

RESEARCH ARTICLE

The association between soluble suppression of tumorigenicity-2 and long-term prognosis in patients with coronary artery disease: A meta-analysis

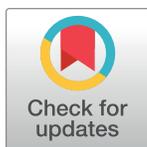
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Abstract

Objective

Findings regarding the prognostic value of soluble suppression of tumorigenicity-2 (sST2) in patients with coronary artery disease (CAD) remain inconsistent. Therefore, we conducted this meta-analysis to investigate the long-term prognostic value of sST2 in patients with CAD.

Methods

A comprehensive literature search was conducted across the PubMed, Embase, and Cochrane Library databases up to June 3, 2020. The primary outcome was major adverse cardiac events (MACEs). The secondary outcomes were all-cause mortality, cardiovascular (CV) death, heart failure (HF), and myocardial infarction (MI). Pooled estimations and 95% confidence intervals (CIs) were assessed using a random-effects model.

Results

Twenty-two articles that enrolled a total of 17,432 patients with CAD were included in the final analysis. CAD patients in the highest categories of baseline sST2 had a significantly higher risk of MACEs (HR: 1.42, 95% CI: 1.09–1.76), all-cause mortality (HR: 2.00, 95% CI: 1.54–2.46), and CV death (HR: 1.42, 95% CI: 1.15–1.68), HF (HR: 2.41, 95% CI: 1.87–2.94), but not that of MI (HR: 1.15, 95% CI: -0.73–3.04), than those in the lowest categories. These results were consistent when baseline sST2 was presented as continuous values in one unit increments. Moreover, subgroup analysis showed that elevated baseline sST2 levels increased the long-term risk of MACEs in the acute coronary syndrome (ACS) population (HR: 1.74, 95% CI: 1.39–2.09) but only showed a trend toward higher risk of MACEs in the non-ACS population (HR: 1.09, 95% CI: 0.87–1.30).

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Conclusions

The findings suggest that a higher concentration of baseline sST2 is associated with a higher risk of MACEs, all-cause mortality, CV death, and HF in patients with CAD. Elevated sST2 levels could significantly predict future MACEs in the ACS population but not in the non-ACS population.

Introduction

Coronary artery disease (CAD), including chronic coronary syndromes (CCS) and acute coronary syndromes (ACS), is one of the leading causes of disability and mortality worldwide [1]. The management of CAD with reperfusion strategies and pharmacological treatment have successfully decreased the incidence of major adverse cardiac events (MACEs) and all-cause mortality [2]. However, patients with CAD still have substantial risk of future cardiovascular events after the onset of the initial episode [3]. Hence, individualized risk prediction in patients with CAD has been applied to identify those with poor prognosis and to optimize cardiovascular disease prevention in high-risk patients [4]. In recent decades, there has been increasing interest in the emerging use of biomarkers as predictors for the prognosis of CAD [5].

Soluble suppression of tumorigenicity-2 (sST2), a member of the interleukin (IL)-1 receptor family, is a decoy receptor of IL-33. sST2 may be released by vascular endothelial and myocardial cells in response to cardiomyocyte biomechanical strain [6]. Excessive circulating sST2 may neutralize the protective effect of IL-33 [7, 8]. sST2 is a promising biomarker for heart failure (HF) compared to natriuretic peptides, especially as sST2 concentrations are less influenced by potential confounders, including age, sex, body mass index, and comorbidities [9, 10]. Moreover, the sST2/IL33 pathway is suggested to be involved in lipid metabolism, and increased IL-33 expression has been detected in human atherosclerotic plaques [11]. Since 2004, epidemiological studies have reported inconsistent results regarding the long-term prognostic value of sST2 in patients with CAD [12, 13]. Therefore, we conducted this meta-analysis to evaluate the relationship between sST2 and long-term clinical outcomes (e.g., MACEs, all-cause mortality, cardiovascular death [CV death], HF, and myocardial infarction [MI] events) in patients with CAD.

Materials and methods

Search strategy

The meta-analysis was conducted according to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses statement [14]. A comprehensive literature search was performed across the PubMed, Embase, and Cochrane Library databases up to June 3, 2020, to identify studies that examined the association between baseline sST2 levels and long-term clinical outcomes using the following terms: ("suppression of tumorigenicity-2" OR "soluble suppression of tumorigenicity-2" OR "ST2" OR "sST2") AND ("coronary heart disease" OR "coronary arterial disease" OR "cardiovascular diseases" OR "coronary disease" OR "myocardial infarction" OR "acute coronary syndrome" OR "stable coronary artery disease"). The detailed search strategy is presented in [S1 Table](#). Furthermore, we also searched the list of references for all related publications to see whether there is other eligible literatures. Two researchers independently searched the database and reviewed all the retrieved papers. The disagreement was resolved by consulting a third researcher. The inclusion criteria were as follows: 1) cohort studies,

including post hoc analyses of randomized clinical trials; 2) studies published in English; 3) studies in which sST2 values related to long-term (>9 months) clinical outcomes in patients with CAD were reported; and 4) studies in which hazard ratio (HR), relative risk (RR), or odds ratio (OR) estimates with the corresponding 95% confidence intervals (CI) for MACEs, all-cause mortality, CV death, HF, or MI events were reported or could be calculated. The exclusion criteria were as follows: 1) studies in which sST2 considered only as an element of a prognostic score and 2) studies with a duplicated population.

Data extraction

The following information was extracted from each eligible study by two authors independently: surname name of the first author, publication year, diagnosis of the patients, sample size, age, number of males, cutoff point of sST2 levels, outcomes evaluated, follow-up duration, and variables adjusted in the multivariate model. When sST2 levels were presented as continuous variables, we standardized the group-level exposure estimates to single units, thereby allowing for the integration of the effects of different sST2 values in different studies.

Definitions of outcomes and quality assessments

The primary outcome in this study was MACEs, which was defined as a composite outcome of death and other fatal or non-fatal cardiovascular events. Given the heterogeneity of MACEs definitions in the included studies, the potential cardiovascular events were distinct, such as MI, HF, stroke, rehospitalization, and revascularization, were included. The secondary outcomes were all-cause mortality, CV death, HF, and MI. A dedicated tool [15, 16] designed for prognostic study assessment was adopted to evaluate the quality of the included publications. This tool consists of five methodological domains to assess for bias: study participation, prognostic factor measurement, outcome measurement, description and measurement of confounding factors, and statistical analysis and reporting. After assessment, every methodological domain was labeled as good, adequate, or unclear for each study.

Statistical analysis

Statistical analyses were performed using STATA software (version 12.0; StatCorp., College Station, TX, USA), and $P < 0.05$ was considered statistically significant. Multivariate adjusted HRs/RRs/ORs with 95% CI for each study were extracted and analyzed. For studies using the sST2 level as the categorical variable, the values of the effects for comparing the highest and lowest categories of sST2 level were extracted. The inverse variance-weighted mean of the logarithm of HR/RRs/ORs and 95% CI with a random effect model was used to calculate the association between sST2 levels and different clinical outcomes among CAD patients.

The I^2 and Cochrane's Q tests were adopted to evaluate the heterogeneity among different studies [17, 18]. Significant heterogeneity was considered if $I^2 > 50\%$ [19]. Meta-regression was performed to explore the study and population characteristics that may contribute to heterogeneity [20]. Publication bias was evaluated by funnel plots using the Egger regression asymmetry test [21].

