

Patent foramen ovale closure or medical therapy for secondary prevention of cryptogenic stroke An update meta-analysis of randomized controlled trials

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Abstract

Background: It was under debate whether cryptogenic stroke patients benefited from patent foramen ovale (PFO) closure. We sought to determine secondary prevention strategy in these patients.

Methods: Scientific databases were searched for randomized controlled trials enrolling cryptogenic stroke patients with PFO who underwent PFO closure or medical therapy. The random-effect model was used to analyze the outcomes.

Results: We identified 6 trials enrolling 3630 participants in this meta-analysis. When compared with medical therapy, PFO closure reduced risks of recurrent stroke (risk ratio [RR] 0.52, 95% confidence interval [CI] 0.29–0.93) and composite of stroke and transient ischemic attack (TIA) (RR 0.60, 95% CI 0.46–0.80). And no differences in all-cause death (RR 0.80, 95% CI 0.37–1.72) and cardiovascular death (RR 1.47, 95% CI 0.36–5.94) between 2 groups were observed. The risks of major bleeding (RR 0.96, 95% CI 0.47–1.96) and any serious adverse event (RR 1.03, 95% CI 0.92–1.16) did not differ between 2 groups. Yet, PFO closure increased risk of atrial fibrillation (RR 4.25, 95% CI 2.10–8.60).

Conclusion: PFO closure, as compared with medical therapy, was associated with decreased risk of recurrent stroke and increased risk of atrial fibrillation in cryptogenic stroke patients with PFO.

Abbreviations: CIs = confidence intervals, PFO = patent foramen ovale, RRs = risk ratios, TIA = transient ischemic attack.

Keywords: cryptogenic stroke, device closure, medical therapy, patent foramen ovale

1. Introduction

Stroke was the third-leading death cause among adults and the main factor of long-term functional impairment and disability.^[1] Approximately 25% to 40% of ischemic strokes had no identifiable cause, which were classified as cryptogenic strokes.^[2,3] A part of cryptogenic strokes might be the results of paradoxical embolism which arose from venous embolus reaching systemic circulation due to right-to-left shunt in congenital heart disease patients. Patent foramen ovale (PFO), observed in 14.9% to 27% of populations,^[4,5] provided anatomic substrate for paradoxical embolism. Numerous observational studies indicated that cryptogenic stroke was associated with PFO^[6,7] and this relationship was more

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Received: 18 May 2018 / Accepted: 27 July 2018 http://dx.doi.org/10.1097/MD.000000000011965 convincingly demonstrated in young patients.^[8] This implied PFO closure might be beneficial to secondary prevention of cryptogenic stroke. But not all studies were in favor of this association. The Northern Manhattan Study^[5] showed PFO was not associated with the increased stroke risk in males and females. In addition, PFO was not the significant and independent predictor of stroke among normal subjects older than 45 years.^[9]

Therefore, physicians have shown great interest in the role of PFO closure in secondary prevention of cryptogenic stroke. Observational studies^[10,11] and meta-analysis of observational studies^[12] demonstrated PFO closure significantly lowered stroke recurrence rate when compared with medical therapy. However, 3 randomized controlled trials^[13-15] which were published in 2012 and 2013 indicated PFO closure was not superior to medical therapy in cryptogenic stroke patients. But there were some limitations in these trials. In CLOSURE I trial,^[13] the number of enrolled patients was lower than expected, and these patients were treated with off-label device. Meanwhile, periprocedural complication rate was relatively high. And the magnitude of effect estimates of RESPECT trial^[14] and PC trial^[15] was low. These drawbacks limited physicians' interpretation of the results into clinical practice. Two meta-analyses^[16,17] of these 3 trials suggested PFO closure was more appropriate than medical therapy in secondary prevention of cryptogenic stroke. But the results of these 2 meta-analyses were confused because results of enrolled trials were negative while the meta-analysis demonstrated positive result. Thus, 2014 AHA/ ASA guideline^[18] and 2016 AAN guideline^[19] recommended existing evidence was not enough to prove that cryptogenic stroke patients could benefit from PFO closure. Recently, CLOSE trial,^[20] Gore REDUCE trial,^[21] long-term result of RESPECT trial,^[22] and DEFENSE-PFO trial^[23] have been published,

YM and DL are first co-authors and contributed equally to this work.

