

# Analgesic Efficacy of Multiple Single-Shot Peripheral Nerve Blocks on Postoperative Short-Term Opioid Usage and Clinical Outcomes in a Suburban Hospital Setting

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# Abstract

**Background:** Preoperative single-shot peripheral nerve blocks (sPNBs) represent promising candidates for controlling postoperative pain, reducing dependence on opioid medications, and reducing postoperative constipation and ileus. However, there is not yet complete consensus regarding their efficacy. The primary aim of this study was to assess the impact of various sPNBs on patient short-term opioid demands and pain management parameters.

**Methods:** This single-center study retrospectively reviewed a cohort of 94 adult, elective surgery inpatients (ASA physical status I-III) scheduled for different operations. Sixty-four (68.1%) were selected for sPNB administration (group 1) and compared to the untreated group (group 0) for different clinical parameters.

**Results:** Contrary to the starting hypothesis, a higher proportion of group 1 patients experienced increasing pain intensities during the immediate postoperative period (P < 0.05, Fisher's exact test), while requiring more bowel care medications (P < 0.05,  $\chi^2$  test). Multiple linear regression modeling, however, showed that recovery time positively correlated with the opioid amount consumed (P < 0.01). Although limited, the results obtained in this study do not support an analgesic efficacy for sPNBs.

**Conclusion:** In conclusion, even though our data must be viewed within the limitations of our retrospective study and small group size, we did not find any compelling evidence for the efficacy of sPNB administration in the perioperative period.

Keywords: Single-shot peripheral nerve block; Opioid; Pain management; Postoperative recovery

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# Introduction

Postoperative recovery is a critical issue in healthcare settings worldwide. Pain management is a key component of this recovery. Adequate pain control improves ambulation, decreases hospitalization costs, lowers the risk of chronic post-surgical pain and persistent opioid use, and improves patient satisfaction [1-8]. Many everyday surgical operations labeled "minor" provoke acute postoperative pain [9]. The relevance of pain appraisal is underscored by the fact that roughly 100 million medical treatments are carried out in the United States annually, half of which are surgeries [10-12]. More concerningly, these numbers are projected to inflate in the near future [10, 12, 13], as orthopedic interventions have already increased steeply in the last decades [14]. Assessing pain intensity is moreover necessary because it is positively correlated with increasing numbers of postoperative impairments [8, 15-17].

Despite this, a substantial number of surgical patients experience insufficient perioperative pain relief [17-19]. Polls conducted in the United States (1993, 2003, and 2012) showed that poor postoperative pain control has persisted in severity and diffusion across three decades. In recent studies, 80-86% of patients reported being afflicted by postoperative pain, and 75-88% of this group experienced moderate to extreme pain intensities [2, 20]. Outside the perioperative period, these issues are paralleled by high incidences of persistent postsurgical pain (PPP), which affects 30-50% of patients and thus represents a significant societal toll [8, 21, 22]. The formulation of the current guidelines for pain management contributes to the ongoing undertreatment of postoperative pain. Often, these instructions are deficient in evidence [10], provide recommendations that are too general to be applied on specific cases [4], or lack unanimous consensus by different institutions [23, 24].

Clinical pain control has nonetheless progressed remarkably in the last few decades, leading to the development of new drugs, advancements in techniques, and innovative guidelines and protocols [1, 4, 25-28]. Even if multimodal analgesia remains vastly underutilized and insufficiently studied [29, 30], this strategy is considered the gold standard for ameliorating pain perceived by patients [1, 2, 29, 31, 32], particularly in postoperative settings [2]. This strategy utilizes the application of multiple different medications and procedures whose additive and synergistic interactions result in a more beneficial outcome when compared to the outcomes of monotypical treatments. With multimodal analgesia, the required dose of each drug used is lower, and the risk of adverse effects is minimized [1, 28, 29]. Multimodal therapies are cost-efficient, yielding higher patient satisfaction with lower side effects, and may lower the incidence of chronic postsurgical pain compared to opioids alone [2, 30].