Results

Literature search and study characteristics

The literature search process is summarized in Fig 1. We identified 159 articles from PubMed, 1,585 from Embase, and 41 from the Cochrane Library. Among them, 22 articles that enrolled a total of 17,432 patients with CAD investigated the relationship between baseline sST2 levels

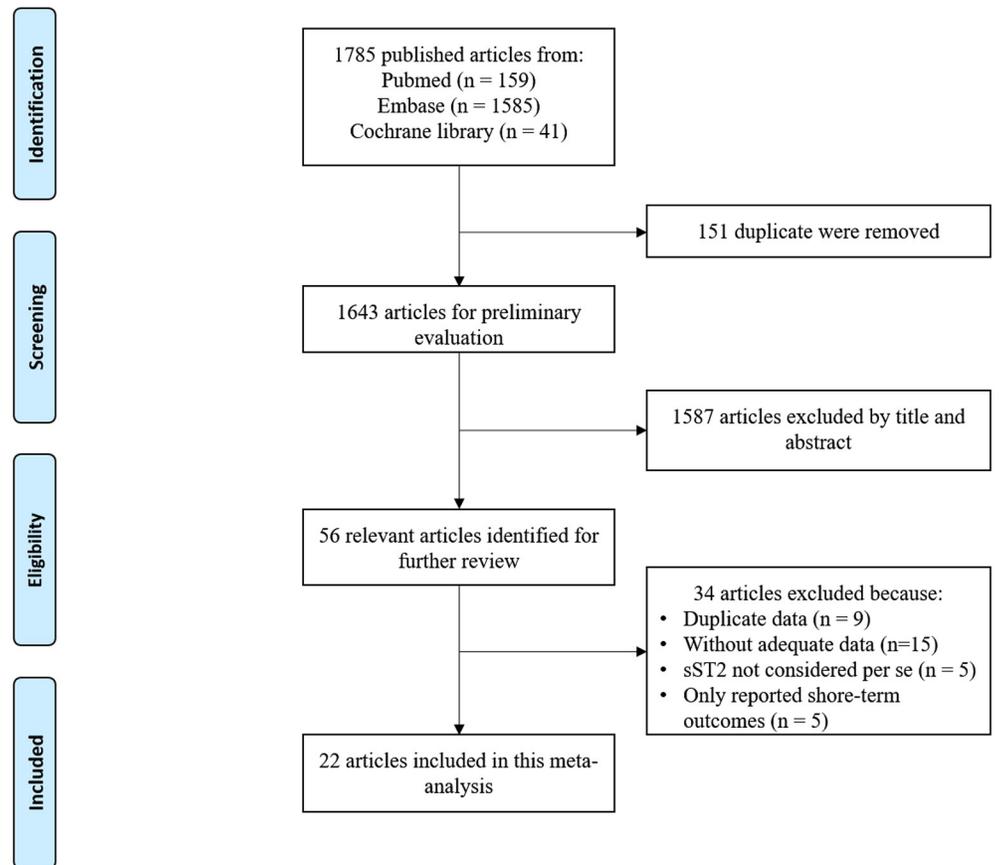


Fig 1. Flow diagram of screened, excluded, and analyzed studies.

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and long-term clinical outcomes in patients with CAD were included [12, 13, 22–41]. In these studies, the baseline sST2 level was provided as a categorical variable (18 articles) or a continuous variable (8 articles). The main characteristics of these studies are summarized in Tables 1 and 2. Regarding the types of CAD, 17 studies were conducted on the ACS population and 5 were conducted on the non-ACS population. MACEs as endpoints were reported in 13 articles, followed by all-cause mortality in 12, CV death in 5, HF in 4 articles, and MI in 2. The duration of follow-up in the included studies varied from 9 months to 13 years. Various confounding factors such as age, sex, medical history, comorbidities, biochemical parameters, and medication were adjusted when presenting the HRs/RRs/ORs in the included studies.

sST2 and long-term clinical outcomes in patients with CAD

CAD patients in the highest categories of baseline sST2 had significantly higher risk of MACEs (HR: 1.42, 95% CI: 1.09–1.76, $I^2 = 37.8\%$) (Fig 2), all-cause mortality (HR: 2.00, 95% CI: 1.54–2.46, $I^2 = 60.9\%$), and CV death (HR: 1.42, 95% CI: 1.15–1.68, $I^2 = 0$), HF (HR: 2.41, 95% CI: 1.87–2.94, $I^2 = 0$), but not of MI (HR: 1.15, 95% CI: -0.73–3.04, $I^2 = 66.3\%$), than those in the lowest categories (Fig 3). The results were consistent when baseline sST2 was presented as continuous with 1-unit increments, with HRs of 1.16 (95% CI: 1.01–1.31, $I^2 = 75.5\%$) in MACEs (Fig 4), 1.56 (95% CI: 1.18–1.95, $I^2 = 88.8\%$) in all-cause mortality, and 1.21 (95% CI: 1.06–1.37, $I^2 = 32.3\%$) in CV death (Fig 5).

Table 1. Characteristics of studies on long-term prognosis value of baseline sST2 as categorical variable in patients with CAD.

Study	Country	Study design	Types of CAD	Patients, n	Age, Mean/Median, years	Males, n	Cut-off value	Endpoints	Follow-up duration	Variables in multivariate model
Dhillon 2011 [22]	UK	Retrospective cohort	NSTEMI	577	70 ± 13	397	638 pg/mL	MACEs (defined as the composite of all-cause mortality, HF hospitalization, and reinfarction), All-cause mortality, HF, MI	532 days (range 150–1059 days)	age, gender, previous angina or AMI, HF, HT, DM, Killip class, eGFR, log10 glucose level, therapy with β -blockers, or statins. log10 troponin I, previous history of hypercholesterolemia, or smoking
Kohli 2012 [23]	Multinational	RCT post-hoc	NSTEMI	4426	NA	2862	24.2 ug/mL	MACEs (defined as the composite of cardiovascular death and new or worsening HF), All-cause mortality, CV death, HF	1 year	TIMI risk score smoking, history of HF, creatinine clearance <60mL/min, cTnI, BNP, myeloperoxidase (MPO), and high-sensitivity C-reactive protein (hsCRP).
Dhillon 2013 [24]	UK	Retrospective cohort	Unselected STEMI	667	64.0 ± 12.2	505	677 pg/mL	All-cause mortality, HF, MI	median 587 days (range 134–2818 days)	age, gender, previous history of angina/AMI, HT, DM, Killip Class, eGFR, peak creatinine phosphokinase level, treatment with thrombolysis, betablockers, statins, ACE inhibitors or ARBs
Demyanets 2014 [25]	Austria	Retrospective cohort	CAD including stable angina, NSTEMI, STEMI	373	64.4 ± 12.1	247	169 pg/mL for SA, 453 pg/mL for STEMI, and 269 pg/mL for NSTEMI	All-cause mortality	mean 43 months	age, sex, HT, smoker, hyperlipidaemia, DM, BMI, serum creatinine
Dieplinger 2014 [26]	Germany	Prospective cohort	Stable CAD	1345	65 (58–71)	1008	19.4 ng/mL	CV death	a median of 9.8 years	sex, age, BMI, HT, dyslipidemia, DM, smoking, prior MI, peripheral artery disease, cerebrovascular disease, resting heart rate, LVEF, eGFR, Cholinesterase, hs-CRP, IL-6, NT-proBNP, hs-cTnT, Galectin-3
Minamisawa 2016 [27]	Japan	Prospective cohort	Prior MI	430	66 ± 12	343	NA	MACEs (defined as the composite of all-cause death, MI, stroke, or hospitalization due to congestive HF)	3 years	age, chronic kidney disease, BNP, GDF-15

