



Non-medical Cannabis Self-Exposure as a Dimensional Predictor of Opioid Dependence Diagnosis: A Propensity Score Matched Analysis

Eduardo R. Butelman^{1*†}, Angelo G. I. Maremmani^{2†}, Silvia Bacciardi², Carina Y. Chen¹, Joel Correa da Rosa³ and Mary Jeanne Kreek¹

¹ Laboratory on the Biology of Addictive Diseases, Rockefeller University, New York, NY, United States, ² "VP Dole" Dual Diagnosis Unit, Azienda Ospedaliera-Universitaria Pisana, Pisa, Italy, ³ Center for Clinical and Translational Science, The Rockefeller University Hospital, Rockefeller University, New York, NY, United States

OPEN ACCESS

Edited by:

Lawrence Toll, Florida Atlantic University, United States

Reviewed by:

Robert F. Leeman, University of Florida, United States Liangsuo Ma, University of Texas Health Science Center at Houston, United States

*Correspondence:

Eduardo R. Butelman butelme@rockefeller.edu

[†]These authors have contributed equally to this work.

Specialty section:

This article was submitted to Addictive Disorders, a section of the journal Frontiers in Psychiatry

Received: 27 March 2018 Accepted: 11 June 2018 Published: 27 June 2018

Citation:

Butelman ER, Maremmani AGI, Bacciardi S, Chen CY, Correa da Rosa J and Kreek MJ (2018) Non-medical Cannabis Self-Exposure as a Dimensional Predictor of Opioid Dependence Diagnosis: A Propensity Score Matched Analysis. Front. Psychiatry 9:283. doi: 10.3389/fpsyt.2018.00283 **Background:** The impact of increasing non-medical cannabis use on vulnerability to develop opioid use disorders has received considerable attention, with contrasting findings. A dimensional analysis of self-exposure to cannabis and other drugs, in individuals with and without opioid dependence (OD) diagnoses, may clarify this issue.

Objective: To examine the age of onset of maximal self-exposure to cannabis, alcohol, cocaine, and heroin, in volunteers diagnosed with OD, using a rapidly administered instrument (the KMSK scales). To then determine whether maximal self-exposure to cannabis, alcohol, and cocaine is a dimensional predictor of odds of OD diagnoses.

Methods: This outpatient observational study examined maximal self-exposure to these drugs, in volunteers diagnosed with DSM-IV OD or other drug diagnoses, and normal volunteers. In order to focus more directly on opioid dependence diagnosis as the outcome, volunteers who had cocaine dependence diagnoses were excluded. Male and female adults of diverse ethnicity were consecutively ascertained from the community, and from local drug treatment programs, in 2002–2013 (n = 574, of whom n = 94 had OD diagnoses). The age of onset of maximal self-exposure of these drugs was examined. After propensity score matching for age at ascertainment, gender, and ethnicity, a multiple logistic regression examined how increasing self-exposure to non-medical cannabis, alcohol and cocaine affected odds of OD diagnoses.

Results: Volunteers with OD diagnoses had the onset of heaviest use of cannabis in the approximate transition between adolescence and adulthood (mean age = 18.9 years), and onset of heaviest use of alcohol soon thereafter (mean age = 20.1 years). Onset of heaviest use of heroin and cocaine was detected later in the lifespan (mean ages = 24.7 and 25.3 years, respectively). After propensity score matching for demographic variables, we found that the maximal self-exposure to cannabis and cocaine, but not to alcohol, was greater in volunteers with OD diagnoses, than in those without this diagnosis. Also, a multiple logistic regression detected that increasing

1

self-exposure to cannabis and cocaine, but not alcohol, was a positive predictor of OD diagnosis.

Conclusions/Importance: Increasing self-exposure to non-medical cannabis, as measured with a rapid dimensional instrument, was a predictor of greater odds of opioid dependence diagnosis, in propensity score-matched samples.

Keywords: opioid, cocaine, cannabis, heroin, alcohol, dimensional, exposure, adolescence

INTRODUCTION

Addictions to heroin or illicitly used prescription opioids (shortacting MOP-r agonists) cause major morbidity and mortality (1, 2), and there is considerable poly-drug use in persons with these diseases (3-6). Use of other substances, especially cannabis and alcohol, often precedes non-medical use of MOP-r agonists. The impact of non-medical cannabis use with respect to vulnerability to develop an opioid use disorder remains under study (7). This has been examined primarily with categorical classifications of cannabis use, such as "any use" vs. "no use" or presence vs. absence of a diagnosed cannabis use disorder. For example, epidemiological studies have shown that any cannabis use is associated with a later increase in odds of non-medical use of opioids and other drugs (8, 9). Another recent report from the NESARC longitudinal study found that any use of cannabis at the "wave 1" time point (2001–2002), was a positive predictor of both non-medical prescription opioid use, and opioid use disorder at the "wave 2" time point (2004-2005) (10).

In this study, we focus on dimensional aspects of drug selfexposure and their relationship to an opioid dependence diagnosis (OD). Dimensional measures are those that characterize a behavioral or biological variable along some form of a continuum. Specifically, we examined the ages of onset of heaviest use of different drugs in volunteers with opioid dependence diagnosis, as well as the level of maximal self-exposure. Dimensional aspects of substance use disorders (SUDs) are receiving recent attention, both for examination of disease progression and for the examination of mechanistic and genetic features (11–13).

Intriguingly, some studies have found that state-wide availability of medical cannabis has resulted in decreases in ageadjusted opioid overdose mortality (14), and other apparent protective effects (15). Experimental studies in humans do not detect a protective effect of the main psychoactive component of cannabis (the CB-1 partial agonist Δ 9-tetra-hydro-cannabinol; Δ 9-THC) on MOP-r agonist-induced respiratory depression, which is the underlying cause of overdose mortality (16). Also, cannabis smoking produced a small but significant increase in the abuse potential of a MOP-r agonist, in a recent laboratory study (17). Studies on the effectiveness of Δ 9-THC in decreasing severity of withdrawal from MOP-r agonists have yielded mixed results, possibly due to different methods used (18, 19).

Some preclinical data show that exposure to Δ 9-THC in adolescence can increase vulnerability to the addiction-related effects of MOP-r agonists in adulthood (20–22). Some, but not all, preclinical studies suggest that greater exposure to a CB1-r agonist could cause neurobiological effects that increase subsequent vulnerability to addictive-like effects of MOP-r agonists (23–25).

At least two major theories have been proposed to account for the sequence of first use of drugs, and also for specific patterns of poly-drug exposure in persons with specific SUD. Two of these major theories have been termed the "gateway theory" and the "shared vulnerability theory," and their relative impact remains an area of controversy (6, 26–28). An examination of dimensional, as opposed to categorical, aspects of drug self-exposure could also provide a framework to further understand the aforementioned phenomena (11, 29).

Given the changes in cannabis availability and use, and the ongoing epidemic of opioid use disorders, this controversy is of current importance (30, 31). It has been suggested that dimensional data at the individual level would be of value to address this issue (32–34). However, few studies have examined dimensionally, how exposure to non-medical cannabis and other drugs can affect odds of a clinically diagnosed opioid use disorder, at the individual level (10, 35–37). Furthermore, in most studies where such data was examined, the instruments used are not suitable for general clinical or preventive practice, due primarily to their length. In this study, we therefore examined dimensionally how different amounts of self-exposure to major drugs of abuse including non-medical cannabis and alcohol, affected odds of developing an opioid dependence diagnosis, using a relatively rapid and simple instrument (38, 39).

