patients with sepsis or septic shock. The results of this meta-analysis suggest that ulinastatin may be an effective treatment for sepsis and septic shock.

Keywords: sepsis, ulinastatin, mortality, inflammatory cytokine, immune system

INTRODUCTION

Sepsis is life-threatening organ dysfunction caused by a dysregulated host response to infection (Singer et al., 2016). It is the major cause of death in intensive care units (Martin et al., 2003). Epidemiological studies in the United States have shown that 750,000 cases are diagnosed with severe sepsis annually, and 215,000 deaths occur every year (Angus et al., 2001). Owing to advances in the management of sepsis, such as early fluid resuscitation, early administration of antibiotics, and advances in supportive care, such lung-protective mechanical ventilation, the risk of sepsis-associated death has been decreasing. However, the mortality of sepsis remains high (Stoller et al., 2016).

The mechanism of sepsis is complicated (Armstrong et al., 2017; Minasyan, 2017). Sepsis initiates a complex interplay of host pro-inflammatory and anti-inflammatory processes (Hotchkiss et al., 2013; Chousterman et al., 2017). Simultaneously, both inflammatory response and immunosuppression are involved in sepsis (Hotchkiss et al., 2013; Hotchkiss and Crouser, 2015). Serum concentrations of tumor necrosis factor-a (TNF-a), interleukin-6 (IL-6), interleukin-8 (IL-8), and many other cytokines or chemokines are increased after the onset of sepsis (Chousterman et al., 2017; Rajaee et al., 2018). Immunoparalysis caused by the apoptosis of many immune cells, including T cells, B cells, dendritic cells, and neutrophils, is another dominant problem in sepsis patients; this state in turn results in a depressed immune system and failure in the elimination of pathogens and maintenance of immune balance (Hotchkiss et al., 2013; Girardot et al., 2017). Furthermore, anti-apoptosis therapy by blocking receptors or inhibitors of apoptotic pathway can reduce mortality in sepsis models (Zhang et al., 2010; Harjai et al., 2013). Blocking of programmed cell death receptor-1 (PD-1) also demonstrated a potential toward the reduction of sepsisassociated mortality (Zhang et al., 2010; Hotchkiss et al., 2013; Patera et al., 2016).

It is known that serine proteases are involved in systemic inflammation and cell apoptosis (Wong, 1998; Wiedow and Meyer-Hoffert, 2005). Urinary trypsin inhibitor (also called ulinastatin or UTI) is an important protease inhibitor found in human urine, blood, and other tissues (Linder and Russell, 2014). It has been shown that UTI plays an anti-inflammatory role by decreasing the phosphorylation of p38 mitogen-activated protein kinase (p38-MAPK) and nuclear factor-κB (NF-κB) activation as well as an anti-apoptotic role by protecting the mitochondria and scavenging oxygen free radicals (Shu et al., 2014; Li et al., 2016). The study of UTI mechanism revealed that UTI can decrease the level of inflammatory mediators and reduce the frequency of immune cell apoptosis in sepsis models. Therefore, UTI has been proposed as a potentially new therapeutic option for the treatment of sepsis and multiple organ dysfunction syndrome (MODS) (Linder and Russell, 2014; Atal and Atal, 2016).

Recently, clinical trials in sepsis patients treated with UTI or UTI combined with thymosin α 1 showed a survival benefit trend (Wu et al., 2013a; Karnad et al., 2014). Meanwhile, five metaanalysis of studies in sepsis patients either UTI administration alone or a combination of UTI and thymosin α 1 have been published (Han et al., 2015; Li et al., 2015; Feng et al., 2016; Wang et al., 2016; Liu et al., 2017). Among these five meta-analysis, two meta-analysis (Feng et al., 2016; Liu et al., 2017) analyzed the effect of using UTI alone and no significant difference between UTI group and control group in the 28-day mortality. Given that the results of these two meta-analysis are based on subgroup analysis, which included the same two trials (Wu et al., 2013a; Karnad et al., 2014), it is difficult to prove the effect of UTI alone. At present, it remains unclear whether the beneficial impact is rendered by UTI, thymosin α 1, or the combination. Therefore, we pooled the randomized controlled trials that involved the use of UTI alone in order to clarify the efficacy of UTI in sepsis.

MATERIALS AND METHODS

In accordance to the PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-Analysis) statement for reporting systematic reviews and meta-analysis (Liberati et al., 2009), two groups of the present authors [(HW and LY) and (YT and BH)] independently conducted literature searches, established the study inclusion and exclusion criteria, performed quality assessment, and extracted data. If a consensus was not reached, it was resolved by the senior authors (ZL and PC). The flow graph is shown in **Figure 1**.

Study Registration

Registration number in PROSPERO: CRD42018110751, an international prospective register of systematic reviews.

Search Strategy

We searched for randomized controlled trials in sepsis published on or before September 01, 2018, regardless of language, publication type, or study region. The electronic database includes PubMed, Medline, Embase, and China National Knowledge Infrastructure (CNKI). The key words that were searched and their combinations in (title/abstract) are shown in **Table 1**.