(Continued)

Table 1. (Continued)

Study	Country	Study design	Types of CAD	Patients, n	Age, Mean/Median, years	Males, n	Cut-off value	Endpoints	Follow-up duration	Variables in multivariate model
Jenkins 2017 [28]	US	Prospective cohort	Prior MI	1401	67.3 ± 14.9	854	The highest tertile versus lowest tertile	All-cause mortality, HF	3.6 ± 1.8 years	age, sex, Charlson comorbidity index, Killip class, and maximum troponin T
Pfetsch 2017 [29]	Germany	Prospective cohort	Stable CAD	1081	58.9 ± 8.0	915	The highest quartile versus lowest quartile	MACEs (defined as the composite of fatal and non-fatal cardiovascular events), All-cause mortality, CV death	13 years	age, gender, school education, rehabilitation clinic, smoking status, history of DM, PCI, left ventricular function, HDL-cholesterol, LDL-cholesterol
Yu 2017 [30]	South Korea	Prospective cohort	STEMI undergoing PCI	323	59.1±13.1	272	75.8 ng/mL	MACEs (defined as the composite of cardiovascular death, non-fatal MI, non-fatal stroke, and ischemia-driven revascularization)	1 year	age, LVEF, final TIMI flow grade, NT-proBNP, troponin I, hs-CRP, glucose level
Liu 2018 [31]	China	Prospective cohort	STEMI receiving PCI	305	NA	243	32.2 ng/mL	MACEs (defined as the composite of all-cause death, heart failure, and nonfatal myocardial infarction), All-cause mortality	1 year	age, sex, time from onset to ER, Killip class, LVEF, NT-proBNP, cTnI-peak, hs-CRP
Lepojarvi 2018 [32]	Finland	Prospective cohort	Stable CAD	1946	NA	NA	21.5 ng/mL	All-cause mortality	76 ± 20 months	age, sex, BMI, CCS grading of angina pectoris, LVEF, eGFR, albumin-creatinine-ratio, glycated hemoglobin, and diabetes
Huang 2018 [33]	China	Prospective cohort	Unselected STEMI	186	68.5 (32–72)	119	56 ng/mL	MACEs (defined as the composite of cardiovascular death, worsening HF and recurrent MI)	1 year	SBP, Log NT-proBNP, Final low BB status
Heydari 2018 [34]	US	RCT post-hoc	AMI	317	NA	NA	35 ng/mL	MACEs (defined as the composite of cardiovascular death and hospitalization for ADHF)	3 years	baseline and cardiac magnetic resonance variables
Jha 2018 [35]	India	Prospective cohort	ACS	122	NA	90	36.5 ng/mL	MACEs (defined as the composite of unstable angina/NSTEMI, stroke, recurrent MI, re-hospitalization, and death)	9 months	age, gender, types of ACS, LVEF, DM, HT, and eGFR

(Continued)

Table 1. (Continued)

Study	Country	Study design	Types of CAD	Patients, n	Age, Mean/Median, years	Males, n	Cut-off value	Endpoints	Follow-up duration	Variables in multivariate model
Jacobs 2018 [36]	US	Prospective cohort	CAD receiving CABG	1047	NA	801	3.8 ng/mL	MACEs (defined as the composite of mortality or readmission)	1 year	The Society of Thoracic Surgeons 30-day readmission model
Zagidullin 2020 [13]	Russia	Prospective cohort	STEMI receiving PCI	156	NA	118	27.2 ng/mL	CV death	734.2 ± 61.2 days	Age, gender, hs-Troponin, LVEF
Kim 2020 [12]	South Korea	Retrospective cohort	STEMI receiving PCI	184	61.4 ± 11.8	156	The highest tertile versus lowest tertile	MACEs (defined as the composite of cardiovascular death, non-fatal MI, and non-fatal stroke)	1 year	SBP, symptom to door time, TIMI risk score, and CACS score
Somuncu 2020 [37]	Turkey	Prospective cohort	STEMI receiving PCI	380	NA	279	35 ng/mL	CV death	1 year	age, DM, IRA, Killip >1, SBP, LVEF

Data was presented as n, mean with SD or median with IQR.

Abbreviations: CAD: coronary artery disease; STEMI: ST-segment elevation myocardial infarction; MACEs: major cardiac adverse events; PCI: percutaneous coronary intervention; MI: Myocardial infarction; NSTEMI: non-ST elevation myocardial infarction; LVEF: left ventricular ejection fraction; hs-CRP: High sensitivity C protein; BMI: body mass index; eGFR: estimated glomerulus filtration rate; DM: diabetes mellitus; HT: hypertension; HF: Heart failure; TIMI: Thrombolysis in myocardial infarction; NT-proBNP: N-terminal pro-B-type natriuretic peptide; hs-cTnT: high-sensitivity cardiac troponin T; AMI: acute myocardial infarction; ADHF: acute decompensated heart failure; SBP: systolic blood pressure; CCS: Canadian cardiovascular society; BB: Beta blocker.

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Source of heterogeneity and subgroup analysis

There was medium ($I^2 = 37.8\%$) and substantial ($I^2 = 75.5\%$) heterogeneity for the association of sST2 with the primary outcome when sST2 was used as a categorical variable and a continuous variable, respectively. We conducted meta-regression with geographic locations, number of participants, study design, or types of CAD to investigate the potential sources of heterogeneity. The results suggested that the type of CAD might be the main source of heterogeneity ($P = 0.009$ for CAD patients with sST2 as categorical variables; $P = 0.574$ for those with sST2 as continuous variables) (S1 Fig). The subgroup analysis stratified by type of CAD is summarized in Table 3. For the primary outcome, the bias risks were similar in the ACS group (HR: 1.74, 95% CI: 1.39–2.09, $I^2 = 0$; S2 Fig) and the non-ACS group (HR: 1.09, 95% CI: 0.87–1.30, $I^2 = 9.6\%$; S2 Fig) when sST2 was used as a categorical variable. However, the effects in the non-ACS group (HR: 1.12, 95% CI: 1.01–1.23, $I^2 = 0$) was more robust than that in the ACS group (HR: 1.31, 95% CI: 0.89–1.72, $I^2 = 84.2\%$) when sST2 was used as a continuous variable.

Quality assessment

An overview of the methodological quality of all included studies is presented in S2 Table. Nine studies [12, 13, 23, 25, 29, 30, 32, 34, 38, 39] did not report well-defined details of the methods and results in the follow-up, which might have contributed to possible bias in their outcomes. Additionally, one [30] study had a high risk of population selection and another [38] did not provide definitions of the outcomes. All studies provided adequate descriptions of sST2 measurement, control of confounding factors, and statistical analysis.

Table 2. Characteristics of studies on long-term prognosis value of baseline sST2 as continuous variable in patients with CAD.