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suggesting stroke recurrence rate of patients undergoing device closure was lower than that of participants accepting medical treatment.

Therefore, the secondary prevention strategy in cryptogenic stroke patients with PFO was still controversial. The aim of this meta-analysis was to compare outcomes of PFO closure and medical therapy in these patients, offering physicians a more comprehensive picture of management strategies in these patients.

2. Methods

2.1. Search strategy and eligibility criteria

Two investigators (YM and DL) independently searched PubMed, Embase, and Cochrane Library for randomized controlled trials enrolling cryptogenic stroke patients with PFO undergoing device closure or medical therapy as a management strategy, which were published before May 18, 2017 and restricted to English. Notably, medical therapy included antiplatelet therapy, anticoagulation therapy, or both.^[18,19] Our search strategy in PubMed incorporated the Cochrane Highly Sensitive Search Strategy for identification of randomized controlled trials.^[24] The main search terms were "patent foramen ovale," "cryptogenic stroke," "device closure," and "medical therapy." The details of the search strategies were listed in the Supplementary Materials, http://links.lww.com/MD/ C412. Two authors (YM and DL), respectively, performed the screening of titles and abstracts, reviewed full texts of articles if needed, and determined their eligibility. We also searched the reference lists of the original articles identified for full text review to find other eligible studies. Divergences were resolved by discussion.

2.2. Data extraction and quality assessment

Two independent reviewers (YM and FQ) extracted data from included studies. Data extracted from studies included study characteristics, patient characteristics, details regarding closure and medical therapy groups, and outcome measures. Outcomes of interest for this meta-analysis were classified as procedurerelated or device-related endpoints, efficacy endpoints, and safety endpoints. Procedure-related or device-related endpoints included implantation success rate, effective closure rate, and procedure-related or device-related complication rate. Efficacy endpoints included recurrent stroke, transient ischemic attack (TIA), a composite of stroke and TIA, all-cause death, and cardiovascular death. Safety endpoints included atrial fibrillation, major bleeding, and any serious adverse event. Specifically, major bleeding was defined as a reduction in the hemoglobin level of at least 20 g/L, transfusion of at least 2 units of packed red cells or bleeding occurred at a critical site.^[25] And any serious adverse event was defined as the adverse event that resulted in permanent impairment of a body function or permanent damage to a body structure, prolongation of hospitalization, life-threatening events, or death.^[13–15,20–22]

Quality of studies was assessed through the Cochrane Collaboration Tool. More specific, sequence generation for randomization, allocation concealment, masking of outcome assessors, incomplete outcome data, selective outcome reporting, and other sources of bias were evaluated in detail.

2.3. Statistical analysis

Intention-to-treat meta-analysis was performed by using Review Manager 5.3 and Stata 12.0. We did not perform per-protocol and as-treated meta-analyses. Because some trials^[15,20,22] did not

incorporate the data of per-protocol and as-treated analyses. Many investigators^[24] and we thought bias would occur if we used data of intention-to-treat analyses of these trials in per-protocol and as-treated meta-analyses like some authors did.^[16]

The categorical variables were reported as the counts and percentages and continuous variables were presented as mean and standard deviation or median and interquartile range, as appropriate. Outcome data were extracted as risk ratios (RRs) and 95% confidence intervals (CIs) for device closure versus medical therapy among cryptogenic stroke patients with PFO. We reported unadjusted RRs since adjusted RR was presented in only 1 trial^[13] and 4 trials^[14,15,20–23] did not report the adjusted RRs. RRs for each outcome were calculated using the randomeffect model.^[24] Given the heterogeneity in the study design and included populations and variability in the definition of medical therapy, a random-effect model rather than a fixed-effect model was considered more appropriate. The Cochran Q test and I^2 test were performed to assess the heterogeneity of the summary effects. If the *P*-value of Cochran O test was <.10 and I^2 was >50%, heterogeneity was considered to exist. Publication bias and bias associated with small study effects were assessed with funnel plot, Begg test and Egger test, respectively. Whether effect sizes were distributed symmetrically or not was judged visually. In addition, if *P*-values of Begg test and Egger test were <.10, and publication bias was considered to exist.