Selection Criteria

Studies were included if they met the following criteria: 1) participants: patients who were diagnosed with sepsis, severe sepsis, or septic shock; 2) type of interventions: use of UTI alone regardless of treatment duration; 3) research design: either randomized controlled trial (RCT) or prospective cohort study. Review articles, animal experimental studies, case reports, and letters that did not describe outcomes or were not published as full reports were excluded. In addition, we only included the most updated and completed studies in case of duplicated publication.

Outcomes and Data Extraction

The primary outcome was all-cause mortality. The secondary outcomes were changes in the serum levels of IL-6, IL-10, and TNF- α , the incidence rate of MODS, and changes in Acute Physiology, Age, Chronic Health Evaluation II (APACHE II) scores. We also collected the following information: study



TABLE 1 | Search terms and phrases used in the meta-analysis.

```
#1 ulinastatin
#2 UTI
#3 urinary protease inhibitor
#4 sepsis
#5 sept*mia
#6 effect*
#7 treatment
#8 therap*
#9 Systemic Inflammatory Response Syndrome
#10 SIRS
#11 MODS
#12 Multiple organ dysfunction syndrome
#11 #1or #2 or #3
#12 #4 or #5 or #9 or #10 or #11 or #12
#13 #11 and #12 and #6
#14 #11 and #12 and #7
```

The means of "*" is truncation.

design, year of study, country, study period, the number of patients included, intervention methods, and adverse events. The main characteristics of the included studies are shown in **Table 2**. For the continuous variables, we acquired data according to the following method. For calculating the mean in this meta-analysis, we employed the formula $X = X_2 - X_1$, where X

represents the mean applied in this meta-analysis, X1 represents the baseline mean, and X2 represents the endpoint mean. For calculating the standard deviation (SD) in this metaanalysis, we chose to employ the formula $S^2 = S_1^2 + S_2^2 - 2 \times \mathbb{R} \times S_1 \times S_2$, where S represents the standard deviation applied in this meta-analysis, S1 represents the baseline SD, and S₂ represents the endpoint SD. R = 0.5 in the meta-analysis, which was described in the Cochrane Handbook. All data were independently extracted by two authors (HW and LY). HW entered data into the computer and LY checked them.

Quality Assessment

First, two reviewers independently assessed the eligibility of articles identified during the initial search strategy. Then, the quality of all included studies was evaluated according to the modified Jadad scale, which can intuitively assess the quality of the included RCTs (Jadad et al., 1996). The studies were rated as low quality and high quality under scores of 1–3 and 4–7, respectively. In this meta-analysis, three studies were regarded as low quality and 12 studies were regarded as high quality. Moreover, detailed scoring results are shown in **Table 3**.

TABLE 2 | The characteristics of the included studies.