Study	Country	Study design	Types of CAD	Patients, n	Age, Mean/Median, years	Males, n	Cut-off value	Endpoints	Follow-up, months or years	Variables in multivariate model
Eggers 2010 [38]	Multinational	RCT post-hoc	NSTEMI	403	69 (59–77)	261	per 1-SD increase	All-cause mortality	1 year	age, gender, HT, congestive HF, DM, hypercholesterolemia, smoking, previous MI, previous stroke, previous coronary revascularization, cTnT, NT-proBNP, CRP, IL-6, fibrinogen, and eGFR
Dieplinger 2014 [26]	Germany	Prospective cohort	Stable CAD	1345	65 (58–71)	1008	per 1-SD increase	CV death	a median of 9.8 years	sex, age, BMI, HT, dyslipidemia, DM, smoking, prior MI, peripheral artery disease, cerebrovascular disease, resting heart rate, LVEF, eGFR, Cholinesterase, hs-CRP, IL-6, NT-proBNP, hs-cTnT, Galectin-3
Pfetsch 2017 [29]	Germany	Prospective cohort	Stable CAD	1081	58.9 ± 8.0	915	per 1 log-unit increase	MACEs (defined as the composite of fatal and non-fatal cardiovascular events), All-cause mortality, CV death	13 years	age, gender, school education, rehabilitation clinic, smoking status, history of DM, PCI, left ventricular function, HDLcholesterol, LDL-cholesterol
Wang 2017 [39]	China	Retrospective cohort	AMI receiving PCI	180	61.4 ± 8.9	110	per 1 unit increase	MACEs (defined as the composite of all-cause death, stent thrombosis, myocardial infarction and target vessel revascularization)	1 year	IL-33, BNP, Gensini score, hs-CRP
Tyminska 2017 [40]	Poland	Prospective cohort	STEMI receiving PCI	117	61 (50.7–67)	82	per 10 units increase	MACEs (defined as the composite of all-cause death, HF, and non-fatal AMI)	1 year	age and NT-proBNP
Liu 2018 [31]	China	Prospective cohort	STEMI receiving PCI	305	NA	243	per 1 unit increase	MACEs (defined as the composite of all-cause death, heart failure, and nonfatal myocardial infarction), All-cause mortality	1 year	age, sex, time from onset to ER, Killip class, LVEF, NT-proBNP, cTnI-peak, hs-CRP
Gerber 2018 [41]	US	Prospective cohort	Prior MI	1401	67.3 ± 15	853	per 1 log-unit increase	All-cause mortality	1 year	GRACE score and Charlson comorbidity index
Jacobs 2018 [36]	US	Prospective cohort	CAD receiving CABG	1047	NA	801	per 1 log-unit increase	MACEs (defined as the composite of mortality or readmission)	1 year	The Society of Thoracic Surgeons 30-day readmission model

Data was presented as n, mean with SD or median with IQR. Abbreviations as in Table 1.

<https://doi.org/10.1371/journal.pone.0238775.t002>

Sensitivity analysis and publication bias

Sensitivity analysis showed that no single study had a significant effect on the pooled effect, indicating that our results were statistically robust. Based on the results of Egger's test, when

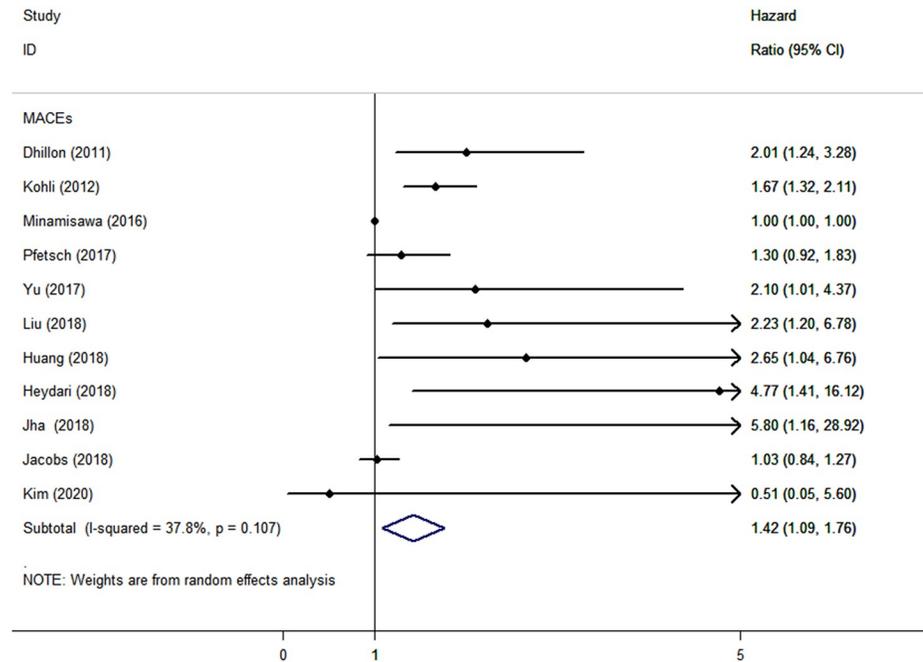


Fig 2. The forest plot between sST2 level as categorical variable and primary outcome in patients with CAD.

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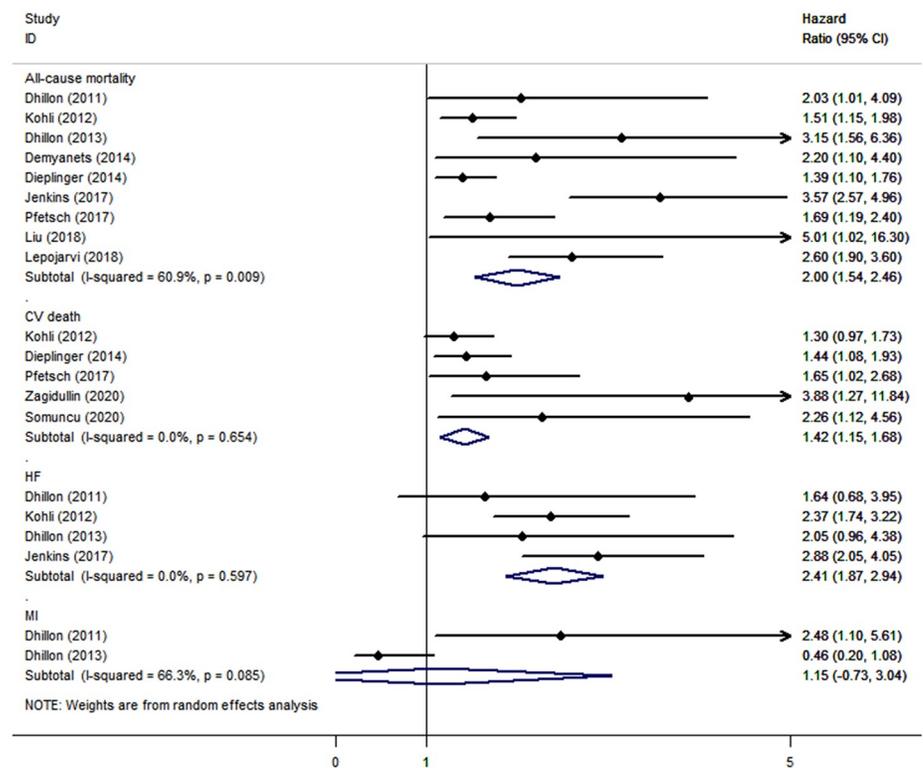


Fig 3. The forest plot between sST2 level as continuous variable and primary outcome in patients with CAD.

