



Review

Depression as a Risk Factor for Alzheimer's Disease: A Systematic Review of Longitudinal Meta-Analyses

Olalla Sáiz-Vázquez ¹, Patricia Gracia-García ², Silvia Ubillos-Landa ^{3,*}, Alicia Puente-Martínez ⁴,
Silvia Casado-Yusta ⁵, Beatriz Olaya ^{6,7} and Javier Santabárbara ^{7,8,9}

- ¹ Department of Occupational Therapy, Faculty of Health Science, University of Burgos, C/Paseo de los Comendadores, Hospital Militar, 1, 09001 Burgos, Spain; osaiz@ubu.es
 - ² Psychiatry Service, Hospital Universitario Miguel Servet, 50009 Zaragoza, Spain; pgraciag@salud.aragon.es
 - ³ Department of Social Psychology, Faculty of Health Science, University of Burgos, C/Villadiego, 1, 09001 Burgos, Spain
 - ⁴ Department of Social Psychology, Faculty of Health Science, University of Burgos, C/Paseo de los Comendadores, Hospital Militar, 1, 09001 Burgos, Spain; alicia.puente@ehu.es
 - ⁵ Department of Applied Economy, Faculty of Economics and Business Sciences, University of Burgos, Pza. De la Infanta D^a Elena, 09001 Burgos, Spain; scasado@ubu.es
 - ⁶ Research, Innovation and Teaching Unit, Parc Sanitari Sant Joan de Déu, Universitat de Barcelona, Carrer Doctor Pujadas 42, 08830 Sant Boi de Llobregat, Spain; beatriz.olaya@pssjd.org
 - ⁷ Centro de Investigación Biomédica en Red de Salud Mental (CIBERSAM), Ministry of Science and Innovation, Av. Monforte de Lemos 3-5, Pabellón 11, Planta 0, 28029 Madrid, Spain; jsantabarbara@unizar.es
 - ⁸ Department of Microbiology, Pediatrics, Radiology and Public Health, University of Zaragoza, C/Domingo Miral s/n, 50009 Zaragoza, Spain
 - ⁹ Aragonese Institute of Health Sciences (IIS Aragón), 50009 Zaragoza, Spain
- * Correspondence: subillos@ubu.es; Tel.: +34-947-258-074



Citation: Sáiz-Vázquez, O.; Gracia-García, P.; Ubillos-Landa, S.; Puente-Martínez, A.; Casado-Yusta, S.; Olaya, B.; Santabárbara, J. Depression as a Risk Factor for Alzheimer's Disease: A Systematic Review of Longitudinal Meta-Analyses. *J. Clin. Med.* **2021**, *10*, 1809. <https://doi.org/10.3390/jcm10091809>

Academic Editor: Marco Bozzali

Received: 10 March 2021

Accepted: 15 April 2021

Published: 21 April 2021

Publisher's Note: MDPI stays neutral with regard to jurisdictional claims in published maps and institutional affiliations.



Copyright: © 2021 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (<https://creativecommons.org/licenses/by/4.0/>).

Abstract: Alzheimer's disease (AD) is the most frequent cause of dementia, linked to morbidity and mortality among elderly patients. Recently, several clinical studies suggested that depression is a potential risk factor for cognitive decline and AD. A review of meta-analyses was performed, calculating pooled odds ratios to estimate the risk of AD in people with a prior diagnosis (or clinically significant symptoms) of depression. A total of six meta-analyses which represented 28 individual studies were analyzed. A significant association between depression and AD was found (OR = 1.54, 95% CI [1.02–2.31]; $p = 0.038$). The results showed that heterogeneity across studies was substantial. We found a significant positive effect size for clinical measures of depression, but not for symptomatic rating scales, in the association of depression with risk of AD. The type of rating scale used to assess depression and the cut-off criteria selected also moderated the relationship between depression and AD risk. We found that studies that used clinically significant criteria for diagnosis of depression had more consistent and significant results than studies that used symptomatic scales.

Keywords: depression; Alzheimer's disease; clinical and symptomatic criteria; meta-meta-analysis

1. Introduction

Alzheimer's disease (AD) is the most frequent cause of dementia and is considered one of the main causes of morbidity and mortality among elderly people [1]. The World Alzheimer's Report revealed that 46.8 million people worldwide were living with dementia in 2015, and the total global social cost of dementia was estimated to be \$818 billion [2]. Estimates of dementia incidence in population-based studies range from 5 to 10 cases per 1000 person-years in people aged 64 to 69, and up to 40 to 60 cases per 1000 person-years in people aged 80 to 84 [3]. In 2017 in Europe, prevalence rates of AD were reported to be 5.05%, with 3.31% in men and 7.13% in women [4]. Given the personal and social consequences of dementia and AD demand, we accelerate the global effort to understand this complex disorder [5].

Decades of research revealed that the pathophysiological mechanisms underlying this neurodegenerative disease include accumulation of amyloid-beta peptide (A β) in brain tissues and cytoskeletal changes related to the hyperphosphorylation of microtubule-associated Tau protein in neurons. As a consequence, neuritic plaques and neurofibrillary tangles are accumulated, mostly in the medial temporal lobe and associative neocortical areas [6], and resulting in several cognitive deficits. The clinical manifestation of AD is progressive, from unnoticeable brain changes to brain changes that cause cognitive deterioration and eventually physical disability [7]. AD usually begins with memory difficulties followed by other cognitive problems such as visuospatial abnormalities, navigation difficulties, executive problems, and language disturbance [2].

Evidence seems to suggest that the etiology of AD is multifactorial, with genetics, older age, and a family history of AD being the greatest contributors to a higher risk of AD [7]. Furthermore, AD is often associated with other chronic diseases (diabetes, cholesterol, cardiovascular diseases, obesity, and hypertension) [8]. Although these risk factors are unchangeable, other risk factors can be modified to reduce the risk of dementia and cognitive decline. This is particularly important, since there is no currently available way to stop the damage and destruction of neurons linked to AD.

Depressive symptoms are common in AD and occur in approximately 20–30% of patients [9]. Depression is a serious medical illness that affects about 300 million people worldwide and which might aggravate existing medical conditions and increase functional disability [9,10]. Clinical evidence suggests a relationship between depression and AD [11–14]. However, it remains unclear whether depression represents a risk factor for AD, is an early symptom of neurodegeneration, or is a reaction to early cognitive deficits [14,15]. Some studies have suggested that depressive symptoms immediately follow the onset of AD rather than precede it [16]. Moreover, evidence from other studies indicates that depression has only a mild effect on dementia [17] and does not increase the risk for developing AD [18]. However, other authors suggest that the presence of depression in patients with AD increases the risk of behavioral disturbance and accelerates functional decline [12]. Hudon et al. [19], for example, found that depression was the most consistent risk factor associated with behavioral or psychological symptoms and cognitive decline in patients with AD. In addition, several studies concluded that late-life depression is related to an increased risk for all-cause dementia, vascular dementia, and Alzheimer's disease [20–22], and late-life depression was shown to be associated consistently with a two-fold increased risk of dementia [23,24].

In order to clarify the role of depression as a risk factor of AD, several meta-analyses were conducted [19,20,22,23]. However, some limitations were pointed out. Cherbuin et al. [24], for example, indicated that, in general, results from previous studies that focused on depression as a risk factor of AD might be biased due to the type of instrument used to assess depression. Results are frequently based on different tools. Some of these studies are based on symptomatic rating scales with cut-off points (e.g., CESD), while others are based on clinical criteria (e.g., DSM). Thus, the pooled estimates of the risk for AD in depressed people might be unreliable, because these meta-analyses combined effect sizes from studies using different instruments to assess depression (i.e., symptomatic rating scales and clinical diagnoses). Additionally, these previous meta-analyses did not pool findings separately for studies using clinical criteria and studies using depressive symptom rating scales with specified cut-off points.

Based on these limitations and the inconclusive evidence, we aimed to perform a meta-meta-analysis of longitudinal studies to assess the effect of depression on the risk of a subsequent diagnosis of AD. Given the expected heterogeneity among studies, we also aimed to pool findings separately from studies using clinical criteria and those using depression symptom rating scales, and to test the association between depression and risk of AD according to the different instruments used.

2. Materials and Methods

2.1. Data Collection

This meta-meta-analysis was performed in accordance with the Preferred Reporting for Systematic Reviews and Meta-analysis (PRISMA) Statement [25]. For data collection, we searched meta-analyses that measured depression at baseline and reported outcomes in individuals with diagnoses of AD at follow-up. ISI Web of Science (WOS), Scopus, Pubmed, Elsevier Science Direct, and Google scholar were searched from inception up to 31 July 2020. Combinations of the following search terms were used: “depression” AND “Alzheimer’s disease” AND “meta-analysis”. The data search was done in English (four studies) and Spanish (one study). When necessary, corresponding authors were contacted to provide full text details of the study outcome measures.

2.2. Inclusion Criteria

By consensus of the authors, studies were included if they met the following criteria:

1. Longitudinal studies that investigated the effect of depression or depressive symptoms (at baseline) as an antecedent to AD (follow-up).
2. Studies including patients with a diagnosis of AD according to diagnosis criteria (e.g., Related Disorders Association criteria, N-ADRDA, the Diagnostic and Statistical Manual of mental Disorders, DSM-III or the National Institute of Neurological and Communication Disorders-Alzheimer’s Disease).
3. Studies that clinically assessed levels of depression by means of a clinical diagnosis (e.g., DSM-IV, ICD-10), or a symptomatic diagnostic tool with a cut-off score (e.g., Geriatric Mental State Schedule, GMS) that identifies clinically significant levels of depression.
4. Studies reporting sufficient information to calculate common effect size statistics (i.e., mean and SD, exact P-, t-, or z-values).
5. Original, peer-reviewed meta-analyses that were published in English and Spanish.

2.3. Exclusion Criteria

By consensus of the authors, the following were excluded:

1. Studies investigating the association of depression and risk of AD using a sample of patients with AD and other dementia (non-independent or overlapping data for AD).
2. Studies not reporting quantitative data to calculate the association between depression and AD, or not published as meta-analyses in peer-reviewed journals (i.e., conference abstracts, book chapters).
3. Meta-analyses about other topics or those that included the same primary studies.

