

Diagnostic role of the neutrophil-to-lymphocyte ratio and the platelet-to-lymphocyte ratio in breast cancer: A systematic review and meta-analysis

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Abstract. Neutrophil-to-lymphocyte ratio (NLR), platelet-to-lymphocyte ratio (PLR) and lymphocyte-to-monocyte ratio (LMR) may be indicative of breast cancer (BC); however, this remains inconclusive. With the aim to assess the current literature to evaluate the diagnostic roles of NLR, PLR and LMR in BC, a systematic literature search was performed using the PubMed, Embase, Cochrane Library, China National Knowledge Infrastructure, Wanfang Database, VIP database and China Biology Medicine disc databases up to August 29, 2023. The standardized mean deviation and 95% confidence intervals (CI) for each outcome was reported, and heterogeneity and publication bias were assessed. Overall, 39 studies were included in the present study. Pooled analysis with the random-effects model demonstrated that patients with BC had significantly higher NLR and PLR, and a lower LMR, compared with non-BC subjects. The pooled sensitivities of the

NLR and PLR were 0.68 (95% CI, 0.59-0.75) and 0.55 (95% CI, 0.36-0.72), respectively, and the pooled specificities of the NLR and PLR were 0.75 (95% CI, 0.68-0.81) and 0.80 (95% CI, 0.51-0.94), respectively. However, the limited number of studies included hindered the evaluation of the diagnostic role of LMR. In summary, a higher NLR and PLR and lower LMR were associated with the presence of BC. NLR and PLR may be potential blood-based biomarkers for the differentiation of BC. Despite these findings, further studies are needed to validate their clinical applicability and practicality. International Prospective Register of Systematic Reviews registration no. CRD42024522226.

Introduction

Breast cancer (BC) is one of the most common malignancies affecting the health of women worldwide with an estimated 2.3 million new cases and 685,000 deaths in 2020. It is also the leading cause of cancer-related death in women (1). Early screening and diagnosis of BC has positive impacts on treatment outcomes and the psychology of the patient as well as decreasing the economic burden of this cancer (2). Widespread BC screening in the USA and other high-income countries has contributed to a decreased number of mortalities from BC in these populations over recent decades (3). It has also helped to identify contraindications to medication, e.g. BC is a contraindication for estrogen plus progestogen (4). However, there are ethical challenges and economic and demographic differences that hinder early screening in underdeveloped countries and regions, which, for example, makes it difficult to systematically implement BC screening in sub-Saharan Africa (5). Furthermore, the contradiction between a large population and limited resources poses a huge challenge for China to increase the national coverage of BC screening (3). For BC, breast self-examination (BSE) and clinical breast examination

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Abbreviations: NLR, neutrophil-to-lymphocyte ratio; PLR, platelet-to-lymphocyte ratio; LMR, lymphocyte-to-monocyte ratio; BC, breast cancer; SMD, standardized mean differences; CI, confidence interval; PRISMA, Preferred Reporting Items for Systematic Review and Meta-Analysis; NOS, Newcastle-Ottawa Scale; ROC, receiver operating characteristic

Key words: NLR, PLR, LMR, BC, meta-analysis

(CBE) can catch the first physical changes in the breasts and, subsequently, a mammography should be performed (6). However, in resource-limited settings, a mammography is assessed as not cost-effective (5). In addition, current research does not indicate that there is an improved detection and diagnosis rate of early BC using BSE and CBE (5,6). In recent decades, the serum concentration of tumor markers has been used to detect tumor activity, as suggested by the updated recommendations of the American Society of Clinical Oncology (7). Tumor markers are minimally invasive, readily available and low-cost, providing an alternative approach to BC screening (7,8). However, the efficacy of mainstream clinical tumor markers has been questioned due to their low diagnostic sensitivity of the disease at early stages, such as carcinoembryonic antigen (7). Thus, there is a need for affordable, accurate and sensitive markers for the monitoring of BC. Research on potential tumor markers may be of significance for the screening of BC, especially in low- and middle-income countries (9).

Cancer development is, among other factors, driven by a tumor-mediated disorder of immunity, along with immune disorders in all cell populations (10). There is evidence suggesting that the neutrophil-to-lymphocyte ratio (NLR), platelet-to-lymphocyte ratio (PLR) and lymphocyte-to-monocyte ratio (LMR), among those derived from peripheral whole blood cell count, are useful indicators of BC onset, development and prognosis (11-13). Despite systematic reviews of peripheral whole blood cell count-derived indicators of BC in the efficacy of drug therapy for BC and the disease prognosis (13-15), no meta-analyses have reported associations between peripheral whole blood cell count-derived indicators (NLR, PLR and LMR) and BC, to the best of our knowledge. Disordered neutrophils, overactivated platelets and reduced lymphocytes create an optimal environment for tumor growth, progression and metastasis (13,14,16,17). NLR and PLR are positively associated with risk for multiple types of cancer while LMR is negatively associated (18). In addition, these biomarkers change prior to diagnosis, and they can be used to predict the presence of malignancy (16,18). Moreover, these markers are low-cost, accessible and sensitive, making them particularly suitable for BC screening in underdeveloped countries and regions (3,5,16,18). However, previous studies have come to different conclusions on the differences in NLR, PLR and LMR between patients with BC, and non-BC and healthy subjects and patients with benign breast disease (17,19-22). This difference has led to uncertainty on the diagnostic role of NLR, PLR and LMR in BC screening and earlier identification. Therefore, the present study performed a meta-analysis to assess the current literature to evaluate the diagnostic role of NLR, PLR and LMR in BC.

Materials and methods

Literature search. The methods of the present study were based on the updated guidelines for systematic review reports of the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) 2020 statement (23). 'Breast neoplasms', 'neutrophils and lymphocytes', 'NLR', 'neutrophil-lymphocyte ratio', 'blood platelets and lymphocytes', 'PLR', 'platelet-lymphocyte ratio', 'lymphocytes

and monocytes', 'LMR' and 'lymphocyte-monocyte ratio' were used as medical subject headings terms and keywords to search in PubMed (<https://pubmed.ncbi.nlm.nih.gov/>), EMBASE (<https://www.embase.com/>), Cochrane Library (<https://www.cochranelibrary.com/>), China National Knowledge Infrastructure (<https://www.cnki.net/>), Wanfang Database (<https://www.wanfangdata.com.cn/>), VIP database (<http://www.cqvip.com/>) and China Biology Medicine disc (<http://www.sinomed.ac.cn/index.jsp>), for a time frame starting from database establishment to August 29, 2023 (24). Articles were limited to English and Chinese versions only. Additional manual searches of relevant journals were performed and the relevant documents were tracked in the references. A total of two authors (DY and HW) independently screened the research literature, and any differences were discussed and resolved with a third author (DA).