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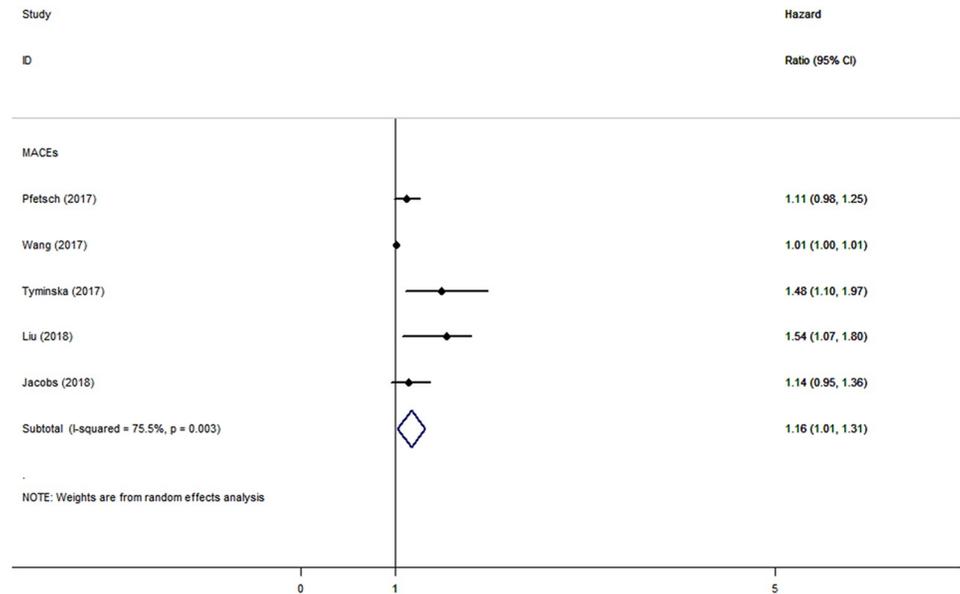


Fig 4. The forest plot between sST2 level as categorical variable and secondary outcomes in patients with CAD.

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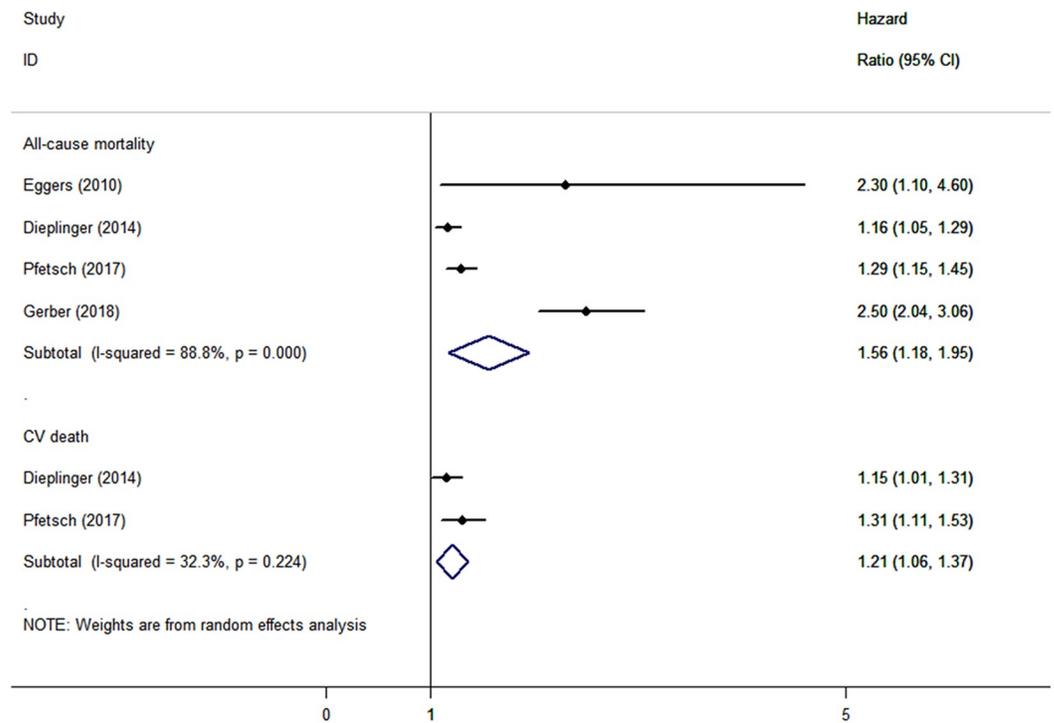


Fig 5. The forest plot between sST2 level as continuous variable and secondary outcomes in patients with CAD.

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Table 3. Subgroup analysis stratified by types of CAD.

Outcomes	ACS group		Non-ACS group	
	HR (95% CI)	I ² /%	HR (95% CI)	I ² /%
Categorical variable	(n = 6870)		(n = 2128)	
MACEs	1.74 (1.39–2.09)	0	1.09 (0.87–1.30)	9.6
All-cause mortality	2.48 (1.32–3.63)	67.4	1.82 (1.27–2.37)	59.8
CV death	1.37 (0.95–1.79)	1.6	1.48 (1.11–1.86)	0
HF	2.41 (1.87–2.94)	0	-	-
MI	1.16 (-0.73–3.04)	66.3	-	-
Continuous variable	(n = 602)		(n = 2128)	
MACEs	1.31 (0.89–1.72)	84.2	1.12 (1.01–1.23)	0
All-cause mortality	2.48 (1.99–2.97)	0	1.22 (1.09–1.34)	43.2
CV death	1.21 (1.06–1.37)	32.2	-	-
HF	-	-	-	-
MI	-	-	-	-

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Sst2 was used as a categorization variable, there was no significant publication bias for the primary outcome ($P = 0.059$). However, a combination of visual assessment of funnel plots (S2 Fig) and Egger's test provided statistical evidence of publication bias for MACEs when sST2 was expressed as a continuous variable ($P = 0.015$). This may indicate that the positive results in CAD patients with sST2 as a continuous variable might be overrepresented, potentially leading to a relative overestimation of the pooled HR in the meta-analysis.

Discussion

In this study, elevated baseline sST2 levels in patients with CAD were positively associated with increased long-term risk of MACEs, all-cause mortality, CV death, and HF, but not of MI events. Moreover, subgroup analysis showed that the predictive role of sST2 for long-term prognosis in the ACS population was consistent with the main results, while the prognostic value of sST2 was not verified in the non-ACS population. This study is the largest meta-analysis to assess the relationship between sST2 and long-term prognosis in patients with CAD, which allows a much greater possibility of reaching reliable conclusions.

sST2 is a novel biomarker in inflammatory conditions and cardiovascular disease [6]. Over the past decade, several meta-analyses have suggested that sST2 is a prognostic indicator of aortic valve replacement [42], acute HF [9], chronic HF [10], and pulmonary hypertension [43]. Gu et al. also demonstrated that a higher circulating baseline sST2 level could predict poor clinical outcomes in patients with ACS in a meta-analysis that included 12 studies [44]. Our updated results with 17 articles involving patients with ACS are consistent with the findings of the above study, which confirmed the positive association between sST2 and increased risk of adverse events in the long-term prognosis of patients with ACS. Nonetheless, the study conducted by Gu et al. may have classified CV death events as all-cause mortality in some studies, which might bias their estimations [23, 26, 29]. Moreover, our study adds new evidence for the prognostic value of sST2 in the non-ACS population. The elevated baseline sST2 levels showed a trend toward an increased long-term risk of MACEs, all-cause mortality, and CV death in these patients, although without statistical significance.