2.4. Data Extraction and Quality Assessment

Titles and abstracts of potential meta-analyses about depression and incident AD were independently analyzed by three researchers (OS, SU, PG). After exclusion of irrelevant articles, the remaining meta-analyses were critically inspected to check data accuracy. Then, full texts of all primary studies included in each meta-analysis were screened according to the inclusion criteria. In the event of ambiguity, two authors (SU, JS) additionally reviewed the study to reach a consensus regarding its eligibility.

Data related to the diagnosis/assessment of depression and AD were collected directly from the text or from statistical tables. The lead author and either the third or fourth author independently extracted data from each study, including study characteristics (year, country, total sample size, and length of follow-up period), sample characteristics (mean age, % of women), measures of depression and AD, and the cut-off point used for depression in each individual study.

Diagnoses of AD were based on the following accepted clinical criteria: Revised criteria and the National Institute of Neurological and Communication Disorders-Alzheimer’s Disease and Related Disorders Association criteria (N-ADRDA), the Diagnostic and Sta-

tistical Manual of Mental Disorders in different editions (DSM-III, DSM-III-R, DSM-IV, DSM-V), and the International Classification of diseases (ICD-10). Additionally, studies established different cut-off scores on neuropsychological tests for the purposes of screening out cognitive impairment and dementia at baseline (see Table 1). Participants with scores above the cut-off on cognitive domains were excluded on the basis that this level of test performance indicates the presence of dementia or cognitive impairment. The most frequently used measures to describe the cognitive characterization of the participants at baseline were the Mini Mental State Examination (MMSE) ($n = 14$) and the Clinical Rating Scale (CRS) ($n = 6$). Diagnoses of depression were based on either symptomatic rating scales or clinical diagnoses. Clinical criteria for depression included the DSM-III, DSM-III-R, DSM-IV, DSM-V, and the Automated Geriatric Examination for Computer Assisted Taxonomy (GMS-AGECAT). Diagnoses of depression were based on symptomatic rating scales on valid cut-off points (SGDS/15/30, CES-D/10/11/20, HRSD-17).

In addition, the quality of the included studies was reported using the Assessment of Multiple Systematic Reviews (AMSTAR) tool [26], which was previously shown to have good inter-rater agreement, reliability, and content validity [26,27].

2.5. Statistical Analysis

Crude odds ratios (ORs) (and 95% confidence intervals (CIs)) were used to calculate the risk of developing AD associated with previous depression. When the number of cases of depression and AD were not provided, the effect sizes were calculated using reported data in the meta-analysis according to Lipsey and Wilson [28]. We considered HRs and ORs as equivalent, since it was previously shown that for rare events, they can be considered equivalent (incidence < 15%) [29]. Seventeen studies provided data that could be used in calculating crude ORs (odds of an outcome in the intervention arm divided by the odds of an outcome in the control). Eleven additional studies provided data on AD risk in samples as HR or ORs with 95% confidence intervals that could be used in pooling estimates.

Summary statistics were calculated using Comprehensive Meta-Analysis software (CMA; Version 3) (Biostat Inc., Englewood, NJ, USA) [30,31]. Initially, we performed an analysis summarizing all data available, including all studies with validated cut-offs or clinical diagnoses in a single pooled estimate [31]. For each study, we calculated: (a) 95% CI of the effect, (b) Z value and p (two-tailed significance), and (c) k or number of studies [32]. Presence of publication bias was assessed through visual inspection of funnel plots and with Egger's test [16].

The level of heterogeneity was assessed with the I² statistic, which describes the percentage of total variation across studies due to heterogeneity rather than chance alone. An I² value of 25% indicates low heterogeneity, 50% moderate heterogeneity, and 75% high heterogeneity [31]. Random-effect models were used to determine statistically significant heterogeneity. Additionally, the Cochran Q test was applied to assess significant heterogeneity (p -value < 0.05). Moderating variables were selected on the basis of substantive considerations and the availability of data across studies included in the meta-analysis. Subgroup analyses were performed according to how depression was assessed: by clinical diagnosis (e.g., DSM-V) or by symptomatic rating scales (e.g., CES-D). Additionally, because the studies included different symptomatic rating scales, we also considered the instrument and the specific cut-off criteria as moderating variables. Therefore, we calculated the effect sizes of the association between depression and risk of AD separately for studies using different cut-off points. Finally, meta-regression analyses were conducted to obtain the proportion of variance explained for each moderator (the R-square analog). The scatter plot represents the mean effect for each level of covariate.

Table 1. Summary of demographic and study information.

| Study | Year | Country | AD Measure ¹ | Cognitive Measure ² | Cut-Off Criteria Cognition | Depression Measure ³ | Cut-Off Criteria Depression ⁴ | n ⁵ | Follow-Up Length (Years) M (SD) | Age M (SD) | Female (%) (Total) | AMSTAR2 ⁶ |
|----------------------------------|------|--------------|-------------------------|--------------------------------|----------------------------|---------------------------------|--|----------------|---------------------------------|-------------|--------------------|----------------------|
| Bae et al. [33] | 2015 | AS | N-ADRDA | CERAD-K | ≥60 | GDS15 | ≥8 | 540 | 3.5 (0.3) | 71.7 (5.1) | 55.2 | HIGH |
| Bartolini et al. [34] | 2005 | EU | N-ADRDA | MMSE | >26 | DSM-III-R | - | 222 | 1 | 69.2 (4.8) | 63.5 | HIGH |
| Becker et al. [18] | 2009 | USA | N-ADRDA | MMSE | >26 | CES-D20 | ≥10 | 729 | 7.1 (NR) | 70 | 69 | HIGH |
| Blasko et al. [35] | 2010 | Austria | N-ADRDA | CERAD | ≥60 | DSM-IV | - | 648 | 2.5 (NR) | 78.3 (0.5) | 56.5 | HIGH |
| Burke et al. [36] | 2018 | USA | N-ADRDA | CRS | ≤3 | DSM-V | - | 12,083 | 4.2 (-) | 63.9 | 83 | HIGH |
| Chen et al. [37] | 1999 | USA | DSM-III-R | MMSE | >26 | CES-D20 | ≥16 | 803 | 4.5 (NR) | 73.7 (5.0) | 60 | MODERATE |
| Dal Forno et al. [38] | 2005 | USA | N-ADRDA | BIMC | | CES-D20 | ≥20 | 1357 | 6.1 (-) | 65.5 (12.0) | 45.5 | HIGH |
| Devanand et al. [39] | 1996 | USA | N-ADRDA | CRS | ≤3 | DSM-III R | - | 456 | 2.54 | 72 | 70 | HIGH |
| Dotson, Beydoun & Zonderman [40] | 2010 | USA | DSM-III R | BIMC | | CES-D20 | ≥16 | 2177 | 23.6 (NR) | 52.7 (18.8) | 42.3 | HIGH |
| Fuhrer, Dufouil & Dartigues [41] | 2003 | France | N-ADRDA/DSM-III-R | MMSE | >26 | CES-D20 | ≥16 | 1576 | 8.0 (NR) | 75.2 (6.9) | 58.3 | HIGH |
| Gatz et al. [16] | 2005 | Canada | DSM-III R | MMSE | >26 | CES-D20 | ≥16 | 766 | 5 | 74.5 (6.0) | 61.7 | HIGH |
| Geerlings et al. [42] | 2000 | Países Bajos | DSM-III-R | MMSE | >26 | GMS-AGECAT | - | 1911 | 5.9 (1.6) | 73.5 (7.9) | 49 | MODERATE |
| Geerlings et al. [43] | 2008 | Netherlands | N-ADRDA | MMSE | >26 | CES-D20 | ≥16 | 393 | 5.9 (1.6) | 73.5 (7.9) | 49 | MODERATE |
| Gracia-García et al. [44] | 2015 | EU | DSM-IV | MMSE | >26 | GMS-AGECAT | ≥3 | 3626 | 4.5 | 71.9 (9.0) | 54.4 | HIGH |
| Heser et al. [45] | 2013 | Germany | DSM-IV/ICD-10 | MMSE | >26 | DSM-IV | - | 2969 | 4 | 81 | 64.8 | HIGH |
| Irie et al. [46] | 2008 | USA | N-ADRDA | CRS | ≤3 | CES-D11 | ≥9 | 1585 | 5.1 | 76.3 (3.6) | 0 | HIGH |
| Kim et al. [47] | 2010 | South Korea | N-ADRDA | CRS | ≤3 | GDS30 | 13/14 | 473 | 2.4 (0.3) | 71.8 (5.1) | 54.4 | HIGH |
| Kim et al. [48] | 2011 | South Korea | DSM-IV | CRS | ≤3 | GMS-AGECAT | ≥3 | 563 | 2.4 (0.3) | 71.8 (5.0) | 54.4 | MODERATE |
| Lauriola et al. [49] | 2018 | EU | DSM-V | MMSE | >26 | DSM-V | - | 181 | 4 | 74.5 (7.5) | 59.7 | HIGH |
| Lenoir et al. [50] | 2011 | France | N-ADRDA | MMSE | >26 | CES-D20 | M ≥ 16 W ≥ 22 | 7989 | 4 (NR) | 74.0 (5.4) | 61.3 | HIGH |
| Li et al. [51] | 2011 | USA | N-ADRDA | CASI | ≥78 | CES-D11 | ≥10/ | 3410 | 7.1 (NR) | 74.9 (6.2) | 59.9 | HIGH |
| Luchsinger et al. [52] | 2008 | USA | N-ADRDA | CRS | ≤3 | HRS17 | ≥10 | 1138 | 5.1 (3.3) | 75.1 (6.4) | 67.7 | HIGH |
| Reding, Haycox & Blass [53] | 1985 | USA | ICD-10 | MSQ | 0–2 errors | DSM-III | - | 60 | 3 | - | - | MODERATE |
| Richard et al. [54] | 2013 | USA | DSM-III R | MMSE | >26 | CES-D10 | ≥4 | 2160 | - | 76.9 (7.1) | 75 | MODERATE |
| Saczynski et al. [55] | 2010 | USA | N-ADRDA | MMSE | >26 | CES-D20 | ≥16 | 949 | 8 (NR) | 79.3 (5.0) | 63.6 | MODERATE |