Eligibility criteria. The inclusion criteria were as follows: i) Study type: Observational studies, including cross-sectional studies, cohort studies, case-control studies or case series; ii) subjects: Patients with BC that had received no treatment (including surgery, drugs and radiation therapy); iii) interventions: NLR, PLR and LMR; iv) controls: Healthy and benign controls; and v) outcomes: Diagnosis.

The exclusion criteria were as follows: i) Cellular experiments, *in vitro* studies; ii) studies assessing NLR, PLR and LMR data of patients with BC after treatment (surgery, drugs and radiotherapy); iii) literature reviews, comments, correspondence letters and case reports; iv) duplicate publications; v) literature with unavailable full text, incomplete data, unavailable raw data and unavailable synthetically extracted data; and vi) relatively low-quality literature [Newcastle-Ottawa scale (NOS) score <6] (25).

Literature screening, quality assessment and data extraction. A total of two investigators (DY and HW) reviewed the titles, abstracts, keywords and full text of the literature separately, and then screened and analyzed them and assessed their quality against the inclusion and exclusion criteria. Any differences arising during the study were resolved through discussion with the third investigator (DA).

The NOS was used to assess the quality of each cohort and case-control study based on the following components: i) Selection of the cohort; ii) comparability of cohorts based on the design or analysis; and iii) how the exposure was ascertained (25). The cross-sectional study evaluation criteria of the Agency for Healthcare Research and Quality (AHRQ) was used (25). The data were then extracted according to an independently pre-defined information extraction form (15) and reviewed by two investigators (DY and HW). Any discrepancy between data extractions was resolved through discussion with the third investigator (DA). The data extracted included the surname of the first author, year of publication, country, age and sex of the patient, as well as the sample size, disease stage, NLR, PLR and LMR.

Statistical analysis. Statistical analysis was performed using RevMan 5.3 (<https://www.cochrane.org/>) and STATA 12.0 software (StataCorp LP). The mean and standard deviation values were extrapolated from the median and

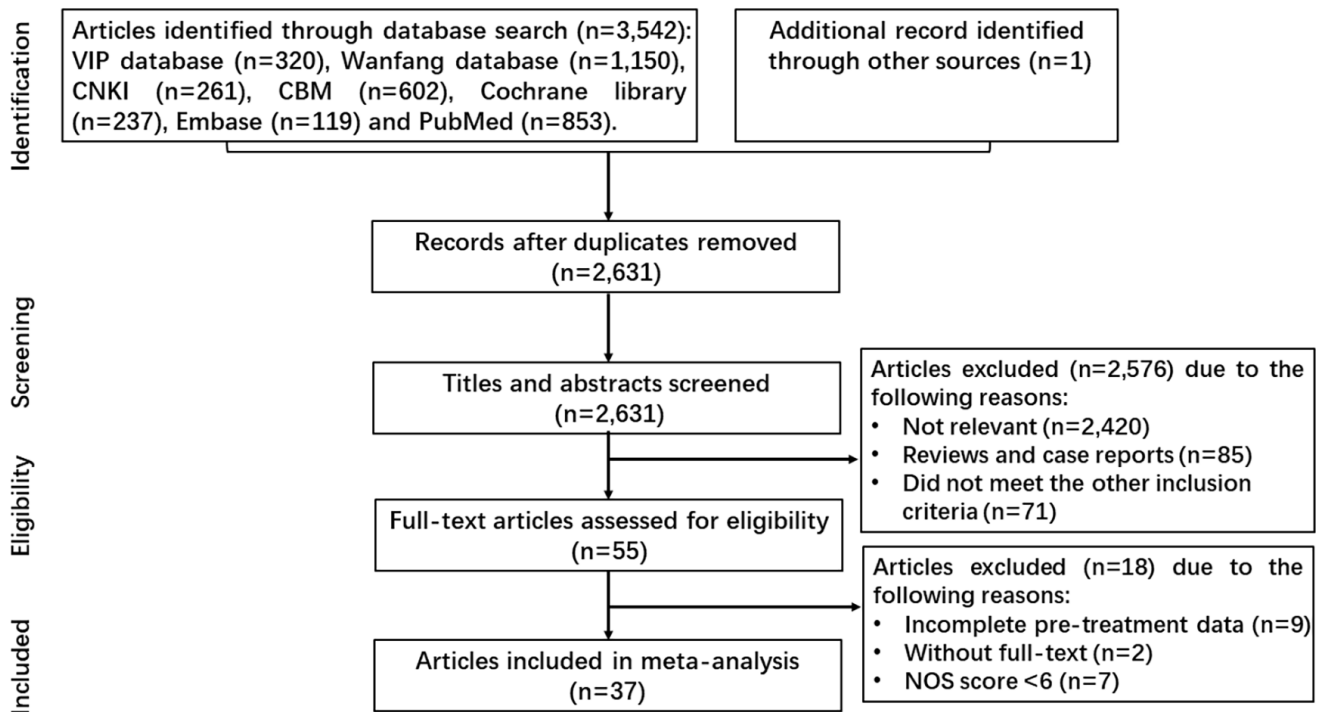


Figure 1. Flowchart of the literature search and study selection in the present study. CNKI, China National Knowledge Infrastructure; NOS, Newcastle-Ottawa Scale; CBM, China Biology Medicine disc.

interquartile range/range values. NLR, PLR and LMR were analyzed using the standardized mean difference (SMD) and 95% confidence intervals (CI). A random-effects model was used in the present study according to the Cochrane Handbook for Systematic Reviews of Interventions, as a systematic review and meta-analysis including multiple studies from different groups (26). $P < 0.05$ was considered to indicate a statistically significant difference. The I^2 metric and χ^2 test were used to assess the heterogeneity among studies. If there was significant heterogeneity ($P < 0.1$, $I^2 \geq 50\%$), subgroup analysis was performed to identify the causes of heterogeneity.

The command 'metandi' was used to calculate the diagnostic odds ratio (DOR), pooled specificity, specificity, positive likelihood ratio and negative likelihood ratio in STATA 12.0. A summary receiver operating characteristic (ROC) curve was also generated. Sensitivity analysis was performed using STATA 12.0 using the 'leave-one-out' method. Publication bias was assessed using funnel plots, Begg's test and Egger's test. The present study is fully compliant with the PRISMA guidelines.