| Author (year published) | Country | Study period | Study Type | Total No. of patients | Num pati | ber of ients | Mean A | Age, yrs | Diagnosis | Inter | Interventions | | Outcomes | Adverse effects | Quality score |
|------------------------------|---------|---------------------|---------------|-----------------------------|--------------|------------------|------------------|------------------|-------------------------------|--------------------------|------------------------------|----------------------------------|--|--------------------------------|------------------|
| | | | | | UTI group | Control group | UTI group | Control group | | UTI group | Control group | Primary outcome | Secondary outcomes | | |
| Fang and Zhao, (2017) | China | 2013.03- 2015.05 | RCT | 96 | 49 | 47 | 56.7 ± 12.5 | 59.3 ± 11.6 | severe sepsis | 30,0000 IU q8h×5d | Antibiotics standard care | 28-day all-cause mortality | PCT, CRP, IL-6, TNF-a | none | ***** |
| Choudhuri et al. (2015) | India | 2012.10- 2014.05 | RCT | 104 | 68 | 36 | P > | 0.05 | sepsis | NR | NR | 28-day all-cause mortality | VDs, length of ICU stay, VASDs, occurrence of MODS | none | ☆☆☆ |
| Karnad et al. (2014) | India | 2009.09- 2010.06 | RCT | 114 | 55 | 59 | 37.5 ± 12.9 | 36.7 ± 12.5 | sepsis | 20,0000 IU q12h×5d | equivalent normal saline | 28-day all-cause mortality | VDs VFDs hospital stay, APACHE II score | none | ***** |
| Wu et al. (2013b) | China | 2011.10- 2012.10 | RCT | 60 | 30 | 30 | 54.3 : | ± 16.2 | sepsis | 30,0000 IU q8h×5d | equivalent normal saline | 28-day all-cause mortality | MODS,IL-10,IL-6 CD4,CD25,IL- 17,HLA-DR | none | ☆☆☆☆☆ |
| Sung et al. (2009) | Korea | 2005.01- 2008.06 | PC | 169 | 43 | 126 | 61 ± 18 | 61 ± 17 | severe sepsis septic shock | 100,0000 IU qd | Antibiotics standard care | mortality | SOFA score | none | ☆☆☆ |
| Shao et al. (2005) | China | NR | RCT | 60 | 30 | 30 | 43.3 | ± 9.2 | sepsis | 10,0000 IU q8h×5d | Antibiotics standard care | mortality | IL-6, IL-10, TNF-a, CRP | none | **** |
| Pavan Kumar et al. (2017) | India | 2014.10- 2017.10 | PO | 225 | 87 | 138 | P > | 0.05 | sepsis | 20,0000 IU q12h×5d | Antibiotics standard care | all-cause mortality | VFDs VASFDs | none | ☆☆☆ |
| Chen et al. (2015) | China | 2013.07- 2014.06 | RCT | 50 | 25 | 25 | 43.6 ± 5.8 | 41.7 ± 3.8 | severe sepsis | 20,0000 IU q12h×7 d | equivalent normal saline | 28-day all-cause mortality | IL-8,TNF-α,IL-6,IL-10 | none | ☆☆☆☆ |
| Tang et al. (2013) | China | NR | RCT | 74 | 37 | 37 | 31 | -52 | severe sepsis | 20,0000 IU q12h×7d | equivalent normal saline | all-cause mortality | IL-8, TNF-α,IL-6,IL-10 | none | ☆☆☆☆ |
| Jiang et al. (2006) | China | 2001.12- 2005.12 | RCT | 78 | 39 | 39 | 56 ± 21 | 54 ± 16 | Severe sepsis septic shock | 20,0000 IU qd×3d | equivalent normal saline | NR | IL-8,IL-1, TNF-α,IL-6 | none | **** |
| Ni et al. (2008) | China | 2006.1- 2007.2 | RCT | 42 | 21 | 21 | 60.18 ± 19.08 | 59.39 ± 21.11 | severe sepsis | 10000 IU/kg/d q12h×5d | equivalent normal saline | 28-day mortality | IL-10, TNF-α, APACHE II score | none | ☆☆☆☆☆ |
| Fang et al. (2005) | China | 2003.09- 2004.02 | RCT | 56 | 28 | 28 | 57 ± 16 | 61 ± 16 | sepsis | 20,0000 IU q12h×5d | equivalent normal saline | 28-day mortality | IL-8, TNF-α,IL-6, APACHE II score | rash | ☆☆☆☆ |
| Dai et al. (2016) | China | 2013.07- 2014.06 | RCT | 86 | 43 | 43 | 59.45 ± 6.54 | 59.32 ± 6.15 | severe sepsis | 20,0000 IU q12h×5d | equivalent normal saline | NR | IL-8, TNF-α,IL-10, APACHE II score | nausea, fatigue and rash | ☆☆☆☆☆ |
| Wang et al. (2007) | China | 2004.1- 2006.12 | RCT | 84 | 44 | 40 | 55.3 ± 24.5 | 52.1 ± 16.3 | sepsis | 20,0000 IU q12h×7d | equivalent normal saline | NR | TNF-α,IL-6,IL-10,IL- 8,IL-1 | none | *** |
| Wu et al. (2016) | China | 2011-2012 | RCT | 60 | 31 | 29 | 48.71 ± 30.15 | 50.09 ± 29.11 | sepsis | 20,0000 IU q8h×8d | equivalent normal saline | 28-day mortality | TNF-α,IL-10 APACHE II score | None | *** |

RCT, Randomized controlled study; PC, prospective case–control study, PO, prospective observational study; p > 0.05, no difference in the baseline of mean age; VFDs, ventilator-free days; VDs, ventilator days; VASFDs, vasopressor-free days; VASDs, vasopressor days; NR, It was not given in the original article.

| Study, year (Station) | Random | sequence | production | | Allocation | concealment | | - | Blinding met | po | Withdrawals | and dropouts | Score |
|-----------------------------|-------------------|---------------|------------------|----------------|---------------|------------------|--------------|----------------|---------------|------------------|-------------------|-------------------|-------|
| | Adequate 2′ | Unclear 1' | Inadequate 0' | Adequate 2' | Unclear 1' | Inadequate 0' | Unused 0' | Adequate 2' | Unclear 1' | Inadequate 0′ | Description 1' | Undescribed 0' | |
| Wang et al., 2007 | | - | | | - | | | 5 | | | | 0 | 4 |
| Jiang et al. (2006) | | | | | - | | | 2 | | | | 0 | 4 |
| Karnad et al., 2014 | | 0 | | | 0 | | | 2 | | | - | | 7 |
| Wu et al., 2013b | | 0 | | | - | | | 2 | | | | 0 | Ŋ |
| Sung et al., 2009 | | | | | - | | | | ۲ | | | 0 | က |
| Wu et al., 2016 | | | | | - | | | 2 | | | | 0 | 4 |
| Dai et al. (2016) | | 0 | | | - | | | 2 | | | | 0 | Ŋ |
| Tang et al. (2013) | | | | | - | | | 2 | | | | 0 | 4 |
| Chen et al., 2015 | | | | | - | | | 2 | | | | 0 | 4 |
| Fang et al. (2005) | | | | | - | | | 2 | | | | 0 | 4 |
| Ni et al. (2008) | | 0 | | | - | | | 2 | | | | 0 | Ŋ |
| Shao et al., 2005 | 2 | | | 0 | | | | 2 | | | | 0 | 9 |
| Fang and Zhao, 2017 | 2 | | | 0 | | | | 2 | | | | 0 | 9 |
| Pavan Kumar et al., 2017 | | - | | | | | | | - | | | 0 | ო |
| Choudhuri et al., 2015 | | - | | | - | | | | - | | | 0 | с |
| 1–3 score. low auality: 4- | -7 score. hiah au | Jalitv. | | | | | | | | | | | |