Table 1. Cont.

| Study | Year | Country | AD Measure ¹ | Cognitive Measure ² | Cut-Off Criteria Cognition | Depression Measure ³ | Cut-Off Criteria Depression ⁴ | n ⁵ | Follow-Up Length (Years) M (SD) | Age M (SD) | Female (%) (Total) | AMSTAR2 ⁶ |
|----------------------------|------|---------|-------------------------|--------------------------------|----------------------------|---------------------------------|--|----------------|---------------------------------|------------|--------------------|----------------------|
| Tyas et al. [56] | 2001 | Canada | N-ADRDA | MMSE | >26 | CES-D20 | ≥16 | 694 | 3 to 5 | 65 | 67 | MODERATE |
| Vilalta-Franch et al. [57] | 2013 | EU | DSM-IV | CAMCOG | ≥79 | DSM-IV | - | 451 | 5 | 76.7 (5.4) | 63.7 | HIGH |
| Wilson et al. [58] | 2003 | USA | N-ADRDA | VARIOUS | - | CES-D10 | ≥4 | 142 | 3.9 (NR) | 81.0 (6.6) | 52.3 | HIGH |

Note: Meta-analyses analyzed were: Cherbuin et al. [24], Diniz et al. [22], Gao et al. [23], Kuring et al. [20], Kuring et al. [59], Santabárbara et al. [21]. ¹ AD: Alzheimer’s disease. DSM-III-R, DSM-IV, DSM-V = Diagnostic and Statistical Manual of Mental Disorders; N-ADRDA = National Institute of Neurological and Communicative Disorders and Stroke–Alzheimer’s Disease and Related Disorders Association; N-AIREN = National Institute of Neurological and Communicative Disorders and Stroke–Association Internationale pour la Recherche et l’ Enseignement en Neurosciences; ICD-10 = International Classification of Diseases. Total of diagnoses are $k = 30$. ² Cognitive measures: CERAD/K: Consortium to Establish a Registry for Alzheimer’s Disease; MMSE: Mini-Mental State Examination; CRS: Clinical Rating Scale; BIMC: Blessed Information-Memory-Concentration; CASI: Cognitive Abilities Screening Instrument; MSQ: Mental Status Questionnaire; CAMCOG: Cambridge Cognitive Examination. ³ Depression. DSM-III, DSM-III-R, DSM-IV, DSM-V: Diagnostic and Statistical Manual of Mental Disorders; HRSD17: Hamilton M. Rating Scale for DP; GMS-AGECAT: Geriatric Mental State-Automated Geriatric Examination for Computer Assisted Taxonomy; GDS-15/30: Geriatric Depression Scale; CES/-D10 (10 items)/- D11 (11 items)/-D20 (20 items) = Center for Epidemiologic Studies–DP Scale. ⁴ Cut-off criteria for categorial depression measures: HRSD-17, Hamilton et al. [60]; Williams et al. [61]; GMS-AGECAT, Copeland et al. [62]; GDS 15/30, Jung et al. [63]; Yesavage et al. [64]; SGDS, Kim et al. [65]; CES-D/D20, Radloff [66]; CES-D10/11, Kohout et al. [67]. ⁵ Follow-up: Total sample size for controls and healthy indicated; separate sample sizes for those with AD and depression and healthy controls were not reported. Study based on registry data. ⁶ AMSTAR 2 identifies quality of randomized controlled clinical trials. Rating overall confidence in the results: High = Zero or one non-critical weakness; Moderate = More than one non-critical weakness; Low = One critical flaw with or without non-critical weaknesses; Critically low = More than one critical flaw with or without non-critical weaknesses, Shea et al. [27] (https://amstar.ca/Amstar_Checklist.php accessed on 19 April 2021).

3. Results

The search strategy produced a total of 443 meta-analyses (see Table 1). Initially, 37 meta-analyses were eligible for inclusion. Of these, 31 were excluded: (a) 3 did not report an effect size; (b) 6 did not provide information on the relationship between depression and AD; (c) 8 were duplicates; (d) 9 were systematic reviews about other topics; (e) 4 aimed to study the effect of medication on AD; and (f) 1 included the same primary studies as another. Finally, a total of six meta-analyses were analyzed (k = 28 pooled effect sizes), representing data from $n = 28$ individual studies (see Figure 1).

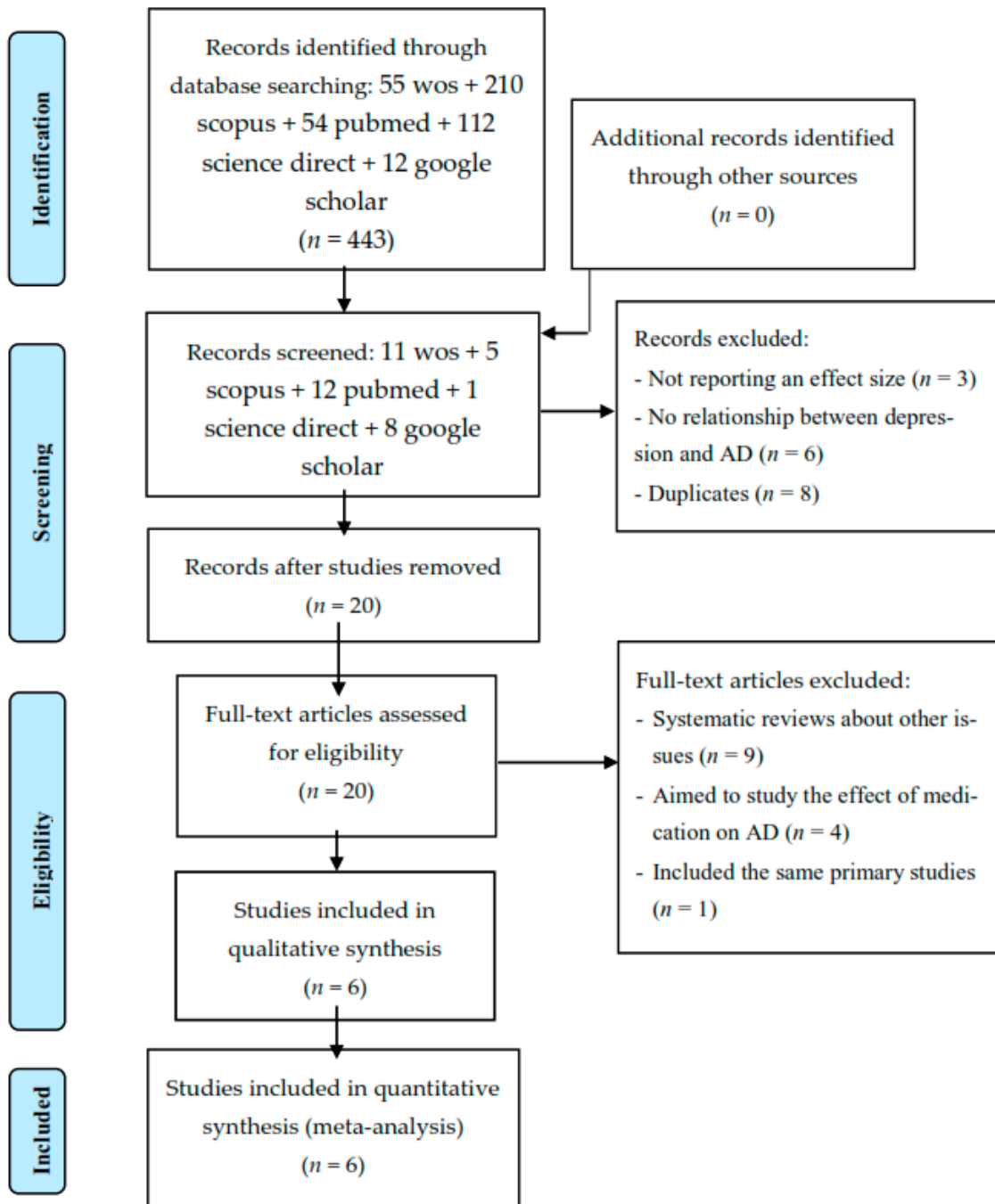


Figure 1. Flow chart depicting the selection of articles for our meta-analysis. Note: AD: alzheimer’s disease; n: number of studies.

Since the effect estimated from a biased collection of studies might overestimate the true effect, we assessed the likely extent of this bias and its potential impact on the

conclusions. The result of Egger’s test was not significant: The intercept (B0) was 0.53, 95% CI (−1.88 to 2.95), with $t = 0.45$, $df = 26$, $p = 0.65$, indicating no publication bias.

3.1. Overall Results from the Meta-Analysis

A total of 28 individual studies reported the association between depression at baseline and AD at follow-up with a total of 101,881 participants (Nbaseline = 51,830; Nfollow-up = 50,051). Individual sample sizes ranged from 60 to 12,083. The majority of subjects was female. The mean age was 71.95, ranging from 52.7 to 81 years. One study did not report gender and age [53]. The mean follow-up length was 4.90 years (range from 1 to 23.6), with one study not reporting the number of years [54]. Characteristics of the 28 individual studies are presented in Table 1.

A total of 17 and 11 studies were based on symptomatic rating scales and clinical criteria to assess depression, respectively: CES-D ($n = 14$) (50%), DSM-III/III-R/IV/V ($n = 8$) (28.6%), GMS-AGECAT ($n = 3$) (10.7%), GDS ($n = 2$) (7.1%), and HAM-D ($n = 1$) (3.6%). AD diagnosis was established based on the N-ADRDA ($n = 17$) (56.7%) or DSMIII-R/IV/V ($n = 10$) (33.3%), ICD10 ($n = 2$) (6.7%), and N-AIREN ($n = 1$) (3.3%) scales.