Results

Search results and included studies. A total of 3,542 articles were retrieved through the initial screening, and one was added by tracking references. After removing 912 duplicates, 2,631 articles remained after the initial screening. Following literature screening by title, abstract and keywords, a total of 2,576 irrelevant studies were also excluded. After full-text reading, an additional 18 studies were excluded due to incomplete pre-treatment data (9 articles), without full-text (2 articles), and NOS score < 6 points (7 articles). Finally, 37 articles were included in the meta-analysis (Fig. 1).

Characteristics of the population and quality assessment. The 37 included studies in the present meta-analysis involved in 8 countries: Greece (n=1), Iraq (n=1), Denmark (n=1), Italy (n=1), Iran (n=2), Egypt (n=2), Turkey (n=4) and China (n=25) (Table I). Of these studies, 37 had cohort or case-control designs with NOS score 6-8, classifying them as moderate or high-quality studies. The other two studies were cross-sectional studies with AHRQ scores of 9 and 10 points, respectively (Table I). Furthermore, 16 studies analyzed ROC curves for NLR, seven for PLR and two for LMR (Table II).

Differences in NLR level between patients with BC, and non-BC and healthy subjects or patients with benign breast disease. A total of 7,479 patients with BC vs. 7,018 with non-BC (3,628 healthy and 3,390 patients with benign breast disease) subjects were included in the meta-analysis. The random effect analysis revealed that NLR was significantly higher in the BC group compared with the non-BC (SMD=0.59; 95% CI, 0.47-0.71; $P < 0.00001$; Fig. 2), healthy (SMD=0.56; 95% CI, 0.39, 0.73; $P < 0.00001$; Fig. S1) and patients with benign breast disease (SMD=0.70; 95% CI, 0.51, 0.90; $P < 0.00001$; Fig. S2) groups. Due to heterogeneity, further subgroup analysis was performed and the results demonstrated that the hematology analyzer (in non-BC and healthy subjects, and patients with benign breast disease) and study design and NOS score (in non-BC subjects and patients with benign breast disease) were the sources of heterogeneity (Tables SI-SIII).

Diagnostic value of NLR for differentiating between patients with BC and non-BC subjects. A total of 15 studies had a pooled sensitivity of 0.68 (95% CI, 0.59-0.75), and a pooled specificity of 0.75 (95% CI, 0.68-0.81). The pooled positive

Table I. Characteristics of the enrolled studies.

| First author/s, year | Region | Study design | BC | | | Control | | | NOS/ AHRQ scores | Outcome | (Refs.) |
|-----------------------------------|---------|-----------------|------|-----------|-------------------------|---------------------------|-------------------|-----------------------|---|-------------|---------|
| | | | n | Sex (M/F) | Age, years | Type | n | Sex (M/F) | | | |
| Seretis <i>et al.</i> , 2012 | Greece | Cross-sectional | 35 | 0/35 | 45.5±11.5 ^a | Benign ^c | 44 | 0/44 | 60.2±12.5 | NLR | (21) |
| Ozyalvacili <i>et al.</i> , 2014 | Turkey | Case control | 120 | 0/120 | 54.02±13.45 | Benign | 50 | 0/50 | 51.90±10.26 | NLR | (28) |
| Okuturlar <i>et al.</i> , 2015 | Turkey | Case control | 178 | 0/178 | 53.8±11.5 | Healthy | 107 | 0/107 | 53.7±14.7 | NLR | (29) |
| Qian <i>et al.</i> , 2015 | China | Case control | 82 | 0/82 | 53.5±10.5 ^a | Healthy | 41 | 0/41 | - | NLR | (30) |
| Zhang <i>et al.</i> , 2016 | China | Case control | 104 | 0/104 | 51.08±10.21 | Healthy | 50 | 0/50 | 45.68±11.37 | NLR/PLR | (43) |
| Sun <i>et al.</i> , 2017 | China | Case control | 110 | 0/110 | 54.34±12.28 | Healthy | 78 | 0/78 | 51.54±10.37 | NLR | (17) |
| Wu <i>et al.</i> , 2017 | China | Case control | 53 | 0/53 | 51.21±12.04 | Benign | 122 | 0/122 | 45.75±12.48 | NLR | (44) |
| Huan <i>et al.</i> , 2018 | China | Case control | 126 | 0/126 | 56±9 | Benign | 102 | 0/102 | 52±9 | NLR | (45) |
| Pan <i>et al.</i> , 2018 | China | Case control | 52 | 0/52 | 53.5±12 ^a | Healthy | 47 | 0/47 | 53±11 ^a | NLR | (46) |
| Fang <i>et al.</i> , 2018 | China | Case control | 1540 | 0/1540 | 50.75±10.36 | Benign | 1540 | 0/1540 | 50.75±10.36 | NLR | (11) |
| Cao <i>et al.</i> , 2018 | China | Case control | 118 | 0/118 | 51.25±9.24 | Healthy | 60 | 0/60 | 50.72±9.13 | NLR/PLR | (47) |
| Zhang <i>et al.</i> , 2018 | China | Case control | 92 | 0/92 | 51±10 | Healthy | 50 | 0/50 | 51±12 | NLR | (48) |
| Zhong <i>et al.</i> , 2018 | China | Case control | 115 | 0/115 | 50±7.25 ^a | Healthy | 120 | 0/120 | - | PLR | (49) |
| Zhao <i>et al.</i> , 2018 | China | Case control | 131 | 0/131 | 53±14 ^a | i) Healthy; ii) Benign | i) 95; ii) 120 | i) 0/95; ii) 0/120 | i) 48±9.75 ^a ; ii) 52.5±13.5 ^a | NLR | (31) |
| Pei <i>et al.</i> , 2019 | China | Case control | 412 | 0/402 | 48.17±11.09 | Benign | 412 | 0/412 | 47.67±10.33 | NLR | (50) |
| Alsaadi and Younus, 2019 | Iraq | Case control | 55 | 0/55 | 52.44±8.8 | Healthy | 28 | 0/28 | 47.13±12.79 | NLR/PLR | (51) |
| Xie <i>et al.</i> , 2019 | China | Case control | 136 | 0/136 | 47.04±9.76 | Benign | 127 | 0/127 | 43.13±4.94 | NLR/PLR | (52) |
| Said, 2019 | Egypt | Case control | 84 | 0/84 | 32.7±18.3 | Healthy | 71 | 0/71 | 36.5±15.5 | PLR | (12) |
| Yan <i>et al.</i> , 2019 | China | Case control | 86 | 6/80 | 56.14±11.98 | Benign | 167 | 2/165 | 41.77±11.77 | NLR | (32) |
| Gao <i>et al.</i> , 2019 | China | Case control | 196 | 17/179 | 55.16±12.32 | Healthy | 392 | 34/358 | 55.53±12.54 | NLR/PLR | (53) |
| Liu <i>et al.</i> , 2020 | China | Case control | 433 | 0/433 | 50 ^b | Healthy | 631 | 0/631 | 44 ^b | NLR/PLR | (54) |
| Chi <i>et al.</i> , 2020 | China | Case control | 70 | 0/70 | 39±12 | Benign | 123 | 0/123 | 53±9 | NLR/PLR/LMR | (33) |
| Jørgensen <i>et al.</i> , 2021 | Denmark | Case control | 22 | 0/22 | 62.8±12 ^a | Healthy | 30 | 0/30 | 51.8±9.25 ^a | NLR | (10) |
| Velidedeoglu <i>et al.</i> , 2021 | Turkey | Case control | 50 | 0/50 | 44.3±7.55 | Healthy | 50 | 0/50 | 44.92±8.02 | NLR/PLR | (20) |
| Peng <i>et al.</i> , 2021 | China | Case control | 49 | 0/49 | - | Benign | 48 | 0/48 | - | NLR | (34) |
| Divsalar <i>et al.</i> , 2021 | Iran | Case control | 160 | 0/160 | 51±12 | Healthy | 160 | 0/160 | 50±13 | NLR | (35) |
| Baselice <i>et al.</i> , 2021 | Italy | Case control | 77 | 0/77 | 63.06±11.8 ^a | Benign | 50 | 0/50 | 33.39±12.7 ^a | NLR | (55) |
| Youssry <i>et al.</i> , 2022 | Egypt | Case control | 82 | 0/82 | 49.58±7.7 | i) Healthy; ii) Benign | i) 40; ii) 44 | i) 0/40; ii) 0/44 | i) 48.97±7.1; ii) 47.36±8.43 | NLR/PLR | (41) |