Heterogeneity is common in meta-analyses. Therefore, it is an essential component of a meta-analysis to explore the sources of heterogeneity. In this study, we used meta-regression to explore the sources of heterogeneity. When sST2 was expressed as a categorical variable, the results showed that the type of CAD (ACS or non-ACS) might be the main source of

heterogeneity for the primary outcome. In the subsequent subgroup analysis, the bias risks were similar in the ACS and non-ACS groups. Additionally, significant publication bias was detected in studies using sST2 as a continuous variable. This result might be ascribed to the fact that relatively few studies adopted continuous sST2 values and thus did not report potentially negative results.

The biological mechanism underlying the relationship between sST2 and clinical outcomes in patients with CAD is not fully understood. One plausible pathway is that sST2 leads to adverse clinical outcomes by counteracting the protective effect of IL-33. ST2 is expressed by cardiac myocytes in two forms: the soluble subtype (sST2) and the membrane-bound subtype (ST2L) [45]. Its ligand, IL-33, which is released by cardiac cells, normally protects the myocardium under pressure overload [46]. sST2 neutralizes this protective mechanism by acting as a "decoy receptor" for IL-33. When bound with sST2, IL-33 can no longer interact with membrane-ST2L to prevent the activation of downstream signaling pathways, leading to cardiomyocyte hypertrophy and remodeling. The indirect resistance to IL-33 might activate adverse cardiac hypertrophy and remodeling through lymphocyte function [47]. Another possible mechanism linking the ST2/IL-33 axis to CAD outcomes is its correlation with atherosclerosis progression and plaque destabilization [48]. It has been speculated that IL-33 reduces the risk of plaque rupture by inhibiting the immune response involved in atherosclerosis, an effect that can be reversed by sST2. Additionally, the IL-33/ST2 system might be a potential pathophysiological mediator in the inflammatory and fibrotic response to tissue injury and has been identified in inflammatory conditions and fibroproliferative diseases [11]. Accordingly, elevated serum sST2 levels in patients with CAD suggest the presence of active inflammation and fibrosis, factors that have been associated with poor prognosis [49, 50]. More importantly, the differential predictive value of sST2 on the occurrence of HF and MI in patients with CAD further suggests that sST2 is predominantly involved in the pathophysiology of HF and independent of CAD progression per se.

Our study has important strengths. First, multivariate-adjusted results were available from all included studies, reducing the likelihood of confounders. Second, all results were derived from cohorts or randomized controlled studies, which would enable us to detect the temporal relationship between sST2 and prognosis in patients with CAD. Nevertheless, several limitations of this study should be acknowledged. First, the heterogeneous definitions of MACEs in different studies might have contributed to potential bias. Second, some degree of heterogeneity was observed, which could be due to different types of CAD. Third, the use of different cutoff values for sST2 may lead to different estimated effects. However, due to the limited information available, we could not address the optimum cutoff value of sST2 for the prognosis of patients with CAD. Further studies with larger sample sizes are warranted to address this issue.

Conclusions

This meta-analysis suggests that elevated baseline sST2 level is associated with a higher risk of MACEs, all-cause mortality, CV death, and HF in patients with CAD. Elevated sST2 levels could significantly predict future MACEs in the ACS population but not in the non-ACS population.

Supporting information

S1 Table. Keywords and search strategy in the umbrella review.
(DOCX)

S2 Table. Quality assessment of studies included in the meta-analysis.
(DOCX)

S1 Fig. Results of meta-regression analysis for MACEs among ACS and non-ACS in patients with CAD.
(DOCX)

S2 Fig. Funnel plot for publication bias for primary outcome in patients with CAD.
(DOCX)

S1 Checklist. PRISMA 2009 checklist.
(DOC)

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References

1. Moran AE, Forouzanfar MH, Roth GA, Mensah GA, Ezzati M, Murray CJ, et al. Temporal trends in ischemic heart disease mortality in 21 world regions, 1980 to 2010: the Global Burden of Disease 2010 study. *Circulation*. 2014; 129(14):1483–92. Epub 2014/02/28. <https://doi.org/10.1161/CIRCULATIONAHA.113.004042> PMID: 24573352.
2. Jernberg T, Johanson P, Held C, Svennblad B, Lindback J, Wallentin L. Association between adoption of evidence-based treatment and survival for patients with ST-elevation myocardial infarction. *Jama*. 2011; 305(16):1677–84. Epub 2011/04/28. <https://doi.org/10.1001/jama.2011.522> PMID: 21521849.
3. Hilvo M, Meikle PJ, Pedersen ER, Tell GS, Dhar I, Brenner H, et al. Development and validation of a ceramide- and phospholipid-based cardiovascular risk estimation score for coronary artery disease patients. *European heart journal*. 2020; 41(3):371–80. Epub 2019/06/19. <https://doi.org/10.1093/eurheartj/ehz387> PMID: 31209498.
4. Al'Aref SJ, Maliakal G, Singh G, van Rosendaal AR, Ma X, Xu Z, et al. Machine learning of clinical variables and coronary artery calcium scoring for the prediction of obstructive coronary artery disease on coronary computed tomography angiography: analysis from the CONFIRM registry. *European heart journal*. 2020; 41(3):359–67. Epub 2019/09/13. <https://doi.org/10.1093/eurheartj/ehz565> PMID: 31513271.
5. Trindade F, Vitorino R, Leite-Moreira A, Falcão-Pires I. Pericardial fluid: an underrated molecular library of heart conditions and a potential vehicle for cardiac therapy. *Basic research in cardiology*. 2019; 114(2):10. Epub 2019/01/20. <https://doi.org/10.1007/s00395-019-0716-3> PMID: 30659359.
6. Altara R, Ghali R, Mallat Z, Cataliotti A, Booz GW, Zouein FA. Conflicting vascular and metabolic impact of the IL-33/sST2 axis. *Cardiovascular research*. 2018; 114(12):1578–94. Epub 2018/07/10. <https://doi.org/10.1093/cvr/cvy166> PMID: 29982301.
7. Gaggin HK, Szymonifka J, Bhardwaj A, Belcher A, De Berardinis B, Motiwala S, et al. Head-to-head comparison of serial soluble ST2, growth differentiation factor-15, and highly-sensitive troponin T measurements in patients with chronic heart failure. *JACC Heart failure*. 2014; 2(1):65–72. Epub 2014/03/14. <https://doi.org/10.1016/j.jchf.2013.10.005> PMID: 24622120.
8. Anand IS, Rector TS, Kuskowski M, Snider J, Cohn JN. Prognostic value of soluble ST2 in the Valsartan Heart Failure Trial. *Circulation Heart failure*. 2014; 7(3):418–26. Epub 2014/03/14. <https://doi.org/10.1161/CIRCHEARTFAILURE.113.001036> PMID: 24622243.