In contrast to other analgesic medications, opioid drugs have a narrow therapeutic window. They have numerous detrimental effects, including phenomena of hyperalgesia, tolerance development, and rehospitalization [2, 29, 33-36]. Furthermore, their over-prescription has spurred an epidemic of opioid dependency in North America [2, 3, 20, 37]. Despite the adverse effects, however, these substances are still routinely administered intraoperatively as potent analgesics acting in the postoperative phase, and continue to represent a cornerstone of postoperative pain management [2, 4, 29, 34-37]. Fortunately, comprehensive clinical recommendations and novel protocols have recently been developed to minimize the risks associated with the injudicious use of opioids [38, 39]. Nevertheless, discarding opioids altogether is currently unfeasible [40], because the efficacy of opioid-free anesthesia (OFA) remains relatively untested and controversial to date [2, 36, 41-43]. Additionally, modern OFA regimens cannot provide analgesia throughout the entire perioperative period [44], and some nonopioid medications pose safety risks of their own [4, 43, 45]. Nonetheless, opioid-related complications have spurred a growing interest in opioid-sparing treatments [29, 43]. These therapies are ideally intended to replace the clinical use of opioids in time [36].

Considering the multimodal anesthesia paradigm, there is a wealth of medications and adjuvants that can be employed for opioid-free pain control [2, 29]. Among these, regional anesthesia driven by nerve blocks (peripheral and neuraxial blocks) appears to be a promising candidate for perioperative pain management [46, 47]. In many surgical specialties, peripheral nerve blocks have been increasingly adopted. However, both peripheral nerve blocks and neuraxial blocks remain reported as highly underutilized [14, 30, 47-49].

Peripheral nerve blocks are widely used on disparate body parts, are less invasive than neuraxial techniques, and have yielded successful results for many interventions in which other practices have been less effective [2, 49-51]. However, their usage by physicians has remained relatively low, with most anesthesiologists reporting performing less than five per month [52, 53]. Single-shot peripheral nerve blocks (sPNBs) involve the administration of a one-time dose of local anesthetic. They allow for pain relief with minimal invasion. The potent local analgesia conveyed by these drugs has a variable duration (< 1 h to > 1 day), depending on the pharmacology of the block used, its concentration, the injection's area, and the patient's response [50, 54, 55].

Although opioid consumption seems to be decreased in most cases, research has not yet established the exact relationship between nerve block administration and opioid consumption for postoperative pain management [49]. Considering an improvement in health care quality, the primary aim of this study was to investigate retrospectively how different sPNBs impact the opioid consumption of patients undergoing different surgical operations. In this study, it is hypothesized that sPNBs will decrease opioid consumption while improving patient satisfaction in the immediate postoperative period.

# **Materials and Methods**

This study has been reviewed by the IRB at Marchand Institute and was found to be exempt from IRB review (November 2021). Data used were exempt from consent to participate or publish secondary to the nature of the study being a retrospective cohort, retrospectively looking at collected data. This study was conducted in compliance with the ethical standards of the responsible institution on human subjects as well as with the Helsinki Declaration. This study did not involve the use of any animals.

This single-center study reviewed retrospectively a cohort of 94 (33 males and 61 females, 35.1% and 64.9% of the total, respectively) adult, elective surgery inpatients (ASA physical status I-III), aged 32 - 91 years of age, scheduled for different operations. All surgical patients during this time period were recorded. The time period was preset secondary to the availability of the staff to volunteer for survey completion and data collection. The population examined did not include intubated, emergent, or intensive care unit (ICU) patients. Of the patients, 96.8% (91 of 94) were operated under general anesthesia. Date and time of discharge from the post-anesthesia care unit (PACU) and hospital were recorded. The discharge date from the PACU ranged from February 25, 2021 to May 5, 2021. The discharge date from the hospital ranged from March 10, 2021 to May 7, 2021. Surveys were taken shortly after discharge from PACU, with dates ranging from February 26, 2021 to May 6, 2021. The "all-comers" approach adopted by our researchers for this study was agreed to approximate the surgical volume and variety of the local community. The variety of surgical operations reported in the present study would have hindered an informative statistical analysis, therefore, the surgeries were grouped into homogeneous categories according to the German procedure classification (OPS) and the scheme used by Gerbershagen et al (2013) [17] to allow for comparisons.