Risk estimates were pooled across the 28 studies. The random effect of the relationship between depression and AD was significant (OR = 2.46, 95% CI [1.81–3.35], $Z = 5.72$, $p < 0.001$). Figure 2 shows the forest plot of the effect sizes and their 95% CI. Heterogeneity across studies was substantial (Q-value = 284.53, $df = 27$, $I^2 = 90.51$, $p < 0.001$), suggesting the presence of potential moderators (Table 2).

Table 2. Summary details for individual studies that examined the risk of dementia (OR) associated with depression.

| Study Name | Statistics for Each Study | | | | | Exposed (AD)/Total | Exposed (AD)/Total |
|----------------------------|---------------------------|-------------|-------------|---------|---------|--------------------------|--------------------------|
| | Odds Ratio | Lower Limit | Upper Limit | Z-Value | p-Value | Cases (Depression) | Controls (No Depression) |
| Bae et al. [33] | 9.84 | 0.57 | 170.00 | 1.57 | 0.116 | 9/359 | 0/181 |
| Bartolini et al. [34] | 16.00 | 3.72 | 68.76 | 3.73 | <0.001 | 31/124 | 2/98 |
| Becker et al. [18] | 1.33 | 0.49 | 3.65 | 0.56 | 0.578 | HR = 1.33 (0.49–3.65) | |
| Blasko et al. [35] | 1.09 | 0.77 | 1.53 | 0.47 | 0.637 | 77/242 | 122/406 |
| Burke et al. [36] | 4.15 | 3.49 | 4.94 | 15.98 | <0.001 | 205/1214 | 507/10,869 |
| Chen et al. [37] | 3.37 | 1.33 | 8.54 | 2.56 | 0.011 | 6/52 | 28/751 |
| Dal Forno et al. [38] | 2.63 | 1.28 | 5.40 | 2.63 | 0.008 | HR = 2.63 (1.28–5.40) | |
| Devanand et al. [39] | 5.07 | 3.02 | 8.52 | 6.13 | <0.001 | 57/173 | 25/283 |
| Dotson et al. [40] | 1.02 | 0.77 | 1.35 | 0.11 | 0.911 | 96/938 | 125/1239 |
| Fuhrer et al. [41] | 1.04 | 0.69 | 1.58 | 0.19 | 0.849 | 30/203 | 196/1373 |
| Gatz et al. [16] | 3.49 | 1.08 | 11.28 | 2.09 | 0.037 | OR = 3.49 (1.08–11.28) | |
| Geerlings et al. [42] | 2.21 | 1.09 | 4.48 | 2.20 | 0.028 | OR = 2.21 (1.09–4.48) | |
| Geerlings et al. [43] | 1.41 | 0.55 | 3.58 | 0.71 | 0.475 | 6/35 | 44/343 |
| Gracia-García et al. [44] | 1.81 | 0.98 | 3.36 | 1.89 | 0.059 | 13/452 | 51/3174 |
| Heser et al. [45] | 2.70 | 1.80 | 4.03 | 4.84 | <0.001 | 34/306 | 118/2663 |
| Irie et al. [46] | 9.94 | 3.16 | 31.22 | 3.93 | <0.001 | 6/146 | 6/1397 |
| Kim et al. [47] | 1.33 | 0.62 | 2.85 | 0.74 | 0.463 | HR = 1.33(0.62–2.85) | |
| Kim et al. [48] | 2.33 | 0.97 | 5.56 | 1.90 | 0.057 | 7/45 | 38/518 |
| Lauriola et al. [49] | 130.73 | 7.90 | 2162.50 | 3.40 | 0.001 | 57/115 | 0/66 |
| Lenoir et al. [50] | 1.01 | 0.69 | 1.49 | 0.05 | 0.960 | HR = 1.0 (0.7–1.6) | |
| Li et al. [51] | 1.43 | 1.05 | 1.94 | 2.28 | 0.022 | HR = 1.43 (1.05–1.94) | |
| Luchsinger et al. [52] | 3.40 | 1.46 | 7.90 | 2.85 | 0.004 | HR = 3.4 (1.5–8.1) | |
| Reding et al. [53] | 19.00 | 12.42 | 29.06 | 13.59 | <0.001 | HR = 19.00 (12.40–27.90) | |
| Richard et al. [54] | 2.03 | 1.44 | 2.86 | 4.06 | <0.001 | 55/452 | 109/1708 |
| Saczynski et al. [55] | 1.76 | 1.03 | 3.01 | 2.07 | 0.039 | HR = 1.76 (1.03–3.01) | |
| Tyas et al. [56] | 2.00 | 1.01 | 3.95 | 2.00 | 0.046 | 21/36 | 271/658 |
| Vilalta-Franch et al. [57] | 2.36 | 1.11 | 5.03 | 2.23 | 0.026 | 13/116 | 17/335 |
| Wilson et al. [58] | 1.33 | 1.01 | 1.76 | 2.01 | 0.044 | OR = 1.33 (1.01–1.76) | |
| Random effects | 2.46 | 1.81 | 3.35 | 5.72 | <0.001 | | |

Note: AD: Alzheimer’s disease; NO-AD: No Alzheimer’s disease. Ns are based on total participant data available for depression or AD (not entire sample). Some data (N at baseline and follow-up) were not available for the depression and control groups, because studies did not provide them. In those cases, we reported the effect given in primary studies.

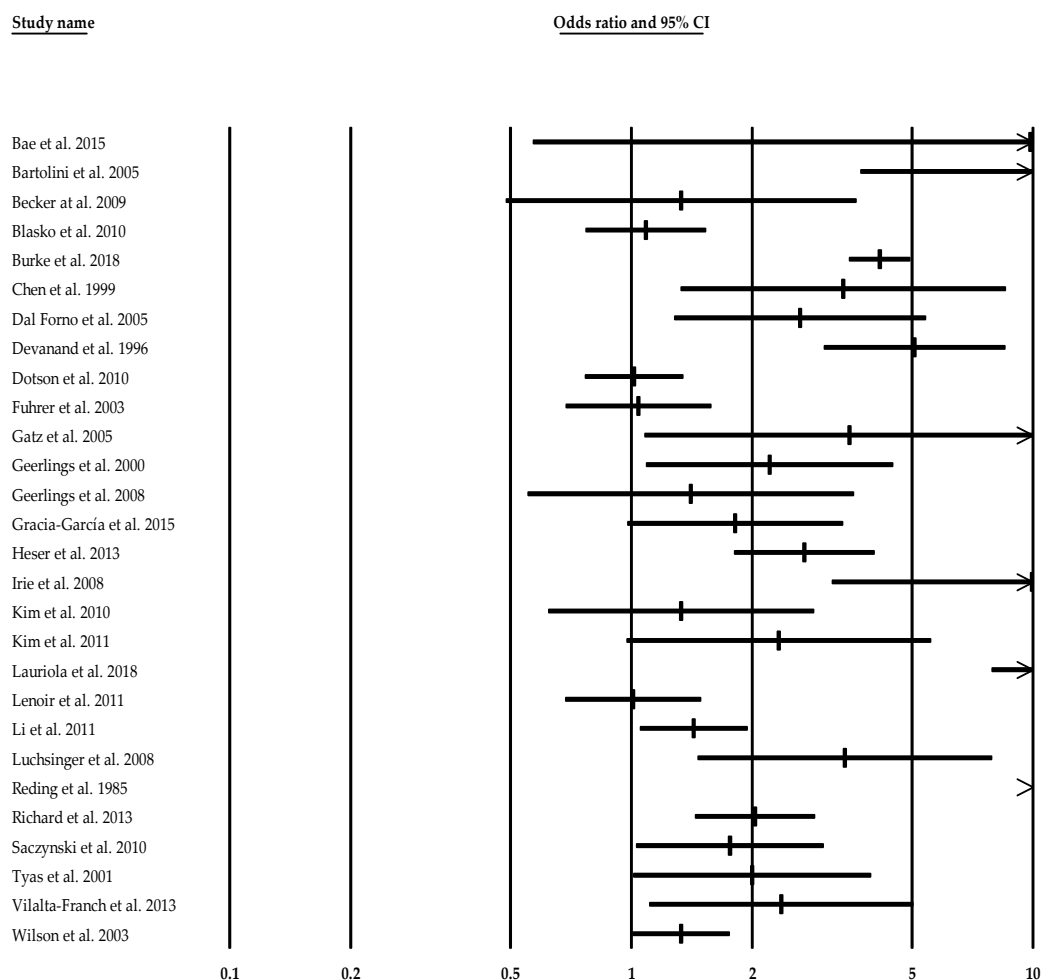


Figure 2. Forest plot of studies investigating the risk of Alzheimer’s disease (Time2) associated with depression (including all instruments).

3.2. Clinical Criteria and Symptomatic Rating Scales to Assess Depression

We tested three different models that reflected a combination of methodological moderators (see Table 3). Random effect models revealed a significant positive effect size of the association between depression and risk of AD for clinical ($k = 11$) and symptomatic ($k = 17$) measures of depression. Heterogeneity was substantial for the depression criteria ($I^2 = 90.51$), indicating that the OR was greater for clinical than symptomatic measures.

Then, we performed an additional sub-group analysis distinguishing between types of symptomatic rating scale used to assess depression. The total effect (OR) was significant (1.80, 95% CI: 1.16–2.78, $Z = 2.62$, $p = 0.009$), and heterogeneity was moderate ($I^2 = 61.84$). Sub-group analysis yielded a significant effect of depression on the development of AD for studies using the CES-D scales and HSRD, although this effect was non-significant when studies used the GDS scale. Only one study included the HSRD scale, and no additional subsample analyses were conducted. However, sufficient data were available for the CES-D ($k = 14$). We conducted further sub-analyses according to different cut-off points of the CES-D scale to define presence of depression. ORs were pooled across 14 studies (OR = 1.68, IC95% 1.24–2.27, $Z = 3.36$, $p = 0.001$). Heterogeneity was moderate across these studies ($I^2 = 63.95$), indicating that the effect of depression on the risk of AD may differ according to the cut-off points used. Estimates were significant for ≥ 10 and ≥ 16 cut-offs, whereas the effect of depression on AD was not significant when studies used a cut-off of ≥ 4 and ≥ 20 (Table 3).