MATERIALS AND METHODS

This was an observational study, with consecutively ascertained adult volunteers who were examined in an outpatient research hospital setting, in the New York City area. This cohort was originally recruited and ascertained as part of genetic association studies of SUD (40–42).

Volunteers

The main outcome under examination was the presence or absence of a DSM-IV opioid dependence (OD) diagnosis. Many

Abbreviations: 95%CI, 95% Confidence interval; Δ 9-THC, delta9tetrahydrocannabinol; CB1-r, Cannabinoid-1 receptor; IQR, Inter-quartile range; KMSK scale, Kreek-McHugh-Schluger-Kellogg scale for maximal selfexposure to specific drugs; MOP-r, mu-opioid receptors; N.S., Non-significant; OD, opioid dependence diagnosis (DSM-IV criteria); OUD, Opioid use disorder (DSM-5 criteria); ROC curve, Receiver operating characteristic curve; SUD, Substance use disorders.

of the volunteers with SUD also had other diagnoses in addition to OD, but the presence of a cocaine dependence diagnosis was an exclusion criterion for this study. However, volunteers with the relatively less severe DSM-IV diagnosis of cocaine *abuse* were not excluded. Volunteers were ascertained sequentially from a number of addictive disease treatment clinics in the greater New York City area, and from the local community in the same area.

Recruitment, Inclusion and Exclusion Criteria

This study was carried out in accordance with the recommendations for Human Subjects Policies and Guidance of the National Institutes of Health. The protocol was approved by the Rockefeller University Hospital Institutional Review Board (IRB). All subjects gave written informed consent in accordance with the Declaration of Helsinki.

Male and female volunteers (≥ 18 years of age) were recruited through IRB-approved posted notices and newspaper advertisements in the community. Volunteers were required to be competent to understand study procedures and understand and sign the IRB-approved informed consent in English. The presence of uncontrolled schizophrenia or other psychotic signs during the interview were exclusion criteria. In order to focus more directly on the impact of cannabis or alcohol exposure on the odds of OD diagnoses, we excluded from this study volunteers who had a cocaine dependence diagnosis. Volunteers who had used cocaine, but did not meet the DSM-IV diagnostic criteria for cocaine dependence diagnosis were not excluded.

Persons were excluded from the normal volunteer category if they had any lifetime drug abuse or dependence diagnosis by DSM-IV criteria, or any of the following: (a) any instance of drinking to a level of intoxication during the previous 30 days, (b) any use of illicit drugs including opiates, cocaine, and amphetamines during the 30 days prior to ascertainment, (c) if they had used cannabis on more than 12 days during the 30 days prior to ascertainment, (d) had used illicit drugs (with the exception of cannabis) for at least three times a week for a period of at least 1 month, in their lifetime (40). This therefore allowed for examination of a range of normative self-exposure to cannabis and alcohol, also in the normal volunteers.

The three diagnostic groups in this study are volunteers with opioid dependence (OD), volunteers with drug diagnoses except OD, and normal volunteers. These groups are described in further detail in **Table 1**. In further analyses, the latter two groups were combined into an overall "not OD" group, for analyses of patterns of self-exposure to specific drugs.

Instruments Administered During Clinical Interviews

All ascertainments were completed during a standardized private face-to-face interview with a licensed trained clinician (e.g., M.D., D.O., Ph.D. Psychologist, Nurse Practitioner or Registered Nurse).Volunteers underwent the SCID I/P structured interview (Version 2.0; DSM-IV criteria) (43), and received the KMSK TABLE 1 | Description of diagnostic groups (DSM-IV criteria).

	Volunteers with opioid dependence (OD) diagnosis	Volunteers with drug diagnoses, except OD	Normal volunteers
Description	Volunteers with opioid dependence diagnosis, as well as other drug diagnoses (if applicable)	Volunteers with any drug diagnoses, <i>except</i> OD	Volunteers without any drug diagnoses
	Cocaine dependence diagr criterion for this study	nosis was an overall exclusio	n

questionnaires for maximum self-exposure to cannabis, alcohol, heroin, and cocaine (38) (see below).

KMSK Scales for Maximal Self-Exposure to Specific Drugs ("KMSK Score")

The KMSK scales for cannabis, alcohol, heroin, and cocaine provide ordinal measures of maximal self-exposure, thus focusing on the period in the volunteer's life when use was the heaviest. For each drug, the scales start at a minimum "0" score, which denotes that the volunteer has not had any lifetime exposure to the drug (i.e., no use). The scores then increase in integers up to a maximum (13 for heroin and alcohol, 14 for cannabis, and 16 for cocaine) (see Table 2). The KMSK score for each drug is the composite sum of responses on three items: (a) frequency of maximal use (e.g., in times per day or per week), (b) duration of pattern of maximal use (e.g., in months or years) and (c) amount used in one day or sitting (e.g., number of alcoholic drinks or cannabis joints) (38). A separate KMSK scale is also used to characterize illicit use of prescription opioids, but was not analyzed in this study (heroin was the predominant MOP-r agonist used in this cohort). Concurrent validity of KMSK scores with the respective DSM-IV dependence diagnoses has been examined, and yielded optimal "cutpoint" scores for sensitivity and specificity (38, 39).

The KMSK scales have been used to characterize drug exposure in patients with medical and psychiatric conditions (44–46). The scales can be rapidly administered within a clinical interview (e.g., \leq 5 min per drug). Each KMSK form also records age of first use, and age of onset of heaviest use (in whole years; the latter was studied herein). The four KMSK forms used in this study (i.e., for cannabis, alcohol, cocaine, and heroin), are provided in the Supplementary Materials. The full text of the scales for these and other drugs can be freely accessed: http://lab. rockefeller.edu/kreek/assets/file/KMSKquestionnaire.pdf.

Statistical Analyses Missing Data

If there were missing data for specific comparisons for a volunteer, the data for that volunteer was removed from analysis. The cannabis KMSK scale and the age-related items were implemented while cohort ascertainment was in progress. Therefore these items, especially for cannabis, were not available for the complete cohort.

TABLE 2 | KMSK scales for maximal self-exposure to specific drugs^{a,b}.

Drugs		Sub-scores	KMSK score range: (sum			
	Frequency of use ^e	Duration of pattern ^f	Amount used in a sitting or a day ^g	 of sub-scores) 	(males and females combined) ^h	
	Range: never used→multiple daily use	Range: <6 months> → 1year	Range: (see below)			
Cannabis ^c	0→6	0→3	0 → 5 None → >5 joints	0–14	10	
Alcohol ^c	0→5	0→3	0 → 5 None → >10 drinks	0–13	10	
Heroin ^d	0→4	0 → 3	0 → 6 None → >10 doses/bags	0–13	6	
Cocaine ^d	0→7	0 → 3	0→6 \$0→>\$100 (also converted from grams, rocks or vials)	0–16	9	

^aOrdinal integer scales; quantifying drug self-exposure at the time in the volunteer's life when use is heaviest.

^b Scales for alcohol, heroin, and cocaine were published initially("KMSK-1") (38). Scales for cannabis were developed and used subsequently ("KMSK-2") (39).