9. Aimo A, Vergaro G, Ripoli A, Bayes-Genis A, Pascual Figal DA, de Boer RA, et al. Meta-Analysis of Soluble Suppression of Tumorigenicity-2 and Prognosis in Acute Heart Failure. *JACC Heart failure*. 2017; 5(4):287–96. Epub 2017/02/13. <https://doi.org/10.1016/j.jchf.2016.12.016> PMID: 28189578.
10. Aimo A, Vergaro G, Passino C, Ripoli A, Ky B, Miller WL, et al. Prognostic Value of Soluble Suppression of Tumorigenicity-2 in Chronic Heart Failure: A Meta-Analysis. *JACC Heart failure*. 2017; 5(4):280–6. Epub 2016/11/07. <https://doi.org/10.1016/j.jchf.2016.09.010> PMID: 27816512.
11. Kakkar R, Lee RT. The IL-33/ST2 pathway: therapeutic target and novel biomarker. *Nature reviews Drug discovery*. 2008; 7(10):827–40. Epub 2008/10/02. <https://doi.org/10.1038/nrd2660> PMID: 18827826.
12. Kim M, Lee DI, Lee JH, Kim SM, Lee SY, Hwang KK, et al. Lack of prognostic significance for major adverse cardiac events of soluble suppression of tumorigenicity 2 levels in patients with ST-segment elevation myocardial infarction. *Cardiology journal*. 2020. Epub 2020/02/28. <https://doi.org/10.5603/CJ.a2020.0028> PMID: 32104902.
13. Zagidullin N, Motloch LJ, Gareeva D, Hamitova A, Lakman I, Krioni I, et al. Combining Novel Biomarkers for Risk Stratification of Two-Year Cardiovascular Mortality in Patients with ST-Elevation Myocardial Infarction. *Journal of clinical medicine*. 2020; 9(2). Epub 2020/02/23. <https://doi.org/10.3390/jcm9020550> PMID: 32085400.
14. Moher D, Liberati A, Tetzlaff J, Altman DG. Preferred reporting items for systematic reviews and meta-analyses: the PRISMA statement. *PLoS medicine*. 2009; 6(7):e1000097. Epub 2009/07/22. <https://doi.org/10.1371/journal.pmed.1000097> PMID: 19621072.
15. Hayden JA, van der Windt DA, Cartwright JL, Cote P, Bombardier C. Assessing bias in studies of prognostic factors. *Annals of internal medicine*. 2013; 158(4):280–6. Epub 2013/02/20. <https://doi.org/10.7326/0003-4819-158-4-201302190-00009> PMID: 23420236.
16. Hayden JA, Cote P, Bombardier C. Evaluation of the quality of prognosis studies in systematic reviews. *Annals of internal medicine*. 2006; 144(6):427–37. Epub 2006/03/22. <https://doi.org/10.7326/0003-4819-144-6-200603210-00010> PMID: 16549855.
17. Higgins J, Green S. *Cochrane Handbook for Systematic Reviews of Interventions* Version 5.1.0. The Cochrane Collaboration. 2011; www.cochranehandbook.org.
18. Higgins JP, Thompson SG. Quantifying heterogeneity in a meta-analysis. *Statistics in medicine*. 2002; 21(11):1539–58. Epub 2002/07/12. <https://doi.org/10.1002/sim.1186> PMID: 12111919.
19. Higgins JP, Thompson SG, Deeks JJ, Altman DG. Measuring inconsistency in meta-analyses. *BMJ (Clinical research ed)*. 2003; 327(7414):557–60. Epub 2003/09/06. <https://doi.org/10.1136/bmj.327.7414.557> PMID: 12958120.
20. Higgins JP, Thompson SG. Controlling the risk of spurious findings from meta-regression. *Statistics in medicine*. 2004; 23(11):1663–82. Epub 2004/05/26. <https://doi.org/10.1002/sim.1752> PMID: 15160401.
21. Egger M, Davey Smith G, Schneider M, Minder C. Bias in meta-analysis detected by a simple, graphical test. *BMJ (Clinical research ed)*. 1997; 315(7109):629–34. Epub 1997/10/06. <https://doi.org/10.1136/bmj.315.7109.629> PMID: 9310563.
22. Dhillon OS, Narayan HK, Quinn PA, Squire IB, Davies JE, Ng LL. Interleukin 33 and ST2 in non-ST-elevation myocardial infarction: comparison with Global Registry of Acute Coronary Events Risk Scoring and NT-proBNP. *American heart journal*. 2011; 161(6):1163–70. Epub 2011/06/07. <https://doi.org/10.1016/j.ahj.2011.03.025> PMID: 21641364.
23. Kohli P, Bonaca MP, Kakkar R, Kudinova AY, Scirica BM, Sabatine MS, et al. Role of ST2 in non-ST-elevation acute coronary syndrome in the MERLIN-TIMI 36 trial. *Clin Chem*. 2012; 58(1):257–66. Epub 2011/11/19. <https://doi.org/10.1373/clinchem.2011.173369> PMID: 22096031.
24. Dhillon OS, Narayan HK, Khan SQ, Kelly D, Quinn PA, Squire IB, et al. Pre-discharge risk stratification in unselected STEMI: is there a role for ST2 or its natural ligand IL-33 when compared with contemporary risk markers? *International journal of cardiology*. 2013; 167(5):2182–8. Epub 2012/07/28. <https://doi.org/10.1016/j.ijcard.2012.05.073> PMID: 22835988.
25. Demyanets S, Speidl WS, Tentzeris I, Jarai R, Katsaros KM, Farhan S, et al. Soluble ST2 and interleukin-33 levels in coronary artery disease: relation to disease activity and adverse outcome. *PloS one*. 2014; 9(4):e95055. Epub 2014/04/23. <https://doi.org/10.1371/journal.pone.0095055> PMID: 24751794.
26. Dieplinger B, Egger M, Haltmayer M, Kleber ME, Scharnagl H, Silbernagel G, et al. Increased soluble ST2 predicts long-term mortality in patients with stable coronary artery disease: results from the Ludwigshafen risk and cardiovascular health study. *Clin Chem*. 2014; 60(3):530–40. <https://doi.org/10.1373/clinchem.2013.209858> PMID: 24401186.
27. Minamisawa M, Motoki H, Izawa A, Kashima Y, Hioki H, Abe N, et al. Comparison of Inflammatory Biomarkers in Outpatients With Prior Myocardial Infarction. *International heart journal*. 2016; 57(1):11–7. Epub 2016/01/09. <https://doi.org/10.1536/ihj.15-197> PMID: 26742699.

