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OPEN The association between hepatitis C virus infection and renal cell cancer, prostate cancer, and bladder cancer: a systematic review and meta-analysis

Yucheng Ma^{1,2}, Zhongli Huang^{1,2}, Zhongyu Jian^{1,2} & Xin Wei^{1⊠}

To update the current evidence on whether hepatitis C virus (HCV) infection represents a possible risk factor for renal cell cancer (RCC), prostate cancer (PCa), and bladder cancer (BC). We searched the literature on Pubmed, Web of Science, and Embases before April 2021. A systematic review and meta-analysis were performed. Finally, we extracted 12 studies based on the eligible criteria. Across 11 studies for HCV and RCC, the incorporated RR was 1.28 (95% CI 1.05-1.55), which meant that participants with HCV infection were associated with higher RCC risk. The pooled RR in hazard ratio (HR) subgroup (HR 1.59, 95% CI 1.22-2.08), cohort studies subgroup (RR 1.47, 95% CI 1.18-1.82), and North America subgroup (RR 1.71, 95% CI 1.40-2.09) detected a stronger association between HCV and RCC risk. Although an inverse association was seen for PCa (RR 0.75, 95% CI 0.54-1.03) across seven studies, it was not statistically significant (P = 0.075). There was no significant association between HCV and BC with an incorporated RR of 0.92 (95% CI, 0.82-1.03) across five studies. Our study demonstrated that HCV infection was significantly associated with increased RCC risk. There appeared to be an inverse association for HCV in PCa risk but not statistically significant. No significant association was found between HCV and BC risk. Prospective, large-scale, and welldesigned cohort studies are required to validate the association between HCV and RCC, and to investigate the role of HCV on PCa.

Urologic cancers such as renal cell cancer (RCC), prostate cancer (PCa), and bladder cancer (BC) are most common diagnosed cancers in humans¹. Both non-modifiable and environmental risk factors are identified associated with these cancers. However, the role of hepatitis C virus (HCV) infection on these cancers is still controversial.

HCV infection is a public health problem. It was estimated that approximately 180 million people infected worldwide². HCV infection has been found to be involved in both hepatic diseases and a variety of extra-hepatic diseases³. Although the mechanism of how HCV could affect these cancers was not clearly comprehended, the association between HCV and RCC, PCa, and BC has been extended researched.

A meta-analysis published in 2016 reported an association in HCV infection for RCC risk⁴. However, the state of the two studies used in meta-analysis was estimated and crude, which was one of the limitations. In addition, more studies were published afterward, and the results were still inconclusive. To conclude a more robust estimation of their associations, we updated the current evidence on whether HCV infection represents a possible risk factor for RCC. There was no meta-analysis focusing on the association between HCV and PCa or BC to our knowledge. Although HCV infection rates are on the decline globally, there are still a large number of people living with HCV infection in absolute terms. If the association between HCV and urological tumors can be established, regular urological tumor screening in this population may significantly extend their survival time.

In light of these controversial roles of HCV on RCC, PCa, and BC risk, the purpose of the present metaanalysis is to explore whether HCV represents a possible risk factor for RCC, PCa, and BC by taking all available studies meeting our inclusion criteria into consideration.

¹Department of Urology, Institute of Urology, Laboratory of Reconstructive Urology), West China Hospital, Sichuan University, No. 37 Guo Xue Xiang, Chengdu 610041, Sichuan, People's Republic of China. ²These authors contributed equally: Yucheng Ma, Zhongli Huang, and Zhongyu Jian. [™]email: weixinscu@scu.edu.cn

Materials and methods

To perform this study, we followed the guide of Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines⁵.

Search strategy. Two independent reviewers (YM and ZJ) searched literature on Pubmed, Web of Science, and Embases before April 2021. Relevant studies restricted to the English language were identified according to the eligibility criteria as follows. The Search terms include Hepatitis C Virus, HCV, Prostatic Neoplasms, Prostate Cancer, Renal Cell Cancer, Renal Cancer, Kidney Neoplasm, Kidney Cancer, Renal Neoplasm, Renal Cell Cancer, Bladder Neoplasm, Bladder Cancer. Negotiated between 2 reviewers or consulted with a third author were used to resolve the disagreements.