Statistical Analysis

In this meta-analysis, all statistical calculations and analysis were performed using Review Manager 5.3 (Cochrane collaboration, Oxford, UK). According to the results of statistical analysis, we divided the type of data into dichotomous and continuous. For the dichotomous data, such as mortality, the incidence of MODS, and adverse events, we calculated the odds ratios (OR), 95% confidence intervals (CIs) of every included study, and the overall Mantel–Haenszel (M-H). For the continuous data, such as IL-6, IL-10, and TNF- α levels and the APACHE II score, we calculated the mean difference (MD) and 95% CIs. The statistical heterogeneity was examined using chi-square and I² statistical tests as well as P values. At first, we used a fixed-effects model, but then chose to employ the random-effects model if I² was \geq 50%.

Because patients with sepsis have a high mortality rate and not all studies report a 28-day mortality rate, we chose all-cause mortality to be our primary outcome. Sensitivity analysis was used to judge whether the study results were statistically significant. For eliminating publication bias, we used the funnel plot method.

RESULTS

Description of Eligible Studies

We identified 15 (Fang and Chen, 2005; Jiang et al., 2006; Wang et al., 2007; Ni et al., 2008; Sung et al., 2009; Tang, 2013; Wu et al., 2013b; Karnad et al., 2014; Chen et al., 2015; Dai and Wang, 2016; Wu et al., 2016; Fang and Zhao, 2017; Choudhuri et al., 2015; Shao et al., 2005; Pavan Kumar et al., 2017) potential studies that included a total of 1358 patients: 630 patients in the UTI group and 728 patients in the control group. Thirteen RCTs and two prospective studies were included in this meta-analysis. The specific method for identifying studies and establishing the inclusion and exclusion criteria is shown in **Figure 1**. Eleven studies were published from China, three from India, and one from Korea.

Primary Outcomes All-Cause Mortality

We extracted the data from 12 studies (Fang and Chen, 2005; Shao et al., 2005; Ni et al., 2008; Sung et al., 2009; Tang, 2013; Wu et al., 2013b; Karnad et al., 2014; Chen et al., 2015; Choudhuri et al., 2015; Wu et al., 2016; Fang and Zhao, 2017; Pavan Kumar Rao et al., 2017) and 1110 participants were classified into two groups to assess all-cause mortality. All-cause mortality was significantly lower in the UTI group than in the control group (OR = 0.48, 95% CI [0.35, 0.66], p < 0.00001), and heterogeneity was low ($x^2 = 12.57$, p = 0.32, $I^2 = 13\%$). The results are shown in **Figure 2**.

Secondary Outcomes Levels of IL-6

We obtained the related data from eight studies (Fang and Chen, 2005; Shao et al., 2005; Jiang et al., 2006; Wang et al., 2007; Tang, 2013; Wu et al., 2013b; Chen et al., 2015; Fang and Zhao, 2017) (558 participants in two groups) to analyze the serum levels of IL-6. The serum level of IL-6 at the time of hospital admission was not different between the UTI and control groups. After