Table I. Continued.

| First author/s, year | Region | Study design | BC | | | Control | | | NOS/AHRQ scores | Outcome | (Refs.) |
|------------------------------|--------|-----------------|------|-----------|-----------------------|---------------------------|-------------------|-----------------------|------------------------------------|-------------|---------|
| | | | n | Sex (M/F) | Age, years | Type | n | Sex (M/F) | | | |
| Dal <i>et al</i> , 2022 | Turkey | Case control | 28 | 28/0 | 60.6±10.6 | Healthy | 22 | 22/0 | 61.0±8.3 | NLR/PLR/LMR | (19) |
| Xu <i>et al</i> , 2022 | China | Case control | 102 | 0/102 | 47.61±9.25 | Healthy | 48 | 0/48 | 48.12±9.47 | NLR | (56) |
| Wang <i>et al</i> , 2022 | China | Case control | 174 | 0/174 | 50±11.5 ^a | Healthy | 181 | 0/181 | 48±10 ^b | NLR/PLR/LMR | (36) |
| Zou <i>et al</i> , 2022 | China | Case control | 653 | 1/652 | 49±15.75 ^b | Benign | 100 | 0/100 | - | NLR | (37) |
| Ding <i>et al</i> , 2022 | China | Cohort study | 286 | 0/286 | 52.0±12.2 | Benign | 143 | 0/143 | 42.5±14.7 | NLR | (57) |
| Alizamir <i>et al</i> , 2022 | Iran | Cross-sectional | 103 | 0/103 | - | Benign | 94 | 0/94 | - | NLR/PLR | (16) |
| Guo <i>et al</i> , 2022 | China | Case control | 278 | 0/278 | 50.79±11.03 | Healthy | 278 | 0/278 | 50.79±11.03 | NLR | (38) |
| Li <i>et al</i> , 2023 | China | Case control | 1224 | 0/1224 | 54±10.37 ^b | Healthy | 1180 | 0/1180 | 56±8.89 ^b | NLR/PLR/LMR | (22) |
| Tang <i>et al</i> , 2023 | China | Case control | 62 | 0/62 | 47.25±9.56 | i) Healthy; ii) Benign | i) 60; ii) 104 | i) 0/60; ii) 0/104 | i) 46.23±10.98; ii) 45.39±10.36 | NLR | (39) |

^aMean and standard deviation were estimated from formulas using the median and range. ^bMean and standard deviation were estimated from formulas using the median and interquartile range. ^cBenign diseases include benign proliferative breast disease, fibroadenomas and intraductal papilloma. BC, breast cancer; NOS, Newcastle-Ottawa Scale; M, male; F, female; NLR, neutrophil-to-lymphocyte ratio; PLR, platelet-to-lymphocyte ratio; LMR, lymphocyte-to-monocyte ratio; -, no data or not available.

Table II. General sensitivity and specificity of the includes studies.

| First author/s, year | NLR | | | PLR | | | LMR | | | (Refs.) |
|---------------------------------|---------------|-------------|-------------|---------------|-------------|-------------|---------------|-------------|-------------|---------|
| | Cut-off point | Sensitivity | Specificity | Cut-off point | Sensitivity | Specificity | Cut-off point | Sensitivity | Specificity | |
| Ozyalvacli <i>et al.</i> , 2014 | 2.96 | 76 | 80 | - | - | - | - | - | - | (28) |
| Okuturlar <i>et al.</i> , 2015 | 2.56 | 30 | 85 | - | - | - | - | - | - | (29) |
| Qian <i>et al.</i> , 2015 | 4.5 | 71 | 81 | - | - | - | - | - | - | (30) |
| Wu <i>et al.</i> , 2017 | 1.9 | 77 | 67 | - | - | - | - | - | - | (44) |
| Zhao <i>et al.</i> , 2018 | 1.995 | 85 | 87 | - | - | - | - | - | - | (31) |
| Yan <i>et al.</i> , 2019 | 1.713 | 79 | 60 | - | - | - | - | - | - | (32) |
| Chi <i>et al.</i> , 2020 | 1.659 | 63 | 77 | 144.339 | 30 | 93 | - | - | - | (33) |
| Peng <i>et al.</i> , 2021 | 1.78 | 69 | 90 | 143.57 | 35 | 98 | - | - | - | (34) |
| Divsalar <i>et al.</i> , 2021 | 2.29 | 32 | 91 | 98.5 | 37 | 84 | - | - | - | (35) |
| Baselice <i>et al.</i> , 2021 | 1.598 | 84 | 38 | - | - | - | - | - | - | (55) |
| Wang <i>et al.</i> , 2022 | 1.85 | 74 | 62 | 131.62 | 74 | 68 | 1.56 | 100 | 2.2 | (36) |
| Zou <i>et al.</i> , 2022 | 1.58 | 69 | 74 | - | - | - | - | - | - | (37) |
| Alizamir <i>et al.</i> , 2022 | 1.24 | 74 | 81 | 96 | 85 | 99 | - | - | - | (16) |
| Guo <i>et al.</i> , 2022 | 1.742 | 55 | 60 | - | - | - | - | - | - | (38) |
| Li <i>et al.</i> , 2023 | - | - | - | 119.43 | 59 | 30 | 5.64 | 38 | 55 | (22) |
| Tang <i>et al.</i> , 2023 | 2 | 58 | 76 | - | - | - | - | - | - | (39) |