Eligible criteria. Eligible criteria are described as follows. (1) Population: Patients with RCC, PCa, or BC; (2) Exposure: HCV infection; (3) Comparison: participants control; (4) Outcomes: diagnosis any kind of these three cancers; (5) Study: all study designs. Studies meeting the eligibility criteria listed above and the any kind of relative risk (RR) estimates (standardized incidence ratio [SIR], hazard ratio [HR] or odds ratios [OR]) and their 95% confidence interval (CI) could be directly extracted were included in our meta-analysis. Besides, review, meta-analysis, letters, abstracts, case reports, meeting comments, editorials, and congress reports were excluded from our study.

Data extraction and quality assessment. Two independent reviewers (YM and ZJ) assessed study quality, followed by extracting data from the individual study. Discussion and reevaluation of the methodology would be adopted to resolve the disagreements. Information including the name of the first author, study location, study period, study design, publication year, sample size, measurement of cancers (RCC, PCa, and BC), and the RR (95% CIs) for each category of cancers. The checklist was used to evaluate the quality used was the Newcastle–Ottawa Scale⁶ (NOS) tool. The total score of each study was nine, and a score of more than six was classified into the high-quality group.

Statistical analyses. Effect measures (SIR, OR, and HR) with 95% CI was extracted from the eligible study, and a final combined RR and its 95% CI was obtained by the inverse-variances method. The I² statistics with Q test was used to assess heterogeneity among included studies⁷. For pooled RR, the level of statistical heterogeneity was used to choose a primary statistical model. When heterogeneity was significant (I²>50% or P < 0.1), a random effect model would be used. Otherwise, the fixed-effect model was applied when substantial heterogeneity was not observed. Next, subgroup analyses were performed to identify which clinical factor might contribute to the potential source of heterogeneity. Then, we performed sensitivity analysis by omitting individual studies one by one. Finally, both Begg and Egger test was conducted to identify publication bias. A two-sided P < 0.05 in our study indicated a significant statistical difference. Statistical analyses were performed with Stata version 14 software (Stata Corporation, College Station, TX, USA).

Results

A total of 644 studies using the search criteria were identified. Finally, we extracted 12 studies^{8–19} based on the eligible criteria (Supplementary Fig. S1). The characteristics were demonstrated in Table 1. We found seven cohort studies and five case–control studies that investigated the role of HCV for RCC. There were seven studies reporting the association between HCV and PCa, of which 5 were cohort studies. For studies exploring the role of HCV on BC, 3 were cohort studies, and 2 were case–control studies. The detailed NOS score of each study belongs to cohort studies, and the case–control studies design was shown in Supplementary Tables S1 and S2, respectively.

HCV and RCC. For HCV and RCC, 11 studies^{8-13,15-19} had sufficient data for meta-analysis. As shown in Fig. 1, the summary of RR obtained by the random-effects model (I^2 =70.1%; P <0.001) was 1.28 (95% CI 1.05–1.55), which meant participants with HCV infection had a significantly higher risk of RCC.

We further assessed the clinical factors that might potential contribute to the heterogeneity (Table 2). The pooled RR in HR subgroup (HR 1.59, 95% CI 1.22–2.08; I^2 = 13.5%, P = 0.315), cohort studies subgroup (RR 1.47, 95% CI 1.18–1.82; I^2 = 45.6%, P = 0.088), and USA subgroup (RR 1.71, 95% CI 1.40–2.09; I^2 = 0, P = 0.991) detected a stronger association between HCV and RCC risk.