 $^{\rm c}$ If the Frequency sub-score ${\leq}2,$ the Duration sub-score is assigned a "0" value.

 d If the Frequency sub-score \leq 1, the Duration sub-score is assigned a "0" value.

^eQuestionnaire text reads: "At the time in your life when you were using the most [drug], were you using it."

^fQuestionnaire text reads: "How long did this pattern of [drug] use last?"

^gQuestionnaire text reads: "During this time when you were using the most, how much [drug] at a sitting [or day] would you typically use?"

^hOptimal cutpoints for concurrent validity in males and females combined, for the respective DSM-IV dependence diagnosis (unpublished data).

Univariate Analyses

Univariate analyses were carried out with GraphPad Prism software. Demographic variables (age at ascertainment, gender, and ethnicity) and KMSK scores were analyzed non-parametrically (Mann–Whitney U-tests or χ^2 analyses, Kruskal–Wallis or Friedman's ANOVAs, and Dunn's *post-hoc* tests).

Propensity Score Matching

As shown in **Table 3**, the overall cohort had a total of n = 574 volunteers, of whom n = 94 had OD diagnoses. Of the volunteers with OD diagnoses, n = 89 had all KMSK scores available, and were used in the propensity score matching procedure. Propensity score matching (47), as implemented in the "MatchIt" package in R software, was applied using a 1:1 "nearest neighbor" algorithm, to minimize heterogeneity in the above demographic variables between volunteers with OD, vs. all the other volunteers. Therefore, the latter comparison group contained the volunteers with drug diagnoses except OD, and also normal volunteers (see **Table 3B**).

Multiple Logistic Regression After Propensity Score Matching, Examining Cannabis, Alcohol and Cocaine KMSK Scores as Dimensional Predictors of Opioid Dependence Diagnosis

A multiple logistic regression was performed with Statistica (TIBCO) software. The predicted outcome was the presence of opioid dependence diagnosis (binary). There were n = 89 volunteers with and n = 89 without OD diagnoses in this sample, after the propensity score matching procedure, described above.

Alpha Level for Rejection of Null Hypotheses

For all analyses, the alpha level was $p \le 0.05$.

RESULTS

Sample Demographics

Sample Demographics are in **Table 3**. In order to provide a complete description of the cohort, **Table 3A** presents data for volunteers with OD diagnoses, volunteers with drug diagnoses except OD, and normal volunteers. **Table 3B** presents the same data, but the latter two groups are combined, as this is the design used in the propensity score matching procedure. See also **Table 5**, for demographic data in the two groups after execution of the propensity score matching procedure.

Age at Ascertainment

Mean age at ascertainment was greater for volunteers with OD diagnoses, and also for volunteers with drug diagnoses except OD, vs. normal volunteers (**Table 3A**).

Gender

A χ^2 analysis of gender was significant, with a greater proportion of males among volunteers with OD diagnoses, or drug diagnoses except OD, vs. normal volunteers (**Table 3A**).

Ethnicity

A χ^2 analysis of ethnicity was also significant, with a relatively greater proportion of African Americans in the normal volunteer group, and a relatively greater proportion of Caucasians and Hispanics in the OD group (**Table 3A**).

Missing Data

Of the 94 volunteers with OD diagnosis (see **Table 3A**), five were removed from further analysis, due to missing data (remaining n = 89). Therefore, the propensity score matching procedure used these n = 89 volunteers with OD diagnoses as the reference

TABLE 3 | Demographics (volunteers sequentially ascertained 4/4/02-8/1/13)^a.

(A) Comparison of volunteers with opioid dependence (OD) diagnoses, volunteers with drug diagnoses except OD, and normal volunteers.

Demographics Mean age at ascertainment(SEM)	Total <i>n</i> = 574	Volunteers with OD diagnosis ^b n = 94		Volunteers with drug diagnoses, except OD^{C} n = 187		Normal volunteers ^C (NV) n = 293		Kruskal-Wallis statistic o χ^2 [df]; p value	
		41.17	(1.23)	39.7	(0.86)	33.4	(0.70)	59.34	<0.0001
Gender	Male	65	69.1%	124	66.3%	132	45.1%	28.91 [2]	< 0.0001
	Female	29	30.1%	63	33.7%	161	55.0%		
Ethnicity	African-American	25	26.6%	78	41.7%	127	43.3%	22.12 [6]	0.0012
	Caucasian	31	33.0%	59	31.6%	81	27.7%		
	Hispanic	32	34.0%	31	16.6%	47	16.0%		
	Other	6	6.4%	19	10.2%	38	13.0%		

(B) Comparison of volunteers with OD diagnosis vs. all volunteers without OD (i.e., combining volunteers with drug diagnoses except OD, and normal volunteers).

Demographics	Total <i>n</i> = 574	Volunteers with OD diagnosis ^b $n = 94$		All volunteers without OD diagnosis n = 480		U or χ ² [df <i>p</i> -value];
Mean age at ascertainment (SEM)		41.17	(1.23)	35.81	(0.5)	16,566	< 0.0001
Gender	Male	65	69.1%	256	53.3%	7.98 [1]	0.0047
	Female	29	30.1%	224	46.7%		
Ethnicity	African-American	25	26.6%	205	42.7%	20.62 [3]	0.0001
	Caucasian	31	33.0%	140	29.2%		
	Hispanic	32	34.0%	78	16.3%		
	Other	6	6.4%	57	11.9%		

^aCocaine dependence diagnosis was an exclusion criterion for this study. See **Table 1** for further description of diagnostic groups.

^b The same data from volunteers with OD diagnosis are presented in (A,B).

^c This group combines the two right-most columns in (A). The two groups presented in (B) are used as the input data for the propensity score matching procedure (see Table 5).

TABLE 4 | Ages of onset of heaviest use of specific drugs, in volunteers with opioid dependence diagnosis (data available for each of the drugs from n = 47).

Drug	Age of onset of heaviest use mean [95%CI] ^a					
Cannabis	18.9 [16.6–21.1] ^b					
Alcohol	20.1 [18.1–22.3] ^c					
Heroin	24.7 [21.9–27.5]					
Cocaine	25.3 [22.6–27.9]					

^aFriedman's ANOVA F₍₄₎ = 29.22; p < 0.0001.

^bDunn's post-hoc tests: cannabis<heroin; cannabis<cocaine.

^cDunn's post-hoc tests: alcohol<heroin; alcohol<cocaine.

group (see below and **Table 5**). Of the 187 volunteers with a drug diagnosis except OD, 13 were removed due to missing data (remaining n = 174). Also, of the 293 normal volunteers, 12 were similarly removed due to missing data (remaining n = 281).

Ages of Onset of Heaviest Use of Different Drugs, in Volunteers With Opioid Dependence Diagnosis

The mean age of onset of heaviest use of cannabis, alcohol, cocaine and heroin are presented in **Table 4**, for volunteers with opioid dependence diagnosis, for whom all these data were available. A Friedman's ANOVA examining these data was

significant [$F_{(4)} = 29.22$; p < 0.0001]. Dunn's *post-hoc* tests show that the age of onset of heaviest use of cannabis use was earlier than that for heroin or cocaine. Likewise, age of onset of heaviest use of alcohol was earlier than that for heroin or cocaine. Ages of onset of heaviest use did not differ between cannabis and alcohol, or between heroin and cocaine.