28. Jenkins WS, Roger VL, Jaffe AS, Weston SA, AbouEzzeddine OF, Jiang R, et al. Prognostic Value of Soluble ST2 After Myocardial Infarction: A Community Perspective. *The American journal of medicine*. 2017; 130(9):1112.e9-.e15. Epub 2017/03/28. <https://doi.org/10.1016/j.amjmed.2017.02.034> PMID: 28344136.
29. Pfetsch V, Sanin V, Jaensch A, Dallmeier D, Mons U, Brenner H, et al. Increased Plasma Concentrations of Soluble ST2 Independently Predict Mortality but not Cardiovascular Events in Stable Coronary Heart Disease Patients: 13-Year Follow-up of the KAROLA Study. *Cardiovascular drugs and therapy*. 2017; 31(2):167–77. Epub 2017/03/12. <https://doi.org/10.1007/s10557-017-6718-1> PMID: 28283847.
30. Yu J, Oh PC, Kim M, Moon J, Park YM, Lee K, et al. Improved early risk stratification of patients with ST-segment elevation myocardial infarction undergoing primary percutaneous coronary intervention using a combination of serum soluble ST2 and NT-proBNP. *PloS one*. 2017; 12(8):e0182829. Epub 2017/08/11. <https://doi.org/10.1371/journal.pone.0182829> PMID: 28796845.
31. Liu X, Hu Y, Huang W, Zhang G, Cao S, Yan X, et al. Soluble ST2 for Prediction of Clinical Outcomes in Patients with ST-Segment Elevation Myocardial Infarction Receiving Primary PCI. *International heart journal*. 2019; 60(1):19–26. <https://doi.org/10.1536/ihj.18-020> PMID: 30464124.
32. Lepojarvi ES, Huikuri HV, Piira OP, Kiviniemi AM, Miettinen JA, Kentta T, et al. Biomarkers as predictors of sudden cardiac death in coronary artery disease patients with preserved left ventricular function (ARTEMIS study). *PloS one*. 2018; 13(9):e0203363. <https://doi.org/10.1371/journal.pone.0203363> PMID: 30226845.
33. Huang WP, Zheng X, He L, Su X, Liu CW, Wu MX. Role of Soluble ST2 Levels and Beta-Blockers Dosage on Cardiovascular Events of Patients with Unselected ST-Segment Elevation Myocardial Infarction. *Chinese medical journal*. 2018; 131(11):1282–8. Epub 2018/05/23. <https://doi.org/10.4103/0366-6999.232819> PMID: 29786039.
34. Heydari B, Abdullah S, Shah R, Francis SA, Feng JH, McConnell J, et al. Omega-3 Fatty Acids Effect on Post-Myocardial Infarction ST2 Levels for Heart Failure and Myocardial Fibrosis. *J Am Coll Cardiol*. 2018; 72(8):953–5. Epub 2018/08/18. <https://doi.org/10.1016/j.jacc.2018.06.018> PMID: 30115235.
35. Jha D, Goenka L, Ramamoorthy T, Sharma M, Dhandapani VE, George M. Prognostic role of soluble ST2 in acute coronary syndrome with diabetes. *Eur J Clin Invest*. 2018; 48(9):e12994. Epub 2018/07/12. <https://doi.org/10.1111/eci.12994> PMID: 29992539.
36. Jacobs JP, Alam SS, Owens SL, Parker DM, Rezaee M, Likosky DS, et al. The Association Between Novel Biomarkers and 1-Year Readmission or Mortality After Cardiac Surgery. *The Annals of thoracic surgery*. 2018; 106(4):1122–8. Epub 2018/06/05. <https://doi.org/10.1016/j.athoracsur.2018.04.084> PMID: 29864407.
37. Somuncu MU, Kalayci B, Avci A, Akgun T, Karakurt H, Demir AR, et al. Predicting long-term cardiovascular outcomes of patients with acute myocardial infarction using soluble ST2. *Horm Mol Biol Clin Investig*. 2020. Epub 2020/03/01. PMID: 32112700.
38. Eggers KM, Armstrong PW, Califf RM, Simoons ML, Venge P, Wallentin L, et al. ST2 and mortality in non-ST-segment elevation acute coronary syndrome. *American heart journal*. 2010; 159(5):788–94. Epub 2010/05/04. <https://doi.org/10.1016/j.ahj.2010.02.022> PMID: 20435187.
39. Wang YP, Wang JH, Wang XL, Liu JY, Jiang FY, Huang XL, et al. Roles of ST2, IL-33 and BNP in predicting major adverse cardiovascular events in acute myocardial infarction after percutaneous coronary intervention. *Journal of cellular and molecular medicine*. 2017; 21(11):2677–84. Epub 2017/06/18. <https://doi.org/10.1111/jcmm.13183> PMID: 28623858.
40. Tymiąska A, Kapton-Cieślícka A, Ozierański K, Budnik M, Wancerz A, Sypień P, et al. Association of galectin-3 and soluble ST2 with in-hospital and 1-year outcomes in patients with ST-segment elevation myocardial infarction treated with primary percutaneous coronary intervention. *Pol Arch Intern Med*. 2019; 129(11):770–80. Epub 2019/10/24. <https://doi.org/10.20452/pamw.15030> PMID: 31642446.
41. Gerber Y, Weston SA, Enriquez-Sarano M, Jaffe AS, Manemann SM, Jiang R, et al. Contemporary Risk Stratification After Myocardial Infarction in the Community: Performance of Scores and Incremental Value of Soluble Suppression of Tumorigenicity-2. *Journal of the American Heart Association*. 2017; 6(10). Epub 2017/10/22. <https://doi.org/10.1161/jaha.117.005958> PMID: 29054840.
42. Tse G, Ip C, Luk KS, Gong M, Ting YY, Lakhani I, et al. Prognostic value of soluble ST2 post-aortic valve replacement: a meta-analysis. *Heart Asia*. 2018; 10(1):e010980. Epub 2018/04/11. <https://doi.org/10.1136/heartasia-2017-010980> PMID: 29636828.
43. Luk KS, Ip C, Gong MQ, Wong SH, Wu WK, Dong M, et al. A meta-analysis of soluble suppression of tumorigenicity 2 (sST2) and clinical outcomes in pulmonary hypertension. *Journal of geriatric cardiology: JGC*. 2017; 14(12):766–71. Epub 2018/03/28. PMID: 29581716.
44. Gu L, Li J. Short-term and long-term prognostic value of circulating soluble suppression of tumorigenicity-2 concentration in acute coronary syndrome: a meta-analysis. *Biosci Rep*. 2019; 39(6). Epub 2019/05/17. <https://doi.org/10.1042/bsr20182441> PMID: 31092701.

45. Weinberg EO, Shimpo M, De Keulenaer GW, MacGillivray C, Tominaga S, Solomon SD, et al. Expression and regulation of ST2, an interleukin-1 receptor family member, in cardiomyocytes and myocardial infarction. *Circulation*. 2002; 106(23):2961–6. Epub 2002/12/04. <https://doi.org/10.1161/01.cir.0000038705.69871.d9> PMID: 12460879.
46. Sabatine MS, Morrow DA, Higgins LJ, MacGillivray C, Guo W, Bode C, et al. Complementary roles for biomarkers of biomechanical strain ST2 and N-terminal prohormone B-type natriuretic peptide in patients with ST-elevation myocardial infarction. *Circulation*. 2008; 117(15):1936–44. Epub 2008/04/02. <https://doi.org/10.1161/CIRCULATIONAHA.107.728022> PMID: 18378613.
47. Sanada S, Hakuno D, Higgins LJ, Schreiter ER, McKenzie AN, Lee RT. IL-33 and ST2 comprise a critical biomechanically induced and cardioprotective signaling system. *The Journal of clinical investigation*. 2007; 117(6):1538–49. Epub 2007/05/12. <https://doi.org/10.1172/JCI30634> PMID: 17492053.
48. Aimo A, Migliorini P, Vergaro G, Franzini M, Passino C, Maisel A, et al. The IL-33/ST2 pathway, inflammation and atherosclerosis: Trigger and target? *International journal of cardiology*. 2018; 267:188–92. Epub 2018/05/26. <https://doi.org/10.1016/j.ijcard.2018.05.056> PMID: 29793758.
49. Kaptoge S, Seshasai SR, Gao P, Freitag DF, Butterworth AS, Borglykke A, et al. Inflammatory cytokines and risk of coronary heart disease: new prospective study and updated meta-analysis. *European heart journal*. 2014; 35(9):578–89. Epub 2013/09/13. <https://doi.org/10.1093/eurheartj/ehz367> PMID: 24026779.
50. Bentzon JF, Otsuka F, Virmani R, Falk E. Mechanisms of plaque formation and rupture. *Circulation research*. 2014; 114(12):1852–66. Epub 2014/06/07. <https://doi.org/10.1161/CIRCRESAHA.114.302721> PMID: 24902970.