2.4. Sensitivity analysis

Jackknife sensitivity analyses were performed for each outcome of interest to verify the robustness of the results and the impact of each single study on the summary estimate of the effect. Pooled estimates were recalculated multiple times by using a randomeffect model, each time with removal of a single study from the baseline group.

3. Results

3.1. Study selection

The results of study selection process were shown in Supplementary Figure 1, http://links.lww.com/MD/C412. After the screening, we identified 7 articles that fulfilled the inclusion criteria. The short-term result of RESPECT trial was reported in 2013^[14] and long-term result was shown in 2017.^[22] Kaplan–Meier curve of this study published in 2013^[14] continued to diverge, suggesting that postulated benefit might need more time to become obvious. Moreover, investigators were able to make a more thorough evaluation of effects of intervention on patients with extension of follow-up duration.^[26] Therefore, we excluded the article published^[14] in 2013. In the end, we enrolled 6 randomized controlled trials^[13,15,20–23] in this meta-analysis.

3.2. Baseline characteristics

Baseline characteristics of the included patients were shown in Supplementary Table 1, http://links.lww.com/MD/C412. A total of 3560 participants were enrolled in 6 trials and the range of mean duration of follow-up was 2 to 5.9 years. The patients of closure groups underwent antiplatelet or anticoagulation therapy after the device implantation, including aspirin, clopidogrel, and warfarin. But dose and treatment duration were different in each study. The drug regimens of medical therapy groups were different, too. But all the drug regimens were in accord with the guideline recommendations.^[18,19] Enrolled participants were



predominantly men and with stroke risk factors of hypertension, diabetes, hyperlipemia, and smoking. Within each trial, baseline characteristics were similar between the closure and medical therapy groups.

3.3. Procedure-related or device-related endpoints

The procedure success rate was 97% (95% CI 95.3–99.2%) in patients undergoing device implantation. After 6 to 12 months of follow-up, effective closure was documented in 85.9% of patients (95% CI 74.5–97.3%). During the whole follow-up period, procedure-related or device-related complication rate was 3.7% (95% CI 2.3–5.1%).

3.4. Efficacy endpoints

Compared with medical therapy, closure reduced risk of recurrent stroke by 48% (RR 0.52, 95% CI 0.29–0.93) (Fig. 1). There was no difference in TIA between closure and medical therapy groups (RR 0.76, 95% CI 0.52–1.12) (Fig. 2). In addition, closure was associated with a 40% lower risk of composite of stroke and TIA (RR 0.60, 95% CI 0.46–0.80) relative to medical therapy (Fig. 3). However, there were no significant differences in all-cause death (RR 0.80, 95% CI 0.37–1.72) (Supplementary Figure 2, http://links.lww.com/MD/C412)

and cardiovascular death (RR 1.47, 95% CI 0.36–5.94) (Supplementary Figure 3, http://links.lww.com/MD/C412) between closure and medical therapy groups.

3.5. Safety endpoints

As compared with medical therapy, closure significantly increased risk of atrial fibrillation (RR 4.25, 95% CI 2.10–8.60) (Table 1 and Supplementary Figure 4, http://links.lww.com/MD/C412). However, no significant differences were observed in major bleeding (RR 0.96, 95% CI 0.47–1.96) (Table 1 and Supplementary Figure 5, http://links.lww.com/MD/C412) and any serious adverse event (RR 1.03, 95% CI 0.92–1.16) (Table 1 and Supplementary Figure 6, http://links.lww.com/MD/C412) between closure and medical therapy groups.

3.6. Sensitivity analyses

Jackknife sensitivity analyses (Supplementary Table 2, http:// links.lww.com/MD/C412) demonstrated 3 studies significantly affected pooled RRs of recurrent stroke. After removal of PC trial,^[15] Gore REDUCE trial,^[21] or RESPECT trial,^[22] there was no difference in recurrent stroke between closure and medical therapy groups.