Propensity Score Matching Procedure for Demographic Variables

As shown in Table 3B, there were demographic differences between the group with OD diagnoses and the group without OD diagnoses (the latter group being the combination of volunteers with drug diagnoses except OD, and normal volunteers). The goal of the propensity score matching procedure was to minimize the impact of the demographic differences. As is common in propensity score matching procedures, we initially utilized a multiple logistic regression to examine the demographic variables (age at ascertainment, gender, and ethnicity) as predictors of the OD diagnosis outcome. Propensity scores were then generated for each volunteer in the whole cohort, as the predicted values from this regression. These propensity scores were then entered in a matching algorithm as described in the section Materials and Methods. This algorithm selected n = 89 volunteers without OD diagnoses, to match the reference group of n = 89 volunteers with OD diagnoses. This matching procedure was effective,

TABLE 5 | Demographics after the propensity score matching procedure (see Table 3B for data prior to the matching procedure).

Demographics	Total <i>n</i> = 178	Volunteers with OD diagnosis $n = 89^a$		Volunteers without OD diagnosis		U or χ^2 [df]; p-value	
Mean age at ascertainment (SEM)		41.37	(1.30)	41.60	(1.25)	3,894	N.S. p = 0.85
Gender	Male	61	68.5%	56	62.9%	0.985 [1]	N.S. $p = 0.80$
	Female	28	31.5%	33	37.1%		
Ethnicity	African-American	25	28.1%	20	22.5%	0.624 [3]	N.S. $p = 0.43$
	Caucasian	27	30.3%	31	34.8%		
	Hispanic	31	34.8%	33	37.1%		
	Other	6	6.7%	5	5.6%		

^aAs mentioned in text, data from 5 of the n = 94 volunteers with OD diagnoses (as described in **Table 3**) had to be excluded from the matching procedure, due to missing data (thus having a remaining group of n = 89 volunteers with OD diagnoses).

as confirmed by the lack of significant differences in gender, ethnicity and age at ascertainment, for the two groups (**Table 5**).

Maximal Self-Exposure to Cannabis, Alcohol, and Cocaine Compared in Volunteers With and Without Opioid Dependence Diagnosis, After Propensity Score Matching

Volunteers with OD had significantly greater KMSK scores for cannabis and cocaine, compared to propensity score—matched volunteers without OD (**Figure 1**). Alcohol KMSK scores did not differ significantly between these two groups. As expected, volunteers with an OD diagnosis had significantly greater heroin KMSK scores, compared to volunteers without this diagnosis (not shown) (38).

Multiple Logistic Regression Examining Cannabis, Alcohol, and Cocaine KMSK Scores as Predictors of Opioid Dependence Diagnosis, After Propensity Score Matching

This multiple logistic regression was carried out after propensity score matching for the demographic variables, as indicated above. A Wald test for a global null hypothesis was significant $[\chi^2_{(df=3)} = 25.05; p < 0.0001]$, showing that the coefficients for the predictor variables were significantly different from 0. A Hosmer–Lemeshow test was non-significant, suggesting no evidence of lack of fit. Cannabis and cocaine KMSK scores were each detected as significant positive predictors of odds of OD diagnosis (**Figure 2**). By contrast, alcohol KMSK scores were not a significant predictor. Odds ratios are presented per point in each KMSK scale (score ranges in the scales are described in **Table 2**).

DISCUSSION

The impact of non-medical cannabis and alcohol use on vulnerability to develop an opioid use disorder, and to recover therefrom, has received considerable recent attention (3, 14,

15, 48). This is an area of current public health importance, given evolving trends in cannabis status across jurisdictions, the ongoing epidemic of opioid use disorders (37, 49), and the increase in prevalence of alcohol use disorders (50). However, few studies have examined dimensionally how exposure to several major drugs, especially non-medical cannabis and alcohol, impacts odds of opioid dependence diagnosis (6, 10, 35, 36, 51, 52).

Ages of Onset of Heaviest Use of Each Drug in Volunteers With Opioid Dependence Diagnosis

We found the ages of onset of heaviest use of both cannabis and alcohol preceded the onset of heaviest use of heroin, in volunteers with OD diagnosis. The ages of onset of heaviest use of cannabis and alcohol did not differ from each other, and occurred in the period of transition from adolescence to adulthood (27, 53). In this group of volunteers with an OD diagnosis, age of onset of heaviest use of cocaine occurred at a similar age as that for heroin. This overall pattern has some similarity to those previously reported (6, 53), but focuses more directly here on the ages of onset of maximal use, rather than on first use. In the context of this study, the aforementioned data provided a rationale for examining cannabis and alcohol KMSK scores as dimensional predictors of OD diagnosis. We also opted to include cocaine KMSK scores as a predictor in the multiple regression below, in order to control for differing levels of exposure to this drug that could occur in volunteers with OD diagnoses (even after exclusion of volunteers with DSM-IV cocaine dependence diagnoses).

Demographic Variables and Rationale for Propensity Score Matching

We detected significant differences in major demographic variables (age at ascertainment, gender, and ethnicity) between the different diagnostic groups. As mentioned above, this was a study of consecutive volunteers responding to advertisements in the community and in drug treatment programs, from a large ethnically diverse urban area. This may have therefore affected some of the demographic parameters of the sample. For example,

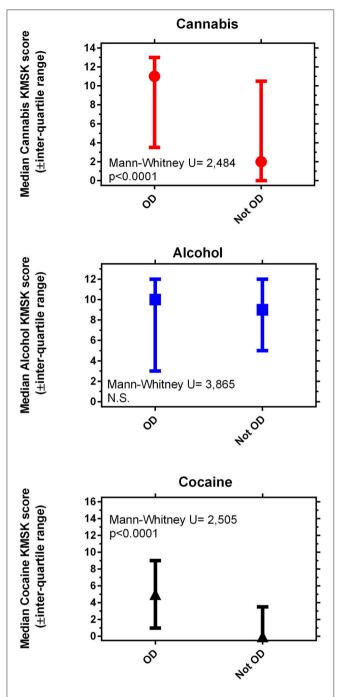
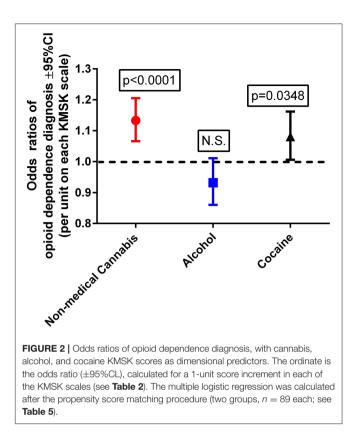


FIGURE 1 | Maximal self-exposure (KMSK scores) to non-medical cannabis (A), alcohol (B), and cocaine (C), upper, middle and lower panels, respectively, in volunteers with and without opioid dependence diagnoses ("OD" and "not OD," respectively). The "not OD" group is the combination of volunteers with drug diagnoses except OD, and normal volunteers. The data represent two groups of n = 89 each, after the propensity score matching procedure (see **Table 5**).

epidemiological studies show that the prevalence of specific SUD can differ based on major demographic factors, including gender (54, 55). We therefore elected to carry out a propensity



score matching procedure for age at ascertainment, gender, and ethnicity, prior to further analysis of maximal self-exposure to specific drugs. This matching procedure was effective in yielding groups with and without OD diagnoses, which did not differ significantly with respect to the aforementioned demographic variables.