Sensitivity analyses were shown in Fig. 2. By omitting individual study yielded nonsignificant RR change ranging from 1.22 (95% CI 1.01–1.49) to 1.36 (95% CI 1.15–1.61). Both Begg rank correlation test (P>|z|=0.938) (Supplementary Fig. S2) and Egger linear regression (P>|t|=0.172) demonstrated that there was no significant publication bias among these studies.

HCV and PCa. Across a total of 7 studies $^{8,10,12-15,17,19}$ which examined the role of HCV for PCa, there was no significant association between HCV and PCa was found, with a pooled RR of 0.75 (95% CI 0.54–1.03). Between-study heterogeneity was significant based on the $I^2 = 93\%$ and Q statistic (P < 0.001) and, so a random-effects model was applied (Supplementary Fig. S3).

Analyses in case–control studies subgroup (RR 1.71, 95% CI 1.40–2.09; I^2 =0, P=0.991), Europe subgroup (RR 1.71, 95% CI 1.40–2.09; I^2 =0, P=0.991), studies published before 2012 subgroup (RR 1.71, 95% CI 1.40–2.09; I^2 =0, P=0.991), and subgroup with lower NOS score (RR 1.71, 95% CI 1.40–2.09; I^2 =0, P=0.991) found a significant inverse association with very low heterogeneity. However, we must treat it with caution because of the limited number and relatively low methodological quality studies included in these subgroups. Similarly,

| Study | Publication year | Study design | Study location | Study period | Total numbers | HCV patient number | Cancer numbers | HCV confirmation | Cancer confirmation | Outcomes RR;95%CI | NOS score |
|------------------------|------------------|-----------------|---------------------|-----------------|---------------|--------------------|---|--|---|---|-----------|
| Nyberg ⁸ | 2019 | CS | California, USA | 2008-2012 | 5,332,903 | 35,712 | 9751 in PCa | ICD-9 code confirma- tion and/or a positive HCV RNA test | Confirmed by KPSC- National Can- cer Institute Surveillance, Epidemiol- ogy and End Results (KPSC-NCI SEER) affili- ated cancer registry | [A] HR 1.71; 1.29, 2.25 [B] HR 1.45; 1.24, 1.69 | 7 |
| Liu ⁹ | 2019 | CCS | China | 2008-2016 | 34,505 | 189 | 1287 in RCC | positive for anti-HCV | Confirmed by pathological and histologi- cal examina- tion | [A] OR 0.71; 0.23, 2.25 | 5 |
| Liu ¹⁰ | 2017 | CS | Sweden | 1990-2010 | 29,271 | 29,271 | 24 in RCC; 35 in BC; 93 in PCa | positive for anti-HCV and HCV-RNA analysis by real-time PCR | Confirmed by nationwide Swedish Can- cer Registry | [A] SIR 1.42; 0.91, 2.12 [B] SIR 0.73; 0.59, 0.90 [C] SIR 1.20; 0.84, 1.67 | 6 |
| Lin ¹¹ | 2017 | CCS | Taiwan | 2000-2011 | 53,241 | 1034 | 17,747 in RCC | positive for anti-HCV | Confirmed by Registry for Catastrophic Illness Patient (RCIP) data- base | [A] OR 1.24; 1.07, 1.44 | 7 |
| Mahale ¹² | 2017 | CCS | United King- dom | 1993-2011 | 1,823,538 | 12,107 | 35,960 in RCC; 90,360 in BC; 283,367 in PCa | NR | Confirmed by Surveillance, Epidemiol- ogy, and End Results (SEER) datasets | [A] OR 0.89; 0.76, 1.04 [B] OR 0.84; 0.48, 1.46 [C] OR 0.88; 0.77, 1.00 | 8 |
| Kamiza ¹³ | 2016 | CCS | Taiwan | 1995–2014 | 44,150 | 35,320 | 1065 in RCC; 1324 in BC; 1820 in PCa | NR | Confirmed by Longitudinal Health Insur- ance Database 2000 and The National Health Insur- ance Research Database | [A] HR 0.98; 0.49, 1.93 [B] HR 1.55; 1.04, 2.32 [C] HR 1.20; 0.75, 1.90 | 6 |
| Gonzalez ¹⁴ | 2015 | ccs | Michigan, USA | 2011–2013 | 240 | 9 | 140 in RCC | positive HCV RNA test | Confirmed by pathological and histologi- cal examina- tion | [A] OR 24.20; 2.4, N/A | N/A |
| Allison ¹⁵ | 2015 | CS | USA | 2006-2010 | 12,126 | 12,126 | 17 in RCC; 46 in PCa | positive HCV RNA test and International Classification of Diseases confirmed HCV infection | Confirmed by Chronic Hepatitis Cohort Study (CHeCS), Surveillance, Epidemiol- ogy, and End Results (SEER) cancer registry data- sts and Mul- tiple Causes of Death (MCOD) datasets | [A] SIR 1.7; 1.1, 2.2 [C] SIR 0.6; 0.5, 0.7 | 8 |
| Hofmann ¹⁶ | 2011 | CS | Sweden | 1990-2006 | 43,000 | 43,000 | 29 in RCC | NR | Confirmed by national Cancer, Cause of Death and Hospital Discharge registers datasets | [A] SIR 1.2; 0.8, 1.7 | 6 |
| Omland ¹⁷ | 2010 | CS | Denmark | 1994-2003 | 4204 | 4204 | 4 in RCC; 2 in BC; 2 in PCa | NR | Confirmed by the Danish National Hospital Reg- istry (DNHR) datasets | [A] SIR 3.6; 0.98, 9.22 [B] SIR 1.02; 0.12, 3.7 [C] SIR 0.92; 0.11, 3.31 | 7 |