The quantity of opioids taken orally for postoperative analgesia was recorded to assess the quality of pain control. The opioid doses were delivered via routine administration as described in our clinical protocols as well as via patientcontrolled analgesia (PCA) pump. The opioid dosages were converted to an opioid oral morphine milligram equivalent (MME) using a standard opiate agent conversion chart to allow for a comparison of the different drugs administered [56]. All opioid doses administered throughout the 4 days following hospital discharge (postoperative days 0 - 3) were recorded, even those not explicitly requested by patients but dispensed as part of our protocol. No medical adjuncts were used in this study. Sixty-four of 94 patients (68.1%) were selected for the administration of a sPNB, and eight of these (8.5%) received two different blocks. The adductor canal (AC) block was the most administered (46.9%), whereas the second-most delivered block, the erector spinae plane (ESP), involved only 11% of the patients. Subsequently, to further appraise the quality of care provided, the following questionnaire was filled out by each patient: 1) What would you rate your average pain score on a 0 - 10 scale (10 = intense) for the first 12 h after surgery? When did it start? 2) Has your pain score changed at any point in the postoperative phase? Has it increased or decreased? What makes it worse or better? 3) Have you required pain medication since surgery? If so, how often? 4) On a 1 - 10 (10 = extremely satisfied) scale, how would you rate the pain management provided to you after surgery? 5) Did you experience postoperative nausea and/or vomiting (PONV)?

The statistical analyses were performed with the R software version 4.0 (R Core Team, 2021) [56] and the Jamovi software version 1.8 (The Jamovi Project, 2021) [57]. To test the efficacy of nerve blocks as analgesic medication on the treated versus untreated group (group 1 and 0, respectively; group 1 also included the patients treated with two nerve blocks), the Mann-Whitney U-test was performed for quantitative variables, while the  $\chi^2$  or Fisher's exact test was carried out for qualitative variables. The absence of inter-group differences was considered the null hypothesis (P < 0.05). For statistical convenience, all patients were considered operated under general anesthesia. The following parameters were treated as quantitative variables: "total consumption of opioids (MME)" (equivalent to the sum of all the opioid doses consumed during the four postoperative days), "pain rate reported (1 - 10)", "satisfaction with pain management (1 - 10)", "pain onset time (h)" (expressed in postoperative elapsed hours), and "age". On the other hand, the following parameters were considered qualitative variables: "progress of pain perception" (increased, decreased, or constant), "opioids requested" (yes, no), "opi-oids received" (yes, no), "bowel care received" (yes, no), and "PONV" (yes, no).

Furthermore, a multiple linear regression model was built to explore potential relationships between the variables. "Total consumption of opioids (MME)" was set as the dependent variable, while the remaining parameters were used as independent variables. Qualitative variables were included by setting one of the possible states of each as the reference level (e.g., "request of opioids (no)" constituted the reference level of "request of opioids (yes)"). Parameters considered independent variables but that were not tested in the previous part of the analysis included "bowel activity (yes)", "nerve block (1)", "nerve block (2)" (in this case, administration of two nerve blocks was initially maintained as a separate category; "nerve block (0)" constituted the reference level of both), "recovery time (h)" (the intervening time between PACU and hospital discharge), and "sex". The null hypothesis expected the absence of any significant regressor among the independent variables selected (P <0.01). However, the regressors can act as confounders of each other, resulting in noise due to multicollinearity. To account for this issue, a model selection procedure was computed to obtain a second optimal model. This process consisted of using the stepwise backward elimination, stepwise forward selection, and best subset selection methods (implemented with the AIC criterion). The best subset selection method generated the optimal model with relatively any number of variables, and from it, the leave-one-out cross-validation (LOOCV) yielded the model with the optimal number of regressors.