NLR, neutrophil-to-lymphocyte ratio; PLR, platelet-to-lymphocyte ratio; LMR, lymphocyte-to-monocyte ratio.

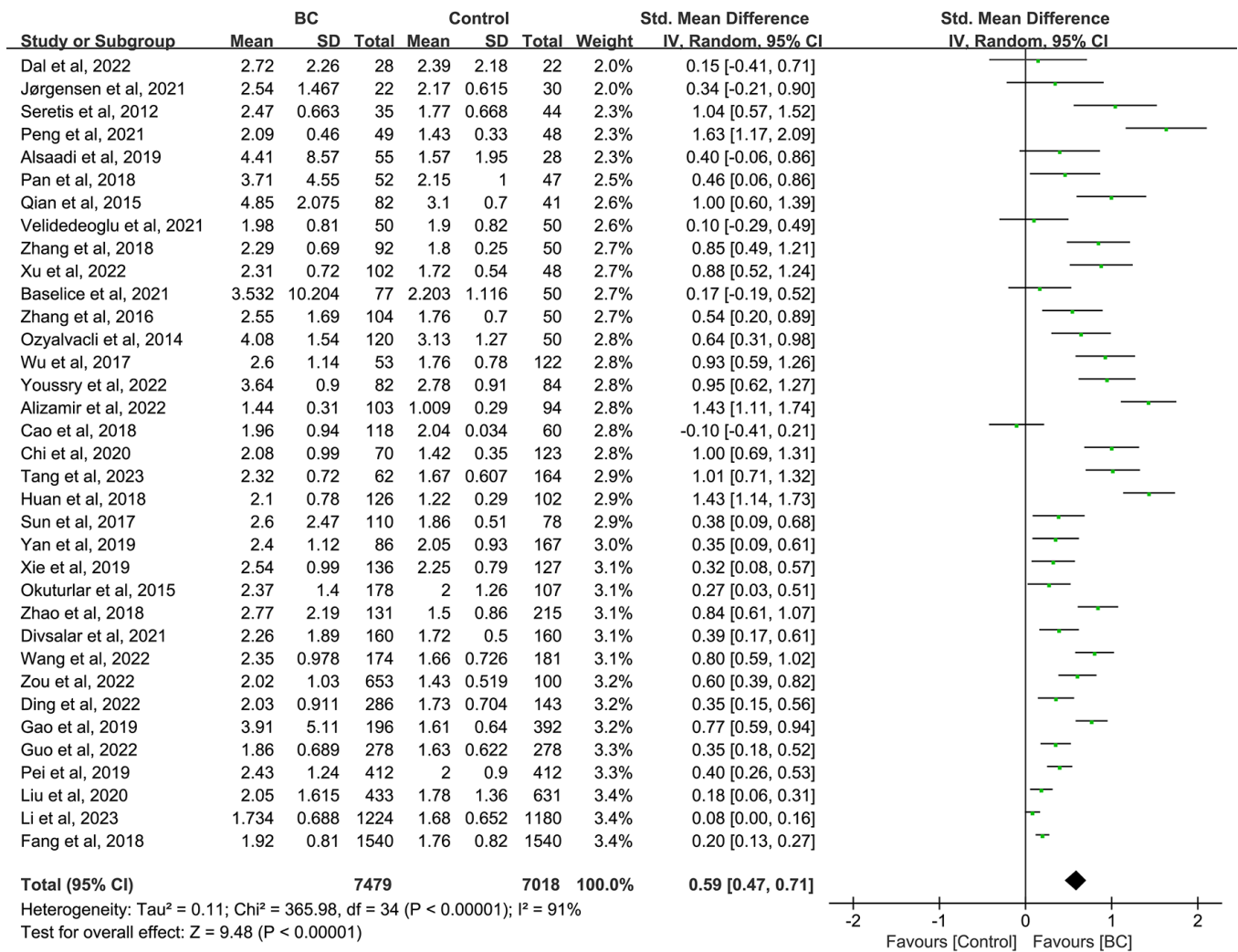


Figure 2. Forest plot of the differences in the neutrophil-to-lymphocyte ratio between patients with BC and non-BC subjects. BC, breast cancer; SD, standard deviation; CI, confidence interval; Std., standard; IV, inverse variance; df, degrees of freedom; Random, random-effects model.

likelihood ratio, negative likelihood ratio and DOR of NLR were 2.75 (95% CI, 2.15-3.51), 0.43 (95% CI, 0.34-0.54) and 6.39 (95% CI, 4.31-9.48), respectively (Fig. 3).

Differences in PLR levels between patients with BC, and non-BC and healthy subjects or patients with benign breast disease. A total of 3,117 patients with BC compared with 3,335 non-BC subjects (2,903 healthy subjects and 432 patients with benign breast disease) from 17 publications were included. The random effect analysis revealed that PLR was significantly higher in the BC group compared with the non-BC (SMD=0.67; 95% CI, 0.41-0.92; P<0.00001; Fig. 4), and healthy (SMD=0.58; 95% CI, 0.35-0.81; P<0.00001; Fig. S3) groups, however it was not significantly higher compared with the benign breast disease group (SMD=0.95; 95% CI, 0.02-1.88; P=0.05; Fig. S4). Further subgroup analysis showed that the hematology analyzer (in non-BC and healthy subjects), study design, NOS score (in non-BC and healthy subjects) and region (in patients with benign breast disease) were the sources of heterogeneity (Tables SIV-SVI), whereas the study by Alizamir *et al* (16) was the source of the heterogeneity in benign subjects, with the results remaining

unchanged after exclusion (SMD=0.45; 95% CI, 0.27-0.63; P<0.0001).