| Study | Publication year | Study design | Study location | Study period | Total numbers | HCV patient number | Cancer numbers | HCV confirmation | Cancer confirmation | Outcomes RR;95%CI | NOS score |
|----------------------|---------------------|-----------------|-------------------|-----------------|------------------|-----------------------|--------------------------------------|---|--|--|-----------|
| Gordon ¹⁸ | 2010 | CS | USA | 1997–2006 | 72,487 | 3057 | 194 in RCC | positive HCV RNA test | Confirmed by pathological and histologi- cal examina- tion | [A] HR 1.77; 1.05, 2.98 | 8 |
| Amin ¹⁹ | 2006 | CS | Sydney | 1990-2002 | 75,834 | 75,834 | 18 in RCC; 11 in BC; 32 in PCa | positive for anti-HCV and HCV-RNA analysis | Confirmed by NSW Central Cancer Reg- istry (CCR) datasets | [A] SIR 0.9; 0.6, 0.4 [B] SIR 0.4; 0.3, 0.6 [C] SIR 0.7; 0.4, 1.2 | 8 |

Table 1. Characteristics of studies included in the systematic review and meta-analysis. [A] = kidney cancer; [B] = bladder cancer; [C] = prostate cancer; N/A = not available; HCV = hepatitis C; RCC = renal cell cancer; RC = bladder cancer

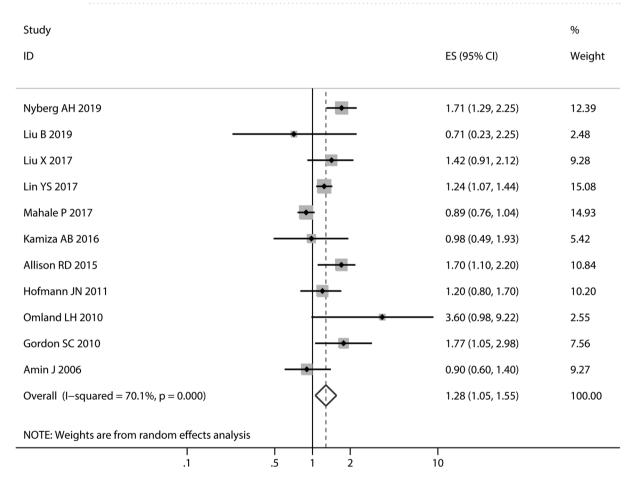


Figure 1. Forest plots of relative risk of studies investigating the association between hepatitis C virus infection and renal cell cancer. Random effects models were used for the primary meta-analysis.

a significant inverse association observed in SIR subgroup (RR 0.58, 95% CI 0.45–0.76) should be interpreted carefully because of the substantial heterogeneity ($I^2 = 66.1$, P = 0.031) (Table 3).