Maximal Self-Exposure to Cannabis, Alcohol, and Cocaine, in Volunteers With and Without Opioid Dependence Diagnoses, After Propensity Score Matching

We found that both cannabis and cocaine KMSK scores were significantly greater in volunteers with OD diagnoses, vs. those without this diagnosis. Of note, the median cannabis KMSK score of volunteers with an OD diagnosis was relatively high, denoting heavy self-exposure to cannabis in this clinical group (based on a prior concurrent validity analysis with the DSM-IV cannabis dependence diagnosis (39) and unpublished data). Alcohol KMSK scores did not differ between volunteers with and without OD diagnosis, and a broad range of alcohol scores was observed in the two propensity score-matched groups. We note that cocaine KMSK scores were significantly greater in volunteers with OD than those without this diagnosis, even though a cocaine dependence diagnosis was an exclusion criterion for this study. As expected, the median cocaine KMSK score in the volunteers with OD was lower than the previously determined optimal "cutpoint" for the cocaine dependence diagnosis (38), due to the aforementioned exclusion criterion. However, the median cocaine KMSK score was even lower in the volunteers without OD (at least 50% of this group reported no lifetime cocaine use).

Multiple Logistic Regression With Cannabis, Alcohol, and Cocaine KMSK Scores as Predictors, After Propensity Score Matching

After propensity score matching for age at ascertainment, gender, and ethnicity, cannabis and cocaine KMSK scores were each positive predictors of odds of an OD diagnosis. The odds in this regression were generated per 1-unit change on each KMSK scale. Therefore, it can be observed that any use of cannabis (i.e., cannabis KMSK score \geq 1) is a predictor of increased odds of OD diagnosis (i.e., odds ratio = 1.13 per point in the cannabis KMSK scale). This study also shows that the odds of an OD diagnosis increase gradually with greater cannabis self-exposure scores. As mentioned above, recent epidemiological data from the NESARC study show that any use of cannabis at the "wave 1" time point (in 2001-2002) was a predictor of greater odds of opioid use disorder at the "wave 2" time point (in 2004-2005) (10). Other studies based on NESARC show that a dimensional measure of cannabis use (i.e., a defined frequency of use in the past year) at "wave 1" was a positive predictor of several SUDs (3), but opioid use disorder was not presented as a specific outcome in that study. Overall, several studies have examined primarily categorical measures of cannabis use as predictors of initiation of other drug use (56, 57). This study therefore provides the first rapid dimensional analysis which detects that self-exposure to non-medical cannabis is a positive predictor of odds of OD diagnosis, in a propensity score-matched sample.

Increasing cocaine self-exposure was also detected as a positive predictor of OD diagnosis, even though volunteers with cocaine dependence diagnoses were excluded from study. Therefore, even relatively smaller amounts of cocaine selfexposure are also associated with an increase in odds of OD diagnoses (further discussed below, in the "Limitations and Design Considerations" section). By contrast, alcohol KMSK scores were not a significant predictor of odds of an OD diagnosis. This study adds to the available literature on different aspects of alcohol use that may be related to opioid use disorders (9). Overall, it can be hypothesized that pharmacological or downstream neurobiological effects of cannabis, but not alcohol, can result in greater later vulnerability to opioid use disorders. An alternative interpretation, in the context of the "common liability" theory (27), is that there is a common pre-existing liability between cannabis and opioid exposure, and that alcohol does not share this liability to the same extent.

In preclinical studies, peri-adolescent exposure to $\Delta 9$ -THC produces long-lasting neurobiological changes to MOP-r and dopaminergic systems, which mediate direct and indirect effects of MOP-r agonists (20, 22, 58–60). There is also evidence that some of the behavioral and downstream neurobiological effects of $\Delta 9$ -THC are partially shared with MOP-r systems (25, 61). Preclinical studies show that the amount and pattern of

exposure to specific drugs of abuse are critical in the emergence of underlying neurobiological changes and of addiction-like behaviors (62–65). Overall, substantial non-medical cannabis exposure in adolescence and early adulthood may result in longlasting disruption in these and other systems, and thus result in increased vulnerability to the later development of opioid use disorders.

Limitations and Design Considerations

In this study, volunteers had to recall and report aspects of their drug exposure history. The possibility that recall bias may have affected these data cannot be excluded with this type of design (66), which is very common in studies of SUD (6, 35). Recalling the age(s) at which heaviest use of a specific drug occurred is also a demand of this scale. Studies with larger cohorts, different sampling methods, as well as longitudinal studies, could be used to further extend these findings. These volunteers were ascertained prior to the passage of the relevant medical marijuana statutes for this community (31). Therefore, these findings are not necessarily relevant to the impact of medically sanctioned cannabis. Studies with later birth cohorts could also investigate possible changes to the age trajectory of exposure to different drugs, due to environmental factors (1, 67–69).

We opted here to focus on the opioid dependence diagnosis as an outcome, and to exclude volunteers who had a cocaine dependence diagnosis. This allowed us to examine more directly the impact of cannabis and alcohol self-exposure on opioid dependence diagnosis as a clinical outcome. It is known that persons with dual severe opioid and cocaine use disorder diagnoses can have a different clinical course from those with only the former diagnosis (5, 70, 71). We observed that the volunteers with OD diagnoses still had significantly greater cocaine KMSK scores than the volunteers without this diagnosis. This is not surprising, as cocaine use is relatively common in persons who use heroin (6), and can occur even in the absence of a diagnosed cocaine dependence diagnosis. As mentioned in the Methods, volunteers with cocaine abuse diagnoses were not excluded from study. Therefore, we included cocaine KMSK scores in the multiple logistic regression, primarily to control for the level of cocaine self-exposure.

We elected to examine two propensity score matched groups: (a) volunteers with OD, and a comparison group: (b) volunteers without OD. The latter group thus included volunteers with drug diagnoses except OD, and normal volunteers. This allowed us to have a propensity score-matched comparison group with a broad range of KMSK scores for the drugs of interest, of value for a more robust dimensional analysis (11, 38, 72).

Propensity score matching studies have become relatively frequent, and have potential strengths and limitations (3, 73). For further examination of the conclusions, we also carried out an overall multiple logistic regression with OD diagnosis as the outcome, controlling for age at ascertainment, gender and ethnicity, but without propensity score matching (i.e., including data from all volunteers in the cohort). In this overall regression, the same KMSK scores were detected as positive predictors, as in the regression in the propensity score matched groups (not shown). Therefore, the results reported above with respect to cannabis, cocaine and alcohol KMSK scores as dimensional predictors of OD diagnosis are not likely to be an artifact of the propensity score matching procedure.