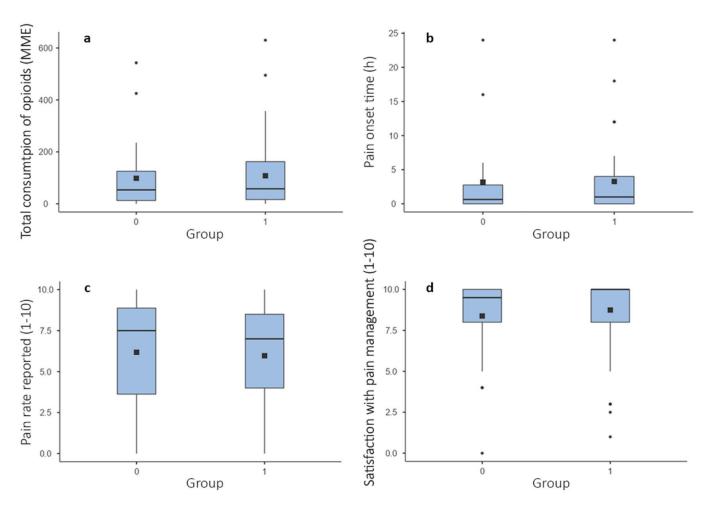
# Results

Cumulative postoperative opioid consumption ("total consumption of opioids (MME)") was virtually equivalent (P =0.483, Mann-Whitney U-test) between group 0 and 1. Therefore, we cannot reject the starting null hypothesis (Fig. 1a). A similar equivalence was recovered with the between-group comparisons in regards to "pain rate reported (1 - 10)" (P = 0.716, Mann-Whitney U-test), and "satisfaction with pain management (1 - 10)" (P = 0.473, Mann-Whitney U-test) (Fig. 1b-d). Conversely, the between-group comparisons concerning "progress of pain perception" and "bowel care received" yielded significant differences (P = 0.011, Fisher's exact test, and P < 0.001,  $\chi^2$  test, respectively) (Fig. 2). The between-group comparison for the qualitative variables "request of opioids" and "opioids received" did not show any significant discrepancy (P = 0.332 and P = 0.442, Fisher's exact test, respectively). Additionally, the Fisher's exact tests performed on each established surgery category (type and respective percentage of total surgeries are the following: "orthopedics, trauma", 62.8%; "general surgery", 22.3%; "neurosurgery", 9.6%; "urology", 3.2%; "gynecology", 2.1%) for the inter-group comparison were not significant. Therefore, these variables were excluded from the multiple linear regression analysis to reduce statistical noise.

The complete multiple linear regression model (P = 0.0003325, F-statistic test) evaluated the variable "recovery time (h)" as the sole significant regressor relative to "total consumption of opioids (MME)" (Table 1, P < 0.01). Thirty-three observations were eliminated due to missing values. All model selection methods resulted in a refined model (P =  $2.79 \times 10^{-6}$ , F-statistic test) with three regressors: "age", "nerve block (1)", and "recovery time (h)" (Table 2, P < 0.01). "Nerve block (2)" was initially included in the refined model, although it was successively included in "nerve block (1)" due to the paucity of observations rendering this separate variable non-significant. In the refined model, "recovery time (h)" was again maintained as the only variable significantly correlated with "total consumption of opioids (MME)" (P =  $7.81 \times 10^{-7}$ ).

## Discussion

As the results show, group 0 and 1 only differ for one of the clinical parameters relevant to the working hypothesis. In Figure 2a, it is evident that a higher proportion of group 1 patients experienced increasing pain intensities throughout the postoperative period compared to group 0 patients, and vice versa for decreasing pain intensities. Therefore, the part of the starting hypothesis expecting an ameliorated condition in patients treated with sPNBs must be rejected and revised. Also, it is to be concluded that the administration of sPNBs did not reduce nor affect opioid consumption, and the corresponding prediction in the starting hypothesis is to be rejected. Comparing the means and medians in Figure 1a renders it possible to appreciate a heightened opioid consumption within group 1, although this is insignificant. As a secondary outcome, it is noteworthy that group 1 patients also required substantially more bowel