One individual study affected the overall results, a significant inverse association was found between HCV and PCa (RR 0.65, 95% CI 0.54–0.77) (Supplementary Fig. S4). However, substantial heterogeneity was still high (I^2 = 64.2, P = 0.016). Publication bias was not significant based on Begg (P>|z|= 0.293) (Supplementary Fig. S5). and Egger (P>|t|= 0.882) tests.

HCV and BC. The pooled RR enrolling five studies 10,12,13,17,19 was 0.92 (95% CI, 0.82–1.03), representing no significant association between HCV infection and BC risk. Heterogeneity was not significant based on the I^2 = 19.2% and Q statistic (P = 0.293), so a fixed-effects model was used (Supplementary Fig. S6).

| | | Heterogeneity | | | | | | |
|-----------------------|------------|--------------------|---------|-----------------|------------------------|--|--|--|
| Category of variables | Studies, n | I ² , % | P value | RR (95%CI) | P value for difference | | | |
| RCC | 11 | 70.1 | < 0.001 | 1.28(1.05-1.55) | 0.013 | | | |
| Type of RR | | | | | | | | |
| HR | 3 | 13.5 | 0.315 | 1.59(1.22-2.08) | 0.001 | | | |
| OR | 3 | 79.0 | 0.009 | 1.03(0.76-1.39) | 0.866 | | | |
| SIR | 5 | 53.4 | 0.072 | 1.37(1.01-1.84) | 0.040 | | | |
| Study design | | | | | | | | |
| Cohort study | 7 | 45.6 | 0.088 | 1.47(1.18-1.82) | <0.001 | | | |
| Case-control study | 4 | 68.7 | 0.023 | 1.02(0.79-1.33) | 0.857 | | | |
| Geographic area | | | | | | | | |
| North America | 3 | 0 | 0.991 | 1.71(1.40-2.09) | < 0.001 | | | |
| Asia | 3 | 0 | 0.521 | 1.22(1.05-1.40) | 0.008 | | | |
| Europe | 3 | 71.9 | 0.014 | 1.23(0.85-1.79) | 0.267 | | | |
| Australia | 1 | - | - | 0.90(0.59-1.37) | 0.626 | | | |
| Publication year | | | | | | | | |
| After 2012 | 7 | 76.6 | < 0.001 | 1.26(0.99-1.59) | 0.057 | | | |
| Before 2012 | 4 | 59.9 | 0.058 | 1.37(0.90-2.07) | 0.142 | | | |
| NOS score | | | | | | | | |
| ≥7 | 7 | 81.0 | < 0.001 | 1.34(1.04-1.73) | 0.024 | | | |
| ≤6 | 4 | 0 | 0.625 | 1.21(0.94-1.56) | 0.144 | | | |
| ≤6 | 2 | 0 | 1.000 | 1.20(0.91-1.58) | 0.196 | | | |

Table 2. Subgroup analysis of association between HCV and RCC risk. HCV = hepatitis C; RCC = renal cell carcinoma; BC = bladder cancer; PCa = prostate cancer; RR = relative risk; HR = hazard ratio; OR = odds ratio; SIR = standardized incidence ratio; NOS = Newcastle–Ottawa Scale;