CONCLUSIONS

We detected that increasing self-exposure to non-medical cannabis was a positive predictor of odds of an OD diagnosis. We also determined that the level of maximal alcohol selfexposure per se was not a predictor of the OD diagnosis outcome. This is one of the few individual-level examinations in which self-exposure to cannabis and alcohol are both examined dimensionally, as predictors of a diagnosed opioid use disorder. Some recent state-wide and epidemiological studies have reported that the legalized status of medical cannabis is associated with decreases in population-wide opioid overdoses and other measures of opioid-related morbidity (14, 15). Other studies have reported divergent findings on the influence of cannabis use on treatment outcomes in opioiddependent volunteers, possibly due to different methods used (48, 74). Non-medical cannabis use has been associated with increased probability of aberrant opioid-taking behaviors in pain patients (75), and it has also been reported recently that some persons substitute cannabis for other substances, including prescription opioids for non-medical use (76). The

REFERENCES

- CDC. Injury Prevention & Control: Prescription Drug Overdose: Centers for Disease Control and Prevention (2015). Available online at: http://www.cdc. gov/DrugOverdose/
- Han B, Compton WM, Jones CM, Cai R. Nonmedical prescription opioid use and use disorders among adults aged 18 through 64 years in the united states, 2003-2013. JAMA (2015) 314:1468–78. doi: 10.1001/jama.2015.11859
- Blanco C, Hasin DS, Wall MM, Florez-Salamanca L, Hoertel N, Wang S, et al. Cannabis use and risk of psychiatric disorders: prospective evidence from a US National Longitudinal Study. *JAMA Psychiatry* (2016) 73:388–95. doi: 10.1001/jamapsychiatry.2015.3229
- Lopez-Quintero C, Anthony JC. Drug use disorders in the polydrug context: new epidemiological evidence from a foodborne outbreak approach. Ann NY Acad Sci. (2015) 1349:119–26. doi: 10.1111/nyas.12868
- Leeman RF, Sun Q, Bogart D, Beseler CL, Sofuoglu M. Comparisons of cocaine-only, opioid-only, and users of both substances in the National Epidemiologic Survey on Alcohol and Related Conditions (NESARC). Subst Use Misuse (2016) 51:553–64. doi: 10.3109/10826084.2015. 1122063
- Woodcock EA, Lundahl LH, Stoltman JJ, Greenwald MK. Progression to regular heroin use: examination of patterns, predictors, and consequences. *Addict Behav.* (2015) 45:287–93. doi: 10.1016/j.addbeh.2015.02.014
- Shmulewitz D, Greene ER, Hasin D. Commonalities and differences across substance use disorders: phenomenological and epidemiological aspects. *Alcohol Clin Exp Res.* (2015) 39:1878–900. doi: 10.1111/acer.12838
- Hasin DS, Kerridge BT, Saha TD, Huang B, Pickering R, Smith SM, et al. Prevalence and correlates of DSM-5 cannabis use disorder, 2012-2013: findings from the national epidemiologic survey on alcohol and related conditions-III. Am J Psychiatry (2016) 173:588–99. doi: 10.1176/appi.ajp.2015.15070907
- Arterberry BJ, Horbal SR, Buu A, Lin HC. The effects of alcohol, cannabis, and cigarette use on the initiation, reinitiation and persistence of non-medical use of opioids, sedatives, and tranquilizers in adults. *Drug Alcohol Depend.* (2016) 159:86–92. doi: 10.1016/j.drugalcdep.2015.11.029

use of categorical vs. dimensional measures of drug use has also been suggested as a possible reason for the apparent discrepancies in this area (48). Future studies could determine whether increasing non-medical cannabis exposure, especially in adolescence and young adulthood, can result in neurobehavioral changes that underlie greater vulnerability to opioid use disorders.

AUTHOR CONTRIBUTIONS

EB, AM, SB, CC, and MK were involved in data collection and organization. All authors were involved in study conceptualization. EB, CC, and JC were involved in data analysis and presentation.

ACKNOWLEDGMENTS

Ascertainment by members of the clinical staff of the Laboratory is gratefully acknowledged. Data collation by Ms. Maya Darst-Campbell is also gratefully acknowledged, as is editing and writing advice from Dr. Orna Levran. Funding by the Dr. Miriam and Sheldon G. Adelson Medical Research Foundation, from the NIH-NIDA, and from the NIH-CTSA (Grant UL1TR001866) to the Rockefeller University Hospital, is gratefully acknowledged.

- Olfson M, Wall MM, Liu SM, Blanco C. Cannabis use and risk of prescription opioid use disorder in the United States. *Am J Psychiatry* (2018) 175:47–53. doi: 10.1176/appi.ajp.2017.17040413
- Kwako LE, Bickel WK, Goldman D. Addiction biomarkers: dimensional approaches to understanding addiction. *Trends Mol Med.* (2018) 24:121–8. doi: 10.1016/j.molmed.2017.12.007
- Sloan ME, Gowin JL, Yan J, Schwandt ML, Spagnolo PA, Sun H, et al. Severity of alcohol dependence is associated with the fatty acid amide hydrolase Pro129Thr missense variant. *Addict Biol.* (2017) 23:474–84. doi: 10.1111/adb.12491
- Shmulewitz D, Keyes K, Beseler C, Aharonovich E, Aivadyan C, Spivak B, et al. The dimensionality of alcohol use disorders: results from Israel. *Drug Alcohol Depend.* (2010) 111:146–54. doi: 10.1016/j.drugalcdep.2010.04.002
- Bachhuber MA, Saloner B, Cunningham CO, Barry CL. Medical cannabis laws and opioid analgesic overdose mortality in the United States, 1999-2010. *JAMA Intern Med.* (2014) 174:1668–73. doi: 10.1001/jamainternmed.2014.4005
- Kim JH, Santaella-Tenorio J, Mauro C, Wrobel J, Cerda M, Keyes KM, et al. State medical marijuana laws and the prevalence of opioids detected among fatally injured drivers. *Am J Public Health* (2016) 106:2032–7. doi: 10.2105/ajph.2016.303426
- Johnstone RE, Lief PL, Kulp RA, Smith TC. Combination of Δ9tetrahydrocannabinol with oxymorphone or pentobarbital: Effects on ventilatory control and cardiovascular dynamics. *Anesthesiology* (1975) 42:674–84.
- Cooper ZD, Bedi G, Ramesh D, Balter R, Comer SD, Haney M. Impact of coadministration of oxycodone and smoked cannabis on analgesia and abuse liability. *Neuropsychopharmacology* (2018). doi: 10.1038/s41386-018-0011-2
- Bisaga A, Sullivan MA, Glass A, Mishlen K, Pavlicova M, Haney M, et al. The effects of dronabinol during detoxification and the initiation of treatment with extended release naltrexone. *Drug Alcohol Depend.* (2015) 154:38–45. doi: 10.1016/j.drugalcdep.2015.05.013
- Epstein DH, Preston KL. No evidence for reduction of opioid-withdrawal symptoms by cannabis smoking during a methadone dose taper. *Am J Addict.* (2015) 24:323–8. doi: 10.1111/ajad.12183