Table 3. Summary effect sizes.

| | | Model Statistics | | | | | | |
|-------------------------------|----------|------------------|-----------|-----------|----------|----------|----------------------|----------------------|
| | <i>k</i> | <i>OR</i> | <i>LL</i> | <i>UL</i> | <i>Z</i> | <i>p</i> | <i>Q_w</i> | <i>Q_b</i> |
| Depression criteria (model 1) | | | | | | | | |
| Clinic | 11 | 3.68 | 2.44 | 5.55 | 6.20 | 0.0001 | 172.78 *** | 6.86 ** |
| Symptomatic | 17 | 1.81 | 1.30 | 2.53 | 3.51 | 0.0001 | | |
| Depression scale (model 2) | | | | | | | | |
| GDS | 2 | 1.63 | 0.64 | 4.15 | 1.03 | 0.303 | 37.83 *** | 1.87 |
| CES-D | 14 | 1.60 | 1.28 | 2.02 | 4.07 | 0.0001 | | |
| HSRD | 1 | 3.40 | 1.19 | 9.71 | 2.29 | 0.022 | | |
| Cut-off (CES-D) (model 3) | | | | | | | | |
| ≥4 | 2 | 1.63 | 0.97 | 2.78 | 1.80 | 0.072 | 28.63 ** | 1.97 |
| ≥10 | 3 | 2.02 | 1.14 | 3.60 | 2.39 | 0.017 | | |
| ≥16 | 8 | 1.44 | 1.04 | 2.00 | 2.19 | 0.028 | | |
| ≥20 | 1 | 2.63 | 0.97 | 7.11 | 1.91 | 0.057 | | |

Note: *** $p \leq 0.001$, ** $p \leq 0.01$, *k*: number of studies; *OR*: Odds ratio; *LL*: Lower limit; *UL*: Upper limit; *Q_w*: heterogeneity within; *Q_b*: heterogeneity between.

3.3. Meta Regression Analysis

We conducted a meta-regression analysis to determine whether the criteria used to measure depression might explain differences across studies in reporting effect size and might also cause heterogeneity. A significant negative effect of the use of symptomatic rating scales on the prediction of AD was found ($b = -0.71$, $Se = 0.27$, 95% CI: $-1.24 / -0.17$, $Z = -2.59$, $p = 0.009$) compared to clinical criteria ($k = 28$, intercept: $b = 1.30$, $se = 0.21$, CI: $0.89 / 1.72$, $Z = 6.14$, $p \leq 0.001$) ($Q = 6.71$, $df = 1$, $p = 0.009$). Together, these explained 26% of the variance. That is, the use of symptomatic rating scales to assess depression was associated with a decreased likelihood of developing AD in the follow-up compared to the use of clinical criteria.

No significant moderating effects were found in meta-regression analyses conducted for the various symptomatic rating scales of depression ($k = 17$) (1 = GDS, intercept: $b = 0.47$, $Se = 0.45$ ($-0.41 / 1.36$), $Z = 1.04$, $p = 0.296$; 2. CES-D: $b = -0.02$, $Se = 0.46$ ($-0.93 / 0.89$), $Z = -0.04$, $p = 0.97$; 3. HSRD: $b = 0.75$, $Se = 0.68$ ($-0.59 / 2.09$), $Z = 1.10$, $p = 0.270$) ($Q = 2.18$, $df = 2$, $p = 0.336$). Differences explained the 28% of variation observed in the association between depression and AD.

When analyzing the differential effect of the CES-D cut-offs on the development of AD ($k = 14$), results showed a greater predictive effect for studies using more restrictive cut-off points (≥ 20) (intercept: $b = 0.97$, $SE = 0.37$, 95% CI: $0.25 / 1.69$, $Z = 2.63$, $p = 0.008$) (≥ 4 : $b = -0.51$, $Se = 0.38$ ($-1.26 / 0.24$), $Z = 1.34$, $p = 0.180$; ≥ 10 : $b = -0.50$, $SE = 0.39$ ($-1.27 / 0.28$), $Z = -1.26$, $p = 0.209$; ≥ 16 : $b = -0.77$, $SE = 0.38$ ($-1.51 / -0.01$), $Z = -2.04$, $p = 0.041$) ($Q = 7.43$, $df = 3$, $p = 0.050$). The different cut-off points of the CES-D explained the 53% of variation in the diagnosis of AD (Figure 3).

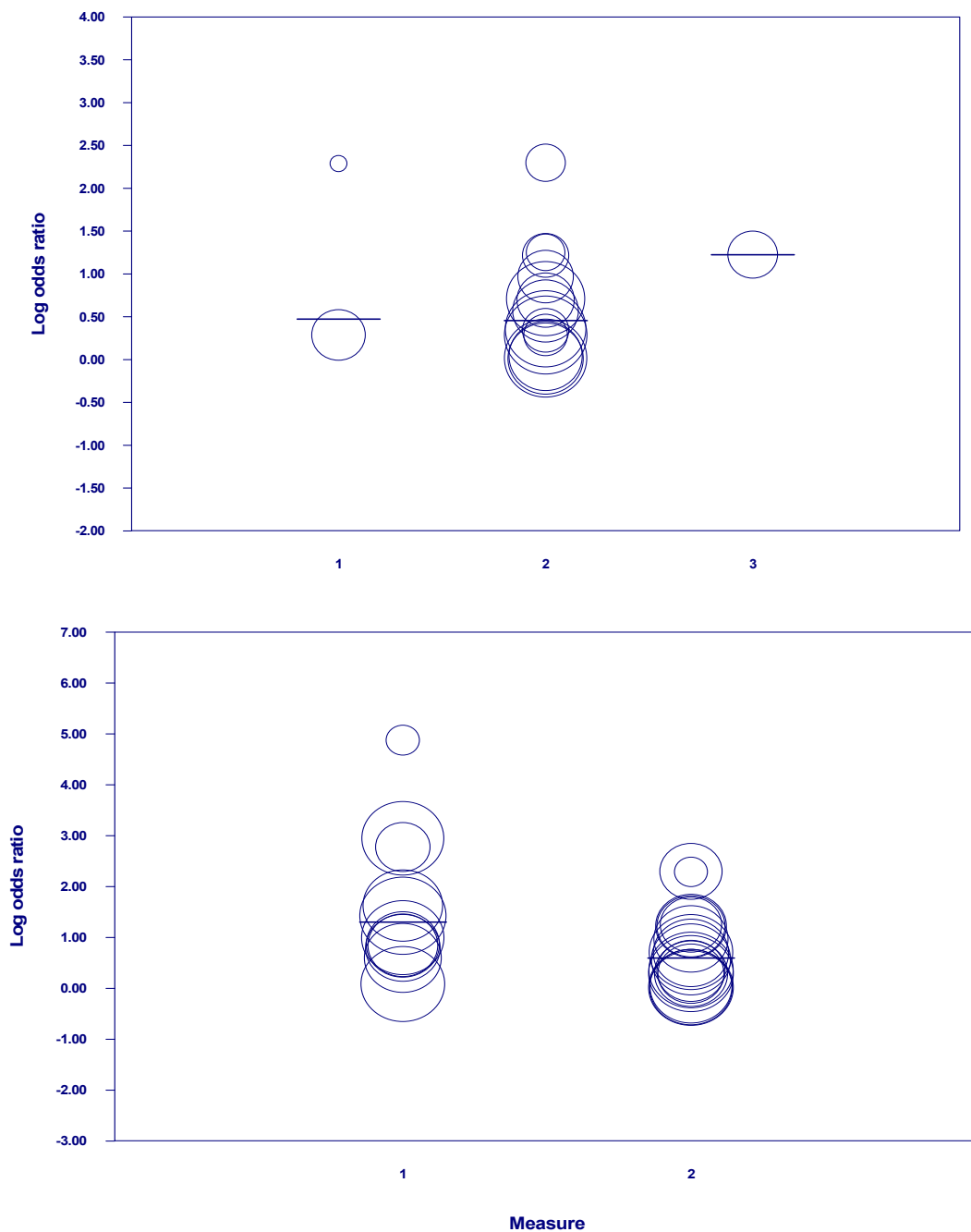


Figure 3. Meta-regression of log odds ratio on type of measure, symptomatic tool, and CES-D cut-offs (95% intervals are simultaneous and based on Z distribution). Scatterplots show the relationship between each covariate and AD.

4. Discussion

The main contribution of this study was to produce precise AD risk estimates associated with different depression criteria, either clinically significant or based on symptomatic scales. Based on the results of 11 cohorts, we found a more than three-fold increased risk of AD for clinically significant depression. Likewise, based on findings of 17 cohort studies, the risk of AD increased almost two-fold in participants diagnosed with symptomatic measures of depression. We found that studies that used clinically significant criteria for diagnosis of depression had more consistent and significant results than those that used symptomatic scales.

However, most included studies used self-reported symptomatic scales for diagnosis of depression, specifically the CES-D. We further analyzed the differential effect of CES-D

cut-off points on AD risk and found that they explained 53% of the variability of results. We found a slightly significant predictive effect in meta-analyzed data of studies using the cut-off CES-D point ≥ 10 and ≥ 16 , but predictive risk of AD was greater for one study using a more restrictive cut-off point (≥ 20). Our results are consistent with those of Cherbuin et al. [24] who found that the meta-analysis of studies using a cut-off previously validated against clinical criteria (≥ 20) demonstrated higher risk estimates than those using a more lenient cut-off (≥ 16).

We found a greater effect of clinically significant depression on AD risk than the MA of Santabábara et al. [21], probably because that meta-study included only three studies with homogeneous criteria for the diagnosis of depression (GMS-AGECAT). We also included eight studies using DSM criteria for depression; all of them but one, Blasko et al. [35], found consistently higher risk of AD compared to any other criteria. However, some of them found relatively large [34,53] or even extreme values of OR [49].