Diagnostic value of PLR for differentiating between patients with BC and non-BC subjects. A total of six studies had a pooled sensitivity of 0.55 (95% CI, 0.36-0.72) and a pooled specificity of 0.88 (95% CI, 0.62-0.97). The pooled positive likelihood ratio, negative likelihood ratio and DOR of NLR were 4.76 (95% CI, 1.17-19.39), 0.51 (95% CI, 0.32-0.81), and 9.30 (95% CI-1.65-56.3), respectively (Fig. 5).

Differences in LMR levels between patients with BC, and non-BC and healthy subjects or patients with benign breast disease. The analysis of the pooled results from four studies revealed that LMR was significantly lower in the BC group compared with the non-BC [SMD=-0.40; 95% CI, -(0.71-0.09); P=0.001; Fig. 6], healthy [SMD=-0.44; 95% CI, -(0.87-0.02); P=0.004; Fig. S5] groups, but was not significantly higher compared with the benign breast disease group [SMD=-0.29; 95% CI, -(0.49-0.00); P=0.06; Fig. S6] groups. Further subgroup analysis demonstrated that the hematology analyzer and NOS score were the sources of heterogeneity in non-BC

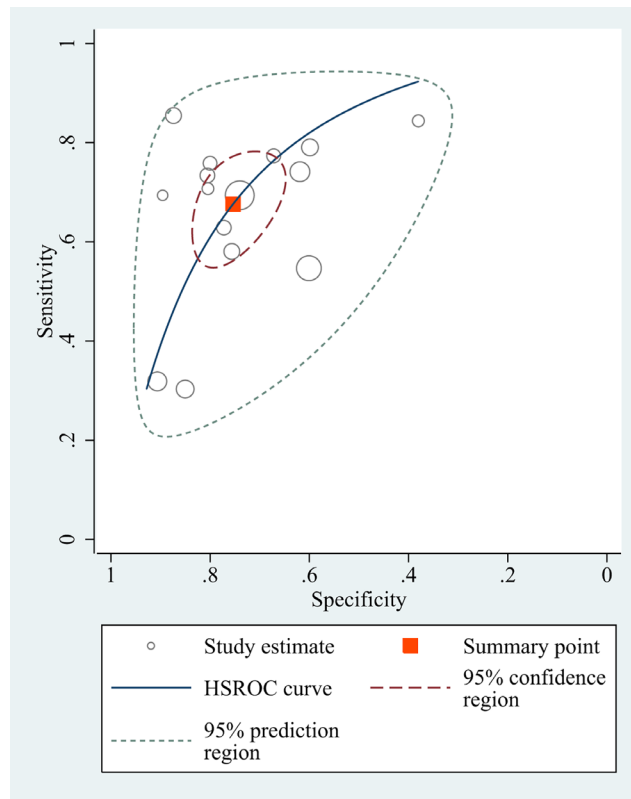


Figure 3. HSROC curve of included studies assessing the diagnostic value of the neutrophil-to-lymphocyte ratio of patients with breast cancer. HSROC, hierarchical summary receiver operating characteristic.

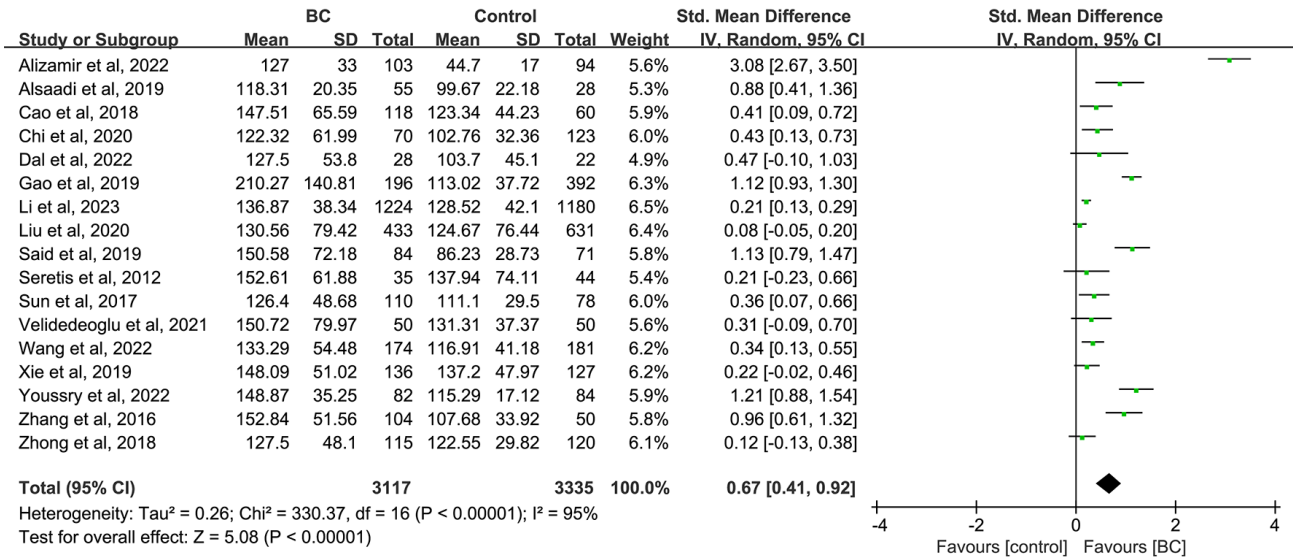


Figure 4. Forest plot of the differences in the platelet-to-lymphocyte ratio between patients with BC and non-BC subjects. BC, breast cancer; SD, standard deviation; CI, confidence interval; Std., standard; IV, inverse variance; df, degrees of freedom; Random, random-effects model.

and healthy subjects, whilst patients with benign breast disease was only included in one study (Tables SVII and SVIII). Only two studies analyzed both sensitivity and specificity, which meant it was not possible to evaluate the diagnostic value of LMR. More research on LMR is required to assess its value.

Sensitivity analysis. The present study performed a sensitivity analysis to evaluate the robustness of the results. The pooled

SMD values did not significantly differ when single studies were removed, suggesting that the results of the meta-analysis were stable (Fig. 7 and Table SIX).