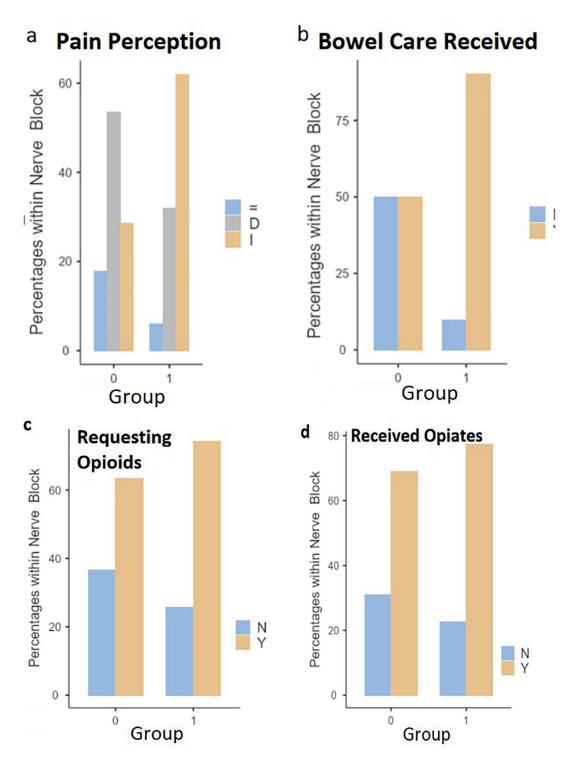
**Figure 1.** Box plots illustrating the results of the Mann-Whitney U-tests performed on four relevant quantitative variables for the between-group comparison: (a) total consumption of opioids (MME), (b) pain onset time, (c) pain rate reported (1 - 10), and (d) satisfaction with pain management (1 - 10). All tests are non-significant. In-box partitions represent means. Square dots represent medians. Box edges represent 25th and 75th percentiles. Whiskers represent 5th and 95th percentiles. Round dots represent outlying values. P < 0.05.

care medications (Fig. 2b). To our knowledge, a correlation between sPNBs and the need for bowel care has not been reported elsewhere. Furthermore, the multiple linear regression analysis was uninformative considering the working hypothesis (Tables 1 and 2). Although the "nerve block (1)" variable is included in the refined model (Table 2), it is insignificantly correlated with "total consumption of opioids (MME)" (P < 0.01). Even if P < 0.05 is considered, the significance level of "nerve block (1)" is not sufficiently low enough to infer a robust correlation. Additionally, the relative estimate contradicts the starting hypothesis. As the only highly significant variable, "recovery time (h)" is uninformative for this study since opioid reliance is known to be correlated with prolonged hospitalization (section "Introduction").

The results of the present study were unexpected, because much evidence has been gathered on the efficacy of peripheral nerve blocks for postoperative pain relief while avoiding the deleterious side effects provoked by other systemic medications [2, 49-51]. However, high-quality evidence is scarce.

Competent studies have shown continuous peripheral nerve blocks (cPNBs) to be superior to sPNBs in terms of therapeutic window, particularly for cases of severe orthopedic interventions, like knee arthroplasty. Concomitantly, the efficacy of nerve blocks in alleviating pain is mirrored by a reduced need to consume opioids [31, 49, 58-65]. More specifically, these assertions have been verified for the analgesic power of AC [66-70], femoral [71, 72], femoral triangle [70], ESP [73], transverse abdominis plane (TAP) [74, 75], and other nerve blocks. Despite exerting a protracted analgesic effect compared to sPNBs, cPNBs have been restricted to inpatient settings for safety purposes secondary to the possible complications arising from the use of catheters [49]. Moreover, the related trials have mostly revolved around a few specific surgeries (e.g., knee arthroplasty), and many studies reported contrasting results.

For instance, the prospective, randomized trial of Elkassabany et al (2019) [76] that examined the efficacy of continuous versus single-shot AC blocks found no clinically meaningful



**Figure 2.** Bar plots illustrating the between-group comparisons concerning (a) pain perception, (b) bowel care received, (c) requesting opioids, and (d) received opioids variables. I: increased; D: decreased; =: constant (P = 0.011, Fisher's exact test); Y: yes, N: no (P < 0.001,  $\chi^2$  test). P < 0.05.

differences, except for some marginal advantages provided by the continuous block.