Furthermore, our study includes recent references [20], and it did not analyze data from studies of patient groups with mixed psychiatric histories or all types of dementia. Even though the meta-analysis of Kuring et al. [20] analyzed 36 independent studies for all types of dementia, they only pooled $k = 8$ studies for AD (OR = 2.23). This inclusion criterion may explain why our results show a greater OR risk from depression to AD than previous meta-analyses [20–24]. Furthermore, they did not analyze variability arising from the type of measure (clinical or symptomatic criteria) and from cut-off points used to assess depression across studies. Another strength of our study is that it includes a selection of prospective cohort studies to provide more evidence in establishing the cause and effect, and the relationship between depression and AD [21]. We analyzed a long follow-up period (4.9 years, range 1–23.6 years) to observe the potential association between depression (as an antecedent) and risk of AD, avoiding cross-sectional studies [20,24]. This analysis covers a gap in the previous literature, adding new information about the association between depression and AD. Finally, previous meta-analyses limited the literature search to biomedical databases. In our study, we included five databases in order to provide coverage of publications from different countries, reducing the likelihood of publication bias [22,23].

Overall, our study is the first to review all previously available meta-analyses of depression as a risk factor of incident AD systematically. Moreover, we included individual studies when they assessed clinically significant depression or a validated cut-off score in a symptomatic depression scale, and we conducted differential meta-analysis of specific AD risk estimates according to depression criteria. Our study demonstrates how depression criteria can explain variability between studies in the association between depression and incident AD. We agree with Cherbuin et al. [24] about the importance of using objective and specific measures of risk in evidence-based clinical practice.

A number of different hypotheses on the association between depression and dementia were suggested, yet the ways in which depression influences AD are as yet unclear. For instance, antidepressant use (i.e., anticholinergic drugs) was shown to be associated with an increased risk of dementia [68–71]. Furthermore, the $\epsilon 4$ allele of apolipoprotein E (APOE) was associated with the development of AD [46,47]. However, the idea that $\epsilon 4$ and dementia may be linked has little support [72–75]. In this vein, some risk factors, such as brain-vascular [76], cortisol, hippocampal atrophy [77], and neuroinflammation, could involve a possible common pathway to explain the association between depression and AD [78].

We should also recognize some limitations of our study. Firstly, as the studies included in the meta-analyses reported either the odds ratio or the hazard ratio for the association between late-life depression and dementia, we calculated the pooled OR for the association between depression and AD separately. Odds ratio is a measure of association between two conditions (such as in logistic regression models), whereas the hazard ratio is a measure of the strength of the association between two conditions in time-to-event statistical analysis. Given this, we should interpret the results from the pooled risk analysis with caution, as

we included studies that reported hazard ratios and odds ratios together. Nonetheless, the results are very consistent across all analyses for AD. Secondly, individual studies assessing depressive symptoms by self-rating scales used pre-established cut-off scores, and no structured interviews were conducted for the diagnosis of depressive disorders, which may have introduced significant heterogeneity into the classification of cases and non-cases, in particular in individuals with mild depression; according our results, this may explain a good deal of the variability in results between studies. In addition, some of the studies included in this meta-analysis were not representative of the entire population (such as studies including only men) [46]. Although we did not find a moderator effect of observation time (results not shown), and the results support the hypothesis that clinical depression is a risk factor for later development of Alzheimer's disease, the influence of prodromal symptoms should not be discounted, and it remains to be determined. Furthermore, we did not examine the influence of any single study on the overall risk estimates with sensitive analysis that omitted them one by one. Moreover, by choosing to include studies that allowed us to calculate crude ORs, we implicitly included studies that provided estimates of the relation between depression and AD risk in the form of unadjusted ORs, so other study-related factors may have affected the outcomes of these studies (age or sex). Inclusion of these studies may have biased our results. Another possible limitation of this meta-analysis is that our search was limited to certain databases. We did a careful review of all references in potentially relevant publications, previous meta-analyses, and systematic reviews published on depression and AD. Nevertheless, a search of other international databases (such as EMBASE and PsycINFO) might have led to the identification of additional studies that could have been included in this meta-analysis.

5. Conclusions

Although we cannot yet assert an etiological basis of the association, our study provides consistent data pointing to an increased risk of AD for clinically significant depression. Our findings highlight the importance of using more stringent and objective measures of depression in future studies. Depression should be assessed by clinicians with standardized, validated measures, and preventive strategies targeting at-risk individuals should be designed. Further studies need to assess the potential for treatment of clinically significant depression to decrease the risk of AD.

Author Contributions: All authors take full responsibility for the data, analysis, and interpretation, and the conducting of the research, and had full access to all of the data and the right to publish any and all data. The contributions of each author are as follows: O.S.-V. conceived and designed the analysis, collected the data, contributed data or analysis tools, performed analysis, and wrote the paper; P.G.-G. helped write the paper; S.U.-L. contributed data and analysis tools, performed analysis, and helped write the paper; A.P.-M. conceived and designed the analysis, collected the data, contributed data and analysis tools, performed analysis, and helped write the paper; S.C.-Y., B.O., and J.S. helped write the paper. All authors have read and agreed to the published version of the manuscript.

Funding: This study was supported by Grants from the Fondo de Investigación Sanitaria, Instituto de Salud Carlos III, Spanish Ministry of Economy and Competitiveness, Madrid, Spain (grants 94/1562, 97/1321E, 98/0103, 01/0255, 03/0815, 06/0617, 12/02254, 16/00896, PI/19/01874, G03/128) and from the Fondo Europeo de Desarrollo Regional (FEDER) of the European Union "Una manera de hacer Europa" (Project number PI16/00896) and Gobierno de Aragón (grant B15_17R). BO's work is supported by the Miguel Servet program (CP20/00040), funded by the Instituto de Salud Carlos III and co-funded by the European Union (ERDF/ESF, "Investing in your future"). This study was partially supported by FEDER funds and the Spanish Ministry of Economy and Competitiveness (Projects ECO2016-76567-C4-2-R and PID2019-104263RB-C44), the Regional Government of "Castilla y León", Spain (Project BU329U14 and BU071G19), the Regional Government of "Castilla y León", and FEDER funds (Project BU062U16).

Institutional Review Board Statement: Not applicable.

Informed Consent Statement: Not applicable.

Conflicts of Interest: We declare that P.G.-G. has received financial support to attend scientific meetings from Servier, Pfizer, Lundbeck, Nutrición Médica, and Esteve. None of these activities is related to the current project. For the remaining authors, no conflicts of interest are declared.

References

- Villain, N.; Dubois, B. Alzheimer's Disease Including Focal Presentations. *Semin. Neurol.* **2019**, *39*, 213–226. [[CrossRef](#)]
- Huang, L.-K.; Chao, S.-P.; Hu, C.-J. Clinical trials of new drugs for Alzheimer disease. *J. Biomed. Sci.* **2020**, *27*, 1–13. [[CrossRef](#)] [[PubMed](#)]
- De la Salud, A.M. *Proyecto de Plan de Acción Mundial Sobre La Respuesta de Salud Pública a La Demencia: Informe de La Directora General*; WHO: Geneva, Switzerland, 2017.
- Niu, H.; Álvarez-Álvarez, I.; Guillén-Grima, F.; Aguinaga-Ontoso, I. Prevalencia e incidencia de la enfermedad de Alzheimer en Europa: Metaanálisis. *Neurología* **2017**, *32*, 523–532. [[CrossRef](#)]
- Selkoe, D.J. Alzheimer disease and aducanumab: Adjusting our approach. *Nat. Rev. Neurol.* **2019**, *15*, 365–366. [[CrossRef](#)] [[PubMed](#)]
- Harris, J.R. *Protein Aggregation and Fibrillogenesis in Cerebral and Systemic Amyloid Disease*; Springer Science and Business Media: Berlin/Heidelberg, Germany, 2012; Volume 65, ISBN 94-007-5416-7.
- Migliaccio, R.; Agosta, F.; Rascovsky, K.; Karydas, A.; Bonasera, S.; Rabinovici, G.D.; Miller, B.L.; Gorno-Tempini, M.L. Clinical syndromes associated with posterior atrophy: Early age at onset AD spectrum. *Neurology* **2009**, *73*, 1571–1578. [[CrossRef](#)]
- Surguchov, A. Caveolin: A New Link Between Diabetes and AD. *Cell. Mol. Neurobiol.* **2020**, *40*, 1059–1066. [[CrossRef](#)] [[PubMed](#)]
- Tsuno, N.; Homma, A. What is the association between depression and Alzheimer's disease? *Expert Rev. Neurother.* **2009**, *9*, 1667–1676. [[CrossRef](#)] [[PubMed](#)]
- WHO. *Depression and Other Common Mental Disorders: Global Health Estimates*; World Health Organization: Geneva, Switzerland, 2017; pp. 1–24.
- Drevets, W.C.; Rubin, E.H. Psychotic symptoms and the longitudinal course of senile dementia of the Alzheimer type. *Biol. Psychiatry* **1989**, *25*, 39–48. [[CrossRef](#)]
- Lyketsos, C.G.; Tune, L.E.; Pearlson, G.; Steele, C. Major Depression in Alzheimer's Disease. *Psychosomatics* **1996**, *37*, 380–384. [[CrossRef](#)]
- De Oliveira, F.; Bertolucci, P.F.; Chen, E.; Smith, M. Assessment of risk factors for earlier onset of sporadic Alzheimer's disease dementia. *Neurol. India* **2014**, *62*, 625. [[CrossRef](#)]
- Cantón-Habas, V.; Rich-Ruiz, M.; Romero-Saldaña, M.; Carrera-González, M.D.P. Depression as a Risk Factor for Dementia and Alzheimer's Disease. *Biomedicines* **2020**, *8*, 457. [[CrossRef](#)]
- Kuo, C.-Y.; Stachiv, I.; Nikolai, T. Association of Late Life Depression, (Non-) Modifiable Risk and Protective Factors with Dementia and Alzheimer's Disease: Literature Review on Current Evidences, Preventive Interventions and Possible Future Trends in Prevention and Treatment of Dementia. *Int. J. Environ. Res. Public Health* **2020**, *17*, 7475. [[CrossRef](#)]
- Gatz, J.L.; Tyas, S.L.; John, P.S.; Montgomery, P. Do Depressive Symptoms Predict Alzheimer's Disease and Dementia? *J. Gerontol. Ser. A Biol. Sci. Med. Sci.* **2005**, *60*, 744–747. [[CrossRef](#)]
- Tapiainen, V.; Hartikainen, S.; Taipale, H.; Tiihonen, J.; Tolppanen, A.-M. Hospital-treated mental and behavioral disorders and risk of Alzheimer's disease: A nationwide nested case-control study. *Eur. Psychiatry* **2017**, *43*, 92–98. [[CrossRef](#)]
- Becker, J.T.; Chang, Y.-F.; Lopez, O.L.; Dew, M.A.; Sweet, R.A.; Barnes, D.; Yaffe, K.; Young, J.; Kuller, L.; Reynolds, C.F. Depressed Mood is Not a Risk Factor for Incident Dementia in a Community-Based Cohort. *Am. J. Geriatr. Psychiatry* **2009**, *17*, 653–663. [[CrossRef](#)]
- Hudon, C.; Escudier, F.; De Roy, J.; Croteau, J.; Cross, N.; Dang-Vu, T.T.; Zomahoun, H.T.V.; Grenier, S.; Gagnon, J.-F.; Parent, A.; et al. Behavioral and Psychological Symptoms that Predict Cognitive Decline or Impairment in Cognitively Normal Middle-Aged or Older Adults: A Meta-Analysis. *Neuropsychol. Rev.* **2020**, *30*, 558–579. [[CrossRef](#)]
- Kuring, J.; Mathias, J.; Ward, L. Risk of Dementia in persons who have previously experienced clinically-significant Depression, Anxiety, or PTSD: A Systematic Review and Meta-Analysis. *J. Affect. Disord.* **2020**, *274*, 247–261. [[CrossRef](#)]
- Serrano, J.S.; Pérez, A.S.; Olaya, B.; García, P.G.; Antón, R.L. Depresión tardía clínicamente relevante y riesgo de demencia: Revisión sistemática y metaanálisis de estudios prospectivos de cohortes. *Revista de Neurología* **2019**, *68*, 493–502. [[CrossRef](#)] [[PubMed](#)]
- Diniz, B.S.; Butters, M.A.; Albert, S.M.; Dew, M.A.; Reynolds, C.F. Late-life depression and risk of vascular dementia and Alzheimer's disease: Systematic review and meta-analysis of community-based cohort studies. *Br. J. Psychiatry* **2013**, *202*, 329–335. [[CrossRef](#)] [[PubMed](#)]
- Gao, Y.; Huang, C.; Zhao, K.; Ma, L.; Qiu, X.; Zhang, L.; Xiu, Y.; Chen, L.; Lu, W.; Huang, C.; et al. Retracted: Depression as a risk factor for dementia and mild cognitive impairment: A meta-analysis of longitudinal studies. *Int. J. Geriatr. Psychiatry* **2012**, *28*, 441–449. [[CrossRef](#)] [[PubMed](#)]
- Cherbuin, N.; Kim, S.; Anstey, K.J. Dementia risk estimates associated with measures of depression: A systematic review and meta-analysis. *BMJ Open* **2015**, *5*, e008853. [[CrossRef](#)] [[PubMed](#)]
- PRISMA Declaration: A Proposal to Improve the Publication of Systematic Reviews and Meta-Analyses. Available online: <https://europepmc.org/article/med/20206945> (accessed on 16 February 2020).