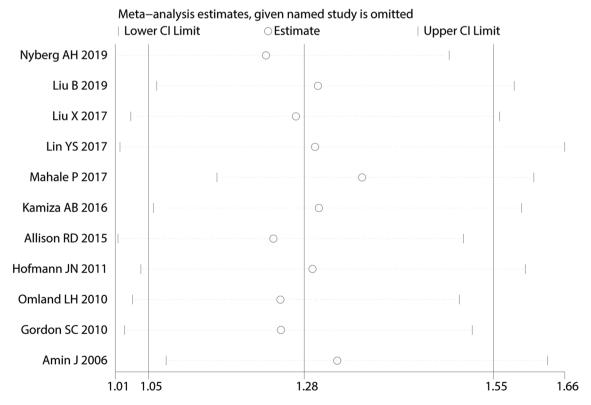


Figure 2. Sensitivity analyses by omitting individual study yielded nonsignificant RR change.

| | | Heterogeneity | | | | | | | | |
|-----------------------|-----------------|--------------------|---------|-----------------|------------------------|--|--|--|--|--|
| Category of variables | Studies, n | I ² , % | P value | RR (95%CI) | P value for difference | | | | | |
| PCa | 7 | 93.0 | < 0.001 | 0.75(0.54-1.03) | 0.075 | | | | | |
| Type of RR | | | | | | | | | | |
| HR | 2 | 70.6 | 0.065 | 1.18(0.71-1.98) | 0.529 | | | | | |
| OR | 1 | - | - | 0.73(0.66-0.81) | < 0.001 | | | | | |
| SIR | 4 | 66.1 | 0.031 | 0.58(0.45-0.76) | < 0.001 | | | | | |
| Study design | Study design | | | | | | | | | |
| Cohort study | 5 | 95.1 | < 0.001 | 0.74(0.44-1.22) | 0.231 | | | | | |
| Case-control study | 2 | 0 | 0.627 | 0.73(0.66-0.82) | < 0.001 | | | | | |
| Geographic area | Geographic area | | | | | | | | | |
| North America | 2 | 98.2 | < 0.001 | 0.93(0.39-2.21) | 0.873 | | | | | |
| Asia | 1 | - | - | 0.84(0.48-1.47) | 0.539 | | | | | |
| Europe | 3 | 0 | 0.930 | 0.73(0.66-0.81) | < 0.001 | | | | | |
| Australia | 1 | - | - | 0.40(0.28-0.57) | < 0.001 | | | | | |
| Publication year | | | | | | | | | | |
| After 2012 | 5 | 94.3 | < 0.001 | 0.83(0.59-1.16) | 0.274 | | | | | |
| Before 2012 | 2 | 9.1 | 0.294 | 0.43(0.26-0.71) | 0.001 | | | | | |
| NOS score | | | | | | | | | | |
| ≥7 | 5 | 95.3 | < 0.001 | 0.73(0.48-1.13) | 0.160 | | | | | |
| ≤6 | 2 | 0 | 0.644 | 0.74(0.61-0.91) | 0.003 | | | | | |

Table 3. Subgroup analysis of association between PCa risk. HCV = hepatitis C; RCC = renal cell carcinoma; BC = bladder cancer; PCa = prostate cancer; RR = relative risk; HR = hazard ratio; OR = odds ratio; SIR = standardized incidence ratio; NOS = Newcastle-Ottawa Scale;