Publication bias. Begg's and Egger's tests and funnel plots were used to determine publication bias. The results demonstrated that there was no publication bias for PLR between BC and benign subjects (Fig. S7 and Table SX). The other

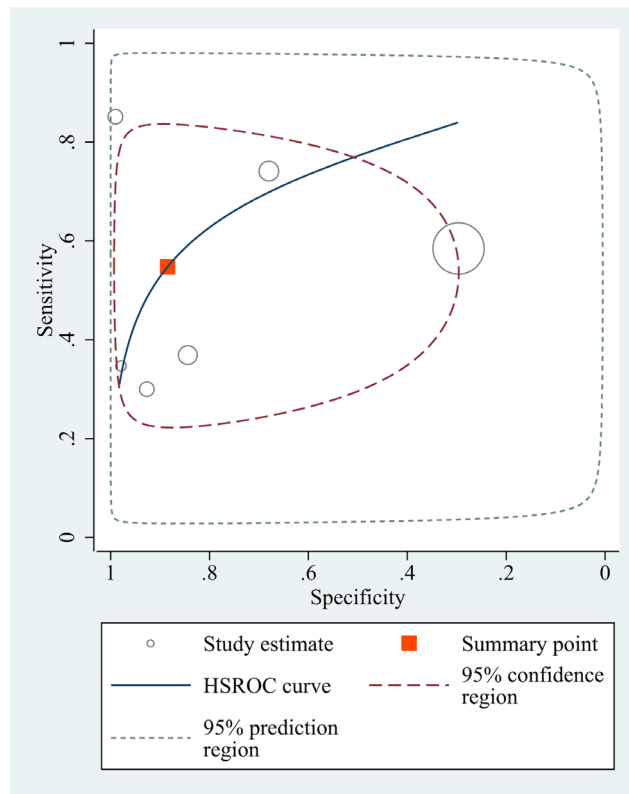


Figure 5. HSROC curve of included studies assessing the diagnostic value of the platelet-to-lymphocyte ratio of patients with breast cancer. HSROC, hierarchical summary receiver operating characteristic.

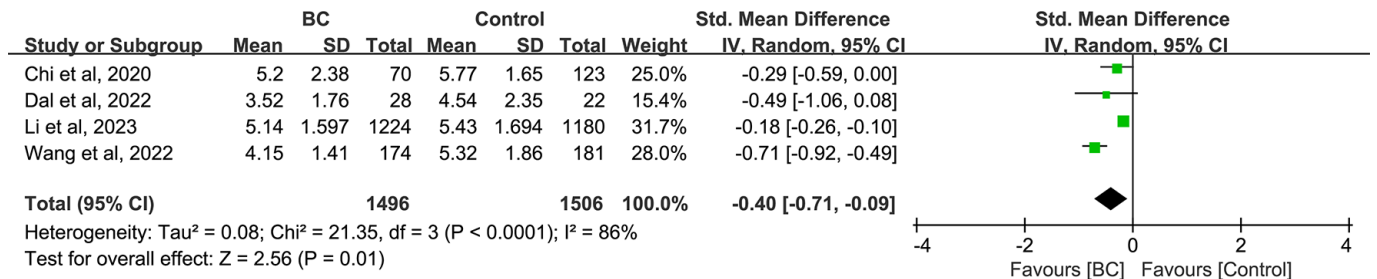


Figure 6. Forest plot of the differences in the lymphocyte-to-monocyte ratio between patients with BC and non-BC subjects. BC, breast cancer; SD, standard deviation; CI, confidence interval; Std., standard; IV, inverse variance; df, degrees of freedom; Random, random-effects model.

asymmetric funnel plots were further processed by trimming and filling, respectively, with no significant differences observed (Fig. S8 and Table SXI), indicating stable results. As ≤5 studies were included, the level of publication bias for LMR was not assessed.

Discussion

The underlying mechanisms of BC are currently unknown, but a notable number of studies have reported that tumor initiation, progression and metastasis are influenced by the host cancer-related inflammatory response as well as tumor micro-environment (6,7,11,20). Therefore, as the derived parameters of peripheral whole blood cell counts are less invasive, more readily available and less expensive compared with mainstream tumor markers (7), their role in cancer-associated inflammatory responses and tumors has become a research topic

of interest. Previous systematic reviews and meta-analyses have demonstrated that peripheral blood cell-derived parameters are notably associated with the efficacy of neoadjuvant chemotherapy for BC and its prognosis (6,13-15). A cohort study also reported that NLR and PLR are associated with an increased incidence of multiple types of cancer, including BC, after 10 years of follow-up (27). Researchers have retrospectively assessed the use of the NLR (17,22,28-39) and PLR (22,34-36) in differentiating between BC, and healthy subjects and patients with benign breast disease, with different conclusions. However, to the best of our knowledge, no study has performed a systematic review and meta-analysis of the association between BC and peripheral blood cell-derived parameters. Therefore, the present study was performed to address the varying results.

The current meta-analysis demonstrated that patients with BC are associated with a higher NLR and PLR, to a medium or