| | UTI | | Contro | ol | | Odds Ratio | Odds Ratio |
|---|---|----------|-----------|----------------------|---------|--------------------|-------------------------|
| _Study or Subg | roup Events | Total | Events | Total | Weight | M-H, Random, 95% C | CI M-H, Random, 95% CI |
| Chen 2015 | 5 | 25 | 12 | 25 | 5.7% | 0.27 [0.08, 0.95] | |
| Choudhuri 2015 | 26 | 68 | 15 | 36 | 12.0% | 0.87 [0.38, 1.97] | |
| Fang 2005 | 8 | 28 | 16 | 28 | 7.2% | 0.30 [0.10, 0.91] | · · · · · · |
| Fang 2017 | 13 | 49 | 17 | 47 | 10.9% | 0.64 [0.27, 1.52] | - _ |
| Karnad 2014 | 4 | 55 | 12 | 59 | 6.2% | 0.31 [0.09, 1.02] | |
| Ni 2008 | 6 | 21 | 14 | 21 | 5.3% | 0.20 [0.05, 0.74] | |
| Pavan Kumar R | ao N 2017 18 | 87 | 50 | 138 | 18.4% | 0.46 [0.25, 0.86] |] |
| Shao 2005 | 1 | 30 | 6 | 30 | 2.0% | 0.14 [0.02, 1.23] |] |
| Sung 2009 | 8 | 43 | 34 | 126 | 11.1% | 0.62 [0.26, 1.47] | |
| Tang 2013 | 8 | 37 | 19 | 37 | 8.4% | 0.26 [0.09, 0.72] | |
| Wu 2013 | 5 | 30 | 6 | 30 | 5.3% | 0.80 [0.22, 2.97] | |
| Wu 2016 | 11 | 31 | 9 | 29 | 7.6% | 1.22 [0.42, 3.59] | |
| Total (95% CI) | | 504 | | 606 | 100.0% | 0.48 [0.35, 0.66] | ▲ |
| Total events | 113 | | 210 | | | | |
| Heterogeneity: | Tau ² = 0.04; Chi ² = 12. | 57, df = | 11 (P = 0 | .32); l ^a | ² = 13% | | |
| Test for overall | effect: Z = 4.56 (P < 0. | 00001) | | | | | UTI group Control group |
| IGURE 2 All-cause mortality of t | he included studie: | s. | | | | | |

treatment, IL-6 was significantly less in the UTI group than in the control group (MD = -53.00, 95% CI [-95.56,-10.05], p = 0.02), and a obvious heterogeneity in the results was observed ($x^2 = 410.24$, p < 0.00001, I² = 98%). The results are shown in **Figure 3A**.

Levels of TNF- $\!\alpha$

We collected the related data from 10 studies (Fang and Chen, 2005; Shao et al., 2005; Jiang et al., 2006; Wang et al., 2007; Ni et al., 2008; Tang, 2013; Chen et al., 2015; Dai and Wang, 2016; Wu et al., 2016; Fang and Zhao, 2017) (686 participants in two groups)