| | | | geneity | | | | | | |
|-----------------------|-----------------|--------------------|---------|-----------------|------------------------|--|--|--|--|
| Category of variables | Studies, n | I ² , % | P value | RR (95%CI) | P value for difference | | | | |
| BC | 5 | 19.2 | 0.293 | 0.92(0.82-1.03) | 0.154 | | | | |
| Type of RR | | | | | | | | | |
| HR | 1 | - | - | 1.03(0.77-1.37) | 0.851 | | | | |
| OR | 1 | - | - | 0.88(0.77-1.37) | 0.055 | | | | |
| SIR | 3 | 25.2 | 0.262 | 1.37(1.01-1.84) | 0.442 | | | | |
| Study design | | | | | | | | | |
| Cohort study | 3 | 25.2 | 0.262 | 1.37(1.01-1.84) | 0.442 | | | | |
| Case-control study | 2 | 36.9 | 0.208 | 0.90(0.79-1.02) | 0.102 | | | | |
| Geographic area | Geographic area | | | | | | | | |
| North America | 0 | - | - | - | - | | | | |
| Asia | 1 | - | - | 1.20(0.75-1.91) | 0.442 | | | | |
| Europe | 3 | 26.9 | 0.255 | 0.92(0.81-1.03) | 0.154 | | | | |
| Australia | 1 | - | - | 0.70(0.40-1.21) | 0.203 | | | | |
| Publication year | | | | | | | | | |
| After 2012 | 3 | 49.4 | 0.138 | 0.93(0.83-1.05) | 0.238 | | | | |
| Before 2012 | 2 | 0 | 0.765 | 0.72(0.43-1.21) | 0.215 | | | | |
| NOS score | | | | | | | | | |
| ≥7 | 3 | 0 | 0.728 | 0.87(0.77-0.99) | 0.031 | | | | |
| ≤6 | 2 | 0 | 1.000 | 1.20(0.91-1.58) | 0.196 | | | | |

Table 4. Subgroup analysis of association between BC risk. HCV = hepatitis C; RCC = renal cell carcinoma; BC = bladder cancer; PCa = prostate cancer; RR = relative risk; HR = hazard ratio; OR = odds ratio; SIR = standardized incidence ratio; NOS = Newcastle-Ottawa Scale;

Subgroup analyses revealed that the pooled RR in higher NOS score subgroup was statistic significant (RR 0.87, 95% CI 0.77–0.99; $I^2 = 0$, P = 0.728) (Table 4). Influence analysis revealed that none study had a greater impact on the pooled RR (Supplementary Fig. S7). No significant publication bias was found according to the Begg (P > |z| = 1.000) (Supplementary Fig. S8). and Egger (P > |z| = 0.536) tests.

Discussion

This meta-analysis had several important findings with respect to the role of HCV infection on RCC, PCa, and BC. First, we found a significantly increased risk of association between HCV and RCC (Fig. 1). This association was detected stronger in the HR subgroup, cohort study subgroup, and North America subgroup accompany with the obviously decreased heterogeneity (Table 2). As we omitted each study in sensitivity analyses (Fig. 2) but nevertheless a nonsignificant trend appeared. Second, although an inverse association was seen for PCa (RR 0.75) in meta-analysis enrolling seven studies, it was not statistically significant (P = 0.075). In subsequent subgroup analyses.

HCV infection was associated with a significantly decreased risk of PCa in the case–control studies subgroup, Europe subgroup, studies published before the 2012 subgroup, and lower NOS score subgroup (Table 2). However, we must interpret it carefully because of the limited number and relatively low methodological quality studies included in these subgroups. Third, regardless of sensitivity analyses, pooled RR remained nonsignificant between HCV infection and BC risk.

RCC incidence has increased over the past two decades, the role of HCV infection on RCC was still inconclusive. A recent meta-analysis found that HCV infection had an increased risk of RCC, but the state of two studies used in meta-analysis were estimated and crude. Therefore, our study provided a more precise estimation of the role of HCV for RCC. The association was most strong (RR = 1.71) in USA subgroup analyses seemed to be intriguing. As we know, an estimated 4.1 million individuals in the USA have been exposed to HCV²⁰. Besides, RCC incidence has increased, particularly among African Americans²¹. Therefore, it seems to be reasonable to consider screening newly diagnosed RCC for HCV infection in the USA.