Additionally, the three study groups displayed a similar opioid consumption rate, albeit this outcome did not represent

the primary goal of the overall trial. A comparable investigation carried out by Lee et al (2018) [77] inferred a similar overall conclusion. Correspondingly, the randomized, double-blinded trial by Dixit et al (2018) [78] concerning the comparison of

	Estimate	Std. error	t value	Pr (> ltl)
Intercept	80.2868	120.1329	0.668	0.507346
Age	-2.7692	1.1701	-2.367	0.022312
Bowel activity (yes)	-5.7327	42.5860	-0.135	0.893516
Bowel care received (yes)	-5.1818	41.7323	-0.124	0.901736
Nerve block (1)	46.7676	38.6682	1.209	0.232803
Nerve block (2)	-30.0692	77.9230	-0.386	0.701401
Opioids received (yes)	73.7575	46.3381	1.592	0.118448
Opioids requested (yes)	-20.8222	49.7985	-0.418	0.677841
Pain onset time (h)	-5.5409	3.6244	-1.529	0.133316
Pain rate reported (1-10)	0.5114	6.0608	0.084	0.933130
PONV (yes)	-31.0436	34.7748	-0.893	0.376765
Progress of pain perception (D)	-57.6477	59.2166	-0.974	0.335506
Progress of pain perception (I)	25.3993	53.3519	0.476	0.636327
Recovery time (h)	1.3808	0.3500	3.946	0.000276
Satisfaction with pain management (1-10)	10.9298	7.1683	1.525	0.134323
Sex (M)	-31.5215	32.4362	-0.972	0.336345

**Table 1.** Complete Multiple Linear Regression Model With "Total Consumption of Opioids (MME)" as Regress (P = 0.0003325, F-statistic test)

"Recovery time (h)" is the only significant regressor. P < 0.01.

single-shot and continuous femoral nerve blocks did not yield any significant difference between the two block types with respect to pain management, opioid consumption, and other clinical parameters. In contrast with part of these findings, the recent meta-analysis by Ma et al (2020) [79] revealed the association between continuous femoral blocks and a decreased need for opioids in the immediate postoperative period. The pain scores, incidence of nausea, and duration of hospitalization were insignificantly different. In opposition to this interpretation, a similar systematic study by Li et al (2020) [80] inferred a substantial superior analgesic potency of the continuous femoral nerve block compared to its single-shot counterpart. Dissimilar conclusions like those mentioned above might be generated by the different parameters associated with statistical clinical relevance or the different methodologies used. Some discrepancies highlighted between the utilization of the two block typologies help to circumscribe their ideal hospital setting. The double-blinded, randomized, controlled trial conducted by Turner et al (2018) [81] also retrieved an

equivalence between the two types of AC blocks, despite the continuous one proving to be more beneficial after 36 - 42 h from its application. As secondary appraisals, no discrepancies concerning opioid consumption, patient satisfaction, or incidence of PONV were noted. These considerations warrant further clinical trials to assess the differences between cPNBs and sPNBs. Studies like the above mentioned add to the body of knowledge pointing to sPNBs as suitable candidates for early dismissal fast track surgical operations.

Some studies even discarded the existence of improvements provided by nerve blocks versus placebo. Jaeer et al (2012) [82] performed a double-blinded, randomized trial whereby a catheter was applied in the AC of all patients undergoing total knee arthroplasty during general anesthesia. However, only a portion of the group received the block, whereas the remainder received a sham injection. This study inferred decreased pain during knee flexion in the group treated with the block, while pain at rest and morphine consumption did not differ. Similarly, Goytizolo et al (2019) [83] regarded opioid

**Table 2.** Refined Multiple Linear Regression Model With "Total Consumption of Opioids (MME)" as Regress (P = 2.79 × 10<sup>-6</sup>, F-statistic test)