| Study or Subgroup | Mean | 0.1 | SD T | otal | Mean | SD | Tota | Weidl | t IV Random | 95% CI | N Bandom 95% Cl | |
|-----------------------------------|-------------|--------------------------|--------------|--------|----------|----------------------|------|---------|----------------------|------------|---|-----|
| Chen 2015 | -167.2 | 6 | 4 33 | 25 | -114 | 71 58 | 24 | 123 | -53 20 L 00 03 | 3 -15 471 | | |
| Fang 2005 | -136 //8 | 13 | 4.33 7 NG | 23 | -79.95 | 00.30 | 29 | 10.6 | -56 53 [.110 | 25 6 10 | | |
| Fang 2003 | -95.7 | 26 32850 | 167 | 10 | -19.6 | 30.07378 | 47 | 13.5 | -3610[4743 | 2 - 24 771 | - | |
| liang 2006 | -03.7 | 20.52050 | 7 47 | 30 | -43.0 | 109 845 | 30 | 11.8 | .3930685 | 30 6 701 | | |
| Shao 2005 | -171.8 | 23 10713 | 526 | 30 | -34.2 | 27 01 20 | 30 | 13.5 | -137 60 L150 32 | -124 881 | ÷ * | |
| Tang 2013 | -164.03 | 20.107.10 | 7 04 | 37 | 115 22 | 72.89 | 37 | 12.7 | .40 70 [.91 61 | 1 .17 70 | | |
| Wang 2013 | -119.5 | a | 7 27 | 44 | -74 6 | 105.43 | 40 | 12.0 | -44 90 1-88 4 | 11 -1 301 | | |
| Wu 2013b | -11.96 | 3 | 6.58 | 30 | -6.75 | 7.4 | 30 | 13.6 | | 751.671 | - | |
| | | | | | | | | | | -,, | | |
| Total (95% CI) | | | | 282 | | | 276 | 100.0 | -53.00 [-95.96 | 6, -10.05] | - | |
| Heterogeneity: Tau ² = | = 3520.11; | Chi ² = 410 | .24, df= | = 7 (P | < 0.0000 | 1); I² = 98% | | | | | -200 -100 0 100 | 200 |
| Test for overall effect | Z = 2.42 (| (P = 0.02) | | | | | | | | | UTI Control | 200 |
| 2 | | | | | | | | | | | | |
|) | | UTI | | | C | Control | | | Mean Differer | nce | Mean Difference | |
| Study or Subgroup | Mean | _ | SD T | otal | Mean | SD | Tota | Weigh | nt IV, Random | n, 95% CI | IV, Random, 95% Cl | |
| Chen 2015 | -230 | 102.6115 | 5491 | 25 | -128.5 | 95.31574 | 25 | 5.7 | % -101.50 [-156.40 | 0, -46.60] | | |
| Dai 2016 | -55.81 | 15.96499 | 9922 | 43 | -25.66 | 15.43758 | 43 | 20.5 | % -30.15 [-36.79 | 9, -23.51] | - | |
| Fang 2005 | -204.4 | 195.6633 | 3351 | 28 | 116.47 | 203.666 | 28 | 1.9 | % -87.93 [-192.5 | 4, 16.68] | | |
| Fang 2017 | -70.2 | 23.73120 | 309 | 49 | -44.9 | 25.43305 | 47 | 19.6 | % -25.30 [-35.16 | 5, -15.45] | • | |
| Jiang 2006 | -151.9 | 156.1759 | 9585 | 39 | -49.2 | 159.7861 | 39 | 3.9 | % -102.70 [-172.82 | 2, -32.58] | | |
| Ni 2008 | -90.48 | 64.30217 | 803 | 21 | -38.97 | 71.86714 | 21 | 8.4 | % -51.51 [-92.76 | 6, -10.26] | | |
| Shao 2005 | -41.66 | 17.2192 | 2102 | 30 | -10.43 | 18.98478 | 30 | 19.8 | % -31.23 [-40.40 | 0, -22.06] | • | |
| Tang 2013 | -262.3 | 104.5263 | 8407 | 37 | 147.47 | 114.234 | 37 | 6.5 | % -114.83 [-164.72 | 2, -64.94] | | |
| Wang 2007 | -203.1 | 155.9627 | 199 | 44 | -27 | 158.433 | 40 | 4.2 | % -176.10 [-243.44, | -108.76] | | |
| Wu 2016 | -145.06 | 71.25186 | 6594 | 31 | -91.76 | 73.30297 | 29 | 9.6 | % -53.30 [-89.92 | 2, -16.68] | | |
| Total (95% CI) | | | | 347 | | | 330 | 100.0 | 63 05 [68 36 | 37 731 | • | |
| Heterogeneitr Tau ² | - 286 88. (| hi ² = 43.59 | R df = 0 | (P < f | 000011 | 1 ² = 70% | 555 | 100.0 | -55.05 [-00.50 | , | -+-++++++++++++++++++++++++++++++++++++ | -+ |
| Test for overall effect | 7 = 6.79 | P < 0.0000 | 1) | 0 - 0 | | 1 - 1 5 /0 | | | | | -200 -100 0 100 | 200 |
| | | | ., | | | | | | | | UTI Control | |
| 2 | | | | | | | | | | | | |
| - | | UTI | | | Contro | bl | | | lean Difference | | Mean Difference | |
| Study or Subgroup | Mean | SD | Total | Mean | | SD Tota | I We | eight I | V, Random, 95% Cl | | IV, Random, 95% Cl | _ |
| Chen 2015 | 81.2 | 58.33 | 25 | 34.1 | 2 | 5.67 2 | 5 11 | 0.1% | 47.10 [22.12, 72.08] | | | |
| Dai 2016 | 61.66 | 22.99 | 43 | 27.03 | 1 | 8.22 4 | 3 11 | .5% : | 34.63 [25.86, 43.40] | | - | |
| Fang 2005 | 90.5 | 57.9 | 28 | 11.9 | 1 | 6.85 2 | 8 10 | 0.4% 7 | 3.60 [56.26, 100.94] | | | |
| Ni 2008 | 77.17 | 30.21 | 21 | 15.13 | 1 | 9.45 2 | 1 11 | .1% (| 62.04 [46.67, 77.41] | | | |
| Shao 2005 | 7.18 | 17.83984 | 30 | 1.56 | 18.965 | 5398 3 | 0 11 | .5% | 5.62 [-3.70, 14.94] | | * | |
| Tang 2013 | 75.2 | 59.84 | 37 | 28 | 2 | 1.91 3 | 7 10 | 0.6% | 7.20 [26.67, 67.73] | | | |
| Wang 2007 | 46.3 | 19.96 | 44 | 34.5 | 2 | 1.27 4 | 0 11 | .5% | 11.80 [2.96, 20.64] | | - | |
| Wu 2013b | -9.57 | 7.18 | 30 | -5.08 | | 7.88 3 | 0 1 | .7% | -4.49 [-8.30, -0.68] | | • | |
| Wu 2016 | 70.9 | 20.92 | 31 | 6.1 | | 13.2 2 | 9 11 | .5% (| 64.80 [56.01, 73.59] | | - | |
| Total (95% CI) | | | 289 | | | 28 | 3 10 | 0.0% 3 | 7.73 [16.92, 58.54] | | • | |
| Heterogeneity: Tau ² : | = 956.74 0 | Chi ² = 329 ! | 55. df = | 8 (P < | 0.00001 |): 2 = 98% | | | | | | _ |
| Test for overall effect | Z = 3.55 | (P = 0.0004) | 4) | - 0 | | | | | | -1 | 100 -50 0 50 100 | |
| | | | | | | | | | | | Control UTI | |

inflammatory factors: IL-10 (C).

to analyze the serum levels of TNF- α . The level of TNF- α at the time of hospital admission was not different between the UTI and control groups. After treatment, TNF- α was significantly less in the UTI group than in the control group (MD = -53.05, 95%CI [-68.36,-37.73], p < 0.00001), and an obvious heterogeneity was observed in the results (x² = 43.58, p < 0.00001, I² = 79%). The results are shown in **Figure 3B**.