Although the mechanism by which HCV increased the risk of RCC is not completely understood, the most important explanation is that HCV-associated chronic kidney disease (CKD) might play an important role. HCV infection was found to be associated with developing CKD and end-stage renal disease²². In addition, several other hypotheses have been proposed. First, the HCV virus core protein and the NY-REN-54 protein contribute to the influence of HCV on RCC²³. Second, serine protease inhibitor Kazal (SPIK) protein can inhibit serine protease-related apoptosis, and HCV was found that could increase the expression of SPIK²⁴. Third, cytotoxic T-cell-dependent apoptosis plays a pilot role in host immunity and normal tissue. HCV can disturb this process and lead to renal oncogenesis²⁵. In conclusion, our evidence supported the promoting role of HCV on RCC, and the mechanism needs to be elucidated in future studies.

For HCV and PCa, there was no significant association according to our overall meta-analysis (Table 3). Up to now, few explanations about this issue was reported. Amin et al. ¹⁹ and Lee et al. ²⁶ were the first to report an association between HCV and PCa incidence, HCV and PCa mortality, respectively. However, both did not discuss the potential reasons. Krystyna et al. ²⁷ reported that HCV infection increased the risk of PCa because they believed more serologic testing would be done if their physician suspected cancer. On the contrary, Mahale et al. ¹² reported an inverse association between HCV and PCa. They concluded that those HCV-infected individuals often come from lower socioeconomic status groups, so the rate of prostate cancer screening is relatively low. As we know, controlling confounding factors well was methodologically challenged in studies which tried to conclude the modifiable risk factors for cancers because the carcinogenesis of cancers was multifactorial. In fact, there will be more meaningful to explore the HCV based on PCa grade or stage, because many of PCa were not clinically relevant. In summary, we did not find that HCV was associated with PCa. Although some subgroup analyses showed an inverse association, further prospective, large-scale, long follow-up, and well-controlled confounders cohort studies are needed to investigate this association.

At the design stage of this study, we assumed that HCV infection might be positively associated with the development of bladder cancer as there were some reported molecular biological processes that might be involved²⁸. However, there was no significant association found between them according to this meta-analysis results (Table 4). As an important urine storage organ of the human body, the bladder has been in contact with a large amount of stationary and non-flowing urine for a long time, so it has also become a vital target organ for many carcinogens²⁹. Several other viruses have also been shown to be significantly associated with bladder cancer, and the possible mechanism is that the genetic material of the virus, such as RNA, travels through the blood into the urine and affects the transitional cells of the bladder^{30–32}. The reasons for the failure to find a significant association between HCV and bladder cancer in this study may be varied, and further validation of this conclusion may require epidemiological investigation with larger sample size, mainly covering the high-risk population of HCV infection.

Several potential limitations should not be ignored. First, the heterogeneity in meta-analysis for RCC and PCa was high. On the one hand, random-effects models were applied for making results to be conservative. On the other hand, we explored the potential clinical factors which might contribute to heterogeneity by subgroup analysis methods. Nevertheless, the residual heterogeneity could not be interpreted sufficiently. Second, the small number of studies included in the meta-analysis for BC was a limitation. Third, this study was limited by pooling a few low NOS score (<7) studies. Lastly, the absence of stratification on grade or stage in these cancers (RCC, PCa, and BC) was a limitation that prevented us from better investigating the clinical significance of HCV on these cancers.

Conclusion

With or without sensitivity analyses, the summary estimates from our meta-analysis demonstrated that HCV infection was significantly associated with increased RCC risk, especially enrolling studies in USA locations. Although an inverse association was seen for HCV and PCa risk, it was not statistically significant. There was no significant association between HCV infection and BC risk. Prospective, large-scale, and well-designed cohort studies are needed to validate the association between HCV and RCC and to investigate the role of HCV on PCa.

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Author contributions

MY: project development, data collection and management, manuscript writing and revising; JZ: data collection, data analysis; LH: manuscript editing and revising; HZ: project design and development, data interpretation, manuscript editing and revising. All authors read and approved the final manuscript. WX: project design and development, data interpretation, manuscript editing and revising. All authors read and approved the final manuscript.

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Competing interests

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Correspondence and requests for materials should be addressed to X.W.

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