26. Shea, B.J.; Grimshaw, J.M.; Wells, A.G.; Boers, M.; Andersson, N.; Hamel, C.; Porter, A.C.; Tugwell, P.; Moher, D.; Bouter, L.M. Development of AMSTAR: A measurement tool to assess the methodological quality of systematic reviews. *BMC Med. Res. Methodol.* **2007**, *7*, 10. [\[CrossRef\]](#)
27. Shea, B.J.; Hamel, C.; Wells, G.A.; Bouter, L.M.; Kristjansson, E.; Grimshaw, J.; Henry, D.A.; Boers, M. AMSTAR is a reliable and valid measurement tool to assess the methodological quality of systematic reviews. *J. Clin. Epidemiol.* **2009**, *62*, 1013–1020. [\[CrossRef\]](#) [\[PubMed\]](#)
28. Lipsey, M.W.; Wilson, D.B. *Practical Meta-Analysis*; SAGE Publications, Inc.: Thousand Oaks, CA, USA, 2001; ISBN 0-7619-2167-2.
29. Shrier, I.; Steele, R. Understanding the Relationship Between Risks and Odds Ratios. *Clin. J. Sport Med.* **2006**, *16*, 107–110. [\[CrossRef\]](#) [\[PubMed\]](#)
30. Borenstein, M.; Hedges, L.; Higgins, J.P.T.; Rothstein, H. *Comprehensive Meta-Analysis*; Version 3; Biostat: Englewood, NJ, USA, 2013.
31. Borenstein, M.; Hedges, L.V.; Higgins, J.P.T.; Rothstein, H.R. *Introduction to Meta-Analysis*; John Wiley & Sons, Ltd.: Hoboken, NJ, USA, 2009; p. 452.
32. Reis, H.T.; Judd, C.M. *Handbook of Research Methods in Social and Personality Psychology*; Cambridge University Press: Cambridge, UK, 2000.
33. Bin Bae, J.; Kim, Y.J.; Ghang, J.H.; Kim, T.H.; Park, J.H.; Lee, S.B.; Lee, J.J.; Jeong, H.G.; Kim, J.L.; Jhoo, J.H.; et al. Incidence of and Risk Factors for Alzheimer’s Disease and Mild Cognitive Impairment in Korean Elderly. *Dement. Geriatr. Cogn. Disord.* **2015**, *39*, 105–115. [\[CrossRef\]](#)
34. Bartolini, M.; Coccia, M.; Luzzi, S.; Provinciali, L.; Ceravolo, M.G. Motivational Symptoms of Depression Mask Preclinical Alzheimer’s Disease in Elderly Subjects. *Dement. Geriatr. Cogn. Disord.* **2004**, *19*, 31–36. [\[CrossRef\]](#)
35. Blasko, I.; Kemmler, G.; Jungwirth, S.; Wichart, I.; Krampla, W.; Weissgram, S.; Jellinger, K.; Tragl, K.H.; Fischer, P. Plasma Amyloid Beta-42 Independently Predicts Both Late-Onset Depression and Alzheimer Disease. *Am. J. Geriatr. Psychiatry* **2010**, *18*, 973–982. [\[CrossRef\]](#)
36. Burke, S.L.; Cadet, T.; Alcide, A.; O’Driscoll, J.; Maramaldi, P. Psychosocial risk factors and Alzheimer’s disease: The associative effect of depression, sleep disturbance, and anxiety. *Aging Ment. Health* **2018**, *22*, 1577–1584. [\[CrossRef\]](#)
37. Chen, P.; Ganguli, M.; Mulsant, B.H.; DeKosky, S.T. The Temporal Relationship Between Depressive Symptoms and Dementia. *Arch. Gen. Psychiatry* **1999**, *56*, 261–266. [\[CrossRef\]](#)
38. Forno, G.D.; Palermo, M.T.; Donohue, J.E.; Karagiozis, H.; Zonderman, A.B.; Kawas, C.H. Depressive symptoms, sex, and risk for Alzheimer’s disease. *Ann. Neurol.* **2005**, *57*, 381–387. [\[CrossRef\]](#)
39. Devanand, D.P.; Sano, M.; Tang, M.-X.; Taylor, S.; Gurland, B.J.; Wilder, D.; Stern, Y.; Mayeux, R. Depressed Mood and the Incidence of Alzheimer’s Disease in the Elderly Living in the Community. *Arch. Gen. Psychiatry* **1996**, *53*, 175–182. [\[CrossRef\]](#)
40. Dotson, V.M.; Beydoun, M.A.; Zonderman, A.B. Recurrent depressive symptoms and the incidence of dementia and mild cognitive impairment. *Neurology* **2010**, *75*, 27–34. [\[CrossRef\]](#)
41. Fuhrer, R.; Dufouil, C.; Dartigues, J.F.; For the PAQUID Study. Exploring Sex Differences in the Relationship Between Depressive Symptoms and Dementia Incidence: Prospective Results From the PAQUID Study. *J. Am. Geriatr. Soc.* **2003**, *51*, 1055–1063. [\[CrossRef\]](#) [\[PubMed\]](#)
42. Geerlings, M.I.; Bouter, L.M.; Schoevers, R.; Beekman, A.T.F.; Jonker, C.; Deeg, D.J.H.; Van Tilburg, W.; Adèr, H.J.; Schmand, B. Depression and risk of cognitive decline and Alzheimer’s disease. *Br. J. Psychiatry* **2000**, *176*, 568–575. [\[CrossRef\]](#)
43. Geerlings, M.I.; Heijer, T.D.; Koudstaal, P.J.; Hofman, A.; Breteler, M. History of depression, depressive symptoms, and medial temporal lobe atrophy and the risk of Alzheimer disease. *Neurology* **2008**, *70*, 1258–1264. [\[CrossRef\]](#)
44. Gracia-García, P.; De-La-Cámara, C.; Santabàrbara, J.; Lopez-Anton, R.; Quintanilla, M.A.; Ventura, T.; Marcos, G.; Campayo, A.; Saz, P.; Lyketsos, C.; et al. Depression and Incident Alzheimer Disease: The Impact of Disease Severity. *Am. J. Geriatr. Psychiatry* **2015**, *23*, 119–129. [\[CrossRef\]](#)
45. Hesper, K.; Tebarth, F.; Wiese, B.; Eisele, M.; Bickel, H.; Kohler, M.; Mosch, E.; Weyerer, S.; Werle, J.; König, H.-H.; et al. Age of major depression onset, depressive symptoms, and risk for subsequent dementia: Results of the German Study on Ageing, Cognition, and Dementia in Primary Care Patients (AgeCoDe). *Psychol. Med.* **2012**, *43*, 1597–1610. [\[CrossRef\]](#)
46. Irie, F.; Masaki, K.H.; Petrovitch, H.; Abbott, R.D.; Ross, G.W.; Taaffe, D.R.; Launer, L.J.; White, L.R. Apolipoprotein E ϵ 4 Allele Genotype and the Effect of Depressive Symptoms on the Risk of Dementia in Men. *Arch. Gen. Psychiatry* **2008**, *65*, 906–912. [\[CrossRef\]](#)
47. Kim, J.-M.; Kim, S.-Y.; Bae, K.-Y.; Kim, S.-W.; Shin, I.-S.; Yang, S.-J.; Song, Y.-H.; Yoon, J.-S. Apolipoprotein E4 Genotype and Depressive Symptoms as Risk Factors for Dementia in an Older Korean Population. *Psychiatry Investig.* **2010**, *7*, 135–140. [\[CrossRef\]](#)
48. Kim, J.-M.; Stewart, R.; Kim, S.-Y.; Kim, S.-W.; Bae, K.-Y.; Yang, S.-J.; Shin, I.-S.; Yoon, J.-S. Synergistic associations of depression and apolipoprotein E genotype with incidence of dementia. *Int. J. Geriatr. Psychiatry* **2010**, *26*, 893–898. [\[CrossRef\]](#)
49. Lauriola, M.; Mangiacotti, A.; D’Onofrio, G.; Cascavilla, L.; Paris, F.; Ciccone, F.; Greco, M.; Paroni, G.; Seripa, D.; Greco, A. Late-Life Depression versus Amnesic Mild Cognitive Impairment: Alzheimer’s Disease Incidence in 4 Years of Follow-Up. *Dement. Geriatr. Cogn. Disord.* **2018**, *46*, 140–153. [\[CrossRef\]](#)
50. Lenoir, H.; Dufouil, C.; Auriacombe, S.; Lacombe, J.-M.; Dartigues, J.-F.; Ritchie, K.; Tzourio, C. Depression History, Depressive Symptoms, and Incident Dementia: The 3C Study. *J. Alzheimer’s Dis.* **2011**, *26*, 27–38. [\[CrossRef\]](#)