Levels of IL-10

We gained the related data from nine studies (Fang and Chen, 2005; Shao et al., 2005; Wang et al., 2007; Ni et al., 2008; Tang, 2013; Wu et al., 2013b; Chen et al., 2015; Dai and Wang, 2016; Wu et al., 2016) (572 participants in two groups) to analyze the serum levels of IL-10. The serum level of IL-10 at the time of hospital admission was not different between the UTI and control groups. After treatment, IL-10 was significantly greater in the UTI group than in the control group (MD = 37.73, 95% CI [16.92, 58.54], p = 0.0004), and an obvious heterogeneity was observed in the results ($x^2 = 329.55$, p < 0.0001, $I^2 = 98\%$). The results are shown in **Figure 3C**.

The Apache II Score

The APACHE II score at the time of hospital admission were not different between the UTI and control groups. We extracted the data from four studies (Ni et al., 2008; Fang et al., 2005; Dai and Wang, 2016; Wu et al., 2016) (244 participants in two groups) to assess this change. After treatment, the APACHE II scores were significantly less in the UTI group than in the control group MD = -3.18, 95%CI [-4.01, -2.35], p < 0.00001), and heterogeneity was low ($x^2 = 4.51$, p = 0.21, I² = 33%). The results are shown in **Figure 4**.

The Incidence of MODS

We extracted the data from three studies (Shao et al., 2005; Karnad et al., 2014; Pavan Kumar Rao et al., 2017) (399

participants in included in two groups) to assess the incidence of MODS. After treatment, the incidence of MODS was significantly less in the UTI groups than in the control groups (OR = 0.3, 95% CI [0.18–0.49], p < 0.00001), and heterogeneity was not observed in the results ($x^2 = 0.58$, p = 0.75, I² = 0%). The result is shown in **Figure 5**.

Publication Bias and Sensitivity Analysis

All included studies involved the use of UTI for treating sepsis patients. Using all-cause mortality as the main variable, the included studies were evaluated for the effect of study size. The funnel plot demonstrated a balanced and or a symmetrical shape, suggesting no significant publication bias. In addition, Egger's test also demonstrated a statistically significant symmetry (p = 0.183). Therefore, the potential publication bias had no significant influence on the results (**Figure 6**).

Sensitivity analysis was conducted by the leave-one-out method and checking the consistency of the overall effect estimate. For IL-6, we found that the I² value decreased to 0% after excluding the studies conducted by Shao et al. (2005) and Wu et al. (2013b). For TNF- α , we found that the I² value decreased to 54% after excluding the studies conducted by Tang et al. (2013) and Wang et al. (2007). For IL-10, we found that the I² value decreased to 84% after removing the study by Shao et al. (2005), Wang et al. (2007) and Wu et al. (2013b). We believe that the high heterogeneity may arise from factors such as sample size, different measuring instruments, and design methods.

DISCUSSION

UTI is a multifunctional Kunitz-type serine protease inhibitor found in human urine and blood. UTI is a member of intera-inhibitor (IaI) family, which is produced by hepatocytes (Linder and Russell, 2014). It was originally used to treat acute pancreatitis or hyperthermia (Itaba et al., 2013;







Zhang et al., 2016). Subsequently, more studies revealed its use in anti-inflammation and the protection of liver function as well as in cardiopulmonary bypass and lung disease treatment (Song et al., 2011; Li et al., 2016). In such conditions, UTI can inhibit the inflammatory response, scavenge oxygen free radicals, and shorten the time of tracheal intubation and ventilation (Yang et al., 2011; Hui et al., 2014). At present, UTI has been evidenced to provide an attractive "rescue" therapeutic option for endotoxin-related inflammatory disorders such as disseminated intravascular coagulation (DIC), acute lung injury, and acute liver injury (Inoue and Takano, 2010). Recently, UTI has been demonstrated to play a vital role in sepsis. It is well known that the immune state of sepsis patients undergoes complex changes from the onset of hyper-inflammatory response in the early phase to the immune paralysis in the late phase. To date, no drug has been specifically approved to treat sepsis in human. Recent studies show that UTI has the capacity to reduce inflammation and protect cells and has a potential survival benefit in sepsis and MODS (Linder and Russell, 2014; Atal and Atal, 2016; Chang et al., 2017a). In the past, the research on UTI treatment for sepsis has mainly been conducted in China, and the results suggested UTI administration can reduce sepsis patient mortality. In 2014, a randomized, controlled, double-blind, and multi-center trial was conducted in India, which revealed the survival benefit of UTI in patients with sepsis (Karnad et al., 2014). A meta-analysis of the effects of UTI combined with thymosin a1 revealed a reduction in mortality. To better clarify the efficacy of UTI or thymosin $\alpha 1$ administration on sepsis patients, we pooled the RCTs involving

treatment with UTI alone. In this meta-analysis, 13 relevant RCTs from three countries and two prospective studies were included. The results showed that in patients with sepsis or septic shock UTI was associated with a significant decrease in all-cause mortality and improvements in both inflammatory cytokine profiles and APACHE II scores. Mortality is the most important index for efficacy evaluation. Studies conducted in sepsis models support that UTI is capable of reducing sepsis-related mortality. The survival benefits were also observed in most clinical trials conducted in different countries. Several trials failed to achieve positive results, probably owing to an insufficient sample size and differences in patients, trials design, and other clinical factors.