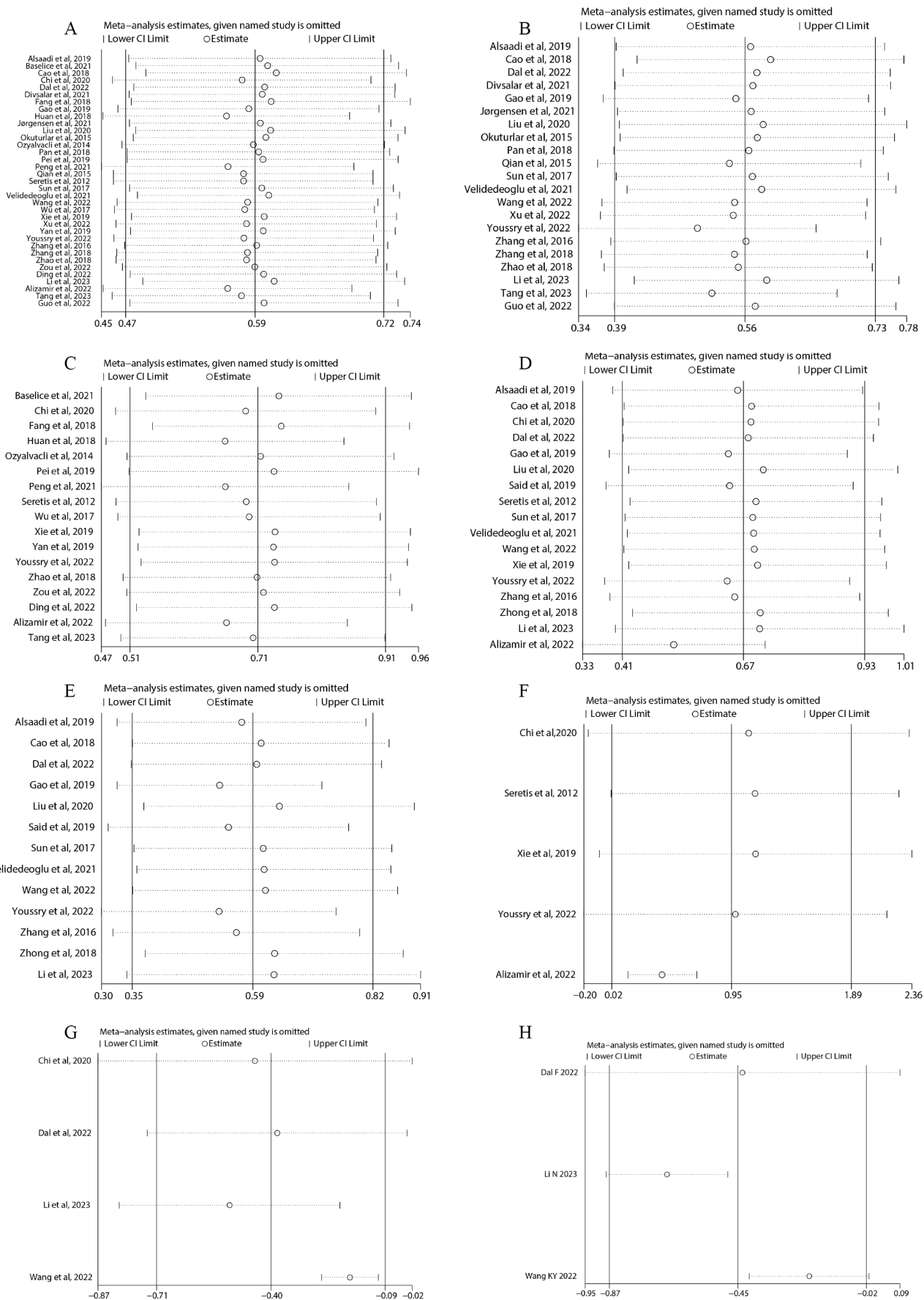


Figure 7. Sensitivity analysis results of blood indexes. Neutrophil-to-lymphocyte ratio of patients with BC and (A) non-BC, (B) healthy subjects and (C) patients with benign breast disease. Platelet-to-lymphocyte ratio of patients with BC and (D) non-BC, (E) healthy subjects and (F) patients with benign breast disease. Lymphocyte-to-monocyte ratio of patients with BC and (G) non-BC and (H) healthy subjects. BC, breast cancer; CI, confidence interval.

large effect, and with lower LMR, to a small effect compared with non-BC individuals (40). The results suggest that NLR, PLR and LMR levels may influence the pathogenesis of BC. As reported by Youssry *et al* (41), altered peripheral blood cells and the cytokines they release may result in a disordered immune response in patients with BC.

Neutrophils are associated with the release of ectopic interleukin-8 in tumor proliferation, progression and metastasis, whereas cancer-associated cytokines, such as tumor necrosis factor- α and interleukin-6, contribute to neutrophilia in solid cancers (7). Neutrophils inhibit the cytotoxic activity of immune cells, such as lymphocytes, natural killer cells and T cells, and reduce regulatory T cells, leading to immune escape (7,10). Activated platelets stimulate cancer-associated inflammation by regulating the migration of hematopoietic and immune cells to the tumor site and promoting metastasis (16). In contrast, lymphocytes activate the host immune response to malignancy by inducing cancer cell death and inhibiting proliferation and migration (17). It has been reported that elevated NLR and PLR and lowered LMR may have potential as biomarkers for predicting the presence of malignancy (22,38), which may help to improve the diagnostic sensitivity for early BC on the basis of common clinical tumor markers, and use of this data may facilitate and improve clinical decision-making for treatment (17). Therefore, NLR and PLR are prospective biomarkers for predicting the pathogenesis of BC. However, these results should be interpreted with caution due to heterogeneity. Given that these indicators are simple, inexpensive, readily available and less invasive, they are especially suitable for BC screening in underdeveloped countries.

The present study has certain limitations: i) The funnel plot and Egger's tests indicate a slight publication bias, with no significant change in direction or magnitude, suggesting that the results are still acceptable after trimming and filling; ii) the meta-analysis had high heterogeneity, and the hematology analyzer was the most important source of heterogeneity, but it had no impact on the robustness of the results. The direction and significance of results for NLR, PLR and LMR did not change in subgroups of hematology analysis, but PLR did not show significance when compared with the benign group. The possible reason is the use of different measurement methods to measure blood cell counts (42), but still provide evidence of a meaningful benefit of a higher NLR and PLR, and a lower LMR in BC as possible potential markers; iii) the geographic concentration of the literature was skewed towards the East Asian region, which may limit the generalizability of the findings. However, in subgroup analysis, the direction of the results did not change, regardless of whether the focus was on East Asian populations. Furthermore, the consistency of the results makes the findings more generalizable; and iv) most of the included studies excluded patients with diseases affecting indices, such as acute or chronic infection, hepatic and renal dysfunction, steroid therapy, inflammatory diseases and hematological disorders. This exclusion criterion increases the validity of the present results. Meanwhile, this exclusion may limit the generalizability of the present findings. Based on the study populations, the NLR and PLR may be used in clinical practice to distinguish patients with BC; however, more real-world application data are still required to support this conclusion.

In summary, the present systematic review and meta-analyses demonstrated that higher NLR and PLR

and lower LMR were associated with the presence of BC. These findings indicate that NLR and PLR may be potential blood-based biomarkers for the differentiation of BC. However, further research is needed to validate their clinical applicability and use.

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Availability of data and materials

The data generated in the present study may be requested from the corresponding author.

Authors' contributions

DY, HW, DA, JZ, QZ, JL, HL and XG contributed to the conception and design of the study. Material preparation and data collection and analysis were performed by DY, HW, DA, QZ, XG and JZ. The first draft of the manuscript was written by DY, QZ, JL and HL, and all authors commented on previous versions of the manuscript. DY and HW confirm the authenticity of all the raw data. All authors have read and approved the final manuscript.

Ethics approval and consent to participate

Not applicable.

Patient consent for publication

Not applicable.

Competing interests

The authors declare that they have no competing interests.

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