Figure 2. Forest plot with individual and summary estimates of the risk ratio (RR) and 95% confidence interval (CI) of transient ischemic attack (TIA). CI = confidence interval.



Figure 3. Forest plot with individual and summary estimates of the risk ratio (RR) and 95% confidence interval (CI) of composite of stroke and transient ischemic attack (TIA).

3.7. Risk of bias

All the trials included in this meta-analysis were randomized, open-label trials. Method of randomization was adequately described (computer generated or automated telephone system) in all trials and allocation concealment was promised. All the trials were open-label trials which might impact the results, but the extent of impact was unknown. Therefore, risk of bias of all the trials in binding of participants and personnel was unclear. Masking of outcome assessors was described in all trials and losses to follow-up were reported in all trials. All trials were free of selective outcome reporting. The risks of bias across all studies were summarized in Supplementary Figure 7, http://links.lww. com/MD/C412.

Effect sizes reported in studies were distributed symmetrically (Supplementary Figure 8, http://links.lww.com/MD/C412), and there was no significant bias from small studies (Begg test P=.81; Egger test P=.56).

4. Discussion

Our study indicated that procedure success rate was high in the closure group and effective closure could be achieved in most patients with a low procedure-related or device-related complication rate. When compared with medical therapy, PFO closure significantly reduced the risks of recurrent stroke and composite of stroke and TIA, but there was no difference in TIA between 2 groups. And no significant differences in all-cause death and cardiovascular death between closure and medical therapy groups were observed. The risks of major bleeding and any serious adverse event did not differ significantly between study groups. Yet, PFO closure was associated with increased risk of atrial fibrillation.

All the studies included in this meta-analysis were designed to compare the role of PFO closure and medical therapy in secondary prevention of cryptogenic stroke and it is impossible to hide the fact that participants were receiving device implantation or medical therapy. Thus, the design of double-blind trial was hard to carry out. Design of open-label trial might introduce bias to studies, but as shown in Supplementary Figure 7, http://links. lww.com/MD/C412, the risks of bias in other aspects were low. Of note, randomization was conducted in the subgroups in which patients were allocated in the light of contraindications to oral anticoagulant or closure device in the CLOSE trial.^[20] After verifying the absence of interaction between treatment effect and treatment group investigators performed the analyses with combined data of patients who accepted the same treatment in different subgroups. Hence, the design and statistical analyses of CLOSE trial were reliable with relatively low bias. Therefore, all included studies were of high quality and the results of this metaanalysis based on these trials would be helpful to guide the clinical work.

Although medical treatments that patients received differed between studies, these treatments were consisted with recommendations of guidelines^[18,19] and were approved by the steering committees. This was similar with clinic practice because physicians had to adjust the dosage or switch the medication according to the patient's risks of thrombosis or bleeding. So, it is appropriate to combine these studies where types and doses of medications were different.

If effective PFO closure could not be achieved and residual shunts existed, it would be possible that embolus traversed from venous system and into systemic circulation. It might mask the real efficacy of PFO closure in the secondary prevention of cryptogenic stroke. Meta-analysis of proportions showed that the procedure success rate and effective closure rate were 97% and

Table 1

	Summary of data	a comparing closure	with medical therapy in the i	intention-to-treat populations
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	Test for he	Test for heterogeneity		Test for overall effect		
	ŕ	Р	Ζ	Р	RR	95% CI
AF	33%	.20	3.63	.00	4.25	2.10-8.60
Major bleeding	37%	.17	0.12	.91	0.96	0.47-1.96
ASE	13%	.33	0.55	.58	1.03	0.92-1.16

AF = atrial fibrillation, ASAE = any serious adverse event, CI = confidence interval, RR = risk ratio.

85.9%, respectively. Our results were consistent with previous study data,^[27] suggesting that the conditions of closure procedures in the included studies approached to clinic practice and the internal and external validity of this meta-analysis was high.