Anti-inflammation is one of the most important properties of UTI. It is well known that systemic inflammatory response plays a key role in organ damage or death in sepsis. Agents directed at a single inflammatory mediator have not been shown to have a protective effect in sepsis patients. These results suggest that a single anti-inflammatory agent cannot disrupt the complicated inflammatory network. However, the removal of blood mediators by continuous renal replacement therapy (CRRT) facilitated the achievement of a survival benefit in patients with sepsis or septic shock (Liu et al., 2011; Servillo et al., 2013). Similarly with CRRT, UTI demonstrated a capacity to decrease diverse inflammatory mediator factors such as IL-1, IL-8, IL-6, HMGB1, and other mediators. UTI also inhibits inflammation by suppressing the infiltration of neutrophils and release of elastase and inflammatory mediators from neutrophils. UTI can also suppress MAPKsignaling pathway, which mediates the release of inflammatory

cytokines such as TNF- α , IL-1, and IL-6 (Inoue and Takano, 2010; Fang et al., 2018). Recently, a retrospective study of 263 critically ill patients with sepsis found that 28-day mortality decreased significantly with UTI (Xu et al., 2018). The authors concluded that 35% of the total effect of UTI was associated with the reduction in C-reactive protein (CRP), a major marker of inflammation.

Anti-apoptosis is another property of UTI. It has been observed that UTI can reduce apoptosis of endothelial cells, lymphocytes, intestinal epithelium, neurons, and renal cells during different diseases and in animal models (Li et al., 2014). It is well known that there are many immune cells, including lymphocytes, monocytes, and dendritic cells that undergo apoptosis during sepsis. It has been confirmed that UTI can protect cells from apoptosis though antioxidation and reduction of mitochondrial damage. It is known that apoptosis contributes to immunoparalysis and death of sepsis patients (Hotchkiss and Nicholson, 2006; Hotchkiss et al., 2013; Chang et al., 2017b). Anti-apoptosis in sepsis models, via increase in Bcl-2 expression or blocking of CD95, reduced the incidence of sepsis-related mortality (Hotchkiss and Nicholson, 2006; Zhang et al., 2010; Sun et al., 2011; Liu et al., 2013). In clinical trials, antiimmune cell apoptosis with anti-PD-1 or anti-PD-L1 also showed potential in sepsis treatment (Zhang et al., 2010; Patera et al., 2016). These studies suggest that cell protection may also be involved in UTI-related survival benefit in patients with sepsis.

LIMITATIONS

Although this meta-analysis reveals the potential benefits of UTI inpatients with sepsis, these trials were conducted mainly in single centers and the sample sizes were small. Recently, a retrospective observational study conducted in a single intensive care unit (ICU) by Uchida et al. (Uchida et al., 2018) found that UTI was not associated with a mortality benefit in elderly patients with established multiple organ failure from a variety of causes, only a minority of which were sepsis related. However, UTI use was associated with reduced time on both mechanical ventilators and vasoactive drugs. Thus, multicenter, large sample, randomized clinical trials are still urgently needed to further evaluate the effects of UTI in patients

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with sepsis. At present, ADJunctive Ulinastatin in Sepsis Treatment in China (ADJUST study), a large sample, multi-center, doubleblind, randomized, parallel-group, placebo-controlled trial is being conducted in mainland China (Jiang et al., 2018). The aim of this trial is to further evaluate the efficacy and safety profiles of UTI.

CONCLUSIONS

UTI is associated with reductions in both all-cause mortality and the incidence of MODS, and improvements in both APACHE II scores and inflammatory cytokine profiles in patients with sepsis, severe sepsis, or septic shock. Large high quality RCTs are needed to confirm these promising results of UTI in sepsis and septic shock.

DATA AVAILABILITY STATEMENT

The data analyzed in this study was obtained from PubMed, Medline, Embase, and China National Knowledge Infrastructure (CNKI), the following licenses apply. Requests to access these datasets should be directed to HW, 970092671@qq.com.

AUTHOR CONTRIBUTIONS

ZL conceived and designed the study. HW, LY, YT, BH, ZL, and PC conducted the literature search, read initial abstracts, extracted data from potential eligible studies, and conducted the statistical analyses. HW and BL wrote the first draft of the manuscript. ZL, PC, RL, and BL contributed with manuscript writing, concrete suggestions, and manuscript revision.

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Conflict of Interest: The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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