- Ellgren M, Artmann A, Tkalych O, Gupta A, Hansen HS, Hansen SH, et al. Dynamic changes of the endogenous cannabinoid and opioid mesocorticolimbic systems during adolescence: THC effects. *Eur Neuropsychopharmacol.* (2008) 18:826–34. doi: 10.1016/j.euroneuro.2008.06.009
- Stopponi S, Soverchia L, Ubaldi M, Cippitelli A, Serpelloni G, Ciccocioppo R. Chronic THC during adolescence increases the vulnerability to stress-induced relapse to heroin seeking in adult rats. *Eur Neuropsychopharmacol.* (2014) 24:1037–45. doi: 10.1016/j.euroneuro.2013.12.012
- Cadoni C, Simola N, Espa E, Fenu S, Di Chiara G. Strain dependence of adolescent Cannabis influence on heroin reward and mesolimbic dopamine transmission in adult Lewis and Fischer 344 rats. *Addict Biol.* (2015) 20:132– 42. doi: 10.1111/adb.12085
- Mori T, Shibasaki M, Abe M, Udagawa Y, Suzuki T. Effects of dronabinol on morphine-induced dopamine-related behavioral effects in animals. *Synapse* (2012) 66:931–7. doi: 10.1002/syn.21586
- Solinas M, Panlilio LV, Tanda G, Makriyannis A, Matthews SA, Goldberg SR. Cannabinoid agonists but not inhibitors of endogenous cannabinoid transport or metabolism enhance the reinforcing efficacy of heroin in rats. *Neuropsychopharmacology* (2005) 30:2046–57. doi: 10.1038/sj.npp.1300754
- 25. Le Merrer J, Befort K, Gardon O, Filliol D, Darcq E, Dembele D, et al. Protracted abstinence from distinct drugs of abuse shows regulation of a common gene network. *Addict Biol.* (2012) 17:1–12. doi: 10.1111/j.1369-1600.2011.00365.x
- Kandel DB, Yamaguchi K, Chen K. Stages of progression in drug involvement from adolescence to adulthood: further evidence for the gateway theory. J Stud Alcohol (1992) 53:447–57.
- Vanyukov MM, Tarter RE, Kirillova GP, Kirisci L, Reynolds MD, Kreek MJ, et al. Common liability to addiction and "gateway hypothesis": theoretical, empirical and evolutionary perspective. *Drug Alcohol Depend.* (2012) 123(Suppl. 1):S3–17. doi: 10.1016/j.drugalcdep.2011.12.018
- Secades-Villa R, Garcia-Rodriguez O, Jin CJ, Wang S, Blanco C. Probability and predictors of the cannabis gateway effect: a national study. *Int J Drug Policy* (2015) 26:135–42. doi: 10.1016/j.drugpo.2014.07.011
- Silins E, Horwood LJ, Patton GC, Fergusson DM, Olsson CA, Hutchinson DM, et al. Young adult sequelae of adolescent cannabis use: an integrative analysis. *Lancet Psychiatry* (2014) 1:286–93. doi: 10.1016/s2215-0366(14)70307-4
- Hall W, West R, Marsden J, Humphreys K, Neale J, Petry N. It is premature to expand access to medicinal cannabis in hopes of solving the US opioid crisis. *Addiction* (2018) 113:987–8. doi: 10.1111/add.14139
- Carliner H, Brown QL, Sarvet AL, Hasin DS. Cannabis use, attitudes, and legal status in the U.S.: a review. *Prev Med.* (2017) 104:13–23. doi: 10.1016/j.ypmed.2017.07.008
- Finney JW, Humphreys K, Harris AS. What ecologic analyses cannot tell us about medical marijuana legalization and opioid pain medication mortality. *JAMA Intern Med.* (2015) 175:655–6. doi: 10.1001/jamainternmed.2014.8006
- Wilkinson ST, Yarnell S, Radhakrishnan R, Ball SA, D'Souza DC. Marijuana Legalization: impact on physicians and public health. *Annu Rev Med.* (2016) 67:453–66. doi: 10.1146/annurev-med-050214-013454
- Volkow ND, Baler RD, Compton WM, Weiss SR. Adverse health effects of marijuana use. N Engl J Med. (2014) 370:2219–27. doi: 10.1056/NEJMra1402309
- Sartor CE, Kranzler HR, Gelernter J. Rate of progression from first use to dependence on cocaine or opioids: a cross-substance examination of associated demographic, psychiatric, and childhood risk factors. *Addict Behav.* (2014) 39:473–9. doi: 10.1016/j.addbeh.2013.10.021
- Beseler CL, Hasin DS. Cannabis dimensionality: dependence, abuse and consumption. *Addict Behav.* (2010) 35:961–9. doi: 10.1016/j.addbeh.2010.06.011
- 37. Hill KP. Cannabis use and risk for substance use disorders and mood or anxiety disorders. *JAMA* (2017) 317:1070–1. doi: 10.1001/jama.2016.19706
- Kellogg SH, McHugh PF, Bell K, Schluger JH, Schluger RP, LaForge KS, et al. The Kreek-McHugh-Schluger-Kellogg scale: a new, rapid method for quantifying substance abuse and its possible applications. *Drug Alcohol Depend*. (2003) 69:137–50. doi: 10.1016/S0376-8716(02)00308-3
- Tang YL, Khoury L, Bradley B, Gillespie CF, Ressler KJ, Cubells JF. Substance use disorders assessed using the Kreek-McHugh-Schluger-Kellogg

(KMSK) scale in an urban low-income and predominantly African American sample of primary care patients. *Am J Addict.* (2011) 20:292-9. doi: 10.1111/j.1521-0391.2011.00121.x

- 40. Bond C, LaForge KS, Tian M, Melia D, Zhang S, Borg L, et al. Single-nucleotide polymorphism in the human mu opioid receptor gene alters beta-endorphin binding and activity: possible implications for opiate addiction. *Proc Natl Acad Sci USA*. (1998) 95:9608–13.
- 41. Levran O, Peles E, Randesi M, Correa da Rosa J, Ott J, Rotrosen J, et al. Synaptic plasticity and signal transduction gene polymorphisms and vulnerability to drug addictions in populations of European or African Ancestry. CNS Neurosci Ther. (2015). doi: 10.1111/cns.12450.
- Levran O, Peles E, Randesi M, Correa da Rosa J, Ott J, Rotrosen J, et al. Susceptibility loci for heroin and cocaine addiction in the serotonergic and adrenergic pathways in populations of different ancestry. *Pharmacogenomics* (2015) 16:1329–42. doi: 10.2217/pgs.15.86.
- First MB, Spitzer RL, Gibbon M, Williams JBW. Structured Clinical Interview for DSM-IV-TR Axis I Disorders, Research Version, Patient Edition (SCID-I/P; version 2.0). Biometrics Research Department; New York State Psychiatric Institute (2002).
- Jackson CB, Varon J, Ho A, Marks KM, Talal AH, Kreek MJ. Identification of substance use and dependence among patients with viral hepatitis. *Dig Liver Dis.* (2010) 42:650–6. doi: 10.1016/j.dld.2010.03.002
- 45. Vassileva J, Ahn WY, Weber KM, Busemeyer JR, Stout JC, Gonzalez R, et al. Computational modeling reveals distinct effects of HIV and history of drug use on decision-making processes in women. *PLoS ONE* (2013) 8:e68962. doi: 10.1371/journal.pone.0068962
- Khoury L, Tang YL, Bradley B, Cubells JF, Ressler KJ. Substance use, childhood traumatic experience, and posttraumatic stress disorder in an urban civilian population. *Depress Anxiety* (2010) 27:1077-86. doi: 10.1002/da.20751.
- Austin PC. An introduction to propensity score methods for reducing the effects of confounding in observational studies. *Mult Behav Res.* (2011) 46:399–424. doi: 10.1080/00273171.2011.568786
- Zielinski L, Bhatt M, Sanger N, Plater C, Worster A, Varenbut M, et al. Association between cannabis use and methadone maintenance treatment outcomes: an investigation into sex differences. *Biol Sex Differ*. (2017) 8:8. doi: 10.1186/s13293-017-0130-1
- 49. CDC. Understanding the Epidemic (2016). Available online at: https://www.cdc.gov/drugoverdose/epidemic/index.html
- 50. Grant BF, Chou S, Saha TD, Pickering RP, Kerridge BT, Ruan WJ, et al. Prevalence of 12-month alcohol use, high-risk drinking, and DSM-IV alcohol use disorder in the united states, 2001-2002 to 2012-2013: results from the national epidemiologic survey on alcohol and related conditions. *JAMA Psychiatry* (2017) 74:911–23. doi: 10.1001/jamapsychiatry.2017.2161
- Hernandez-Avila CA, Rounsaville BJ, Kranzler HR. Opioid-, cannabis- and alcohol-dependent women show more rapid progression to substance abuse treatment. *Drug Alcohol Depend.* (2004) 74:265–72
- Chen CY, Storr CL, Anthony JC. Early-onset drug use and risk for drug dependence problems. *Addict Behav.* (2009) 34:319–22. doi: 10.1016/j.addbeh.2008.10.021
- Cleveland HH, Wiebe RP. Understanding the association between adolescent marijuana use and later serious drug use: gateway effect or developmental trajectory? *Dev Psychopathol.* (2008) 20:615–32. doi: 10.1017/s0954579408000308
- Kerridge BT, Pickering R, Chou P, Saha TD, Hasin DS. DSM-5 cannabis use disorder in the National Epidemiologic Survey on Alcohol and Related Conditions-III: gender-specific profiles. *Addict Behav.* (2017) 76:52–60. doi: 10.1016/j.addbeh.2017.07.012
- Hasin DS, Grant BF. The National Epidemiologic Survey on Alcohol and Related Conditions (NESARC) Waves 1 and 2: review and summary of findings. Soc Psychiatry Psychiatr Epidemiol. (2015) 50:1609–40. doi: 10.1007/s00127-015-1088-0
- Mayet A, Legleye S, Falissard B, Chau N. Cannabis use stages as predictors of subsequent initiation with other illicit drugs among French adolescents: Use of a multi-state model. *Addict Behav.* (2012) 37:160–6. doi: 10.1016/j.addbeh.2011.09.012
- 57. Fergusson DM, Horwood LJ. Does cannabis use encourage other forms of illicit drug use? *Addiction* (2000) 95:505–20. doi: 10.1046/j.1360-0443.2000.9545053.x