51. Li, G. Temporal Relationship Between Depression and Dementia. *Arch. Gen. Psychiatry* **2011**, *68*, 970–977. [[CrossRef](#)] [[PubMed](#)]
52. Luchsinger, J.A.; Reitz, C.; Honig, L.S.; Tang, M.-X.; Shea, S.; Mayeux, R. Aggregation of vascular risk factors and risk of incident Alzheimer disease. *Neurology* **2005**, *65*, 545–551. [[CrossRef](#)]
53. Reding, M.; Haycox, J.; Blass, J. Depression in Patients Referred to a Dementia Clinic. *Arch. Neurol.* **1985**, *42*, 894–896. [[CrossRef](#)]
54. Richard, E.; Reitz, C.; Honig, L.H.; Schupf, N.; Tang, M.X.; Manly, J.J.; Mayeux, R.; Devanand, D.; Luchsinger, J.A. Late-Life Depression, Mild Cognitive Impairment, and Dementia. *JAMA Neurol.* **2013**, *70*, 380–382. [[CrossRef](#)]
55. Saczynski, J.S.; Beiser, A.; Seshadri, S.; Auerbach, S.; Wolf, P.A.; Au, R. Depressive symptoms and risk of dementia: The Framingham Heart Study. *Neurology* **2010**, *75*, 35–41. [[CrossRef](#)]
56. Tyas, S.L.; Manfreda, J.; A Strain, L.; Montgomery, P.R. Risk factors for Alzheimer's disease: A population-based, longitudinal study in Manitoba, Canada. *Int. J. Epidemiol.* **2001**, *30*, 590–597. [[CrossRef](#)]
57. Vilalta-Franch, J.; López-Pousa, S.; Llinàs-Reglà, J.; Calvó-Perxas, L.; Merino-Aguado, J.; Garre-Olmo, J. Depression subtypes and 5-year risk of dementia and Alzheimer disease in patients aged 70 years. *Int. J. Geriatr. Psychiatry* **2012**, *28*, 341–350. [[CrossRef](#)]
58. Wilson, R.; Schneider, J.; Bienias, J.; Arnold, S.; Evans, D.; Bennett, D. Depressive symptoms, clinical AD, and cortical plaques and tangles in older persons. *Neurology* **2003**, *61*, 1102–1107. [[CrossRef](#)]
59. Kuring, J.K.; Mathias, J.L.; Ward, L. Prevalence of Depression, Anxiety and PTSD in People with Dementia: A Systematic Review and Meta-Analysis. *Neuropsychol. Rev.* **2018**, *28*, 393–416. [[CrossRef](#)]
60. Hamilton, M. Development of a Rating Scale for Primary Depressive Illness. *Br. J. Soc. Clin. Psychol.* **1967**, *6*, 278–296. [[CrossRef](#)]
61. Williams, J.B.W. A Structured Interview Guide for the Hamilton Depression Rating Scale. *Arch. Gen. Psychiatry* **1988**, *45*, 742–747. [[CrossRef](#)]
62. Copeland, J.R.M.; Dewey, M.E.; Griffiths-Jones, H.M. A computerized psychiatric diagnostic system and case nomenclature for elderly subjects: GMS and AGE-CAT. *Psychol. Med.* **1986**, *16*, 89–99. [[CrossRef](#)] [[PubMed](#)]
63. Jung, I.K.; Kwak, D.I.; Joe, S.H.; Lee, H.S. A Study of Standardization of Korean Form of Geriatric Depression Scale (KGDS). *J. Korean Geriatr. Psychiatry* **1997**, *1*, 61–72.
64. Yesavage, J.A.; Brink, T.; Rose, T.L.; Lum, O.; Huang, V.; Adey, M.; Leirer, V.O. Development and validation of a geriatric depression screening scale: A preliminary report. *J. Psychiatr. Res.* **1982**, *17*, 37–49. [[CrossRef](#)]
65. Kim, J.Y.; Park, J.H.; Lee, J.J.; Huh, Y.; Lee, S.B.; Han, S.K.; Choi, S.W.; Lee, N.Y.; Kim, K.W.; Woo, J.I. Standardization of the Korean Version of the Geriatric Depression Scale: Reliability, Validity, and Factor Structure. *Psychiatry Investig.* **2008**, *5*, 232–238. [[CrossRef](#)]
66. Radloff, L.S. The CES-D Scale. *Appl. Psychol. Meas.* **1977**, *1*, 385–401. [[CrossRef](#)]
67. Kohout, F.J.; Berkman, L.F.; Evans, D.A.; Cornoni-Huntley, J. Two Shorter Forms of the CES-D Depression Symptoms Index. *J. Aging Health* **1993**, *5*, 179–193. [[CrossRef](#)] [[PubMed](#)]
68. Green, R.C.; Cupples, L.A.; Kurz, A.; Auerbach, S.; Go, R.; Sadovnick, D.; Duara, R.; Kukull, W.A.; Chui, H.; Edeki, T.; et al. Depression as a Risk Factor for Alzheimer Disease. *Arch. Neurol.* **2003**, *60*, 753–759. [[CrossRef](#)] [[PubMed](#)]
69. Kessing, L.V.; Søndergård, L.; Forman, J.L.; Andersen, P.K. Antidepressants and dementia. *J. Affect. Disord.* **2009**, *117*, 24–29. [[CrossRef](#)]
70. Richardson, K.; Fox, C.; Maidment, I.; Steel, N.; Loke, Y.K.; Arthur, A.; Myint, P.K.; Grossi, C.M.; Mattishent, K.; Bennett, K.; et al. Anticholinergic drugs and risk of dementia: Case-control study. *BMJ* **2018**, *361*, k1315. [[CrossRef](#)]
71. Wang, C.; Gao, S.; Hendrie, H.C.; Kesterson, J.; Campbell, N.L.; Shekhar, A.; Callahan, C.M. Antidepressant Use in the Elderly Is Associated with an Increased Risk of Dementia. *Alzheimer's Dis. Assoc. Disord.* **2016**, *30*, 99–104. [[CrossRef](#)] [[PubMed](#)]
72. Henderson, A.S.; Korten, A.E.; Jacomb, P.A.; MacKinnon, A.J.; Jorm, A.F.; Christensen, H.; Rodgers, B. The course of depression in the elderly: A longitudinal community-based study in Australia. *Psychol. Med.* **1997**, *27*, 119–129. [[CrossRef](#)]
73. Class, C.A.; Unverzagt, F.W.; Gao, S.; Sahota, A.; Hall, K.S.; Hendrie, H.C. The association between Apo E genotype and depressive symptoms in elderly African-American subjects. *Am. J. Geriatr. Psychiatry* **1997**, *5*, 339–343. [[CrossRef](#)] [[PubMed](#)]
74. Forsell, Y.; Corder, E.H.; Basun, H.; Lannfelt, L.; Viitanen, M.; Winblad, B. Depression and dementia in relation to apolipoprotein E polymorphism in a population sample age 75+. *Biol. Psychiatry* **1997**, *42*, 898–903. [[CrossRef](#)]
75. Zubenko, G.S.; Henderson, R.; Stiffler, J.S.; Stabler, S.; Rosen, J.; Kaplan, B.B. Association of the APOE ϵ 4 allele with clinical subtypes of late life depression. *Biol. Psychiatry* **1996**, *40*, 1008–1016. [[CrossRef](#)]
76. Puy, L.; Jouvent, E. Accidente cerebrovascular en el paciente anciano. *EMC Tratado de Medicina* **2020**, *24*, 1–6. [[CrossRef](#)]
77. Ouanes, S.; Popp, J. High Cortisol and the Risk of Dementia and Alzheimer's Disease: A Review of the Literature. *Front. Aging Neurosci.* **2019**, *11*, 43. [[CrossRef](#)]
78. Hong, H.; Kim, B.S.; Im, H.-I. Pathophysiological Role of Neuroinflammation in Neurodegenerative Diseases and Psychiatric Disorders. *Int. Neurol.* **2016**, *20*, 52–57. [[CrossRef](#)] [[PubMed](#)]