Our analysis demonstrated that PFO closure was associated with significant improvement in recurrent stroke and composite of stroke and TIA, which was similar with results of some previous observational studies,^[10,11,28] randomized controlled trials^[21,22] and meta-analyses.^[12,16] However, results of some studies were contradictory to ours. Windecker et al^[10] reported that PFO closure could not prevent stroke recurrence (RR 0.36, 95% CI 0.08-1.74). This might be because patients treated with PFO closure, as compared with medical therapy group, had suffered from more than one stroke event before they were enrolled in this study. And they were more prone to stroke. Investigators of CLOSURE I^[13] did not observe sufficient evidence to support the use of closure device in patients with PFO for secondary prevention of cryptogenic stroke. This could be explained by the utilization of the STARFlex device. Thrombi on the device surface were observed in approximately 3.6% of patients of the STARFlex group though they accepted adequate antiplatelet treatment.^[29] But no thrombus was observed in participants undergoing implantation of other types of device.^[29] It was possible for patients in STARFlex group to suffer from stroke when thrombus detached from device surface, which might dilute the actual beneficial effect of device closure. The short-term results of the RESPECT trial^[14] did not show beneficial effect of closure for stroke prevention, but the rate of stroke in device closure versus medical therapy appeared to separate after follow-up of 2.1 years and continued to diverge. And the long-term results of RESPECT trial^[22] confirmed this trend, demonstrating that PFO closure was effective for secondary prevention of cryptogenic stroke.

Atrial fibrillation was common in cryptogenic stroke patients.^[30] Most included studies^[13,20–23] showed the rates of atrial fibrillation in the closure groups were higher than those in the medical therapy groups and majority of these cases in the closure groups occurred within periprocedural periods, implying catheter was associated with the increased risk of atrial fibrillation. Notably, atrial fibrillation was transient in most patients and could be terminated with or without pharmacologic cardioversion. And few atrial fibrillation patients suffered from stroke in most included studies.^[13,15,20,22] Perhaps, atrial arrhythmias might be transient without adverse effect if treated properly. Furthermore, the rates of atrial fibrillation after the periprocedural period did not differ between 2 groups.^[22] This might be because extra electrocardiograph monitoring contributed to higher diagnosis rates in closure groups during periprocedural periods and the rates of atrial fibrillation in the medical therapy groups were underestimated. Thus, it was confusing for us to evaluate the real effects of atrial fibrillation in the trials.

Different management strategies did not make difference in the occurrence of major bleeding and any serious adverse event. In some ways, the safety of closure was comparable with that of medical therapy. Although the included trials^[13,15,20–23] did not reported the occurrence time of major bleeding events, the observational study^[30] showed most major bleeding events occurred during the period of anticoagulation therapy. This suggested drug was major risk factor of major bleeding and appropriate anticoagulation or antiplatelet therapies with shorten duration might prevent major bleeding events. There

was no difference in adverse event rate, but great variability was observed in types of adverse events between groups. Procedurerelated and device-related complications were observed only in closure group and STARFlex device significantly increased risk of adverse events.^[20] Thus, it is helpful in the prevention of this kind of adverse events to choose safer device and improve the implantation procedures. Most adverse events in medical therapy groups were related to the medicine; hence, rational use of the medicine might reduce the incidence of adverse events.

4.1. Study limitation

Our study had some limitations inherent to the included studies and to meta-analysis. Firstly, the trials differed meaningfully in study devices, medical regimens, and durations of follow-up. However, it can be argued that it extends the generalizability of our findings to a wider range of population. Secondly, this metaanalysis lacked specified individual data to conduct metaregression or subgroup analyses to explore the source of heterogeneity. In addition, even a small change in the rate of event might lead to a dramatic change in the results because of relatively low event rates.

5. Conclusion

The PFO closure, as compared with medical therapy, reduced risks of stroke and the composite of stroke and TIA in patients with PFO after cryptogenic stroke. But PFO closure did not lower the mortality risk. They showed comparable safety in terms of major bleeding and any serious adverse event. However, PFO closure was associated with increased risk of atrial fibrillation.

Author contributions

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