- Bloomfield MA, Ashok AH, Volkow ND, Howes OD. The effects of Delta9tetrahydrocannabinol on the dopamine system. *Nature* (2016) 539:369–77. doi: 10.1038/nature20153
- Gorelick DA, Kim YK, Bencherif B, Boyd SJ, Nelson R, Copersino M, et al. Imaging brain mu-opioid receptors in abstinent cocaine users: time course and relation to cocaine craving. *Biol Psychiatry* (2005) 57:1573–82. doi: 10.1016/j.biopsych.2005.02.026
- Di Chiara G, Imperato A. Drugs abused by humans preferentially increase synaptic dopamine concentrations in the mesolimbic system of freely moving rats. *Proc Natl Acad Sci USA*. (1988) 85:5274–8.
- Ghozland S, Matthes HW, Simonin F, Filliol D, Kieffer BL, Maldonado R. Motivational effects of cannabinoids are mediated by muopioid and kappa-opioid receptors. *J Neurosci.* (2002) 22:1146-54. doi: 10.1523/JNEUROSCI.22-03-01146.2002
- Panlilio LV, Justinova Z. preclinical studies of cannabinoid reward, treatments for cannabis use disorder, and addiction-related effects of cannabinoid exposure. *Neuropsychopharmacology* (2017) 43:116–41. doi: 10.1038/npp.2017.193
- Kruzich PJ, Chen AC, Unterwald EM, Kreek MJ. Subject-regulated dosing alters morphine self-administration behavior and morphine-stimulated [35S]GTPgammaS binding. Synapse (2003) 47:243–9. doi: 10.1002/syn.10173
- Unterwald EM, Kreek MJ, Cuntapay M. The frequency of cocaine administration impacts cocaine-induced receptor alterations. *Brain Res.* (2001) 900:103–9. doi: 10.1016/S0006-8993(01)02269-7
- Szutorisz H, Hurd YL. High times for cannabis: epigenetic imprint and its legacy on brain and behavior. *Neurosci Biobehav Rev.* (2017) 85:93–101. doi: 10.1016/j.neubiorev.2017.05.011
- Birrell L, Newton NC, Teesson M, Slade T. Early onset mood disorders and first alcohol use in the general population. J Affect Disord. (2016) 200:243–9. doi: 10.1016/j.jad.2016.04.032
- Paulozzi LJ. Prescription drug overdoses: a review. J Safety Res. (2012) 43:283– 9. doi: 10.1016/j.jsr.2012.08.009
- Banerjee G, Edelman EJ, Barry DT, Becker WC, Cerda M, Crystal S, et al. Non-medical use of prescription opioids is associated with heroin initiation among US veterans: a prospective cohort study. *Addiction* (2016) 111:2021– 31. doi: 10.1111/add.13491
- 69. Wu LT, Woody GE, Yang C, Mannelli P, Blazer DG. Differences in onset and abuse/dependence episodes between prescription opioids and heroin: results

from the National Epidemiologic Survey on Alcohol and Related Conditions. Subst Abuse Rehabil. (2011) 2011:77–88. doi: 10.2147/sar.s18969

- Borg L, Broe DM, Ho A, Kreek MJ. Cocaine abuse sharply reduced in an effective methadone maintenance program. J Addict Dis. (1999) 18:63–75. doi: 10.1300/J069v18n04_06
- Bandettini Di Poggio A, Fornai F, Paparelli A, Pacini M, Perugi G, Maremmani I. Comparison between heroin and heroin-cocaine polyabusers: a psychopathological study. *Ann N Y Acad Sci.* (2006) 1074:438–45. doi: 10.1196/annals.1369.044
- Cuthbert BN. The RDoC framework: facilitating transition from ICD/DSM to dimensional approaches that integrate neuroscience and psychopathology. *World Psychiatry* (2014) 13:28–35. doi: 10.1002/wps. 20087
- Okoli GN, Sanders RD, Myles P. Demystifying propensity scores. BJA: Br J Anaest. (2014) 112:13–5. doi: 10.1093/bja/aet290
- Hill KP, Bennett HE, Griffin ML, Connery HS, Fitzmaurice GM, Subramaniam G, et al. Association of cannabis use with opioid outcomes among opioid-dependent youth. *Drug Alcohol Depend* (2013) 132:342–5. doi: 10.1016/j.drugalcdep.2013.02.030
- 75. Reisfield GM, Wasan AD, Jamison RN. The prevalence and significance of cannabis use in patients prescribed chronic opioid therapy: a review of the extant literature. *Pain Med.* (2009) 10:1434–41. doi: 10.1111/j.1526-4637.2009.00726.x
- Corroon JM Jr, Mischley LK, Sexton M. Cannabis as a substitute for prescription drugs - a cross-sectional study. J Pain Res. (2017) 10:989–98. doi: 10.2147/jpr.s134330

Conflict of Interest Statement: The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

Copyright © 2018 Butelman, Maremmani, Bacciardi, Chen, Correa da Rosa and Kreek. This is an open-access article distributed under the terms of the Creative Commons Attribution License (CC BY). The use, distribution or reproduction in other forums is permitted, provided the original author(s) and the copyright owner are credited and that the original publication in this journal is cited, in accordance with accepted academic practice. No use, distribution or reproduction is permitted which does not comply with these terms.