	Estimate	Std. error	t value	Pr (> ltl)
Intercept	132.9930	75.5701	1.760	0.0840
Age	-2.8915	1.1146	-2.594	0.0121
Nerve block (1)	70.0594	32.0716	2.184	0.0332
Recovery time (h)	1.7344	0.3113	5.572	7.81× 10 <sup>-7</sup>

The model was obtained by using the stepwise backward elimination, stepwise forward selection, and best subset selection methods (implemented with the AIC criterion). "Recovery time (h)" is the only significant regressor. P < 0.01.

consumption to be unaffected by administering the AC block as part of a multimodal regimen. Some studies even refused the use of PNBs as a viable alternative to other analgesic therapies. In the study conducted by Patel et al (2020) [84], a database of more than 300,000 patients subjected to primary total knee arthroplasty were queried to test the efficacy of both continuous and single-shot femoral blocks for postoperative pain control. Concerningly, patients treated with either type of block experienced more postoperative complications in comparison to the untreated sample. These adverse outcomes affected particularly the group using the continuous block, and were represented by a higher incidence of postoperative falls, inpatient readmission, and several systemic adverse conditions. Despite inaccurately counting the opioid doses consumed, the authors argued for an overall equivalent opioid consumption between all three study groups. Given its design, this study presents some limitations, such as the impossibility of accounting for adjuvants' potential administration. Similarly, Lyngeraa et al (2019) [85] observed a negligible effect from both typologies of nerve blocks on opioid consumption decrease, although other parameters improved.

How sPNBs alone impact opioid consumption has been investigated by an insufficient number of studies (e.g., [86]). Abdallah et al (2015) [87] addressed this particular gap in their meta-analysis examining the efficacy of the single-shot interscalene block for successfully attenuating pain following shoulder surgery. Their study recovered a tangible short-lasting analgesic effect delivered by the block, accompanied by decreased postoperative opioid consumption, reduced PONV, and swifter recovery with early discharge. Instead, Seelam et al (2020) [86] assessed the analgesic efficacy of single-shot ultrasound-guided ESP block on patients undergoing mastectomies. The authors observed a reduced opioid consumption in the patient group treated with the block, although the incidence of PONV was equal across the whole study population. Conversely, Meftah et al (2020) [88] failed to detect a significant effect of sPNBs on opioid consumption compared to other analgesic treatments (e.g., periarticular injection). On the same side, Perlas et al (2013) [89] suggested implementing singleshot AC blocks as rescue analgesics.

The present study may add slightly to the body of knowledge expressing skepticism on using sPNBs over other analgesics. However, the conclusions reached in this work are far from firm. Several limitations and different methodologies hinder comparison with other clinical evidence. For instance, the "recovery time (h)" variable was defined based on the available time point data, and the distinction between pain on movement and at rest was not considered [12]. Furthermore, our investigation did not use a randomized, double-blinded, placebo-controlled approach. Missing values influenced the purity of the statistical analysis, whose low resolution was also influenced by the overabundance of categories compared to the small sample size. The insignificance relative to most of our parameters represented a strong liability because factors such as preoperative pain and sex are known to affect postoperative pain [12]. Additionally, the refined regression model was fitted to the data on which the model selection was performed, and this most likely rendered the inference estimates biased and distorted. Considering the plethora of often unaccounted-for clinical parameters and individual differences among patients [4, 14, 20], future clinical trials should correct all these shortcomings by focusing on a larger population and by including a restricted category number per parameter.

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#### **Financial Disclosure**

None to declare.

# **Conflict of Interest**

None to declare.

# **Informed Consent**

Data used were exempt from consent to participate or publish secondary to the nature of the study being a retrospective cohort, retrospectively looking at collected data. No identifiable patient data was used.

## **Author Contributions**

The study was designed by JS, NS, SB and CR. The data were collected by CR and GM. The initial draft was written by SB, CR and GM. The calculations, data analysis, tables and figures were done by JS, SB, CR, and GM. Final draft and discussion were written by JS, NS, SB, and CR. All authors attest to significant contributions to this work.

# **Data Availability**

The authors declare that data supporting the findings of this study are available within the article.

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